Statistics

From probability to statistics

In probability:

$$X_i \sim F(x) = \mathbb{P}(X \leqslant x)$$

In statistics:

 $\{X_1,\ldots,X_n\} + ext{hypothesis} o ext{information}$

What kind of information?

- Distribution
- Distribution property Functional
- Hypothesis correctness

What we will do in this course

- Estimation
 - Point estimates = best guess
- Bayesian estimate = probabilistic guess • Confidence sets (another way to provide probabilistic guess)
- Hypothesis testing (confidence sets inversed)

Parametric vs non-parametric statistics • In probability, we studied different distributions, that were parametrized by one or more parameters.

- In statistics, we can think of these distributions as of models of the data, and try to find the parameters for these models best describing the data.
- Example: you own a chain of restaurants, and due to budget cuts, you would like to close one restaurant that is the least popular. In order to do that, for each restaurant
- you collect the data of visitor counts per day. Then you can fit a distribution to find the actual rates of visitors per day and then proceed to close the restaurant with the lowest rate. This is called parametric statistics.

• We can also go without any model assumption. How?

Parametric vs non-parametric statistics

- This will be non-parametric statistics.

Estimation

Estimation

Consider sample X_1, X_2, \ldots, X_n i.i.d. (simple sample). Consider its true distribution $F_{\theta}(x)$ (unknown) parametrized by θ . In case of parametric statistics this will be the set of distribution parameters, in case of non-parametric statistics it will be the distribution itself (next time). Consider an estimate $\hat{\theta}$. It is a **random variable**, since it is a function of sample, which is random. How **good** is this estimate?

Bias

 $ext{bias} \Big(\hat{ heta} \Big) = \mathbb{E} \left[\hat{ heta}
ight] - heta$

Bias is one measure of goodness of fit, describing the systematic error (not related to randomness).

$$\mathbb{E}\left[\hat{ heta}
ight]= heta$$

An estimate is called **unbiased** if bias is zero:

Example 1

Example with restaurant. Consider simple sample $X_1,\ldots,X_n\sim Pois(\lambda)$. We would like to estimate the parameter of Poisson distribution and we chose the following

estimator: $\widehat{\lambda} = rac{1}{n} \sum_{k=1}^n X_k$

Solution 1

What would be its bias?

$$\mathrm{bias}\Big(\widehat{\lambda}\Big) = \mathbb{E}\left[\widehat{\lambda}
ight] - \lambda = 0$$

 $\mathbb{E}\left[\widehat{\lambda}
ight] = \mathbb{E}\left[rac{1}{n}\sum_{k=1}^n X_k
ight] = rac{1}{n}\sum_{k=1}^n \mathbb{E}[X_k] = rac{1}{n}\sum_{k=1}^n \lambda = \lambda$

import scipy.stats as sts

import seaborn as sns

import matplotlib.pyplot as plt

In [3]: import numpy as np

0.0008

Standard error

Standard error is very important. We can have a perfectly unbiased estimator with large standard error.

Bias describes the tendency of our estimate, but what about its noisiness? This is where standard error comes in handy.

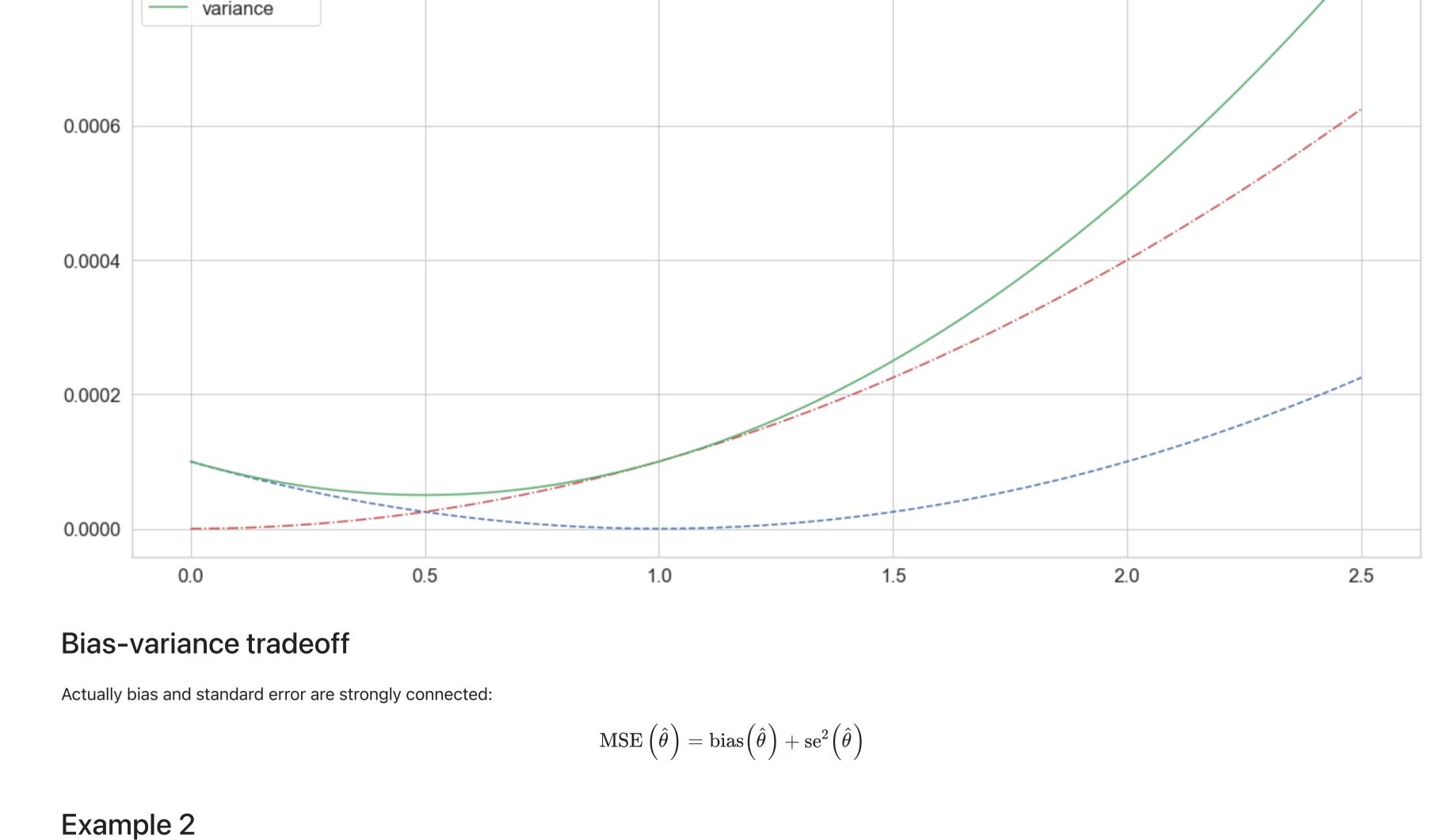
 $ext{se}ig(\hat{ heta}ig) = \sqrt{ ext{Var}ig(\hat{ heta}ig)}$

sns.set(style="whitegrid", font_scale=1.5) sns.despine()

%matplotlib inline In [21]: | 1bd = 1/100 |n = 100x = sts.poisson(lbd).rvs(n)lbd_hat = np.mean(x) scaling factors = np.linspace(0, 2.5, 100)

variance

biases = lbd * (scaling factors - 1) ses = scaling_factors * np.sqrt(lbd / n) fig, ax = plt.subplots(figsize=(20,10)) ax.plot(scaling_factors, biases ** 2, "b--", label="bias squared") ax.plot(scaling_factors, ses ** 2, "r-.", label="variance") ax.plot(scaling_factors, biases ** 2 + ses ** 2, 'g-', label="variance") ax.legend() <matplotlib.legend.Legend at 0x13b342640> bias squared



Example with restaurant. Consider simple sample $X_1,\dots,X_n\sim Pois(\lambda)$. We would like to estimate the parameter of Poisson distribution and we chose the following estimator:

What would be its standard error?

Solution 2

 $\mathbb{V}\mathrm{ar}\left(\widehat{\lambda}
ight) = \mathbb{E}\left[\widehat{\lambda}^2
ight] - \left(\mathbb{E}\left[\widehat{\lambda}
ight]
ight)^2$

 $\widehat{\lambda} = rac{1}{n} \sum_{k=1}^n X_k$

 $\mathbb{E}\left[\widehat{\lambda}^2
ight] = \mathbb{E}\left[\left(rac{1}{n}\sum_{k=1}^n X_k
ight)^2
ight] = \mathbb{E}\left[rac{1}{n^2}\sum_{k,m=1}^n X_k X_m
ight] = 0$ $=rac{1}{n^2}\Biggl(\sum_{k=1}^n\mathbb{E}\left[X_k^2
ight]+\sum_{k
eq m}E[X_k]E[X_m]\Biggr)=$ $=rac{1}{n^2}ig(n\cdot(\lambda^2+\lambda)+n(n-1)\cdot\lambda^2ig)=rac{1}{n}ig(n\lambda^2+\lambdaig)$ $\mathrm{se}ig(\widehat{\lambda}ig) = \sqrt{\mathbb{V}\mathrm{ar}\left(\widehat{\lambda}
ight)} = \sqrt{rac{1}{n}ig(n\lambda^2 + \lambdaig) - \lambda^2} = \sqrt{\lambda/n}$

Consistency and strong consistency $\hat{ heta} \stackrel{P}{ o} heta$

What is the base of the limit? What type of convergence is it? An estiamtor $\hat{\theta}$ is called strongly consistent if

An estiamtor $\hat{ heta}$ is called consistent if

What type of convergence is it?

Example 3

estimator:

 $\widehat{\lambda} = rac{1}{n} \sum_{k=1}^{n} X_k$ Is this estimator consistent?

Example with restaurant. Consider simple sample $X_1,\dots,X_n\sim Pois(\lambda)$. We would like to estimate the parameter of Poisson distribution and we chose the following

As I hope everyone remembers, convergence in probability is that convergence in probability is the weakest of all, and follows from other types of convergence.

We can see that $ext{MSE}\left(\widehat{\lambda}\right) o 0$, as $n o \infty$. This means convergence in mean-squared, and convergence in probability follows from it unconditionally. Strong consistency follows from strong LLN.

Solution 3

We have $\mathrm{bias}\Big(\widehat{\lambda}\Big)=0$ and $\mathrm{se}\Big(\widehat{\lambda}\Big)=\sqrt{\lambda/n}.$ From bias-variance tradeoff, we can see that $ext{MSE}\left(\widehat{\lambda}
ight) = ext{bias}^2\Big(\widehat{\lambda}\Big) + ext{se}^2\Big(\widehat{\lambda}\Big) = \lambda/n$