

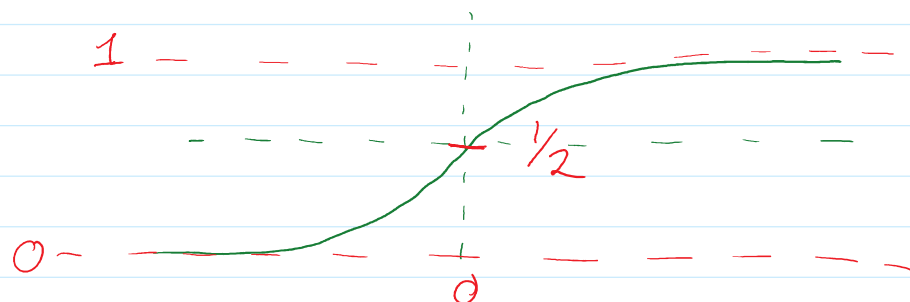
Additional Office Hours : Friday 10-11 am
Monday 12-1 pm

Theorem: F is the CDF of some RV iff

- ① $\lim_{x \rightarrow \infty} F(x) = 1$ and $\lim_{x \rightarrow -\infty} F(x) = 0$
- ② F is non-decreasing
- ③ F is right-continuous

Ex.

$$F(x) = \frac{1}{1 + e^{-x}} \quad \forall x \in \mathbb{R}$$



Is this a valid CDF?

Check our 3 conditions:

$$\textcircled{1} \lim_{x \rightarrow \infty} F(x) = \frac{1}{1 + e^{-\infty}} = \frac{1}{1 + 0} = 1$$

$$\lim_{x \rightarrow -\infty} F(x) = \frac{1}{1 + e^{-(-\infty)}} = \frac{1}{\infty} = 0$$

$$(2) \frac{d}{dx} F(x) = \frac{e^{-x}}{(1+e^{-x})^2} > 0$$

so increasing function

(3) Right cts.?

It is a continuous function.

So, yes, this is a valid CDF.

Defn: Identically Distributed RVs
(equal in distribution)

Two RVs X and Y are equal in distribution if $\forall A \subset \mathbb{R}$

$$P(X \in A) = P(Y \in A).$$

We denote this as $X \stackrel{d}{=} Y$.

This is different than $X = Y$ as functions.

Ex. 3 coin flips

$X = \# \text{ heads}$ and $Y = \# \text{ tails}$

notice:

$$X(\text{HTT}) = 1 \quad \text{but} \quad Y(\text{HTT}) = 2$$

however:

$$P(X=0) = 1/8 = P(Y=0)$$

$$P(X=1) = 3/8 = P(Y=1)$$

So $X \stackrel{d}{=} Y$ but $X \neq Y$ (as functions).

Theorem: $X \stackrel{d}{=} Y$ iff $F_X = F_Y$ (as fns).

\uparrow CDF of X \uparrow CDF of Y

Ex. Toss coins independently until a H appears.

$$S = \{H, TH, TTH, TTTH, \dots\}$$

$$\text{so } |S| = \infty.$$

Let p as the prob. of a H on any flip.

Let $X = \#$ of flips to get H.

$\omega \in S$	$X(\omega)$
H	1
TH	2
TTH	3

TH	2
TTH	3
TTTH	4
⋮	⋮

Q: What is the CDF of X ?

$$F(x) = P(X \leq x)$$

To determine F let's consider

$$P(X=x) \text{ for } x \in \mathbb{R}$$

it takes x flips to get H

Let $H_i = H$ on i^{th} flip and $T_i = H_i^c$.

Then

$$"X=x" = T_1 T_2 T_3 \cdots T_{x-1} H_x \quad \leftarrow \text{independent}$$

so

$$\begin{aligned} P(X=x) &= P(T_1 \cdots T_{x-1} H_x) \\ &= P(T_1) \cdots P(T_{x-1}) P(H_x) \\ &= \underbrace{(1-p) \cdots (1-p)}_{x-1 \text{ of these}} p \end{aligned}$$

$$= (1-p)^{x-1} p$$

Let $W_i = "X=i"$ so that $P(W_i) = (1-p)^{i-1} p$

notice that $W_i \cap W_j = \emptyset \quad i \neq j$

and

$$X \leq x = W_1 \cup W_2 \cup W_3 \cup \dots \cup W_x$$

so $F(x) = P(X \leq x) = P(W_1 \cup \dots \cup W_x)$

recall:

$$\sum_{i=0}^{n-1} r^i = \frac{1-r^n}{1-r}$$

$$= \sum_{i=1}^x P(W_i)$$

$$= \sum_{i=1}^x (1-p)^{i-1} p$$

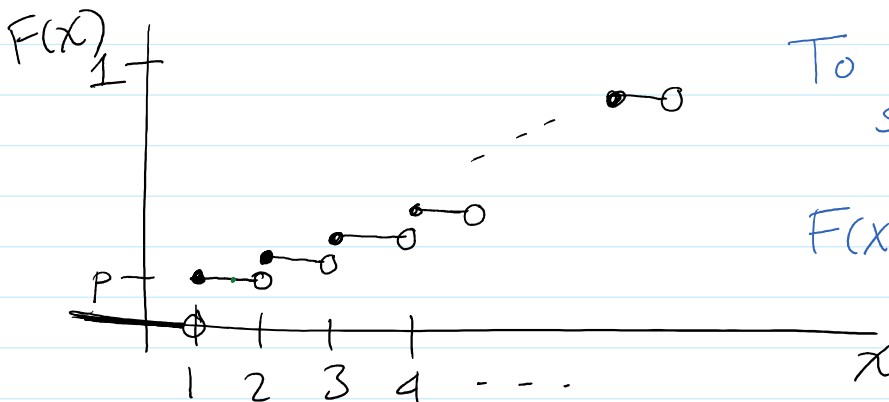
$$= p \sum_{i=0}^{x-1} (1-p)^i$$

$$r = 1-p$$

$$= p \left(\frac{1-(1-p)^x}{1-(1-p)} \right)$$

$$\frac{1-(1-p)}{1-(1-p)} = p$$

$$= 1 - (1-p)^x$$



To actually encode
step fn behavior

$$F(x) = \begin{cases} 1 - (1-p)^{\lfloor x \rfloor} & , x \geq 1 \\ 0 & , x < 1 \end{cases}$$

$\lfloor x \rfloor$ = round x
down to
nearest
integer

This type of RV is
called a geometric RV.

$$\text{e.g. } \lfloor 3.5 \rfloor = 3$$

Defn: Discrete / Continuous RVs

A discrete RV is one whose CDF is a step fun.

A continuous RV is one whose CDF is continuous.

Defn: Probability Mass Function

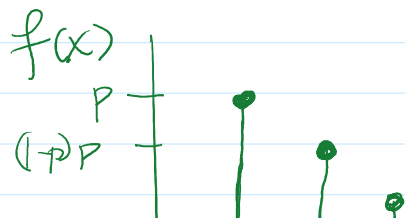
For a discrete RV X the prob. mass function (PMF) is the function $f: \mathbb{R} \rightarrow \mathbb{R}$ where

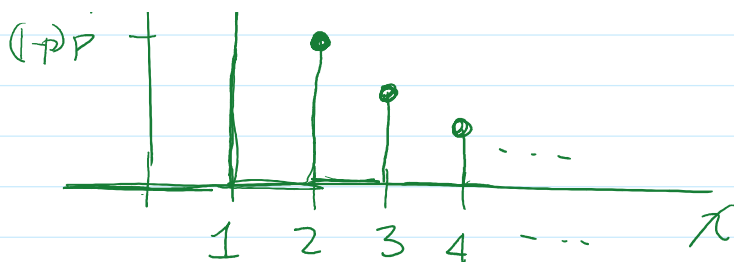
$$f(x) = P(X=x) \quad \forall x \in \mathbb{R}.$$

Sometimes called the distribution of X .

Ex. For our geometric RV

$$f(x) = P(X=x) = \begin{cases} (1-p)^{x-1} p, & x=1, 2, 3, 4, \dots \\ 0, & \text{else} \end{cases}$$





Theorem: For discrete RVs

$$F(x) = \sum_{i \leq x} f(i)$$

\uparrow $P(X \leq x)$ \uparrow $P(X = i)$

Pf. " $X \leq x$ " = $\bigcup_{i \leq x} "X = i"$

\uparrow disjoint union

hence

$$\begin{aligned}
 F(x) &= P(X \leq x) = P\left(\bigcup_{i \leq x} "X = i"\right) \\
 &= \sum_{i \leq x} P(X = i) \\
 &= \sum_{i \leq x} f(i) .
 \end{aligned}$$

Ex. Say X has a uniform distribution on $1, \dots, n$

denoted $X \sim \underline{U(\{1, \dots, n\})}$

\nwarrow read: distributed as

$$X \sim \underline{U(\{1, \dots, n\})}$$

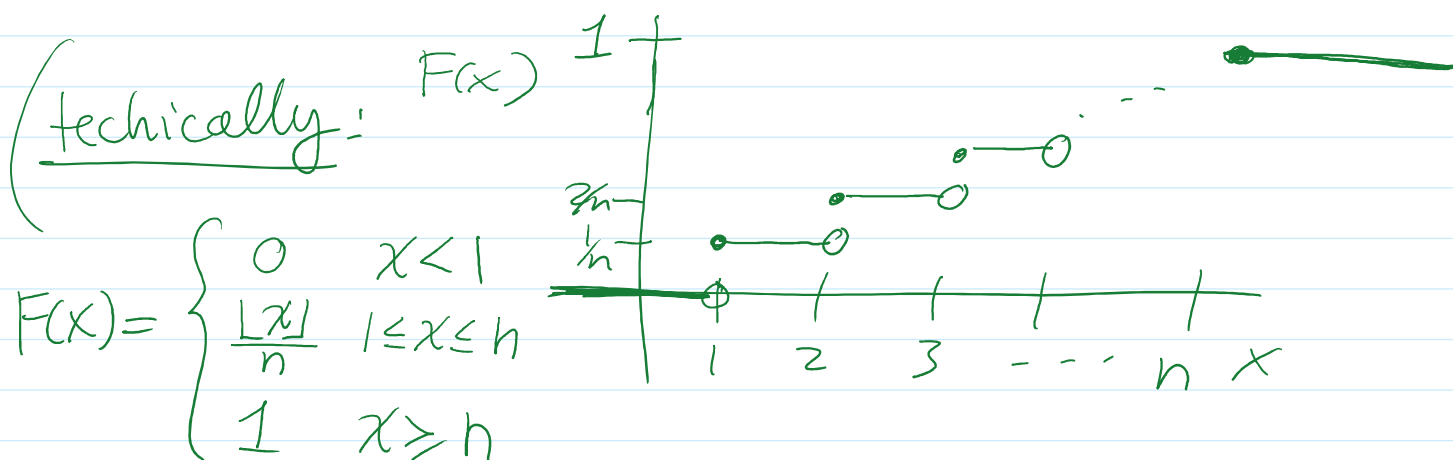
if

$$f(i) = \begin{cases} \frac{1}{n} & \text{for } i=1, \dots, n \\ 0 & \text{else} \end{cases}$$



Q: What is the CDF?

$$F(x) = \sum_{i \leq x} f(i) = \sum_{i=1}^x \frac{1}{n} = \frac{x}{n}$$



Said: $F(x) = \sum_{i \leq x} f(i)$
 \Downarrow
 $P(X \leq x)$

More general fact:

$$P(X \in A) = \sum_{i \in A} f(i) \quad (\textcircled{x})$$

Ex. X uniform over $1, \dots, 7$
($X \sim U(\{1, \dots, 7\})$)

$$\begin{aligned} P(2 \leq X \leq 5) \\ &= P(X \in \{2, 3, 4, 5\}) \\ &= \sum_{i=2,3,4,5} f(i) \\ &= \sum_{i=2}^5 1/7 = 4/7. \end{aligned}$$

Ex. Roll a die 60 times (independently)
 $X = \#$ of 6s I get
lets get the PMF of X .

$$f(x) = P(X=x) = \text{prds. I get } x \text{ 6s among 60 rolls.}$$

$$f(0) = P(X=0) = \underbrace{(5/6)(5/6)(5/6) \dots (5/6)}_{60 \text{ times}} = (5/6)^{60}$$

60 times

$$\begin{aligned} f(1) = P(X=1) &= \binom{60}{1} \left(\frac{1}{6}\right) \underbrace{\left(\frac{5}{6}\right) \left(\frac{5}{6}\right) \dots \left(\frac{5}{6}\right)}_{59 \text{ times}} \\ &= \binom{60}{1} \left(\frac{1}{6}\right) \left(\frac{5}{6}\right)^{59} \end{aligned}$$

$$\begin{aligned} f(2) = P(X=2) &= \binom{60}{2} \left(\frac{1}{6}\right) \left(\frac{1}{6}\right) \underbrace{\left(\frac{5}{6}\right) \dots \left(\frac{5}{6}\right)}_{58} \\ &= \binom{60}{2} \left(\frac{1}{6}\right)^2 \left(\frac{5}{6}\right)^{58} \end{aligned}$$

general pattern

$$f(x) = P(X=x) = \binom{60}{x} \left(\frac{1}{6}\right)^x \left(\frac{5}{6}\right)^{60-x}$$

We call this a Binomial RV.

Any experiment where I do n actions
independently each w/ a prob. p of
occurring ad

$X = \#$ occurrences

then

$$X \sim \text{Bin}(n, p).$$

