

## planetmath.org

Math for the people, by the people.

## Fisher information matrix

Canonical name FisherInformationMatrix

Date of creation 2013-03-22 14:30:15 Last modified on 2013-03-22 14:30:15

Owner CWoo (3771) Last modified by CWoo (3771)

Numerical id 14

Author CWoo (3771)
Entry type Definition
Classification msc 62H99
Classification msc 62B10
Classification msc 62A01

Synonym information matrix
Defines Fisher information

Defines information

Defines Cramer-Rao inequality
Defines Cramer-Rao lower bound

Given a statistical model  $\{f_{\mathbf{X}}(\mathbf{x} \mid \boldsymbol{\theta})\}\$  of a random vector  $\mathbf{X}$ , the *Fisher information matrix*, I, is the variance of the score function U. So,

$$I = Var[U].$$

If there is only one parameter involved, then I is simply called the *Fisher information* or *information* of  $f_{\mathbf{X}}(\mathbf{x} \mid \theta)$ .

## Remarks

• If  $f_{\mathbf{X}}(\boldsymbol{x} \mid \boldsymbol{\theta})$  belongs to the exponential family,  $I = \mathbb{E}\left[U^{\mathrm{T}}U\right]$ . Furthermore, with some regularity conditions imposed, we have

$$I = -\operatorname{E}\left[\frac{\partial U}{\partial \boldsymbol{\theta}}\right].$$

• As an example, the normal distribution,  $N(\mu, \sigma^2)$ , belongs to the exponential family and its log-likelihood function  $\ell(\boldsymbol{\theta} \mid x)$  is

$$-\frac{1}{2}\ln(2\pi\sigma^2) - \frac{(x-\mu)^2}{2\sigma^2},$$

where  $\boldsymbol{\theta} = (\mu, \sigma^2)$ . Then the score function  $U(\boldsymbol{\theta})$  is given by

$$\left(\frac{\partial \ell}{\partial \mu}, \frac{\partial \ell}{\partial \sigma^2}\right) = \left(\frac{x-\mu}{\sigma^2}, \frac{(x-\mu)^2}{2\sigma^4} - \frac{1}{2\sigma^2}\right).$$

Taking the derivative with respect to  $\theta$ , we have

$$\frac{\partial U}{\partial \boldsymbol{\theta}} = \begin{pmatrix} \frac{\partial U_1}{\partial \mu} & \frac{\partial U_2}{\partial \mu} \\ \\ \frac{\partial U_1}{\partial \sigma^2} & \frac{\partial U_2}{\partial \sigma^2} \end{pmatrix} = \begin{pmatrix} \frac{-1}{\sigma^2} & -\frac{x-\mu}{\sigma^4} \\ \\ -\frac{x-\mu}{\sigma^4} & \frac{1}{2\sigma^4} - \frac{(x-\mu)^2}{\sigma^6} \end{pmatrix}.$$

Therefore, the Fisher information matrix I is

$$-\operatorname{E}\left[\frac{\partial U}{\partial \boldsymbol{\theta}}\right] = \frac{1}{2\sigma^4} \begin{pmatrix} 2\sigma^2 & 0\\ 0 & -1 \end{pmatrix}.$$

• Now, in linear regression model with constant variance  $\sigma^2$ , it can be shown that the Fisher information matrix I is

$$\frac{1}{\sigma^2} \mathbf{X}^{\mathrm{T}} \mathbf{X},$$

where **X** is the design matrix of the regression model.

• In general, the Fisher information meansures how much "information" is known about a parameter  $\theta$ . If T is an unbiased estimator of  $\theta$ , it can be shown that

$$\operatorname{Var}\left[T(X)\right] \ge \frac{1}{I(\theta)}$$

This is known as the *Cramer-Rao inequality*, and the number  $1/I(\theta)$  is known as the *Cramer-Rao lower bound*. The smaller the variance of the estimate of  $\theta$ , the more information we have on  $\theta$ . If there is more than one parameter, the above can be generalized by saying that

$$\operatorname{Var}\left[T(X)\right] - I(\boldsymbol{\theta})^{-1}$$

is positive semidefinite, where I is the Fisher information matrix.