

University of Padua

DEGREE COURSE IN ICT FOR INTERNET AND MULTIMEDIA ACADEMIC YEAR 2017/2018



1^{st} Homework

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PROBLEM 1

Given a single realization of 800 i.i.d. samples of the r.p $\{x(k)\}$ defined by:

$$x(k) = e^{j(2\pi f_1 k + \varphi_1)} + 0.8e^{j(2\pi f_2 k + \varphi_2)} + w(k)$$

where $w(k) \sim \mathcal{CN}(0, \sigma_w^2)$, $f_1 = 0.17$, $f_2 = 0.78$, $\sigma_w^2 = 2.0 \varphi_{1,2} \sim \mathcal{U}(0, 2\pi)$ are statistically independent, we want to estimate the \mathcal{PSD} of $\{x(k)\}$ using the following models:

BLACKMAN AND TUKEY CORRELOGRAM

$$\mathcal{P}(F) = T_c \sum_{n=-L}^{L} w(n)\hat{r}_x(n)e^{-j2\pi f n T_c}$$
(1)

where $\hat{r}_x(n)$ is the autocorrelation of x taken from -L to L and w(n) is a Hamming window of length 2L+1. The autocorrelation function is evaluated using the *unbiased* estimate give by:

$$\hat{r}_x(n) = \frac{1}{K-n} \sum_{k=n}^{K-1} x(k) x^*(k-n) \qquad n = 0, 1, ..., K-1$$

where K is the number of samples. In order to reduce the variance of the estimate, we chose $L = \lfloor \frac{K}{5} \rfloor$. The hamming window is defined by:

$$w(k) = \begin{cases} 0.54 + 0.46 \cos\left(2\pi \frac{k - \frac{D-1}{2}}{D-1}\right) & k = 0, 1, ..., D-1\\ 0 & elsewhere \end{cases}$$
 (2)

PERIODOGRAM

$$\mathcal{P}_{PER}(f) = \frac{1}{KT_c} |\tilde{\mathcal{X}}(f)| \tag{3}$$

where $\tilde{\mathcal{X}}(f) = T_c \mathcal{X}(f)$, $\mathcal{X}(f)$ is the Fourier Transform of $\{x(k)\}, k = 0, 1, ..., K - 1$.

WELCH PERIODOGRAM

$$\mathcal{P}_{WE}(F) = \frac{1}{N_s} \sum_{r=0}^{N_s - 1} \mathcal{P}_{PER}^{(s)}(f)$$
(4)

where $\mathcal{P}_{PER}^{(s)}(f)$ is the periodogram estimate given by Eq.3. The model extracts different subsequences of consecutive D samples which eventually overlap. If $x^{(s)}$ is the s-th subsequence, characterized by S samples in common with the previous subsequence $x^{(s-1)}$ and with the following $x^{(s+1)}$, the total number of subsequences is $N_s = \lfloor \frac{K-D}{D-S} - 1 \rfloor$. In our analysis, we chose S=?? and D=??. More in details, we computed:

$$x^{(s)}(k) = w(k)x(k+s(D_S)) k = 0, 1, ..., D-1 s = 0, 1, ..., N_S - 1$$

$$\tilde{\mathcal{X}}^{(s)}(f) = \mathcal{F}[x^{(s)}(k)] = T_c \sum_{k=0}^{D-1} x^{(s)}(k)e^{-j2\pi f k T_c}$$

$$\mathcal{P}_{PER}^{(s)}(f) = \frac{1}{DT_c M_w} \left| \tilde{\mathcal{X}}^{(s)}(f) \right|^2$$

where w(k) is the hamming window diffined by Eq. 2 and $M_w = \frac{1}{D} \sum_{k=0}^{D-1} w^2(k)$ is the normalized energy of the window.

AR(N) MODEL

$$\mathcal{P}_{AR}(f) = \frac{T_c \sigma_w^2}{|\mathcal{A}(f)|^2} \tag{5}$$

The output process is described by the following recursive e equation

$$x(k) = -\sum_{n=1}^{N} a_n x(k-n) + w(k)$$
(6)

where w is white noise with variance σ_w^2 . The transfer function of this system is given by $H_{AR} = \frac{1}{A(z)}$ where $A(z) = 1 + \sum_{n=1}^{N} a_n z^{-n}$. The ACS of x in z-domain is easily evaluated using

$$P_x(z) = \frac{\sigma_w^2}{A(z)A^*(\frac{1}{z^*})}$$

from which Eq. 5 is computed letting $\mathcal{A}(f) = A(e^{j2\pi fT_c})$.

Coefficients $a_1, a_2, ..., a_N$ are computed exploiting the Yule-Walker equation $\mathbf{a} = -\mathbf{R}^{-1}\mathbf{r}$, where r = [r(1), r(2), ..., r(N)] and the autocorrelation matrix is defined as

$$\mathbf{R} = \begin{pmatrix} r(0) & r(-1) & \dots & r(N+1) \\ r(1) & r(0) & \dots & r(-N+2) \\ \vdots & \vdots & \ddots & \vdots \\ r(N-1) & r(N-2) & \dots & r(0) \end{pmatrix}$$

The noise variance deriving from the model is

$$\sigma_w^2 = r_x(0) + \mathbf{r}^H \mathbf{a} \tag{7}$$

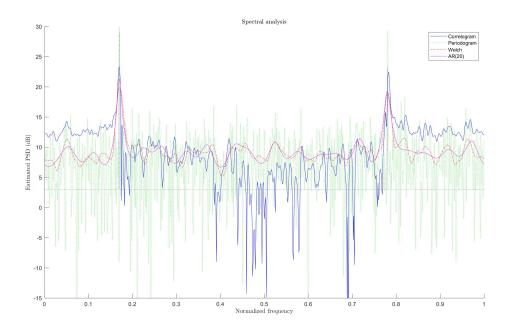


Figure 1: PSD estimates.

CONCLUSIONS

A comparison between different PSD estimates of $\{x(k)\}$ is now discussed. The difficulties in this problem were related to the fact that the noise variance is double with respect to the maximum amplitude of the usefull signal. For this reason, a good analysis can be achieved only by considering a generated sequence $\{x(k)\}$ which is not too corrupted at the carriers frequencies f_1 and f_2 . The four different PSD estimates are shown in Figure [1].

The parameters of the different models are as follow. The order of the autocorrelation estimator for the Correlogram is $N_{corr} = \lfloor K/5 \rfloor = 160$, a the window used is a Hamming window definde in Eq. 2. For the Periodogram the window used is the rectangular one, while for the Welch Periodogram we set D = ?? and S = ?? still using an Hamming window. The order evaluation for the AR model was carried out by analysing the behaviour of the noise variance σ_w^2 with varying N defined in eq. 7 and graphically shown in Fig. 2.

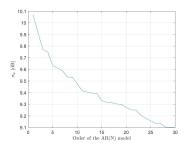


Figure 2: Variance of the AR model white noise as a function of the order N.