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What is Multi-Armed Bandit?

Consider a sequential decision problem.

Suppose..

- A learner is sequentially faced with N actions.
- At each time, the learner can choose only one action.
- The chosen action yields a reward.
- The learner repeats the process and accumulates the rewards.

The goal of the learner is to maximize the sum of rewards.

What is Multi-Armed Bandit?

Multi-Armed Bandits (MAB) [Robbins, 1952, Lai and Robbins, 1985] frame the sequential decision problem.

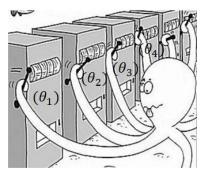


image source: Microsoft Research

- Arms=Actions (# of arms: N)
- At time t, the i-th action yields a random reward r_i(t), such that

$$\mathbb{E}(r_i(t)) = \theta_i(t), \quad i = 1, \cdots, N,$$

where $\theta_i(t)$'s are unknown.

- At time t, the learner chooses one action a(t), and observes the reward $r_{a(t)}(t)$.
- Goal is to maximize $\sum_{t=1}^{T} \theta_{a(t)}(t)$.



Yahoo! front page snapshot

- At each user visit, the web system selects one article from a large pool of articles.
- The system displays it on the Featured tab.
- The user clicks the article if he/she is interested in the contents.
- Based on user click feedback, the system updates its article selection strategy.



Yahoo! front page snapshot

- At each user visit, the web system selects one article from a large pool of articles. Arms=Articles
- The system displays it on the Featured tab.
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Yahoo! front page snapshot

- At each user visit, the web system selects one article from a large pool of articles. Arms=Articles
- The system displays it on the Featured tab. a(t): index of chosen article
- The user clicks the article if he/she is interested in the contents.
- Based on user click feedback, the system updates its article selection strategy.



Yahoo! front page snapshot

- At each user visit, the web system selects one article from a large pool of articles. Arms=Articles
- The system displays it on the Featured tab. a(t): index of chosen article
- The user clicks the article if he/she is interested in the contents.
 r_{a(t)}(t) = 1 or 0.
- Based on user click feedback, the system updates its article selection strategy.



Example of mHealth app

- At specific times in a day, the mHealth system selects one type of message among various types of messages.
- The system pushes the message to the app user.
- The user reacts to the message.
- Based on user reaction, the system updates its message selection strategy.



Example of mHealth app

- At specific times in a day, the mHealth system selects one type of message among various types of messages. Arms=Messages
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Example of mHealth app

- At specific times in a day, the mHealth system selects one type of message among various types of messages. Arms=Messages
- The system pushes the message to the app user. a(t): index of chosen message
- The user reacts to the message. $r_{a(t)}(t) = \text{step counts}.$
- Based on user reaction, the system updates its message selection strategy.

Other Applications

Applications include..

- mobile healthcare system [Tewari and Murphy, 2017]
- news article recommendation algorithms [Li et al., 2010]
- web page ad placement algorithms [Langford et al., 2008]
- revenue management [Ferreira et al., 2017]
- marketing [Schwartz et al., 2017]
- recommendation systems [Kawale et al., 2015]

Multi-armed bandit (MAB)

• Let $a^*(t) = \underset{1 \le i \le N}{\operatorname{argmax}} \ \theta_i(t)$ and $regret(t) := \theta_{a^*(t)}(t) - \theta_{a(t)}(t)$. Goal is to minimize the sum of regrets,

$$R(T) := \sum_{t=1}^{T} regret(t) = \sum_{t=1}^{T} \{\theta_{a^*(t)}(t) - \theta_{a(t)}(t)\}.$$

- Since $\theta_i(t)$'s are unknown, the learner has to **learn** which actions yield maximum mean reward by trying them out.
- Only $r_{a(t)}(t)$ is observed among the whole reward vector $r(t) = [r_1(t), \dots, r_N(t)]^T$.
 - ⇒ Trade-off between **exploitation** and **exploration**.

Contextual MAB

- Contextual MABs assume there is context vector $b_i(t) \in \mathbb{R}^d$ associated with each arm i at time t.
- Reward $r_i(t)$ is assumed to depend on $b_i(t)$:

$$\mathbb{E}(r_i(t)|b_i(t)) = \theta(b_i(t)), \quad i = 1, \dots, N.$$

• $\mathbb{E}(r_i(t)|b_i(t))$ is linear in $b_i(t)$, i.e.,

$$\theta(b_i(t)) = b_i(t)^T \mu, \quad i = 1, \dots, N,$$

where $\mu \in \mathbb{R}^d$ is unknown.

• $\mathbb{E}(r_i(t)|b_i(t))$ is linear in $b_i(t)$, i.e.,

$$\theta(b_i(t)) = b_i(t)^T \mu, \quad i = 1, \dots, N,$$

where $\mu \in \mathbb{R}^d$ is unknown.

• Error $\eta_i(t):=r_i(t)-\mathbb{E}(r_i(t)|b_i(t))$ is R-sub-Gaussian, i.e., for every $\lambda\in\mathbb{R}$,

$$\mathbb{E}\big[\exp(\lambda\eta_i(t))\big] \leq \exp\big(\frac{\lambda^2 R^2}{2}\big).$$

• WLOG, $||b_i(t)|| \le 1$, $||\mu|| \le 1$.

Remarks

• $a^*(t) = \underset{1 \le i \le N}{\operatorname{argmax}} \{b_i(t)^T \mu\}$ and,

$$regret(t) = b_{a^*(t)}(t)^T \mu - b_{a(t)}(t)^T \mu.$$

Lower Bounds

- When N is infinite, [Dani et al., 2008] proved that no algorithm can achieve lower bound than $O(d\sqrt{T})$.
- When N is finite, [Chu et al., 2011] proved a lower bound of $O(\sqrt{dT})$ when $d^2 < T$.

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But, what if $d \gg T$?

In modern applications, contexts are often high-dimensional with only a sparse subset correlated with the reward.

We consider,

• High-dimensional reward model:

$$\mathbb{E}(r_i(t)|b_i(t)) = b_i(t)^T \mu, \quad i = 1, \cdots, N,$$

where $\mu \in \mathbb{R}^d$ and $||\mu||_0 = s_0 \ll d$.

• In low dimension,

$$(\mathsf{Gram\ matrix}) = \sum_{\tau=1}^t b_{\mathsf{a}(\tau)}(\tau) b_{\mathsf{a}(\tau)}(\tau)^T$$

is positive definite after O(d) rounds.

• But if $d \gg T$, Gram matrix is singular until the end.

A weaker condition is ..

• Compatibility Condition [van de Geer and Bühlmann, 2009]. Define $\hat{\Sigma}_t := \frac{1}{t} \sum_{\tau=1}^t b_{a(\tau)}(\tau) b_{a(\tau)}(\tau)^T$ and let $I = supp(\mu)$, the set of indices of non-zero components of μ . Then $\exists \phi_1 > 0$ such that

for
$$\forall v \in \mathbb{R}^d$$
 such that $||v_{I^c}||_1 \leq 3||v_I||_1$, $||v_I||_1^2 \leq \frac{|I|(v^T \hat{\Sigma}_t v)}{\phi_1^2}$.

Lemma (Lemma 11.2 of [van de Geer and Bühlmann, 2009])

Let $x_{\tau} \in \mathbb{R}^d$ and $y_{\tau} \in \mathbb{R}$ be random variables with $y_{\tau} = x_{\tau}^T \beta + \varepsilon_{\tau}, \ \tau = 1, 2, \cdots, t$, where $\beta \in \mathbb{R}^d$, $||\beta||_0 = s_0$, and ε_{τ} 's are i.i.d.

gaussian with mean zero and variance R^2 . Let $\lambda_t = R\sqrt{\frac{2\log(\mathrm{e}d/\delta)}{t}}$ and

$$\hat{\beta}(t) = \underset{\beta}{\operatorname{argmin}} \left\{ \frac{1}{t} \sum_{\tau=1}^{t} (y_{\tau} - x_{\tau}^{T} \beta)^{2} + \lambda_{t} ||\beta||_{1} \right\}.$$

If $\hat{\Sigma}_t := \frac{1}{t} \sum_{\tau=1}^t x_\tau x_\tau^T$ satisfies compatibility condition for some $\phi > 0$, then for $\forall \delta \in (0,1)$, with probability at least $(1-\delta/t^2)$,

$$||\hat{\beta}(t) - \beta||_1 \leq \frac{4\lambda_t s_0}{\phi^2} = \frac{4s_0 R}{\phi^2} \sqrt{\frac{4\log(edt/\delta)}{t}}.$$

Lemma (Lemma EC.6 of [Bastani and Bayati, 2015])

Let x_1, x_2, \dots, x_t be i.i.d. random vectors in \mathbb{R}^d with $||x_\tau||_\infty \le 1$ for all τ . Let $\Sigma = \mathbb{E}[x_\tau x_\tau^T]$ and $\hat{\Sigma}_t = \frac{1}{t} \sum_{\tau=1}^t x_\tau x_\tau^T$. Suppose Σ satisfies compatibility for some $\phi > 0$. Then if $c = \min(0.5, \frac{\phi^2}{256s_0})$ and $t \ge \frac{3}{c^2} \log d$, with probability at least $1 - \exp(-c^2t)$, $\hat{\Sigma}_t$ satisfies compatibility as well for $\phi/\sqrt{2}$.

Lemma (Lemma EC.6 of [Bastani and Bayati, 2015])

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 \Rightarrow But $b_{a(t)}(t)$'s are not i.i.d. !

Recall...

1 (1)	1 (0)	1 (2)	1 (1)
$b_1(1)$	$b_1(2)$	$b_1(3)$	 $b_1(t)$
$b_2(1)$	$b_2(2)$	$b_2(3)$	 $b_2(t)$
$b_3(1)$	$b_3(2)$	$b_3(3)$	 $b_3(t)$
$b_4(1)$	$b_4(2)$	$b_4(3)$	 $b_4(t)$
$b_5(1)$	$b_5(2)$	$b_5(3)$	 $b_5(t)$
$b_6(1)$	$b_6(2)$	$b_6(3)$	 $b_6(t)$
$b_7(1)$	$b_7(2)$	$b_7(3)$	 $b_7(t)$
$b_8(1)$	$b_8(2)$	$b_8(3)$	 $b_8(t)$

Recall...

```
b_1(1)
           b_1(2)
                      b_1(3)
     b_2(2)
                    b_2(3)
                                             b_2(t)
b_3(1) r_3(1) b_3(2) b_3(3)
                                             b_3(t)
    b_4(2)
                    b_4(3)
                                             b_4(t)
b_5(1)
    b_5(2) b_5(3)
                                             b_5(t)
    b_6(2) b_6(3)
                                             b_7(t)
b_7(1)
    b_7(2) b_7(3)
           b_8(2)
                      b_8(3)
```

$$a(1)=3$$

Recall...

```
b_1(1)
           b_1(2)
                       b_1(3)
      b_2(2) b_2(3)
                                               b_2(t)
b_3(1) r_3(1) b_3(2) b_3(3)
                                               b_3(t)
    b_4(2)
                     b_4(3)
                                               b_4(t)
    b_5(2)
                     b_5(3)
                                               b_5(t)
    b_7(2) r_7(2) b_7(3)
                                               b_7(t)
b_7(1)
```

$$a(1)=3$$
 $a(2)=7$

Recall...

$$a(1) = 3$$

$$a(2)=7$$

$$a(3)=2$$

Recall...

$$a(1) = 3$$

$$a(2) = 7$$

$$a(3)=2$$

$$a(t)=4$$

- We can assume $\{\bar{b}(t) = \frac{1}{N} \sum_{i=1}^{N} b_i(t)\}$ is i.i.d.
- Bandit data is missing data!

a(2)=7

a(1)=3

$b_1(1)$ $b_2(1)$ $b_3(1)$ $r_3(1)$	$b_1(2)$ $b_2(2)$ $b_3(2)$	$b_1(3)$ $b_2(3)$ $r_2(3)$ $b_3(3)$		$b_1(t) \\ b_2(t) \\ b_3(t)$
$b_4(1)$ $b_5(1)$	$b_4(2) \\ b_5(2)$	$b_4(3)$ $b_5(3)$		$b_4(t) r_4(t)$ $b_5(t)$
$b_6(1)$ $b_7(1)$ $b_8(1)$	$b_6(2)$ $b_7(2)$ $r_7(2)$ $b_8(2)$	$b_6(3)$ $b_7(3)$ $b_8(3)$	•••	$b_6(t) \\ b_7(t) \\ b_8(t)$

a(3)=2

a(t)=4

• Let \mathcal{H}_{t-1} be history until time t-1,

$$\mathcal{H}_{t-1} = \{a(\tau), r_{a(\tau)}(\tau), \{b_i(\tau)\}_{i=1}^N, \tau = 1, \cdots, t-1\}.$$

- Define filtration $\mathcal{F}_{t-1} = \{\mathcal{H}_{t-1}, \{b_i(t)\}_{i=1}^N\}$.
- Let $\pi_i(t)$ be action selection probability,

$$\pi_i(t) = \mathbb{P}(a(t) = i | \mathcal{F}_{t-1}).$$

Remark $\pi_i(t)$ is observation probability. In bandits, $\pi_i(t)$ is controlled by the learner, i.e., the value is known.

We construct a doubly-robust pseudo reward.

$$egin{aligned} \hat{r}(t) &= rac{1}{N} \sum_{i=1}^{N} \hat{r}_i(t) \ &= rac{1}{N} \sum_{i=1}^{N} \Big\{ rac{I(a(t) = i)}{\pi_i(t)} r_i(t) + \Big(1 - rac{I(a(t) = i)}{\pi_i(t)}\Big) b_i(t)^T \hat{\mu}(t-1) \Big\}, \end{aligned}$$

where $\hat{\mu}(t-1)$ is the Lasso estimate of μ obtained from last step. Whether or not $\hat{\mu}(t-1)$ is a valid estimate, this value has conditional expectation $\bar{b}(t)^T \mu$ given that $\pi_i(t) > 0$ for all i:

$$\mathbb{E}[\hat{r}(t)|\mathcal{F}_{t-1}] = \bar{b}(t)^T \mu.$$

 \Rightarrow Apply Lasso on $\{(\bar{b}(\tau), \hat{r}(\tau))\}_{\tau=1}^t$ instead of $\{(b_{a(\tau)}(\tau), r_{a(\tau)}(\tau))\}_{\tau=1}^t$.

Note that

$$\hat{r}(t) = \frac{1}{N} \sum_{i=1}^{N} \left\{ b_i(t)^T \hat{\mu}(t-1) + \frac{I(a(t)=i)}{\pi_i(t)} \left\{ r_i(t) - b_i(t)^T \hat{\mu}(t-1) \right\} \right\}
= \frac{1}{N} \sum_{i=1}^{N} \left\{ b_i(t)^T \hat{\mu}(t-1) + \frac{I(a(t)=i)}{\pi_i(t)} \left\{ \eta_i(t) + b_i(t)^T \left(\mu - \hat{\mu}(t-1) \right) \right\} \right\}$$

If we have $||\hat{\mu}(t-1) - \mu||_1 \leq O\left(\sqrt{\frac{\log t}{t}}\right)$, and if we set $\pi_i(t) \geq O\left(\frac{1}{N}\sqrt{\frac{\log t}{t}}\right)$, the variance of $\hat{r}(t)$ is constant scale!

To ensure $\pi_i(t) \geq O\left(\frac{1}{N}\sqrt{\frac{\log t}{t}}\right)$, we do:

Generate
$$m_t \sim Ber\Big(O\big(\sqrt{\frac{\log t}{t}}\big)\Big).$$

- if $m_t = 1$ then
 - Pull arm a(t) = i with probability $\frac{1}{N}$ $(i = 1, \dots, N)$
- else

Pull arm
$$a(t) = \underset{1 \le i \le N}{\operatorname{argmax}} \{b_i(t)^T \hat{\mu}(t-1)\}.$$
 (2)

(1)

Remark: Step (1) is exploration, step (2) is exploitation.

- Step (1) induces suboptimal choice of arms.
- Let R(T,1) be sum of regrets due to step (1). Then $R(T,1) \leq \sum_{t=1}^{T} m_t$. Due to Hoeffding's inequality, with probability at least 1δ ,

$$R(T,1) \leq \sum_{t=1}^T m_t \leq \sum_{t=1}^T \mathbb{E}(m_t) + \sqrt{T \log(1/\delta)/2}.$$

where

$$\begin{split} \sum_{t=1}^T \mathbb{E}(m_t) &= O\Big(\sum_{t=1}^T \sqrt{\frac{\log t}{t}}\Big) \leq O\Big(T\frac{1}{T}\sum_{t=1}^T \sqrt{\frac{\log T}{t}}\Big) \\ &\leq O\Big(T\sqrt{\frac{1}{T}\sum_{t=1}^T \frac{\log T}{t}}\Big) \quad (\because \text{ Jensen's inequality}) \\ &= O(\sqrt{T}\log T). \end{split}$$

• In Step (2),

$$\begin{aligned} b_{a^*(t)}(t)^T \mu &\leq b_{a^*(t)}(t)^T \hat{\mu}(t-1) + ||\hat{\mu}(t-1) - \mu||_1 \\ &\leq b_{a(t)}(t)^T \hat{\mu}(t-1) + ||\hat{\mu}(t-1) - \mu||_1 \\ &\leq b_{a(t)}(t)^T \mu + ||\hat{\mu}(t-1) - \mu||_1 + ||\hat{\mu}(t-1) - \mu||_1 \\ \Rightarrow \textit{regret}(t) &\leq 2||\hat{\mu}(t-1) - \mu||_1 \end{aligned}$$

• Let R(T,2) be sum of regrets from step (2). With probability at least $1-\delta$,

$$R(T,2) \le 2 \sum_{t=1}^{T} ||\hat{\mu}(t-1) - \mu||_1$$
$$\le 2 \sum_{t=1}^{T} \frac{4s_0 R}{\phi^2} \sqrt{\frac{4\log(edt/\delta)}{t}}$$
$$\le O\left(s_0 \sqrt{T}\log(dT)\right)$$

Theorem

For $\forall \delta \in (0,1)$, with probability at least $1-\delta$,

$$R(T) \leq O(s_0 \sqrt{T} \log(dT)).$$

Related Works

- [Abbasi-Yadkori et al., 2012]: given any online prediction algorithm for the regression parameter, their algorithm constructs a high-probability confidence set for the true parameter using a bound on the prediction loss, and then pulls arms according to the *optimism in the face of uncertainty* rule. Their high-probability regret bound is proportional to \sqrt{d} instead of $\log d$, so is not sublinear in T when d scales with T.
- [Carpentier and Munos, 2012]: used an explicit exploration phase to identify the support of the regression parameter using techniques from compressed sensing. Their regret bound is tight scaling with $\log d$, but the algorithm is specific to the case where the set of arms is the unit ball for the $||\cdot||_2$ norm and fixed over time.
- [Gilton and Willett, 2017]: leveraged ideas from linear Thompson Sampling and Relevance Vector Machines [Tipping, 2001]. The theoretical results are weak since they derived the regret bound under the assumption that a sufficiently small superset of the support for the regression parameter is known in advance.

Related Works

• [Bastani and Bayati, 2015]: proposed the Lasso Bandit under a different reward model,

$$E(r_i(t)|b_i(t)) = b(t)^T \beta_i$$
 (1)

with $||\beta_i|| = s_0$, $i = 1, \dots, N$. They imposed compatibility through forced-sampling of each arm. The upper bound of the *expected* regret is proportional to number of arms, N:

$$\mathbb{E}[R(T)] \le O(Ns_0^2[\log T + \log d]^2).$$

An application of the Hoeffding's inequality gives an additional term of order $O(\sqrt{T})$ for the high-probability bound.

• [Wang et al., 2018]: proposed the Minimax Concave Penalized (MCP) Bandit algorithm for the reward model (1), which uses forced-sampling along with the MCP estimator [Zhang, 2010]. They obtained,

$$\mathbb{E}[R(T)] \leq O(Ns_0^2[s_0 + \log d]\log T).$$

Experiments

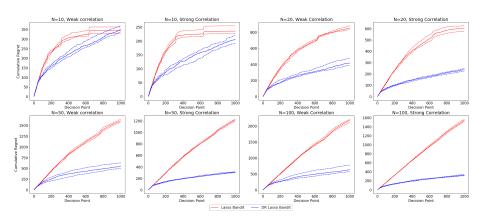
We compare the Doubly-Robust Lasso Bandit with Lasso Bandit [Bastani and Bayati, 2015].

- Number of arms: N = 50 or 100.
- Dimension of context vector: d = 100.
- Distribution of context vector: for fixed $j=1,\cdots,d$, $[b_{1j}(t),\cdots,b_{Nj}(t)]^T\sim \mathcal{N}(0_N,V)$, where V(i,i)=1 for every i and $V(i,k)=\rho^2$ for every $i\neq k$. $\rho^2=0.3$ (weak correlation) or $\rho^2=0.7$ (strong correlation).
- Regression parameter: $s_0 = 5$ and $\beta_{supp(\beta)} \sim U([0,1])^5$.
- Distribution of the reward:

$$r_i(t) = b_i(t)^T \beta + \eta_i(t),$$

where $\eta_i(t) \stackrel{i.i.d.}{\sim} \mathcal{N}(0, 0.05^2)$.

Experiments



Thank You!