

# **Lecture 10:**

# **Transformations**

# Transformations

We'll talk about **transformations** of r.v.s. After applying a function to a r.v.  $X$  (or random vector  $\mathbf{X}$ ), the goal is to find the distribution of the transformed variable.

Some examples of transformations:

- **Unit conversion:** In 1D, this is a loc-scale transform:  $Y = aX + b$
- **Sums and averages:**  $X_1, \dots, X_n \rightarrow \bar{X}_n = (X_1 + \dots + X_n)/n$
- **Convolutions:** we'll cover these in this lecture
- **Extreme values:**  $X_1, \dots, X_n \rightarrow \min / \max(X_1, \dots, X_n)$ , quantiles

Remember, if we just need the expectation of a transformed variable, we have LOTUS. But we want the whole distribution now.

# Change of variables

# Change of variables

**Theorem 8.1.1** (Change of variables in 1D). Let  $X$  be a continuous r.v. with PDF  $f_X$  and let  $Y = g(X)$ , where  $g$  is 1) differentiable and 2) strictly increasing (/decreasing). Then the PDF of  $Y$  is

$$f_Y(y) = f_X(x) \left| \frac{dx}{dy} \right|$$

where  $x = g^{-1}(y)$ . The support of  $Y$  is all  $g(x)$  with  $x$  in  $\text{Supp}(X)$ .

**8.1.2** When finding the distribution of  $Y$ , be sure to:

- 1) Check the assumptions of the theorem
- 2) Express the final answer for the PDF of  $Y$  as a function of  $y$
- 3) Specify the support of  $Y$

# Change of variables

**Example 8.1.3** (Log-Normal PDF). Let  $X \sim \mathcal{N}(0,1)$  and  $Y = e^X$ .

The distr. of  $Y$  is called **log-normal**. The PDF of  $Y$  is:

$$f_Y(y) = f_X(x) \left| \frac{dx}{dy} \right| = \varphi(x) \frac{1}{e^x} = \varphi(\log y) \frac{1}{y}, \quad y > 0$$

where  $\varphi(x) = e^{-x^2/2}/\sqrt{2\pi}$  is the normal PDF.

We can rather work with the CDF, of course:

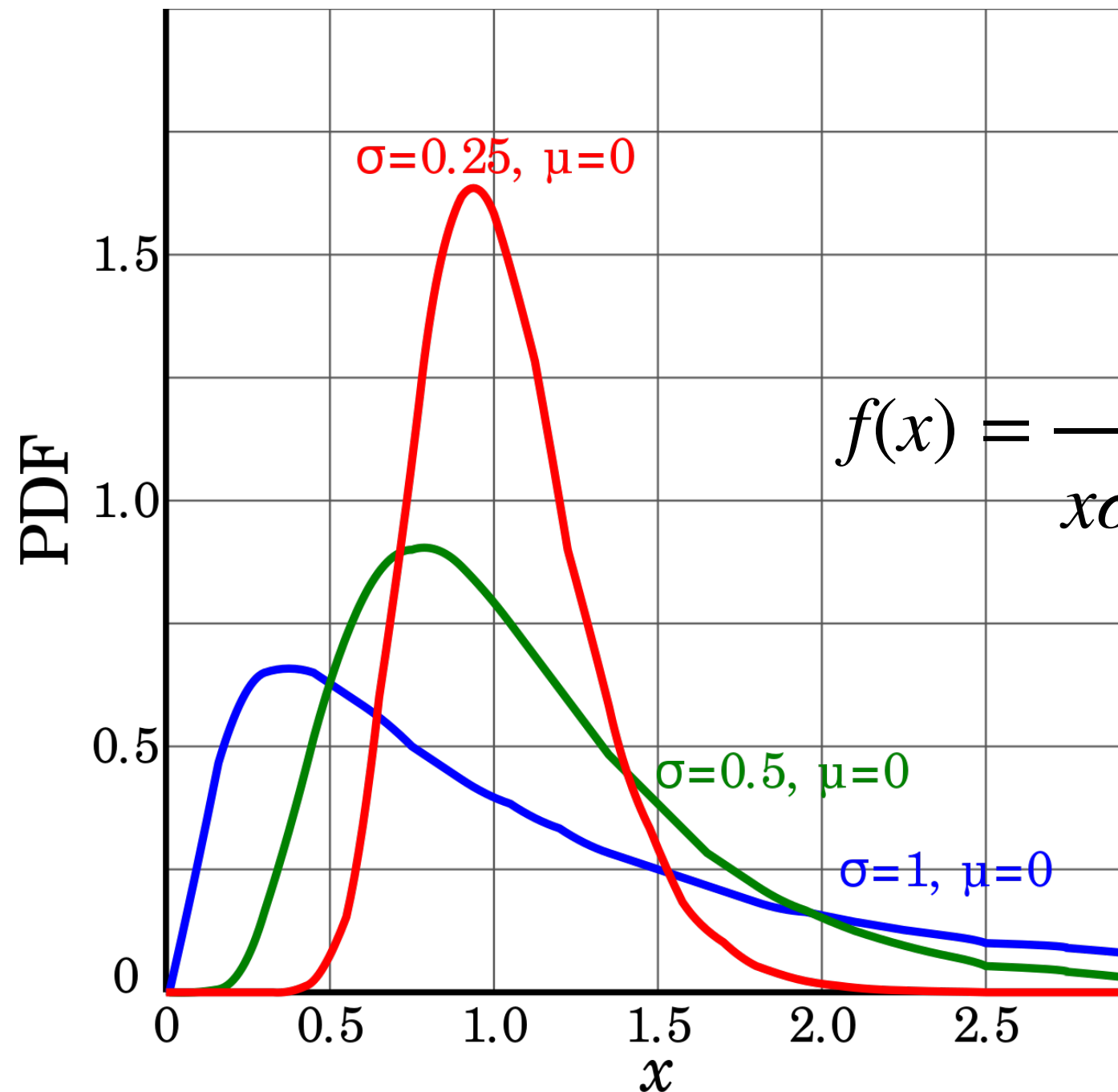
$$Y \leq y \rightarrow e^X \leq y \rightarrow X \leq \log Y, \text{ so:}$$

$$F_Y(y) = P(Y \leq y) = P(e^X \leq y) = P(X \leq \log y) = \Phi(\log y)$$

$$\text{and the PDF is, again, } f_Y(y) = \frac{d}{dy} \Phi(\log y) = \varphi(\log y) \frac{1}{y}, \quad y > 0$$

# Change of variables

General log-normal: let  $X \sim \mathcal{N}(\mu, \sigma)$  and  $Y = e^X$ .



$$f(x) = \frac{1}{x\sigma\sqrt{2\pi}} \exp\left(-\frac{((\log x) - \mu)^2}{2\sigma^2}\right)$$

# Change of variables

**Example 8.1.4** (Chi-Squared PDF). Let  $X \sim \mathcal{N}(0,1)$  and  $Y = X^2$ .

The distr. of  $Y$  is called **chi-squared**, and we can get it from CDF:

$$X^2 \leq y \quad \rightarrow \quad -\sqrt{y} \leq X \leq \sqrt{y},$$

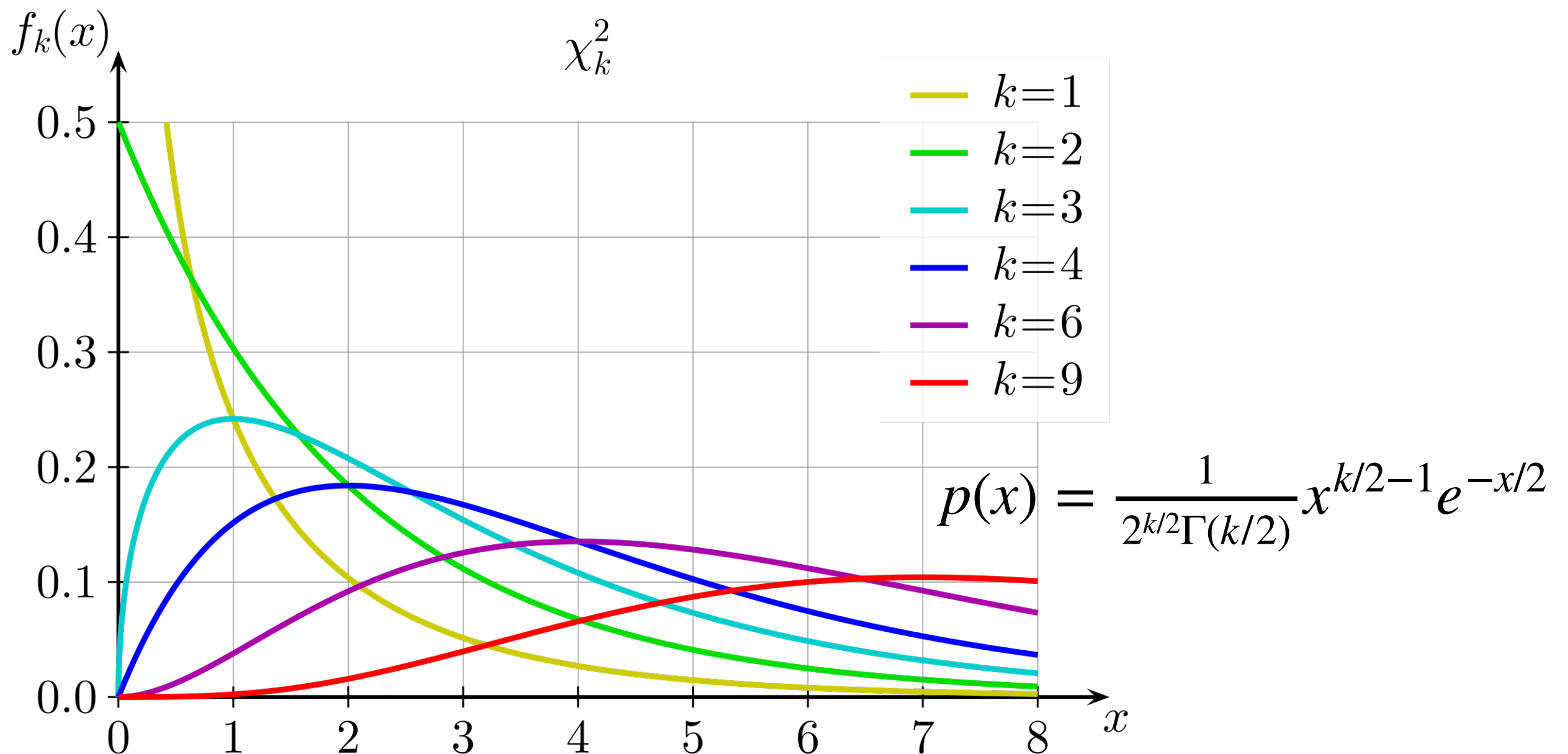
$$\begin{aligned} \text{so } F_Y(y) &= P(X^2 \leq y) = \\ &= P(-\sqrt{y} \leq X \leq \sqrt{y}) = \Phi(\sqrt{y}) - \Phi(-\sqrt{y}) = 2\Phi(\sqrt{y}) - 1 \end{aligned}$$

$$\text{so } f_Y(y) = 2\varphi(\sqrt{y}) \cdot \frac{1}{2}y^{-1/2} = \varphi(\sqrt{y})y^{-1/2}, \quad y > 0$$

# Change of variables

General chi-squared: let  $Z_i \sim \mathcal{N}(0,1)$ , then  $Q = \sum_{i=1}^k Z_i^2 \sim \chi_k^2$

- is said to have  $k$  degrees of freedom:





# Change of variables

**Theorem 8.1.7** (Change of multiple variables). Let  $\mathbf{X} = (X_1, \dots, X_n)$  be a random vector with joint PDF  $f_{\mathbf{X}}$ . Let  $g : A_0 \rightarrow B_0$  be an invertible function where  $A_0, B_0$  are open subsets of  $\mathbb{R}^n$ , and  $\text{Supp}(\mathbf{X}) \subset A_0$  and  $B_0$  is the range of  $g$ .

Let  $\mathbf{Y} = g(\mathbf{X})$ , so  $\mathbf{y} = g(\mathbf{x})$ . Since  $g$  is invertible,  $\mathbf{X} = g^{-1}(\mathbf{Y})$ .

Suppose that all partial derivatives  $\frac{\partial x_i}{\partial y_j}$  are continuous, so we

can make the **Jacobian matrix**:  $\left[ \frac{\partial \mathbf{x}}{\partial \mathbf{y}} \right]_{ij} = \frac{\partial x_i}{\partial y_j}$

Then the joint PDF of  $\mathbf{Y}$  is:  $f_{\mathbf{Y}}(\mathbf{y}) = f_{\mathbf{X}}(g^{-1}(\mathbf{y})) \cdot \left| \det \frac{\partial \mathbf{x}}{\partial \mathbf{y}} \right|$

We have  $\det \frac{\partial \mathbf{x}}{\partial \mathbf{y}} = \left( \det \frac{\partial \mathbf{y}}{\partial \mathbf{x}} \right)^{-1}$ , so use the one that is easier to find

# Change of variables

**Example 8.1.9** (Box-Muller). Let  $U \sim \text{Unif}(0, 2\pi)$  and let  $T \sim \text{Expo}(1)$  – independent of  $U$ . Now define:

$$X = \sqrt{2T} \cos U, \quad Y = \sqrt{2T} \sin U$$

The joint PDF of  $U, T$  is  $f_{U,T}(u, t) = \frac{1}{2\pi} e^{-t}$ , since  $U, T$  are independent.

Let's view  $(X, Y)$  as a point in the plane, then we have:

$$X^2 + Y^2 = 2T(\cos^2 U + \sin^2 U) = 2T$$

so  $(\sqrt{2T}, U)$  is the polar coordinates form of  $(X, Y)$ . This is an invertible transformation, let's find its Jacobian

# Change of variables

**Example 8.1.9** (Box-Muller).  $U \sim \text{Unif}(0, 2\pi)$ ,  $T \sim \text{Expo}(1)$  – indep.

$X = \sqrt{2T} \cos U$ ,  $Y = \sqrt{2T} \sin U$ , PDF of  $U, T$  is  $f_{U,T}(u, t) = \frac{1}{2\pi} e^{-t}$

$$\frac{\partial(x, y)}{\partial(u, t)} = \begin{pmatrix} -\sqrt{2t} \sin u & \frac{1}{\sqrt{2t}} \cos u \\ \sqrt{2t} \cos u & \frac{1}{\sqrt{2t}} \sin u \end{pmatrix} \quad \text{– the abs. determinant is}$$

$|\det \dots| = |-\sin^2 u - \cos^2 u| = 1$  is always 1.

So,  $f_{X,Y}(x, y) = f_{U,T}(u, t) \cdot |\det \dots| = \frac{1}{2\pi} e^{-t} = \frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)}$  – a

product of two st. normals! ( $X, Y$  are i.i.d.  $\mathcal{N}(0, 1)$  r.v.s!)

Hence the Box-Muller method for generating Normal r.v.s!

# Convolutions

# Convolutions

**Theorem 8.2.1** (Convolution). Let  $X$  and  $Y$  be independent r.v.s and  $T = X + Y$ . If  $X, Y$  are discrete, the PMF of  $T$  is

$$P(T = t) = \sum_x P(Y = t - x) P(X = x) = \sum_y P(X = t - y) P(Y = y)$$

if  $X, Y$  are continuous, then the PDF of  $T$  is

$$f_T(t) = \int_{-\infty}^{+\infty} f_Y(t - x) f_X(x) dx = \int_{-\infty}^{+\infty} f_X(t - y) f_Y(y) dy$$

This is called **convolution** of PMF-s/PDF-s. **Proof** is straightforward from LOTP.

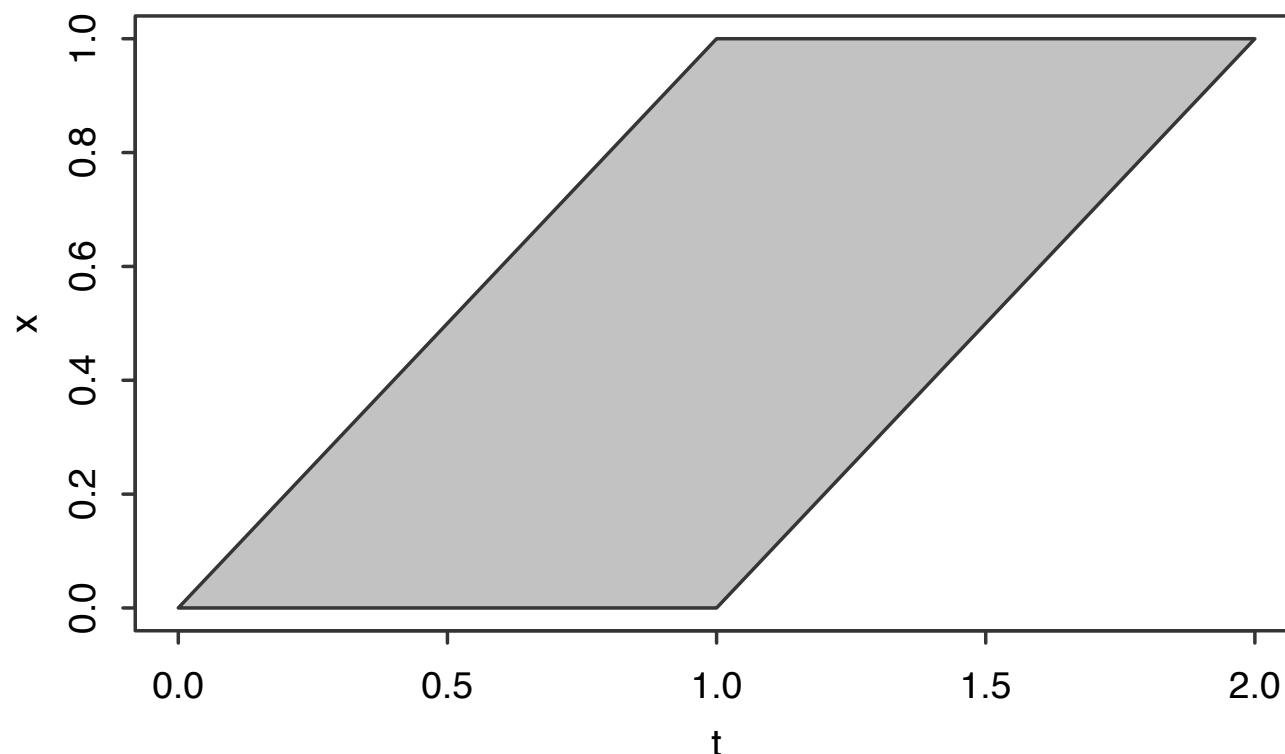
# Convolutions

**Example 8.2.5** (Uniform convolution). Let  $X, Y \stackrel{i.i.d}{\sim} \text{Unif}(0,1)$ .

Find the distribution of  $T = X + Y$

**Solution:** The PDF of both  $X$  and  $Y$  is  $g(x) = \begin{cases} 1, & x \in (0,1) \\ 0, & \text{otherwise} \end{cases}$

$$\text{Convolution: } f_T(t) = \int_{-\infty}^{+\infty} f_Y(t-x) f_X(x) dx = \int_{-\infty}^{+\infty} g(t-x) g(x) dx$$



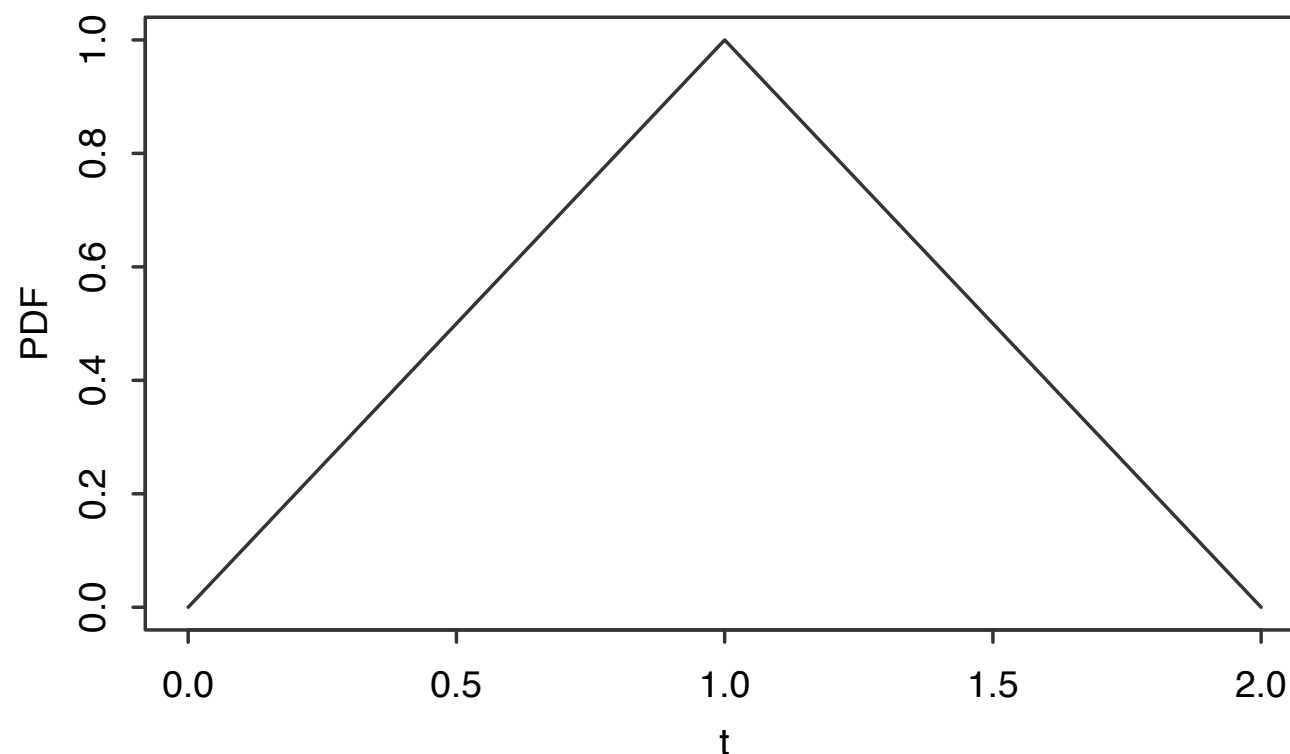
← area where  
 $g(t-x) g(x)$   
is nonzero

# Convolutions

**Example 8.2.5** (Uniform convolution). Let  $X, Y \stackrel{i.i.d}{\sim} \text{Unif}(0,1)$ .

Find the distribution of  $T = X + Y$

**Solution:** So  $f_T(t) = \begin{cases} \int_0^t dx = t, & 0 < t < 1, \\ \int_{t-1}^1 dx = 2 - t, & 1 < t < 2 \end{cases}$



← PDF  $f_T(t)$

# Convolutions

**Example 8.2.4** (Exponential convolution). Let  $X, Y \stackrel{i.i.d}{\sim} \text{Expo}(\lambda)$ .

Find the distribution of  $T = X + Y$

**Solution:** For  $t > 0$ ,

$$f_T(t) = \int_{-\infty}^{+\infty} f_Y(t-x)f_X(x) dx = \int_0^t \lambda e^{-\lambda(t-x)} \lambda e^{-\lambda x} dx$$

where we restricted the integral to be from 0 to  $t$  since both  $t-x$  and  $x$  should be  $> 0$ . That gives

$$f_T(t) = \lambda^2 \int_0^t e^{-\lambda t} dx = \lambda^2 e^{-\lambda t}, \quad t > 0 \quad - \text{this is Gamma}(2, \lambda)$$