

# BSCMA1004

STATISTICS II NOTES



**WEEK 4 NOTES**

IITM B.S Degree

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## Statistics II - Week IV Notes

### ★ Continuous Random Variables

A drv  $X$  takes values in alphabet  $X$ .  $|X|$  is growing very large, is unwieldy for calculations.

Eg. Meteorite weight, Binomial ( $n, p$ )

Meteorite Data:

① Preprocessing: Take  $\log_2 \rightarrow [0.01, 6000000]$   
↓  
[-6.6, 25.8]

We still have 45000+ data

② Move from individual values to intervals.

Divide [-6.6, 25.8]  $\approx$  100 intervals.

[-6.6, -6.3], [-6.3, -6], ..., [25.5, 25.8]

Count no. of values falling inside each interval.

### ★ (Cumulative) Distribution Function [CDF]

The CDF of a rv  $X$ ,  $F_X(x)$  is a fn. from  $\mathbb{R}$  to  $[0, 1]$

$$F_X(x) = P(X \leq x)$$

$$F_X(b) - F_X(a) = P(a < X \leq b)$$

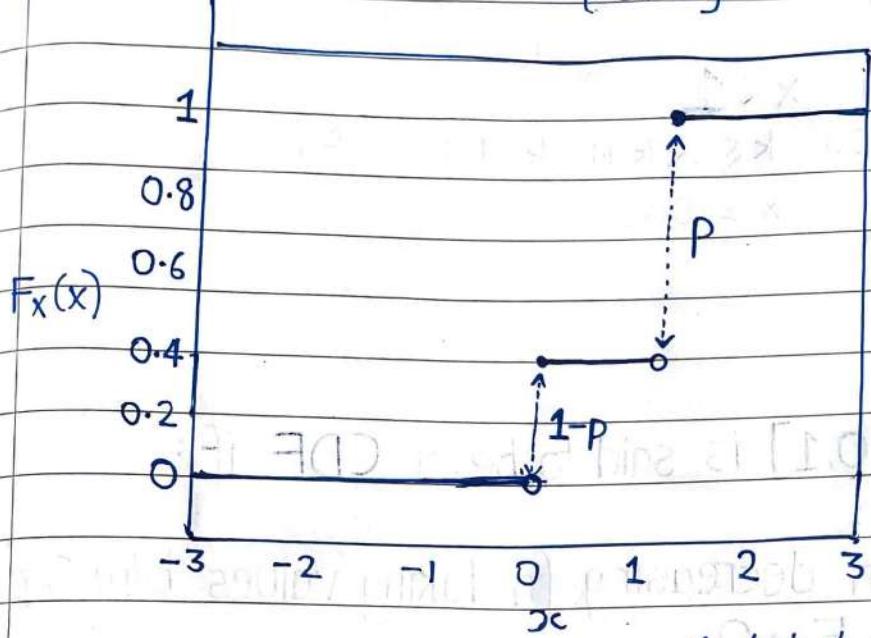
$F_X$ : non decreasing, non negative values

As  $x \rightarrow -\infty$ ;  $F_X \rightarrow 0$

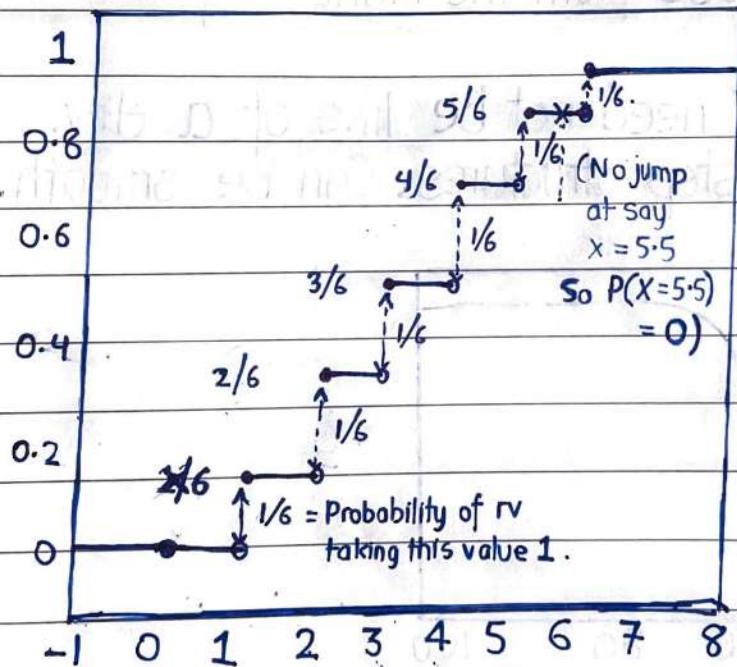
- As  $x \rightarrow \infty$ ,  $F_X \rightarrow 1$

## Eg. Bernoulli RV

$$X \sim \{0, 1\}$$



$$F_X(x) = \begin{cases} 0 & x < 0 \\ 1-p & 0 \leq x < 1 \\ 1 & x \geq 1 \end{cases}$$

Eg. Die throw  $X \sim \{1, 2, 3, 4, 5, 6\}$ 

# Computing Probability of Intervals using CDF



$X \sim \text{Uniform } \{1, 2, \dots, 100\}$

$$F_X(x) = \begin{cases} 0 & x < 1 \\ k/100 & 1 \leq x < k+1, k=1, 2, \dots, 99. \\ 1 & x \geq 100. \end{cases}$$

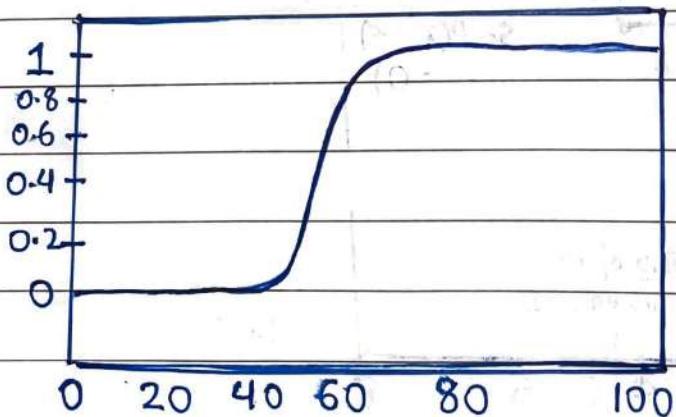


## CDF

A fn.  $F: \mathbb{R} \rightarrow [0, 1]$  is said to be a CDF if:

- ①  $F$  is a non decreasing fn taking values b/w 0 and 1.
- ② As  $x \rightarrow -\infty$ ,  $F \rightarrow 0$
- ③ As  $x \rightarrow \infty$ ,  $F \rightarrow 1$
- ④  $F$  is continuous from the right

A general CDF need not be like of a drv.  
 • No need for step structures. Can be smooth.



\* Binomial using continuous CDF

$$X \sim \text{Binomial}(100, 0.6)$$

$$F_X(k) = \sum_{j=0}^k {}^{100}C_j (0.6)^j (0.4)^{100-j}$$

$$F(x) = \frac{1}{1 + \exp\left(\frac{-1.65451(x-60)}{\sqrt{24}}\right)}$$

- $P(40 < X \leq 50) = 0.0271$        $F(50) - F(40) = 0.0318$        $(np(1-p))$
- $P(50 < X \leq 60) = 0.5108$        $F(60) - F(50) = 0.4670$
- $P(60 < X \leq 70) = 0.4473$        $F(70) - F(60) = 0.4670$
- $P(70 < X \leq 80) = 0.0148$        $F(80) - F(70) = 0.0318$ .

\* CDF & Random Variables

Given a valid CDF  $F(x)$ ,  $\exists$  rv  $X$  taking values in  $\mathbb{R}$  such that

$$P(X \leq x) = F(x)$$

- If  $F(x)$  rises from  $F_1$  to  $F_2$  at  $x_1$ ,  $P(X = x_1) = F_2 - F_1$
- If  $F(x)$  is continuous at  $x_0$ ,  $P(X = x_0) = 0$

\*  $P(1.99 < X < 2.01) = F(2.01) - F(1.99) = 0.002$  (say)

• Value with finite precision taken with +ve probability

\*  $P(1.999999 < X < 2.00000001) = 0.00000002$

\*  $P(X = 2.000000 \dots) = 0$

• Cannot take value with infinite precision when fn is cont.

## \* Continuous Random Variables

A random var.  $X$  with CDF  $F_X(x)$  is said to be a continuous RV if  $F_X(x)$  is continuous at every  $x$ .

- CDF has no jumps/steps
- $P(X=x) = 0$  for all  $x$ .
- $P(a < X \leq b) = F(b) - F(a)$

$$\therefore P(X=a) = 0 = P(X=b)$$

$$P(a \leq X \leq b) = P(a < X \leq b) = P(a \leq X < b) = P(a < X < b)$$

Many drvs are well approximated by crv's.

## \* Probability Density Function (PDF)

A crv  $X$  with CDF  $F_X(x)$  is said to have a PDF  $f_X(x)$ ; if  $\forall x_0$ ,

$$F_X(x_0) = \int_{-\infty}^{x_0} f_X(x) dx$$

- CDF is the integral of PDF
- Value of PDF around  $f_X(x_0) \equiv X$  takes a value around  $x_0$ .
- Higher the PDF, higher this chance
- The PDF can be  $> 1$ , if range of values is shorter.

A func.  $f : R \rightarrow R$  is said to be a density function if:

- ①  $f(x) \geq 0$
- ②  $\int_{-\infty}^{\infty} f(x) dx = 1$

- ③  $f(x)$  is piecewise continuous

- Given a density fn  $f$ , there is a crv  $X$  with PDF as  $f$ .

- Support of the random variable  $X$  with PDF  $f_x$  is  $\text{supp}(x) = \{x : f_x(x) > 0\}$

- $\text{supp}(x)$  contains intervals in which  $X$  can fall with +ve probability.  $\int_{\text{supp}(x)} f_x(x) dx = 1$ .

- Remember  $P(X=x) = 0$  for a CRV.

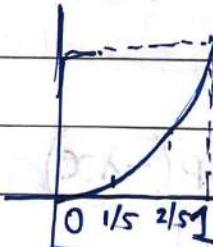
For rv  $X$  with PDF  $f_x$ , event  $A$  is a subset of real line and prob. is  $P(A) = \int_A f(x) dx$

Q. Consider  $f(x) = \begin{cases} 3x^2 & 0 < x < 1 \\ 0 & \text{otherwise} \end{cases}$

Show  $f$  is a density fn. Consider rv  $X$  with density  $f$ . Find  $P(X=1/5)$ ,  $P(X=2/5)$ ,  $P(X \in [1/5 - \epsilon, 1/5 + \epsilon])$ .

$$P(X=1/5) = 0$$

$$P(X=2/5) = 0$$



$$F(1/5) = 3/25$$

$$f(2/5) = 12/25$$

$$P(1/5 - \epsilon, 1/5 + \epsilon) = \int_{1/5 - \epsilon}^{1/5 + \epsilon} 3x^2 dx = x^3 \Big|_{1/5 - \epsilon}^{1/5 + \epsilon}$$

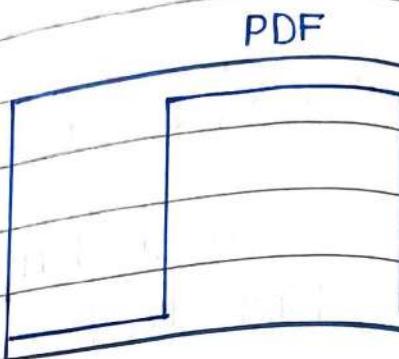
$$= \left(\frac{1}{5} + \epsilon\right)^3 - \left(\frac{1}{5} - \epsilon\right)^3 = \frac{6}{25}\epsilon + 2\epsilon^3 \ll \epsilon$$

## ★ Common Distributions

$X \sim \text{Uniform } [a, b]$

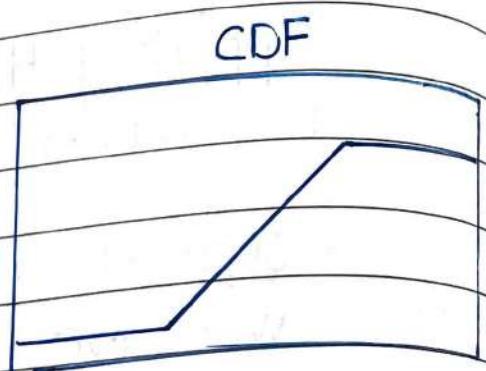
• PDF

$$f_x(x) = \begin{cases} \frac{1}{b-a} & a < x < b \\ 0 & \text{otherwise} \end{cases}$$



• CDF

$$F_x(x) = \begin{cases} 0 & x \leq a \\ \frac{x-a}{b-a} & a < x < b \\ 1 & x \geq b \end{cases}$$



Suppose  $X \sim \text{Uniform } [-10, 10]$ . Find  $P(-3 \leq X \leq 2)$ ,  $P(5 < |X| < 7)$ ,  $P(1 - \epsilon < X < 1 + \epsilon)$ ,  $P(X > 7 | X > 3)$

$$f_x(x) = \begin{cases} 1/20 & -10 < x < 10 \\ 0 & \text{otherwise} \end{cases}$$

$\uparrow 5/20$

$\downarrow 4/20$

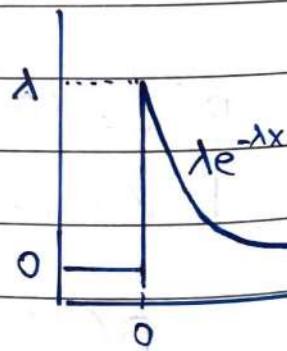
$\downarrow 2\epsilon/20 \text{ for } x_0 \in [-9, 9]$

$\downarrow 3/7$

$X \sim \text{Exp}(\lambda)$

• PDF

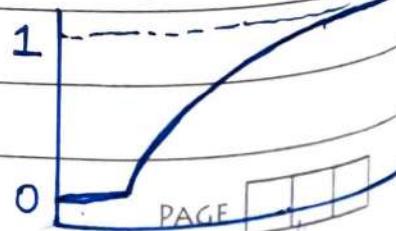
$$f_x(x) = \begin{cases} \lambda e^{-\lambda x} & x > 0 \\ 0 & \text{otherwise} \end{cases}$$



• CDF

$$F_x(x) = \begin{cases} 0 & x \leq 0 \\ 1 - \exp(-\lambda x) & x > 0 \end{cases}$$

classmate



$$X \sim \text{Exp}(2)$$

$$P(1-\varepsilon < X < 1+\varepsilon) = e^{-2(1-\varepsilon)} - e^{-2(1+\varepsilon)}$$

$P(X > s+t | X > s)$  =  $e^{-t}$  → Unique for exp. distribution.  
"memoryless"

## Normal Distribution / Gaussian Distribution

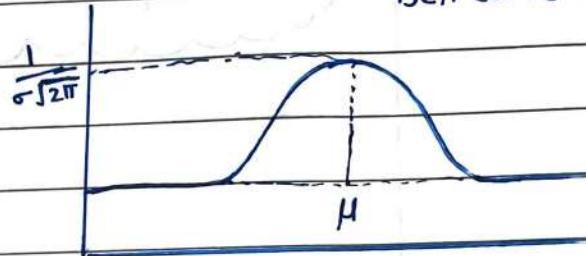
$$X \sim \text{Normal}(\mu, \sigma^2)$$

$\mu \rightarrow$  any real num  
 $\sigma \rightarrow$  +ve real num

"Bell curve"

- PDF  $\text{supp}(X) = \mathbb{R}$

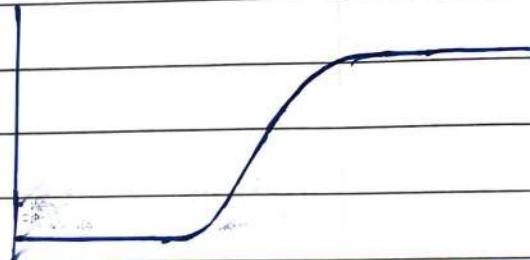
$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$



- CDF

$$F_X(x) = \int_{-\infty}^x f_X(u) du.$$

↓  
no closed form



Standard Normal :  $\text{Normal}(0,1)$

- CDF of  $X \sim \text{Normal}(\mu, \sigma^2)$  does not have closed form.
- Standardisation: If  $X \sim \text{Normal}(\mu, \sigma^2)$  then: →  
 $(X - \mu/\sigma) \sim \text{Normal}(0,1)$

$$Z \sim \text{Normal}(0,1) \quad \text{PDF: } f_Z(z) = \frac{1}{\sqrt{2\pi}} \exp(-z^2/2)$$

$$\text{CDF: } F_Z(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} \exp(-u^2/2) du$$

Normal tables: Tabulation of  $F_Z(z)$   
So convert prob. to a standard normal. Use normal tables.



$$X \sim \text{Normal}(\mu, \sigma^2)$$

$$Z = X - \mu \sim \text{Normal}(0, 1) \quad X = Z + \mu$$

$$\begin{aligned} P(5 < X < 7) &\leftrightarrow 5 - \mu < Z + \mu - \mu < 7 - \mu \Rightarrow 2 < Z < 4 \\ &= P(2 < Z < 4) = F_Z(4) - F_Z(2) \end{aligned}$$

$$P(X > \mu) = P(X < \mu) = 1/2$$