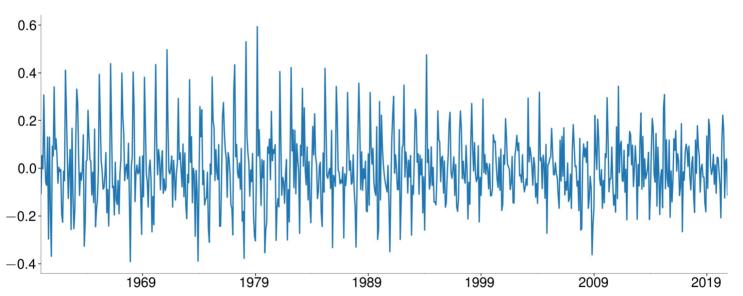
Univariate Time Series Analysis

Kevin Sheppard

Seasonality

• A seasonal time series has a deterministic pattern that repeats on an annual basis

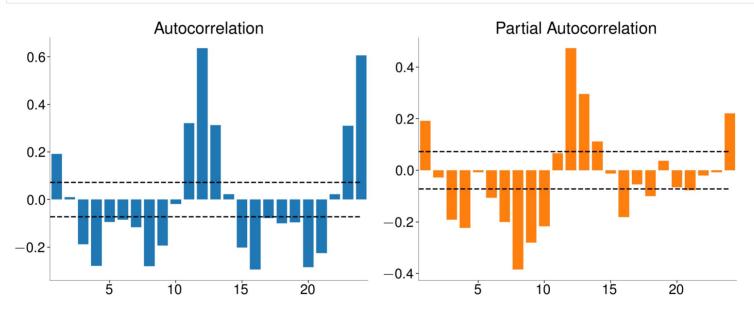
In [5]: plot(housing, y=-2)



Seasonal Autocorrelation Pattern

• Seasonal data has dynamics at the annual frequency

```
In [6]: acf_pacf_plot(housing, 24, size=-2)
```



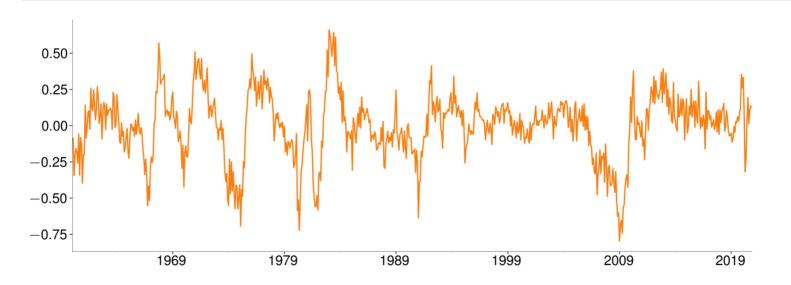
Seasonal Differencing

ullet Seasonal differencing differences using the seasonal period s

$$\Delta_s Y_t = Y_t - Y_{t-s}$$

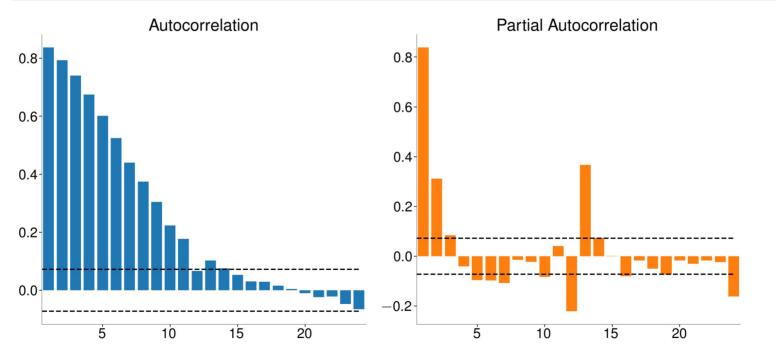
• Can reduce or eliminate seasonalities

In [7]: plot(housing yoy, y=12)



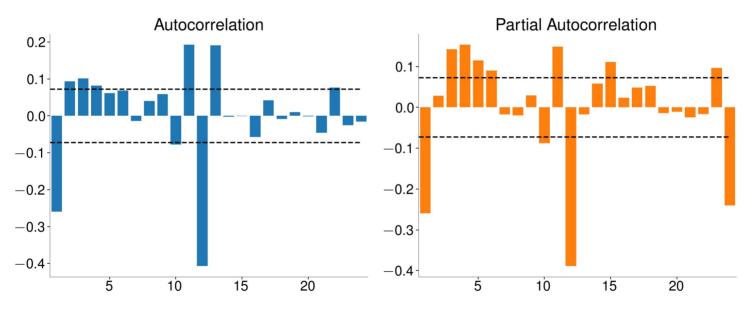
Seasonal Difference Autocorrelation Pattern

```
In [8]: acf_pacf_plot(housing_yoy, 24)
```



Modeling Seasonal Differences

```
In [9]: res = SARIMAX(housing_yoy, order=(1, 0, 0)).fit()
    resids = res.resid.iloc[1:]
    acf_pacf_plot(resids, 24, size=-2)
```



Seasonal ARMA Models

- SARMA models data at both the observational and seasonal frequency
- In lag polynomial representation

$$(1-\phi_1L-\phi_2L^2)Y_t=\phi_0+(1+ heta_sL^s)\epsilon_t$$

In standard ARMA representation

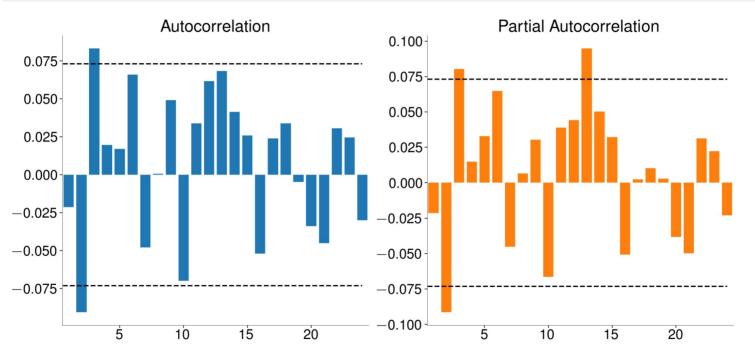
$$Y_{t} = \phi_{0} + \phi_{1}Y_{t-1} + \phi_{2}Y_{t-2}\theta_{s}\epsilon_{t-s} + \epsilon_{t} \ SARMA\$(P,0,Q) \times (P_{s},0,Q_{s},s)$$

res = SARIMAX(housing_yoy, order=(2, 0, 0), seasonal_order=(0, 0, 1, 12)).fit()
summary(res)

	coef	std err	Z	P> z	[0.025	0.975]
ar.L1	0.6809	0.034	20.284	0.000	0.615	0.747
ar.L2	0.2824	0.034	8.233	0.000	0.215	0.350
ma.S.L12	-0.8795	0.017	-50.520	0.000	-0.914	-0.845
sigma2	0.0083	0.000	21.791	0.000	0.008	0.009

SARMA Residual Diagnostics

```
In [11]: acf_pacf_plot(res.resid.iloc[13:], 24)
```



Extended Dynamics $(1,0,1) \times (1,0,1)_{12}$

• In lag polynomial representation

$$(1-\phi_1 L)(1-\phi_s L^s)Y_t = \phi_0 + (1+\theta_1 L)(1+\theta_s L^s)\epsilon_t$$

In standard ARMA representation

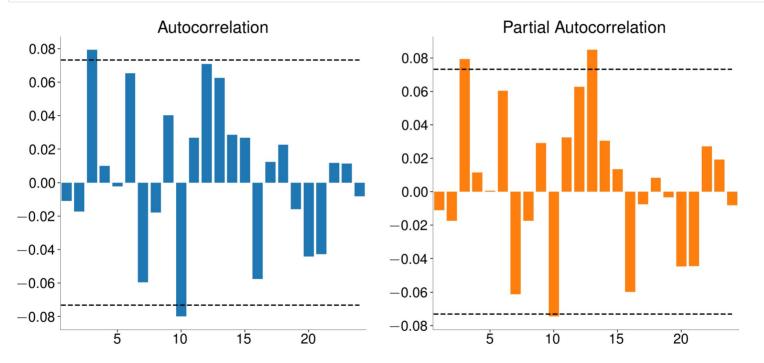
$$Y_{t} = \phi_{0} + \phi_{1}Y_{t-1} + \phi_{s}Y_{t-s} - \phi_{1}\phi_{s}Y_{t-s-1} + \theta_{1}\epsilon_{t-1} + \theta_{s}\epsilon_{t-s} + \theta_{1}\theta_{s}\epsilon_{t-s-1} + \epsilon_{t}$$

In [12]: res = SARIMAX(housing raw, order=(1, 0, 1), seasonal order=(1, 0, 1, 12)).fit() summary(res)

	coef	std err	Z	P> z	[0.025	0.975]
ar.L1	0.9994	0.001	1441.218	0.000	0.998	1.001
ma.L1	-0.3196	0.031	-10.367	0.000	-0.380	-0.259
ar.S.L12	0.9987	0.001	1469.955	0.000	0.997	1.000
ma.S.L12	-0.8957	0.016	-54.402	0.000	-0.928	-0.863
sigma2	0.0083	0.000	22.277	0.000	0.008	0.009

Extended Model Residual Diagnostics

```
In [13]: acf_pacf_plot(res.resid.iloc[13:], 24)
```



Incorporating differencing

• SARIMA - Seasonal Autoregressive Integrated Moving Average

$$ext{SARIMA}(P,D,Q) imes(P_s,D_s,Q_s,s) \ \Phi(L)\Phi_s(L)(1-L)^D(1-L^s)^{D_s}Y_t=\phi_0+\Theta(L)\Theta_s(L)\epsilon_t$$

• AR

$$\Phi(L)=\$1-\phi_1L-\phi_2L^2-\ldots-\phi_PL^P$$

Seasonal AR:

$$\Phi_s(L)=1-\phi_sL^s-\phi_{2s}L^{2s}-\ldots-\phi_{P_ss}L^{P_ss}$$

• MA:

$$\Theta(L) = 1 + heta_1 L + heta_2 L^2 + \ldots + \phi_Q L^Q$$

• Seasonal MA:

$$\Theta_s(L) = 1 + heta_s L^s + heta_{2s} L^{2s} + \ldots + heta_{Q_s s} L^{Q_s s}$$

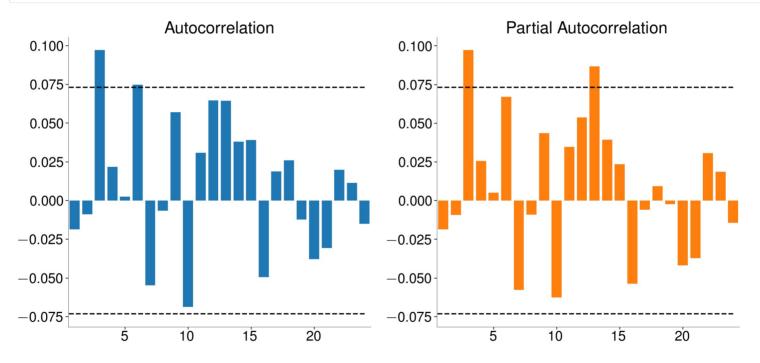
Estimating SARIMA Models

In [14]: res = SARIMAX(housing_raw, order=(1, 0, 1), seasonal_order=(0, 1, 1, 12)).fit() summary(res)

	coef	std err	Z	P> z	[0.025	0.975]
ar.L1	0.9777	0.008	126.932	0.000	0.963	0.993
ma.L1	-0.3128	0.033	-9.352	0.000	-0.378	-0.247
ma.S.L12	-0.8775	0.018	-48.061	0.000	-0.913	-0.842
sigma2	0.0083	0.000	21.771	0.000	0.008	0.009

$\mathsf{SARIMA}(1,0,1) imes (0,1,1)_{12}$ Residual Analysis

In [15]: acf_pacf_plot(res.resid.iloc[13:], 24)



Enforcing Unit Roots

- Set both differences to 1
 - $\blacksquare D = 1$
 - $D_s = 1$
- Model is

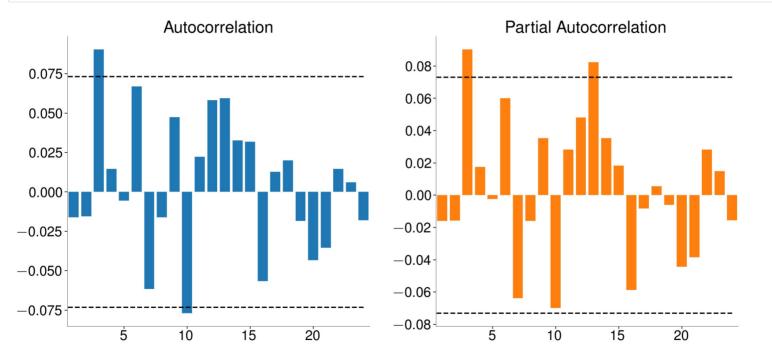
$$(1-L)(1-L^s)Y_t = heta_s\epsilon_{t-s} + \epsilon_t \ \Delta\Delta_s Y_t = heta_s\epsilon_{t-s} + \epsilon_t$$

In [16]: res = SARIMAX(housing_raw, order=(0, 1, 1), seasonal_order=(0, 1, 1, 12)).fit()
 summary(res)

	coef	std err	Z	P> z	[0.025	0.975]
ma.L1	-0.3246	0.031	-10.362	0.000	-0.386	-0.263
ma.S.L12	-0.8774	0.019	-47.035	0.000	-0.914	-0.841
sigma2	0.0084	0.000	22.025	0.000	0.008	0.009

Double Difference Residual Diagnostics

```
In [17]: acf_pacf_plot(res.resid.iloc[13:], 24)
```



Seasonal Dummies

- Alternative approach to SARIMA
- Model shifts as a deterministic processes
- ullet One dummy for each of s-1 periods
 - Avoid dummy variable trap

Out[18]:

	Feb	Mar	Apr	May	Jun	Jul	Aug	Sep	Oct	Nov	Dec
1960-01-01	0	0	0	0	0	0	0	0	0	0	0
1960-02-01	1	0	0	0	0	0	0	0	0	0	0
1960-03-01	0	1	0	0	0	0	0	0	0	0	0
1960-04-01	0	0	1	0	0	0	0	0	0	0	0
1960-05-01	0	0	0	1	0	0	0	0	0	0	0

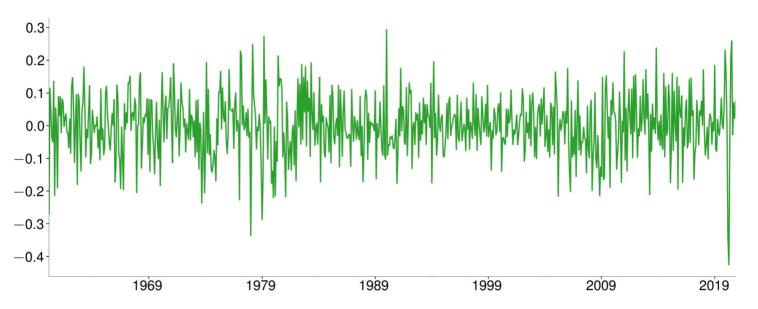
Incorporating Dummies

In [19]: res = SARIMAX(housing_raw, order=(2, 1, 0), exog=dummies, trend="c").fit() summary(res)

	coef	std err	z	P> z	[0.025	0.975]
intercept	0.0002	0.004	0.048	0.962	-0.007	0.008
Feb	0.0358	0.012	2.965	0.003	0.012	0.059
Mar	0.3075	0.012	24.777	0.000	0.283	0.332
Apr	0.4289	0.015	29.516	0.000	0.400	0.457
May	0.4669	0.018	26.260	0.000	0.432	0.502
Jun	0.4697	0.019	25.308	0.000	0.433	0.506
Jul	0.4328	0.019	23.265	0.000	0.396	0.469
Aug	0.4117	0.019	22.227	0.000	0.375	0.448
Sep	0.3657	0.017	21.802	0.000	0.333	0.399
Oct	0.3921	0.015	26.252	0.000	0.363	0.421
Nov	0.2169	0.013	16.942	0.000	0.192	0.242
Dec	0.0502	0.010	5.240	0.000	0.031	0.069
ar.L1	-0.2675	0.033	-8.115	0.000	-0.332	-0.203
ar.L2	-0.1107	0.034	-3.276	0.001	-0.177	-0.044
sigma2	0.0087	0.000	21.344	0.000	0.008	0.010

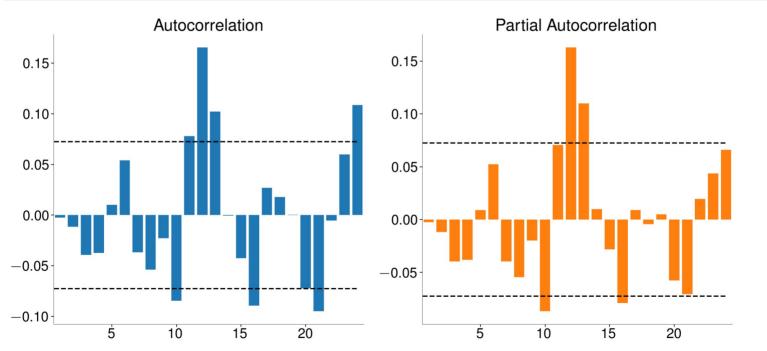
Seasonal Dummy Model Residuals

```
In [20]: resid = res.resid.iloc[2:]
   plot(resid, y=-2)
```



Seasonal Dummy Model Residuals Diagnostics

```
In [21]: acf_pacf_plot(resid, 24)
```



Unit Root Testing

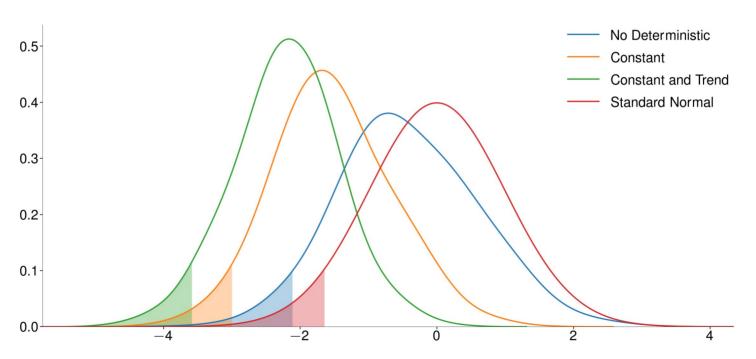
- Test Y_t for a unit root
- If null is not rejected, ensure that all required deterministic terms are included
- ullet If null is still *not* rejected, difference and test ΔY_t
- ullet Positive Y_t series should usually be transformed with \ln
- ullet Select lag length of ΔY_{t-j} terms using IC, usually AIC
- $H_0: \gamma = 0, H_1: \gamma < 0$

$$Y_t = \delta_0 + \delta_1 t + \gamma Y_{t-1} + \sum_{i=1}^P \Delta Y_{t-i} + \epsilon_t$$

ullet Deterministic terms are needed to remove deterministic components from Y_{t-1}

The Dickey-Fuller distributions

```
In [23]: adf_cv_plot()
```



Testing Default Premium

```
In [24]: from arch.unitroot import ADF
adf = ADF(default)
adf.summary()

Out[24]: Augmented Dickey-
Fuller Results
```

 Test Statistic
 -3.866

 P-value
 0.002

 Lags
 10

Trend: Constant

Critical Values: -3.44 (1%), -2.87 (5%), -2.57 (10%) Null Hypothesis: The process contains a unit root.

Testing Curvature

```
In [25]: adf = ADF(curve)
   adf.summary()
```

Out [25]: Augmented Dickey-Fuller Results

Test Statistic -4.412
P-value 0.000
Lags 19

Trend: Constant

Critical Values: -3.44 (1%), -2.87 (5%), -2.57 (10%) Null Hypothesis: The process contains a unit root.

Testing Industrial Production

Trend: Constant

Critical Values: -3.44 (1%), -2.87 (5%), -2.57 (10%) Null Hypothesis: The process contains a unit root.

ADF Regression Results

In [27]: | summary(adf.regression)

	coef	std err	t	P> t	[0.025	0.975]
Level.L1	-0.0017	0.001	-2.186	0.029	-0.003	-0.000
Diff.L1	0.3637	0.037	9.860	0.000	0.291	0.436
Diff.L2	-0.1110	0.039	-2.832	0.005	-0.188	-0.034
Diff.L3	0.0447	0.039	1.138	0.255	-0.032	0.122
Diff.L4	0.0217	0.037	0.589	0.556	-0.051	0.094
const	0.0083	0.003	2.579	0.010	0.002	0.015

Increasing the deterministic order

```
In [28]: adf = ADF(np.log(orig.INDPRO), trend="ct")
   adf
```

Out [28]: Augmented Dickey-Fuller Results

Test Statistic -1.831
P-value 0.690
Lags 6

Trend: Constant and Linear Time Trend

Critical Values: -3.97 (1%), -3.42 (5%), -3.13 (10%) Null Hypothesis: The process contains a unit root.

Testing the difference

ullet The Industrial Productivity Index is I(1) since 1 difference makes it stationary

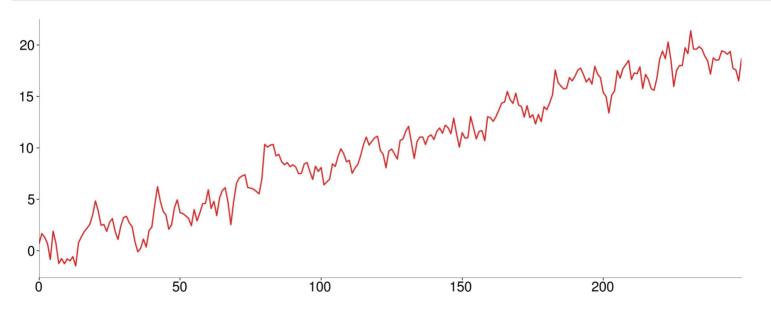
Trend: Constant

Critical Values: -3.44 (1%), -2.87 (5%), -2.57 (10%) Null Hypothesis: The process contains a unit root.

The importance of time trends

- Incorrect trend specification results in no power to reject false nulls
- Always find a unit root in trending time series

```
In [31]: plot(trending, 13)
```



Under-specifying Deterministic Terms

ullet Do not use model with no deterministic terms on real data unless Y_t is known to be mean 0

```
In [32]: ADF(trending, trend="n")

Out[32]: Augmented Dickey-
Fuller Results

Test Statistic 1.934
P-value 0.988
Lags 9
```

Trend: No Trend

Critical Values: -2.58 (1%), -1.94 (5%), -1.62 (10%) Null Hypothesis: The process contains a unit root.

Under-specifying Deterministic Terms

Trend: Constant

Critical Values: -3.46 (1%), -2.87 (5%), -2.57 (10%) Null Hypothesis: The process contains a unit root.

Correctly specifying Deterministic Terms

Trend: Constant and Linear Time Trend

Critical Values: -4.00 (1%), -3.43 (5%), -3.14 (10%) Null Hypothesis: The process contains a unit root.

Over-specifying Deterministic Terms

```
In [35]: ADF(y, trend="ctt")

Out[35]: Augmented Dickey-
Fuller Results

Test Statistic -6.885

P-value 0.000

Lags 0
```

Trend: Constant, Linear and Quadratic Time Trends Critical Values: -4.42 (1%), -3.86 (5%), -3.57 (10%) Null Hypothesis: The process contains a unit root.

Incorporating Time Trends into ARMA Models

• Can jointly estimate trends and stationary components of ARMA models

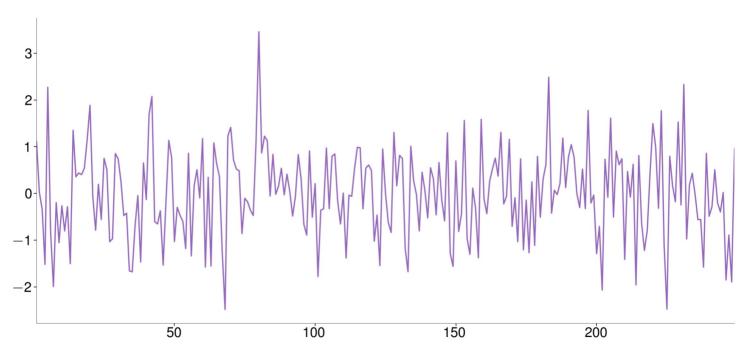
$$Y_t = \phi_0 + \delta_1 t + \phi_1 Y_{t-1} + \epsilon_t$$

```
In [36]: res = SARIMAX(trending, order=(1, 0, 0), trend="ct").fit()
         summary(res)
```

	coef	std err	Z	P> z	[0.025	0.975]
intercept	0.0796	0.116	0.687	0.492	-0.148	0.307
drift	0.0255	0.004	6.149	0.000	0.017	0.034
ar.L1	0.6822	0.050	13.534	0.000	0.583	0.781
sigma2	0.9213	0.080	11.451	0.000	0.764	1.079

Time Trend Model Residuals

```
In [37]: plot(res.resid.iloc[1:])
```



Next Week

• Univariate Volatility Modeling