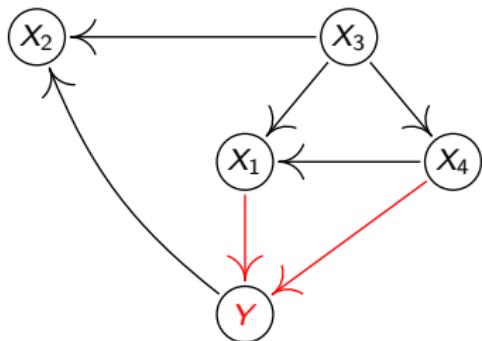


# Causality



Jonas Peters

University of Copenhagen

LxMLS, Virtual, 21.7.–29.7.2020



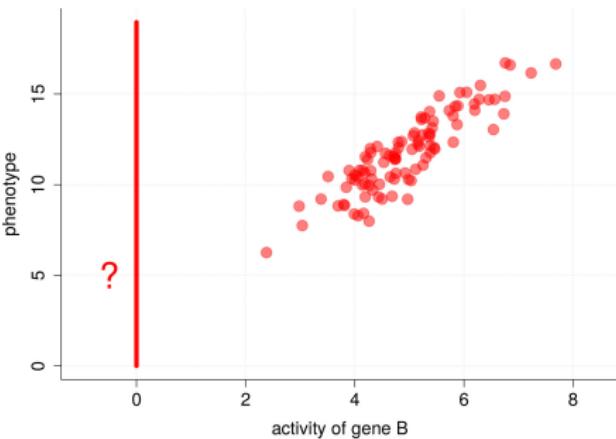
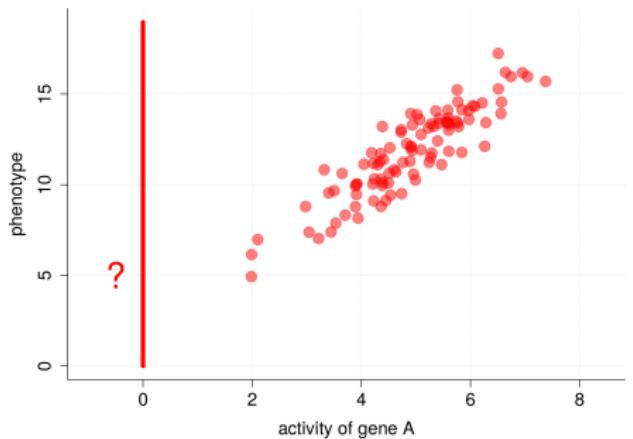
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Peters et al.: Elements of Causal Inference, MIT Press 2017.

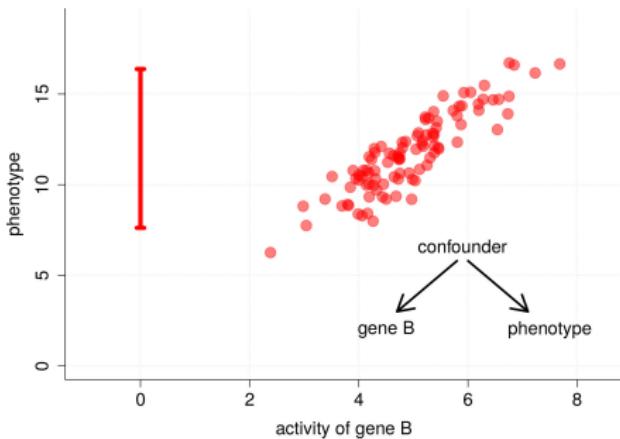
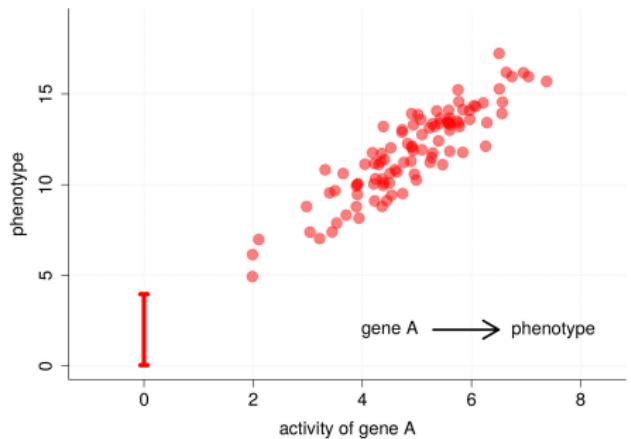
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Peters et al.: Elements of Causal Inference, MIT Press 2017.
- The presentation is biased. In particular, there is little statistics and nothing about potential outcomes. Good books include  
Hernan & Robins: Causal Inference, Chapman & Hall/CRC 2019,  
Imbens & Rubin: Causal Inference for Statistics, Cambridge Univ. Press 2015,  
Pearl: Causality, Cambridge Univ. Press 2009,  
... and others.

# Consider the following problem.



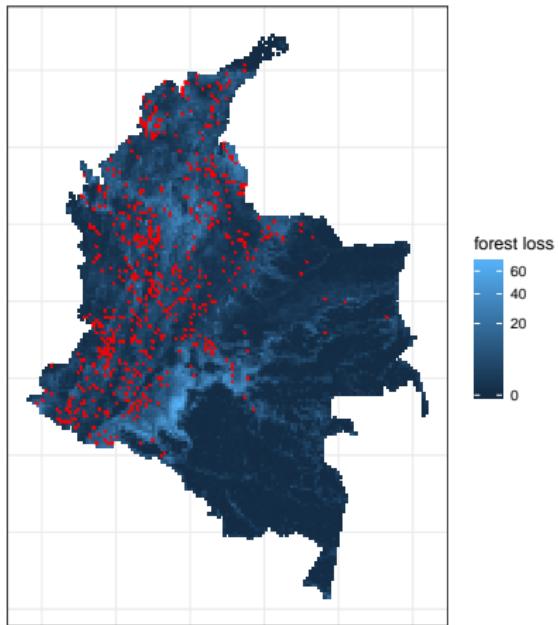
# Causality matters!



# Example: Forest Loss

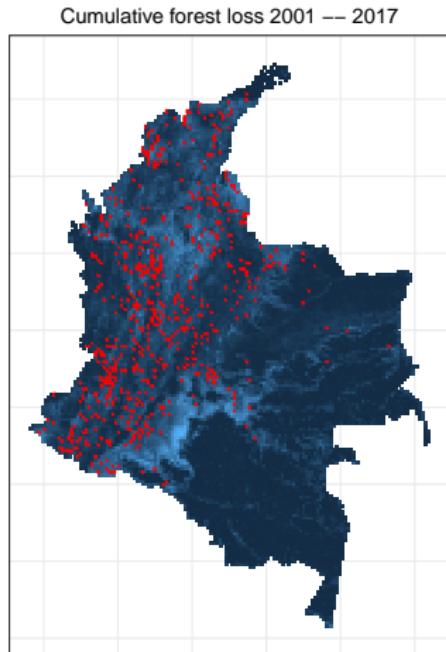
Colombia

Cumulative forest loss 2001 -- 2017

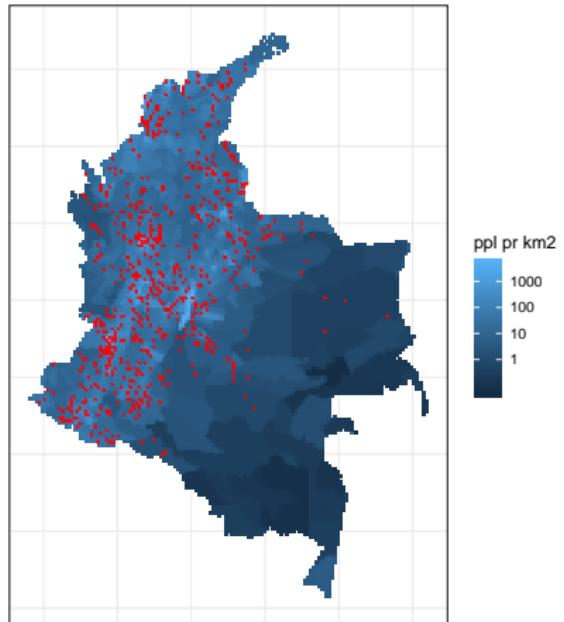


# Example: Forest Loss

Colombia



Population density in 2000



Christiansen et al. 2019

## Example: kidney stones

	Treatment A	Treatment B
	$\frac{273}{350} = 0.78$	$\frac{289}{350} = 0.83$
		$\frac{562}{700} = 0.80$

Charig et al.: *Comparison of treatment of renal calculi by open surgery, (...)*, British Medical Journal, 1986

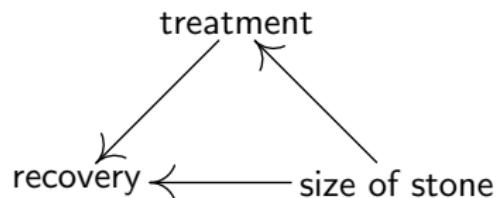
## Example: kidney stones

	Treatment A	Treatment B
Small Stones ( $\frac{357}{700} = 0.51$ )	$\frac{81}{87} = 0.93$	$\frac{234}{270} = 0.87$
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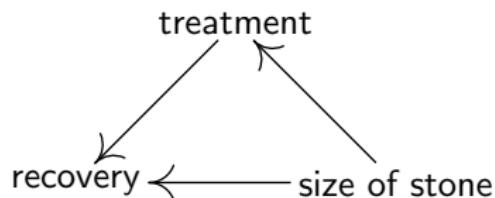
# Example: kidney stones

underlying ground truth:



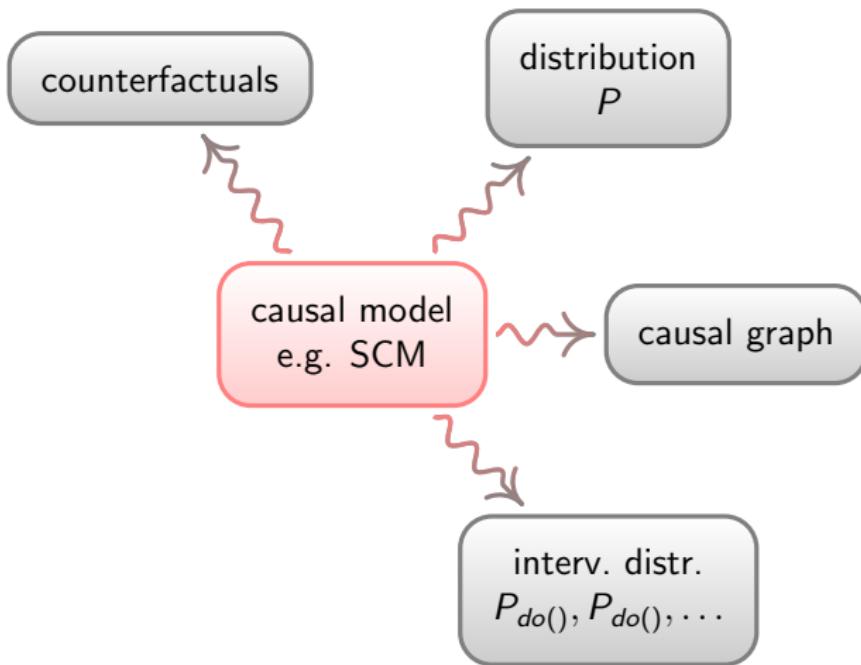
# Example: kidney stones

underlying ground truth:



What is the expected recovery if all get treatment A?

# What is a causal model?



## **Part I: Causal Models**

## Example: Two variables

SCMs model observational distributions.

$$X := N_x$$

$$Y := -6X + N_Y$$

$$N_X, N_Y \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1)$$



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$$X := N_x$$

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$$P : \quad (X, Y) \sim \mathcal{N} \left( \begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 1 & -6 \\ -6 & 37 \end{pmatrix} \right)$$

## Example: Two variables

SCMs model interventions, too.

$$X := N_X \quad X := 3$$

$$Y := -6X + N_Y$$

$$N_X, N_Y \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1)$$



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$$P_{do(X:=3)} : \quad P_{do(X:=3)}(X = 3) = 1 \quad \text{and} \quad Y \sim \mathcal{N}(-18, 1)$$

## Example: Two variables

SCMs model interventions, too.

$$X := N_x$$

$$Y := -6X + N_Y \quad Y := \mathcal{N}(2, 2)$$

$$N_X, N_Y \stackrel{\text{iid}}{\sim} \mathcal{N}(0, 1)$$

altitude



temperature



## Example: Two variables

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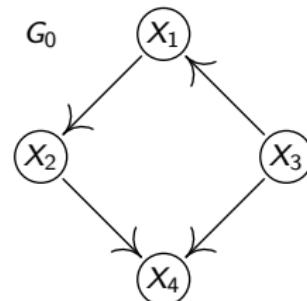


$$P_{do(Y:=\mathcal{N}(2,2))} : \quad (X, Y) \sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 2 \end{pmatrix}, \begin{pmatrix} 1 & 0 \\ 0 & 2 \end{pmatrix}\right)$$

SCMs model **observational distributions** over  $X_1, \dots, X_d$ . Call it:  $P$ .

$$\begin{aligned}X_1 &:= X_3 + N_1 \\X_2 &:= 2X_1 + N_2 \\X_3 &:= N_3 \\X_4 &:= -X_2 - X_3 + N_4\end{aligned}$$

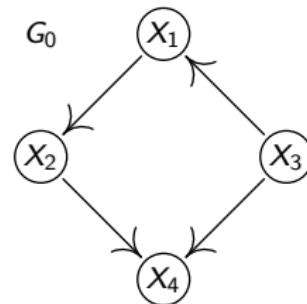
- $N_i$  jointly independent  $\mathcal{N}(0, 1)$
- $G_0$  has no cycles



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$$\begin{pmatrix} X_1 \\ X_2 \\ X_3 \\ X_4 \end{pmatrix} = \mathcal{N} \left( \begin{pmatrix} 0 \\ 0 \\ 0 \\ 0 \end{pmatrix}, \begin{pmatrix} 2 & 4 & 1 & -5 \\ 4 & 9 & 2 & -11 \\ 1 & 2 & 1 & -3 \\ -5 & -11 & -3 & 15 \end{pmatrix} \right)$$

SCMs model **interventions**, too. Call it:  $P_{do(X_1:=0)}$ .

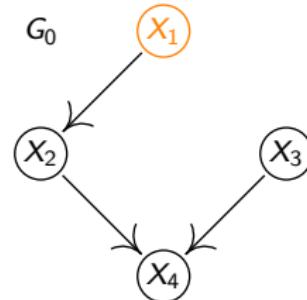
$$X_1 := 0$$

$$X_2 := f_2(X_1, N_2)$$

$$X_3 := f_3(N_3)$$

$$X_4 := f_4(X_2, X_3, N_4)$$

- $N_i$  jointly independent
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# Example: kidney stones

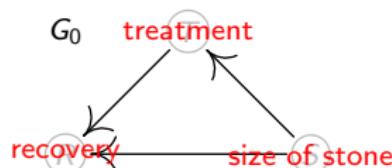
Given: graph and  $P$ , i.e., only the structure, not the functions.

$$T := f_1(S, N_1)$$

$$R := f_2(T, S, N_2)$$

$$S := f_3(N_3)$$

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# Example: kidney stones

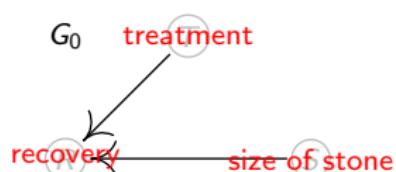
Given: graph and  $P$ . We want to compute  $P_{\text{do}(T:=A)}$ .

$$T := f_1(S, N_1) \quad T := A$$

$$R := f_2(T, S, N_2)$$

$$S := f_3(N_3)$$

- $N_i$  jointly independent
- $G_0$  has no cycles



IMPORTANT: modularity, autonomy: Aldrich 1989, Pearl 2009, Schölkopf et al. 2012, ...

If you intervene only on  $X_j$ , you intervene only on  $X_j$  (MUTE).

# Example: kidney stones

	Treatment A	Treatment B
Small Stones ( $\frac{357}{700} = 0.51$ )	$\frac{81}{87} = 0.93$	$\frac{234}{270} = 0.87$
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wanted:

$$\text{use: } P(R | S, T) \quad = \quad P_{do(T:=A)}(R | S, T)$$

$$P_{do(T:=A)}(R = 1)$$

## Example: kidney stones

$$\begin{aligned}E_{do(T:=A)}R &= P_{do(T:=A)}(R = 1) \\&= \sum_s P_{do(T:=A)}(R = 1, S = s, T = A) \\&= \sum_s P_{do(T:=A)}(R = 1 | S = s, T = A)P_{do(T:=A)}(S = s, T = A) \\&= \sum_s P_{do(T:=A)}(R = 1 | S = s, T = A)P_{do(T:=A)}(S = s) \\&= \sum_s P(R = 1 | S = s, T = A)P(S = s) \\&= 0.832 \\&> 0.782 \\&= \dots \\&= P_{do(T:=B)}(R = 1) = E_{do(T:=B)}R\end{aligned}$$

## Example: kidney stones

$$\begin{aligned}E_{do(T:=A)}R &= P_{do(T:=A)}(R = 1) \\&= \sum_s P_{do(T:=A)}(R = 1, S = s, T = A) \\&= \sum_s P_{do(T:=A)}(R = 1 | S = s, T = A)P_{do(T:=A)}(S = s, T = A) \\&= \sum_s P_{do(T:=A)}(R = 1 | S = s, T = A)P_{do(T:=A)}(S = s) \\&= \sum_s P(R = 1 | S = s, T = A)P(S = s) \\&= 0.832 \quad \neq P(R = 1 | T = A) \\&> 0.782 \\&= \dots \\&= P_{do(T:=B)}(R = 1) = E_{do(T:=B)}R\end{aligned}$$

This idea holds more generally.

## Definition

Given an SCM over  $(X, Y, W)$ . We call  $Z \subseteq W$  a valid adjustment set for  $(X, Y)$  if

$$p_{do(X:=x)}(y) = \sum_z p(y|x, z)p(z) \neq p(y|x)$$

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## Proposition (Parent Adjustment)

Assume  $Y \notin PA(X)$ . Then

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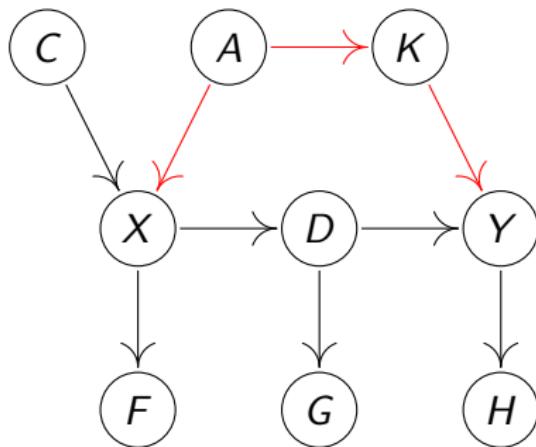
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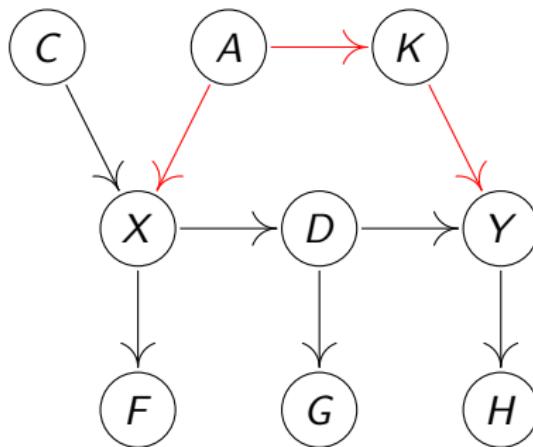
In particular, if  $\emptyset$  is valid adjustment set, then

$$p_{do(X:=x)}(y) = p(y|x).$$

## Adjusting in Linear Gaussian Models



$X \leftarrow A \rightarrow K \rightarrow Y$  is a “backdoor path” from  $X$  to  $Y$ .



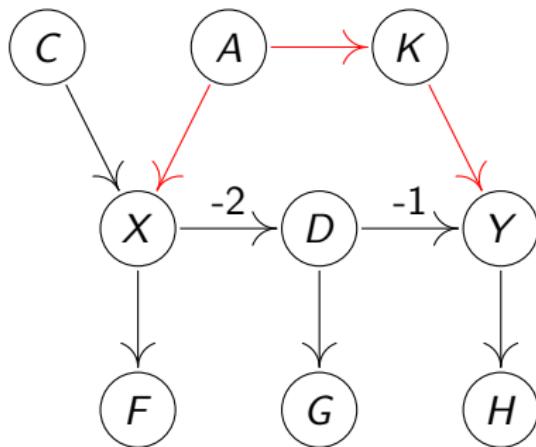
$X \leftarrow A \rightarrow K \rightarrow Y$  is a “backdoor path” from  $X$  to  $Y$ .

$$Z = \{C, A\},$$

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$$Z = \{F, C, K\}$$

are valid adjustment sets for  $(X, Y)$  (no proof).



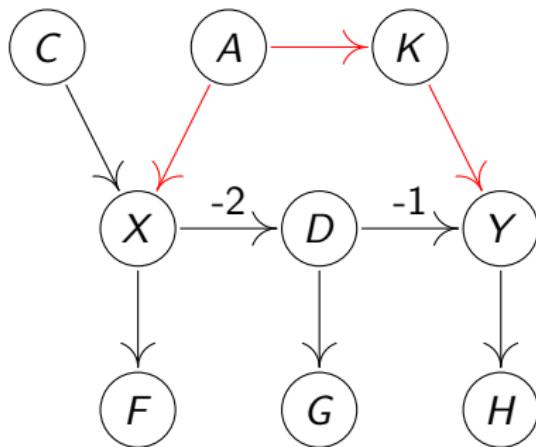
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are valid adjustment sets for  $(X, Y)$  (no proof).

Thus (no proof):  $\text{Im}(Y \sim X + K)$  yields consistent estimator for

$$\frac{\partial}{\partial x} E_{do(X:=x)} Y = (-2) \cdot (-1) = 2.$$

```
1 n <- 500
2
3 # generate a sample from the distr. ent. by the SCM
4 set.seed(1)
5 C <- rnorm(n)
6 A <- 0.8*rnorm(n)
7 K <- A + 0.1*rnorm(n)
8 X <- C - 2*A + 0.2*rnorm(n)
9 F <- 3*X + 0.8*rnorm(n)
10 D <- -2*X + 0.5*rnorm(n)
11 G <- D + 0.5*rnorm(n)
12 Y <- 2*K - D + 0.2*rnorm(n)
13 H <- 0.5*Y + 0.1*rnorm(n)
14
15 lm(Y~X)$coefficients
16 lm(Y~X+K)$coefficients
17 lm(Y~X+F+C+K)$coefficients
18 lm(Y~X+F+C+K+H)$coefficients
```

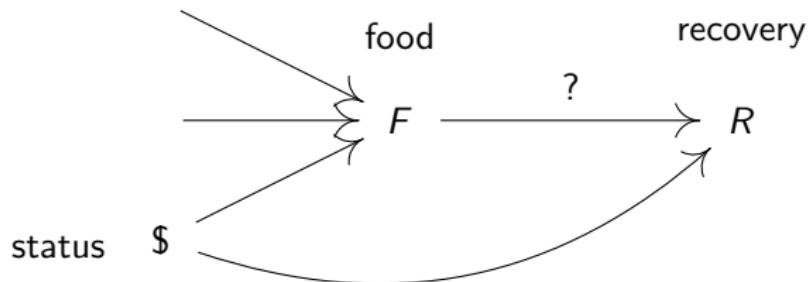
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```

Do not simply throw in as many variables as possible.



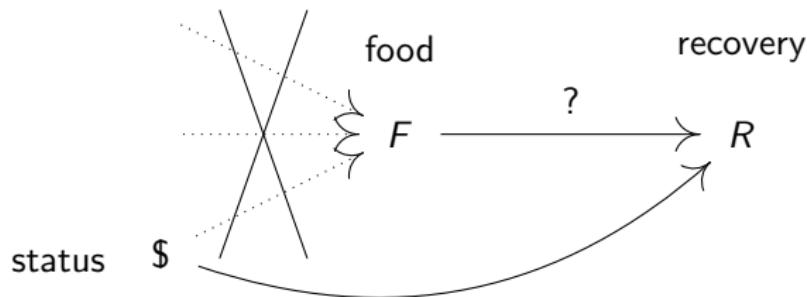
## **James Lind (1716–94):**

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Causal relationship unclear.



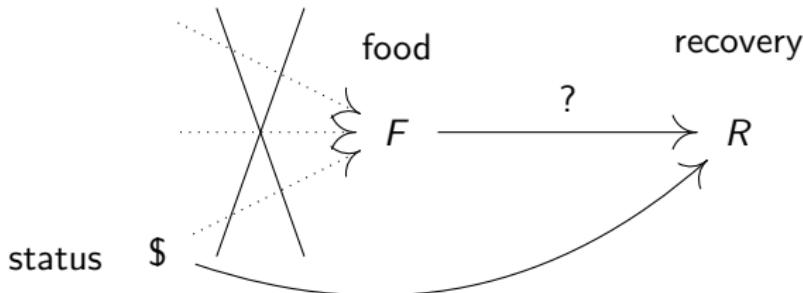
# James Lind (1716–94):

Randomize!  $F$  and  $R$  dependent  $\implies$  there is a causal link!



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Randomize!  $F$  and  $R$  dependent  $\implies$  there is a causal link!



"On the 20th of May 1747, I selected twelve patients in the scurvy, on board the Salisbury [...] Two were ordered each a quart of cyder a day. Two others took twenty-five drops of elixir vitriol three times a day [...] Two others took two spoonfuls of vinegar three times a day [...] Two of the worst patients were put on a course of sea-water [...] Two others had each two oranges and one lemon given them every day [...] The two remaining patients, took [...] an electuary recommended by a [...] surgeon [...] The consequence was, that the most sudden and visible good effects were perceived from the use of oranges and lemons;"

## Example: smoking

# BRITISH MEDICAL JOURNAL

LONDON SATURDAY SEPTEMBER 30 1950

---

## SMOKING AND CARCINOMA OF THE LUNG PRELIMINARY REPORT

BY

**RICHARD DOLL, M.D., M.R.C.P.**

*Member of the Statistical Research Unit of the Medical Research Council*

AND

**A. BRADFORD HILL, Ph.D., D.Sc.**

*Professor of Medical Statistics, London School of Hygiene and Tropical Medicine; Honorary Director of the Statistical Research Unit of the Medical Research Council*

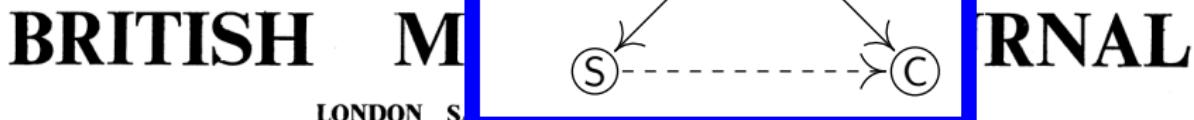
In England and Wales the phenomenal increase in the number of deaths attributed to cancer of the lung provides one of the most striking changes in the pattern of mortality recorded by the Registrar-General. For example, in the quarter of a century between 1922 and 1947 the annual number of deaths recorded increased from 612 to

whole explanation, although no one would deny that it may well have been contributory. As a corollary, it is right and proper to seek for other causes.

### Possible Causes of the Increase

Two main causes have from time to time been put for-

# Example: smoking



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"One of the most important books of the year . . .  
What it has to say needs to be heard." —The Christian Science Monitor

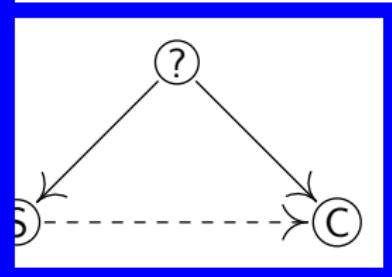
The book that inspired the film  
**MERCHANTS OF DOUBT**

# Merchants of DOUBT



How a Handful of Scientists Obscured  
the Truth on Issues from  
Tobacco Smoke to Global Warming

NAOMI ORESKES  
& ERIK M. CONWAY



# JOURNAL

## NOMA OF THE LUNG SYMPOSIUM REPORT

BY

**L, M.D., M.R.C.P.**

*Unit of the Medical Research Council*

AND

**HILL, Ph.D., D.Sc.**

*Head of Tropical Medicine; Honorary Director of the Statistical  
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whole explanation, although no one would deny that it may well have been contributory. As a corollary, it is right and proper to seek for other causes.

### Possible Causes of the Increase

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## Definition (Equivalence of causal models)

Two models are called

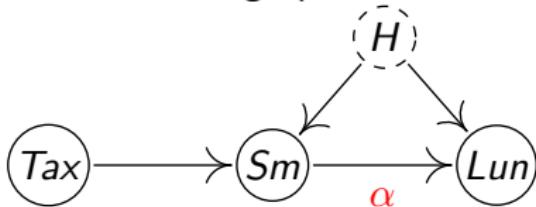
{**probabilistically / interventionally**} equivalent

if they entail the same

{observational / observational & interventional}

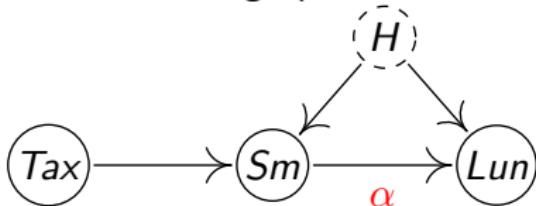
distributions. Here, it suffices to consider interventions that set a variable  $X_j$  to a fully supported  $\tilde{N}_j$  ("randomized experiments").

Consider this graph



$$Lun = \alpha Sm + \beta H + N$$

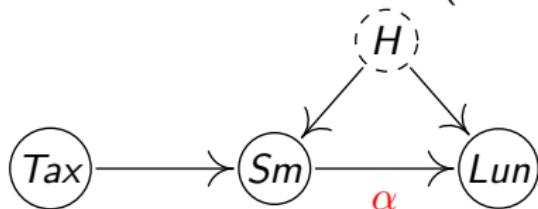
Consider this graph



$$Lun = \alpha Sm + \beta H + N$$



An **instrumental variable** (here: tax) can fix the problem!



$$Lun = \alpha Sm + \beta H + N$$



## Summary Part I:

- What if interested in iid prediction, i.e., **observational data**? Don't worry (too much) about causality!

## Summary Part I:

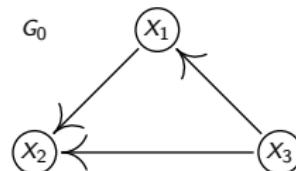
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- SCMs entail graphs, obs. distr., interventions and counterfactuals.

$$\begin{aligned}X_1 &:= f_1(X_3, N_1) \\X_2 &:= f_2(X_1, X_3, N_2) \\X_3 &:= f_3(N_3)\end{aligned}$$

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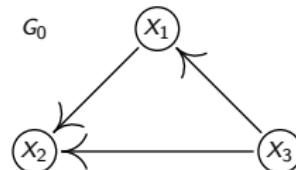


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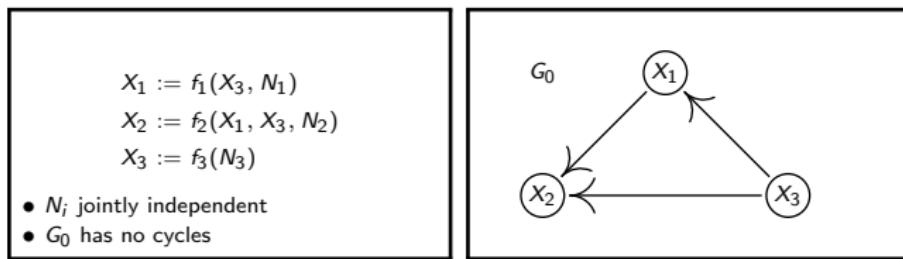
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- **adjusting: graph + observational distribution  $\rightsquigarrow$  interventions**  
ComputeInterventions.ipynb (skip Exercise 1 and maybe Exercise on d-sep.)

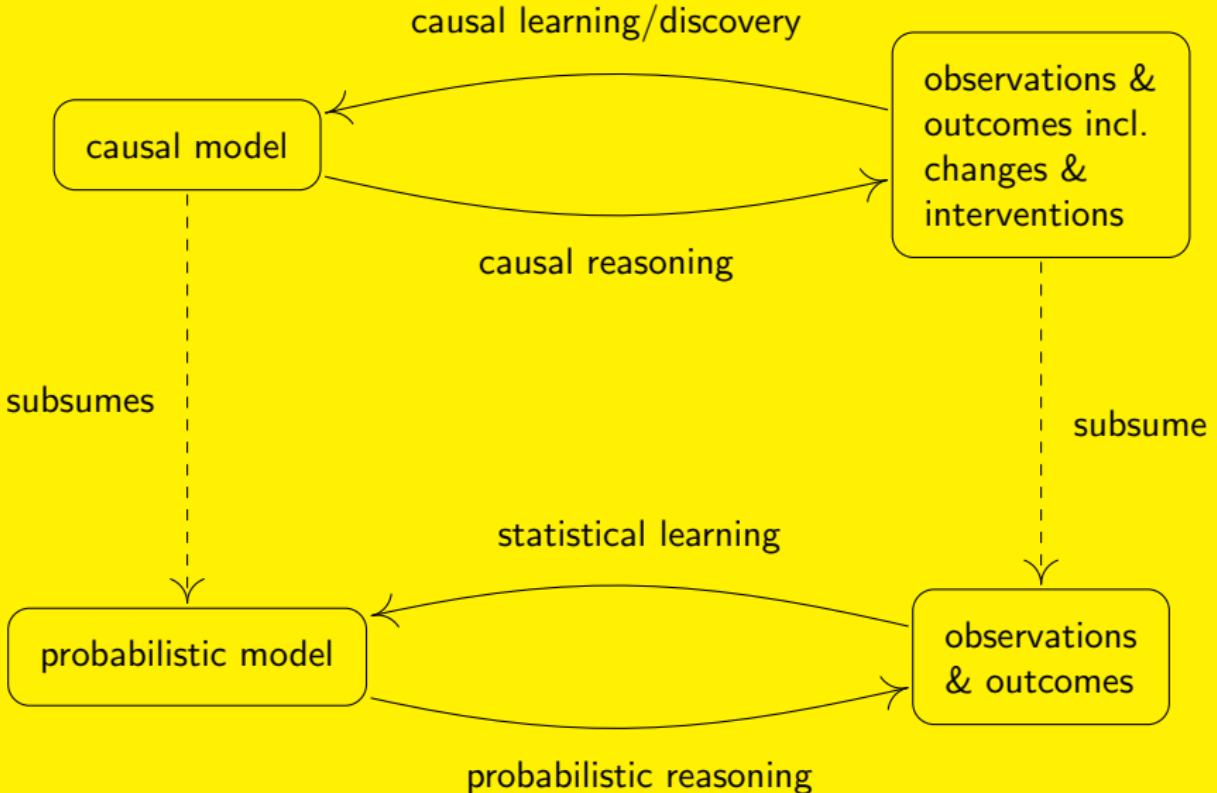
## Summary Part I:

- What if interested in iid prediction, i.e., **observational data**? Don't worry (too much) about causality!
- But often, we are interested in a system's behaviour **under intervention**.
- SCMs entail graphs, obs. distr., interventions and counterfactuals.



- **adjusting: graph + observational distribution  $\rightsquigarrow$  interventions**  
ComputeInterventions.ipynb (skip Exercise 1 and maybe Exercise on d-sep.)
- **instrumental variables: may help if there are hidden variables**  
InstrumentalVariables.ipynb (skip Exercise 1)

## **Part II: Structure Learning or Causal Discovery**



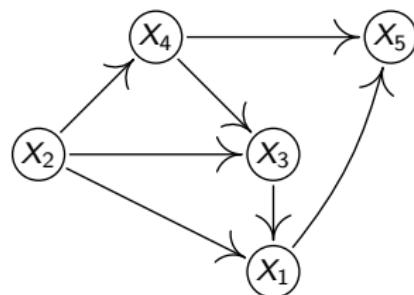
## The Problem of Causal Discovery:

observed iid data  
from  $P(X_1, \dots, X_5)$



causal model, e.g. DAG  $\mathcal{G}$

$X_1$	$X_2$	$X_3$	$X_4$	$X_5$
3.4	-0.3	5.8	-2.1	2.2
1.7	-0.2	7.0	-1.2	0.4
-2.4	-0.1	4.3	-0.7	3.5
2.3	-0.3	5.5	-1.1	-4.4
3.5	-0.2	3.9	-0.9	-3.9
⋮	⋮	⋮	⋮	⋮



Correlation (Dependence) does not imply causation

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### **Reichenbach's common cause principle.**

Assume that  $X \perp\!\!\!\perp Y$ . Then

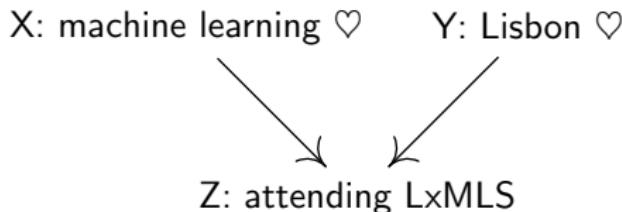
- $X$  “causes”  $Y$ ,
- $Y$  “causes”  $X$ ,
- there is a hidden common “cause” or
- combination of the above.

Correlation (Dependence) does not imply causation ... but:

### Reichenbach's common cause principle.

Assume that  $X \not\perp\!\!\!\perp Y$ . Then

- $X$  “causes”  $Y$ ,
- $Y$  “causes”  $X$ ,
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- combination of the above.
- (In practice implicit conditioning also happens:



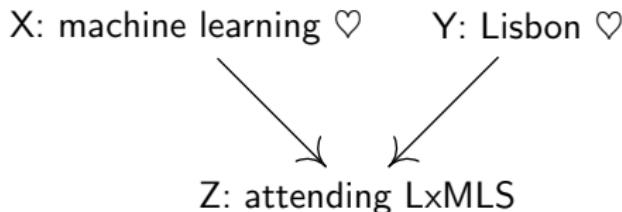
aka “selection bias”).

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Assume that  $X \perp\!\!\!\perp Y$ . Then

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- $Y$  “causes”  $X$ ,
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- (In practice implicit conditioning also happens:



aka “selection bias”). Formalization of this idea...

## Definition

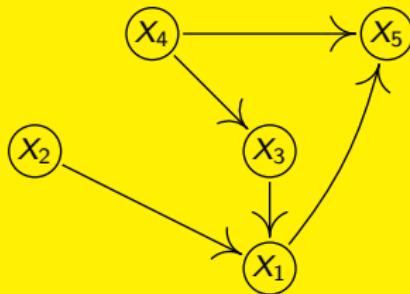
$P$  is Markov w.r.t.  $G$  if

$$\underbrace{X \text{ and } Y \text{ are } d\text{-separated by } \mathcal{S} \text{ in } G}_{\text{properties of graph}} \Rightarrow \underbrace{X \perp Y | \mathcal{S}}_{\text{properties in } P}$$

# Definition: graphs

$G = (V, E)$  with  $E \subseteq V \times V$ . The rest is as in real life!

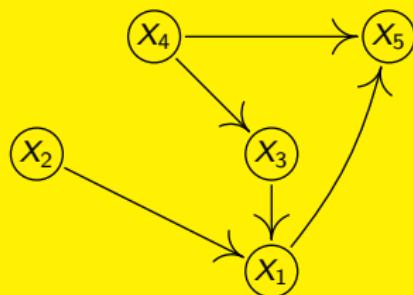
- parents, children, descendants, ancestors, ...
- paths, directed paths
- immoralities (or v-structures)
- $d$ -separation (see next)
- ...



## Definition: $d$ -separation

$X_i$  and  $X_j$  are  $d$ -separated by  $\mathcal{S}$  if all paths between  $X_i$  and  $X_j$  are blocked by  $\mathcal{S}$ .

Check, whether all paths blocked!!



$X_2$  and  $X_5$  are  $d$ -sep. by  $\{X_1, X_4\}$

$X_4$  and  $X_1$  are  $d$ -sep. by  $\{X_2, X_3\}$

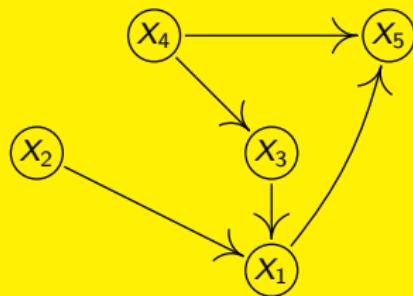
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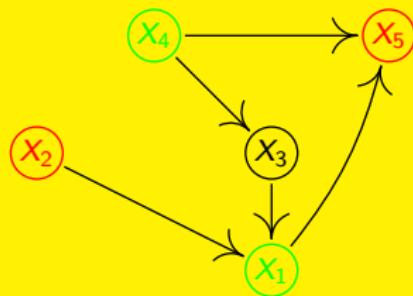
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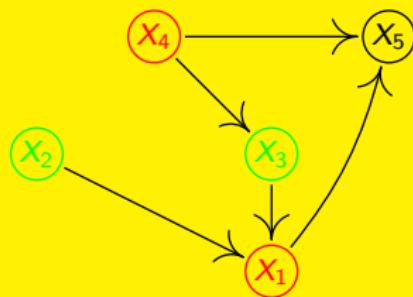
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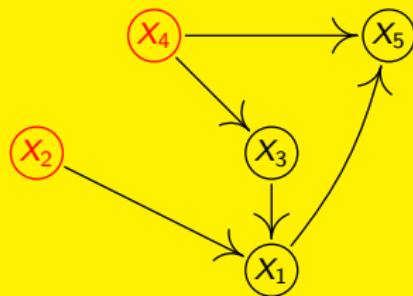
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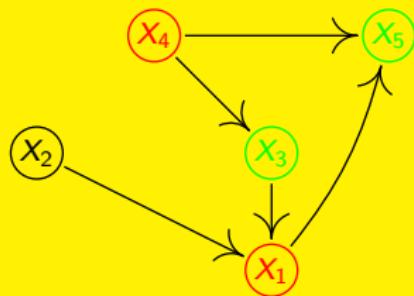
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## Definition

$P$  satisfies the (global) Markov condition w.r.t.  $G$  if

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## Proposition

Let the distribution  $P$  be Markov wrt a causal graph  $G$ . Then, Reichenbach's common cause principle is satisfied.

Proof: dependent variables must be  $d$ -connected.

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## Definition

$P$  satisfies faithfulness w.r.t.  $G$  if

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# Idea 1: independence-based methods

Exercise:

- a) Assume  $P^{(X,Y,Z)}$  is Markov and faithful wrt.  $G$ . Assume all(!) conditional independences are

$$X \perp\!\!\!\perp Z | \emptyset$$

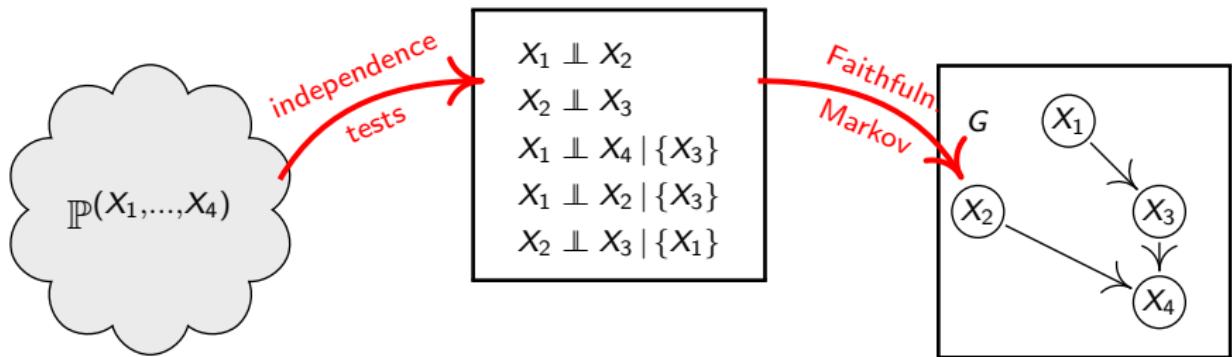
(plus symmetric statements). What is  $G$ ?

- b) Assume  $P^{(W,X,Y,Z)}$  is Markov and faithful wrt.  $G$ . Assume all(!) conditional independences are

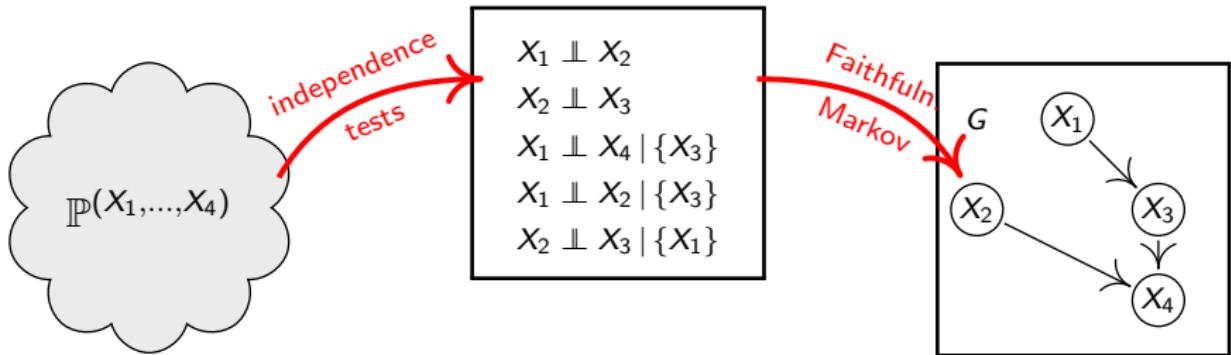
$$\begin{aligned} (Y, Z) &\perp\!\!\!\perp W | \emptyset \\ W &\perp\!\!\!\perp Y | (X, Z) \\ (W, X) &\perp\!\!\!\perp Y | Z \end{aligned}$$

(plus symmetric and trivially implied statements). What is  $G$ ?

# Idea 1: independence-based methods



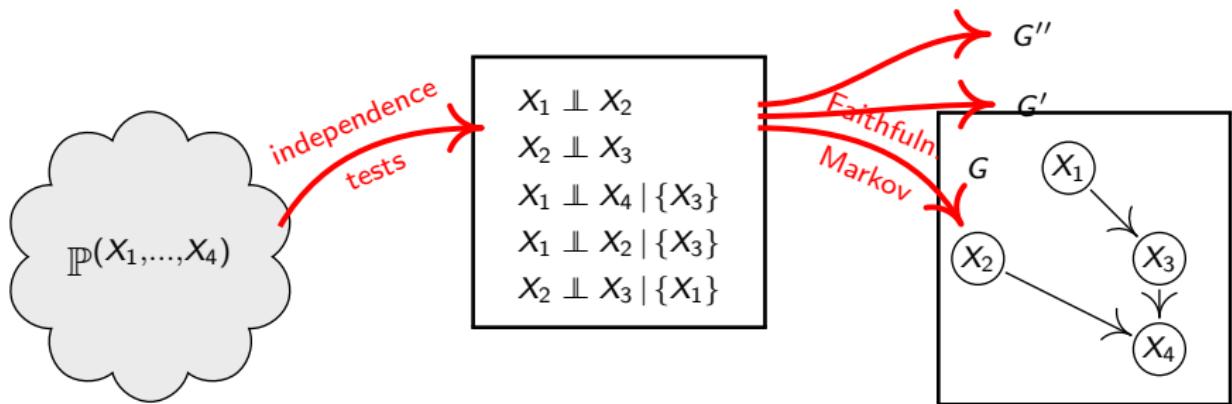
# Idea 1: independence-based methods



Method: IC (Pearl 2009); PC, FCI (Spirtes et al., 2000)

- ① Find all (cond.) independences from the data.
- ② Select the DAG(s) that corresponds to these independences.

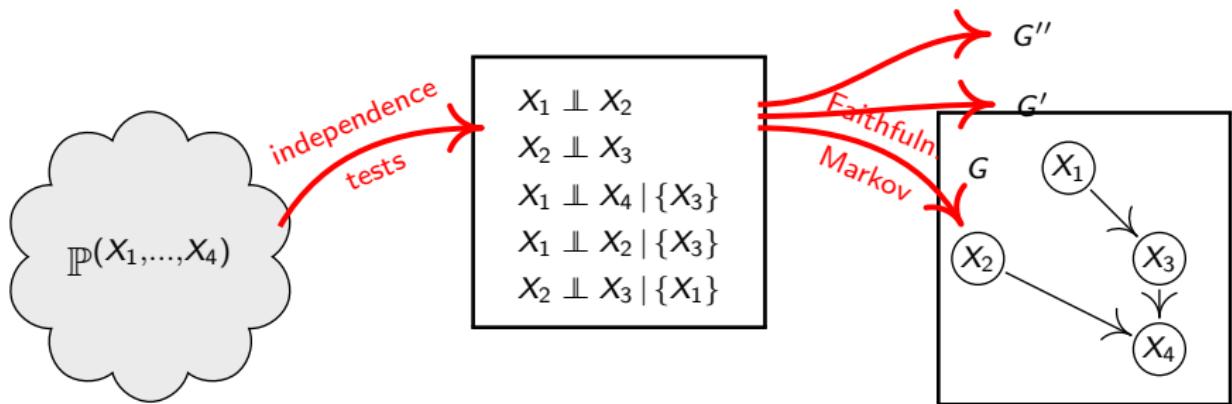
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# Idea 1: independence-based methods



Method: IC (Pearl 2009); PC, FCI (Spirtes et al., 2000)

- ① Find all (cond.) independences from the data. Be smart.
- ② Select the DAG(s) that corresponds to these independences.



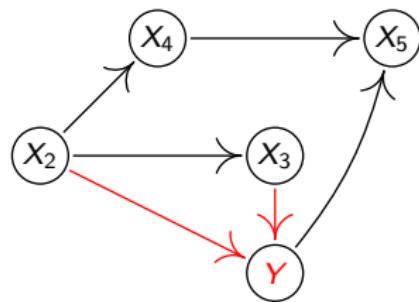
# The Problem of Causal Discovery:

observed data

$Y$	$X_2$	$X_3$	$X_4$	$X_5$
3.4	-0.3	5.8	-2.1	2.2
1.7	-0.2	7.0	-1.2	0.4
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:	:	:	:	:

?

causal model



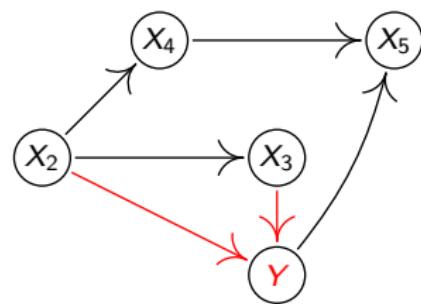
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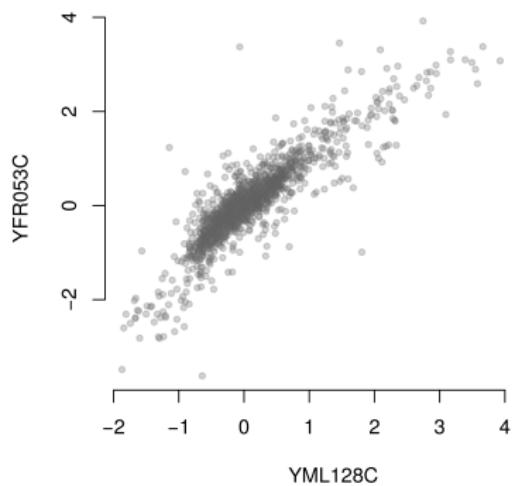
?  
→

causal model



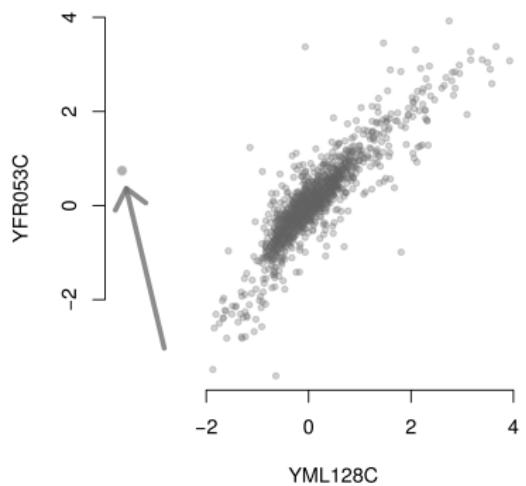
Here: Find the direct causes of  $Y$ !

Choose the predictor with the strongest correlation...



data from: Kemmeren et al. 2014

...and check the corresponding intervention on that predictor:



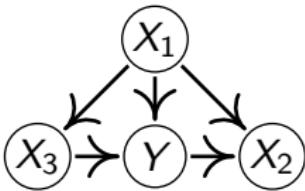
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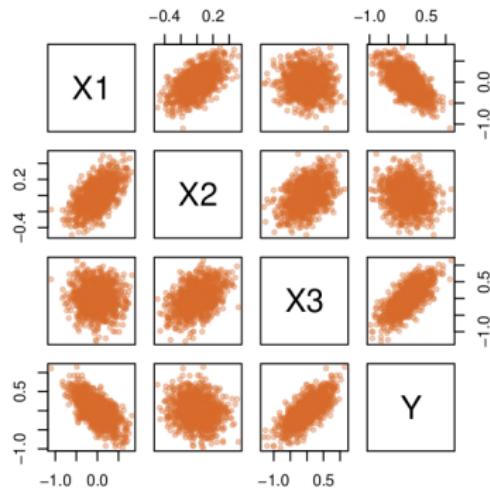
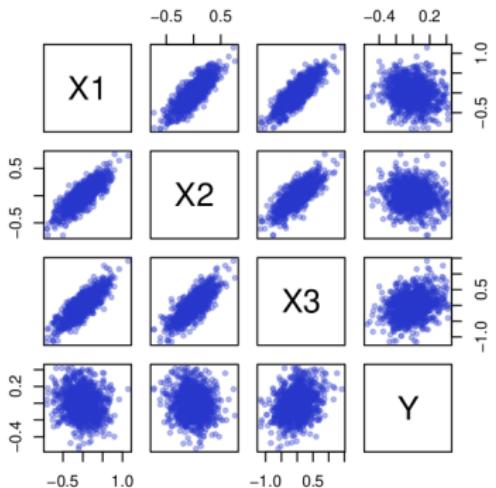


Invariant Causal Prediction

unknown:



known:



## linear model

```
> linmod <- lm( Y ~ X)
> summary(linmod)
```

Coefficients:

	Estimate	Std. Error	t value	Pr(> t )
(Intercept)	8.305e-05	2.067e-03	0.04	0.968
X1	-5.490e-01	9.725e-03	-56.46	<2e-16 ***
X2	-4.078e-01	1.810e-02	-22.52	<2e-16 ***
X3	6.821e-01	6.896e-03	98.91	<2e-16 ***

## ICP (R-package InvariantCausalPrediction)

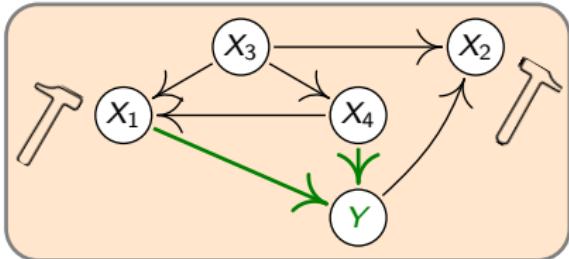
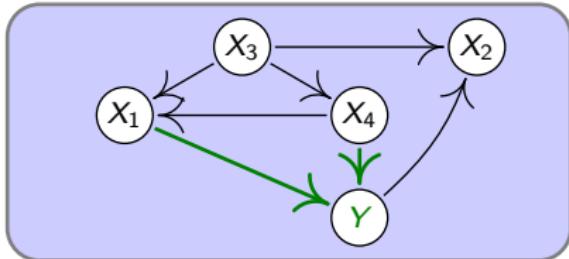
```
> ExpInd
```

```
[1]1111111111111111111111111111111111111111111111111111111111111111...2222222222222222...
```

```
> icp <- ICP(X,Y,ExpInd)
```

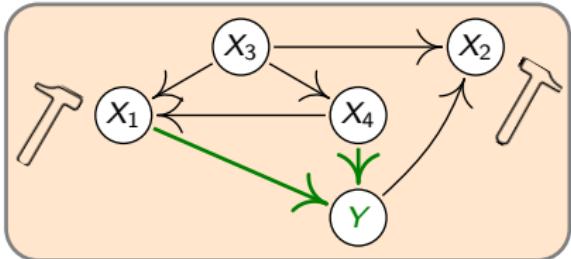
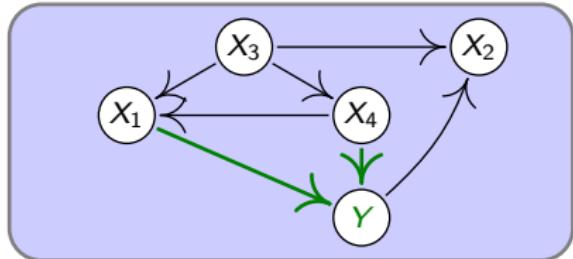
	LOWER BOUND	UPPER BOUND	MAXIMIN	EFFECT	P-VALUE
X1	-0.71	-0.52		-0.52	<1e-09 ***
X2	-0.46	0.00		0.00	0.55
X3	0.58	0.70		0.58	<1e-09 ***
<hr/>					
Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1					

**Fundamental assumption:**  $X_1, X_4 \rightarrow Y$  is invariant under interventions.



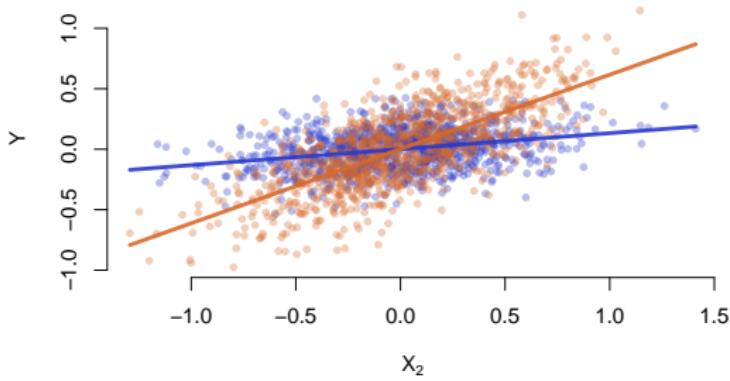
cf. modularity, autonomy, Haavelmo 1944, Aldrich 1989, Pearl 2009, ...

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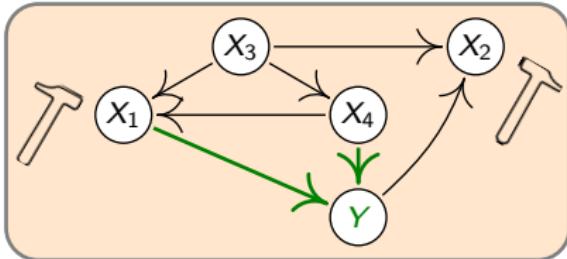
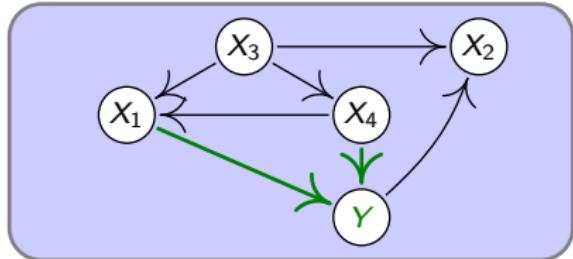


cf. modularity, autonomy, Haavelmo 1944, Aldrich 1989, Pearl 2009, ...

Not all sets of predictors yield an invariant model. Here: {2}.



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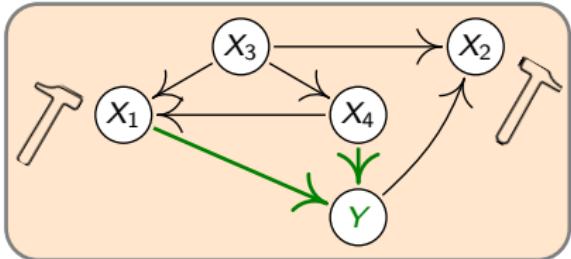
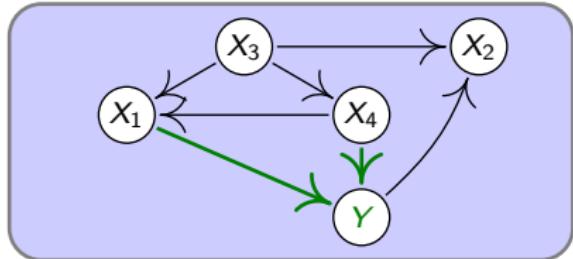
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**Key idea:** Use and data and search for invariant models.

set	$\emptyset$	$\{1\}$	$\{2\}$	$\{3\}$	$\dots$	$\{1, 4\}$	$\{2, 4\}$	$\dots$	$\{1, 3, 4\}$
invariance	$\times$	$\times$	$\times$	$\times$	$\dots$	$\checkmark$	$\times$	$\dots$	$\checkmark$

$$\hat{S} := \bigcap_{S \text{ invariant}} S = \{1, 4\}$$

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$$\hat{S} := \bigcap_{S \text{ invariant}} S = \{1, 4\}$$

JP, Bühlmann, Meinshausen, JRSS-B 2016 (with discussion):  $P(\hat{S} \subseteq S^*) \geq 1 - \alpha$ . (ICP.ipynb)

**Given**  $(X_1, Y_1), \dots, (X_n, Y_n) \in \mathbb{R}^d \times \mathbb{R}$  and environments  $\mathcal{E}$ .

**Invariance**  $H_{0,S}$ :

- for all  $i = 1, \dots, n$ :  $Y_i = X_{S,i} \cdot \gamma + \varepsilon_i$ .
- $\varepsilon_1, \dots, \varepsilon_n$  are i.i.d.
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**Environments**  $\mathcal{E}$  have elements

$e_1 = \{1, 2, 3, \dots, 40\}$ ,  $e_2 = \{41, \dots, 100\}$ ,  $e_3 = \{101, \dots, n\}$ , for example.

**Given**  $(X_1, Y_1), \dots, (X_n, Y_n) \in \mathbb{R}^d \times \mathbb{R}$  and environments  $\mathcal{E}$ .

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**Environments**  $\mathcal{E}$  have elements

$e_1 = \{1, 2, 3, \dots, 40\}$ ,  $e_2 = \{41, \dots, 100\}$ ,  $e_3 = \{101, \dots, n\}$ , for example.

**Relation to causality:**

Environments: different interventions (not on  $Y$ ). Then,  $H_{0,PA(Y)}$  holds.

cf. modularity, autonomy, Haavelmo 1944, Aldrich 1989, Pearl 2009, Schölkopf et al. 2012, ...

**Given**  $(X_1, Y_1), \dots, (X_n, Y_n) \in \mathbb{R}^d \times \mathbb{R}$  and environments  $\mathcal{E}$ .

**Invariance**  $H_{0,S}$ :

- for all  $i = 1, \dots, n$ :  $Y_i = X_{S,i} \cdot \gamma + \varepsilon_i$ .
- $\varepsilon_1, \dots, \varepsilon_n$  are i.i.d.
- $X_i$  can have an arbitrary distribution

**Environments**  $\mathcal{E}$  have elements

$e_1 = \{1, 2, 3, \dots, 40\}$ ,  $e_2 = \{41, \dots, 100\}$ ,  $e_3 = \{101, \dots, n\}$ , for example.

**Relation to causality:**

Environments: different interventions (not on  $Y$ ). Then,  $H_{0,PA(Y)}$  holds.

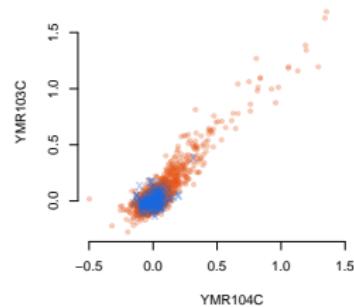
cf. modularity, autonomy, Haavelmo 1944, Aldrich 1989, Pearl 2009, Schölkopf et al. 2012, ...

**Theorem (PBM 2016)**

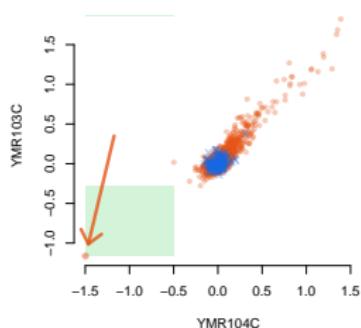
Assume  $H_{0,S^*}$  satisfied for some  $S^*$ . For any test level  $\alpha$  we obtain

$$P(\hat{S} \subseteq S^*) \geq 1 - \alpha.$$

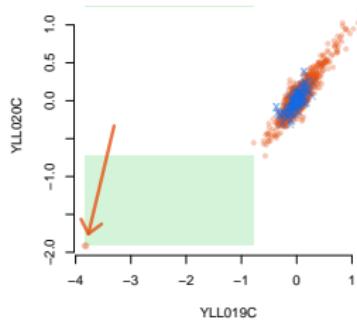
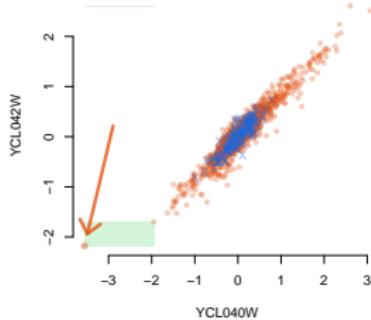
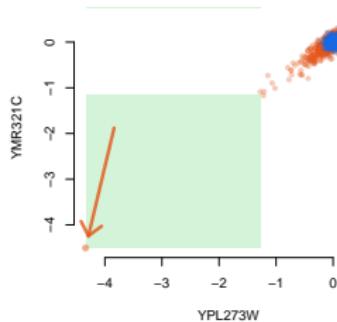
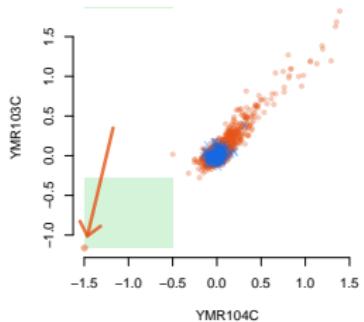
# Predictors that are inferred to be causal by ICP...



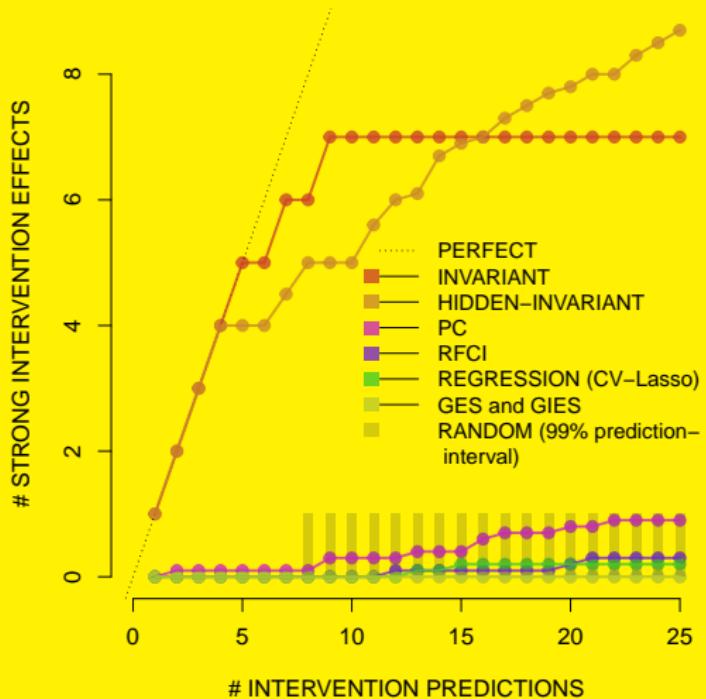
...and the corresponding interventions on the predictor



...and the corresponding interventions on the predictor



# Yeast data (Kemmeren et al., 2014)



So far: invariance with respect to



anchor = environments

Also possible:

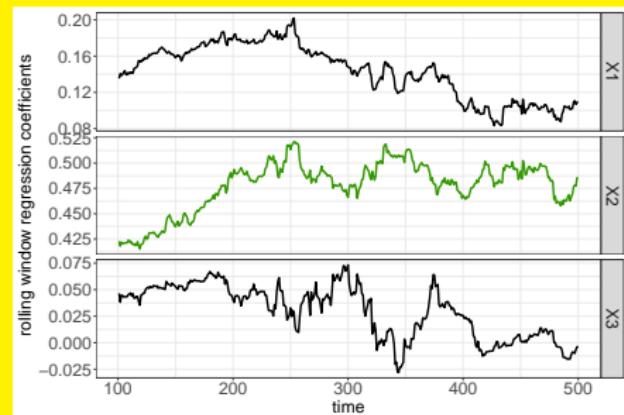
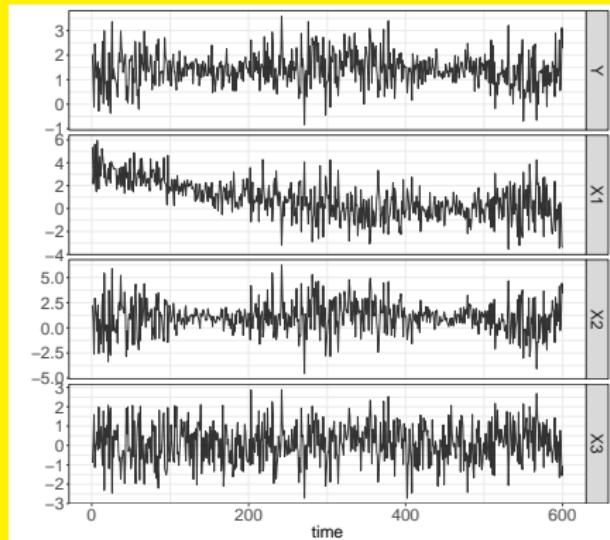


anchor = time

Suppose there is time.

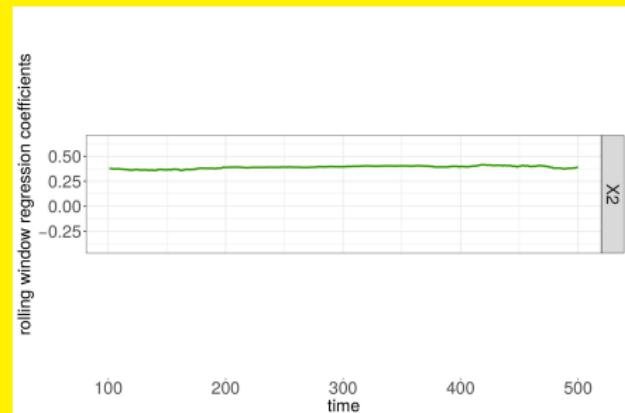
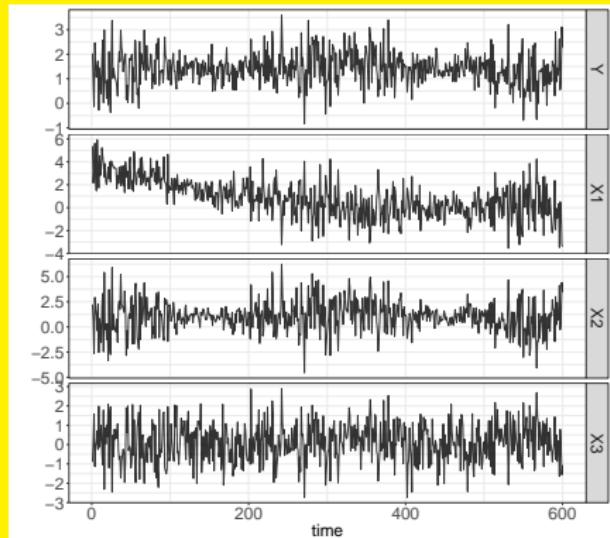
$t$	$X_1$	$X_2$	$X_3$	$X_4$	$Y$
1	3.4	-0.3	5.8	-2.1	2.2
2	1.7	-0.2	7.0	-1.2	0.4
3	-2.4	-0.1	4.3	-0.7	3.5
4	2.3	-0.3	5.5	-1.1	-4.4
5	3.5	-0.2	3.9	-0.9	-3.9
:	:	:	:	:	:

Regressing on  $(X_1, X_2, X_3)$ :

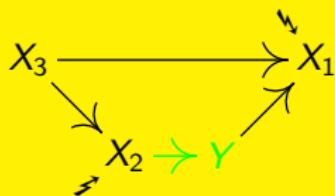


The coefficients change.

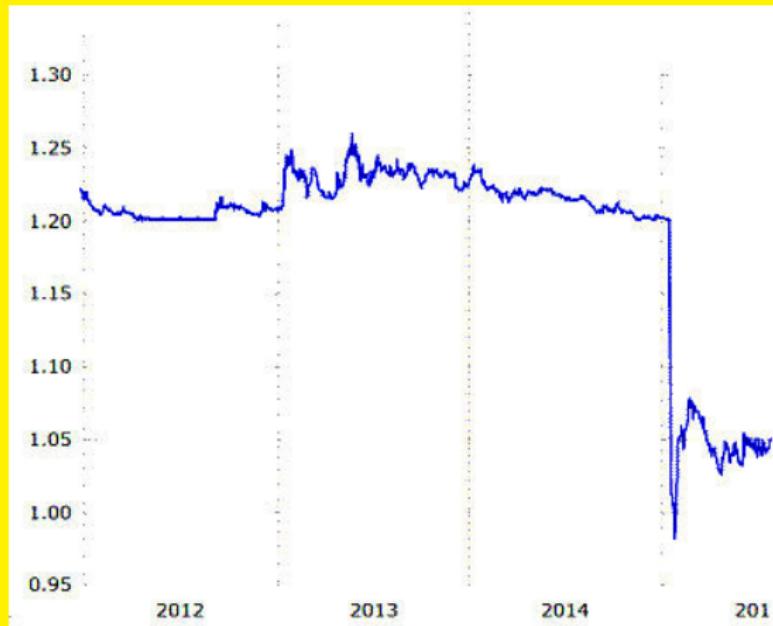
Regressing on  $X_1$ ,  $X_2$ , and  $X_3$ :



$X_2$  yields an invariant model. Ground truth:



How much CHF do I need to pay for buying 1 EUR?



<http://www.fremdwaehrungskonto.info/wp-content/uploads/2015/07/CHF-EUR-Kursentwicklung-2011-2015.gif>

monthly data Swiss National Bank Jan 1999 - Jan 2017

<b>description</b>	
$Y$	exchange rate Euro to Swiss Franks
$X^1$	change in average call money rate
$X^2$	log returns of foreign currency investments of the SNB
$X^3$	log returns of reserve positions at Intern. Monetary Fund of the SNB
$X^4$	log returns of monetary assistance loans of the SNB
$X^5$	log returns of Swiss Frank securities of the SNB
$X^6$	log returns of remaining assets of the SNB
$X^7$	log returns of Swiss GDP
$X^8$	log returns of Euro zone GDP
$X^9$	inflation rate for Switzerland

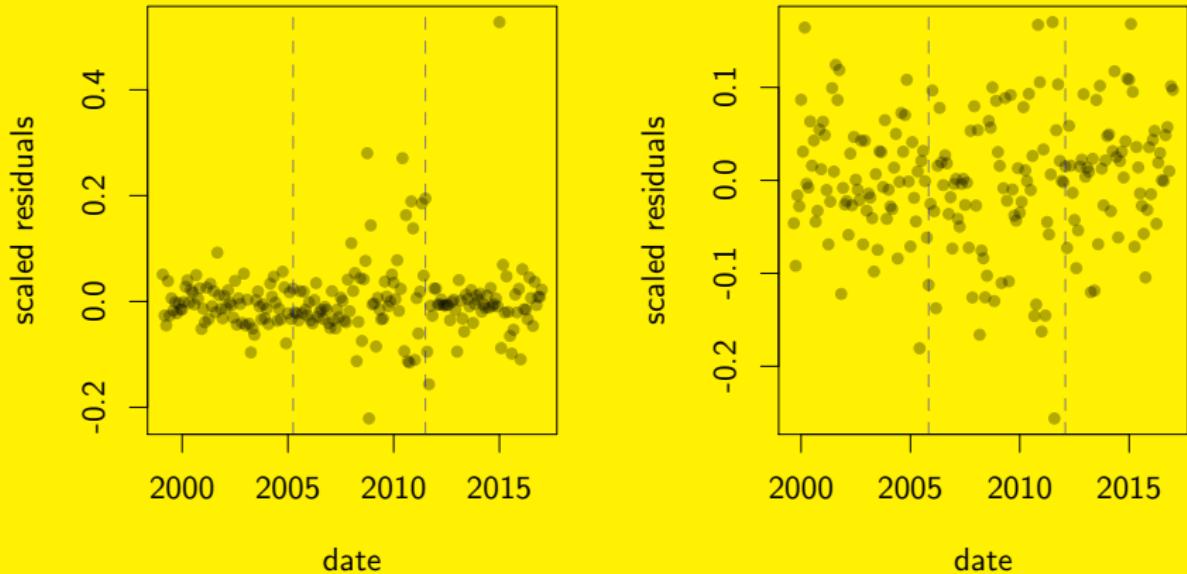


Figure: left plot (not invariant) and right plot (invariant)

monthly data Swiss National Bank Jan 1999 - Jan 2017

---

### **description**

- $Y$  exchange rate Euro to Swiss Franks
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Pfister, Bühlmann, JP, JASA 2018:

Non-inv. models rejected if  $\sqrt{\log n/n} = o(a_n)$ , where  $a_n$  is largest difference in noise variances.

## Discrete environments (gene data):

JP, Meinshausen, Bühlmann: *Causal inference using invariant prediction: identification and confidence intervals*, JRSSB 2016

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No environments (finance data):  = time

Pfister, Bühlmann, JP: *Invariant causal pred. for seq. data*, JASA 2018

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Heinze-Deml, JP, Meinshausen: *Invariant Causal Prediction for Nonlinear Models*, Journal of Causal Inference 2018

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Discrete hidden variables (Earth system data):

Christiansen, JP: *Invariance-based Causal Discovery in the Presence of Discrete Hidden Variables*, JMLR 2020

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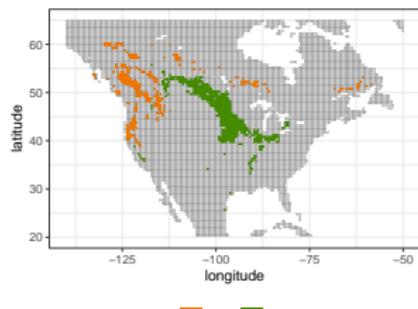
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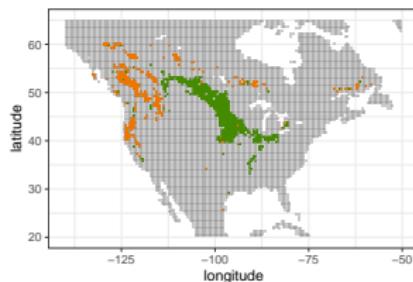
D

Chr

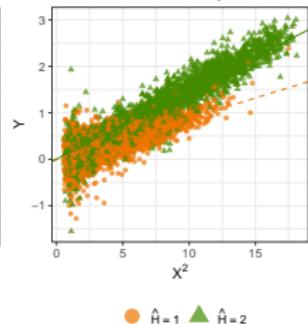
IGBP land cover classification



Classification based on reconstruction of  $H$



Fluorescence yield



Discrete environments (gene data):

JP, Meinshausen, Bühlmann: *Causal inference using invariant prediction: identification and confidence intervals*, JRSSB 2016

No environments (finance data):  = time

Pfister, Bühlmann, JP: *Invariant causal pred. for seq. data*, JASA 2018

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Heinze-Deml, JP, Meinshausen: *Invariant Causal Prediction for Nonlinear Models*, Journal of Causal Inference 2018

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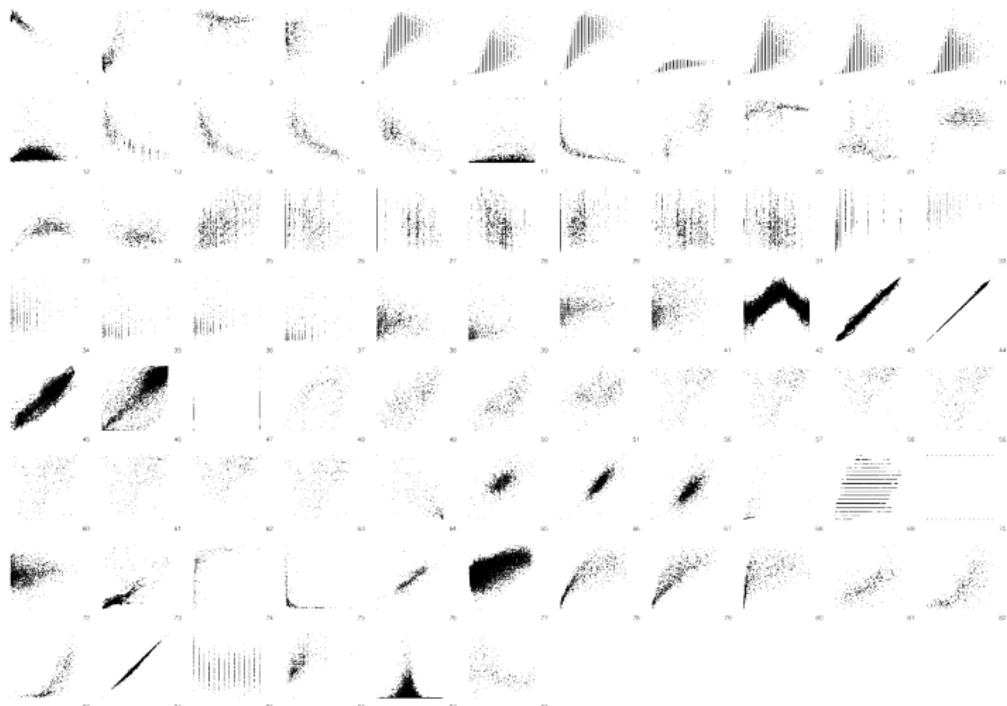
Survival data (registry data):

Laksafoss, JP: *Causal Methods for Survival Analysis*, work in progress



What can we do with two variables and no environments?  
(In general, nothing is possible.)

# Idea 3: restricted structural causal models



Mooij, JP, Janzing, Zscheischler, Schölkopf: *Disting. cause from effect using obs. data: methods and benchm.*, JMLR 2016

## Idea 3: restricted structural causal models

Consider a distribution entailed by

$$\boxed{Y = f(X) + N_Y}$$

with  $N_Y, X \stackrel{ind}{\sim} \mathcal{N}$



## Idea 3: restricted structural causal models

Consider a distribution entailed by

$$\boxed{Y = f(X) + N_Y}$$

with  $N_Y, X \stackrel{\text{ind}}{\sim} \mathcal{N}$



Then, if  $f$  is nonlinear, there is no

$\cancel{X = g(Y) + Mx}$

with  $M_x, Y \stackrel{\text{ind}}{\sim} \mathcal{N}$

```
graph LR; Y((Y)) --> X((X))
```

JP, J. Mooij, D. Janzing and B. Schölkopf: *Causal Discovery with Continuous Additive Noise Models*, JMLR 2014

## Idea 3: restricted structural causal models

Consider a distribution entailed by

$$\boxed{Y = \textcolor{red}{X}^3 + N_Y}$$

with  $N_Y, X \stackrel{\textit{ind}}{\sim} \mathcal{N}$

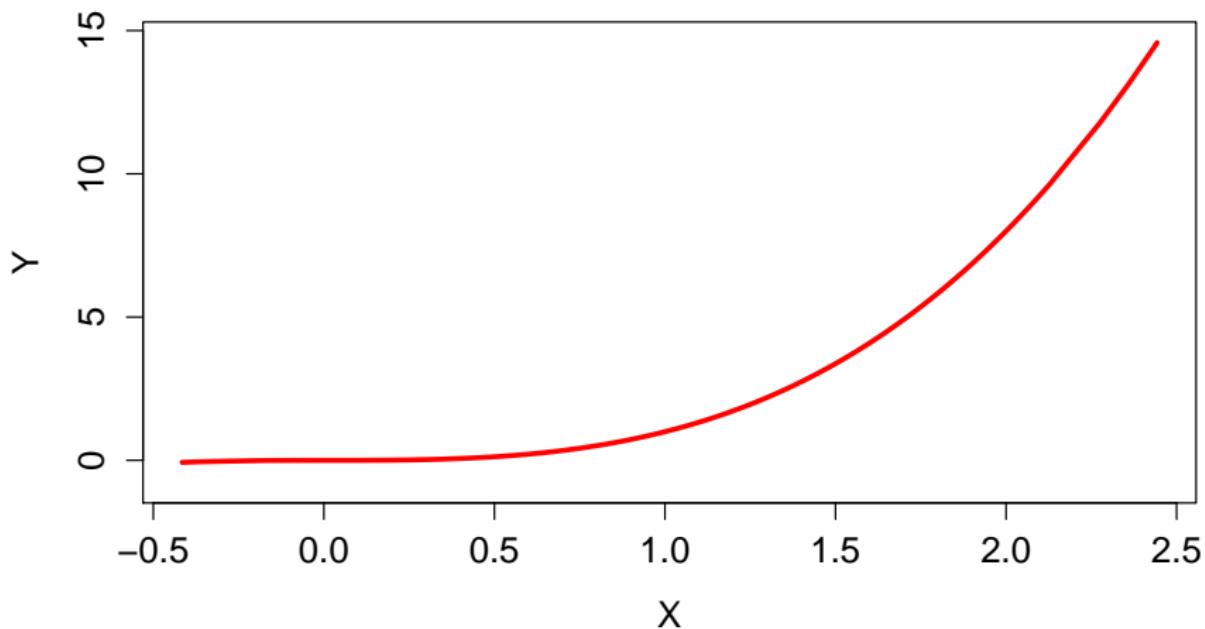
$$\textcircled{X} \longrightarrow \textcircled{Y}$$

with

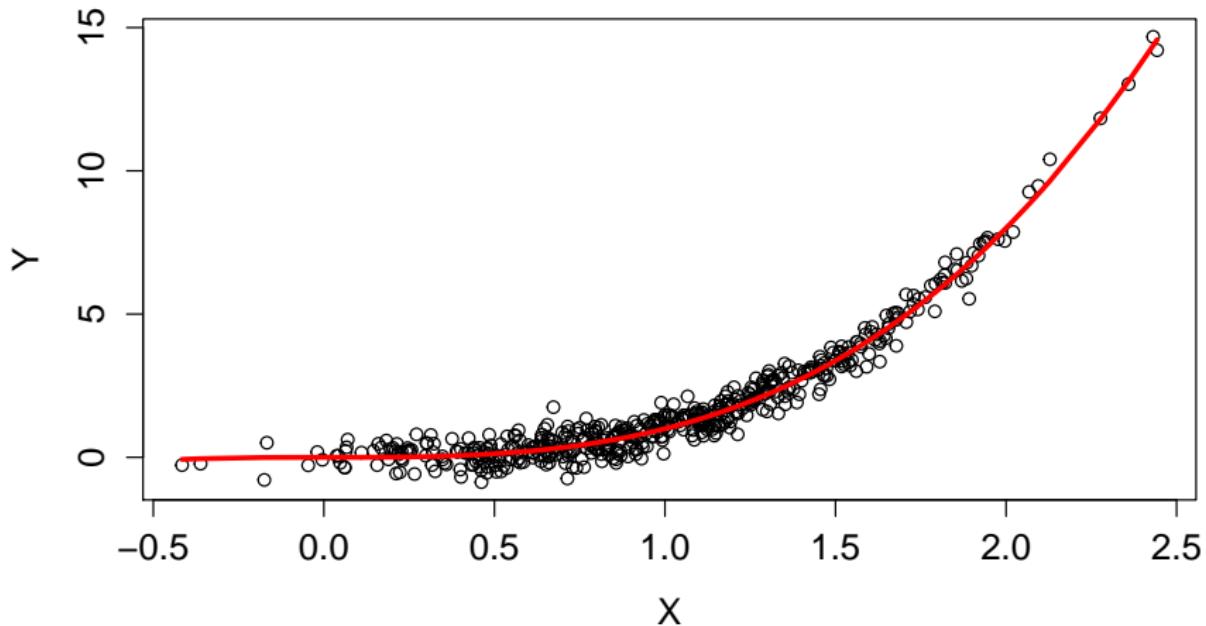
$$X \sim \mathcal{N}(1, 0.5^2)$$

$$N_Y \sim \mathcal{N}(0, 0.4^2)$$

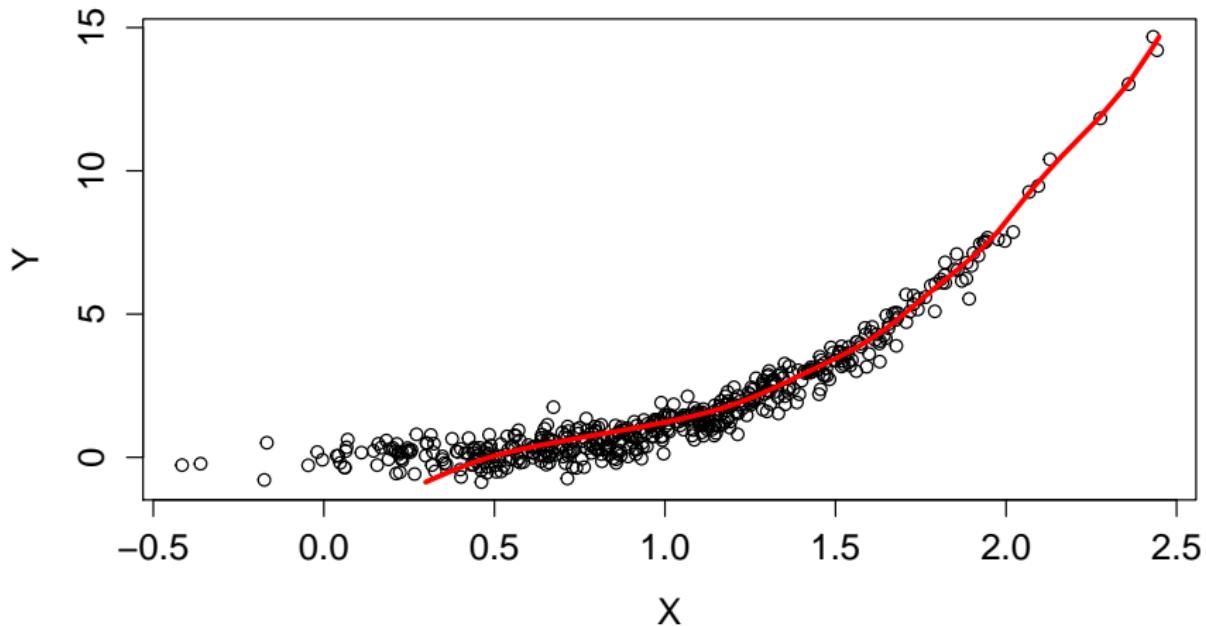
## Idea 3: restricted structural causal models



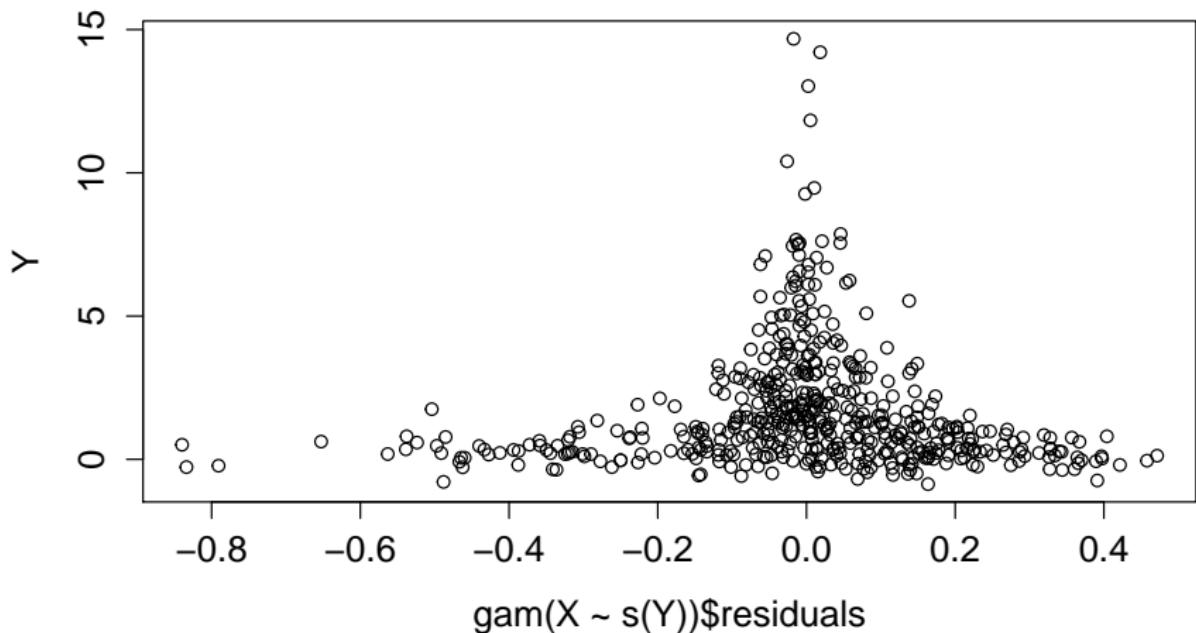
## Idea 3: restricted structural causal models



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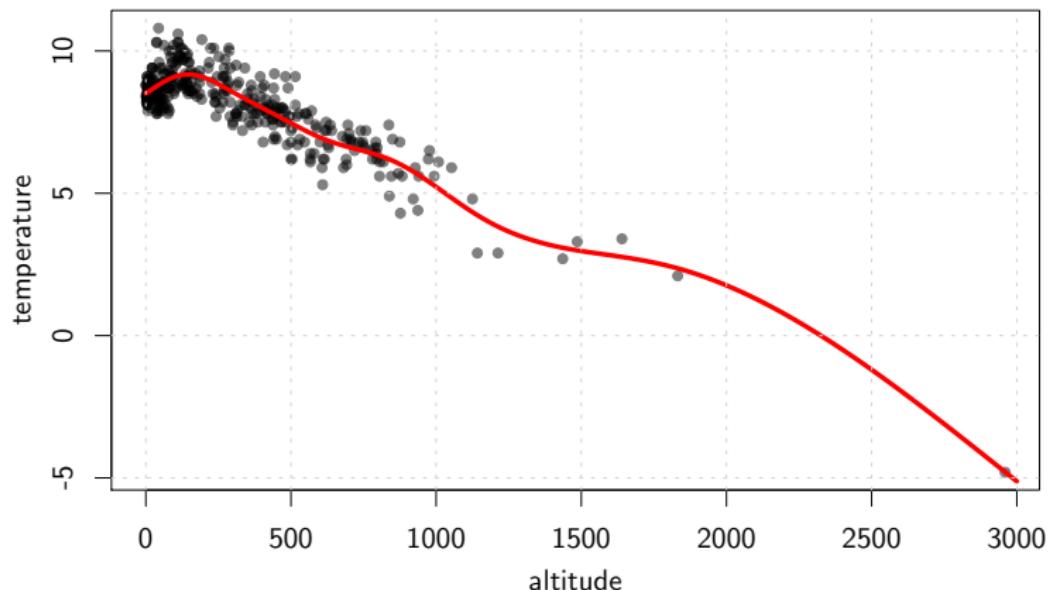
## Idea 3: restricted structural causal models



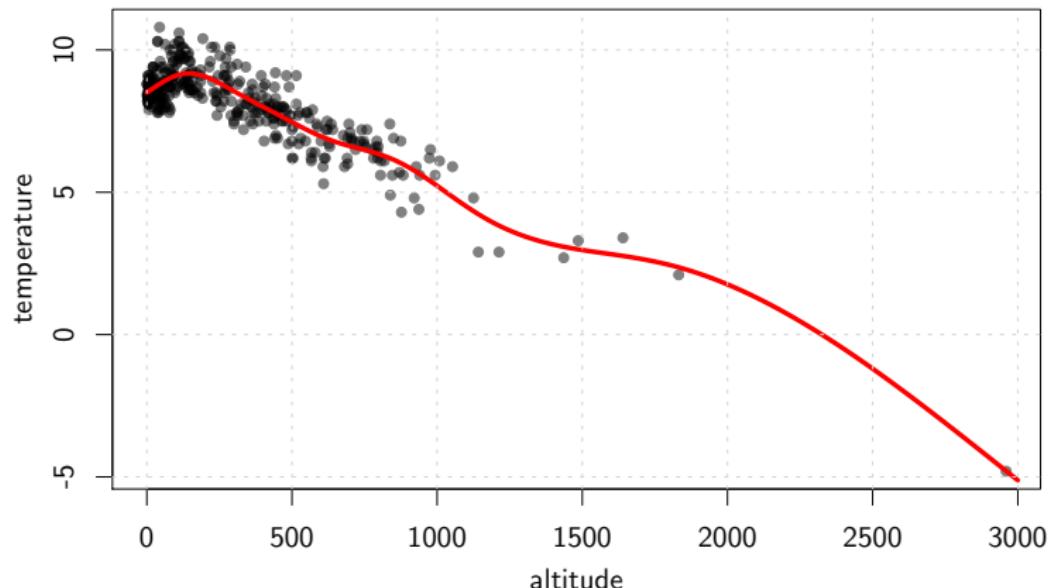
## Idea 3: restricted structural causal models

Method... (jupyter notebook: RestrictedSCMs.ipynb)

## Example: altitude and temperature



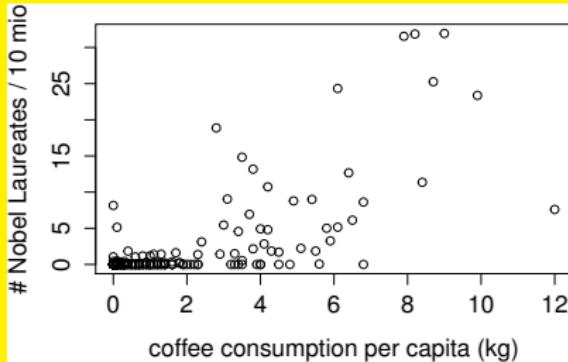
## Example: altitude and temperature



p-value forward: 0.024

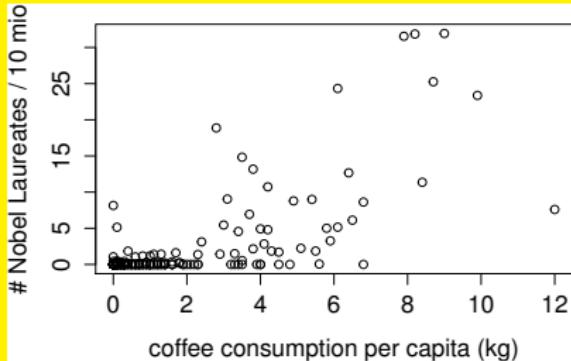
p-value backward: 0.0000000000019

# Example: coffee



Correlation: 0.698  
 $p$ -value:  $< 2.2 \cdot 10^{-16}$

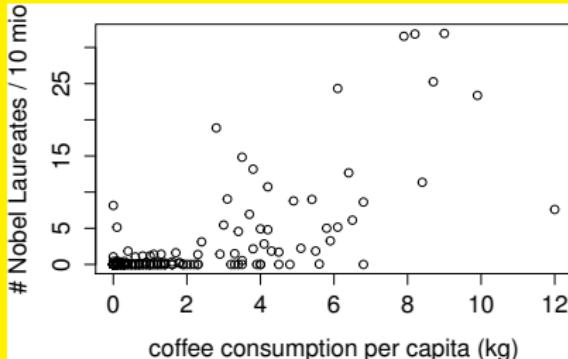
# Example: coffee



Correlation: 0.698

p-value:  $< 2.2 \cdot 10^{-16}$

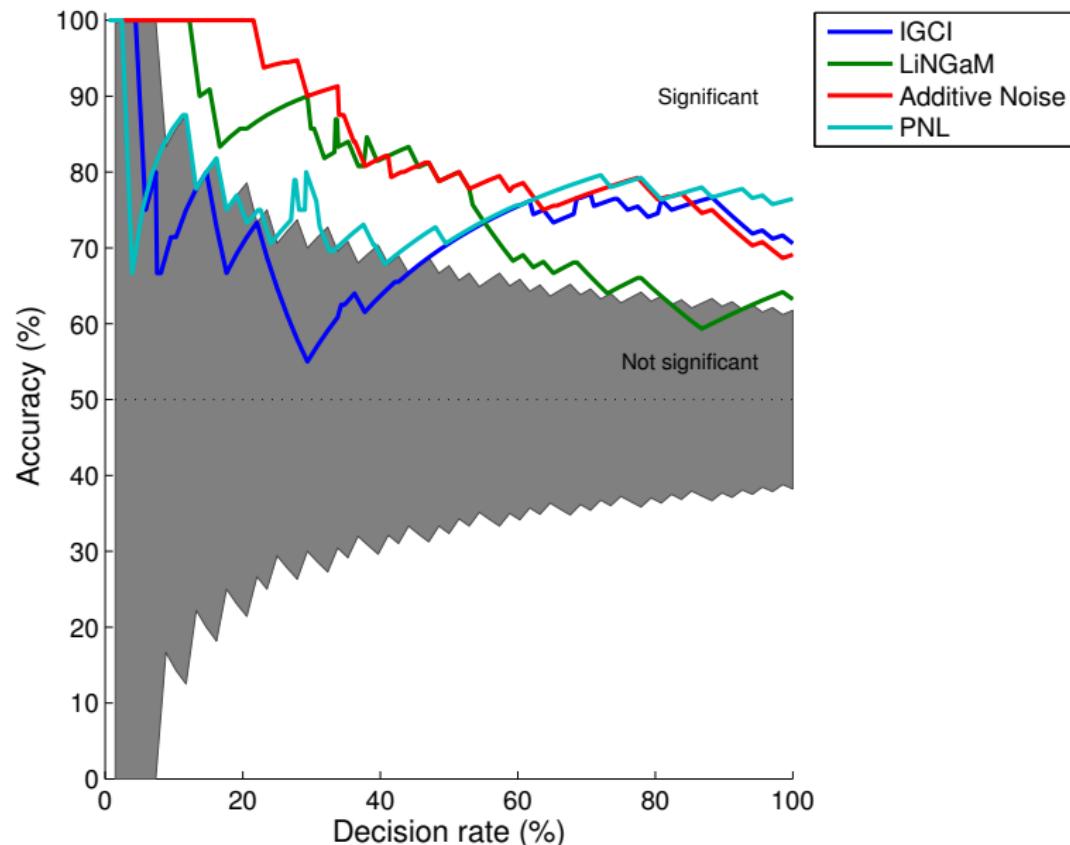
# Example: coffee



Correlation: 0.698

$p$ -value:  $< 2.2 \cdot 10^{-16}$

# Real Data: cause-effect pairs



## Idea 3: restricted structural causal models

Slightly surprising:

identifiability for two variables  $\rightsquigarrow$  identifiability for  $d$  variables

Peters et al.: *Identifiability of Causal Graphs using Functional Models*, UAI 2011

## Idea 3: restricted structural causal models

Slightly surprising:

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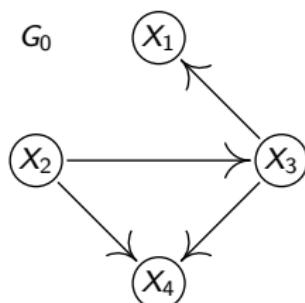
Most important counterexample: linear Gaussian.

## Idea 3: restricted structural causal models

Assume  $P(X_1, \dots, X_4)$  has been entailed by

$$\begin{aligned}X_1 &= f_1(X_3, N_1) \\X_2 &= N_2 \\X_3 &= f_3(X_2, N_3) \\X_4 &= f_4(X_2, X_3, N_4)\end{aligned}$$

- $N_i$  jointly independent
- $G_0$  has no cycles



Structural equation model.

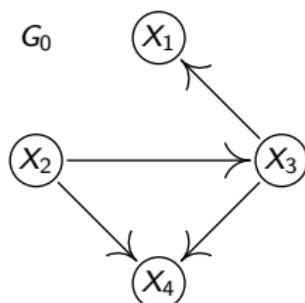
Can the DAG be recovered from  $P(X_1, \dots, X_4)$ ?

## Idea 3: restricted structural causal models

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Structural equation model.

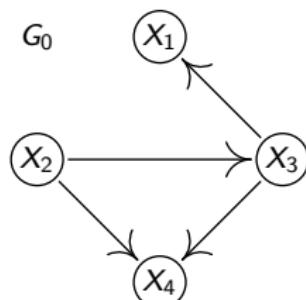
Can the DAG be recovered from  $P(X_1, \dots, X_4)$ ? **No.**

# Idea 3: restricted structural causal models

Assume  $P(X_1, \dots, X_4)$  has been entailed by

$$\begin{aligned}X_1 &= f_1(X_3) + N_1 \\X_2 &= N_2 \\X_3 &= f_3(X_2) + N_3 \\X_4 &= f_4(X_2, X_3) + N_4\end{aligned}$$

- $N_i \sim \mathcal{N}(0, \sigma_i^2)$  jointly independent
- $G_0$  has no cycles



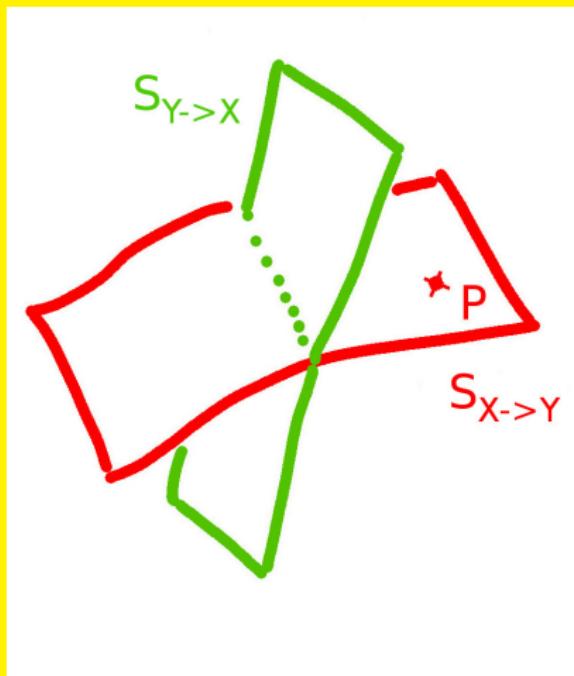
Additive noise model with Gaussian noise.

Can the DAG be recovered from  $P(X_1, \dots, X_4)$ ? Yes iff  $f_i$  nonlinear.

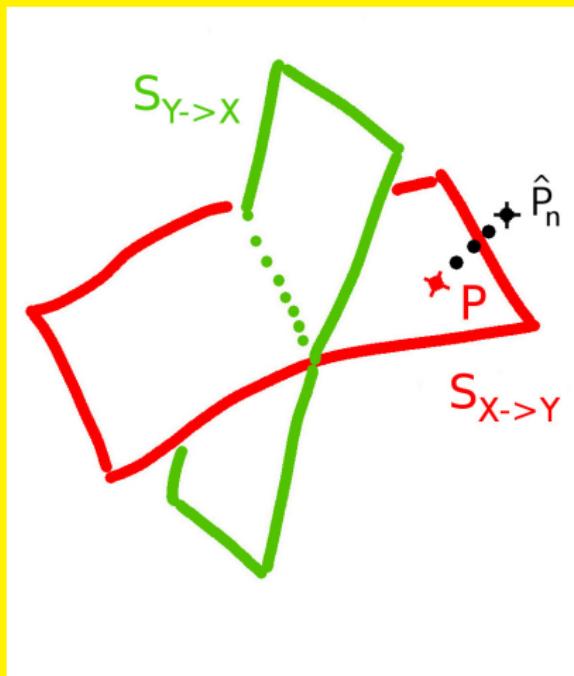
JP, J. Mooij, D. Janzing and B. Schölkopf: *Causal Discovery with Continuous Additive Noise Models*, JMLR 2014

P. Bühlmann, JP, J. Ernest: *CAM: Causal add. models, high-dim. order search and penalized regr.*, Annals of Statistics 2014

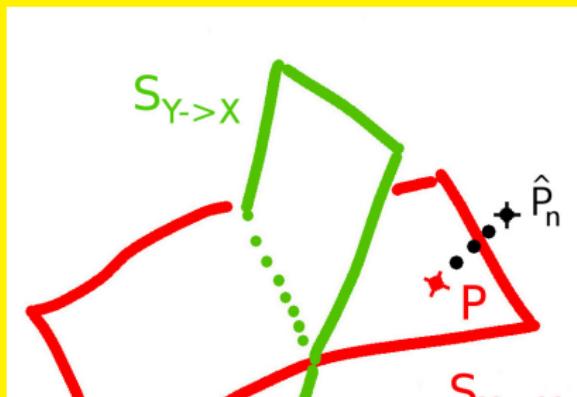
## Idea 3: restricted structural causal models



## Idea 3: restricted structural causal models



# Idea 3: restricted structural causal models



Method: Minimizing KL

Choose the direction that corresponds to the closest subspace...



## Idea 3: restricted structural causal models

Consider model classes

$$\mathcal{S}_G := \{Q : Q \text{ entailed by a causal additive model (CAM) with DAG } G\}$$

Define

$$\hat{G}_n := \underset{\text{DAG } G}{\operatorname{argmin}} \inf_{Q \in \mathcal{S}_G} \text{KL}(\hat{P}_n || Q)$$

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$$\max_{\substack{\text{likelihood}}} \underset{\substack{\text{DAG } G}}{\operatorname{argmin}} \sum_{i=1}^p \log \hat{\text{var}}(\text{residuals}_{\text{PA}_i^G \rightarrow X_i})$$

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$$\stackrel{\substack{\text{max.} \\ \text{likelihood}}}{=} \underset{\substack{\text{DAG } G}}{\operatorname{argmin}} \sum_{i=1}^p \log \hat{\text{var}}(\text{residuals}_{\text{PA}_i^G \rightarrow X_i})$$

Wait, there is no penalization on the number of edges!

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Wait, there is no penalization on the number of edges!

Wait again, there are too many DAGs!

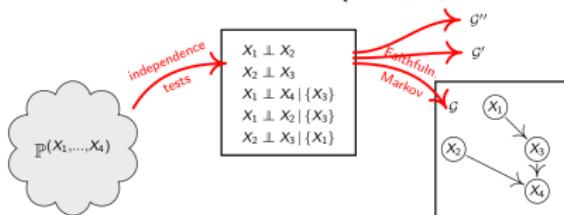
# Idea 3: restricted structural causal models

$d$	number of DAGs with $d$ nodes
1	1
2	3
3	25
4	543
5	29281
6	3781503
7	1138779265
8	783702329343
9	1213442454842881
10	4175098976430598143
11	31603459396418917607425
12	521939651343829405020504063
13	18676600744432035186664816926721
14	1439428141044398334941790719839535103
15	237725265553410354992180218286376719253505
16	83756670773733320287699303047996412235223138303
17	62707921196923889899446452602494921906963551482675201
18	99421195322159515895228914592354524516555026878588305014783
19	332771901227107591736177573311261125883583076258421902583546773505
20	2344880451051088988152559855229099188899081192234291298795803236068491263
21	34698768283588750028759328430181088222313944540438601719027559113446586077675521
22	107582292172576149365295617932762432657372766280918521810409000500559527511693495107583

<https://oeis.org/A003024/b003024.txt>

## Summary Part II:

- Idea 1: independence-based methods (single environment)



- Idea 2: invariant prediction (the more heterogeneity the better!)



InvariantCausalPrediction.ipynb

- Idea 3: additive noise (single environment)

$$X_1 = f_1(X_3) + N_1$$

$$X_2 = N_2$$

$$X_3 = f_3(X_2) + N_3$$

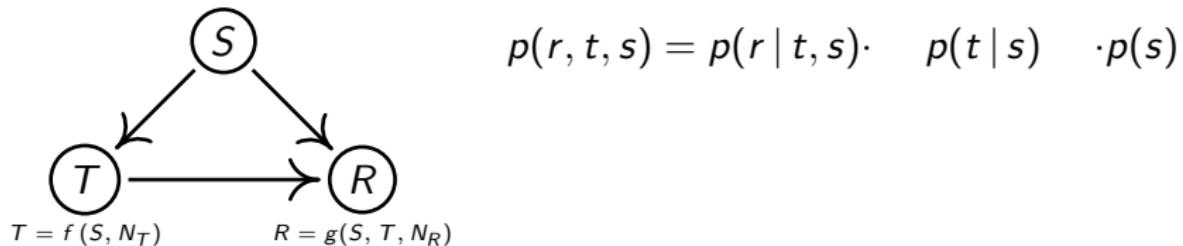
$$X_4 = f_4(X_2, X_3) + N_4$$

RestrictedSCMs.ipynb

## **Part III: Applications to Machine Learning (short)**

# Idea: RL

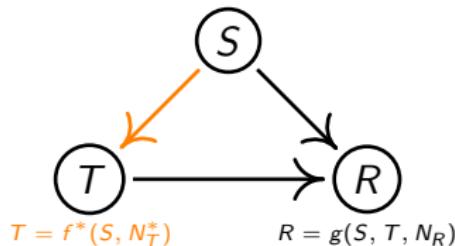
Recall the kidney stones:



Question: What would happen if...?

# Idea: RL

Recall the kidney stones:



$$p(r, t, s) = p(r | t, s) \cdot p(t | s) \cdot p(s)$$
$$p_3^*(r, t, s) = p(r | t, s) \cdot \underbrace{p^*(t | s)}_{p^*(t | s) = ?} \cdot p(s)$$

Question: What would happen if...?

What is  $\sup_{p^*} E_{p^*} R$ ?

# Idea: anchor regression



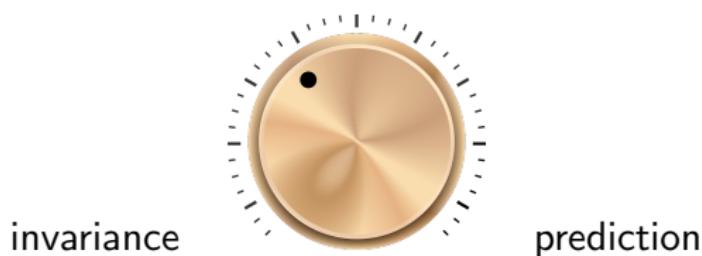
# Idea: anchor regression



# Idea: anchor regression



# Idea: anchor regression



Find a trade-off between

- invariance with respect to 
- AND • predictive power

# Idea: anchor regression

$Y \in \mathbb{R}^1$ : target

$X \in \mathbb{R}^{1 \times d}$ : predictors

$A \in \mathbb{R}^{1 \times q}$ : anchors,  $EA^t A = Id$

$$b^\gamma := \underset{b}{\operatorname{argmin}} \underbrace{\mathbb{E}(Y - Xb)^2}_{\text{prediction}} + \gamma \underbrace{\|EA^t(Y - Xb)\|_2^2}_{\text{invariance}}$$

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$\gamma \rightarrow 0$ : OLS

$\gamma \rightarrow \infty$ : IV solution (if identifiable)

$\gamma \rightarrow \infty$ : best invariant predictor (if not identifiable)

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- Anchor regression minimizes worst case prediction error under shift interventions.

Rothenhäusler, Bühlmann, Meinshausen, JP (arXiv:1801.06229)

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Rothenhäusler, Bühlmann, Meinshausen, JP (arXiv:1801.06229)
- The finite sample estimator is known as a  $k$ -class estimator for IV solution.  
Theil (1958), Nagar (1959), Jakobsen and JP (arXiv:2005.03353)

# Idea: anchor regression

$$\begin{pmatrix} X \\ Y \\ H \end{pmatrix} \leftarrow B \cdot \begin{pmatrix} X \\ Y \\ H \end{pmatrix} + \varepsilon + MA,$$

shifted:  $\begin{pmatrix} X^\nu \\ Y^\nu \\ H^\nu \end{pmatrix} \leftarrow B \cdot \begin{pmatrix} X^\nu \\ Y^\nu \\ H^\nu \end{pmatrix} + \varepsilon + \nu.$

$Id - B$  invertible

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$Id - B$  invertible

## Theorem

For any  $b \in \mathbb{R}^d$  we have

$$\operatorname{argmin}_b E(Y - Xb)^2 + \gamma \|EA^t(Y - Xb)\|_2^2 = \max_{\nu \in C^\gamma} \mathbb{E}[(Y^\nu - X^\nu b)^2],$$

where

$$C^\gamma := \{\nu = M\delta \text{ such that } \|\delta\|_2 \leq \sqrt{\gamma}\}.$$

# Idea: anchor regression



[http://www.srfcdn.ch/radio/modules/dynimages/624/srf-1/2015/01/diverses/264377.150114\\_raclette\\_key.jpg](http://www.srfcdn.ch/radio/modules/dynimages/624/srf-1/2015/01/diverses/264377.150114_raclette_key.jpg)

# Idea: anchor regression

Example: Maillard reaction

Glu, Mel, C5, ForAc, Triose, Cn, AcAc, Amad, lysR, Fru, AMP

# Idea: anchor regression

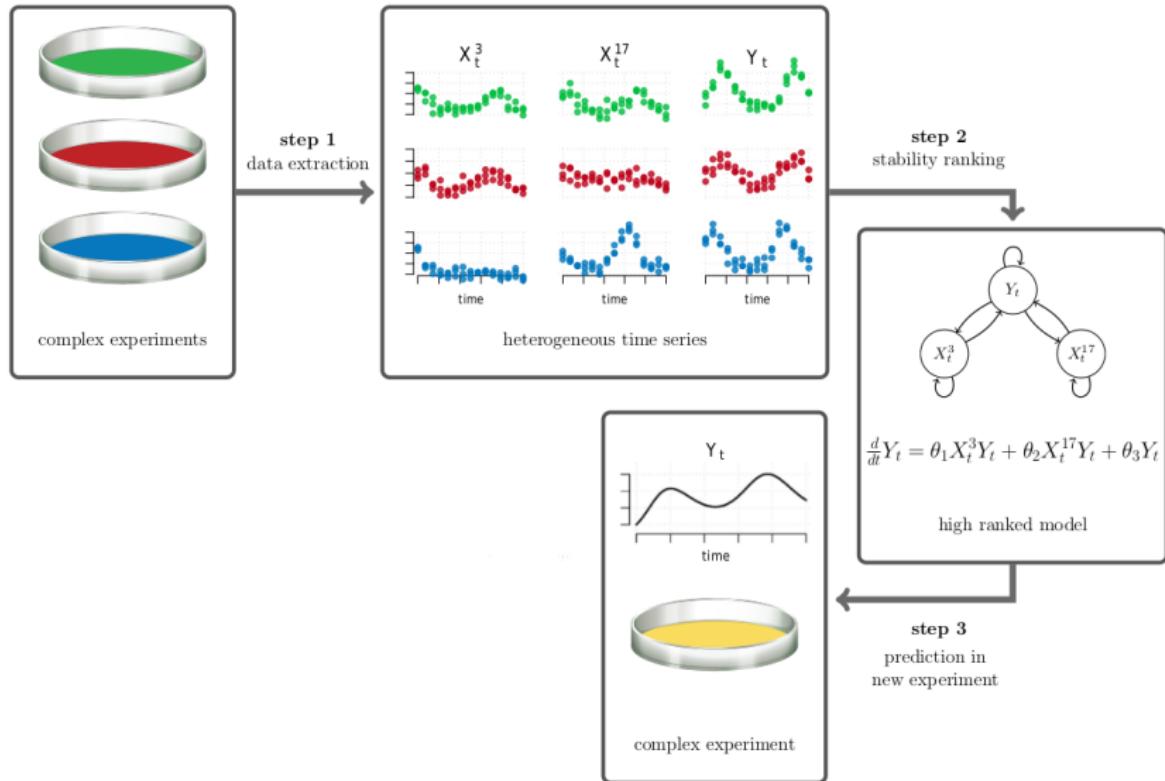
Example: Maillard reaction

Glu, Mel, C5, ForAc, Triose, Cn, AcAc, Amad, lysR, Fru, AMP

$$\frac{d}{dt}[\text{Glu}]_t = -\theta_1[\text{Glu}]_t + \theta_2[\text{Fru}]_t - \theta_3[\text{lysR}]_t[\text{Glu}]_t$$

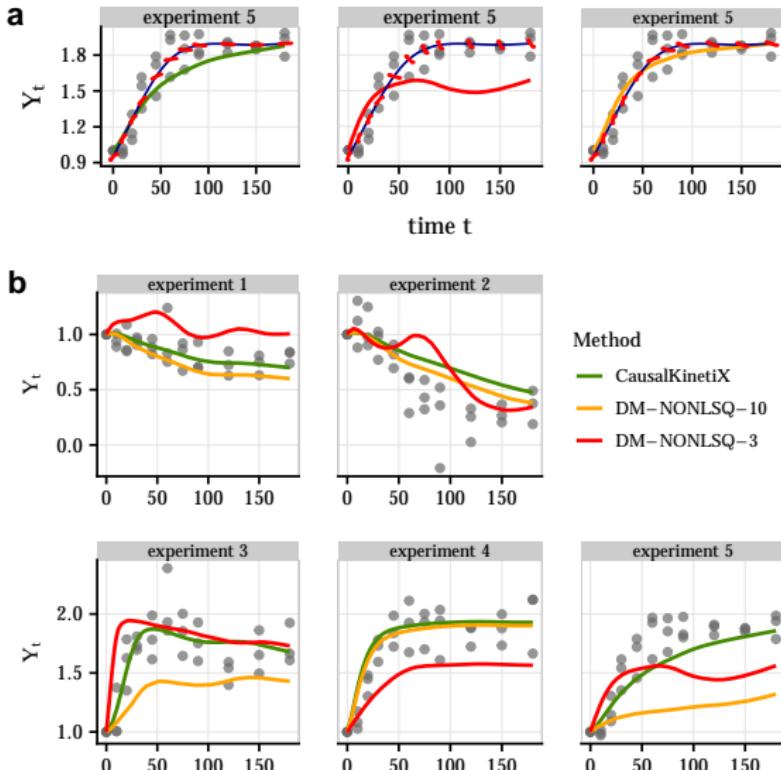
$$\frac{d}{dt}[\text{Mel}]_t = \theta_4[\text{AMP}]_t \quad \dots$$

# Idea: anchor regression



N. Pfister, S. Bauer, JP: *Identifying Causal Structure in Large-Scale Kinetic Systems*, PNAS 2019

# Idea: anchor regression

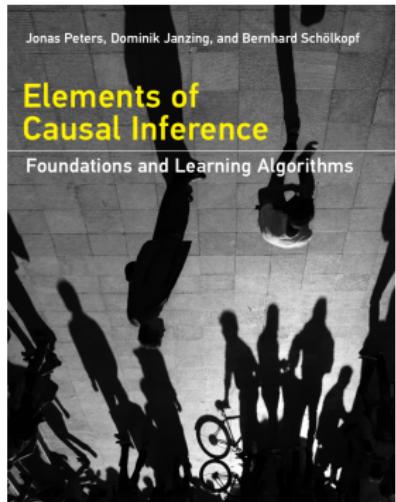


## **Summary Part III:**

- Idea 1: reformulate reinforcement learning,  
use causal structure
- Idea 2: semi-supervised learning from cause  
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- Idea 3: anchor regression

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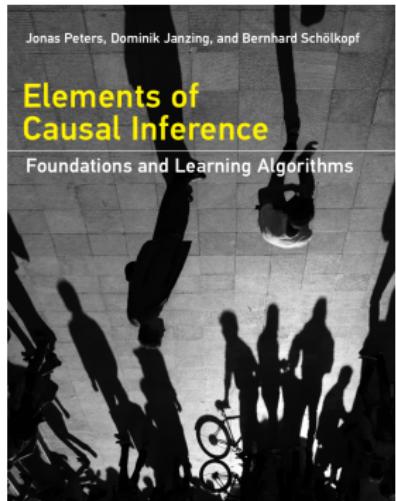
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For an exhaustive list of references, download pdf of  
JP, D. Janzing, B. Schölkopf: *Elements of Causal Inference: Foundations and Learning Algorithms*, MIT Press 2017.

## Summary Part III:

- Idea 1: reformulate reinforcement learning, use causal structure
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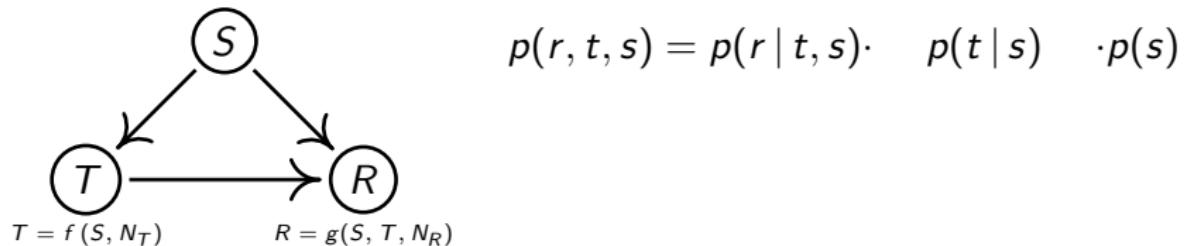
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— Tusind tak!

## **Part III: Applications to Machine Learning (long)**

# Idea 1: Blackjack

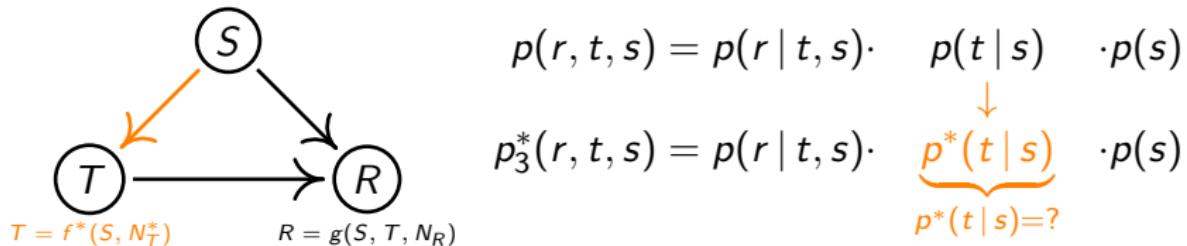
Recall the kidney stones:



Question: What would happen if...?

# Idea 1: Blackjack

Recall the kidney stones:



Question: What would happen if...?

What is  $\sup_{p^*} E_{p^*} R$ ?

# Idea 1: Blackjack

(some) Rules:

- **Dealing:** player two cards, dealer one card (all face up).
- **Goal:** more points in hand. Face cards: 10, ace either 1 or 11 points.
- **Player's moves:** *hit* (take card, but try  $\leq 21$ ), *stand*, *double down*, *split* (in case of pair).
- **Dealer's moves:** deterministic, does not stand before  $\geq 17$  points.
- **Blackjack:** ace and face card  $\rightarrow 1.5 \cdot \text{bet}$ .

# Idea 1: Blackjack



[https://de.wikipedia.org/wiki/Black\\_Jack.JPG](https://de.wikipedia.org/wiki/Black_Jack.JPG)

# Idea 1: Blackjack

When can we learn?

Objects of Interest:

- sample from  $p = p(X, Y, Z)$  (games),
- function of interest  $\ell = \ell(X, Y, Z)$  (money) and
- $p^*$  replacing  $p(y | x) \rightarrow p^*(y | x)$  (strategy = decisions | game state).

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Questions:

- What is  $E_{p^*} \ell$ ?

Needed:

- Values of  $X_i$ ,  $Y_i$  and  $\ell(X_i, Y_i, Z_i)$  (under  $p$ )

$X_i$	$Y_i$	$Z_i$	$\ell(X_i, Y_i, Z_i)$
-1.4	2.0	?	2.1
-0.5	0.7	?	2.5
-0.8	1.5	?	2.6
:	:	:	:

$X_i$	$Y_i$	$Z_i$	$\ell(X_i, Y_i, Z_i)$
$\heartsuit K, \heartsuit 9$	hit	?	-1
$\clubsuit A, \spadesuit J$	stand	?	1.5
$\spadesuit 10, \heartsuit 8$	stand	?	-1
:	:	:	:

# Idea 1: Blackjack

## Computation: Means

Assume  $p(y | x) \rightarrow p^*(y | x)$ .

$$\begin{aligned}\eta := \mathsf{E}_{p^*} \ell &= \int \ell(x, y, z) \, p^*(x, y, z) \, dx \, dy \, dz \\ &= \int \ell(x, y, z) \, \frac{p^*(x, y, z)}{p(x, y, z)} \, p(x, y, z) \, dx \, dy \, dz\end{aligned}$$

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Estimate  $\eta$  by

$$\hat{\eta} = \frac{1}{N} \sum_{i=1}^N \ell(X_i, Y_i, Z_i) \underbrace{\frac{p^*(Y_i | X_i)}{p(Y_i | X_i)}}_{w_i} = \frac{1}{N} \sum_{i=1}^N M_i, \quad \mathbb{E}_p \hat{\eta} = \eta$$

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Confidence intervals available!

# Idea 1: Blackjack

$$p(y | x) \rightarrow p^*(y | x)$$

Which  $p^*$  is best?

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$$p(y | x) \rightarrow p^*(y | x)$$

Which  $p^*$  is best? Parameterize and estimate

$$\nabla_{\theta} E_{p_{\theta}}|_{\theta=\tilde{\theta}}$$

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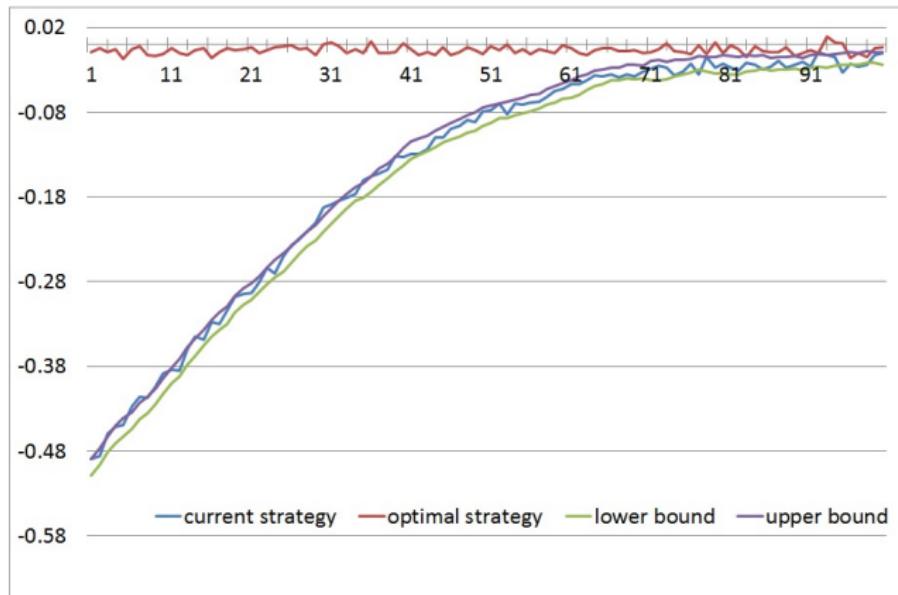
$$\nabla_{\theta} E_{p_{\theta}}|_{\theta=\tilde{\theta}}$$

Goal: Optimize  $E_{p_{\theta}} \ell$

Idea: Use gradient  $\nabla_{\theta} E_{p_{\theta}} \ell$  and optimize step-by-step.

Issues: confidence intervals, step size, . . . .

# Idea 1: Blackjack

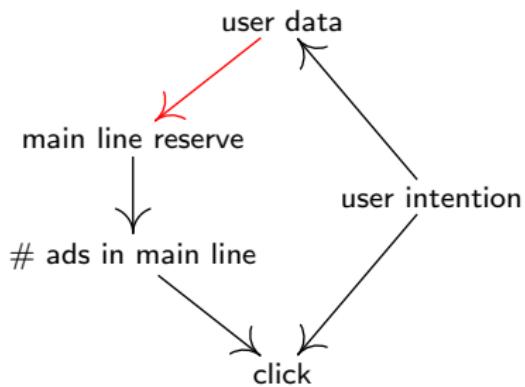


# Idea 1: Blackjack

What can we do with 100,000 samples?

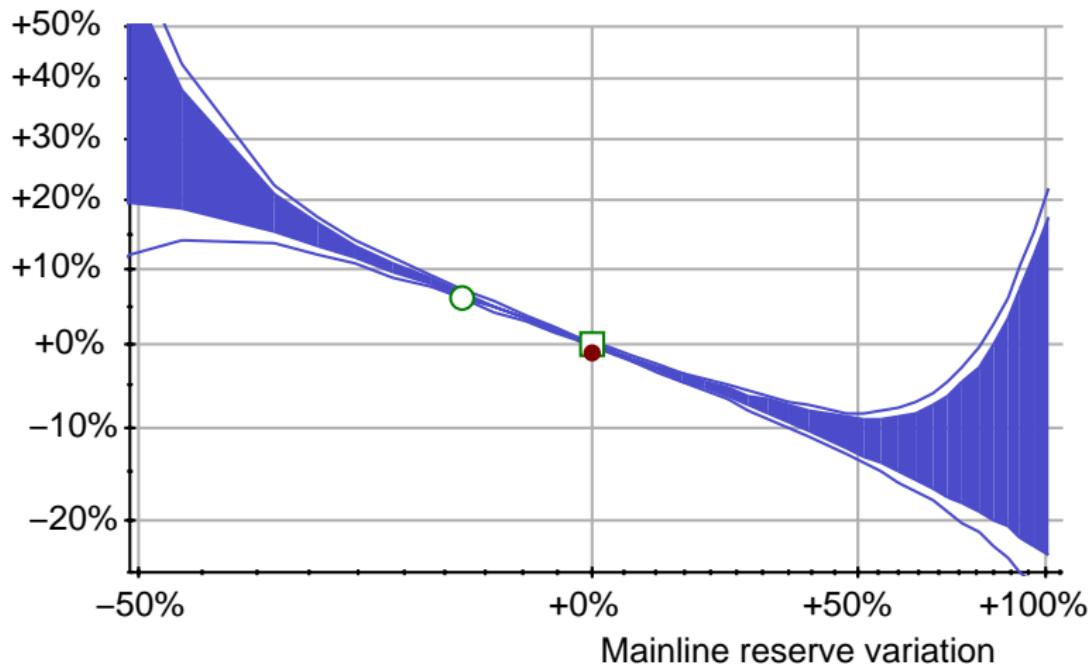
	Online	Offline
reached strategy	$E_{p^*} \ell \approx -5.1 C t$	$E_{p^*} \ell \approx -5.8 C t$
irrelevant games	33,653	61,048
costs	\$29,300	\$51,500
speed	slow: probabilities	even slower: gradients

# Idea 1: advertisement

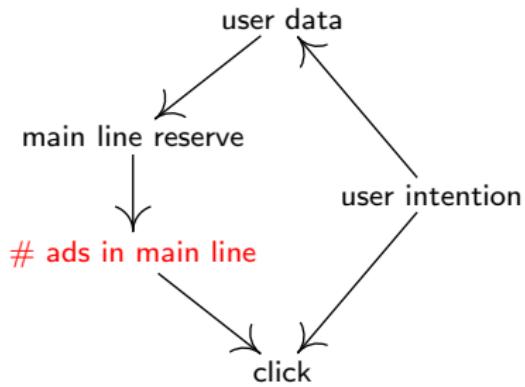


# Idea 1: advertisement

Average clicks per page

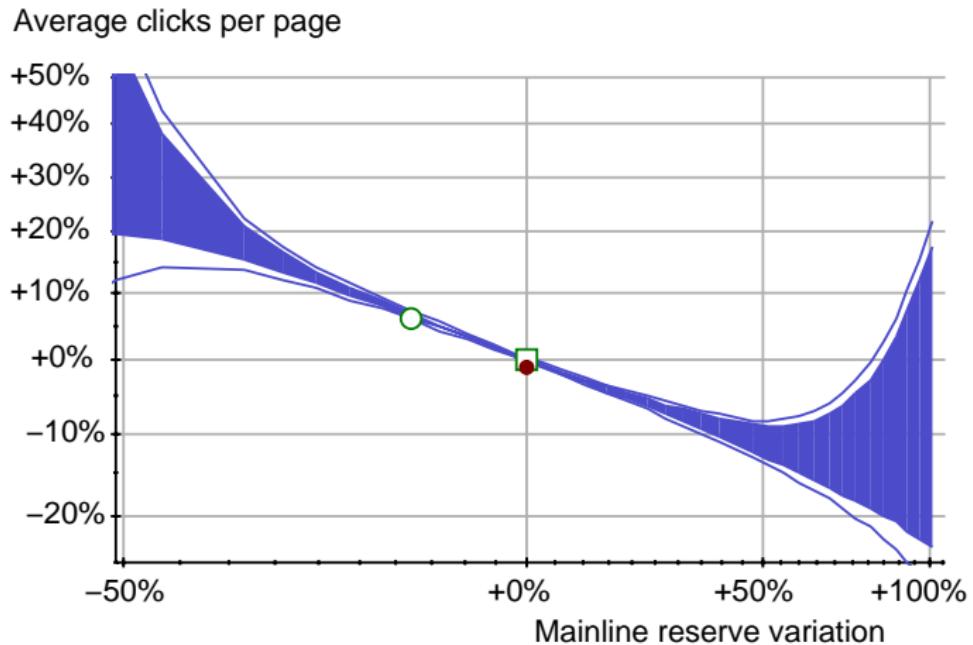


# Idea 1: advertisement



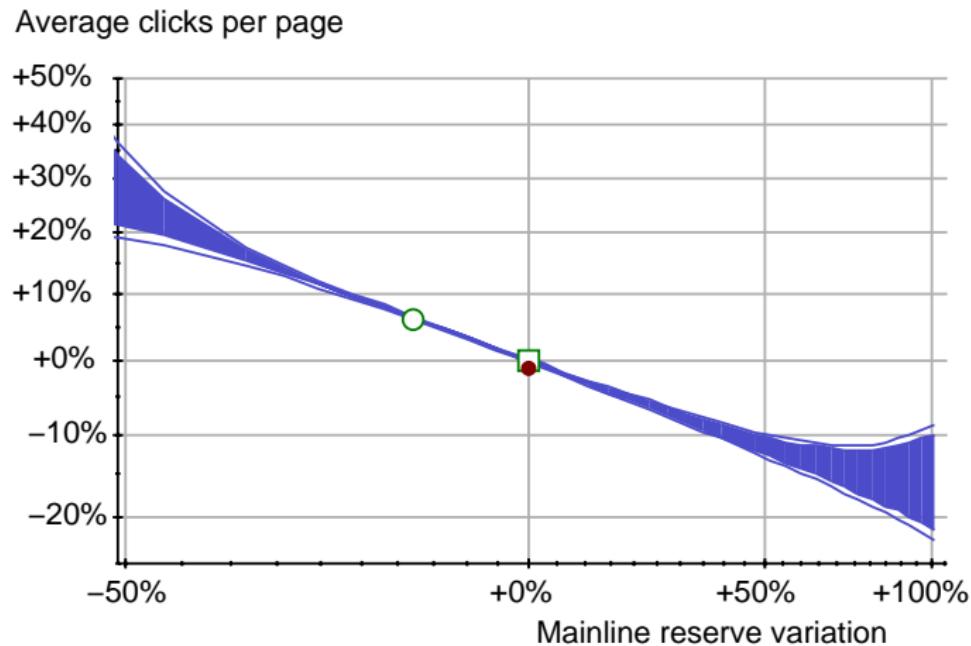
# Idea 1: advertisement

Old:



# Idea 1: advertisement

Using discrete variable (ads shown in mainline):





## Idea 2: semi-supervised learning

Consider a Markov factorization w.r.t. causal DAG:

$$p(x_1, \dots, x_d) = \prod_{i=1}^d p(x_i | x_{pa(i)})$$

## Idea 2: semi-supervised learning

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Modularity suggests:

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Special case:

$p(\text{cause}), p(\text{effect} | \text{cause})$  are “independent”

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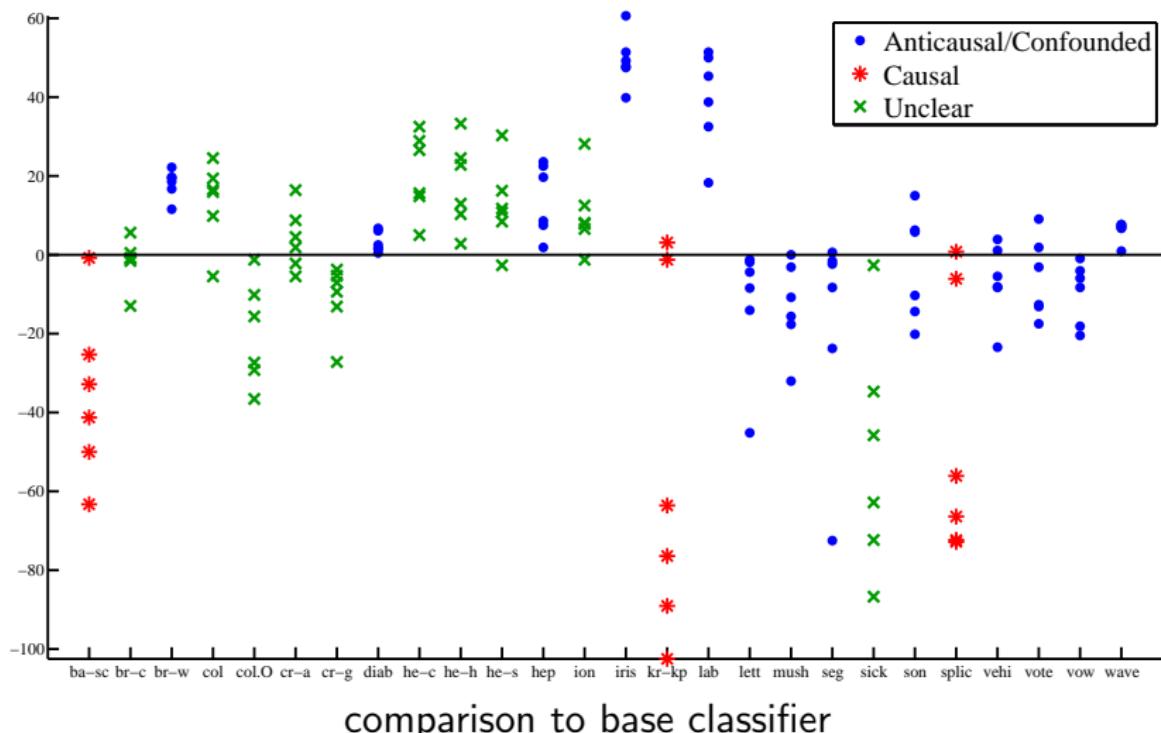
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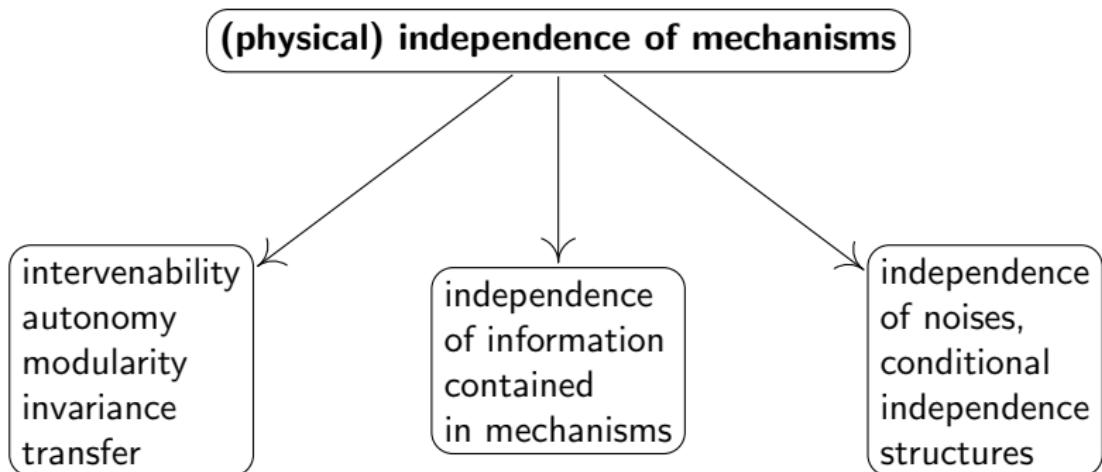
**But then: Semi-supervised Learning does not work from cause to effect.**

## Idea 2: semi-supervised learning

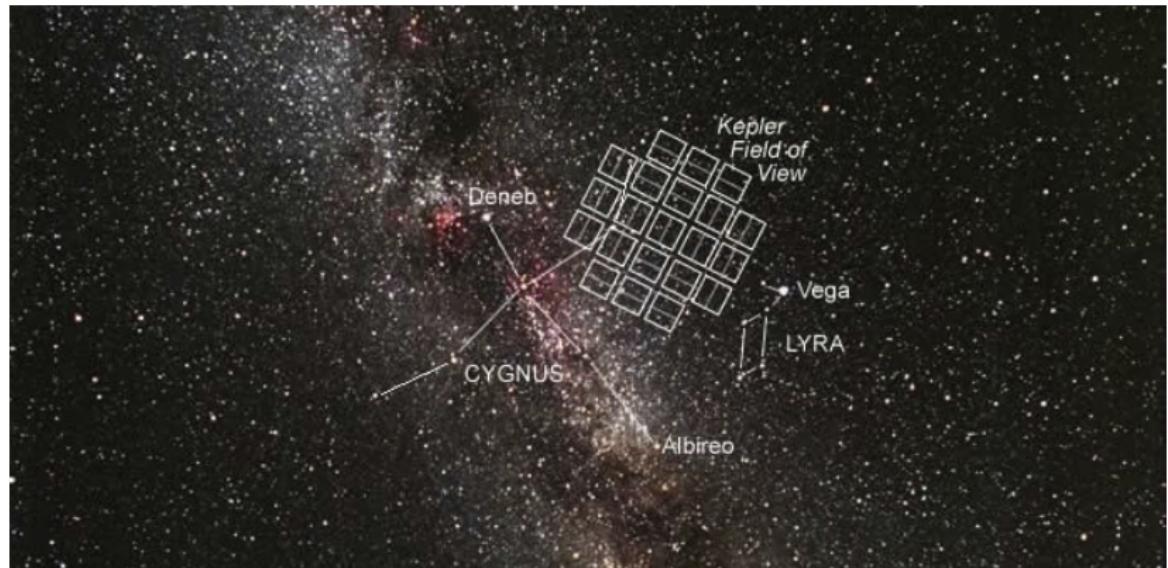


Schölkopf et al.: *On causal and anticausal learning*, ICML 2012

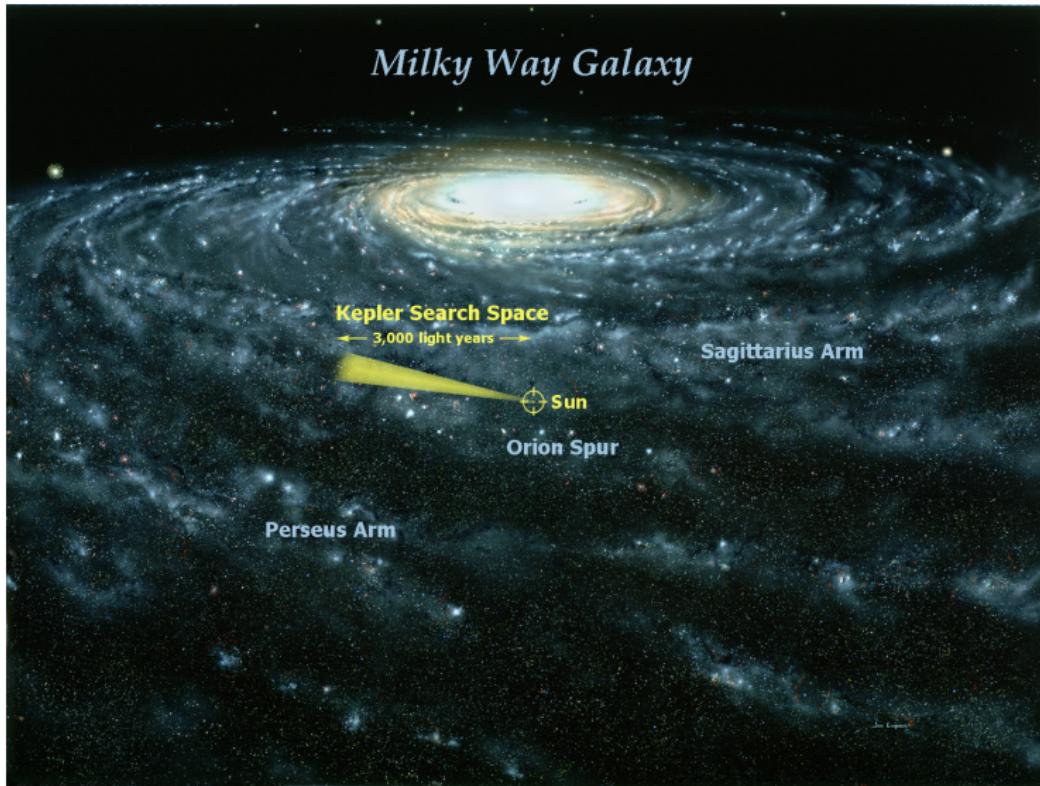
## Idea 2: semi-supervised learning



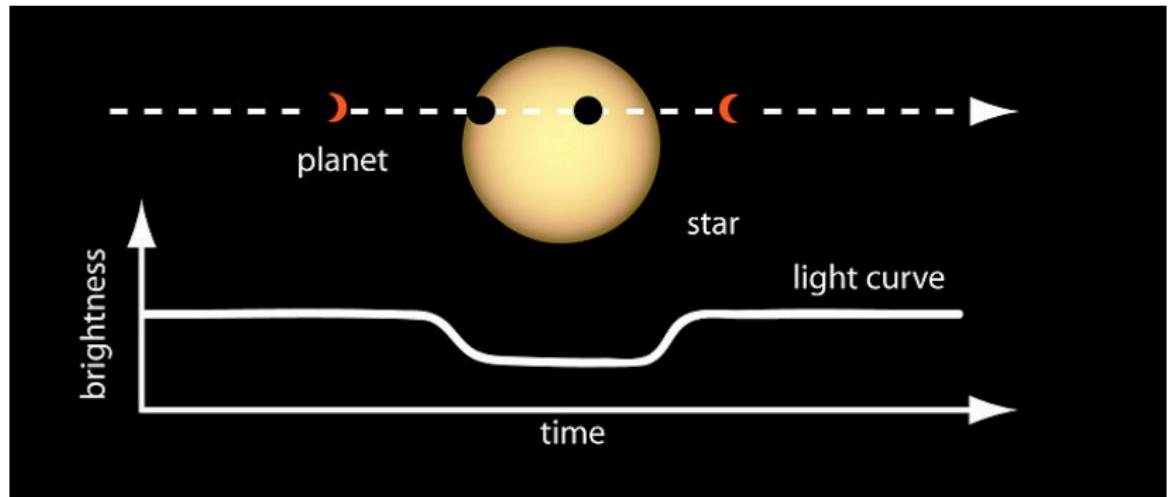
# Idea 3: half-sibling regression



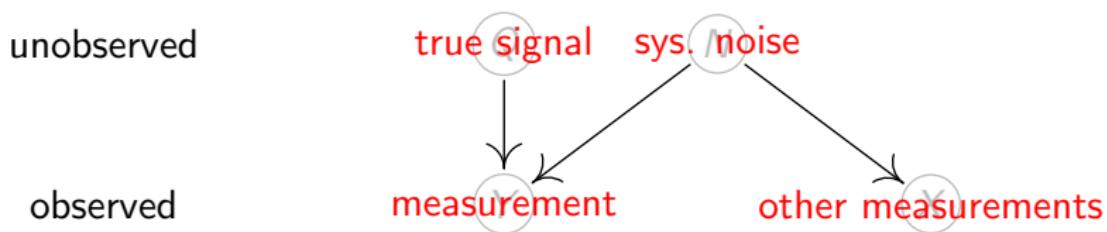
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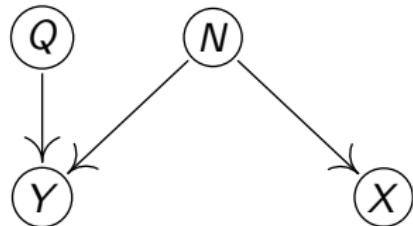
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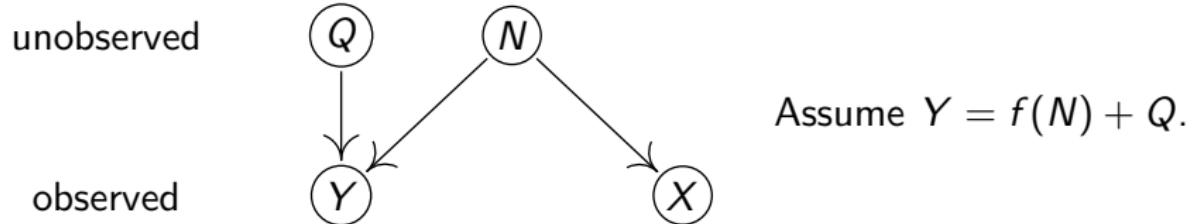
unobserved

observed



Assume  $Y = f(N) + Q$ .

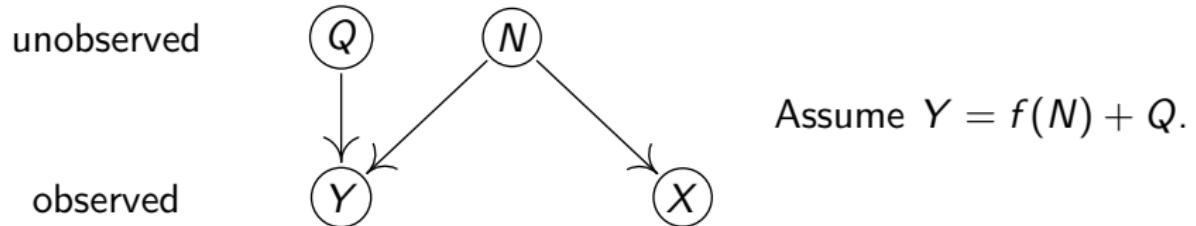
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Proposed idea:

Remove everything from  $Y$  explained by  $X$ .

## Idea 3: half-sibling regression



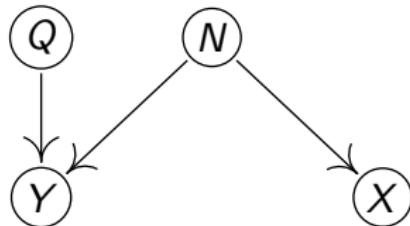
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Remove everything from  $Y$  explained by  $X$ .

Or:  $\hat{Q} := Y - E[Y | X]$ .

# Idea 3: half-sibling regression

unobserved



Assume  $Y = f(N) + Q$ .

observed

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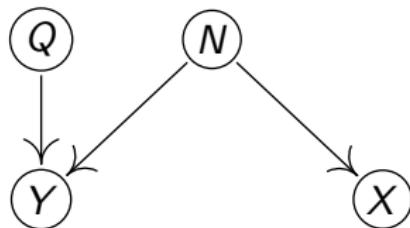
## Proposition

Convergence against “correct” signal  $Q$  (up to reparameterization) if

- perfect reconstruction:  $\exists \psi$  such that  $f(N) = \psi(X)$

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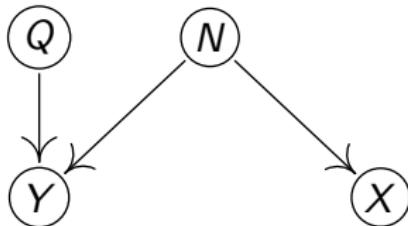
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- low noise:  $X = g(N) + s \cdot R$  and  $s \rightarrow 0$

# Idea 3: half-sibling regression

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Assume  $Y = f(N) + Q$ .

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Proposed idea:

Remove everything from  $Y$  explained by  $X$ .

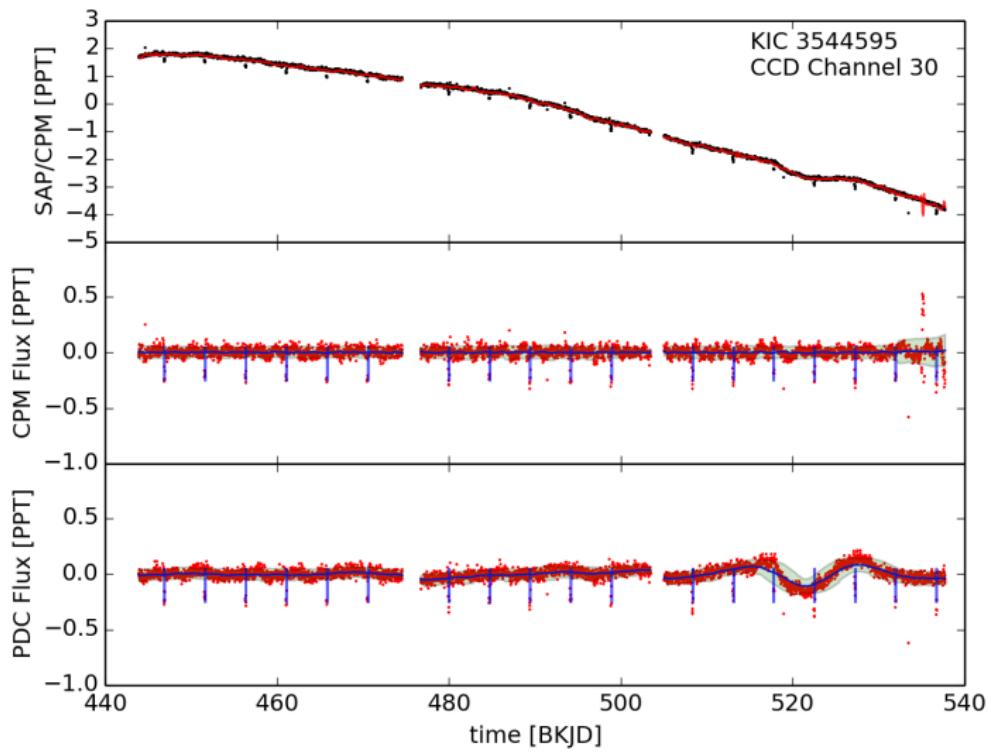
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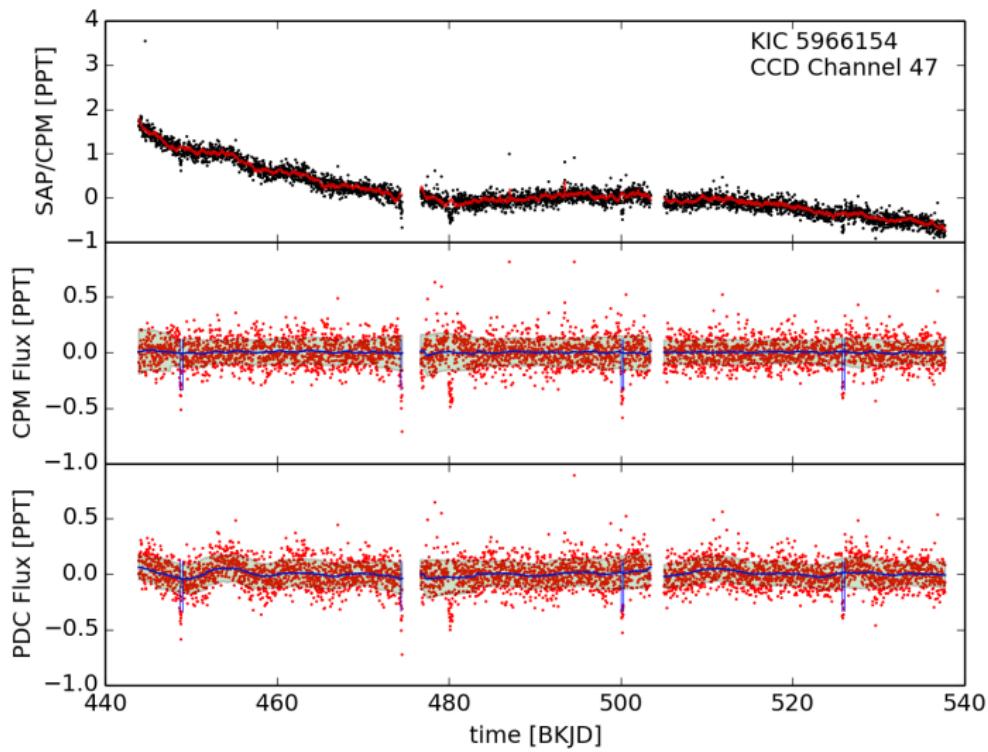
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- many  $X$ ’s:  $X_i = g_i(N) + R_i$ ,  $i = 1, \dots, \infty$

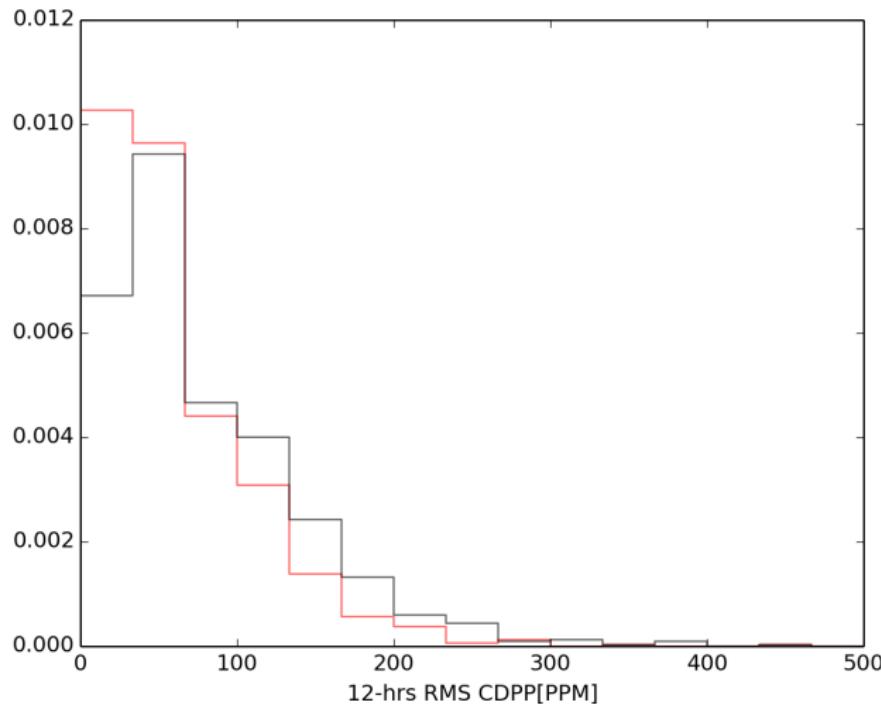
# Idea 3: half-sibling regression



# Idea 3: half-sibling regression



# Idea 3: half-sibling regression



Schölkopf et al.: *Removing systematic errors for exoplanet search via latent causes*, ICML 2015

## Idea 4: anchor regression



## Idea 4: anchor regression



## Idea 4: anchor regression



## Idea 4: anchor regression



Find a trade-off between

- invariance with respect to 
- AND • predictive power

## Idea 4: anchor regression

$Y \in \mathbb{R}^1$ : target

$X \in \mathbb{R}^{1 \times d}$ : predictors

$A \in \mathbb{R}^{1 \times q}$ : anchors,  $EA^t A = Id$

$$b^\gamma := \underset{b}{\operatorname{argmin}} \underbrace{\mathbb{E}(Y - Xb)^2}_{\text{prediction}} + \gamma \underbrace{\|EA^t(Y - Xb)\|_2^2}_{\text{invariance}}$$

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$\gamma \rightarrow 0$ : OLS

$\gamma \rightarrow \infty$ : IV solution (if identifiable)

$\gamma \rightarrow \infty$ : best invariant predictor (if not identifiable)

- Anchor regression minimizes worst case prediction error under shift interventions.

Rothenhäusler, Bühlmann, Meinshausen, JP (arXiv:1801.06229)

## Idea 4: anchor regression

$Y \in \mathbb{R}^1$ : target

$X \in \mathbb{R}^{1 \times d}$ : predictors

$A \in \mathbb{R}^{1 \times q}$ : anchors,  $EA^t A = Id$

$$b^\gamma := \underset{b}{\operatorname{argmin}} \underbrace{\mathbb{E}(Y - Xb)^2}_{\text{prediction}} + \gamma \underbrace{\|EA^t(Y - Xb)\|_2^2}_{\text{invariance}}$$

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- Anchor regression minimizes worst case prediction error under shift interventions.  
Rothenhäusler, Bühlmann, Meinshausen, JP (arXiv:1801.06229)
- The finite sample estimator is known as a  $k$ -class estimator for IV solution.  
Theil (1958), Nagar (1959), Jakobsen and JP (arXiv:2005.03353)

## Idea 4: anchor regression

$$\begin{pmatrix} X \\ Y \\ H \end{pmatrix} \leftarrow B \cdot \begin{pmatrix} X \\ Y \\ H \end{pmatrix} + \varepsilon + MA,$$

shifted:  $\begin{pmatrix} X^\nu \\ Y^\nu \\ H^\nu \end{pmatrix} \leftarrow B \cdot \begin{pmatrix} X^\nu \\ Y^\nu \\ H^\nu \end{pmatrix} + \varepsilon + \nu.$

$Id - B$  invertible

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$Id - B$  invertible

### Theorem

For any  $b \in \mathbb{R}^d$  we have

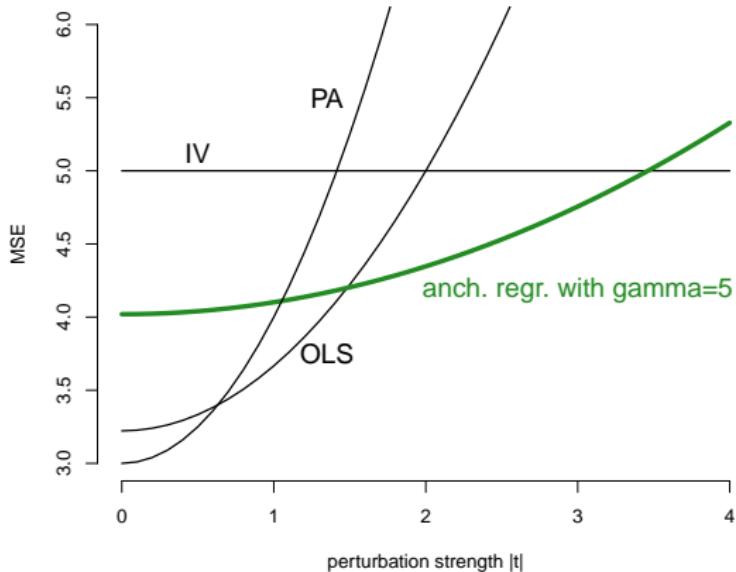
$$\operatorname{argmin}_b E(Y - Xb)^2 + \gamma \|EA^t(Y - Xb)\|_2^2 = \max_{\nu \in C^\gamma} \mathbb{E}[(Y^\nu - X^\nu b)^2],$$

where

$$C^\gamma := \{\nu = M\delta \text{ such that } \|\delta\|_2 \leq \sqrt{\gamma}\}.$$

## Idea 4: anchor regression

MSE under a **shift** of  $X$  ( $\gamma$  fixed):



## Idea 4: anchor regression



[http://www.srfcdn.ch/radio/modules/dynimages/624/srf-1/2015/01/diverses/264377.150114\\_raclette\\_key.jpg](http://www.srfcdn.ch/radio/modules/dynimages/624/srf-1/2015/01/diverses/264377.150114_raclette_key.jpg)

## Idea 4: anchor regression

Example: Maillard reaction

Glu, Mel, C5, ForAc, Triose, Cn, AcAc, Amad, lysR, Fru, AMP

## Idea 4: anchor regression

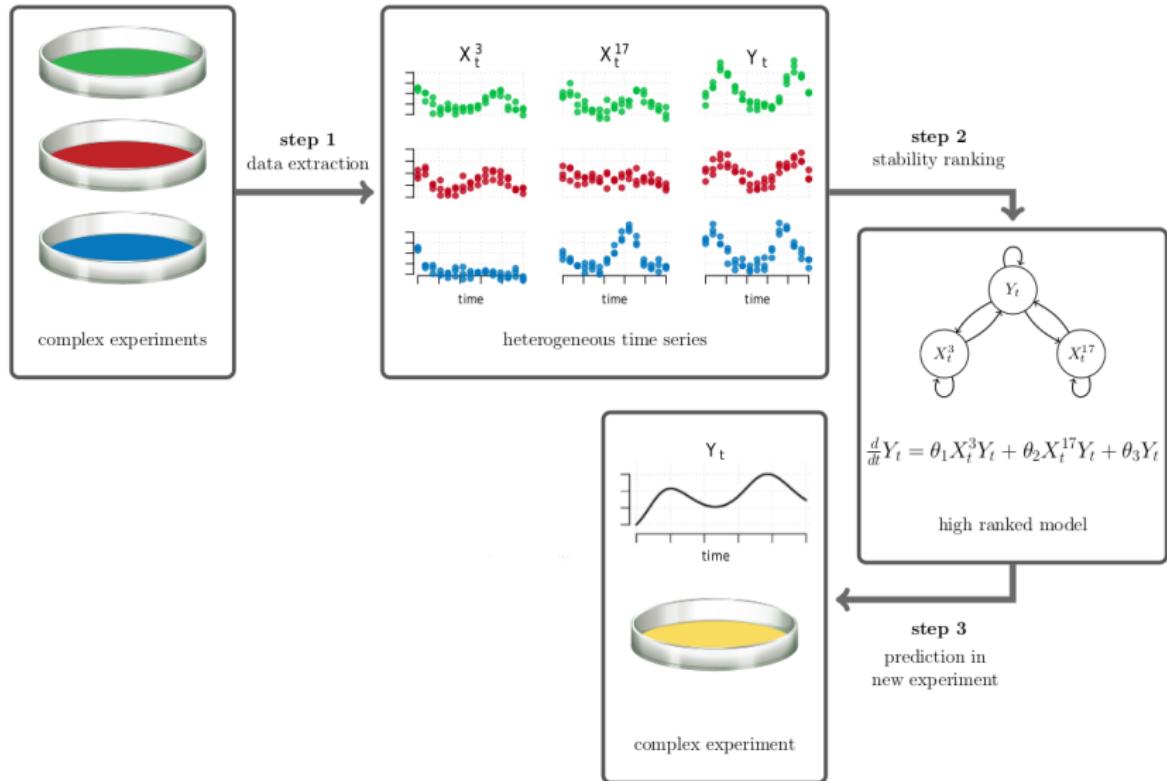
Example: Maillard reaction

Glu, Mel, C5, ForAc, Triose, Cn, AcAc, Amad, lysR, Fru, AMP

$$\frac{d}{dt}[\text{Glu}]_t = -\theta_1[\text{Glu}]_t + \theta_2[\text{Fru}]_t - \theta_3[\text{lysR}]_t[\text{Glu}]_t$$

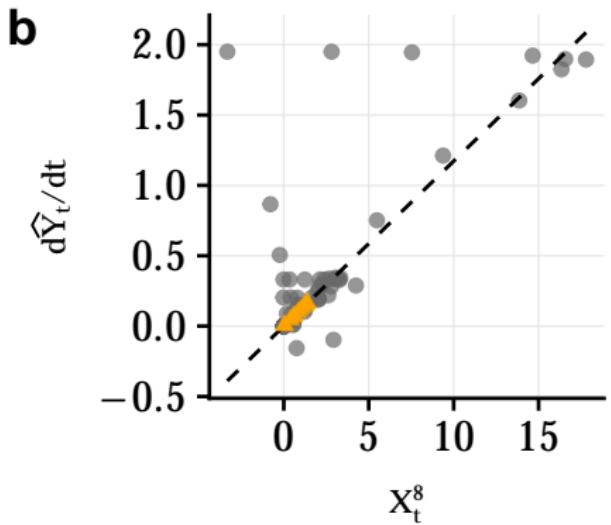
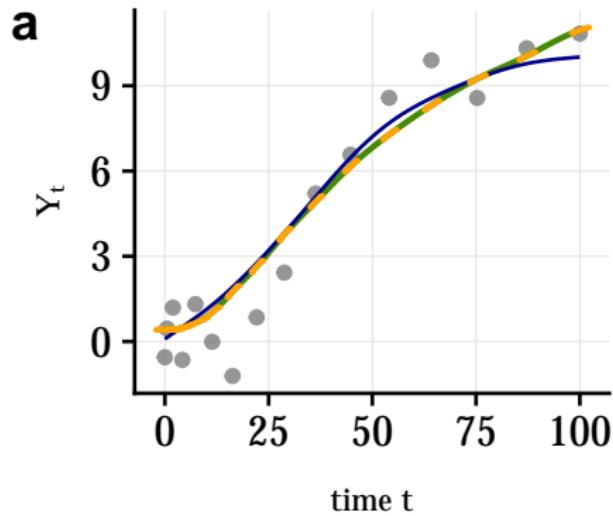
$$\frac{d}{dt}[\text{Mel}]_t = \theta_4[\text{AMP}]_t \quad \dots$$

# Idea 4: anchor regression



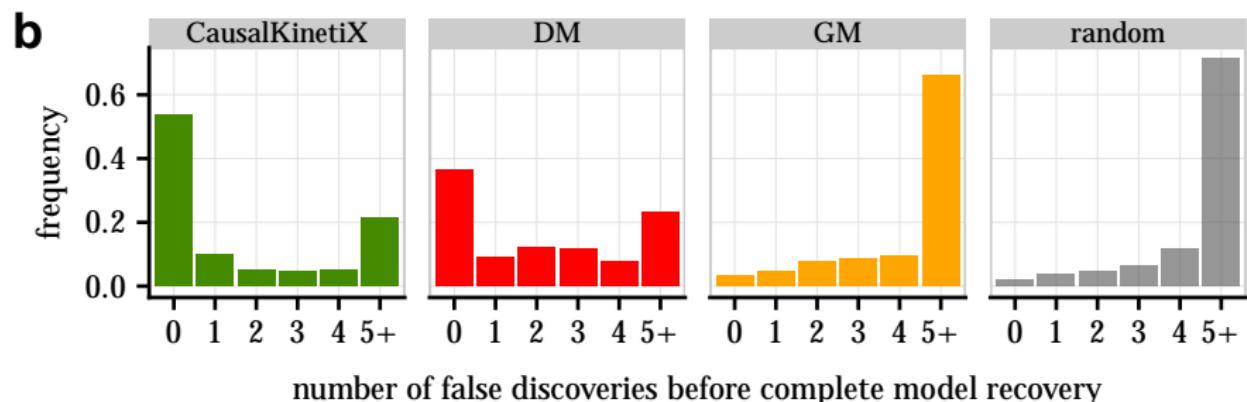
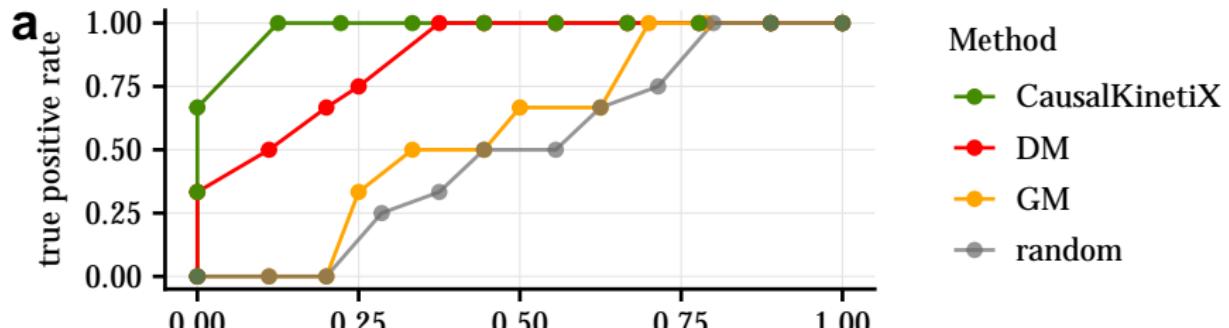
N. Pfister, S. Bauer, JP: *Identifying Causal Structure in Large-Scale Kinetic Systems*, PNAS 2019

## Idea 4: anchor regression

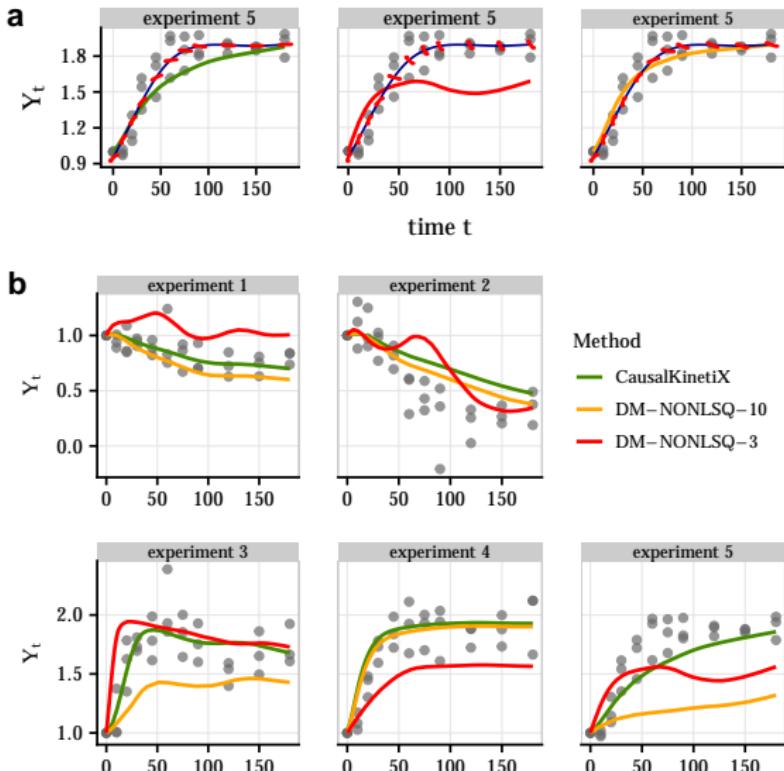


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# Idea 4: anchor regression



# Idea 4: anchor regression

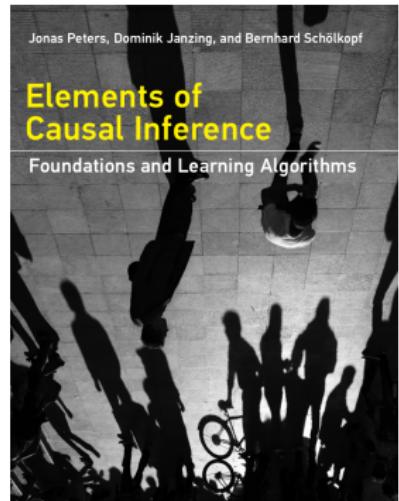


## **Summary Part III:**

- Idea 1: reformulate reinforcement learning,  
use causal structure
- Idea 2: semi-supervised learning from cause  
to effect does not work
- Idea 3: half-sibling regression
- Idea 4: anchor regression

## Summary Part III:

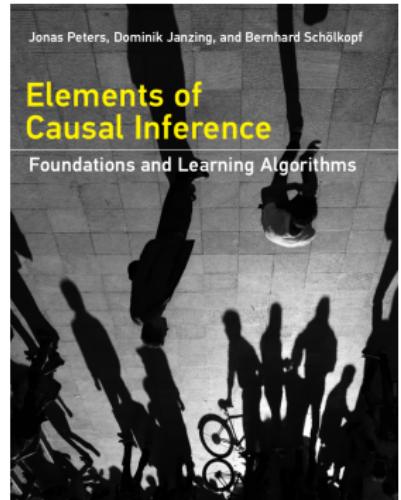
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For an exhaustive list of references, download pdf of  
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## Summary Part III:

- Idea 1: reformulate reinforcement learning, use causal structure
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