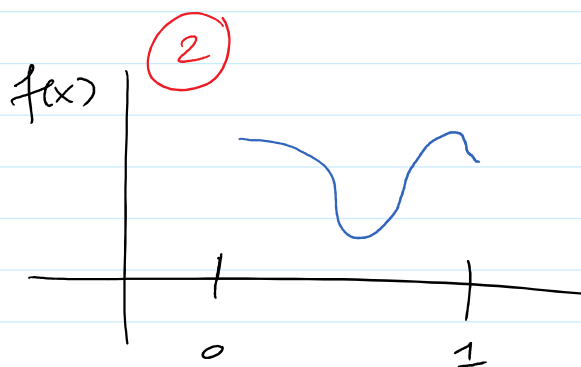
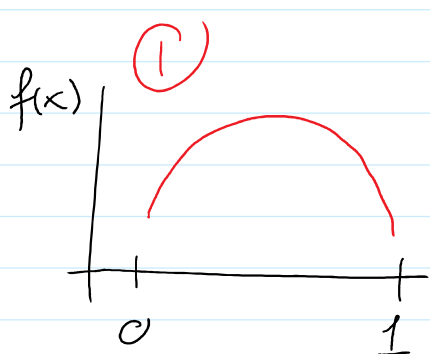


Beta Distribution

— a cts RV w/ support between 0 and 1



Beta Function $a, b \in \mathbb{R}^+$

$$B(a, b) = \int_0^1 x^{a-1} (1-x)^{b-1} dx$$

Fact:

$$B(a, b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

$$\frac{1}{B(a, b)} = \frac{\Gamma(a+b)}{\Gamma(a)\Gamma(b)} \approx \frac{(a+b)!}{a! b!} = \binom{a+b}{a}$$

Beta Dist:

$$X \sim \text{Beta}(a, b)$$

$$f(x) = \frac{x^{a-1} (1-x)^{b-1}}{B(a, b)} \quad \text{for } 0 < x < 1$$

$$B(a, b)$$

Expectation:

$$E[X] = \int_0^1 x \frac{x^{a-1} (1-x)^{b-1}}{B(a, b)} dx$$

$$= \int_0^1 \frac{x^{(a+1)-1} (1-x)^{b-1}}{B(a, b)} dx$$

looks like

$$\propto \int \text{PDF Beta}(a+1, b)$$

goes to 1

$$= \frac{B(a+1, b)}{B(a, b)} \underbrace{\int_0^1 \frac{x^{(a+1)-1} (1-x)^{b-1}}{B(a+1, b)} dx}_1$$

$$\frac{x^{(a+1)-1} (1-x)^{b-1}}{B(a+1, b)}$$

$$= \frac{B(a+1, b)}{B(a, b)} = \frac{\cancel{\Gamma(a+1)} \cancel{\Gamma(b)}}{\Gamma(a+b+1)} \bigg/ \frac{\cancel{\Gamma(a)} \cancel{\Gamma(b)}}{\Gamma(a+b)}$$

$$= \frac{\Gamma(a+b) \Gamma(a+1)}{\Gamma(a+b+1) \Gamma(a)}$$

$$= \frac{\cancel{\Gamma(a+b)} a \cancel{\Gamma(a)}}{(a+b) \cancel{\Gamma(a+b)} \cancel{\Gamma(a)}}$$

$$\Gamma(a+1) = a \Gamma(a)$$

$$(a+b) / (a+b) / (a)$$

$$= \boxed{\frac{a}{a+b} = E[X]}$$

Moments:

$$E[X^r] = \int_0^1 x^r \frac{x^{a-1} (1-x)^{b-1}}{B(a,b)} dx$$

$$= \frac{B(a+r,b)}{B(a,b)} \int_0^1 \frac{x^{(a+r)-1} (1-x)^{b-1}}{B(a+r,b)} dx$$

integrates to 1

$$\boxed{E[X^r] = \frac{B(a+r,b)}{B(a,b)}}$$

$$E[X^2] = \frac{B(a+2,b)}{B(a,b)} = \frac{\Gamma(a+2)\Gamma(b)}{\Gamma(a+b+2)} \bigg/ \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

$$= \frac{(a+1)a \cancel{\Gamma(a)} \cancel{\Gamma(b)} \cancel{\Gamma(a+b)}}{(a+b+1)(a+b) \cancel{\Gamma(a+b)} \cancel{\Gamma(a)} \cancel{\Gamma(b)}}$$

$$= \frac{a(a+1)}{(a+b)(a+b+1)}$$

$$\text{Var}(X) = E[X^2] - E[X]^2$$

$$= \frac{a(a+1)}{(a+b)(a+b+1)} - \left(\frac{a}{a+b} \right)^2$$

$$= \dots = \frac{ab}{(a+b+1)(a+b)^2} = \text{Var}(X)$$

Transformations:

(1) If I know something about X

(2) What do I know about $Y = g(X)$

Discrete RVs

Let X be a discrete RV

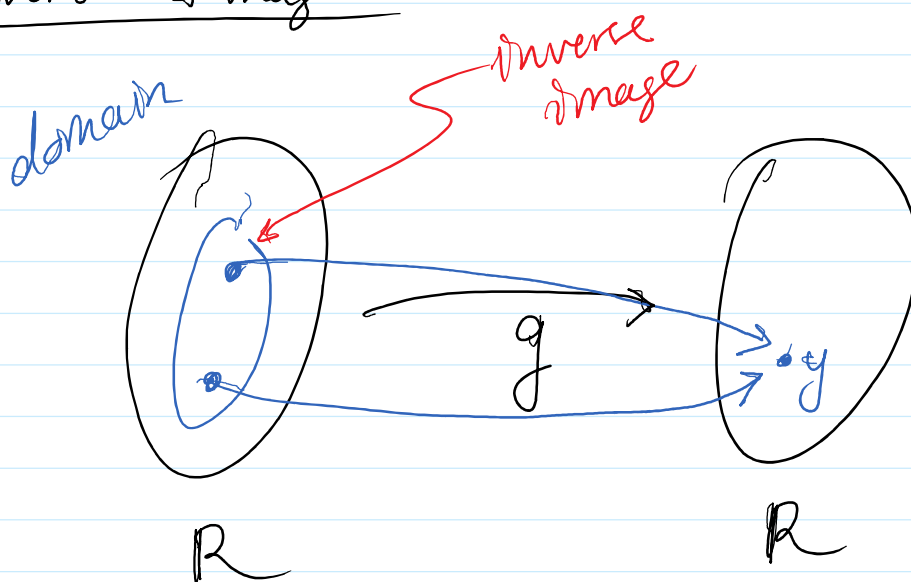
knows f_X

PMF of X

and let $Y = g(X)$

Q: what is f_Y ?

Inverse Image



$$g^{-1}(\{y\}) = \{x : g(x) = y\}$$

if g is really invertible then inverse image = true inverse

$$f_Y(y) = P(Y=y) = P(g(X)=y) \quad (*)$$

want

if g is invertible then

$$\begin{aligned} \textcircled{*} &= P(X = g^{-1}(y)) \\ &= f_X(g^{-1}(y)) \end{aligned}$$

If g isn't invertible then

$$\textcircled{*} = P(X \in g^{-1}(y))$$

$$= \sum_{x \in g^{-1}(y)} f_X(x)$$

$$= \sum_{x: g(x)=y} f_X(x)$$

$$\begin{aligned} P(X \in A) \\ &= \sum_{x \in A} f_X(x) \end{aligned}$$

* Theorem: If X is discrete and $Y = g(X)$ then

$$f_Y(y) = \sum_{x: g(x)=y} f_X(x)$$

Ex.

Let $X \sim \text{Bin}(n, p)$

↑ # of H in n coin flips each w/ a prob. p of getting H

$Y = n - X$ ← # tails

$$Y = g(X) \quad \text{where } y = g(x) = n - x$$

$$\Updownarrow x = n - y$$

$$f_Y(y) = \sum_{x: g(x)=y} f_X(x) = \sum_{x=n-y} f_X(x)$$

$$= f_X(n-y)$$

$$f_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$= \binom{n}{n-y} p^{n-y} (1-p)^{n-(n-y)}$$

note: $\binom{n}{n-y} = \binom{n}{y}$; $q = 1-p$

$$= \binom{n}{y} q^y (1-q)^{n-y}$$

PMF of $\text{Bin}(n, q)$

i.e. $Y \sim \text{Bin}(n, 1-p)$

Let's consider continuous RVs

Theorem: If X is continuous and $Y = g(X)$

then

✓ ① if g is increasing then

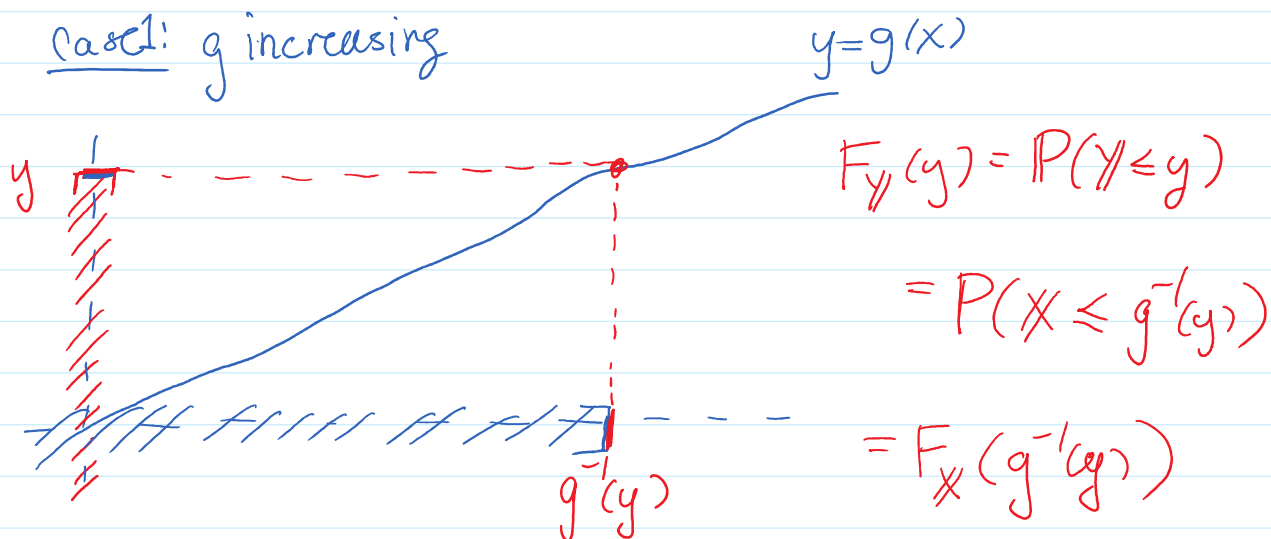
$$F_Y(y) = F_X(g^{-1}(y))$$

inverse exists

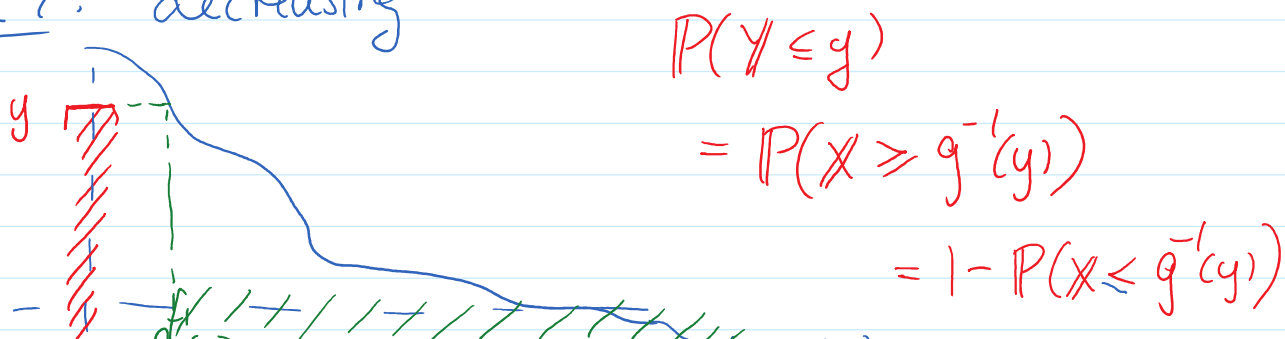
② if g is decreasing then

$$✓ F_Y(y) = 1 - F_X(g^{-1}(y))$$

Pf. case 1: g increasing

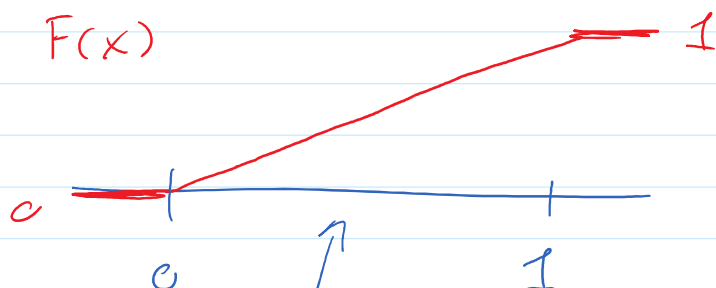
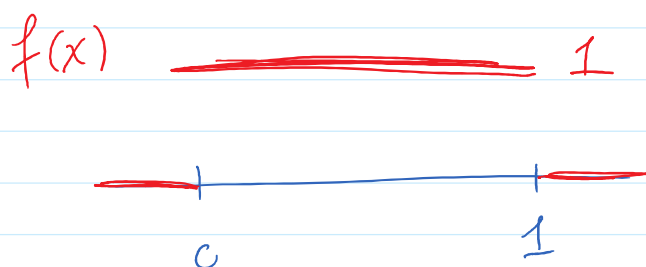


Case 2: decreasing



$$\begin{aligned}
 &= 1 - P(X \leq g^{-1}(y)) \\
 &= 1 - F_X(g^{-1}(y))
 \end{aligned}$$

Ex. $X \sim U(0, 1)$



$$\text{let } Y = -\log X$$

$$g(x) = -\log x \quad (\text{decreasing})$$

$$y = -\log x = g(x)$$

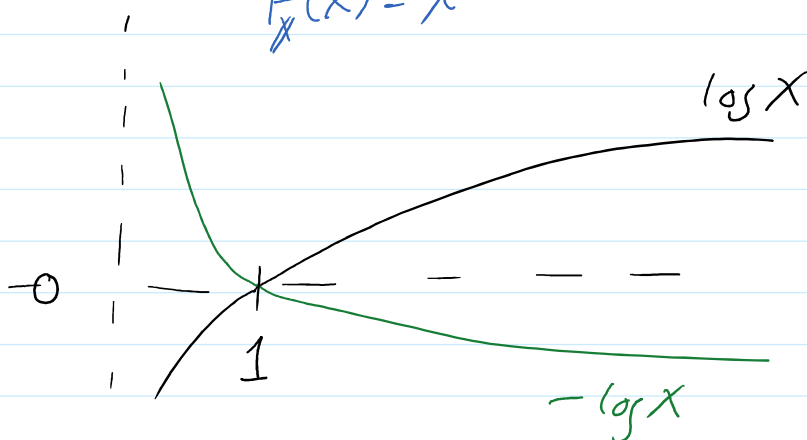
$$\Rightarrow -y = \log x$$

$$\Rightarrow e^{-y} = x = g^{-1}(y)$$

$$F_Y(y) = 1 - F_X(g^{-1}(y))$$

$$= 1 - F_X(e^{-y})$$

$$= 1 - e^{-y}$$



$$0 < x < 1 \Rightarrow -\log x > 0$$

$$\text{if } y > 0 \text{ then } e^{-y} = \frac{1}{e^y} < 1$$

$$\text{so } 0 < e^{-y} < 1$$

$$n \quad \Gamma \dots (1)$$

↑ CDF of a $\text{Exp}(1)$

So $Y \sim \text{Exp}(1)$

$Z \sim \text{Exp}(1)$

$$F(z) = \int_0^z e^{-x} dx = 1 - e^{-z}$$

Theorem: If X is continuous and $Y = g(X)$

and

- ① g is invertible
- ② g^{-1} is differentiable

then

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{dg^{-1}}{dy} \right|$$