

## Lecture 23

### Mutual Independence of RVs

We say that  $X_1, \dots, X_n$  are mutually indep. if for any sets  $A_1, \dots, A_n \subset \mathbb{R}$  we have that

$$\begin{aligned} P(X_1 \in A_1, X_2 \in A_2, \dots, X_n \in A_n) \\ = \\ P(X_1 \in A_1) P(X_2 \in A_2) \dots P(X_n \in A_n) \end{aligned}$$

---

### Theorem: Factorization

If the support of  $\underline{X}$  is a product space then the following 3 statements are equivalent:

- ①  $X_1, \dots, X_n$  are indep.
- ②  $f(x_1, \dots, x_n) = f(x_1) \dots f(x_n)$
- ③  $F(x_1, \dots, x_n) = F(x_1) \dots F(x_n)$

Theorem:

If  $X_1, \dots, X_n$  are indep. then

① If  $g_i: \mathbb{R} \rightarrow \mathbb{R}$  then  
 $g_1(X_1), g_2(X_2), \dots, g_n(X_n)$   
are independent.

②  $E[\underbrace{X_1 X_2 \dots X_n}_{\text{multiply}}] = E[X_1]E[X_2] \dots E[X_n]$

---

Corollary: If  $X_i$  are indep. and

$$Z = \sum_{i=1}^n X_i$$

then

$$M_Z(t) = \prod_{i=1}^n M_{X_i}(t)$$

more generally, if

$$Z = \sum_{i=1}^n (a_i X_i + b_i)$$

then

$$M_z(t) = e^{t \sum_{i=1}^n b_i} \prod_{i=1}^n M_{X_i}(a_i t)$$


---

Ex. If  $X_i \overset{\text{indep}}{\sim} N(\mu_i, \sigma_i^2)$   
and independent.

$$Y = \sum_{i=1}^n (a_i X_i + b_i)$$

$$\sim N\left(\sum_{i=1}^n (a_i \mu_i + b_i), \sum_{i=1}^n a_i^2 \sigma_i^2\right)$$


---

### Multivariate Transformation

If  $g: \mathbb{R}^n \rightarrow \mathbb{R}^n$  and let

$$U = g(\underline{X})$$

$\nwarrow$   $n$ -component rand. vector

and if  $\underline{X}$  has cts components, and

- ①  $g$  is invertible
- ②  $g^{-1}$  is diff'able

then

$$f_u(u) = f_{\underline{\tilde{x}}}(g^{-1}(u)) / |\det J|$$

$J$  is  $n \times n$

$$J_{ij} = \frac{\partial g_i^{-1}}{\partial u_j}$$

## Means / Variances for MV RVs

Uni:  $E[X] \in \mathbb{R}$

$$\text{Var}(X) = E[(X - E[X])^2] \geq 0$$

MV:  $\underline{\tilde{x}} = \begin{bmatrix} X_1 \\ \vdots \\ X_n \end{bmatrix}$

mean:  $\mu = E[\underline{\tilde{x}}] = \begin{bmatrix} E[X_1] \\ E[X_2] \\ \vdots \\ E[X_n] \end{bmatrix} \in \mathbb{R}^n$

Covariance Matrix:

Variance:  $\Sigma = \text{Cov}(\underline{X}) \in \mathbb{R}^{n \times n}$

$$\Sigma_{ij} = \text{Cov}(X_i, X_j)$$

note:  $\Sigma_{ii} = \text{Cov}(X_i, X_i) = \text{Var}(X_i)$

$$\Sigma = \begin{bmatrix} \text{Var}(X_1) & \text{Cov}(X_1, X_2) & \dots \\ \text{Cov}(X_2, X_1) & \text{Var}(X_2) & \dots \\ \vdots & \vdots & \ddots \end{bmatrix}$$

$\underbrace{\hspace{10em}}_{n \times n}$   
 $\underbrace{\hspace{4em}}_{n \times 1} \quad \underbrace{\hspace{4em}}_{1 \times n}$

$$\text{Cov}(\underline{X}) = E[(\underline{X} - E[\underline{X}])(\underline{X} - E[\underline{X}])^T]$$

Uni:  $\text{Var}(X) = E[(X - E[X])^2]$

Theorem:

Assume  $a \in \mathbb{R}^m$ ,  $B \in \mathbb{R}^{m \times n}$  then

$$\textcircled{1} E[a + B\underline{X}] = a + B E[\underline{X}]$$

$\underbrace{\hspace{2em}}_{m \times 1} \quad \underbrace{\hspace{2em}}_{m \times n} \quad \underbrace{\hspace{2em}}_{n \times 1}$

$$\textcircled{2} \quad \underbrace{\text{Cov}(a + B\underline{X})}_{m \times m} = B \underbrace{\text{Cov}(\underline{X})}_{n \times n} B^T$$

$\underbrace{\hspace{10em}}_{m \times m}$

## Multivariate Normal

$$\underline{X} \sim N(\underline{\mu}, \underline{\Sigma})$$

$\swarrow$   $\nearrow$   
 $n$ -vector  $n \times n$

$$f(\underline{x}) = (2\pi)^{-n/2} (\det \underline{\Sigma})^{-1/2} \exp\left(-\frac{1}{2}(\underline{x} - \underline{\mu})^T \underline{\Sigma}^{-1}(\underline{x} - \underline{\mu})\right)$$

Special Case:  $\underline{\mu} = 0$  and  $\underline{\Sigma} = I$   
 called the standard MV norm.

Theorem: If  $\underline{X} \sim N(\underline{\mu}, \underline{\Sigma})$  and  
 $a \in \mathbb{R}^m$ ,  $B \in \mathbb{R}^{m \times n}$  then

$$a + B\underline{X} \sim N(a + B\underline{\mu}, B\underline{\Sigma}B^T)$$



FINAL



The cov. mtx.  $\Sigma$  is

- ① Symmetric ( $\Sigma = \Sigma^T$ )
- ② pos (semi-) definite

A Pos semi-def:  $x^T A x \geq 0$   
gen. of  $A \geq 0$

A pos def:  $x^T A x > 0 \quad \forall x \neq 0$   
gen. of  $A > 0$

### Indicator Functions

$$\mathbb{I}(\text{statement}) = \begin{cases} 1 & \text{statement true} \\ 0 & \text{else} \end{cases}$$

$X \sim \text{Exp}(\lambda)$

$$\begin{aligned} f(x) &= \lambda e^{-\lambda x} \quad \text{for } x > 0 \\ &= \lambda e^{-\lambda x} \mathbb{I}(x > 0) \end{aligned}$$

Often:  $f(x) = \mathbb{1}(x \in \text{Support})$

Checking independence:

$$f(x, y) = \lambda e^{-\lambda x} e^{-y} \quad \text{for } x > 0 \text{ and } y > 0$$
$$= \lambda e^{-\lambda x} e^{-y} \mathbb{I}(x > 0) \mathbb{I}(y > 0)$$

$$= \underbrace{(\lambda e^{-\lambda x} \mathbb{I}(x > 0))}_{f_X(x)} \underbrace{(e^{-y} \mathbb{I}(y > 0))}_{f_Y(y)}$$