

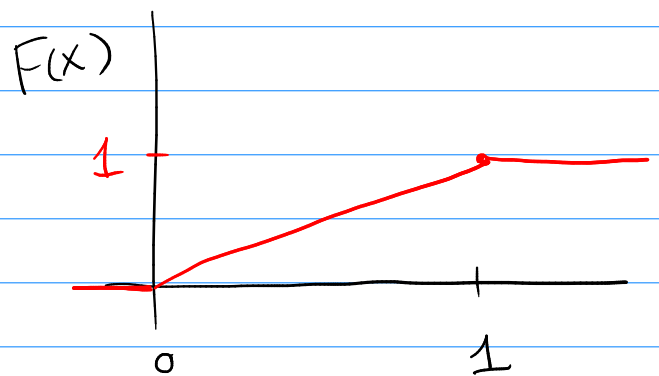
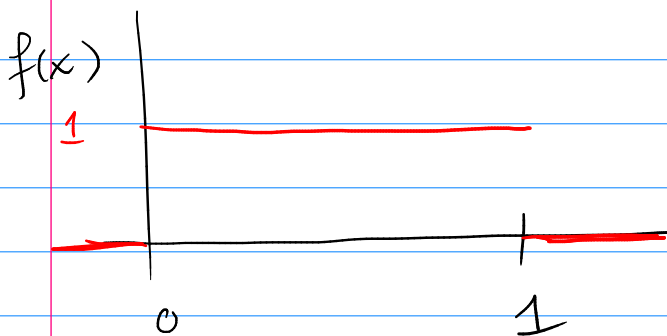
# Lecture 15

OHS: Today 2:30 - 3:30

Weds 1-2 pm

Mon 10-11 am, 3-4 pm

Ex.  $X \sim U(0,1)$



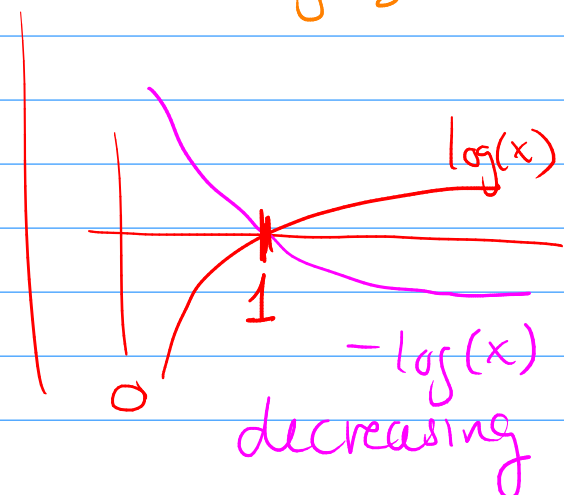
$$\text{let } Y = -\log X$$

$$g(x) = -\log(x)$$

$$y = -\log(x) \Rightarrow -y = \log(x) \Rightarrow e^{-y} = x = g^{-1}(y)$$

Apply theorem:

$$\begin{aligned} F_Y(y) &= 1 - F_X(g^{-1}(y)) \\ &= 1 - F_X(e^{-y}) \end{aligned}$$



If  $0 < x < 1$  then  $\log(x) < 0$

$$y = -\log(x) > 0$$

$$\text{so } 0 < e^{-y} = \frac{1}{e^y} \leq 1$$

$$\rightarrow = 1 - e^{-y} = F_Y(y)$$

CDF of  $\text{Exp}(1)$

So  $Y \sim \text{Exp}(1)$

$$Z \sim \text{Exp}(1)$$

then

$$f(z) = e^{-z} \text{ for } z > 0$$

$$F(z) = \int_0^z e^{-t} dt = 1 - e^{-z}$$

What about PDFs?

Theorem: If  $X$  is cts and  $Y = g(X)$

and if

(1)  $g$  is invertible

(2)  $g^{-1}$  is differentiable

then

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{dg^{-1}}{dy} \right|$$

pf Case 1:  $g$  increasing

Prev. CDF theorem said

$$F_Y(y) = F_X(g^{-1}(y))$$

-  $g$  inc  
-  $g^{-1}$  inc  
-  $\frac{dg^{-1}}{dy} > 0$

$$f_Y(y) = \frac{dF_Y}{dy} = f_X(g^{-1}(y)) \left| \frac{dg^{-1}}{dy} \right|$$

Case 2:  $g$  is decreasing  $\rightarrow g^{-1}$  is dec,  $\frac{dg^{-1}}{dy} < 0$

Prev. thrm said  $F_Y(y) = 1 - F_X(g^{-1}(y))$

$$\begin{aligned} f_Y(y) &= \frac{dF_Y}{dy} = -f_X(g^{-1}(y)) \frac{dg^{-1}}{dy} \\ &= f_X(g^{-1}(y)) \left( -\frac{dg^{-1}}{dy} \right) \\ &= f_X(g^{-1}(y)) \left| \frac{dg^{-1}}{dy} \right| \end{aligned}$$

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Ex. let  $X \sim \text{Gamma}(k, \lambda)$

$$f_X(x) = \frac{\lambda e^{-\lambda x} (x\lambda)^{k-1}}{\Gamma(k)}, \quad x > 0$$

let  $Y = 1/X$

$$g(x) = 1/x$$

$$y = 1/x \Rightarrow x = 1/y = g^{-1}(y)$$

$$\text{so } \frac{dg^{-1}}{dy} = -\frac{1}{y^2}$$

$$f_Y(y) = f_X(g^{-1}(y)) \left| \frac{dg^{-1}}{dy} \right|$$

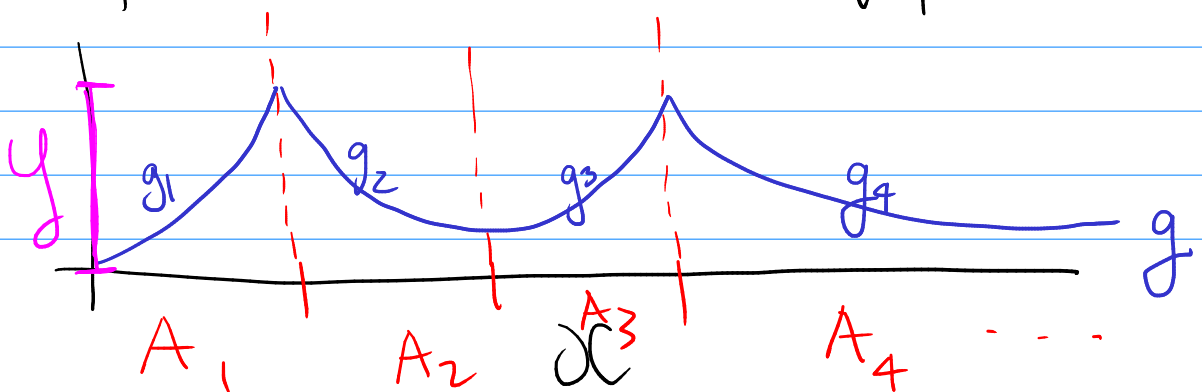
$$= f_X\left(\frac{1}{y}\right) \left| -\frac{1}{y^2} \right|$$

$$= \frac{\lambda e^{-\lambda \frac{1}{y}} \left(\lambda \frac{1}{y}\right)^{k-1}}{\Gamma(k)} \cdot \frac{1}{y^2} \quad \text{for } y > 0$$

↑ called Inverse Gamma dist

What about non-invertible  $g$ ?

Theorem: Let  $X$  be a cts RV w/ support  $\mathcal{X}$  and for  $i=1, \dots, K$  let  $A_i$  partition  $\mathcal{X}$ .



Let  $g_i$  to be  $g$  restricted to  $A_i$

If ① my prev. theorem applies to each  $g_i$  restr. to  $A_i$

[ $g_i$  invertible on  $A_i$   
 $g_i^{-1}$  is diffable on  $A_i$ ]

② The image of  $A_i$  under  $g_i$  is the same for all  $i$

[all  $g_i$  have the range  $Y$ ]

then

$$f_Y(y) = \sum_{i=1}^K f_X(g_i^{-1}(y)) \left| \frac{dg_i^{-1}}{dy} \right|$$

for  $y \in Y$

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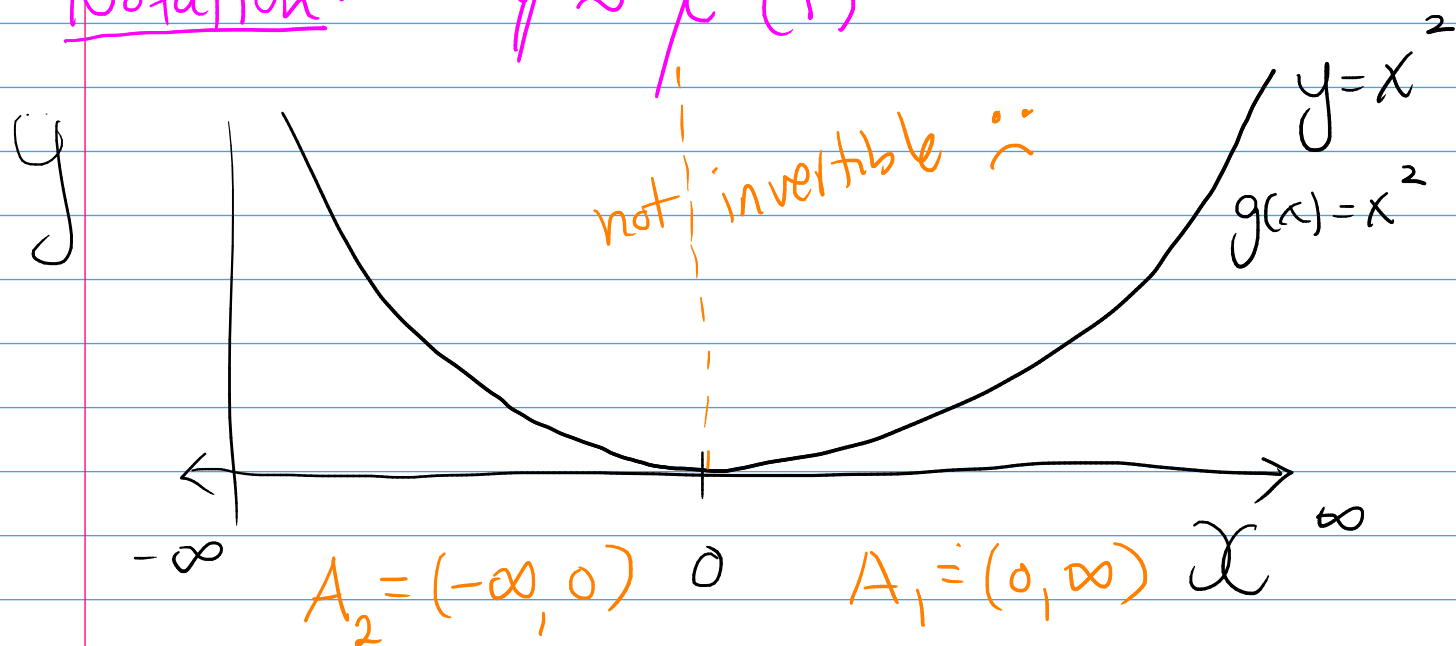
Ex. Chi-Squared Dist

If  $X \sim N(0, 1)$  and  $Y = X^2$

then we say  $Y$  has a Chi-Sq dist

w/ one degree of freedom.

Notation:  $Y \sim \chi^2(1)$



$$A_1 = (0, \infty), \quad g_1(x) = x^2, \quad g_1^{-1}(y) = \sqrt{y};$$
$$\frac{dg_1^{-1}(y)}{dy} = \frac{1}{2\sqrt{y}}$$

$$A_2 = (-\infty, 0), \quad g_2(x) = x^2, \quad g_2^{-1}(y) = -\sqrt{y}$$
$$\frac{dg_2^{-1}}{dy} = \frac{-1}{2\sqrt{y}}$$

$$f_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}x^2\right) \text{ for } x \in \mathbb{R}$$

Apply theorem:

$$\begin{aligned}
 f_Y(y) &= f_X(g_1^{-1}(y)) \left| \frac{dg_1^{-1}}{dy} \right| + f_X(g_2^{-1}(y)) \left| \frac{dg_2^{-1}}{dy} \right| \\
 &= f_X(\sqrt{y}) \left| \frac{1}{2\sqrt{y}} \right| + f_X(-\sqrt{y}) \left| \frac{-1}{2\sqrt{y}} \right| \\
 &= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}(\sqrt{y})^2\right) \frac{1}{2\sqrt{y}} + \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}(-\sqrt{y})^2\right) \cdot \frac{1}{2\sqrt{y}}
 \end{aligned}$$

$$= \frac{1}{\sqrt{2\pi}} \frac{1}{2\sqrt{y}} \left[ e^{-\frac{1}{2}y} + e^{-\frac{1}{2}y} \right]$$

$$f_Y(y) = \frac{1}{\sqrt{2\pi}} \frac{1}{\sqrt{y}} e^{-\frac{1}{2}y} \quad \text{for } y > 0$$

↑ density of  $\chi^2(1)$  [Type of Gamma]

Thrm Probability Integral Transf

If  $X$  is a ctr RV w/ CDF  $F_X$   
then

$$F_X(X) \sim U(0,1).$$

pf. Assume that  $F_X$  is strictly increasing  
so that it is invertible so  $F_X^{-1}$  exists.

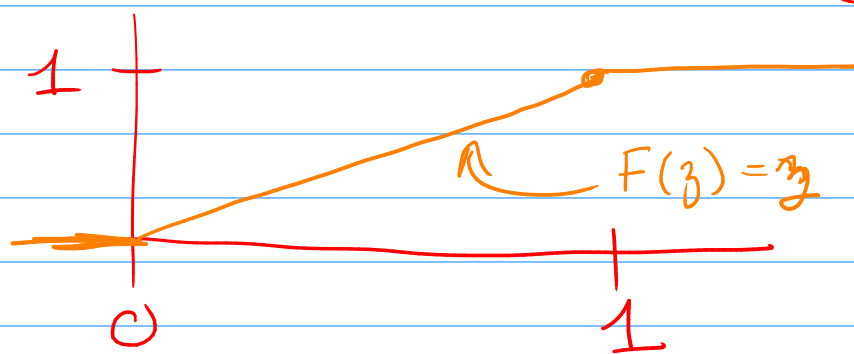
Our CDF thm says that if  
 $Y = g(X)$ ,  $g$  inc, then  $F_Y(y) = F_X(g^{-1}(y))$

Let  $Y = F_X(X)$  i.e.  $g = F_X$   
so  $g^{-1} = F_X^{-1}$

hence

$$F_Y(y) = F_X(F_X^{-1}(y)) = y$$

↑ CDF of  $U(0,1)$



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Know: how to generate a  $U(0,1)$

Want: generate some RV w/ CDF  $F_X$

Let  $Z = F_X^{-1}(U)$  when  $U \sim U(0,1)$

then  $Z \sim F_X$



this happens b/c

$$F_Z(z) = P(Z \leq z)$$

$$= P(F_X^{-1}(u) \leq z)$$

$$= P(u \leq F_X(z))$$

$$= F_u(F_X(z))$$

$$= F_X(z)$$

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Algorithm:

① generate  $u \sim U(0,1)$

②  $z = F_X^{-1}(u)$

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Ex. Want  $X \sim \text{Exp}(1)$

CDF of  $\text{Exp}(1)$  is  $F_X(x) = 1 - e^{-x}$

$$y = 1 - e^{-x} \\ \Rightarrow x = -\log(1-y) = F_X^{-1}(y)$$