

Additional Office Hours : Friday 10-11 am
Monday 12-1 pm

Theorem: F is the CDF of some RV iff

- ① $\lim_{x \rightarrow \infty} F(x) = 1$ and $\lim_{x \rightarrow -\infty} F(x) = 0$
- ② F is non-decreasing
- ③ F is right-continuous

Ex. Let

$$F(x) = \frac{1}{1 + e^{-x}} \quad \forall x \in \mathbb{R}$$

Q: Is this the CDF of some RV?



check 3 conditions:

$$\textcircled{1} \quad \lim_{x \rightarrow \infty} F(x) = \frac{1}{1 + e^{-\infty}} = \frac{1}{1 + 0} = 1$$

$$\lim_{x \rightarrow -\infty} F(x) = \frac{1}{1 + e^{-(-\infty)}} = \frac{1}{1 + \infty} = 0$$

$$\lim_{x \rightarrow -\infty} F(x) = \frac{1}{1 + e^{-(-\infty)}} = \frac{1}{\infty} = 0$$

② Non-decreasing?

$$\frac{dF}{dx} = \frac{d}{dx} \left(\frac{1}{1 + e^{-x}} \right) = \frac{e^{-x}}{(1 + e^{-x})^2} > 0$$

(increasing)

③ Right-continuous?

Yes its a cts fn.

So F is the CDF of some RV.

Defn: Identically Distributed RVs

We say two RVs X and Y are equal in distribution if $A \subset \mathbb{R}$

$$P(X \in A) = P(Y \in A).$$

We denote this as

$$X \stackrel{d}{=} Y$$

This doesn't mean $X = Y$.

Ex. 3 coin flips.

$X = \# \text{ heads}$ and $Y = \# \text{ tails}$.

$$X(\text{HTT}) = 1 \quad \text{and} \quad Y(\text{HTT}) = 2$$

but $X \stackrel{d}{=} Y$.

$$P(X=0) = 1/8 = P(Y=0)$$

$$P(X=1) = 3/8 = P(Y=1)$$

\vdots

Theorem:

$$X \stackrel{d}{=} Y \quad \text{iff} \quad F_X = F_Y.$$

\uparrow
CDF of X

\uparrow
CDF of Y

Ex. Toss coins (independently) until a H appears.

$$S = \{H, TH, TTH, TTTT, \dots\}$$

note $|S| = \infty$

Let p be prob of getting a H on any flip.

Let $X = \# \text{ flips until I get a H}$

$\omega \in S$	$X(\omega)$
H	1
TH	2
TTH	3
TTTH	4
\vdots	\vdots

Q: What is CDF of X ? $F(x) = P(X \leq x)$

we'll look at $P(X = x)$.

Let $H_i = i^{\text{th}}$ toss is a H and $T_i = H_i^c$

then

$$W_i = "X = i" = \underbrace{T_1 T_2 T_3 \dots T_{i-1}}_{\text{independent}} H_i$$

then

$$\begin{aligned} P(X = i) &= P(W_i) = P(T_1 \dots T_{i-1} H_i) \\ &= P(T_1) \dots P(T_{i-1}) P(H_i) \\ &= (1-p) \dots (1-p) p \\ &= (1-p)^{i-1} p \end{aligned}$$

$$X \leq x = W_1 \cup W_2 \cup W_3 \cup \dots \cup W_x$$

disjoint union

$$F(x) = P(X \leq x) = \sum_{i=1}^x P(W_i) = \sum_{i=1}^x (1-p)^{i-1} p$$

geometric sum

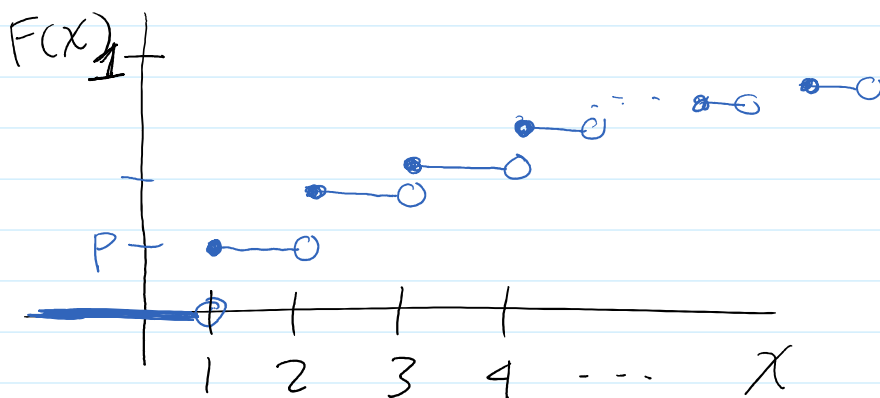
$$\sum_{i=0}^{n-1} r^i = \frac{1-r^n}{1-r}$$

$$= p \sum_{i=0}^{x-1} \frac{(1-p)^i}{1}$$

$$= p \frac{1-(1-p)^x}{1-(1-p)}$$

The RV w/ this CDF is called a Geometric RV.

$$F(x) = 1 - (1-p)^x$$



$$F(x) = \begin{cases} 1 - (1-p)^{\lfloor x \rfloor} & x \geq 1 \\ 0 & x < 1 \end{cases}$$

$\lfloor x \rfloor = \text{round down}$
 $\lfloor 3.14 \rfloor = 3$

Defn: Discrete / Continuous

Defn: Discrete / Continuous

A discrete RV has a CDF that is a step fu.

A continuous RV has a continuous CDF.

Defn: Probability Mass Function

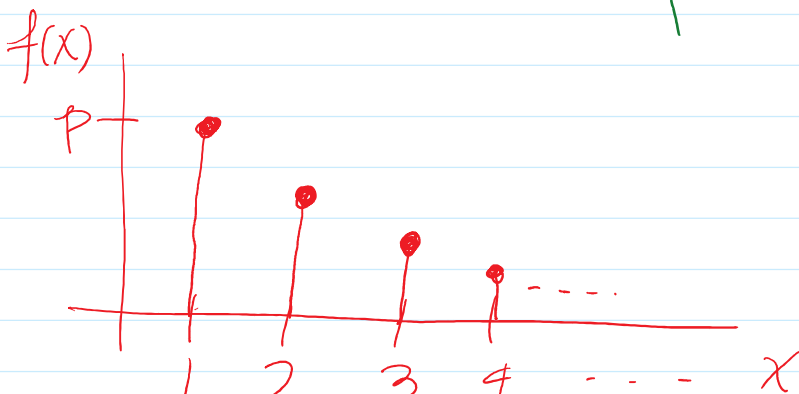
For a discrete RV X the probability Mass function (PMF) is a function $f: \mathbb{R} \rightarrow \mathbb{R}$ where

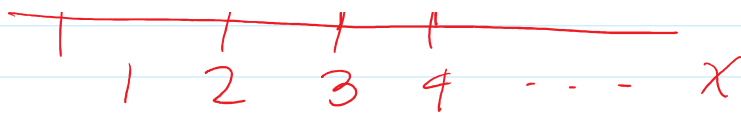
$$f(x) = P(X=x) \quad \forall x \in \mathbb{R}.$$

Also called the distribution of X .

Ex. Prev. ex.

$$f(x) = P(X=x) = \begin{cases} (1-p)^{x-1} p, & x=1, 2, 3, \dots \\ 0, & \text{else} \end{cases}$$





Theorem! For discrete RVs

$$F(x) = \sum_{i \leq x} f(i)$$

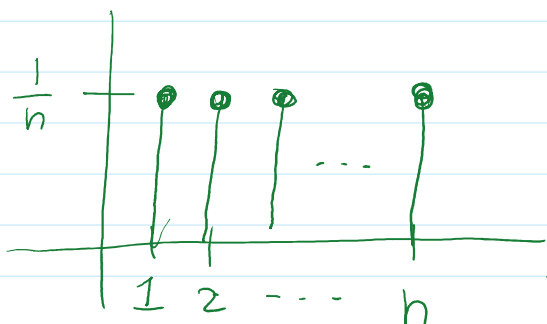
Pf. $"X \leq x" = \bigcup_{i \leq x} "X = i"$
 $\quad \quad \quad \uparrow$ disjoint union

So

$$F(x) = P(X \leq x) = P\left(\bigcup_{i \leq x} "X = i"
$$= \sum_{i \leq x} \underbrace{P(X = i)}_{f(i)}$$$$

Ex. We say X has a discrete uniform distribution over $1, \dots, n$

if $f(i) = \begin{cases} \frac{1}{n}, & \text{for } i = 1, \dots, n \\ 0, & \text{else} \end{cases}$



notation:

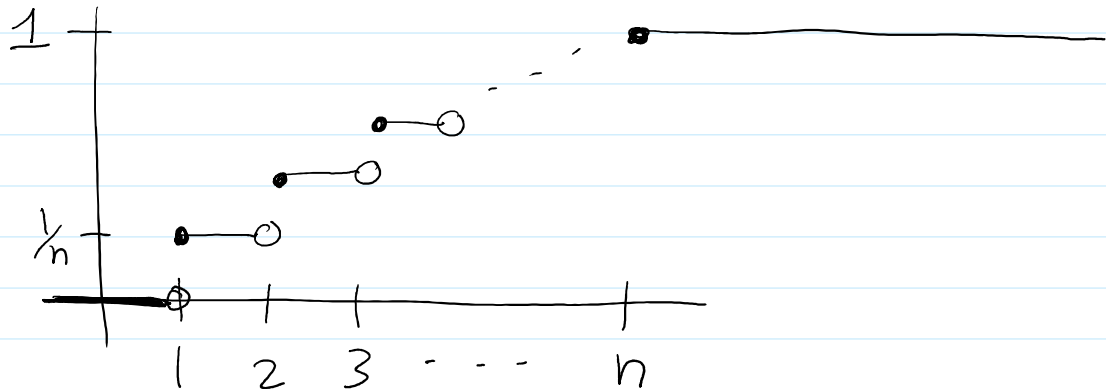
$$X \sim U(\{1, \dots, n\})$$

\uparrow
read! distributed as

Q: what is the CDF?

$$F(x) = \sum_{i \leq x} f(i) = \sum_{i=1}^x \frac{1}{n} = \frac{x}{n}$$

$x \in \{1, \dots, n\}$



$$F(x) = \begin{cases} 0, & x < 1 \\ \lfloor x \rfloor / n, & 1 \leq x \leq n \\ 1, & x \geq n \end{cases}$$

Saw: $F(x) = \sum_{i \leq x} f(i)$
 \parallel
 $P(X \leq x)$

More generally: $A \subset \mathbb{R}$,

$$P(X \in A) = \sum_{i \in A} f(i)$$

Ex. $X \sim U(\{1, \dots, 7\})$

$$\begin{aligned}
 P(2 \leq X \leq 5) &= P(X \in \{2, 3, 4, 5\}) \\
 &= \sum_{x=2}^5 \underbrace{f(x)}_{1/4} = 4/4
 \end{aligned}$$

Ex. Roll a die 60 times. (independently)
 $X = \#$ of 6s I roll.

What is f ?

$$\begin{aligned}
 f(0) &= P(X=0) = \underbrace{\left(\frac{5}{6}\right)\left(\frac{5}{6}\right)\left(\frac{5}{6}\right) \cdots \left(\frac{5}{6}\right)}_{60 \text{ times}} \\
 &= \left(\frac{5}{6}\right)^{60}
 \end{aligned}$$

$$\begin{aligned}
 f(1) &= P(X=1) = \binom{60}{1} \underbrace{\left(\frac{1}{6}\right)\left(\frac{5}{6}\right)\left(\frac{5}{6}\right) \cdots \left(\frac{5}{6}\right)}_{59 \text{ times}} \\
 &= \binom{60}{1} \left(\frac{1}{6}\right) \left(\frac{5}{6}\right)^{59}
 \end{aligned}$$

$$\begin{aligned}
 f(2) &= P(X=2) = \binom{60}{2} \underbrace{\left(\frac{1}{6}\right)\left(\frac{1}{6}\right)\left(\frac{5}{6}\right) \cdots \left(\frac{5}{6}\right)}_{58 \text{ times}} \\
 &= \binom{60}{2} \left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right)^{58}
 \end{aligned}$$

$$f(x) = \binom{60}{x} \left(\frac{1}{6}\right)^x \left(\frac{5}{6}\right)^{60-x}$$

We call this a Binomial RV

If I do a sequence of ⁿ Yes/No experiments (independently) and each a prob of p of being "Yes"

and $X = \#$ of Yes

We call X a Binomial RV

denoted $X \sim \text{Bin}(n, p)$,

$$f(x) = \binom{n}{x} p^x (1-p)^{n-x}$$

above:
 $n = 60$
 $p = 1/6$