Sample answer Lab04: Matrix Operations & Model Diagnostics

**Handed out:** Thursday, October 15, 2020

**Return date:** Friday, October 30, 2020 by midnight into the link **Lab04Submit** in eLearning.

**Grading:** This lab counts 8 % towards your final grade

**Objectives:** This lab practices in part 1 operations with matrices which are relevant to regression analysis and explore properties of different factor coding schemes. Part 2 focuses on model building and standard model diagnostics.

## Part 1: Matrix Operations (4 points)

## Task 1: Manual matrix operations and regression analysis with matrices [1 points]

You are given a vector of the dependent variable and the design matrix

[a] Enter and into . Write your own OLS  function using the dependent vector and the associated design matrix as input. Your function should return the vector of the estimated regression coefficients. (0.5 point)

X <- matrix(c(1,1,1,1,1,1,2,4,2,4,2,6), nrow=6)

y<- matrix(c(2,5,2,5,2,9), nrow=6)

solve(crossprod(X), crossprod(X,y))

[,1]

[1,] -1.5

[2,] 1.7

[b] Use 's matrix operations to calculate for a dependent variable , the design matrix and the diagonal weights matrix the weighted regression coefficients with the formula . (0.5 points)

X <- matrix(c(1,1,1,2,4,6), nrow=3)

y<- matrix(c(2,5,9), nrow=3)

W <- diag(c(3,2,1))

solve(t(X)%\*%W%\*%X, t(X)%\*%W%\*%y)

[,1]

[1,] -1.5

[2,] 1.7

[c] Compare the estimated regression coefficients from task 1 [a] with those from task 1 [b]. Explain why they are identically. Hint: what is the effect of the weights matrix . (0.5 points)

There are only three different sets of observations:

1. = 2 with with 3 observations
2. = 4 with with 2 observations
3. = 6 with with 1 observation

Comment: If we compare the original independent and dependent matrices in question (a) with those of question (c), we see that matrices in question (c) are suppressing duplicate observations and just show the unique observations. The weights matrix 𝐖 provides information how frequent each unique observation is in the full regression system. The weights matrix in-between and between replicates the observations according to their frequency in the original dataset with 6 observations. Consequently, the same results, that is, of regression coefficients, are obtained by both equations.

### Task 2: Coding schemes of categorical variables (3 points)

Provide the  syntax code of your answers. You can either use the **lm(…)** or your coded ordinary least squares function for this task

[a] Enter the matrix and the design matrices to separate matrix objects into  and show these object in your answer (0.5 points):

and are given in the ***indicator coding*** scheme ( codes it as **contrasts(factor) <- "contr.treatment"**) whereas and are given in the ***centered coding*** scheme ( codes it as **contrasts(factor) <- "contr.sum"** and Hamilton p 99 calls it ***effect*** coding). In and the last category is suppressed, whereas in and the second category is suppressed due to the redundancy among a full set of indicator variables.

(y <- matrix(c(7, 5, 3, 1, 3, 2, 9, 5, 7), ncol = 1))

(X1 <- matrix(c(rep(1, 12), rep(0, 9), rep(1, 3), rep(0, 3)), ncol = 3))

(X2 <- matrix(c(rep(1, 12), rep(0, 12), rep(1, 3)), ncol = 3))

(X3 <- matrix(c(rep(1, 12), rep(0, 3), rep(-1, 3), rep(0, 3), rep(1, 3),rep(-1, 3)), ncol = 3))

(X4 <- matrix(c(rep(1, 12), rep(-1, 3), rep(0, 6), rep(-1, 3), rep(1, 3)), ncol = 3))

[b] Calculate the three group means of the observations , and as well as the global mean for all observations . (0.5 points)

(mean.group1 <- mean(y[1:3]))

5

(mean.group2 <- mean(y[4:6]))

2

(mean.group3 <- mean(y[7:9]))

7

(mean.global <- mean(y))

4.67

[c] Find the four sets of estimated regression coefficients for the intercept and group coefficients by regressing on the four design matrices , , and with your linear regression function that you have developed in task 1 [a] and enter these estimates into the table below (see columns *Assign Estimated Regression Coefficients*). (0.5 points)

Best <- function (x,y){solve(t(x)%\*%x)%\*%t(x)%\*%y}

(parameter1 <- Best(X1,y))

(parameter2 <- Best(X2,y))

(parameter3 <- Best(X3,y))

(parameter4 <- Best(X4,y))

Hints: (i) in the ***centered*** coding scheme the coefficient for the missing category can be calculated as the ***negative sum*** of the two other estimated parameters, i.e., . (ii) For the cornered coding scheme the values for the ***dashed*** cells cannot be calculated from the regression results.

|  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  |  | Assign Estimated Regression Coefficients | | | | Give Expressions for the Means in Terms of the Estimate Regression Coefficients | | | |
| Model | **Coding** |  |  |  |  |  |  |  |  |
| y~X1 | ***cornered*** | 7 | -2 | -5 | ─ | ─ |  |  |  |
| y~X2 | ***cornered*** | 2 | 3 | ─ | 5 | ─ |  |  |  |
| y~X3 | ***centered*** | 4.67 | 0.33 | -2.67 | 2.33 |  |  |  |  |
| y~X4 | ***centered*** | 4.67 | 0.33 | -2.67 | 2.33 |  |  |  |  |

Notes:

1. In the cornered coding scheme, the intercept is equal to the mean of the suppressed group.
2. For the centered coding scheme, the global mean   is only equal to the intercept term if all groups have the same number of observations.
3. For the centered coding scheme, the negative sum of the regression coefficients is equal to the coefficient of the suppressed group.

[d] For each design matrix the global mean and group means , and can be expressed as a function of the estimated regression coefficients in the columns *Assign Estimated Regression Coefficients*. (0.5 point)  
Find the expressions for the means and write them into columns labels by “*Give Expressions…*” using the parameter symbols, e.g., and note that this is an invalid expression.

[e] Which coding scheme has a more *intuitive interpretation*? Justify your answer. (0.5 points)

Comment: the centered coding scheme is more intuitive to be interpreted. No matter which reference category is suppressed, the estimated regression coefficients remain the same. In other words, the interpretation of the relationships between dependent variables and the factor levels is invariant of the suppressed reference category. In contrast, the regression coefficients in the cornered coding scheme depend on the suppressed reference category. Moreover, in the centered coding scheme the regression coefficients measure the variation in the mean levels of the categories around the global mean.

## Part 2: Model Building and Diagnostics (4 points)

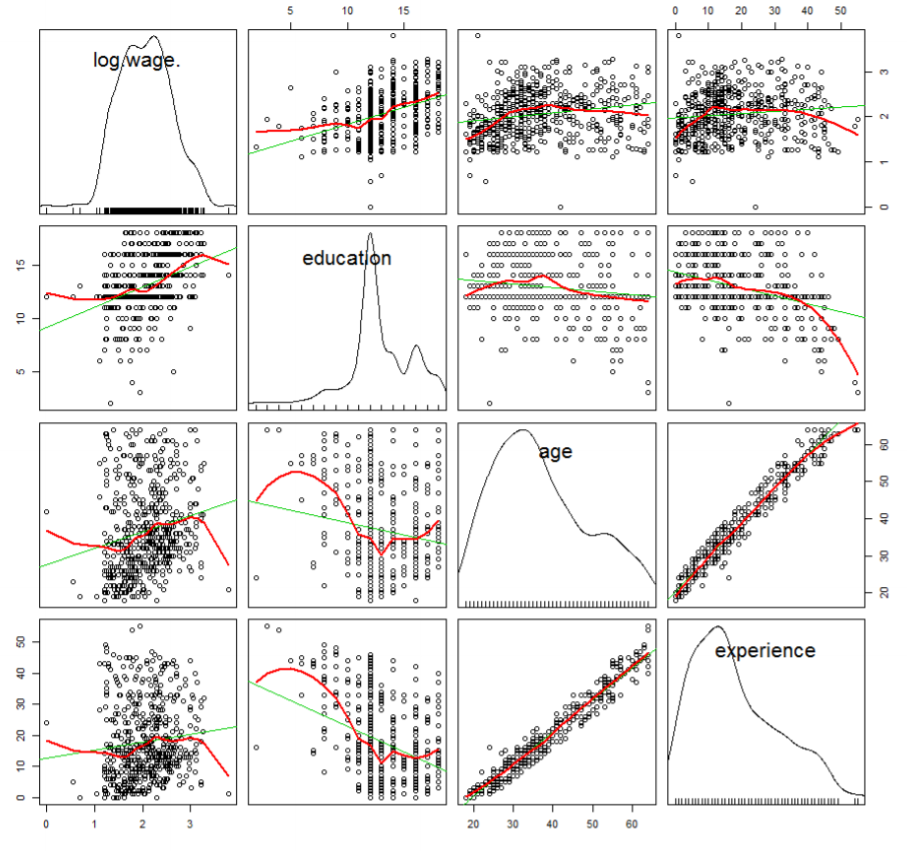
Open the **CPS1985** data-frame with **data("CPS1985",package="AER")**. Assign new row-names with the statement **rownames(CPS1985) <- 1:nrow(CPS1985)** to the data-frame. ***Study the description*** of the variable **experience** in the associated online help.

### Task 3: Multicollinearity diagnostics (2 points)

[a] For the variables **~log(wage)+education+age+experience** generate a scatterplot matrix. (0.5 points)

library(car)

scatterplotMatrix(~log(wage)+education+age+experience, data=CPS1985, spread = FALSE, id.cex=1.5)



Based on the definition of the variables and the scatterplot matrix, which variables do you expect to be multicollinear? Justify your decisions.

Comment: We expect **age, experience** and **education** to be perfectly multicollinear because the definition of experience is **𝑒𝑥𝑝𝑒𝑟𝑖𝑒𝑛𝑐𝑒 = 𝑎𝑔𝑒 − 𝑒𝑑𝑢𝑐𝑎𝑡𝑖𝑜𝑛 − 6**. In other words, **experience** is a linear function of **age** and **education**. Therefore, there the variables measure jointly redundant information. Investigating the bivariate plot of **experience-age** displays a strong linear relationship.

[b] Estimate the model **log(wage)~education+experience** and calculate the ***variance inflation factors***. Fully interpret the estimated model and the ***VIF***s. (0.5 points)

model1 <- lm(log(wage)~education+experience, data=CPS1985)

summary(model1)

Coefficients:

Estimate Std. Error t value Pr(>|t|)

(Intercept) 0.594169 0.124428 4.775 2.33e-06 \*\*\*

education 0.096414 0.008310 11.603 < 2e-16 \*\*\*

experience 0.011774 0.001756 6.707 5.10e-11 \*\*\*

Signif. codes: 0 ‘\*\*\*’ 0.001 ‘\*\*’ 0.01 ‘\*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 0.4695 on 531 degrees of freedom

Multiple R-squared: 0.2115, Adjusted R-squared: 0.2085

F-statistic: 71.21 on 2 and 531 DF, p-value: < 2.2e-16

vif(model1)

education experience

1.142049 1.142049

Comment: Both education and experience are significant and have positive impact on log(wage). Only 21% of the variation within **log(wage)** is explained by **education** and **experience**. The variance inflation factors of **education** and **experience** are smaller than 10. Thus, no multicollinearity exists in this model.

[c] Estimate the augmented model **log(wage)~education+experience+age** and show the output. (1 points)

Address the following points:

1. What do the ***VIF***s tell you?
2. ***What*** happened to the significances of the *t*-tests for the estimated regression parameters of the augmented model and ***why***?
3. Why does the global *F*-test still remain significant?

model2 <- lm(log(wage)~education+experience+age, data=CPS1985)

summary(model2)

Call:

lm(formula = log(wage) ~ education + experience + age, data = CPS1985)

Residuals:

Min 1Q Median 3Q Max

-2.03367 -0.33094 0.04165 0.31958 1.84066

Coefficients:

Estimate Std. Error t value Pr(>|t|)

(Intercept) 0.84480 0.71884 1.175 **0.240**

education 0.13805 0.11791 1.171 **0.242**

experience 0.05353 0.11796 0.454 **0.650**

age -0.04173 0.11786 -0.354 **0.723**

Residual standard error: 0.4699 on 530 degrees of freedom

Multiple R-squared: 0.2117, Adjusted R-squared: 0.2072

F-statistic: 47.44 on 3 and 530 DF, p-value: < 2.2e-16

vif(model2)

education experience age

229.5738 5147.9190 4611.4008

Comment: Due to the substantial degree of collinearity, none of the variables is significant since their standard errors become substantially inflated, but the overall F-statistic remains highly significant. The F -test is significant because jointly the independent variables still influence the dependent variable. The variance inflation factors of **education**, **age**, and **experience** are drastically larger than 10. This example demonstrates that **education**, **age**, and **experience** are highly collinear and at least one of the redundant variables should be dropped.

### Task 4: Refined model specification (1 point)

[a] Estimate the model: **log(wage)~education+experience+gender+occupation+union** and *fully interpret* the estimated regression model. (0.5 point)

model.full <- lm(log(wage)~education+experience+gender+occupation+union, data=CPS1985)

summary(model.full)

Coefficients:

Estimate Std. Error t value Pr(>|t|)

(Intercept) 0.972050 0.132893 7.315 9.74e-13 \*\*\*

education 0.072296 0.009931 7.280 1.23e-12 \*\*\*

experience 0.010775 0.001670 6.454 2.49e-10 \*\*\*

genderfemale -0.203606 0.041860 -4.864 1.52e-06 \*\*\*

occupationtechnical 0.161965 0.069502 2.330 0.02017 \*

occupationservices -0.198521 0.061204 -3.244 0.00126 \*\*

occupationoffice -0.018791 0.063715 -0.295 0.76817

occupationsales -0.150690 0.082108 -1.835 0.06703 .

occupationmanagement 0.209102 0.076316 2.740 0.00635 \*\*

unionyes 0.216589 0.051117 4.237 2.68e-05 \*\*\*

Signif. codes: 0 ‘\*\*\*’ 0.001 ‘\*\*’ 0.01 ‘\*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Residual standard error: 0.4323 on 524 degrees of freedom

Multiple R-squared: 0.3404, Adjusted R-squared: 0.3291

F-statistic: 30.05 on 9 and 524 DF, p-value: < 2.2e-16

Comment: the regression coefficients of all metric variables are significant. People, who have higher education and experience as well as work in a union job, will have higher wages than other people. However, females seem to be discriminated because on average they earn less money. As the **experience** continues to accumulate over the years wages of workers continue to rise.

The occupation factor indicates that technical and management employees earn more than workers (this is the suppressed reference category) whereas service employees earn less. Office workers and salespersons earn approximately the same as workers, who constitute the reference category, since their coefficients are insignificant.

[b] Test whether the factor **occupation** is significant and if necessary refine the model specification accordingly. (0.25 points)

model.test <- lm(log(wage)~education+experience+gender+union, data=CPS1985)

anova(model.test, model.full)

Model 1: log(wage) ~ education + experience + gender + union

Model 2: log(wage) ~ education + experience + gender + occupation + union

Res.Df RSS Df Sum of Sq F Pr(>F)

1 529 105.092

2 524 97.915 5 7.1769 7.6816 **5.535e-07 \*\*\***

Signif. codes: 0 ‘\*\*\*’ 0.001 ‘\*\*’ 0.01 ‘\*’ 0.05 ‘.’ 0.1 ‘ ’ 1

Comment: the p-value of the partial F-test is much smaller than 0.05, so the factor **occupation** is significant and should remain in the model. The wages differ by occupation category. It’s not necessary to refine the model.

[c] Investigate the model with **car::residualPlots( )**. Discuss the output and decide whether it is advisable to refine the model. (0.25 point)

car::residualPlots(model.full, main="Full model")

Test stat Pr(>|Test stat|)

education -0.3367 0.7365

experience -4.2117 2.985e-05 \*\*\*

gender

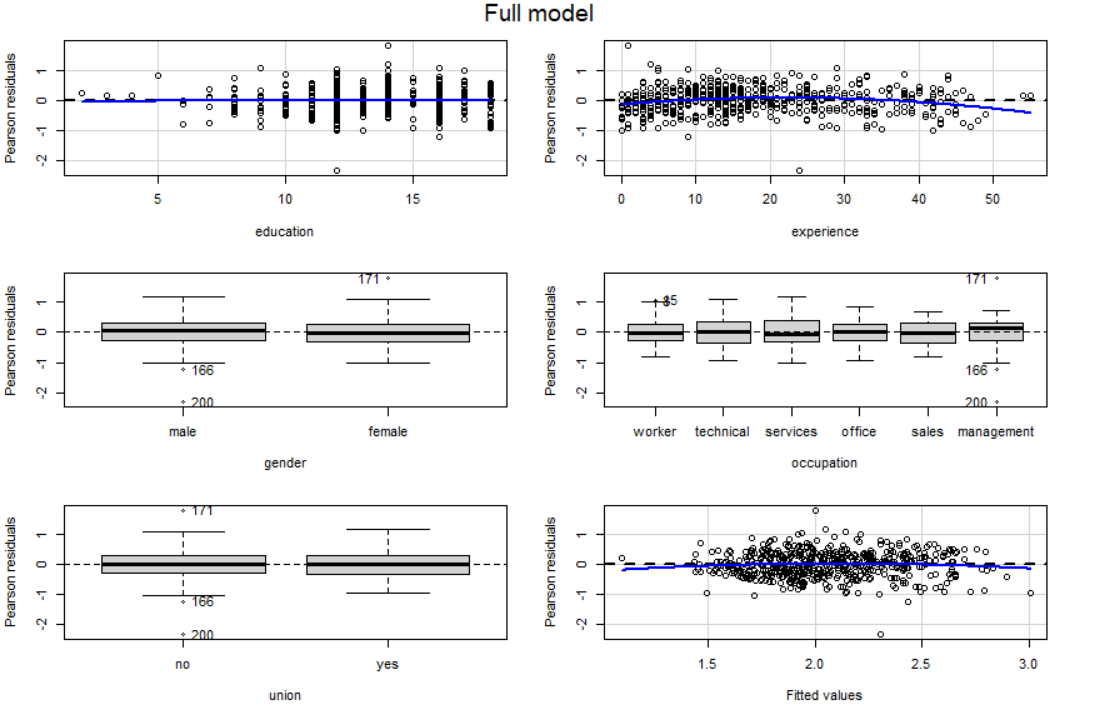
occupation

union

Tukey test -1.3752 0.1691

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Signif. codes: 0 ‘\*\*\*’ 0.001 ‘\*\*’ 0.01 ‘\*’ 0.05 ‘.’ 0.1 ‘ ’ 1



Comment: The significant specification test for **experience+experience2** as well as the quadratic lowess line in the residual plot for **experience** clearly indicate the need for a quadratic specification of models as

### Task 5: Case statistics of the final model (1 point)

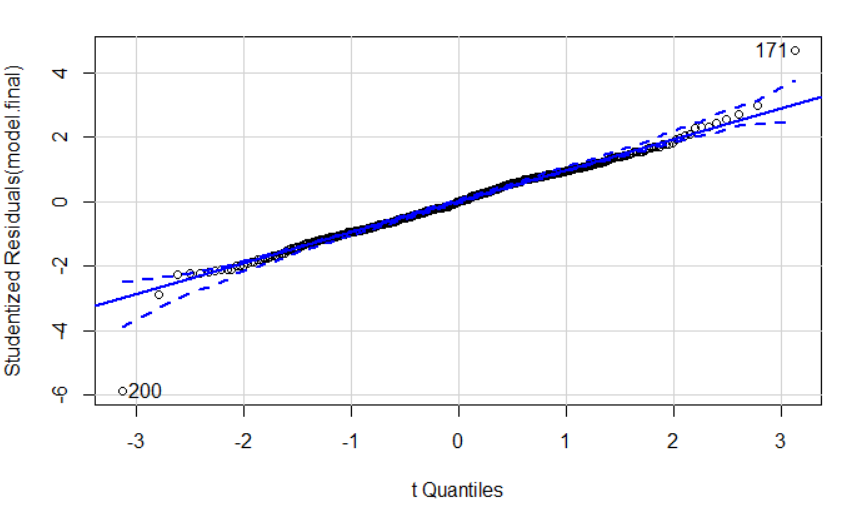
[a] Generate the following plots and *interpret* them for your final model. (0.75 points)

1. Identify the two most extreme observations with a **car::qqPlot( )** and interpret it.

model.final <-lm(log(wage)~education+experience+I(experience^2)+gender+occupation+union, data=CPS1985)

car::qqPlot(model.final, id.n=2)

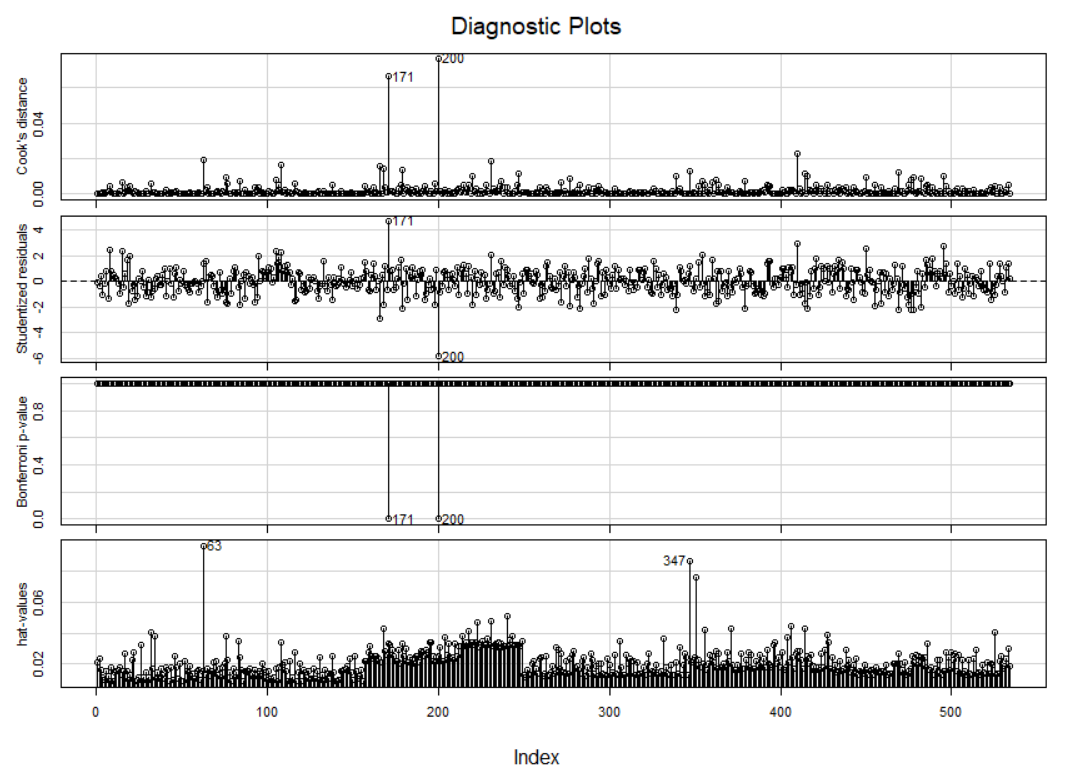
[1] 171 200



Comment: Most standardized residuals are within the confidence interval around straight line associated with equal quantiles for the observed studentized residuals and their theoretical t-distribution. However, the 171st and 200th observations deviate substantially and are therefore potential outliers.

1. Identify *potential* extreme observations with a **car::influenceIndexPlot( )** and interpret the plots.

car::influenceIndexPlot(model.final, id.n=2)



Comments:

Studentized residuals: Two extreme cases 171 and 200 were detected

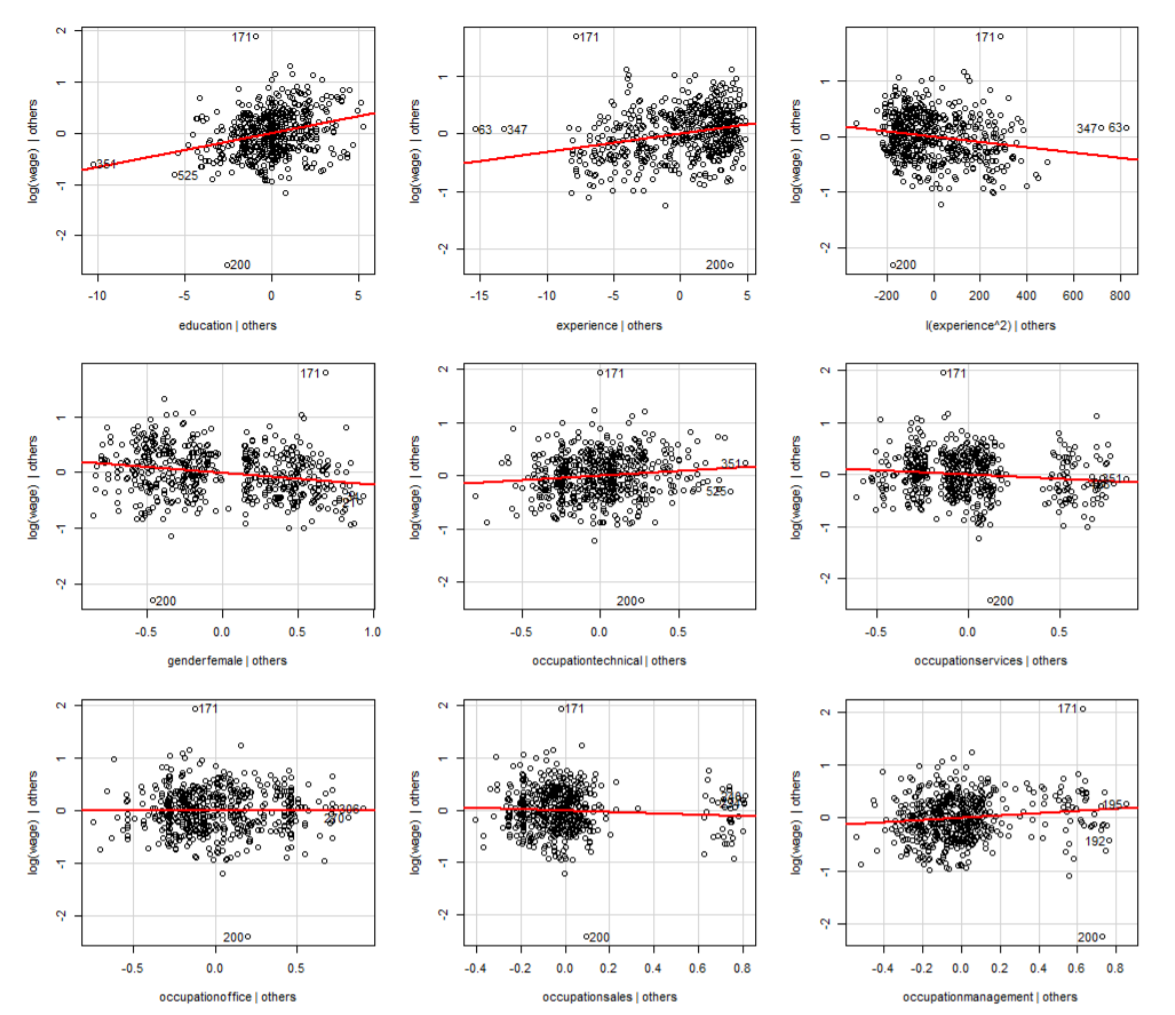
Cook’s distance: cases 171 and 200

Bonferoni adjusted p-values: Significant outliers for cases 171 and 200

Leverage value: Special combinations of independent variables are prominent for cases 63 and 347

1. Identify the two most extreme observation with a **car::avPlots( )** and interpret the plots.

car::avPlots(model.final, id.n=2)





Comment: The 171st and 200th observations are outliers, which affect all added-variable plots most. They exhibit an unusual wage of $44.5 for one year of job experience and being 21 years old for observation 171; perhaps this person is the boss’s daughter on a summer job. For observation #200 makes $1 hourly while being in management; perhaps this person is a family member who for tax reasons only makes $1.

[b] Inspect the ***two*** most extreme observations in the data-frame by examining their records. (0.25 points)

1. Discuss their attributes and argue if they are representative of the underlying population.

CPS1985[c(63,171,200,347), ]

|  |  |  |  |  |  |  |  |  |  |  |  |
| --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- | --- |
|  | wage | education | experience | age | ethnicity | region | gender | occupation | sector | union | married |
| 63 | 7 | 3 | 55 | 64 | hispanic | south | male | worker | manufacturing | no | yes |
| 171 | 44.5 | 14 | 1 | 21 | cauc | other | female | management | other | no | no |
| 200 | 1 | 12 | 24 | 42 | cauc | other | male | management | other | no | yes |
| 347 | 6 | 4 | 54 | 64 | cauc | other | male | services | other | no | yes |

1. Drop them from the data-frame and show your code for doing so.

CPS1985 <- CPS1985[-c(171, 200), ]