## SDS 385: Homework 2

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## Problem 1. SGD for logistic regression

(A) In part (A) of the last homework we proved that the gradient of the negative log-likelihood can be expressed as

$$\nabla l(\beta) = -\sum_{i=1}^{n} \{x_i(y_i - m_i w_i)\} = \sum_{i=1}^{n} g_i(\beta)$$

where

$$g_i(\beta) = x_i(y_i - m_i w_i) = x_i(y_i - \hat{y}_i)$$

and

$$\hat{y}_i = \mathbb{E}(y_i|\beta) = m_i w_i(\beta) = m_i \frac{1}{1 + \exp(-x_i^T \beta)}.$$

A nice interpretation can be given to this latter expression: the gradient is large when the data  $y_i$ 's differ from their maximum likelihood estimates, i.e. the probabilities  $w_i$ 's.

(B) Suppose that we draw a single data point at random from the sample, obtaining the pair  $\{y_i, x_i\}$ . We show that the random vector  $ng_i(\beta)$  is an unbiased estimate of  $\nabla l(\beta)$ , that is,  $E[ng_i(\beta)] = \nabla l(\beta)$ , where the expectation is under random sampling from the set of all  $\{y_i, x_i\}$  pairs. In fact,

$$E[ng_{i}(\beta)] = E[n(x_{i}y_{i} - x_{i}m_{i}\frac{1}{1 + e^{-x_{i}^{T}\beta}})]$$

$$= \sum_{i=1}^{n} \left\{ n\left(x_{i}y_{i} - x_{i}m_{i}\frac{1}{1 + e^{-x_{i}^{T}\beta}}\right) \frac{1}{n}\right\}$$

$$= \sum_{i=1}^{n} \left\{ x_{i}y_{i} - x_{i}m_{i}\frac{1}{1 + e^{-x_{i}^{T}\beta}}\right\}$$

$$= \sum_{i=1}^{n} g_{i}(\beta) = \nabla l(\beta).$$

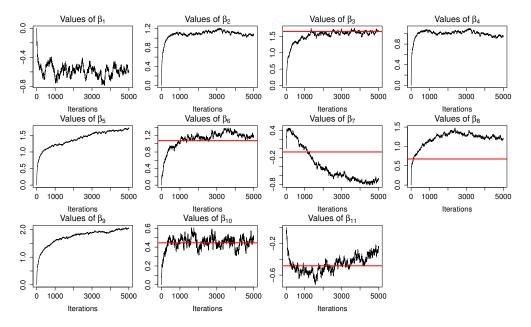
(C) The SGD exploits this fact in order to use an update step which is faster to compute. In fact, instead of using the gradient  $\nabla l(\beta)$  calculated from all n data points to choose the step direction, we use the gradient  $g_i(\beta)$  calculated from a single data point, sampled randomly from the whole data set. In this version of the algorithm, sampling without replacement has been performed.

In order to assess the validity of the implementation, we run the algorithm on the real data for a reasonable number of iterations N=100000. We tried different values for the step size  $\gamma$ , which was kept constant over the iterations. We report in Figures 1 - 3 the traceplots of all of the  $\beta$  parameters (i.e. the values of the parameters over the iterations). In order to facilitate the visualization, we thinned the samples by a factor 20 (we displayed one value of  $\beta$  every 20 iterations).

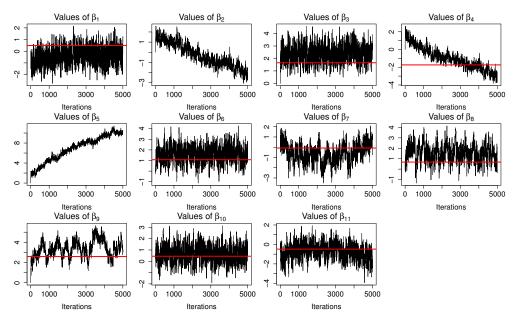
We can see that for small values of the step size  $\gamma$ , convergence is not reached for many of the components of  $\beta$ . When we increase the step size, on the other hand, convergence

is reached more rapidly because the algorithm explores well the parameters space. However, the asymptotic variance of the parameters is higher, since at every iteration the weight attributed to the direction given by any new sampled data is high.

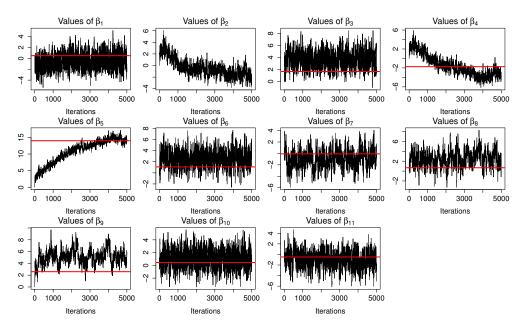
In this first implementation of the algorithm we do not worry about the choice of the tuning parameter, that is instead crucial for SGD. As a convergence diagnostics, we plot the iterates of the running average of  $l_t(\beta)$ , the individual log-likelihood contribution from the data point sampled at step t (see Figure 4).



*Figure 1: Traceplots of the*  $\beta$  *parameters when*  $\gamma = 0.01$ .



*Figure 2: Traceplots of the*  $\beta$  *parameters when*  $\gamma = 0.5$ .



*Figure 3: Traceplots of the*  $\beta$  *parameters when*  $\gamma = 1$ .

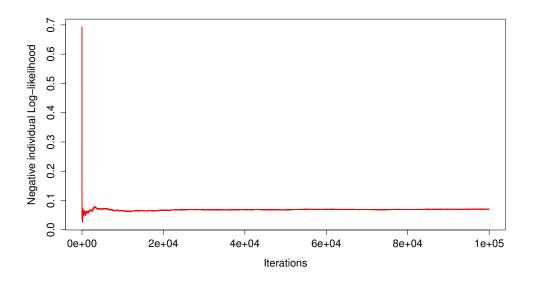


Figure 4: Running average of the single log-likelihood contributions.

(D) Now we try using a decaying step size. Specifically, we use the Robbins–Monro rule for step sizes:

$$\gamma(t) = C(t + t_0) - \alpha,$$

where C > 0,  $\alpha \in [0.5, 1]$ , and  $t_0$  (the "prior number of steps") are constants. The exponent  $\alpha$  is usually called the learning rate.

Ideally, we want to obtain large values of  $\gamma$  at the beginning of the algorithms, so that the "true" values of the parameters are rapidly reached. Afterwards, we want the step sizes to be reduced in order to diminish the variance of the estimates and to, eventually, converge.

After several trials, we set the values  $\alpha=0.8, C=4000, t_0=1000$ , yielding the results displayed in Figure 5.

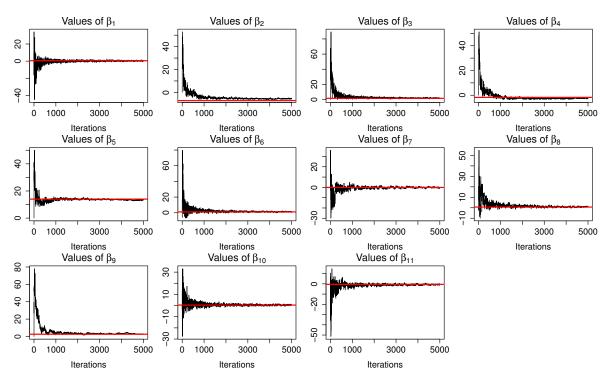


Figure 5: Running average of the single log-likelihood contributions.

(E) At last, in order to estimate the parameters, we report the time-averaged iterates instead of the iterates. The comparison between all the algorithms is found in Listing 1.

```
1
                        Beta0 Beta1 Beta2 Beta3 Beta4 Beta5 Beta6
 2
    Glm
                      0.4870168 - 7.221851 \ 1.654756 - 1.737630 \ 14.00485 \ 1.074953 - 0.07723455
 3
    Gradient
                      0.4830816 - 7.020847 \ 1.654900 - 1.922239 \ 13.97312 \ 1.073116 - 0.06648360
 4
                      0.4870168 -7.221851 1.654756 -1.737630 14.00485 1.074953 -0.07723455
    Newton
 5
                      3.5570133 -3.245963 2.779636 -2.102697 17.71371 1.398131 -0.39040924
    SGD
 6
    SGD Robbins-Monro 0.5527124 -5.426020 2.036747 -2.446775 13.33293 1.152240 0.06014270
    SGD Polyak-Ruppert 0.2865011 -5.538449 1.980589 -2.581991 13.60858 1.281796 -0.05755391
 7
 8
                        Beta7
                                 Beta8
                                           Beta9
                                                    Beta10
 9
    Glm
                     0.6751231 2.592874 0.4462563 -0.4824842
10 Gradient
                      0.6767472 2.595408 0.4459843 -0.4832114
11 Newton
                      0.6751231 2.592874 0.4462563 -0.4824842
12 sgD
                      1.2890052 5.722428 1.7388858 0.4833998
13 SGD Robbins-Monro 1.0255726 2.835010 0.5391915 -0.4159569
14 SGD Polyak-Ruppert 0.9174220 3.249271 0.5015554 -0.6899736
```

*Listing 1: Comparison of the estimates using five different algorithms* 

## Appendix A

## R code

```
1 # Function for the SGD of the logit model
    SGD <- function(y, X, mi, beta0, maxiter, alpha, tol) {
 3
 4
      N \leftarrow dim(X)[1]
 5
      # Sample the data points to calculate the gradient
      idx <- sample(1:N, maxiter, replace = TRUE)</pre>
 8
      betas <- array(NA, dim=c(maxiter, length(beta0)))</pre>
 9
      betas[1,] <- beta0 # Initial guess</pre>
10
11
      av_ll <- array(NA, dim = maxiter)</pre>
12
       av_ll[1] \leftarrow log_lik(betas[1,], matrix(y[idx[1]], nrow=1), matrix(X[idx[1],], nrow=1), mi[idx[1]])
13
14
      for (iter in 2:maxiter) {
15
       gradient <- grad_loglik(betas[iter-1,], matrix(y[idx[iter]],nrow=1), matrix(X[idx[iter],],</pre>
             nrow=1), mi[idx[iter]])
16
         betas[iter,] <- betas[iter-1,] - alpha*gradient</pre>
17
         11 <- log_lik(betas[iter,], matrix(y[idx[iter]],nrow=1), matrix(X[idx[iter],],nrow=1), mi[idx</pre>
18
         av_ll[iter] <- ((av_ll[iter-1]) * (iter-1) + ll) / iter</pre>
19
20
       return(list("ll" = av_ll, "beta" = betas))
```

*Listing A.1: Function implementing the SGD for the logit model.* 

```
# Function for the SGD of the logit model
 2
    SGD_Robbins <- function(y, X, mi, beta0, maxiter, C, t0, alpha, tol){
 3
 4
      N \leftarrow dim(X)[1]
 5
      # Sample the data points to calculate the gradient
 6
      idx <- sample(1:N, maxiter, replace = TRUE)</pre>
 7
 8
      betas <- array(NA, dim=c(maxiter, length(beta0)))</pre>
 9
      betas[1,] <- beta0 # Initial guess</pre>
10
11
      av_ll <- array(NA, dim = maxiter)</pre>
12
      av_ll[1] <- log_lik(betas[1,], matrix(y[idx[1]],nrow=1), matrix(X[idx[1],],nrow=1), mi[idx[1]])</pre>
13
14
    for (iter in 2:maxiter) {
```

```
15
         step_size <- C*(iter+t0)^(-alpha)</pre>
16
17
         gradient <- grad_loglik(betas[iter-1,], matrix(y[idx[iter]],nrow=1), matrix(X[idx[iter],],</pre>
           nrow=1), mi[idx[iter]])
18
         betas[iter,] <- betas[iter-1,] - step_size*gradient</pre>
19
         11 <- log_lik(betas[iter,], matrix(y[idx[iter]],nrow=1), matrix(X[idx[iter],],nrow=1), mi[idx</pre>
             [iter]])
20
         av_ll[iter] <- ((av_ll[iter-1]) * (iter-1) + ll) / iter</pre>
21
22
       return(list("ll" = av_ll, "beta" = betas))
23
```

Listing A.2: Function implementing the SGD for the logit model using the Robbins-Monro rule.