

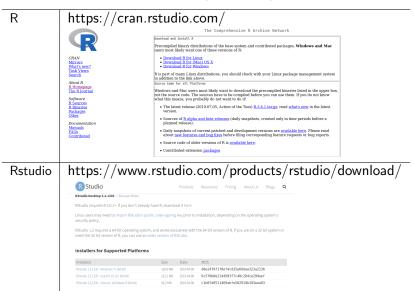
Crash-course on Bayesian statistics with R

BioBit@CDN, zoom October 21, 2020

Plan

PART 1 Facts about Bayesian stats
PART 2 Practical examples with *RJags*

1. Download and install Rstudio (requires R):



- 2. Install the JAGS library.
 - ► linux/ubuntu:
 - from package manager or source: http://mcmc-jags.sourceforge.net/
 - windows:
 - https://sourceforge.net/projects/mcmc-jags/files/JAGS/4.x/Windows/ for more infos about the jags library visit the webpage
 - http://mcmc-jags.sourceforge.net/
- 3. Open Rstudio and install RJAGS by running
 - > install.packages("rjags")
- 4. Lecture notes and R scripts available on github https://github.com/giovannidiana/BioBayes

PART 1

Introduction

- ► Understanding data variability
- Statistical confidence and probability
- ► Building statistical models
- ▶ Inference of "unknown" features from "known" data

Statistical models

Statistical models provide a probabilistic description of how the data have been generated. We will consider three types of models

- ▶ **simple models:** data are modeled by a single known probability distribution (normal, exponential etc..)
- mixed models: data are distributed according to a linear combination of known distributions
- ▶ hierarchical models: the parameters are hierarchically organized

Statistical models can involve *latent* (unobserved), features which are part of the generative model.

Parameters and likelihood

Given a statistical model M characterized by a set of parameters $oldsymbol{ heta}$ the data **likelihood**

$$L_M(\theta) = \operatorname{Prob}(data|\theta, M)$$

measures the probability to generate the data from model M using the given parameter set θ .

In the presence of latent variables ${\bf z}$ alongside the observed variables ${\bf x}$, the likelihood is expresse as the sum over the unobserved variables

$$L_M(\theta) = \sum_{\mathbf{z}} \operatorname{Prob}(\mathbf{z} = data, \mathbf{z} | \theta, M)$$

Bayes theorem

Assume we have a set of data and a statistical model with parameters θ . We are interested in the "posterior" probability

$$\text{Prob}(\theta|\text{data},\text{Model})$$

Meaning: how does my uncertainty about the model parameters change after observing the data?

$$\begin{split} P(\theta, \text{data}|\text{Model}) = & P(\text{data}|\text{Model}) \cdot P(\theta|\text{data}, \text{Model}) \\ = & P(\theta|\text{Model}) \cdot P(\text{data}|\theta, \text{Model}) \end{split}$$

Posterior distribution

$$P(\theta | ext{data}, ext{Model}) = \underbrace{\frac{P(ext{data} | \theta, ext{Model}) \cdot P(\theta | ext{Model})}_{P(ext{data} | ext{Model})} \cdot \underbrace{P(\theta | ext{Model})}_{\text{marginal likelihood}}$$

Expectations

Given the posterior distribution P(x) we can compute posterior averages of any quantity as usual

$$\mathbb{E}[x] = \langle x \rangle_P = \sum_{x \in \Omega} P(x) \cdot x$$

of any function of the model variables

$$\mathbb{E}[g(x)] = \langle g(x) \rangle_P = \sum_{x \in \Omega} P(x) \cdot g(x)$$

Sampling and Monte Carlo approximation

To calculate expectations from the formulas above we need to perform sums or integrals (often in many dimensions) \rightarrow HARD! A much easier way to compute expectations is to generate first a large "sample" of values of x drawn from its probability distribution $\{x_1, \dots, x_N\}$ and use the **Monte Carlo approximation**

Monte Carlo approximation

$$x_i \sim P(x), \quad i = 1, \dots, N$$

$$\mathbb{E}[g(x)] \approx \frac{1}{N} \sum_{i=1}^{N} g(x_i)$$

" \sim " \rightarrow "drawn from"

Summary so far

- ► Bayesian inference corresponds to the computation of posterior probabilities.
- ► In general we need Monte Carlo algorithms to sample efficiently from the posterior distribution
- We can use the Monte Carlo approximation of the posterior distribution to estimate model variables of interest (means, variances, correlations...).
- We still need to focus on appropriate prior distributions to represent our prior knowledge on the model parameters.

PART 2

script_1.R: normal distribution

Estimation of mean and variance from univariate normally distributed data

- 1. Run script
- 2. Plot histograms of sampled mean and variance
- 3. Modify the prior distribution of the mean and resample
- 4. Compare the variance of the posterior distribution of the mean with the SEM estimated from the data
- 5. Calculate credible intervals for mean and variance using the function quantile

script_2.R: linear regression

Model:

$$egin{aligned} y|x &\sim \mathcal{N}(\mathsf{a}x+b, \sigma = 1/\sqrt{t}) \ &a \sim \mathcal{N}(1, 10^{-3}) \ &b \sim \mathcal{N}(1, 10^{-3}) \ &t \sim \mathsf{gamma}(1, 1) \end{aligned}$$

- 1. obtain posterior distributions of a and b
- 2. obtain posterior distribution of $\sigma = 1/\sqrt{t}$
- 3. Are a and b correlated? Why?
- 4. If we define $y^{pred}(x) \equiv ax + b$, what is the distribution of $y^{pred}(x = 13)$?
- 5. What is the difference between the previous calculation and P(y|x=13, data)?

script_3.R: logistic regression

Assume age affects the onset of a desease and we want to quantify this effect. We can model the probability of developing the desease before a given age using the logistic distribution

$$P_{logistic}(A) = \frac{1}{1 + e^{\alpha \cdot A}}.$$

Our statistical model looks like

status
$$\sim$$
 bernoulli $(P_{logistic}(A))$
 $\alpha \sim \mathcal{N}(0, 10^{-3})$

- 1. plot most probable $P_{logistic}(A)$
- 2. what is the probability to develop the desease before the age of 50? With which confidence?

script_4.R: Genetic effect on a desease

Consider the four probabilities $p_{s,A}$, $p_{h,A}$, $p_{s,B}$, $p_{h,B}$ where A and B are variants of a gene and (s,h) stand for "sick" and "healthy" condition of the patient. We can use the model

Model:
$$\{N_{s,A}, N_{h,A}, N_{s,B}, N_{h,B}\} \sim \mathcal{M}(\{p_{s,A}, p_{h,A}, p_{s,B}, p_{h,B}\}; N = 100)$$

Prior: $\{p_{s,A}, p_{h,A}, p_{s,B}, p_{h,B}\} \sim Dirichlet(\vec{\alpha})$

- 1. calculate the distributions of P(s|A) and P(s|B).
- 2. calculate credible intervals
- 3. With which statistical confidence can we claim that the variant of the gene has a role in the desease?

script_5.R: Hierarchical model

Consider measurements of cell counts. We want to model the fraction of small and big cells given data from three trials:

	trial 1	trial 2	trial 3
big	10	25	35
small	90	75	65

We can use a hierarchical model to describe variability across trials

$$n_k^{big} \sim binomial(p_k, 100), \quad k = 1, 2, 3$$
 $p_k \sim Beta(lpha, eta)$ $lpha, eta \sim exponential(10^{-3})$

1. compare the distribution of p^{big} obtained in this model with the distribution obtained from model1 (see the script). Are they different? why?

script_6.R: Hierarchical model of normals

Consider now a similar situation where we measured the expression of a specific gene from single cells. Suppose we have data from three independent trials. Let us consider the hierarchical model

$$egin{aligned} g_i &\sim \mathcal{N}(\mu^T, t^T) \ \mu^T &\sim \mathcal{N}(\mu_0, t_0) \ t^T &\sim \textit{gamma}(2, 10^{-2}) \ \mu_0 &\sim \mathcal{N}(0, 10^{-3}) \end{aligned}$$

- 1. plot distributions of μ and μ_0
- 2. calculate the standard deviation of μ_0 and compare with the naive SEM obtained from the data.

script_7.R: Mixture of Gaussians

Consider the case where we quantified cell polarity in terms of cell orientations and we have reasons to believe that the distribution of orientations is a mixture of two normal distributions. We define a "cell type" variable (latent) t=1,2 so that we have

$$egin{aligned} & \mathsf{Anormal}(\mu_i, t_i) \ & t_i \sim \mathsf{Binomial}(1/2, \mathsf{N}) \ & \mu_1, \mu_2 \sim \mathsf{Normal}(0, 0.001) \ & \sigma_1, \sigma_2 \sim \mathsf{Gamma}(2, 1) \end{aligned}$$

NOTE 1: R multi-dimensional array

RJAGS returns draws from the posterior distribution in the format of a multidimensional array, which is the R container for multivariate data.

► For single numeric parameters:

 $lackbox{ For vector parameters } \theta = \{\theta_1, \theta_2, \cdots, \theta_K\}$

lacktriangle Form matrix parameters $oldsymbol{\sigma}$

NOTE 2: Prediction from a model and propagation of parameter uncertainties

Given a model M with parameters θ and some data, making a prediction on a new data point x correspond to evaluate the probability

$$P(x|data, M) = \sum_{\theta} P(x, \theta|data, M)$$

$$= \sum_{\theta} \underbrace{P(\theta|data, M)}_{posterior} \cdot P(x|\theta, M)$$

where we used the fact that x is conditionally independent of the data given the set of parameter θ . Now we can use the Monte Carlo approximation to get

$$P(x|data, M) \approx \frac{1}{N} \sum_{i} P(x|\theta_i, M)$$

where the sum runs over all the sampled values of the parameters.

NOTE 3: Marginal likelihood

The marginal likelihood (model evidence) is

$$egin{aligned} L_{M}(\textit{data}) &= \sum_{ heta} P(\textit{data}, heta | M) = \ &= \sum_{ heta} P(heta | M) \cdot P(\textit{data} | heta, M) \end{aligned}$$