# souvenir-sales-time-series-analysis

May 4, 2023

# 1 Souvenir Sales - Time Series Analysis

The following is a statistical analysis on a monthly time series which collects data about the sales of a souvenir shop in Australia in the period between 1987 and 1992.

The analysis will roughly follow the Box-Jenkins method and will focus on reaching stationarity for the time series, estimating a SARIMA model and predicting future values.

### 1.1 Data exploration

```
import numpy as np
import pandas as pd
import matplotlib.pyplot as plt
import seaborn as sns
from statsmodels.graphics.tsaplots import plot_acf, plot_pacf
from statsmodels.tsa.stattools import adfuller, kpss
from scipy.stats import boxcox
from scipy.special import inv_boxcox
import pmdarima as pm
import warnings
warnings.filterwarnings("ignore")
```

```
[2]: tseries = pd.read_csv("data/monthly_sales_queensland.csv", header = 0,⊔

⇒parse_dates = ["date"], index_col = 0)
```

Now that we have imported the time series, let's have a first look at its values.

```
[3]: tseries
```

```
[3]:
                    sales
     date
                  1664.81
     1987-01-01
     1987-02-01
                  2397.53
     1987-03-01
                  2840.71
     1987-04-01
                  3547.29
     1987-05-01
                  3752.96
     1992-08-01 19888.61
     1992-09-01
                 23933.38
```

```
1992-10-01 25391.35

1992-11-01 36024.80

1992-12-01 80721.71

[72 rows x 1 columns]

[4]: tseries.index.min(), tseries.index.max()

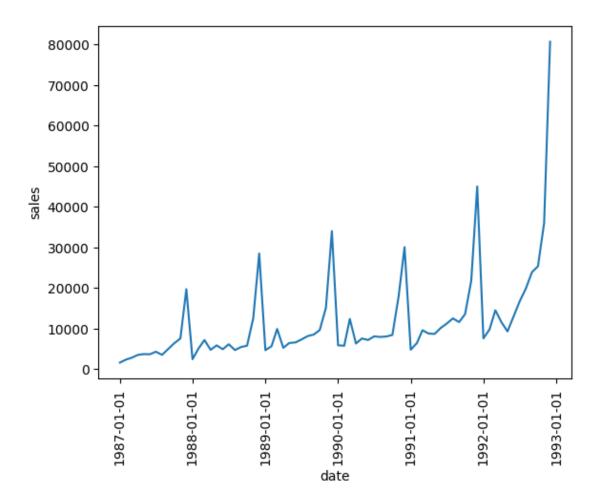
[4]: (Timestamp('1987-01-01 00:00:00'), Timestamp('1992-12-01 00:00:00'))

[5]: tseries.index.max() - tseries.index.min()
```

### 1.2 Check for non-stationarity

Let's make a plot and see what we can say about it.

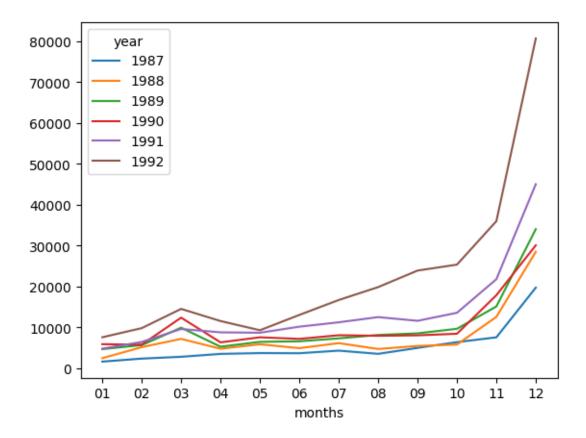
```
[6]: fig, ax = plt.subplots()
   ax.plot(tseries["sales"])
   ax.set_xlabel("date")
   ax.set_ylabel("sales")
   ax.set_xticks(ax.get_xticks()[1::1])
   plt.xticks(rotation = 90)
   plt.show()
```



From the plot of the series, we can already grasp that it is not stationary, as the expected value is not constant over time and neither is variance, which tends to increase. In particular, we can speculate the presence of an upward trend and a seasonal effect, which is comprehensible, considering the turistical vocation of the shop.

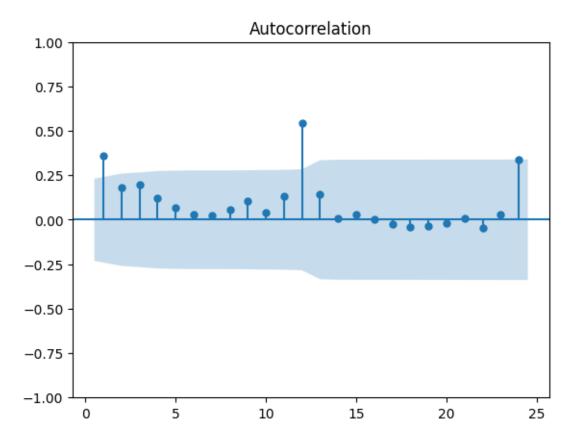
```
[7]: df_years = tseries.copy(deep = True)
    df_years.reset_index(inplace = True)
    df_years["year"] = pd.to_datetime(df_years["date"]).dt.year
    df_years["date"] = pd.to_datetime(df_years["date"]).dt.strftime("%m")
    unstacked = df_years.set_index(["year", "date"])["sales"].unstack(-2)
    unstacked.plot(xlabel = "months", xticks = pd.Series(range(0,12)))
```

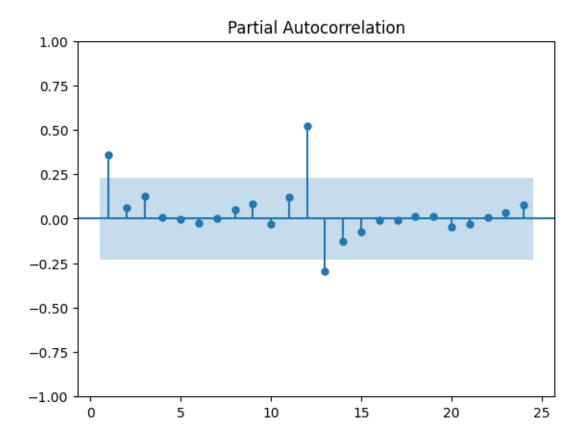
[7]: <Axes: xlabel='months'>



To better appreciate it, this plot shows the data for each year separately. The values of the series are clearly higher for the latest years and there is a recurring peak in the months of March and August, followed by a valley in October.

```
[8]: plot_acf(tseries["sales"], lags = 24, zero = False);
```





Lastly, these are the *global* and *partial autocorrelation functions* for the series. The slow decay for the ACF suggests, once again, the existence of a trend, while the spikes at lag 12 indicate a probable seasonality.

To formalize our guesses, let's resort to two statistical test: - The **Augmented Dickey-Fuller test** tests the null hypothesis of the presence of a unit root in our time series - The **KPSS test** tests the null hypothesis that our data is stationary

```
[11]: (0.9887885592096786,
0.01,
3,
{'10%': 0.347, '5%': 0.463, '2.5%': 0.574, '1%': 0.739})
```

We were expecting to reject H0 for **KPSS** and not be able to reject it for **ADF** and that's exactly what happened, looking at the p-values.

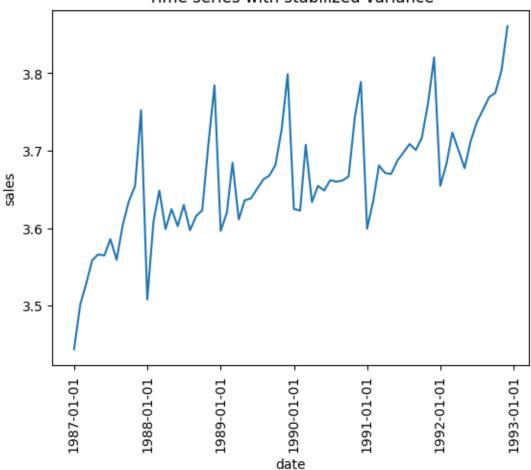
#### 1.3 Reach stationarity

In order to obtain stationarity in our time series, we need to perform a series of operations: we are going to stabilize the variance through the *Box-Cox transformation* and then apply differencing to treat trend and seasonality.

```
[12]: tseries_novar = tseries.copy(deep = True)
bc = boxcox(tseries["sales"])[0]
lmbda = boxcox(tseries["sales"])[1]
tseries_novar["sales"] = bc
```

```
fig, ax = plt.subplots()
ax.plot(tseries_novar["sales"])
ax.set_title("Time series with stabilized variance")
ax.set_xlabel("date")
ax.set_ylabel("sales")
ax.set_xticks(ax.get_xticks()[1::1])
plt.xticks(rotation = 90)
plt.show()
```





```
[14]: tseries_diff = tseries_novar.copy(deep = True)
    tseries_diff["sales"] = tseries_diff["sales"].diff(periods = 1)
    tseries_diff = tseries_diff.iloc[1:]
    tseries_diff
```

```
[14]: sales

date

1987-02-01 0.057891

1987-03-01 0.025229

1987-04-01 0.031510

1987-05-01 0.007729

1987-06-01 -0.001396

...

1992-08-01 0.016046

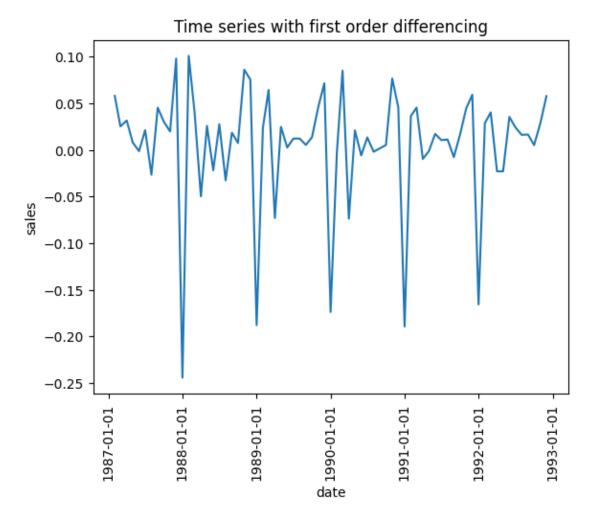
1992-09-01 0.016463

1992-10-01 0.005105
```

```
1992-11-01 0.028748
1992-12-01 0.057713
```

[71 rows x 1 columns]

```
fig, ax = plt.subplots()
ax.plot(tseries_diff["sales"])
ax.set_title("Time series with first order differencing")
ax.set_xlabel("date")
ax.set_ylabel("sales")
ax.set_xticks(ax.get_xticks()[1::1])
plt.xticks(rotation = 90)
plt.show()
```

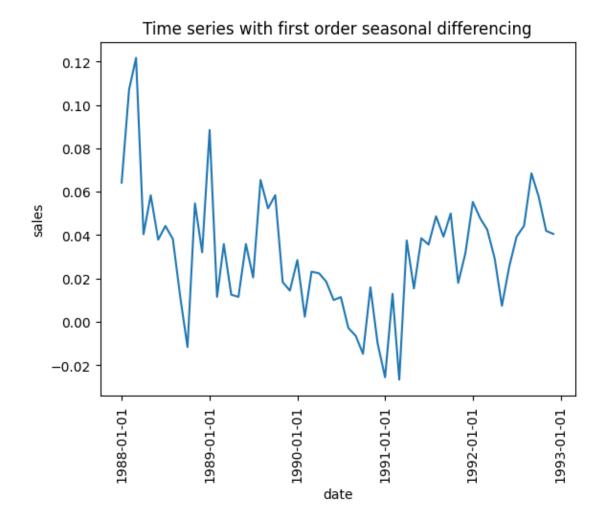


The first order differencing removes trend, but leaves seasonality.

```
[16]: tseries_diff_s = tseries_novar.copy(deep = True)
     tseries_diff_s["sales"] = tseries_diff_s["sales"].diff(periods = 12)
     tseries_diff_s = tseries_diff_s.iloc[12:]
     tseries_diff_s
[16]:
                    sales
     date
     1988-01-01 0.064202
     1988-02-01 0.107133
     1988-03-01 0.121736
     1988-04-01 0.040426
     1988-05-01 0.058380
     1988-06-01 0.037900
     1988-07-01 0.044232
     1988-08-01 0.038239
     1988-09-01 0.011337
     1988-10-01 -0.011622
     1988-11-01 0.054617
     1988-12-01 0.032061
     1989-01-01 0.088451
     1989-02-01 0.011526
     1989-03-01 0.035867
     1989-04-01 0.012516
     1989-05-01 0.011527
     1989-06-01 0.035912
     1989-07-01 0.020505
     1989-08-01 0.065412
     1989-09-01 0.052333
     1989-10-01 0.058421
     1989-11-01 0.018319
     1989-12-01 0.014419
     1990-01-01 0.028504
     1990-02-01 0.002387
     1990-03-01 0.023147
     1990-04-01 0.022415
     1990-05-01 0.018570
     1990-06-01 0.010074
     1990-07-01 0.011416
     1990-08-01 -0.002764
     1990-09-01 -0.006471
     1990-10-01 -0.014722
     1990-11-01 0.015944
     1990-12-01 -0.009979
     1991-01-01 -0.025552
     1991-02-01 0.012920
     1991-03-01 -0.026660
```

1991-04-01 0.037513

```
1991-05-01 0.015415
     1991-06-01 0.038552
     1991-07-01 0.035647
     1991-08-01 0.048681
     1991-09-01 0.039313
     1991-10-01 0.049992
     1991-11-01 0.017970
     1991-12-01 0.031579
     1992-01-01 0.055320
     1992-02-01 0.047692
     1992-03-01 0.042533
     1992-04-01 0.029227
     1992-05-01 0.007447
     1992-06-01 0.025719
     1992-07-01 0.039269
     1992-08-01 0.044277
     1992-09-01 0.068504
     1992-10-01 0.057723
     1992-11-01 0.041973
     1992-12-01 0.040557
[17]: fig, ax = plt.subplots()
     ax.plot(tseries_diff_s["sales"])
     ax.set_title("Time series with first order seasonal differencing")
     ax.set_xlabel("date")
     ax.set_ylabel("sales")
     ax.set_xticks(ax.get_xticks()[1::1])
     plt.xticks(rotation = 90)
     plt.show()
```



The seasonal differencing removes seasonality, but leaves trend.

```
[18]: tseries_diff_final = tseries_novar.copy(deep = True)
    tseries_diff_final["sales"] = tseries_diff_final["sales"].diff(periods = 1).
    diff(periods = 12)
    tseries_diff_final = tseries_diff_final.iloc[13:]
    tseries_diff_final
```

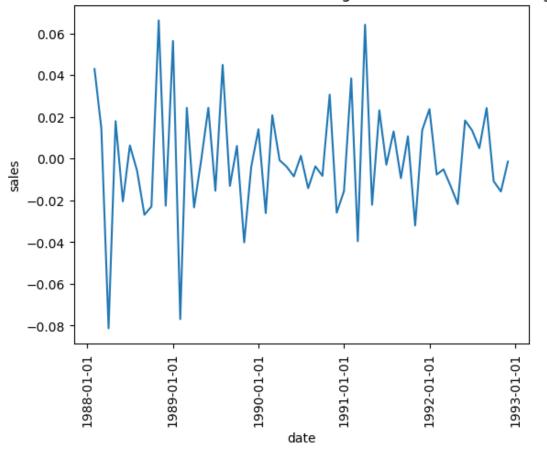
```
[18]: sales

date

1988-02-01 0.042931
1988-03-01 0.014603
1988-04-01 -0.081310
1988-05-01 0.017954
1988-06-01 -0.020480
1988-07-01 0.006332
1988-08-01 -0.005994
```

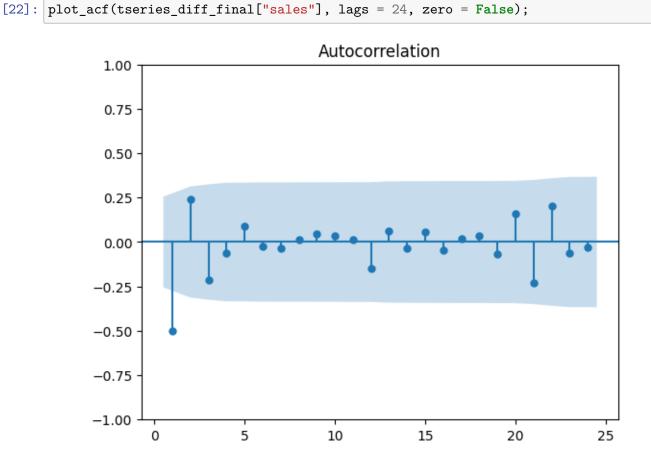
- 1988-09-01 -0.026901
- 1988-10-01 -0.022960
- 1988-11-01 0.066239
- 1988-12-01 -0.022556
- 1989-01-01 0.056390
- 1989-02-01 -0.076924
- 1989-03-01 0.024341
- 1989-04-01 -0.023351
- 1989-05-01 -0.000989
- 1989-06-01 0.024385
- 1989-07-01 -0.015407
- 1989-08-01 0.044907
- 1989-09-01 -0.013079
- 1989-10-01 0.006088
- 1989-11-01 -0.040102
- 1989-12-01 -0.003900
- 1990-01-01 0.014085
- 1990-02-01 -0.026117 1990-03-01 0.020760
- 1990-04-01 -0.000732
- 1990-05-01 -0.003845
- 1330 00 01 0.000040
- 1990-06-01 -0.008495
- 1990-07-01 0.001341
- 1990-08-01 -0.014180
- 1990-09-01 -0.003707
- 1990-10-01 -0.008252
- 1990-11-01 0.030666
- 1990-12-01 -0.025923
- 1991-01-01 -0.015573 1991-02-01 0.038472
- 1991-03-01 -0.039579
- 1991-04-01 0.064173
- 1991-05-01 -0.022099
- 1991-06-01 0.023137
- 1991-07-01 -0.002905
- 1991-08-01 0.013035
- 1001 00 01 0 00000
- 1991-09-01 -0.009368
- 1991-10-01 0.010679
- 1991-11-01 -0.032023
- 1991-12-01 0.013609
- 1992-01-01 0.023742
- 1992-02-01 -0.007628
- 1992-03-01 -0.005159
- 1992-04-01 -0.013307
- 1992-05-01 -0.021780
- 1992-06-01 0.018272
- 1992-07-01 0.013551

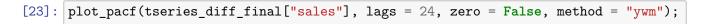
# Time series with first order differencing and seasonal differencing

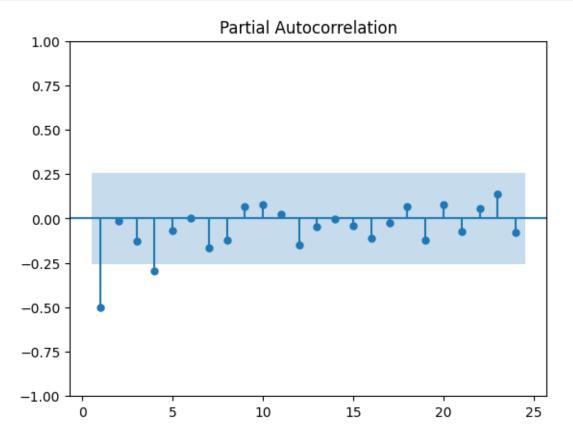


The new series should be stationary now. Let's check with our two tests.

```
[20]: adfuller(tseries_diff_final["sales"])
[20]: (-5.903726178997968,
       2.73961678960973e-07,
       3,
       55,
       {'1%': -3.5552728880540942,
        '5%': -2.9157312396694217,
        '10%': -2.5956695041322315},
       -234.49533536833405)
[21]: kpss(tseries_diff_final["sales"])
[21]: (0.02613487681632163,
       0.1,
       0,
       \{'10\%': 0.347, '5\%': 0.463, '2.5\%': 0.574, '1\%': 0.739\}
     Our conclusions on the null hypothesis have now switched, as we expected.
     Let's look at the ACF and PACF.
```







At this point, we should be able to estimate the values for the parameters of our ARIMA model by looking at these two plots and the spikes in them. However, the most reliable way to actually determine the parameters is using an objective procedure, for example a stepwise-like, and let a computer do it for us by choosing among many ARIMA models the "best" one, in terms of optimizing a certain indicator.

Performing stepwise search to minimize aic

ARIMA(1,1,1)(0,1,1)[12] : AIC=-269.383, Time=1.02 sec ARIMA(0,1,0)(0,1,0)[12] : AIC=-251.964, Time=0.04 sec

```
ARIMA(1,1,0)(1,1,0)[12]
                                    : AIC=-269.841, Time=0.68 sec
ARIMA(0,1,1)(0,1,1)[12]
                                    : AIC=-269.740, Time=0.42 sec
                                    : AIC=-267.544, Time=0.07 sec
ARIMA(1,1,0)(0,1,0)[12]
ARIMA(1,1,0)(2,1,0)[12]
                                    : AIC=-268.583, Time=0.43 sec
ARIMA(1,1,0)(1,1,1)[12]
                                    : AIC=-269.956, Time=0.56 sec
                                    : AIC=-271.244, Time=0.22 sec
ARIMA(1,1,0)(0,1,1)[12]
ARIMA(1,1,0)(0,1,2)[12]
                                    : AIC=-269.821, Time=0.67 sec
                                    : AIC=-267.700, Time=1.65 sec
ARIMA(1,1,0)(1,1,2)[12]
ARIMA(0,1,0)(0,1,1)[12]
                                    : AIC=-252.302, Time=0.38 sec
                                    : AIC=-269.369, Time=0.81 sec
ARIMA(2,1,0)(0,1,1)[12]
ARIMA(2,1,1)(0,1,1)[12]
                                    : AIC=-267.667, Time=1.18 sec
ARIMA(1,1,0)(0,1,1)[12] intercept
                                  : AIC=-269.622, Time=0.45 sec
```

Best model: ARIMA(1,1,0)(0,1,1)[12]

Total fit time: 8.612 seconds

[24]: <class 'statsmodels.iolib.summary.Summary'>

11 11 11

#### SARIMAX Results

========

Dep. Variable: y No. Observations:

72

Model: SARIMAX(1, 1, 0)x(0, 1, [1], 12) Log Likelihood

138.622

Date: Thu, 04 May 2023 AIC

-271.244

Time: 21:22:09 BIC

-265.011

Sample: 01-01-1987 HQIC

-268.811

- 12-01-1992

Covariance Type:

opg

	coef	std err	z	P> z	[0.025	0.975]
ar.L1 ma.S.L12 sigma2	-0.5521 -0.4516 0.0005	0.124 0.181 0.000	-4.442 -2.493 4.347	0.000 0.013 0.000	-0.796 -0.807 0.000	-0.308 -0.097 0.001
						0.001

===

Ljung-Box (L1) (Q): 0.00 Jarque-Bera (JB):

0.02

Prob(Q): 0.99 Prob(JB):

0.99

Heteroskedasticity (H): 0.30 Skew:

-0.04

```
Prob(H) (two-sided): 0.01 Kurtosis: 2.95
```

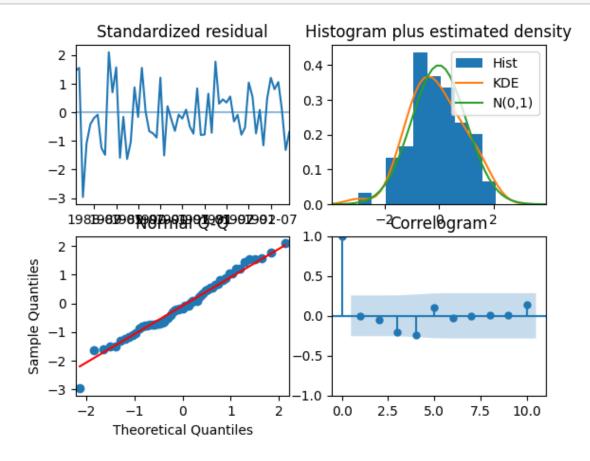
### Warnings:

[1] Covariance matrix calculated using the outer product of gradients (complex-step).  $\footnote{1.5mm}$ 

The chosen model appears to be ARIMA(1,1,0)(0,1,1)[12]. Along with it, we have also got a number of diagnostic tools: we can see that the parameters are significantly different than 0, while none of the residuals is significant.

Let's look at other diagnostic measures through some plots.

### [25]: fit.plot\_diagnostics();



The residuals roughly follow a normal distribution, as deduced from the histogram and the Q-Q plot, but they seem to follow a pattern in their time series, which is not good for our model.

For the last step, let's try to forecast some future values, in particular 12 more observations, and

plot the result.

```
[26]: | forecast = fit.predict(n_periods = 12)
      df_forecast = pd.DataFrame({"sales": forecast.values})
      df_forecast["sales"] = inv_boxcox(df_forecast["sales"], lmbda)
      df_forecast.set_index(forecast.index, inplace = True)
      df_forecast.index.name = "date"
      tseries_forecast = pd.concat([tseries, df_forecast])
      tseries_forecast
[26]:
                          sales
      date
      1987-01-01
                    1664.810000
                    2397.530000
      1987-02-01
                    2840.710000
      1987-03-01
      1987-04-01
                    3547.290000
      1987-05-01
                    3752.960000
      1993-08-01
                   26288.574310
      1993-09-01
                   29161.143853
      1993-10-01
                   32410.080979
      1993-11-01
                   55069.265169
      1993-12-01 137764.601875
      [84 rows x 1 columns]
[27]: fig, ax = plt.subplots()
      ax.plot(tseries_forecast["sales"])
      ax.set_xlabel("date")
      ax.set_ylabel("sales")
      ax.set_xticks(ax.get_xticks()[1::1])
      plt.xticks(rotation = 90)
      plt.show()
```

