

# Discrete RVs

$$f(x) = P(X=x) \quad \text{PMF}$$

$$F(x) = P(X \leq x) \quad \text{CDF}$$

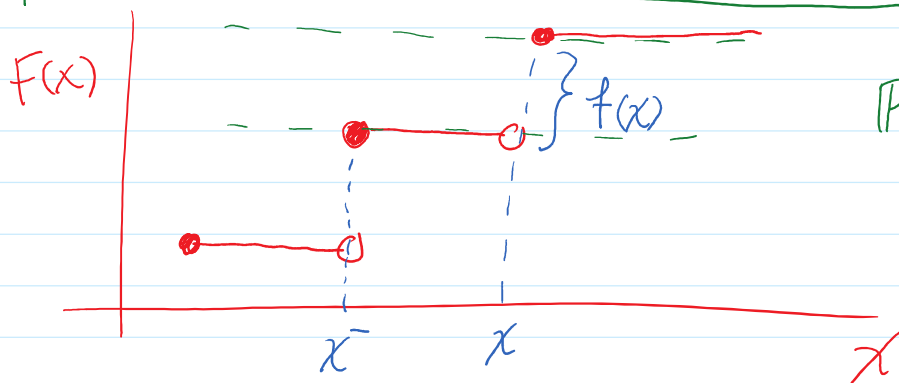
$$F(x) = \sum_{i \leq x} f(i)$$

from PMF to CDF

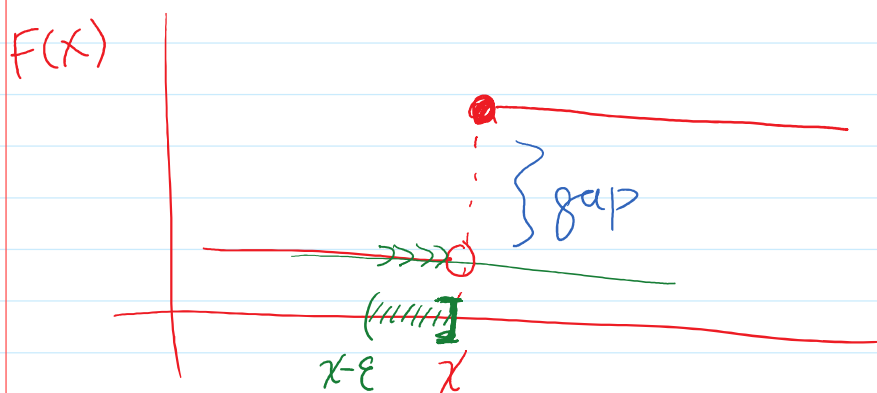
$$P(a < X \leq b) = F(b) - F(a)$$

← from CDF to PMF

$$P(X=x) \leftarrow$$



$$P(x^- < X \leq x) = F(x) - F(x^-)$$

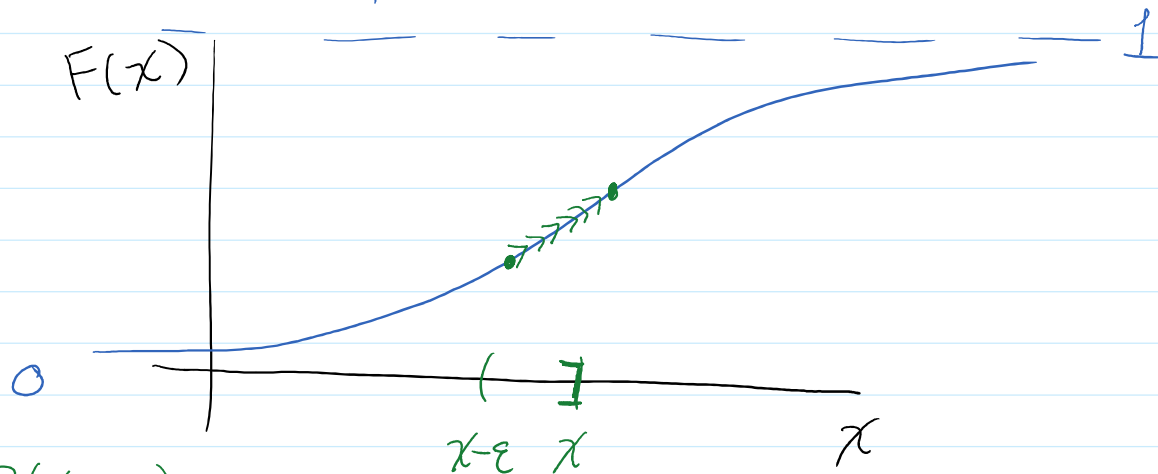


$$P(X=x) = \lim_{\epsilon \downarrow 0} P(x-\epsilon < X \leq x)$$

$$= \lim_{\epsilon \downarrow 0} F(x) - F(x-\epsilon)$$

= gap size at  $x$

What about for Continuous RVs?



$$P(X=x) =$$

$$\lim_{\varepsilon \downarrow 0} P(x-\varepsilon < X \leq x)$$

$$= \lim_{\varepsilon \downarrow 0} F(x) - F(x-\varepsilon)$$

$$= F(x) - \lim_{\varepsilon \downarrow 0} F(x-\varepsilon)$$

↖  $F$  is continuous

$$= F(x) - F(x)$$

$$= 0$$

punch line: - can't define a PMF as a jump size  
-  $P(X=x) = 0 \quad \forall x \in \mathbb{R}$

Can we do something similar for cts RVs?

Want something like:  $F(x) = \sum f(i)$

Want something like:  $F(x) = \sum_{i \leq x} f(i)$

Defn: Probability Density Function (PDF)

Analogy of PMF for cts RVs.

The PDF is a function  $f: \mathbb{R} \rightarrow \mathbb{R}$  defined so that  $\forall x \in \mathbb{R}$

$$F(x) = \int_{-\infty}^x f(t) dt$$

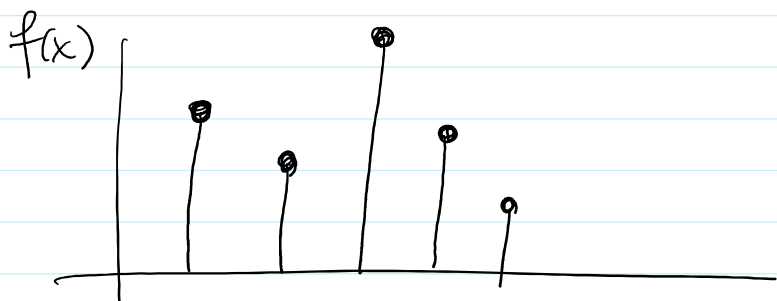
note! by Fundamental Theorem of Calculus

$$\frac{dF}{dx} = \frac{d}{dx} \int_{-\infty}^x f(t) dt = f(x).$$

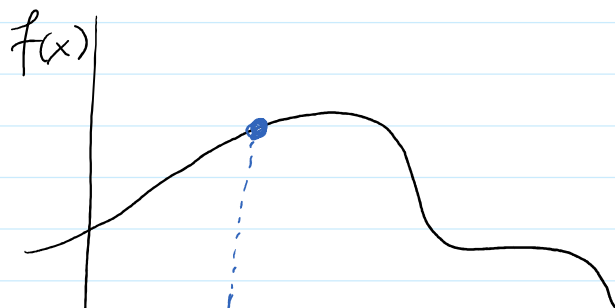
punchline!  $f(x) = \frac{dF}{dx}$

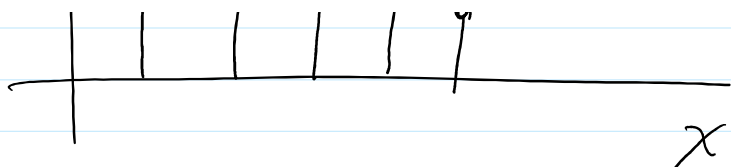
PDF = derivative of CDF.

discrete case (PMF)

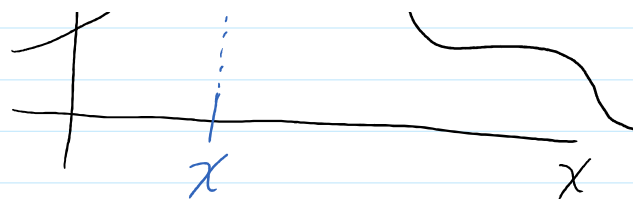


Continuous (PDF)





$$f(x) = P(X=x)$$

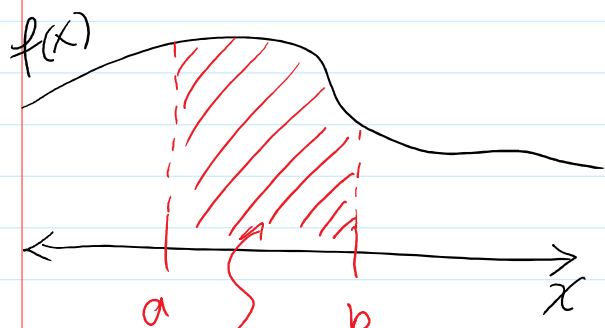


$$f(x) \neq P(X=x) = 0$$

## Properties for cts RVs

$$P(a < X \leq b) = F(b) - F(a)$$

$$= \int_{-\infty}^b f(t) dt - \int_{-\infty}^a f(t) dt$$



$$= \int_a^b f(t) dt$$

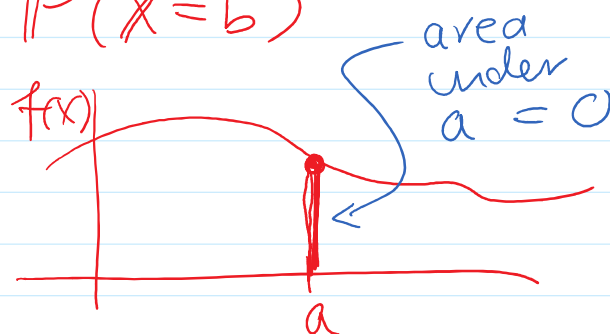
$P(a < X \leq b) = \text{area btwn } a \text{ and } b \text{ under PDF}$

We said:  $P(X=a) = 0 = P(X=b)$

$$P(a < X \leq b)$$

$$= P(a \leq X \leq b)$$

$$= P(a < X < b)$$



$$= P(a \leq X < b)$$

More generally:

For discrete:  $P(X \in A) = \sum_{x \in A} f(x)$

For cts:  $P(X \in A) = \int_A f(x) dx$

Ex.  $P(X \in (2, 3]) = \int_2^3 f(t) dt$

$$P(X \in \{2\}) = P(X=2) = \int_2^2 f(t) dt = 0$$

Ex.  $F(x) = \frac{1}{1+e^{-x}}$

what is the associated PDF?

$$f(x) = \frac{dF}{dx} = \dots = \frac{e^{-x}}{(1+e^{-x})^2}$$

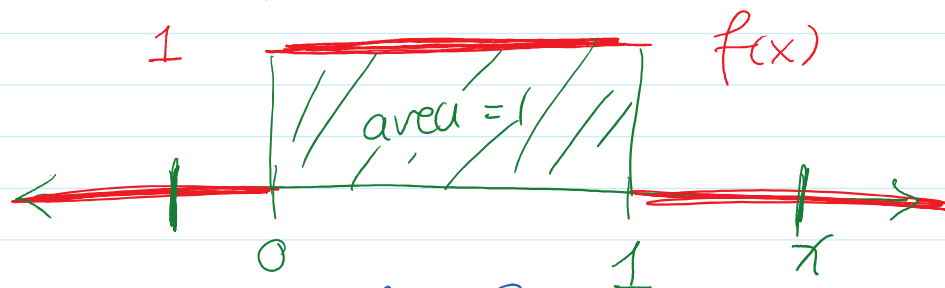
Ex. Continuous Uniform Distribution (on  $[0, 1]$ )

$$X \sim U(0, 1)$$

means

means

$$f(x) = \begin{cases} 1 & \text{if } 0 \leq x \leq 1 \\ 0 & \text{else} \end{cases}$$



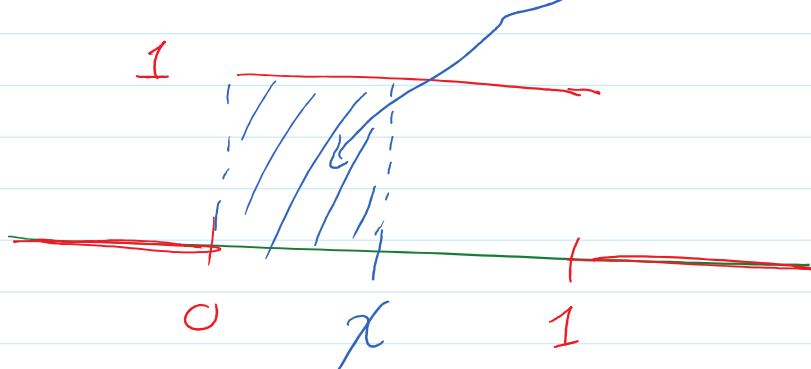
What is the CDF of  $X$ ?

$$F(x) = \int_{-\infty}^x f(t) dt$$

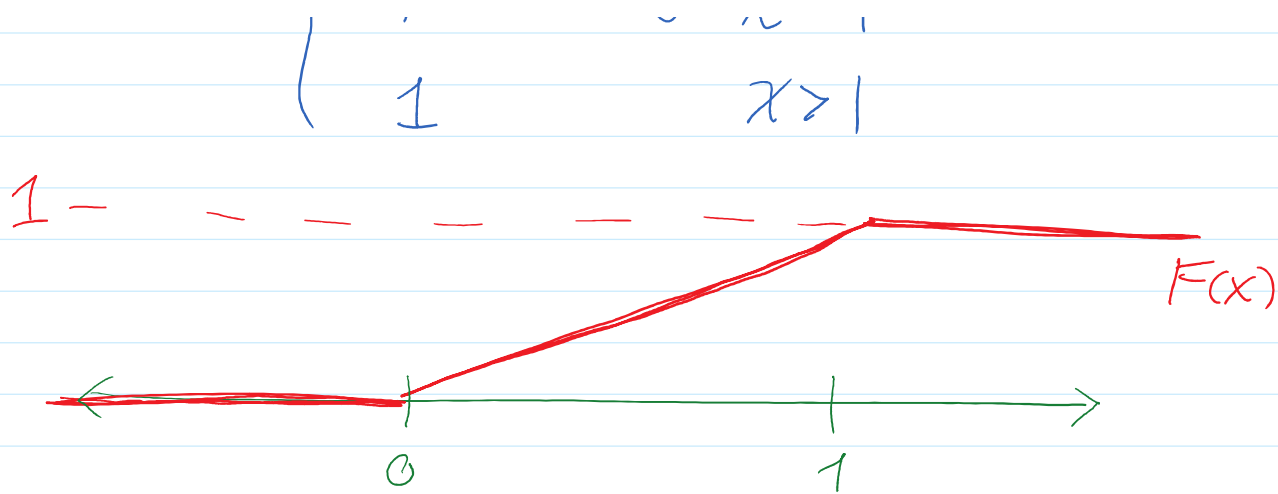
$$\text{If } x < 0 \text{ then } F(x) = \int_{-\infty}^x 0 dt = 0$$

$$\text{If } x > 1 \text{ then } F(x) = \int_{-\infty}^x f(t) dt = 1$$

$$\text{If } 0 < x < 1 \text{ then } F(x) = \text{area} = x$$



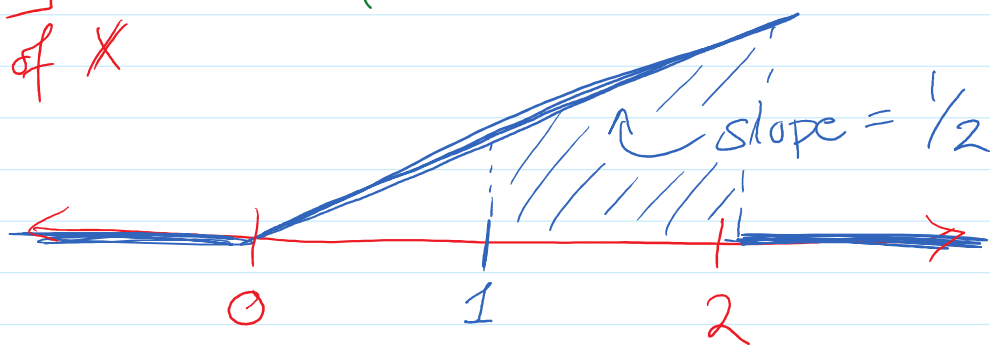
$$F(x) = \begin{cases} 0 & x < 0 \\ x & 0 < x < 1 \\ 1 & x > 1 \end{cases}$$



Ex. Let

$$f(x) = \begin{cases} x/2 & 0 < x < 2 \\ 0 & \text{else} \end{cases}$$

PDF of X

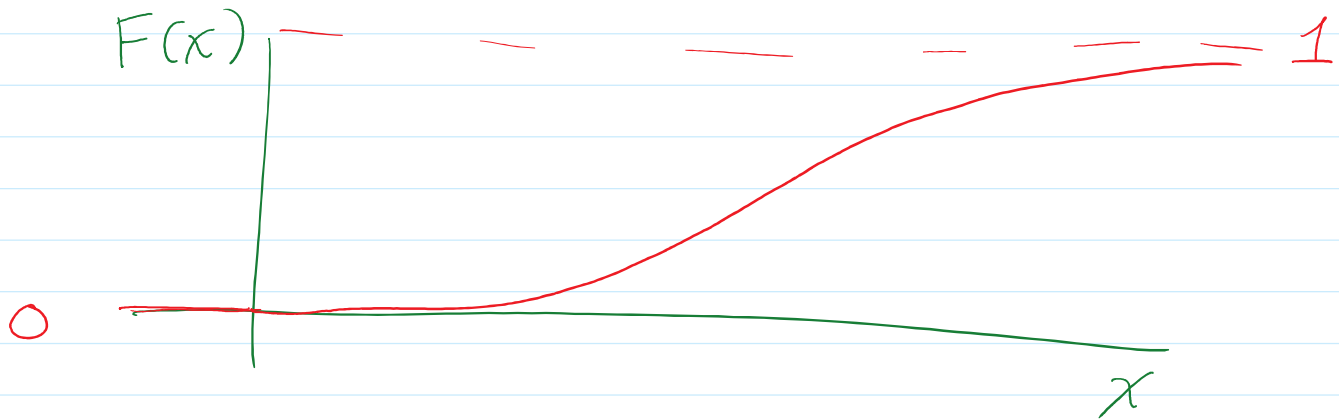


$$P(X > 1) = \int_1^{\infty} f(t) dt = \int_1^2 \frac{t}{2} dt = \left. \frac{t^2}{4} \right|_1^2$$

$$= \frac{1}{4} (2^2 - 1^2) = \frac{3}{4}$$

Ex. Let

$$F(x) = \begin{cases} 1 - e^{-x} & \text{for } x \geq 0 \\ 0 & \text{else} \end{cases}$$



Q:  $P(1 < X < 2)$

Way 1

Theorem:  $P(a < X \leq b) = F(b) - F(a)$

$$\begin{aligned} P(1 < X < 2) &= F(2) - F(1) \\ &= (1 - e^{-2}) - (1 - e^{-1}) \\ &= e^{-1} - e^{-2} \end{aligned}$$

Way 2:

$$f(x) = \frac{dF}{dx} = \frac{d}{dx}(1 - e^{-x}) = e^{-x}$$

and so

$$P(1 < X < 2) = \int_1^2 f(x) dx = \int_1^2 e^{-x} dx$$



$$= -e^{-x} \Big|_1^2 = -e^{-2} - (-e^{-1}) \\ = e^{-1} - e^{-2}$$

## Theorem: Valid PMF/PDFs

There is some RV w/  $f$  as its PMF/PDF iff

- ①  $f(x) \geq 0 \quad \forall x \in \mathbb{R}$
- ②  $\begin{cases} \text{(discrete)} & \sum_{x \in \mathbb{R}} f(x) = 1 \\ \text{(continuous)} & \int_{\mathbb{R}} f(x) dx = 1 \end{cases}$

note: For ① (discrete)  $P(X \in A) = \sum_{x \in A} f(x) \geq 0$   
 (cts)  $P(X \in A) = \int_A f(x) dx \geq 0$

$$\begin{aligned} \text{② } 1 &= P(S) = P(X \in \mathbb{R}) \\ &= \begin{cases} \sum_{\mathbb{R}} f(x) = 1 \\ \int_{\mathbb{R}} f(x) dx = 1 \end{cases} \end{aligned}$$

Fact: if I have a fn  $g(x) \geq 0$

in cts case:  $\int_{\mathbb{R}} g(x) dx = c < \infty$

then

$$f(x) = \frac{1}{c} g(x)$$

we have  $f(x) \geq 0$  and

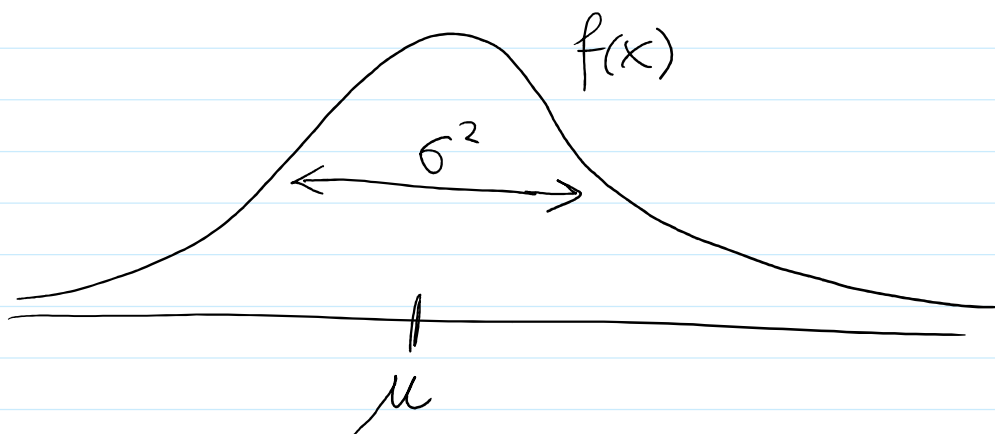
$$\int_{\mathbb{R}} f(x) dx = \frac{1}{c} \int_{\mathbb{R}} g(x) dx = \frac{1}{c} c = 1.$$

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Ex. Normal Distribution (Gaussian Distribution)

$$X \sim N(\mu, \sigma^2)$$

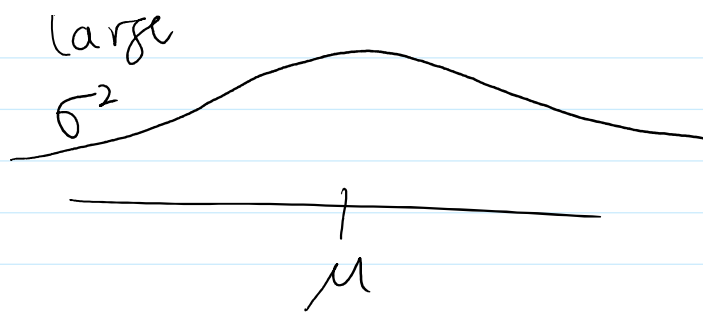
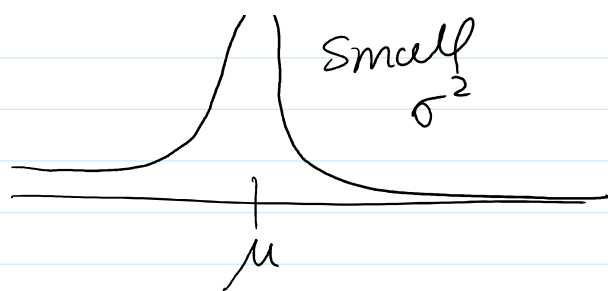
mean:  $\mu \in \mathbb{R}$       Variance:  $\sigma^2 > 0$



 small  $\sigma^2$

large  $\sigma^2$





Special case: standard normal  
when  $\mu = 0$  and  $\sigma^2 = 1$

$$X \sim N(0, 1).$$

the PDF of  $X$  is

$$f(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}x^2\right) \quad \forall x \in \mathbb{R}$$

$$\exp(a) = e^a$$

Is this a valid PDF?

$$(1) f(x) \geq 0 \quad \checkmark$$

$$(2) \int f(x) dx = 1$$



$$\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}x^2\right) dx \stackrel{?}{=} 1$$

a number  $I$

Want to show:  $I = 1$

$$\Leftrightarrow I^2 = 1$$

$\infty$

$\infty$

$$I^2 = I \cdot I = \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}x^2) dx}_I \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}y^2) dy}_I$$

$$\stackrel{\text{Calc III}}{=} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \frac{1}{2\pi} \exp(-\frac{1}{2}x^2) \exp(-\frac{1}{2}y^2) dx dy$$

$$= \iint_{\mathbb{R}^2} \frac{1}{2\pi} \exp(-\frac{1}{2}(x^2+y^2)) dx dy$$

Polar coordinates?

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