



Math for the people, by the people.

likelihood function

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Defines	maximum likelihood estimate
Defines	MLE
Defines	log-likelihood function

Let $\mathbf{X}=(X_1, \dots, X_n)$ be a random vector and

$$\{f_{\mathbf{X}}(\mathbf{x} \mid \boldsymbol{\theta}) : \boldsymbol{\theta} \in \Theta\}$$

a statistical model parametrized by $\boldsymbol{\theta} = (\theta_1, \dots, \theta_k)$, the parameter vector in the *parameter space* Θ . The *likelihood function* is a map $L : \Theta \rightarrow \mathbb{R}$ given by

$$L(\boldsymbol{\theta} \mid \mathbf{x}) = f_{\mathbf{X}}(\mathbf{x} \mid \boldsymbol{\theta}).$$

In other words, the likelihood function is functionally the same in form as a probability density function. However, the emphasis is changed from the \mathbf{x} to the $\boldsymbol{\theta}$. The pdf is a function of the x 's while holding the parameters θ 's constant, L is a function of the parameters θ 's, while holding the x 's constant.

When there is no confusion, $L(\boldsymbol{\theta} \mid \mathbf{x})$ is abbreviated to be $L(\boldsymbol{\theta})$.

The parameter vector $\hat{\boldsymbol{\theta}}$ such that $L(\hat{\boldsymbol{\theta}}) \geq L(\boldsymbol{\theta})$ for all $\boldsymbol{\theta} \in \Theta$ is called a *maximum likelihood estimate*, or *MLE*, of $\boldsymbol{\theta}$.

Many of the density functions are exponential in nature, it is therefore easier to compute the MLE of a likelihood function L by finding the maximum of the natural log of L , known as the log-likelihood function:

$$\ell(\boldsymbol{\theta} \mid \mathbf{x}) = \ln(L(\boldsymbol{\theta} \mid \mathbf{x}))$$

due to the monotonicity of the log function.

Examples:

1. A coin is tossed n times and m heads are observed. Assume that the probability of a head after one toss is π . What is the MLE of π ?

Solution: Define the outcome of a toss be 0 if a tail is observed and 1 if a head is observed. Next, let X_i be the outcome of the i th toss. For any single toss, the density function is $\pi^x(1 - \pi)^{1-x}$ where $x \in \{0, 1\}$. Assume that the tosses are independent events, then the joint probability density is

$$f_{\mathbf{X}}(\mathbf{x} \mid \pi) = \binom{n}{\sum x_i} \pi^{\sum x_i} (1 - \pi)^{\sum (1-x_i)} = \binom{n}{m} \pi^m (1 - \pi)^{n-m},$$

which is also the likelihood function $L(\pi)$. Therefore, the log-likelihood function has the form

$$\ell(\pi \mid \mathbf{x}) = \ell(\pi) = \ln \binom{n}{m} + m \ln(\pi) + (n - m) \ln(1 - \pi).$$

Using standard calculus, we get that the MLE of π is

$$\hat{\pi} = \frac{m}{n} = \bar{x}.$$

2. Suppose a sample of n data points X_i are collected. Assume that the $X_i \sim N(\mu, \sigma^2)$ and the X_i 's are independent of each other. What is the MLE of the parameter vector $\boldsymbol{\theta} = (\mu, \sigma^2)$?

Solution: The joint pdf of the X_i , and hence the likelihood function, is

$$L(\boldsymbol{\theta} \mid \mathbf{x}) = \frac{1}{\sigma^n (2\pi)^{n/2}} \exp\left(-\frac{\sum (x_i - \mu)^2}{2\sigma^2}\right).$$

The log-likelihood function is

$$\ell(\boldsymbol{\theta} \mid \mathbf{x}) = -\frac{\sum (x_i - \mu)^2}{2\sigma^2} - \frac{n}{2} \ln(\sigma^2) - \frac{n}{2} \ln(2\pi).$$

Taking the first derivative (gradient), we get

$$\frac{\partial \ell}{\partial \boldsymbol{\theta}} = \left(\frac{\sum (x_i - \mu)}{\sigma^2}, \frac{\sum (x_i - \mu)^2}{2\sigma^4} - \frac{n}{2\sigma^2} \right).$$

Setting

$$\frac{\partial \ell}{\partial \boldsymbol{\theta}} = \mathbf{0} \text{ See score function}$$

and solve for $\boldsymbol{\theta} = (\mu, \sigma^2)$ we have

$$\hat{\boldsymbol{\theta}} = (\hat{\mu}, \hat{\sigma}^2) = (\bar{x}, \frac{n-1}{n} s^2),$$

where $\bar{x} = \sum x_i / n$ is the sample mean and $s^2 = \sum (x_i - \bar{x})^2 / (n-1)$ is the sample variance. Finally, we verify that $\hat{\boldsymbol{\theta}}$ is indeed the MLE of $\boldsymbol{\theta}$ by checking the negativity of the 2nd derivatives (for each parameter).