

Role extraction for digraphs via neighbourhood pattern similarity

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Foundations of Computational Mathematics

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A motivating example



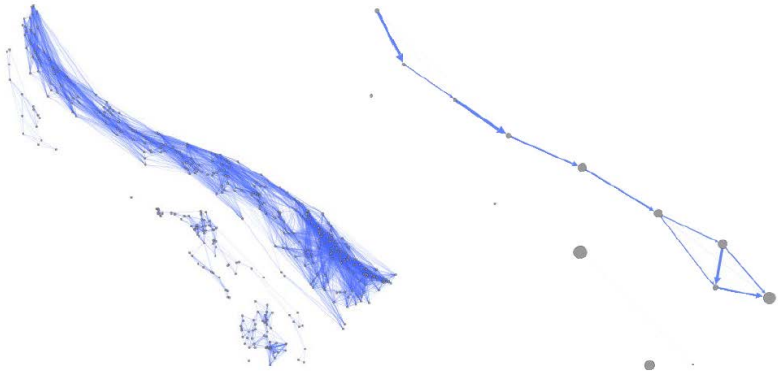
Once a year the corals in the Great Barrier Reef reproduce.

Corals release their eggs and sperm into the water at the same time.

Clouds of coral eggs and sperm float in all directions, carried by the currents, winds, and waves.

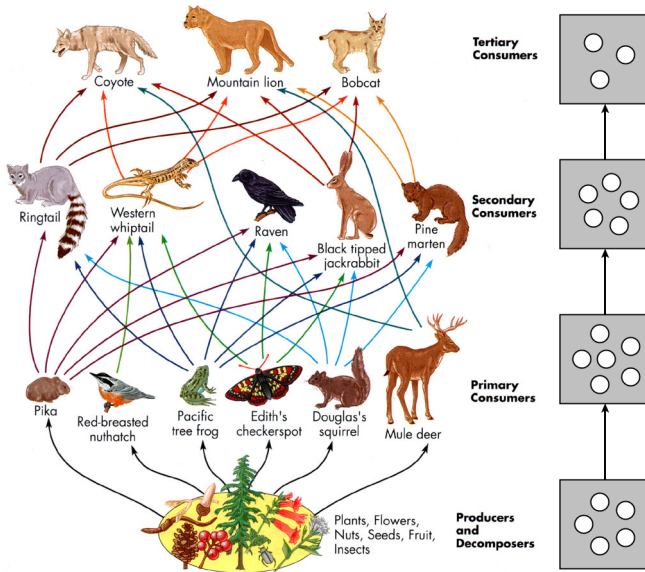
When egg and sperm meet, the resulting larvae continue to drift to find the perfect spot to settle.

A motivating example



Understanding the underlying directed structure lets us control and predict the growth of the reef

A motivating example



Roles of Directed Graph

Given a graph with adjacency matrix A we want to find an assignment function, that partitions the graph into **Roles**

$$A = \begin{array}{|c|} \hline \text{[Dense Matrix]} \\ \hline \end{array} \xRightarrow{P?} PAP^T = \begin{array}{|c|} \hline \begin{array}{|c|} \hline \text{[Dense Matrix]} \\ \hline \end{array} \\ \hline \end{array}$$

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$$A = \begin{array}{c} \text{[Adjacency Matrix Visualization]} \end{array} \xRightarrow{P?} PAP^T = \begin{array}{c} \text{[Partitioned Adjacency Matrix Visualization]} \end{array}$$

We also want a matrix B telling us how the roles are connected

$$B = \begin{pmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{pmatrix}$$

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$$A = \begin{array}{|c|} \hline \text{[Noisy Matrix]} \\ \hline \end{array} \xRightarrow{P?} PAP^T = \begin{array}{|c|} \hline \begin{array}{|c|} \hline \text{[Noisy Matrix]} \\ \hline \end{array} \\ \hline \end{array}$$

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In the ideal case, all nodes in the same role are **structurally equivalent**:

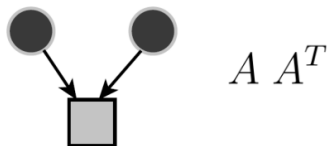
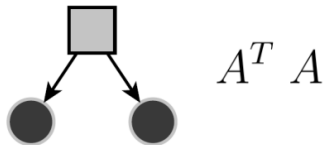
- their children are the same
- their parents are the same

Neighbourhood Pattern Similarity Measure

The **Neighbourhood Pattern Similarity Measure** between two nodes (i, j) takes into consideration all their common 'ancestors', 'descendants'... and 'relatives'

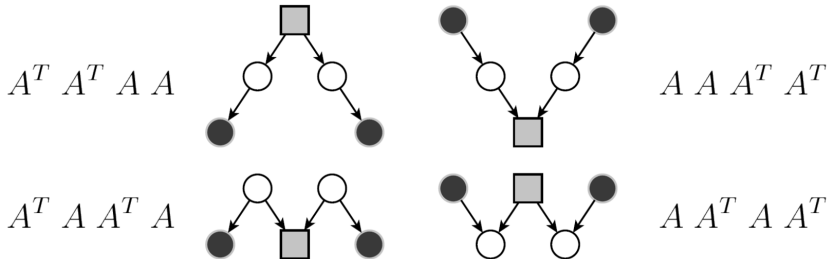
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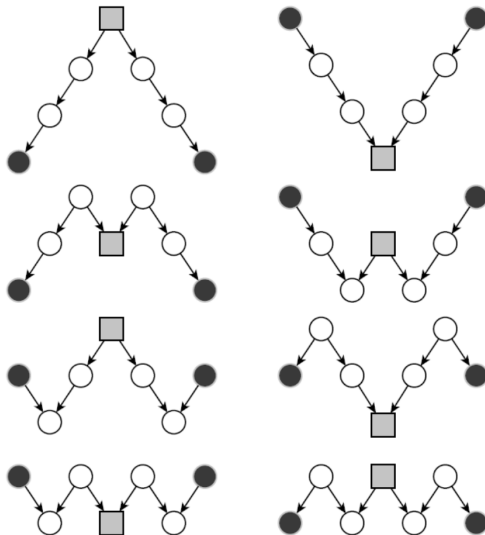
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$$L_1 = AA^T + A^T A = \begin{bmatrix} A & A^T \end{bmatrix} \begin{bmatrix} A^T \\ A \end{bmatrix}$$

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$$\Gamma_A^{k+1}[I] = \Gamma_A[L_k] := L_{k+1} = \begin{bmatrix} A & A^T \end{bmatrix} \begin{bmatrix} L_k & \\ & L_k \end{bmatrix} \begin{bmatrix} A^T \\ A \end{bmatrix}$$

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$$S_{k+1} := \Gamma_A[I] + \beta^2 \Gamma_A^2[I] + \dots + \beta^{2(k+1)} \Gamma_A^{k+1}[I] = \Gamma_A[I + \beta^2 S_k]$$

$$S_{k+1} - S_k = \beta^{2(k+1)} \Gamma_A^{k+1}[I] \succeq 0 \implies S_{k+1} \succeq S_k$$

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- S_k are always PSD matrices and $S_{k+1} \succeq S_k$
- If $\beta^2 < \frac{1}{4\|A\|^2}$, then $S_k \rightarrow S^*$ with

$$\text{vec}(S^*) = \left[I - \beta^2 (A \otimes A + (A \otimes A)^T) \right]^{-1} \text{vec}(S_1)$$

- In the ideal case, if $[B \ B^T]$ has maximum rank, then the rank of each S_k is the number of roles and a spectral method on S_k let us recover the roles

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Warning: The number of roles may not be linked to the rank of A or B

$$B = \begin{bmatrix} 0 & 1 & 0 \\ 1 & 0 & 1 \\ 1 & 0 & 1 \end{bmatrix} \implies \text{rk}(A) = \text{rk}(B) = 2$$

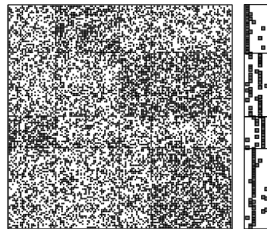
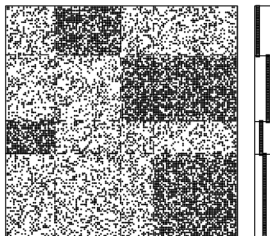
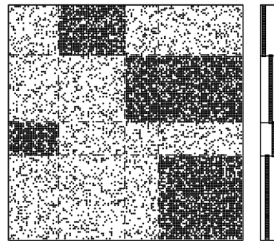
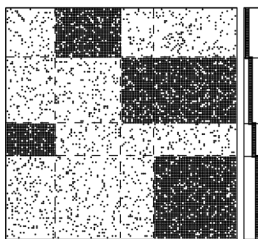
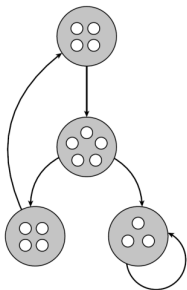
$$[B \ B^T] = \begin{bmatrix} 0 & 1 & 0 & 0 & 1 & 1 \\ 1 & 0 & 1 & 1 & 0 & 0 \\ 1 & 0 & 1 & 0 & 1 & 1 \end{bmatrix} \implies \text{rk}(S_k) = \text{rk}([B \ B^T]) = 3$$

Erdős-Renyi Random Graphs

Suppose now that G is a random graph induced by B where the probability of an edge between nodes from role i to role j is p if $B_{i,j} = 1$ and $1 - p$ if $B_{i,j} = 0$

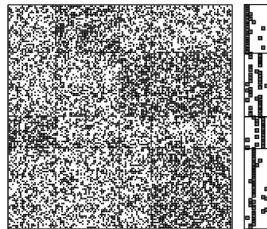
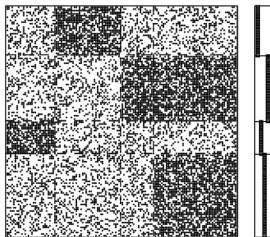
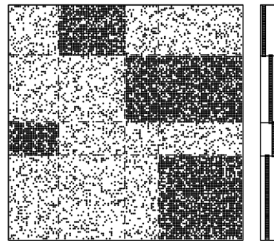
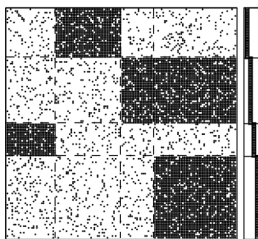
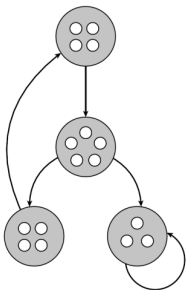
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NPS Measure can empirically identify the roles

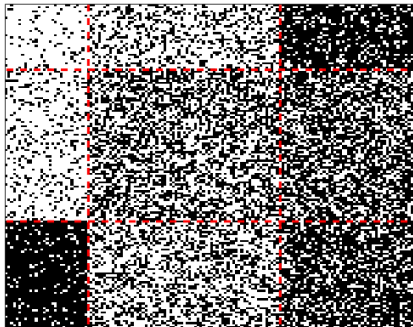
Stochastic Block Model

- The nodes are partitioned in q roles \mathcal{C}_i of size $m_i n$, $m := \sum_i m_i$
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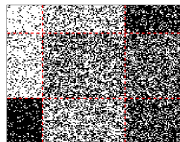
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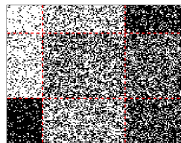
We work asymptotically in n , so we add some additional hypotheses:

- $0 < m_{\min} \leq m_i \leq m_{\max}$ for any i , and m_{\min}, m_{\max} are independent from n
- $B_{i,j} = \Psi_{i,j} f(n)$, $0 \leq \Psi_{i,j} \leq 1$ where the matrix Ψ is independent from n
- $[B \ B^T] = [\Psi \ \Psi^T] f(n)$ has full rank, equal to the number of roles q
- $nf(n) \rightarrow \infty$ that agrees with standard results for which $f(n) = \Omega(1/n)$ for exact recovery

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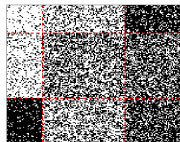
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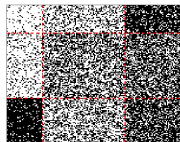
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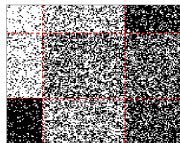
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If $D = \text{diag}(\sqrt{nm_i})$ and $U \Sigma U^T$ is the eigendecomposition of $D \hat{T}_k D$, then

$$T_k = Z \hat{T}_k Z^T = (Z D^{-1}) D \hat{T}_k D (Z D^{-1})^T = (Z D^{-1} U) \Sigma (Z D^{-1} U)^T$$

is the reduced eigendecomposition of T_k

Spectral Method on the Average Case

$$T_{k+1} := \begin{bmatrix} M & M^T \end{bmatrix} \begin{bmatrix} I + \beta^2 T_k & \\ & I + \beta^2 T_k \end{bmatrix} \begin{bmatrix} M^T \\ M \end{bmatrix} = \dots = Z \hat{T}_{k+1} Z^T$$

where all \hat{T}_k are PD $q \times q$ matrices

⇒ The number of roles can be inferred by the rank of all T_k

If $D = \text{diag}(\sqrt{nm_i})$ and $U\Sigma U^T$ is the eigendecomposition of $D\hat{T}_k D$, then

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The reduced orthogonal matrix $ZD^{-1}U \in \mathbb{R}^{nm \times q}$ has only q distinct rows, since

$$Z(D^{-1}U) = \begin{bmatrix} \text{blue} \\ \text{green} \\ \text{red} \end{bmatrix} = \begin{bmatrix} \text{blue} \\ \text{green} \\ \text{red} \end{bmatrix} \quad D^{-1}U \in \mathbb{R}^{q \times q}$$

⇒ The assignment function can be computed from the EVD of all T_k

Perturbation Result

Given $M = \mathbb{E}[A]$, $T_1 = \begin{bmatrix} M & M^T \end{bmatrix} \begin{bmatrix} M^T \\ M \end{bmatrix}$, and $T_{k+1} = \Gamma_M[I + \beta^2 T_k]$,

- The number of roles is the rank of any T_k
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We need to work on A , $S_1 = \begin{bmatrix} A & A^T \end{bmatrix} \begin{bmatrix} A^T \\ A \end{bmatrix}$, and $S_{k+1} = \Gamma_A[I + \beta^2 S_k]$ by

- Extracting the number of roles as the leading rank of S_k
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Theorem (Bai, Silverstein (2010))

For the matrix $Y := A - M$, almost surely

$$\|Y\|^2 \leq \delta^2 := 4mnf(n) \quad \|[Y \ Y^T]\|^2 \leq 2\delta^2$$

where $B = f(n)\Psi$, Ψ independent on n , and $A \in \{0, 1\}^{mn}$

Conjecture: $\|[Y \ Y^T]\|^2 \leq \left(\frac{1+\sqrt{2}}{2}\right)^2 \delta^2$ as suggested by MP distribution

Error Propagation

$$Y = A - M \quad \|[Y \ Y^T]\|^2 \leq 2\delta^2 = 8mnf(n) \quad \|M\| \sim \|B\| \|Z\|^2 = \Theta(\delta^2)$$

Recall that $S_{k+1} = \Gamma_A[I + \beta^2 S_k]$ and $T_{k+1} = \Gamma_M[I + \beta^2 T_k]$ and notice that

- $\Gamma_X[N] = \begin{bmatrix} X & X^T \end{bmatrix} \begin{bmatrix} N & \\ & N \end{bmatrix} \begin{bmatrix} X^T \\ X \end{bmatrix} \Rightarrow \|\Gamma_X\| = \|[X \ X^T]\|^2 \leq 2\|X\|^2$
- $\|\Gamma_A - \Gamma_M\| \sim \|[M \ M^T]\| \|[Y \ Y^T]\| = O(\delta^3)$
- If $\beta^2 < 1/(6\|A\|^2)$, then $\beta^2 \|\Gamma_A\| \leq 1/3$ and $\beta^2 \|\Gamma_M\| \leq 1/2$ almost surely

More precisely,

Theorem (B., N., V-D. (2022))

- $\|\Gamma_A - \Gamma_M\| \leq \delta^3 \|[\Psi \ \Psi^T]\| / \sqrt{2} + 2\delta^2$
- $\beta^2 < 1/(6\|A\|^2) \Rightarrow \gamma := \beta^2 \max\{\|\Gamma_M\|, \|\Gamma_A\|\} \leq 1/2$ a.s.

and they imply that

$$\|S_k - T_k\| \leq \left(\sum_{i=0}^{k-1} \gamma^i \right)^2 \|\Gamma_A - \Gamma_M\| \leq 4\|\Gamma_A - \Gamma_M\| = O(\delta^3)$$

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Number of Roles

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The number of roles q is $\text{rk}(T_k)$, so
how can we infer q from S_k ?

- q is the rank of $[M \ M^T]$ and all its non-zero singular values are $\Theta(\delta^2)$
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$$\begin{aligned} [M \ M^T] &= Z[B \ B^T] \begin{bmatrix} Z^T & \\ & Z^T \end{bmatrix} \\ &= U D f(n) [\Psi \ \Psi^T] \begin{bmatrix} D & \\ & D \end{bmatrix} V^T \end{aligned}$$

where U, V have orthonormal columns

$$U = ZD^{-1} \quad V = \begin{bmatrix} ZD^{-1} & \\ & ZD^{-1} \end{bmatrix}$$

and thus

$$m_{\min} \leq \frac{\sigma_i([M \ M^T])}{nf(n)\sigma_i([\Psi \ \Psi^T])} \leq m_{\max} \quad 1 \leq i \leq q$$

Recall that Ψ , m_{\max} and m_{\min} do not depend on n

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$$\begin{aligned} \|T_{k+1}\| &= \|\Gamma_M[I + \beta^2 T_k]\| \\ &\leq \|\Gamma_M\| + \|T_k\|/2 \end{aligned}$$

- The biggest q eigenvalues of S_k are $\Theta(\delta^4)$ almost surely

$$\leq \|\Gamma_M\| \left(\sum_{i=0}^k \frac{1}{2^i} \right) \leq 4\|M\|^2 = O(\delta^4)$$

- All other eigenvalues of S_k are $O(\delta^2)$

$$T_{k+1} \succeq T_k \succeq \cdots \succeq T_1 \implies \lambda_q(T_k) \geq \lambda_q(T_1)$$

$$\lambda_q(T_1) = \sigma_q([M \ M^T])^2 = \Theta(\delta^4)$$

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By Weyl's perturbation Theorem

$$|\sigma_i(S_k) - \sigma_i(T_k)| \leq \|S_k - T_k\|$$

but

$$\sigma_i(T_k) = \Theta(\delta^4) \gg O(\delta^3) = \|S_k - T_k\|$$

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$$\|S_{k+1}\| = \|\Gamma_A[I + \beta^2 S_k]\| \leq \|\Gamma_A\| + \|S_k\|/2$$

$$\leq \|\Gamma_A\| \left(\sum_{i=0}^k \frac{1}{2^i} \right) \leq 2\|\Gamma_A\|$$

$$S_{k+1} = \Gamma_A[I + \beta^2 S_k] \preceq (1 + \beta^2 \|S_k\|) \Gamma_A[I] \preceq 2S_1$$

By Weyl on $Y = A - M$ and

$$S_1 = [A \ A^T] \begin{bmatrix} A^T \\ A \end{bmatrix} \implies \lambda_i(S_1) = \sigma_i([A \ A^T])^2$$

we conclude that for $i > q$

$$\lambda_i(S_k) \leq 2\lambda_i(S_1) = 2\sigma_i([A \ A^T])^2$$

$$\sigma_i([A \ A^T]) \leq \sigma_i([M \ M^T]) + \|[Y \ Y^T]\| = O(\delta)$$

Number of Roles

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The dominant q eigenvalues of S_k are well separated from the noise

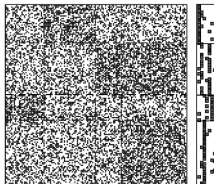
Theorem (B., N., V-D. (2022))

$$\frac{1}{2} \|[\Psi \ \Psi^T]\|^2 \delta^4 \geq \|S_k\| = \lambda_1(S_k)$$

$$\lambda_q(S_k) \geq \frac{1}{2} \left[\frac{\sigma_q([\Psi \ \Psi^T])}{4q} \frac{m_{\min}}{m_{\max}} \right]^2 \delta^4$$

$$4\delta^2 \geq \lambda_{q+1}(S_k)$$

Notice: if $[\Psi \ \Psi^T]$ is almost singular, it is hard to infer the number of roles



Dominant Subspaces

Recall that $\text{rk } T_k = q$ and the clustering assignment corresponds to the different rows of Q_k in the reduced EVD of $T_k = Q_k \Sigma_k Q_k^T$

As a consequence, we work on the q -dominant subspace given by the reduced EVD of $S_k = V_k \tilde{\Sigma}_k V_k^T$, but how are they related?

Theorem (Davis, Kahan (1970))

Given E, F , the q -dominant subspaces of S_k, T_k

$$\|\Pi_E - \Pi_F\| \leq 2 \frac{\|S_k - T_k\|}{\lambda_q(T_k)}$$

If we now plug in our previous estimations

$$\|\Pi_E - \Pi_F\| \leq \frac{4\sqrt{2}\delta^3 \|[\Psi \ \Psi^T]\| + 16\delta^2}{\left[\frac{\sigma_q([\Psi \ \Psi^T])}{4q} \frac{m_{\min}}{m_{\max}} \right]^2 \delta^4} = O(\delta^{-1}).$$

Notice that an almost singular $[\Psi \ \Psi^T]$ produces a small $\sigma_q([\Psi \ \Psi^T])$ and thus a bigger distance between the subspaces

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Misclassification Error

$$\|\Pi_E - \Pi_F\| = O(\delta^{-1}) \quad \|[Y \ Y^T]\|^2 \leq 2\delta^2 = 8mnf(n)$$

Given the q -reduced EVD $T_k = U_k \Sigma_k U_k^T$, K-means on the rows of U_k (whose columns are a basis of F) returns the exact role partition

Given the q -reduced EVD $S_k = V_k \tilde{\Sigma}_k V_k^T$, we apply K-means on the rows of V_k (whose columns are a basis of E) to obtain the roles

Given the exact roles $\mathcal{C}_1, \dots, \mathcal{C}_q$ and $\mathcal{T}_1, \dots, \mathcal{T}_q$ the resulting roles from the algorithm on S_k , let the **misclassification error** \hat{f}_k be

$$\hat{f} := \min_{\pi \in \mathcal{S}_q} \max_{i=1, \dots, q} \frac{|\mathcal{T}_{\pi(i)} \triangle \mathcal{C}_i|}{|\mathcal{C}_i|}$$

where \triangle is the symmetric difference of sets

Theorem

There exists an absolute constant C such that

$$\hat{f} \leq Cq \frac{m_{\max}}{m_{\min}} \|\Pi_E - \Pi_F\|^2 \leq C \frac{q^5}{\delta^2} \frac{m_{\max}^5}{m_{\min}^5} \frac{\|[Y \ Y^T]\|^2}{\sigma_q([Y \ Y^T])^4} = O\left(\frac{1}{nf(n)}\right)$$

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The Idea Behind

Theorem (Sheffet, Awasthi (2012))

Let U, V be $mn \times q$ matrices, where U has only q distinct rows μ_1, \dots, μ_q that identify the roles \mathcal{C}_i and call

$$\Delta_i := \frac{1}{\sqrt{|\mathcal{C}_i|}} \min\{\sqrt{q}\|U - V\|, \|U - V\|_F\}$$

Suppose there exists $\rho \geq 100$ such that $\|\mu_i - \mu_j\| \geq \rho(\Delta_i + \Delta_j)$ for any $i \neq j$. If \mathcal{T}_i are the roles determined by the K-means algorithm on V , then there exists a permutation π and an absolute constant C such that

$$\hat{f} \leq \max_r \frac{|\mathcal{C}_r \Delta \mathcal{T}_{\pi(r)}|}{|\mathcal{C}_r|} \leq \frac{C}{\rho^2}$$

In our case, let $T_k = U_k \Sigma_k U_k^T$ and $S_k = V_k \tilde{\Sigma}_k V_k^T$ be q -reduced EVDs

- $T_k = Z \hat{T}_k Z^T = \dots = (ZD^{-1}W_k) \Sigma_k (ZD^{-1}W_k)^T \implies U_k = ZD^{-1}W_k$
so U_k has q distinct rows of the form $\nu_i / \sqrt{m_i n}$ where ν_i are orthonormal
- There exists a $q \times q$ orthogonal Q s.t. $\|U_k Q - V_k\|_F \leq \sqrt{2q} \|\Pi_E - \Pi_F\|$

$$\rho = \min_{i \neq j} \frac{\left\| \frac{\mu_i}{nm_i} - \frac{\mu_j}{nm_j} \right\|}{\Delta_i + \Delta_j} \geq \frac{\sqrt{\frac{1}{m_i} + \frac{1}{m_j}}}{\frac{1}{\sqrt{m_i}} + \frac{1}{\sqrt{m_j}}} \frac{1}{\sqrt{2q} \|\Pi_E - \Pi_F\|} = \Omega(\delta) \rightarrow \infty$$

$$\implies \hat{f} = O(\rho^{-2}) = O(\delta^{-2}) = O(1/(nf(n)))$$

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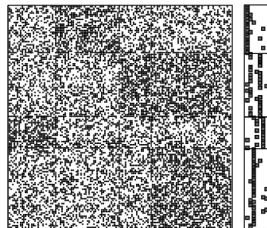
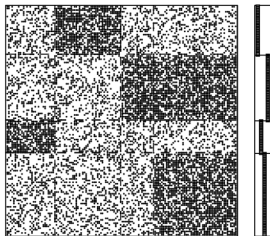
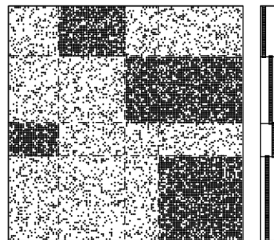
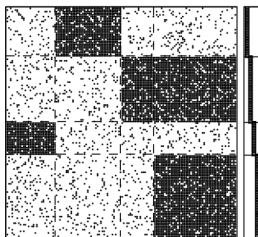
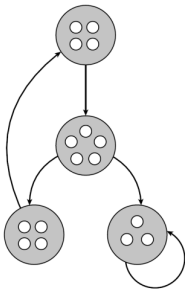
- $T_k = Z \hat{T}_k Z^T = \dots = (ZD^{-1}W_k) \Sigma_k (ZD^{-1}W_k)^T \implies U_k = ZD^{-1}W_k$
so U_k has q distinct rows of the form $\nu_i / \sqrt{m_i n}$ where ν_i are orthonormal
- There exists a $q \times q$ orthogonal Q s.t. $\|U_k Q - V_k\|_F \leq \sqrt{2q} \|\Pi_E - \Pi_F\|$

$$\rho = \min_{i \neq j} \frac{\left\| \frac{\mu_i}{nm_i} - \frac{\mu_j}{nm_j} \right\|}{\Delta_i + \Delta_j} \geq \frac{\sqrt{\frac{1}{m_i} + \frac{1}{m_j}}}{\frac{1}{\sqrt{m_i}} + \frac{1}{\sqrt{m_j}}} \frac{1}{\sqrt{2q} \|\Pi_E - \Pi_F\|} = \Omega(\delta) \rightarrow \infty$$

$$\implies \hat{f} = O(\rho^{-2}) = O(\delta^{-2}) = O(1/(nf(n)))$$

Numerical Example

We have already seen some: for $p = .9, .8, .7, .6$ and S_{10} we have

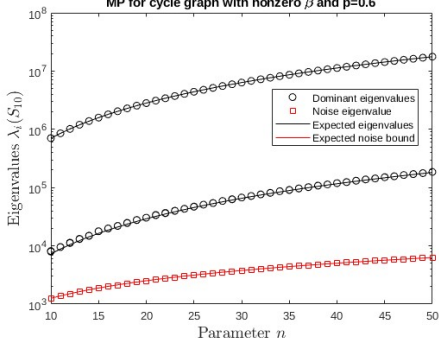


Numerical Example

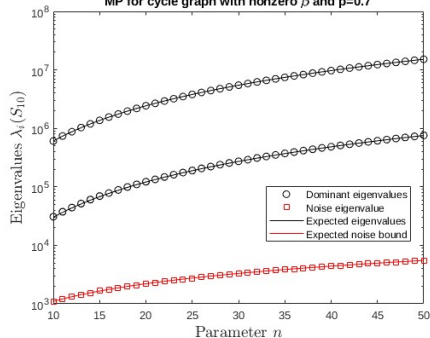
Here instead we compare the eigenvalues of T_{10} with those of S_{10} where the matrix dimension is $30n$ and

$$B = p \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix} + (1-p) \begin{bmatrix} 1 & 0 & 1 \\ 1 & 1 & 0 \\ 0 & 1 & 1 \end{bmatrix}$$

MP for cycle graph with nonzero β and $p=0.6$



MP for cycle graph with nonzero β and $p=0.7$

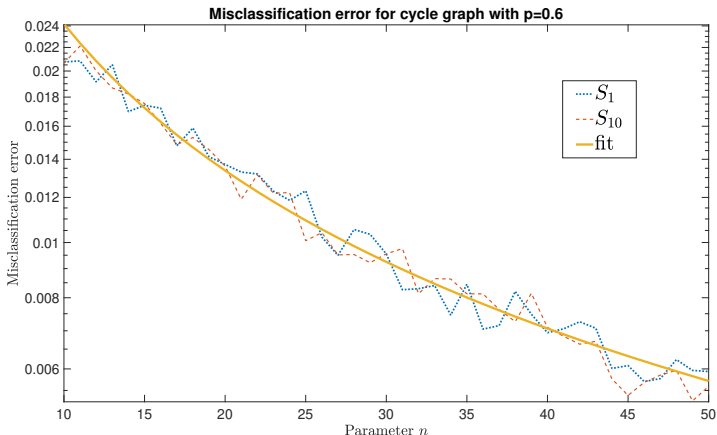


The eigenvalues estimations are more accurate when taking into consideration the **conjecture**: $\| [Y \ Y^T] \|^2 \leq \left(\frac{1+\sqrt{2}}{2} \right)^2 \delta^2$






Numerical Example

Here is the misclassification error for S_1 and S_{10} where the matrix dimension is $30n$, the yellow line is a fit for the estimated bound $\hat{f} \leq C/n$ and

$$B = p \begin{bmatrix} 0 & 1 & 0 \\ 0 & 0 & 1 \\ 1 & 0 & 0 \end{bmatrix} + (1-p) \begin{bmatrix} 1 & 0 & 1 \\ 1 & 1 & 0 \\ 0 & 1 & 1 \end{bmatrix}$$



Thank You!

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