

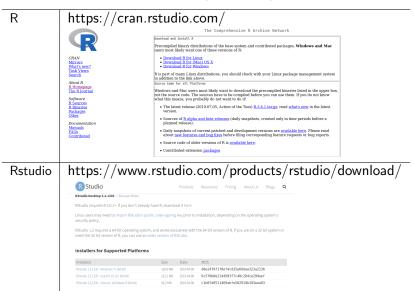
Crash-course on Bayesian statistics with  ${\sf R}$ 

Model-guided data science, Como 5 Sep 2019

#### Plan

PART 1 Facts about Bayesian stats
PART 2 Practical examples with *RJags* 

### 1. Download and install Rstudio (requires R):



- 2. Install the JAGS library.
  - ► linux/ubuntu:

from package manager

windows:

3. Open Rstudio and install RJAGS by running > install.packages("rjags")

https://sourceforge.net/projects/mcmc-jags/files/JAGS/4.x/Windows/

# PART 1

### Introduction

- ► Understanding data variability
- Statistical confidence and probability
- ► Building statistical models
- ▶ Inference of "unknown" features from "known" data

### Statistical models

Statistical models provide a probabilistic description of how the data have been generated. We will consider three types of models

- ▶ **simple models:** data are modeled by a single known probability distribution (normal, exponential etc..)
- mixed models: data are distributed according to a linear combination of known distributions
- ▶ hierarchical models: the parameters are hierarchically organized

Statistical models can involve *latent* (unobserved), features which are part of the generative model.

### Parameters and likelihood

Given a statistical model M characterized by a set of parameters  $oldsymbol{ heta}$  the data **likelihood** 

$$L_M(\theta) = \operatorname{Prob}(data|\theta, M)$$

measures the probability to generate the data from model M using the given parameter set  $\theta$ .

In the presence of latent variables  ${\bf z}$  alongside the observed variables  ${\bf x}$ , the likelihood is expresse as the sum over the unobserved variables

$$L_M(\theta) = \sum_{\mathbf{z}} \operatorname{Prob}(\mathbf{z} = data, \mathbf{z} | \theta, M)$$

# Bayes theorem

Assume we have a set of data and a statistical model with parameters  $\theta$ . We are interested in the "posterior" probability

$$\text{Prob}(\theta|\text{data},\text{Model})$$

Meaning: how does my uncertainty about the model parameters change after observing the data?

$$\begin{split} P(\theta, \text{data}|\text{Model}) = & P(\text{data}|\text{Model}) \cdot P(\theta|\text{data}, \text{Model}) \\ = & P(\theta|\text{Model}) \cdot P(\text{data}|\theta, \text{Model}) \end{split}$$

#### Posterior distribution

$$P(\theta | ext{data}, ext{Model}) = \underbrace{\frac{P( ext{data} | \theta, ext{Model}) \cdot P(\theta | ext{Model})}_{P( ext{data} | ext{Model})} \cdot \underbrace{P(\theta | ext{Model})}_{\text{marginal likelihood}}$$

### **Expectations**

Given the posterior distribution P(x) we can compute posterior averages of any quantity as usual

$$\mathbb{E}[x] = \langle x \rangle_P = \sum_{x \in \Omega} P(x) \cdot x$$

of any function of the model variables

$$\mathbb{E}[g(x)] = \langle g(x) \rangle_P = \sum_{x \in \Omega} P(x) \cdot g(x)$$

# Sampling and Monte Carlo approximation

To calculate expectations from the formulas above we need to perform sums or integrals (often in many dimensions)  $\rightarrow$  HARD! A much easier way to compute expectations is to generate first a large "sample" of values of x drawn from its probability distribution  $\{x_1, \dots, x_N\}$  and use the **Monte Carlo approximation** 

### Monte Carlo approximation

$$x_i \sim P(x), \quad i = 1, \dots, N$$
  
$$\mathbb{E}[g(x)] \approx \frac{1}{N} \sum_{i=1}^{N} g(x_i)$$

" $\sim$ "  $\rightarrow$  "drawn from"

# Summary so far

- ► Bayesian inference corresponds to the computation of posterior probabilities.
- ► In general we need Monte Carlo algorithms to sample efficiently from the posterior distribution
- We can use the Monte Carlo approximation of the posterior distribution to estimate model variables of interest (means, variances, correlations...).
- We still need to focus on appropriate prior distributions to represent our prior knowledge on the model parameters.

# PART 2

### script\_1.R: normal distribution

Estimation of mean and variance from univariate normally distributed data

- 1. Run script
- 2. Plot histograms of sampled mean and variance
- 3. Modify the prior distribution of the mean and resample
- 4. Compare the variance of the posterior distribution of the mean with the SEM estimated from the data
- 5. Calculate credible intervals for mean and variance using the function quantile

# script\_2.R: linear regression

Model:

$$egin{aligned} y|x &\sim \mathcal{N}(\mathsf{a}x+b, \sigma = 1/\sqrt{t}) \ &a \sim \mathcal{N}(1, 10^{-3}) \ &b \sim \mathcal{N}(1, 10^{-3}) \ &t \sim \mathsf{gamma}(1, 1) \end{aligned}$$

- 1. obtain posterior distributions of a and b
- 2. obtain posterior distribution of  $\sigma = 1/\sqrt{t}$
- 3. Are a and b correlated? Why?
- 4. If we define  $y^{pred}(x) \equiv ax + b$ , what is the distribution of  $y^{pred}(x = 13)$ ?
- 5. What is the difference between the previous calculation and P(y|x=13, data)?

# script\_3.R: logistic regression

Assume age affects the onset of a desease and we want to quantify this effect. We can model the probability of developing the desease before a given age using the logistic distribution

$$P_{logistic}(A) = \frac{1}{1 + e^{\alpha \cdot A}}.$$

Our statistical model looks like

status 
$$\sim$$
 bernoulli $(P_{logistic}(A))$   
 $lpha \sim \mathcal{N}(0, 10^{-3})$ 

- 1. plot most probable  $P_{logistic}(A)$
- 2. what is the probability to develop the desease before the age of 50? With which confidence?

### script\_4.R: Genetic effect on a desease

As in exercise 6 from previous lecture, consider the four probabilities  $p_{s,A}, p_{h,A}, p_{s,B}, p_{h,B}$  where A and B are variants of a gene and (s,h) stand for "sick" and "healthy" condition of the patient. We can use the model

$$\begin{split} & \text{Model}: \{N_{s,A}, N_{h,A}, N_{s,B}, N_{h,B}\} \sim \mathcal{M}(\{p_{s,A}, p_{h,A}, p_{s,B}, p_{h,B}\}; N = 100) \\ & \text{Prior}: \{p_{s,A}, p_{h,A}, p_{s,B}, p_{h,B}\} \sim \textit{Dirichlet}(\vec{\alpha}) \end{split}$$

- 1. calculate the distributions of P(s|A) and P(s|B).
- 2. calculate credible intervals
- 3. With which statistical confidence can we claim that the variant of the gene play a role in the desease?

### script\_5.R: Hierarchical model

Consider measurements of cell counts. We want to model the fraction of small and big cells given data from three trials:

	trial 1	trial 2	trial 3
big	10	25	35
small	90	75	65

We can use a hierarchical model to describe variability across trial

$$n_k^{big} \sim binomial(p_k, 100), \quad k = 1, 2, 3$$
  $p_k \sim Beta(\alpha, \beta)$   $\alpha, \beta \sim exponential(10^{-3})$ 

1. compare the distribution of  $p^{big}$  obtained in this model with the distribution obtained from model1 (see the script). Are they different? why?

### script\_6.R: Hierarchical model of normals

Consider now a similar situation where we measured the expression of a specific gene from single cells. Suppose we have data from three independent trials. Let us consider the hierarchical model

$$egin{aligned} g_i &\sim \mathcal{N}(\mu^T, t^T) \ \mu^T &\sim \mathcal{N}(\mu_0, t_0) \ t^T &\sim \textit{gamma}(2, 10^{-2}) \ \mu_0 &\sim \mathcal{N}(0, 10^{-3}) \end{aligned}$$

- 1. plot distributions of  $\mu$  and  $\mu_0$
- 2. calculate the standard deviation of  $\mu_0$  and compare with the naive SEM obtained from the data.

### script\_7.R: Mixture of Gaussians

We can model the angle distribution with a mixture of two normal distributions. We define a "cell type" variable (latent) t = 1, 2 so that we have

$$angle_i \sim Normal(\mu_i, t_i)$$
 $t_i \sim Binomial(1/2, N)$ 
 $\mu_1, \mu_2 \sim Normal(0, 0.001)$ 
 $\sigma_1, \sigma_2 \sim Gamma(2, 1)$ 

### NOTE 1: R multi-dimensional array

RJAGS returns draws from the posterior distribution in the format of a multidimensional array, which is the R container for multivariate data.

► For single numeric parameters:

 $lackbox{ For vector parameters } \theta = \{\theta_1, \theta_2, \cdots, \theta_K\}$ 

ightharpoonup Form matrix parameters  $\sigma$ 

# NOTE 2: Prediction from a model and propagation of parameter uncertainties

Given a model M with parameters  $\theta$  and some data, making a prediction on a new data point x correspond to evaluate the probability

$$P(x|data, M) = \sum_{\theta} P(x, \theta|data, M)$$

$$= \sum_{\theta} \underbrace{P(\theta|data, M)}_{posterior} \cdot P(x|\theta, M)$$

where we used the fact that x is conditionally independent of the data given the set of parameter  $\theta$ . Now we can use the Monte Carlo approximation to get

$$P(x|data, M) \approx \frac{1}{N} \sum_{i} P(x|\theta_i, M)$$

where the sum runs over all the sampled values of the parameters.

### NOTE 3: Marginal likelihood

The marginal likelihood (model evidence) is

$$egin{aligned} L_{M}( extit{data}) &= \sum_{ heta} P( extit{data}, heta | M) = \ &= \sum_{ heta} P( heta | M) \cdot P( extit{data} | heta, M) \end{aligned}$$