

ENGR-E 511; ENGR-E 399

“Machine Learning for Signal Processing”

Module 01: Lecture 01: Probability

Minje Kim

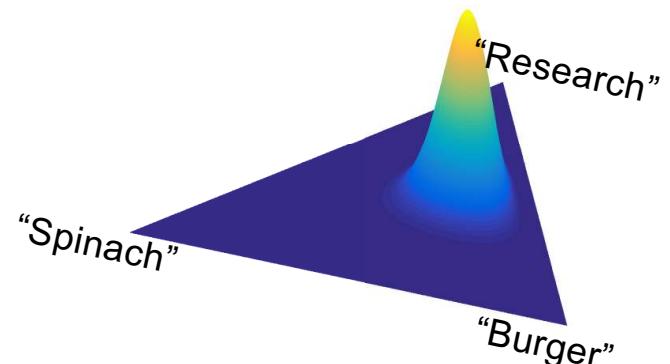
Department of Intelligent Systems Engineering

Email: minje@indiana.edu

Website: <http://minjekim.com>

Research Group: <http://saige.sice.indiana.edu>

Meeting Request: <http://doodle.com/minje>



Minje Kim

January 2 · Bloomington · *

I hate spinach. I like my wife.
I like research. I like my burger.

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Basic Probability Theories

- Mean & Random Variables

- My wife has been worried about my health during my PhD study
- So, after I finished my dissertation she ordered me to lose weight
- I worked out, ate more greens, stayed away from junk food, etc.
- A month later, I weighed and reported a measurement to her
- (Of course) she didn't trust me. Why?
 - Because it was just one observation
- I had to measure my weight five more times in the following mornings, and averaged the results
 - "Are you happy now?"
- What are we doing here?
 - Empirical moments can be used as a summary of your data
 - Measurements are with error (random variable)



$$\frac{1}{5} \sum_{i=1}^5 x_i$$



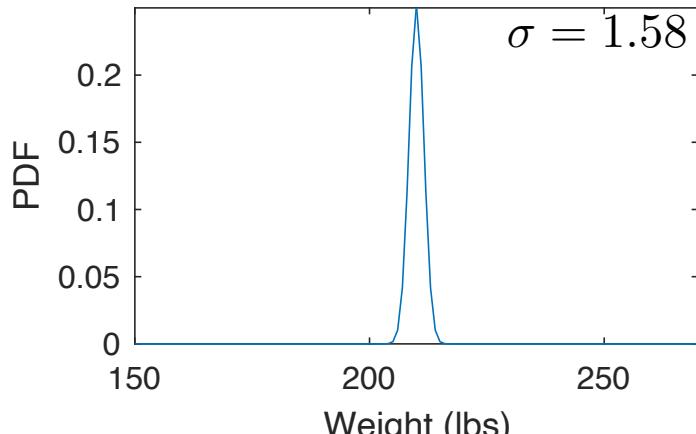
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Basic Probability Theories

- Gaussian (Normal) Distribution

- It turned out that the average is 210 lbs
 - Is that all she wants to know?
 - Gaussian (Normal) distribution with same means can have different meanings



[211, 208, 210, 209, 212]

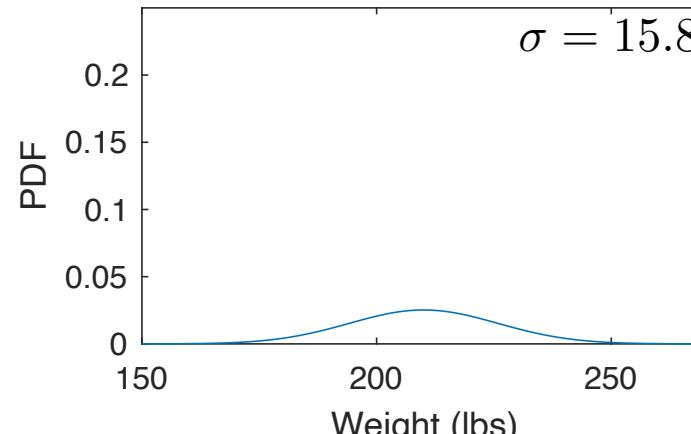
$$\frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

Normalization

How far you are from the mean

How certain you are

The diagram shows the Gaussian probability density function. A red dotted box highlights the term $\frac{1}{\sqrt{2\pi\sigma^2}}$, with an arrow pointing to it labeled "Normalization". Another red dotted box highlights the term $\frac{(x-\mu)^2}{2\sigma^2}$, with an arrow pointing to it labeled "How far you are from the mean". A third red dotted box highlights the term e^{-} , with an arrow pointing to it labeled "How certain you are".



[190, 200, 210, 220, 230]

- Are we good?
 - My weights are following a Gaussian distribution?



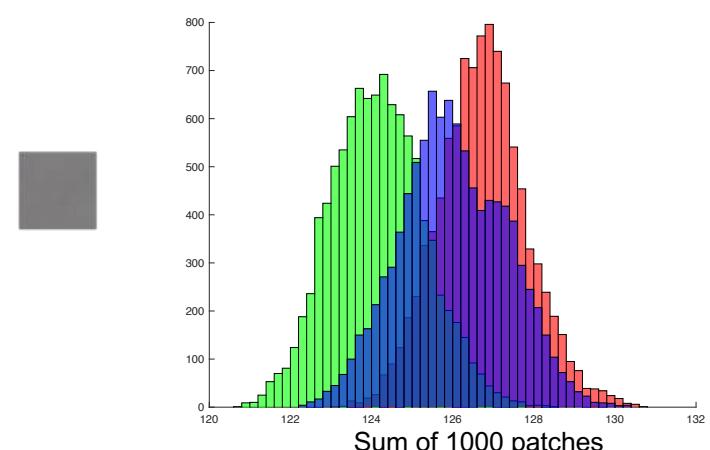
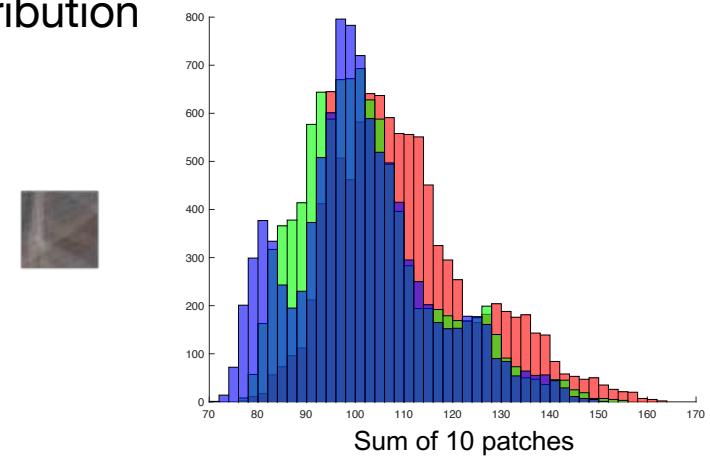
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Basic Probability Theories

- Central Limit Theorem

- Sum of many i.i.d. random samples follows a Gaussian distribution
 - That's why it's the *Normal* distribution



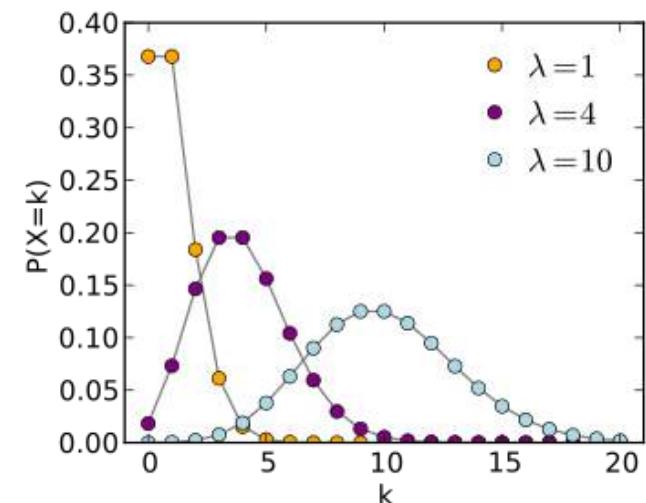
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Basic Probability Theories

- Poisson Distribution

- My wife is worried if I forget to measure my weight everyday
 - So, she decided to ask me at the end of the day, and count it
[1, 0, 1, 2, 1, 1, 0, 2]
 - What's the mean/standard deviation of these counts?
 - Is this Gaussian?
- There are other distributions defined with a different set of parameters
 - i.e. Poisson distribution
$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$$
 - Is it different from Gaussian?
- The counts are actually following a Poisson distribution with $\lambda = 1$
 - Although it's kinda unclear just by watching the samples:
[1, 0, 1, 2, 1, 1, 0, 2]



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https://en.wikipedia.org/wiki/Poisson_distribution

Basic Probability Theories

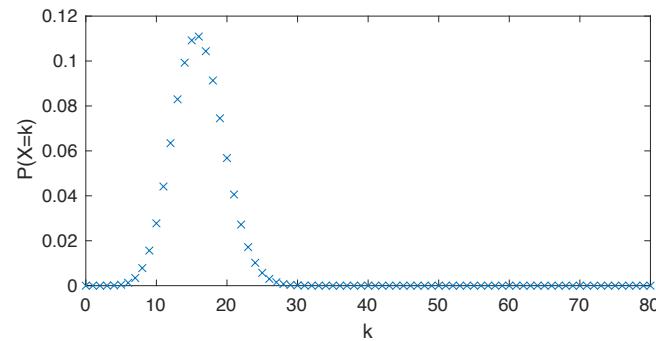
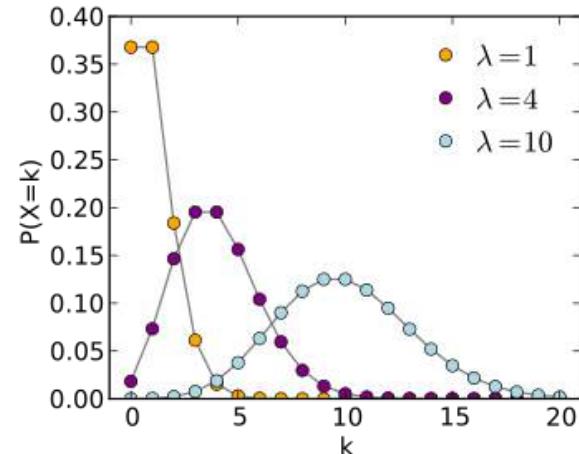
- Binomial Distribution

- She thinks that I'm not checking on my weight frequently enough
 - Because she's doing it more often:
[2, 3, 4, 4, 5, 5, 6, 3]
 - Her $\lambda = 4$
- For given eight days, there were 40 measurements in the house
 - We know that only 20% of them are mine (why?)
- In the next month, we observe another set of 80 measurements in the same family of two people
 - What's the probability that 40 of them are mine?
 - What's the probability that 16 of them are mine?

○ Binomial distribution

$$P(X = k) = \binom{N}{k} p^k (1 - p)^{N-k} \quad p = \frac{\lambda_1}{\lambda_1 + \lambda_2}$$

- Why do we need the combination?



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Basic Probability Theories

- Multinomial Distribution

- Let's say that there's another person in the family
- We need another distribution called **Multinomial distribution**

$$P(X_1 = x_1, X_2 = x_2, \dots, X_K = x_K) = \frac{N!}{x_1!x_2!\dots x_K!} p_1^{x_1} p_2^{x_2} \cdots p_K^{x_K}$$

$$p_k = \frac{\lambda_k}{\lambda_1 + \lambda_2 + \cdots + \lambda_k}$$

- This is a very useful distribution



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What's the probability of seeing the word "my" 2 times,
"like" 3 times, and so on, among the 14 words?

k	p _k
I	.30
my	.20
wife	.13
today	.12
like	.10
research	.05
hate	.04
burger	.04
spinach	.02



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From Probability to Machine Learning

- Maximum Likelihood

- My weighing behavior follows Poisson dist. Can we be more specific?

- IOW, can we estimate λ from the data samples? [1, 0, 1, 2, 1, 1, 0, 2]

- $P^4(X = 1) \cdot P^2(X = 0) \cdot P^2(X = 2) = \left(\frac{\lambda e^{-\lambda}}{1}\right)^4 \cdot \left(\frac{\lambda^0 e^{-\lambda}}{1}\right)^2 \cdot \left(\frac{\lambda^2 e^{-\lambda}}{2}\right)^2 = \lambda^8 e^{-8\lambda} / 4$

- $\lambda = 1.0 : 8.3865 \times 10^{-5}$

- $\lambda = 1.5 : 3.9367 \times 10^{-5}$

- $\lambda = 3.0 : 6.1921 \times 10^{-8}$ An analytic way?

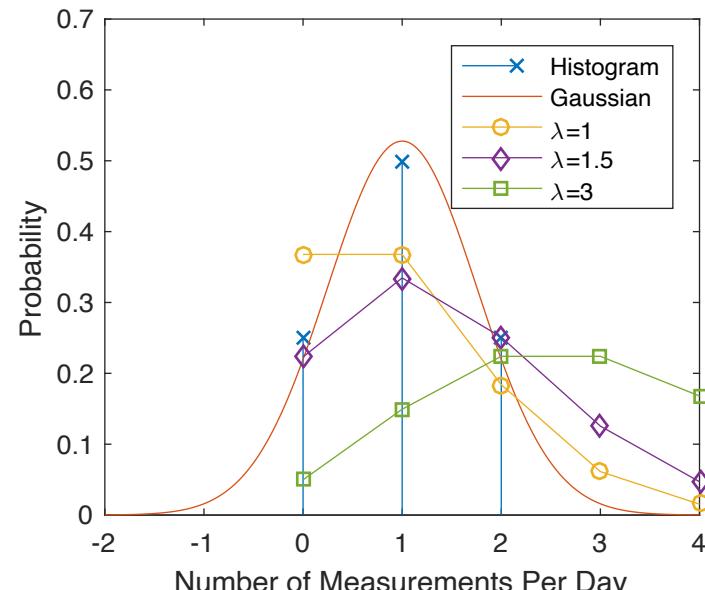
- Maximum likelihood estimation

$$\arg \max_{\lambda} \prod_i P(X = x_i) = \arg \max_{\lambda} \prod_i \frac{\lambda^{x_i} e^{-\lambda}}{x_i!}$$

$$\begin{aligned} \arg \max_{\lambda} \sum_i \ln P(X = x_i) &= \arg \max_{\lambda} \sum_i (x_i \ln \lambda - \lambda - \ln x_i!) \\ &= \arg \max_{\lambda} \ln \lambda \sum_i x_i - N\lambda \end{aligned}$$

$$\frac{\partial \ln \lambda \sum_i x_i - N\lambda}{\partial \lambda} = \frac{\sum_i x_i}{\lambda} - N \quad \lambda^* = \frac{1}{N} \sum_i x_i$$

$$P(X = k) = \frac{\lambda^k e^{-\lambda}}{k!}$$



Why the mismatch between the histogram and the ground-truth distribution?

→ Lack of data



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From Probability to Machine Learning

- Independence, Conditional Probabilities

- It turned out that I lost 5 lbs for the first month (it's a true story)
 - And, I bragged

- My wife argued that (as always she's right)
“You're taller than me, so it's easier for you to lose weight”

1. $P(\text{“weight"} > 200\text{lbs}) = 0.45$
2. $P(\text{“weight"} > 200\text{lbs} \mid \text{“height"} > 6.5\text{ft}) = ?$
3. $P(\text{“weight"} > 200\text{lbs} \mid \text{“eye color"} = \text{“black"}) = ?$

- The probabilities 1 and 3 are same, and we say that the weight and eye color are **independent** $0.15/0.33 \approx 0.45$

- The probability 2 and 3 are called **conditional probabilities** given the observations about the height and eye color, respectively

- $P(\text{“weight"} > 200\text{lbs}, \text{“eye color"} = \text{“black"})$
 $= P(\text{“weight"} > 200\text{lbs} \mid \text{“eye color"} = \text{“black"}) \times P(\text{“eye color"} = \text{“black"})$

- What if they are independent?

- $P(200\text{lbs} < \text{“weight"}, \text{“eye color"} = \text{“black"})$

$= P(200\text{lbs} < \text{“weight"}) \times P(\text{“eye color"} = \text{“black"})$ Check if this is true in the tables!

	W<200	W>=200	P(H)
H>=6.5	0.13	0.27	0.40
H<6.5	0.42	0.18	0.60
P(W)	0.55	0.45	

	W<200	W>=200	P(EC)
EC=BK	0.18	0.15	0.33
EC=BR	0.37	0.30	0.67
P(W)	0.55	0.45	



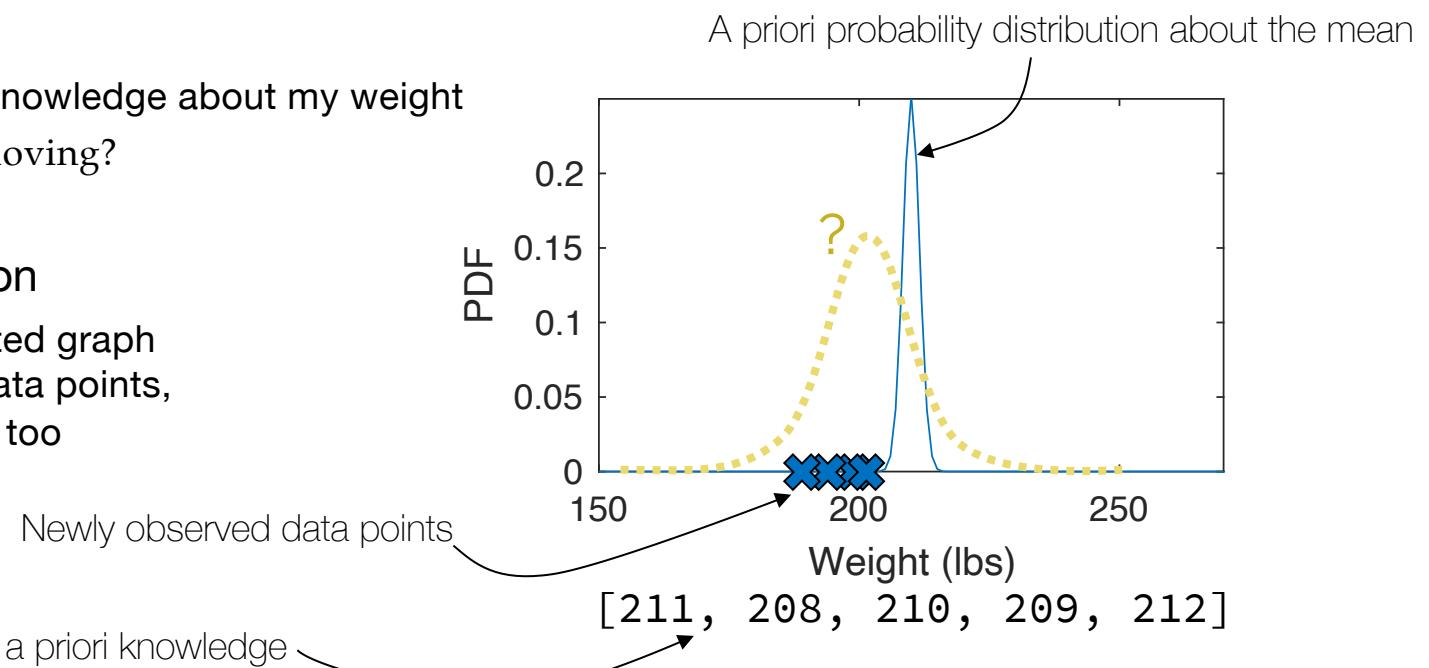
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From Probability to Machine Learning

- Maximum A Posteriori

- We moved into Bloomington after my graduation.
After unpacking, I was asked to weigh myself immediately
 - It turned out that I finally lost 15 lbs
- Once again, she didn't trust me
 - Why is she so grumpy?
 - She has something called **a priori** knowledge about my weight
 - "What if the scale is broken while moving?
It must be 210, not 200."
- **Maximum A Posteriori** estimation
 - We need to estimate the yellow dotted graph
not only from the newly observed data points,
but from the previous observations, too



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From Probability to Machine Learning

- Maximum A Posteriori

- Bayes' theorem

$$P(\mu|[x_1, x_2, \dots, x_5]) = \frac{P([x_1, x_2, \dots, x_5]|\mu)P(\mu)}{P([x_1, x_2, \dots, x_5]|)}$$

{A Posteriori} = $\frac{\text{Likelihood} \times \text{A Priori}}{\text{Normalizing Constant}}$

- MAP estimation:

- To find the parameter that maximizes the **a posteriori distribution**

- We can find an analytic MAP estimation with certain conditions

$$\begin{aligned}\mathcal{LL} &= \ln P([x_1, x_2, \dots, x_5]|\mu)P(\mu) = \sum_{i=1}^5 \ln P(x_i|\mu) + \ln P(\mu) & \frac{\partial \mathcal{LL}}{\partial \mu} &= \sum_i^5 \frac{x_i - \mu}{\sigma^2} - \frac{\mu - \mu_0}{\sigma_0^2} = \frac{\sigma_0^2 \sum_{i=1}^5 x_i + \sigma^2 \mu_0 - (5\sigma_0^2 + \sigma^2)\mu}{\sigma^2 \sigma_0^2} \\ &= \sum_i^5 \ln \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x_i - \mu)^2}{2\sigma^2}} + \ln \frac{1}{\sqrt{2\pi\sigma_0^2}} e^{-\frac{(\mu - \mu_0)^2}{2\sigma_0^2}} & \mu^* &= \frac{\sigma_0^2 \sum_{i=1}^5 x_i + \sigma^2 \mu_0}{5\sigma_0^2 + \sigma^2} & \text{Just another kind of average}\end{aligned}$$

- Does it always work like this?



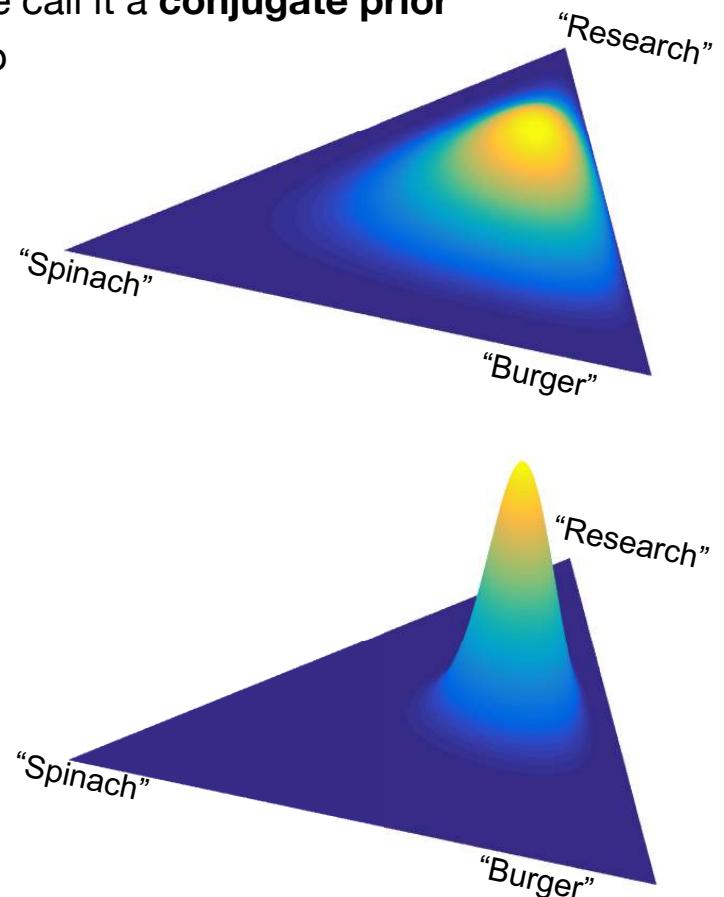
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From Probability to Machine Learning

- Conjugate Priors

- If the a priori distribution has the similar algebraic form with the likelihood, we call it a **conjugate prior**
 - This ensures the posterior distributions have the same algebraic forms, too
- You observed in my tweets that I used
 - “Research”: 5 times
 - “Burger”: 4 times
 - “Spinach”: 2 times
- What if you’ve watched them for a month and accumulated more evidence?
 - “Research”: 19 times
 - “Burger”: 15 times
 - “Spinach”: 7 times
- One of these is the distribution where I sample from to decide what to talk about on the tweet on a given day
- It looks like multinomial, but it’s not
 - The RV is not for the counts, but for the parameter
$$P(\Theta_1 = p_1, \Theta_2 = p_2, \dots, \Theta_K = p_K)$$
 - The counts don’t have to be integers (what?)



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From Probability to Machine Learning

- Conjugate Priors

○ Dirichlet distribution

- $P(\Theta_1 = p_1, \Theta_2 = p_2, \dots, \Theta_K = p_K | \alpha_1, \alpha_2, \dots, \alpha_K) = \frac{\Gamma(\sum_{k=1}^K \alpha_k)}{\prod_{k=1}^K \Gamma(\alpha_k)} \prod_{k=1}^K p_k^{\alpha_k - 1}$
 $\sum_{k=1}^K p_k = 1, \quad p_k \geq 0 \quad \forall p_k, \quad \alpha_k > 0 \quad \forall \alpha_k$

□ We call α_k a **hyperparameter** or **pseudo count**

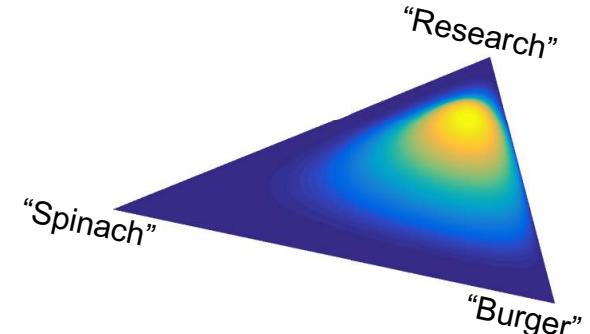
□ The conjugate prior of multinomial distribution

- Because... $P(X_1 = x_1, X_2 = x_2, \dots, X_K = x_K | \Theta_1 = p_1, \Theta_2 = p_2, \dots, \Theta_K = p_K)$
· $P(\Theta_1 = p_1, \Theta_2 = p_2, \dots, \Theta_K = p_K)$

$$= \frac{N!}{\prod_{k=1}^K x_k!} \prod_{k=1}^K p_k^{x_k} \cdot \frac{\Gamma(\sum_{k=1}^K \alpha_k)}{\prod_{k=1}^K \Gamma(\alpha_k)} \prod_{k=1}^K p_k^{\alpha_k - 1}$$

$$= \frac{N!}{\prod_{k=1}^K x_k!} \frac{\Gamma(\sum_{k=1}^K \alpha_k)}{\prod_{k=1}^K \Gamma(\alpha_k)} \prod_{k=1}^K p_k^{x_k + \alpha_k - 1}$$

$$\propto \prod_{k=1}^K p_k^{x_k + \alpha_k - 1} \quad \leftarrow \text{MAP is to maximize this w.r.t. } \Theta$$



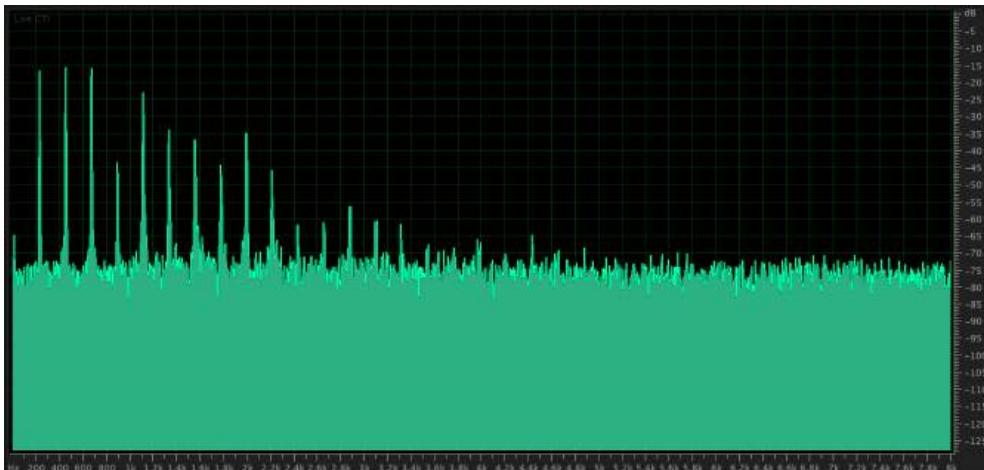
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Information Theory

- Entropy

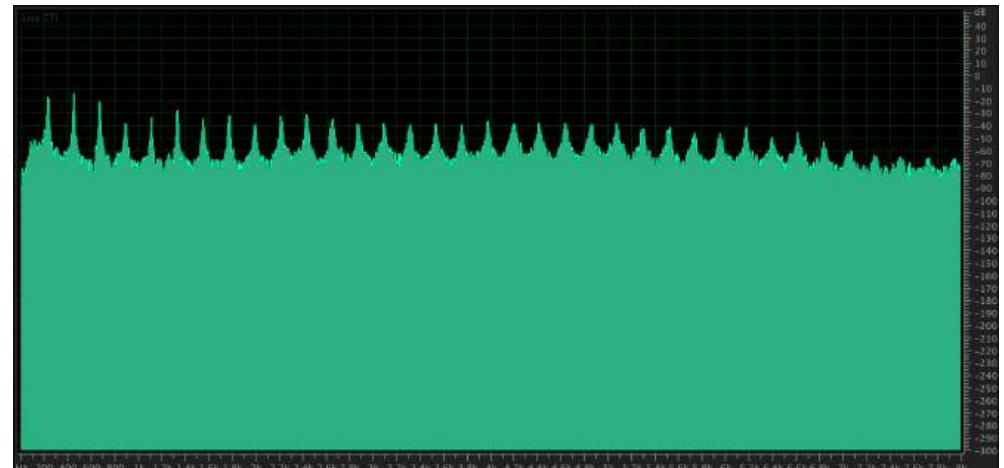
- Let's talk about audio a little bit
 - Flute vs. electric guitar
 - Electric guitar is noisier



Flute



E.G.



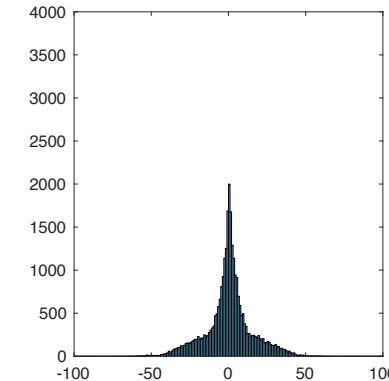
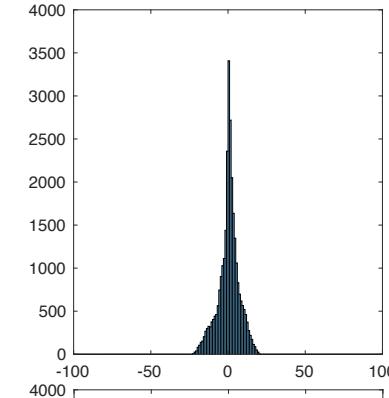
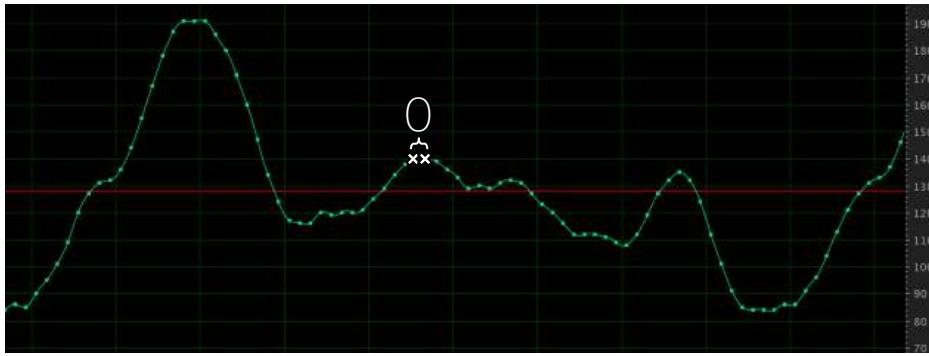
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Information Theory

- Entropy

- Calculate the difference between adjacent samples
 - Flute is with a smoother wave form → sharper distribution of differences
 - Electric guitar is the opposite



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Information Theory

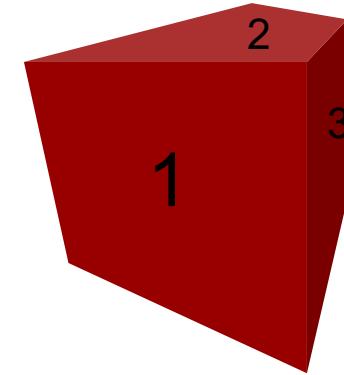
- Entropy coding



Information Theory

- Entropy coding

- The (expected) amount of information in the message
 - Represented in terms of events and their probabilities
- Less probable events have more information
 - A dice with all ones → no information
 - An unfair dice: (1, .9), (2, .05), (3, .02), (4, .015), (5, .01), (6, .005)
 - 1: 0
 - 2: 10
 - 3: 110
 - 4: 1110
 - 5: 11110
 - 6: 11111
 - Expected code length for an event:
$$0.9 \times 1 + 0.05 \times 2 + 0.02 \times 3 + 0.015 \times 4 + 0.01 \times 5 + 0.005 \times 5 = 1.195 < \lceil \log_2 6 \rceil = 3$$
 - Is this the best we can do?



Information Theory

- Entropy coding

- We choose the **information function** to be the negative logarithm of the probability

$$I(x_i) = \log(1/p_i) = -\log(p_i), \quad p_i = P(x_i)$$

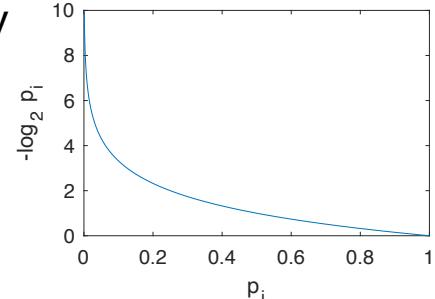
- $-\log(p_i)$ corresponds to the number of bits

- More probable events carry less information (bits)
 - No information when the probability is 1
 - Events that happen at the same time carry the sum of the amount of information each of which carries, i.e.
 $-\log(p_i \cdot p_j) = -\log(p_i) - \log(p_j)$

- If we can assign $-\log(p_i)$ bits to x_i , that's an efficient coding scheme

- (Shannon's) **Entropy** $H(X) = -\sum_i p_i \log(p_i)$

- Lower bound of the average code length for a given distribution
 - Each distribution has a unique entropy
 - e.g. Flute's distribution of sample difference: 4.6889
 - e.g. Electric guitar: 5.7527
 - E.G. needs longer code because its distribution is less ordered



$$\frac{4.6889}{5.7527} = 0.8151$$

$$\frac{136987}{167732} = 0.8167$$



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Information Theory

- Kullback-Leibler Divergence

○ Cross Entropy

- Calculates entropy against a different distribution

- If p_i and q_i are same, $-\sum_i p_i \log q_i = -\sum_i p_i \log p_i$

- Otherwise

- For example, with a wrong $I(X)$ (two slides ago, the unfair dice example)

$$0.9 \times 1 + 0.05 \times 2 + 0.02 \times 3 + 0.015 \times 4 + 0.01 \times 5 + 0.005 \times 5 = 1.195$$

- With an optimal $I(X)$ (i.e. $-\log p_i$), cross entropy is minimal

$$0.9 \times 0.15 + 0.05 \times 4.32 + 0.02 \times 5.64 + 0.015 \times 6.06 + 0.01 \times 6.64 + 0.005 \times 7.64 = 0.6613$$

- This starts to look like a distance (not exactly though)

○ KL Divergence

- The amount of information lost by an approximation
- Frequently used to measure the difference between distributions

$$\mathcal{D}_{KL}(P(X)||Q(X)) = H(P(X), Q(X)) - H(P(X)) = -\sum_i p_i \log q_i + \sum_i p_i \log p_i = \sum_i p_i \log \frac{p_i}{q_i}$$



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Thank You!



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