

Lecture 14

$$X \sim \text{Geom}(p)$$

$$M(t) = \mathbb{E}[e^{tX}] = \sum_{x=1}^{\infty} e^{tx} (1-p)^{x-1} p$$

$$= p \sum_{x=0}^{\infty} e^{t(x+1)} (1-p)^x$$

$$= pe^t \sum_{x=0}^{\infty} \underbrace{((1-p)e^t)^x}_r$$

$$= pe^t \frac{1}{1 - (1-p)e^t}$$

$$\sum_{i=0}^{\infty} r^i = \frac{1}{1-r}$$

$$|r| < 1$$

$$M(t) = \frac{pe^t}{1 - (1-p)e^t}$$

$$\text{for } (1-p)e^t < 1$$

$$\begin{array}{c} \updownarrow \\ t < -\log(1-p) \end{array}$$

$$\mathbb{E}[X^2] = \left. \frac{d^2 M}{dt^2} \right|_{t=0} = \dots = \frac{2-p}{p^2}$$

$$\text{Var}(X) = \mathbb{E}[X^2] - \mathbb{E}[X]^2$$

$$= \frac{2-p}{p^2} - \left(\frac{1}{p}\right)^2 = \dots = \frac{1-p}{p^2}$$

Beta Distribution

- cts RV w/ support of $[0, 1]$



Beta Function: $a, b > 0$

$$\begin{aligned} B(a, b) &= \int_0^1 x^{a-1} (1-x)^{b-1} dx \\ &= \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)} \end{aligned}$$

Beta dist: $X \sim \text{Beta}(a, b)$

$$f(x) = \frac{x^{a-1} (1-x)^{b-1}}{B(a, b)} \quad \text{for } 0 < x < 1.$$

Expectation

$$EX = \int_0^1 x \frac{x^{a-1} (1-x)^{b-1}}{B(a, b)} dx$$

↑

looks like PDF of
 $\text{Beta}(a+1, b)$

$$\frac{x^a (1-x)^{b-1}}{B(a+1, b)}$$

$$= \frac{B(a+1, b)}{B(a, b)} \int_0^1 \frac{x^a (1-x)^{b-1}}{B(a+1, b)} dx$$

└────────────────────────────────┘
integrate to 1

$$= \frac{B(a+1, b)}{B(a, b)}$$

$$= \frac{\Gamma(a+1)\Gamma(b)}{\Gamma(a+b+1)} \bigg/ \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

$$= \frac{a \cancel{\Gamma(a)} \cancel{\Gamma(b)}}{(a+b) \cancel{\Gamma(a+b)}} \bigg/ \frac{\cancel{\Gamma(a)} \cancel{\Gamma(b)}}{\cancel{\Gamma(a+b)}}$$

$$E[X] = \frac{a}{a+b}$$

$$E[X^r] = \int_0^1 x^r \frac{x^{a-1} (1-x)^{b-1}}{B(a,b)} dx \quad \left| \quad \frac{x^{a+r-1} (1-x)^{b-1}}{B(a+r,b)} \right.$$

\nwarrow Beta(a+r, b)

$$= \frac{B(a+r,b)}{B(a,b)} \underbrace{\int_0^1 \frac{x^{a+r-1} (1-x)^{b-1}}{B(a+r,b)} dx}_1$$

$$E[X^r] = \frac{B(a+r,b)}{B(a,b)}$$

$$E[X^2] = \frac{B(a+2,b)}{B(a,b)}$$

$$= \frac{\Gamma(a+2)\Gamma(b)}{\Gamma(a+b+2)} \bigg/ \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}$$

$$= \frac{(a+1)a \cancel{\Gamma(a)} \cancel{\Gamma(b)}}{(a+b+1)(a+b) \cancel{\Gamma(a+b)}} \bigg/ \frac{\cancel{\Gamma(a)} \cancel{\Gamma(b)}}{\cancel{\Gamma(a+b)}}$$

$$E[X^2] = \frac{a(a+1)}{(a+b+1)(a+b)}$$

$$\begin{aligned}
 \text{Var}(X) &= E[X^2] - E[X]^2 \\
 &= \frac{a(a+1)}{(a+b+1)(a+b)} - \left(\frac{a}{a+b}\right)^2 \\
 &= \dots \\
 &= \frac{ab}{(a+b)^2(a+b+1)}
 \end{aligned}$$


EXAM 2


Transformations

If I know something about X ,
 what do I know about $Y = g(X)$?

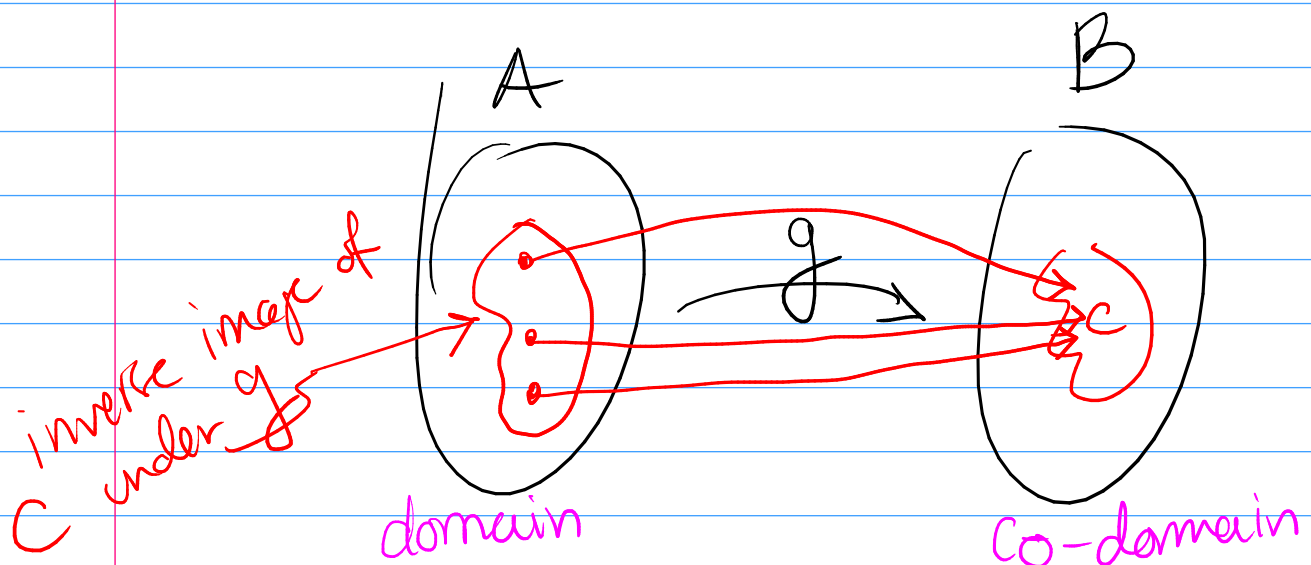
Discrete RVs:

Q: If I know f_X , what is f_Y ?

PMF of X 

 PMF of Y

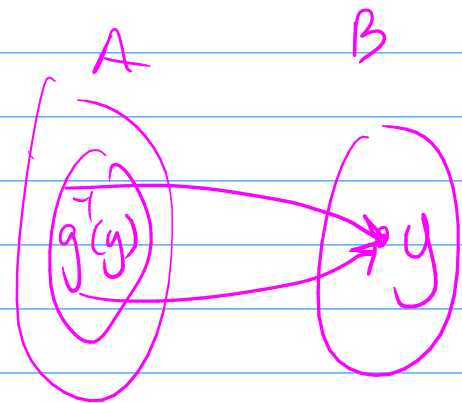
Inverse Images / Preimages



$$g^{-1}(C) = \{x \in A \mid g(x) \in C\}$$

Single pt: $y \in B$

$$g^{-1}(\{y\}) = g^{-1}(y)$$



If g is invertible, then g^{-1} is the inverse image

Notice:

$$f_Y(y) = P(Y=y)$$

$$= P(g(X)=y) \quad (*)$$

Case 1: g is invertible

$$\begin{aligned} (*) &= P(X = g^{-1}(y)) \\ &= f_X(g^{-1}(y)) \end{aligned}$$

Case 2: g isn't invertible

$$(*) = P(X \in \underbrace{g^{-1}(y)}_A)$$

$$= \sum_{x \in g^{-1}(y)} f_X(x)$$

$$= \sum_{x: g(x)=y} f_X(x)$$

$$P(X \in A)$$

$$= \sum_{x \in A} f_X(x)$$

Theorem: If X is discrete and $Y = g(X)$

then

$$f_Y(y) = \sum_{x: g(x)=y} f_X(x).$$

ex. let $X \sim \text{Bin}(n, p)$ n coin tosses
prob. p of H
 $X = \# \text{ heads}$

Consider: $Y = n - X = \# \text{ tails}$

$$y = g(x) = n - x \Leftrightarrow x = g^{-1}(y) = n - y$$

$$f_Y(y) = \sum_{\substack{x: g(x)=y \\ x=g^{-1}(y)}} f_X(x) = \sum_{x=g^{-1}(y)} f_X(x) = f_X(g^{-1}(y))$$

$$f_X(n-y)$$

$$f_X(x) = \binom{n}{x} p^x (1-p)^{n-x}$$

$$= \binom{n}{n-y} p^{n-y} (1-p)^{n-(n-y)}$$

$$\binom{n}{n-y} = \binom{n}{y}$$

$$f_Y(y) = \binom{n}{y} q^y (1-q)^{n-y}$$

$$q = 1-p$$

pmf of Bin(n, q)

$$Y \sim \text{Bin}(n, q=1-p)$$

Cts RVs

Theorem: If X is cts and $Y = g(X)$ then

① if g is increasing then

$$F_Y(y) = F_X(\bar{g}^{-1}(y))$$

true
inverse

② if g is decreasing then

$$F_Y(y) = 1 - F_X(\bar{g}^{-1}(y))$$

pf Case 1: g is inc., then so is \bar{g}^{-1}

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y)$$

$$= P(X \leq \bar{g}^{-1}(y))$$

$$= F_X(\bar{g}^{-1}(y))$$

Case 2: g dec. then so is \bar{g}^{-1}

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \geq \bar{g}^{-1}(y))$$

$$= 1 - P(X < g^{-1}(y))$$

cts
RV



$$= 1 - F_X(g^{-1}(y))$$