

Bayes' Theorem: $f(y|x) = \frac{f(x|y)f(y)}{f(x)}$

Law of Tot. Prob:
$$\begin{cases} f(y) = \sum_x f(y|x)f(x) & (\text{discrete}) \\ f(y) = \int f(y|x)f(x)dx & (\text{cts}) \end{cases}$$

Ex. $X \sim \text{Exp}(\lambda)$

$Y|X=x \sim \text{Pois}(x)$

what is dist of Y ?

$$f(y) = \int f(y|x)f(x)dx \quad \leftarrow$$

$$= \int_0^{\infty} \underbrace{\frac{x^y e^{-x}}{y!}}_{\text{Pois}(x)} \underbrace{\lambda e^{-\lambda x}}_{\text{Exp}(\lambda)} dx$$

$$= \frac{1}{y!} \int_0^{\infty} \lambda e^{-(\lambda+1)x} x^y dx$$

PDF trick notice it's basically $\text{Gamma}(y+1, \lambda+1)$

$$= \frac{\lambda \Gamma(y+1)}{y! (\lambda+1)^{y+1}} \int_0^{\infty} \frac{(\lambda+1)^{y+1} x^y e^{-(\lambda+1)x}}{\Gamma(y+1)} dx$$

distribution of Gamma(y+1, λ+1)

$$= \frac{\lambda \cancel{\Gamma(y+1)}}{\cancel{y!} (\lambda+1)^{y+1}} = \frac{\lambda}{(\lambda+1)^{y+1}} \quad \text{for } y=0, 1, 2, \dots$$

Ex.

→ $Y \sim \text{Pois}(\lambda)$

$X|Y=y \sim \text{Bin}(y, p)$

$p \in [0, 1]$

what is the dist of X ?

$y \geq x$

$$f(x) = \sum_y f(x|y) f(y)$$

$$= \sum_{y=x}^{\infty} \binom{y}{x} p^x (1-p)^{y-x} \frac{\lambda^y e^{-\lambda}}{y!}$$

$$\binom{y}{x} \frac{1}{y!} = \frac{y!}{x!(y-x)!} \frac{1}{y!} = \frac{1}{x!(y-x)!}$$

$$x, y; \quad x!(y-x)!, \quad x!(y-x)!$$

$$\rightarrow = \frac{p^x e^{-\lambda}}{x!} \sum_{y=x}^{\infty} \frac{(1-p)^{y-x}}{(y-x)!} \lambda^{y-x}$$

$$= \frac{p^x e^{-\lambda} \lambda^x}{x!} \left[\sum_{y=0}^{\infty} \frac{(1-p)^y}{y!} \lambda^y \right] = e^{(1-p)\lambda}$$

$$= \frac{(\lambda p)^x e^{-\lambda}}{x!} e^{(1-p)\lambda}$$

$$f(x) = \frac{(\lambda p)^x e^{-\lambda p}}{x!}$$

← Pois(λp)

$$\text{So } X \sim \text{Pois}(\lambda p)$$

Theorem: Iterated Expectation

If X and Y are RVs then

$$E[X] = E[E[X|Y]]$$

La f. y

$$E[X|Y=y] = \int x f(x|y) dx \in \mathbb{R}$$

For each $y \in \mathbb{R}$ I get a potentially different $\mathbb{E}[X|Y=y]$

This defines a function

$$g(y) = \mathbb{E}[X|Y=y]$$

what happens I plug Y into g ?

$$\begin{aligned} g(Y) &= \mathbb{E}[X|Y=Y] \quad \text{weird notation} \\ &= \mathbb{E}[X|Y] \end{aligned}$$

- ① $\mathbb{E}[X|Y=y]$ is a number
- ② $\mathbb{E}[X|Y]$ is a RV (a fun of Y)

pf. Facts: (cts)

$$\textcircled{1} f(x) = \int f(x, y) dy$$

$$\textcircled{2} f(x|y) = \frac{f(x, y)}{f(y)} \Leftrightarrow f(x, y) = f(x|y) f(y)$$

$$\textcircled{3} \mathbb{E}[X|Y=y] = \int x f(x|y) dx$$

$$(3) \mathbb{E}[X|Y=y] = \int x f(x|y) dx$$

$$\mathbb{E}[X] = \int x f(x) dx \stackrel{(1)}{=} \int x \int f(x,y) dy dx$$

$$\stackrel{(2)}{=} \int x \int f(x|y) f(y) dy dx$$

$$= \int \underbrace{x f(x|y)}_{\mathbb{E}[X|Y=y]=g(y)} f(y) dy \quad [\text{rearrange}]$$

$$\mathbb{E}[X|Y=y] = g(y)$$

$$= \int g(y) f(y) dy$$

$$= \mathbb{E}[g(Y)] \quad \swarrow \text{notation}$$

$$= \mathbb{E}[\mathbb{E}[X|Y]]$$

Ex.

$$Y \sim \text{Pois}(\lambda)$$

$$X|Y=y \sim \text{Bin}(y, p)$$

$$\mathbb{E}[X]? \quad (\mathbb{E}X = \mathbb{E}[\mathbb{E}[X|Y]])$$

$$(1) \text{ get } \mathbb{E}[X|Y=y] = yp$$

$$(2) \text{ get } \mathbb{E}[X|Y] = Yp$$

$$\textcircled{3} \text{ get } E[E[X|Y]] = E[Y/p] = p E[Y] \\ \boxed{= p \lambda} \\ \uparrow \text{ answer}$$

ΣX , $P \sim \text{Beta}(\alpha, \beta)$

$$X|P=p \sim \text{Bin}(n, p)$$

$$E[X] ? = E[E[X|P]]$$

$$\textcircled{1} \text{ get } E[X|P=p] = np$$

$$\textcircled{2} \text{ get } E[X|P] = nP$$

$$\textcircled{3} \text{ get } E[E[X|P]] = E[nP] = n E[P]$$

$$\boxed{= n \frac{\alpha}{\alpha + \beta}}$$

Theorem: Law of Total Variance

$$\text{Var}(X) = E[\text{Var}(X|Y)] + \text{Var}(E[X|Y])$$

calc. $\text{Var}(X|Y=y)$...

then promote y to y
in formula

Ex.

$$P \sim \text{Beta}(\alpha, \beta)$$

$$X|P=p \sim \text{Bin}(n, p)$$

Var(X)?

$$(1) E[X|P=p] = np$$

$$\text{Var}(X|P=p) = np(1-p)$$

$$(2) E[X|P] = nP$$

$$\text{Var}(X|P) = nP(1-P)$$

$$(3) \text{Var}(X) = E \text{Var}(X|P) + \text{Var} E[X|P]$$

$$= E[nP(1-P)] + \text{Var}(nP)$$

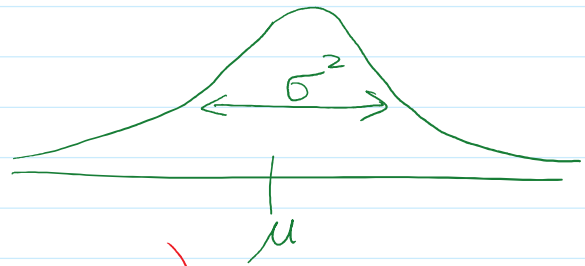
$$= n(\underbrace{E[P]} - \underbrace{E[P^2]}) + n^2 \underbrace{\text{Var}(P)}$$

= ... plug in

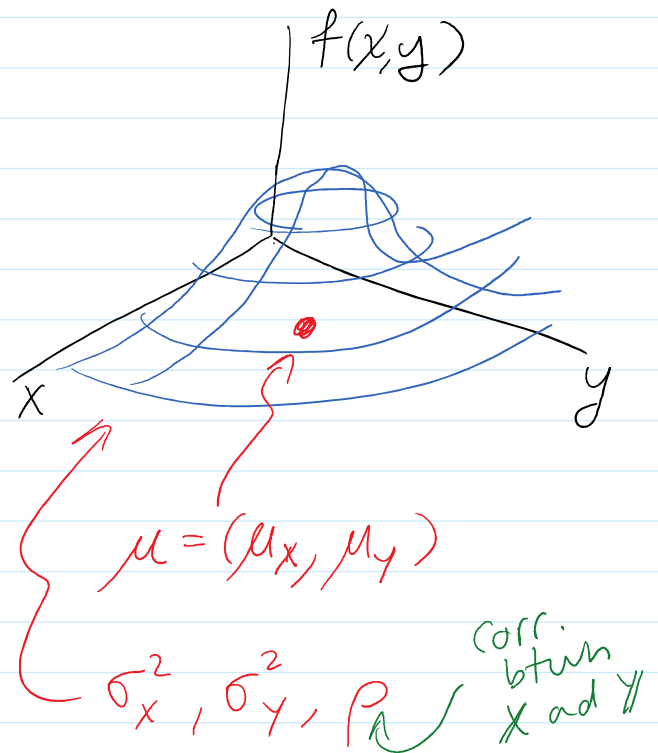
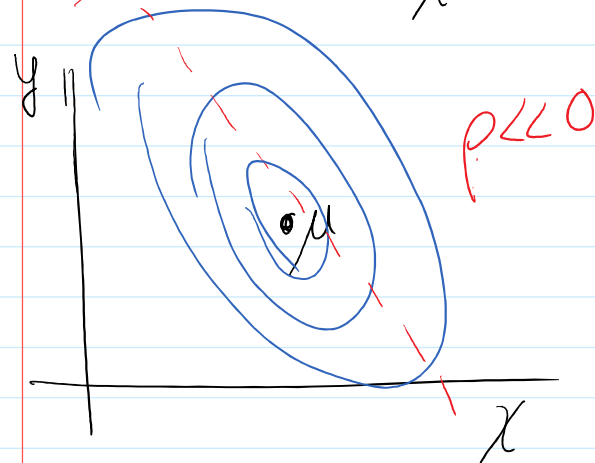
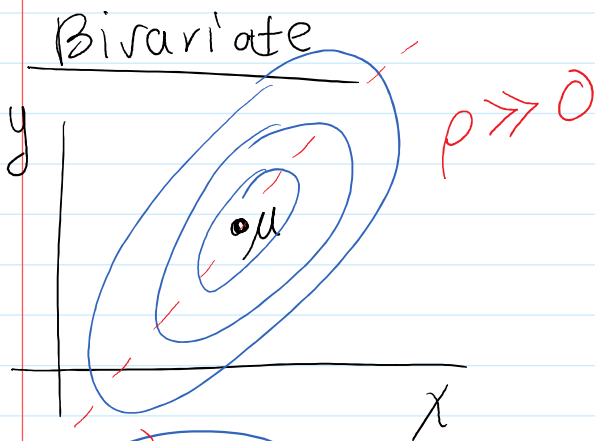
$$= n \frac{\alpha \beta}{(\alpha + \beta)(\alpha + \beta + 1)} + n^2 \frac{\alpha \beta}{(\alpha + \beta)^2 (\alpha + \beta + 1)}$$

Bivariate Normal

univariate: $N(\mu, \sigma^2)$



$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x-\mu)^2\right)$$



PDF: $(X, Y) \sim \text{BiVN}(\mu_x, \mu_y, \sigma_x^2, \sigma_y^2, \rho)$

$$f(x, y) = \frac{1}{2\pi \sigma_x \sigma_y \sqrt{1-\rho^2}} \exp \left\{ -\frac{1}{2} \frac{1}{1-\rho^2} \left[\left(\frac{x-\mu_x}{\sigma_x} \right)^2 + \left(\frac{y-\mu_y}{\sigma_y} \right)^2 - 2\rho \left(\frac{x-\mu_x}{\sigma_x} \right) \left(\frac{y-\mu_y}{\sigma_y} \right) \right] \right\}$$

mean vector

$$\mu = (\mu_x, \mu_y) \in \mathbb{R}^2$$

covariance matrix

$$\Sigma = \begin{bmatrix} \sigma_x^2 & \sigma_x \sigma_y \rho \\ \sigma_x \sigma_y \rho & \sigma_y^2 \end{bmatrix} = \begin{bmatrix} \text{Var}(X) & \text{Cov}(X, Y) \\ \text{Cov}(X, Y) & \text{Var}(Y) \end{bmatrix}$$
$$z = (x, y)$$

Then

$$f(z) = \frac{1}{2\pi} \frac{1}{\sqrt{\det \Sigma}} \exp \left(-\frac{1}{2} (z - \mu)^T \Sigma^{-1} (z - \mu) \right)$$

Uni: $f(x) = \frac{1}{2\pi} \frac{1}{\sqrt{\sigma^2}} \exp \left(-\frac{1}{2} (x - \mu) (\sigma^2)^{-1} (x - \mu) \right)$

Facts: ① $X \sim N(\mu_x, \sigma_x^2)$

$$Y \sim N(\mu_y, \sigma_y^2)$$

$$\textcircled{2} \text{Cor}(X, Y) = \rho$$

$$\textcircled{3} aX + bY \sim N(a\mu_x + b\mu_y, a^2\sigma_x^2 + b^2\sigma_y^2 + 2ab\sigma_x\sigma_y\rho)$$

④ A characterization of BivN

$$(X, Y) \sim \text{BivN} \Leftrightarrow \forall a, b \quad aX + bY \sim N(\dots)$$

⑤ Prev: If $X \perp Y$ then $\text{Cor}(X, Y) = 0$

Partial Converse: if $(X, Y) \sim \text{BivN}$

$$\text{and } \text{Cor}(X, Y) = 0$$

then $X \perp Y$.