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Pairwise stochastic approximation for categorical factor models

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Problem Setting¹

Model setup:

- **Ordinal** data $\mathbf{Y} \in \mathbb{R}^p$, with $Y_i \in \{1, \dots, m_i\}$, $i = 1, \dots, p$.



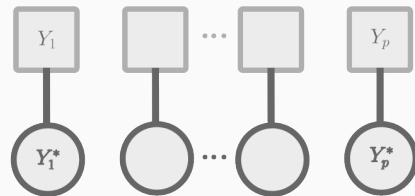
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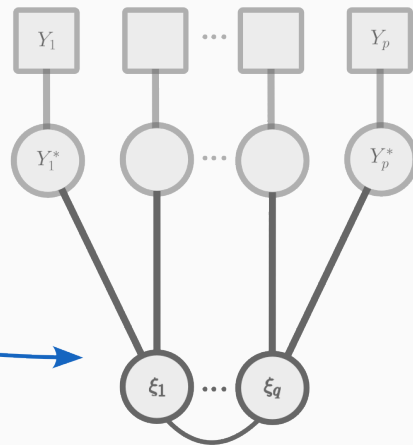
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- **Underlying linear factor model:**

$$\mathbf{Y}^* = \mathbf{\Lambda} \boldsymbol{\xi} + \boldsymbol{\delta},$$

- $\boldsymbol{\delta} \sim \mathcal{N}_p(0, \boldsymbol{\Sigma}_\delta)$ and $\boldsymbol{\xi} \sim \mathcal{N}_q(0, \boldsymbol{\Sigma}_\xi)$;
- $\mathbf{\Lambda} = (\lambda_{ij})_{p \times q}$;
- $\boldsymbol{\Sigma}_\xi = (\sigma_{rs})_{q \times q}$, $\sigma_{rr} = 1$ for $r = 1, \dots, q$



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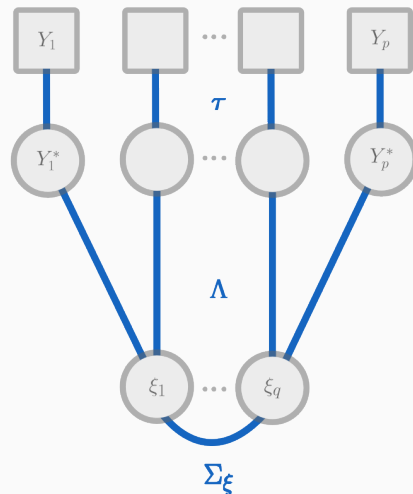
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- θ collects all the **free parameters** in $\mathbf{\Lambda}$, $\boldsymbol{\Sigma}_\xi$ and $\boldsymbol{\tau}$.



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Consider the **joint p-variate probability** of observing the pattern $\mathbf{c} = (c_1, \dots, c_p)$

$$\pi_{\mathbf{c}}(\boldsymbol{\theta}) = P(\mathbf{Y} = \mathbf{c}; \boldsymbol{\theta}) = \int_{\tau_{c_1-1}^{(1)}}^{\tau_{c_1}^{(1)}} \cdots \int_{\tau_{c_p-1}^{(p)}}^{\tau_{c_p}^{(p)}} \phi_p(\mathbf{y}^*; \boldsymbol{\Lambda} \boldsymbol{\Sigma}_{\xi} \boldsymbol{\Lambda}^{\top} + \boldsymbol{\Sigma}_{\delta}) d\mathbf{y}^*,$$

where $\phi_p(x; \boldsymbol{\Sigma})$ is the density of a p -dimensional normal distribution at x , with mean zero and variance $\boldsymbol{\Sigma}$.

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It requires evaluating a **p-dimensional integral!**

Pairwise Likelihood inference² (i)

Pairwise Likelihood Estimator:

$$\hat{\boldsymbol{\theta}}_{PML} = \arg \max_{\boldsymbol{\theta}} pl(\boldsymbol{\theta}), \quad \text{where} \quad pl(\boldsymbol{\theta}) = \sum_{i < j} \sum_{c_i=1}^{m_i} \sum_{c_j=1}^{m_j} n_{c_i c_j}^{(ij)} \log \left\{ \pi_{c_i c_j}^{(ij)}(\boldsymbol{\theta}) \right\},$$

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- ✓ $\sqrt{n} \left(\hat{\boldsymbol{\theta}}_{PML} - \boldsymbol{\theta}^* \right) \xrightarrow{d} \mathcal{N} \left(\mathbf{0}, \mathbf{H}^{-1} \mathbf{J} \mathbf{H}^{-1} \right)$, with $\mathbf{H} = E \{ -\nabla^2 p\ell(\boldsymbol{\theta}^*) / n \}$ and $\mathbf{J} = \text{Var} \{ \nabla p\ell(\boldsymbol{\theta}^*) / n \}$.

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- ✗ The number of summands grows with **$O(K)$** , where K is the number of pairs, i.e. $O(p^2)$.

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Do we need all the pairs?

Pairwise gradient:
$$\nabla p\ell(\boldsymbol{\theta}) = \sum_{i < j} \sum_{c_i=1}^{m_i} \sum_{c_j=1}^{m_j} \frac{n_{c_i c_j}^{(ij)}}{\pi_{c_i c_j}^{(ij)}(\boldsymbol{\theta})} \nabla \pi_{c_i c_j}^{(ij)}(\boldsymbol{\theta}).$$

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Consider a $p \times p$ matrix \mathbf{W} , with generic element $W_{ij} \in \{0, 1\}$ and let $E\{W_{ij}\} = \nu/K$.

Stochastic approximations - Gradients

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Stochastic gradient:
$$S(\boldsymbol{\theta}; \mathbf{W}) = \frac{K}{\nu} \sum_{i < j} \mathbf{W}_{ij} \sum_{c_i=1}^{m_i} \sum_{c_j=1}^{m_j} \frac{n_{c_i c_j}^{(ij)}}{\pi_{c_i c_j}^{(ij)}(\boldsymbol{\theta})} \nabla \pi_{c_i c_j}^{(ij)}(\boldsymbol{\theta}).$$

The key feature is $E_{\mathbf{W}} \{S(\boldsymbol{\theta}; \mathbf{W})\} = \nabla p\ell(\boldsymbol{\theta})$, where $S(\boldsymbol{\theta}; \mathbf{W})$ accounts for ν pairs on average.

Stochastic optimization:³⁴

Given a suitable decreasing scheduling for η_t , at each iteration $t = 1, \dots, T_n$ repeat

1. **Sample** \mathbf{W}_t by drawing ν out of K pairwise components;
2. **Update** the current estimate via $\boldsymbol{\theta}_t = \boldsymbol{\theta}_{t-1} - \frac{\eta_t}{n} S(\boldsymbol{\theta}_{t-1}; \mathbf{W}_t)$;

Trajectories averaging via $\bar{\boldsymbol{\theta}} = \frac{1}{T_n} \sum_{t=1}^{T_n} \boldsymbol{\theta}_t$.

³ Herbert Robbins and Sutton Monro. "A stochastic approximation method". In: *The Annals of Mathematical Statistics* 22.3 (1951), pp. 400–407.

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Stochastic approximation - Asymptotics (i)

Let T_n diverge with $n \rightarrow \infty$. Then,

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Covariance matrix:

$$\Omega_n = \underbrace{\frac{1}{n} \mathbf{H}^{-1} \mathbf{J} \mathbf{H}^{-1}}_{\text{same as } \hat{\theta}_{\text{PML}}} + \frac{\text{NOISE}}{T_n}.$$

The **NOISE** depends on $\text{Var}\{n^{-1}S(\theta; \mathbf{W})\}$; thus on the distribution of \mathbf{W} and the chosen ν .

Recall that
$$\frac{1}{n}S(\boldsymbol{\theta}; \mathbf{W}) = \frac{K}{n\nu} \sum_{i < j} W_{ij} \sum_{c_i=1}^{m_i} \sum_{c_j=1}^{m_j} \frac{n_{c_i c_j}^{(ij)}}{\pi_{c_i c_j}^{(ij)}(\boldsymbol{\theta})} \nabla \pi_{c_i c_j}^{(ij)}(\boldsymbol{\theta}).$$

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When $W_{ij} = 1$, we account for all n observations on the (i, j) pair. It can be shown that such special behaviour implies that **NOISE** = $O(1/n)$.

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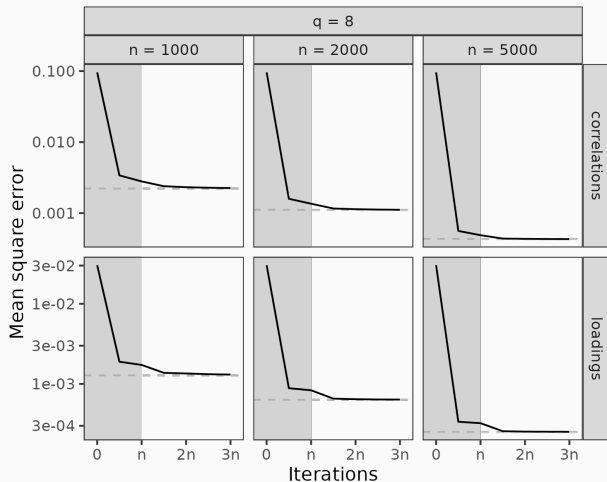
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Asymptotic equivalence between $\bar{\boldsymbol{\theta}}$ and $\hat{\boldsymbol{\theta}}_{\text{PML}}$

$$\boldsymbol{\Omega}_n = \frac{1}{n} \mathbf{H}^{-1} \mathbf{J} \mathbf{H}^{-1} + O\left(\frac{1}{nT_n}\right) \approx \frac{1}{n} \mathbf{H}^{-1} \mathbf{J} \mathbf{H}^{-1}.$$

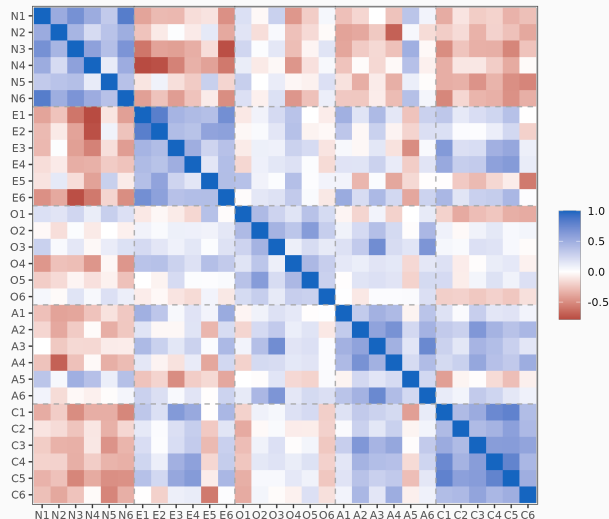
Simulation experiments

- $p = 40$ items, $q = 8$ latent traits with simple loading structure augmented with $q - 1$ cross-loadings;
- $m = 4$ categories per item;
- $n \in \{1000, 2000, 5000\}$;
- $\nu = 8$ i.e. 8 out of $K = 780$ pairs per iteration;
- Replications = 1000.



Big Five Application

- $p = 120$ items, $m = 5$ categories per item;
- $q = 30$ latent traits that we expect to be grouped in **N**euroticism, **A**greeableness, **E**xtraversion, **O**penness to experience, **C**onscientiousness;
- $n = 410,376$. The 60% is used for training;
- $\nu = 1$ out of $K = 7140$ pairs per iteration.



Thank you for your attention!



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🔗 [giuseppealfonzetti](https://github.com/giuseppealfonzetti)

🌐 giuseppealfonzetti.github.io





📖 Alfonzetti et al., "Pairwise stochastic approximation for confirmatory factor analysis of categorical data", *BJMSP* (2024)



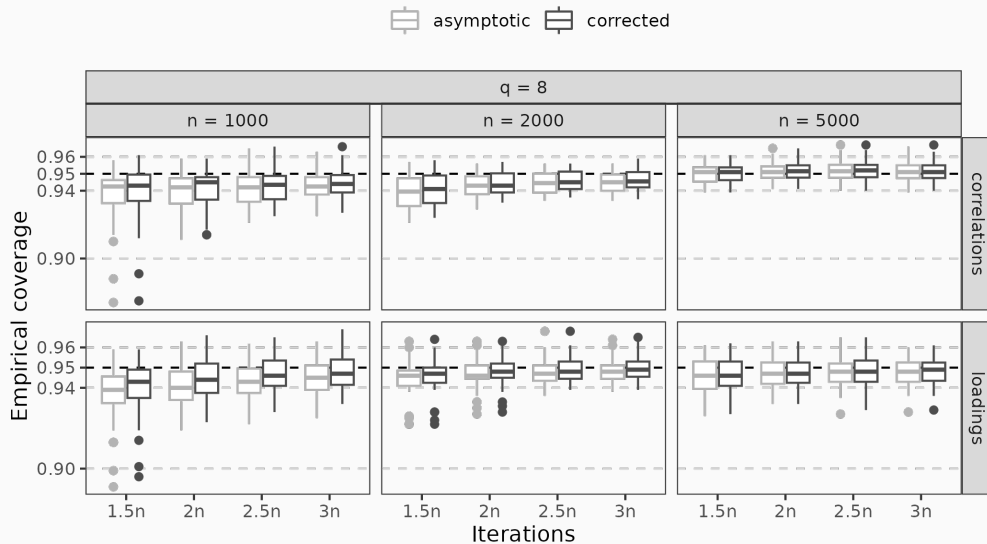
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Simulation experiments - Empirical coverage



Big Five - Estimation trajectories

