

Lecture 10: Rao - Blackwell Theorem

Some facts:

① If $\hat{\theta}$ is unbiased for $T(\theta)$

$$\text{i.e. } E\hat{\theta} = T(\theta)$$

Let W be some function of \underline{X} i.e. $W = W(\underline{X})$
(maybe a stat., maybe not)

Consider $\psi(W) = \psi = E[\hat{\theta} | W]$ ← a RV

notice that

$$E[\psi] = E[E[\hat{\theta} | W]] = E[\hat{\theta}] = T(\theta)$$

$$\begin{aligned} E[X|Y] &= g(Y) \\ E[E[X|Y]] &= EX \end{aligned}$$

If ψ is a statistic then it is unbiased for $T(\theta)$

② $\text{Var}(\psi) \leq \text{Var}(\hat{\theta})$.

Law of Total Variance

$$\text{Var} \hat{\theta} = \underbrace{\text{Var} E[\hat{\theta} | W]}_{\psi} + \underbrace{E \text{Var}(\hat{\theta} | W)}_{\geq 0}$$

So $\text{Var } \hat{\theta} = \text{Var } \varphi + a$ where $a \geq 0$

i.e. $\text{Var}(\hat{\theta}) \geq \text{Var}(\varphi)$

Ex. $X_n \stackrel{\text{iid}}{\sim} N(\theta, 1)$

and let $\hat{\theta} = \frac{1}{2}(X_1 + X_2)$

$\hat{\theta}$ unbiased for θ : $E\hat{\theta} = \frac{1}{2}EX_1 + \frac{1}{2}EX_2 = \frac{1}{2}\theta + \frac{1}{2}\theta = \theta$

$\text{Var } \hat{\theta} = \frac{1}{2} \text{Var}(X_1) = \frac{1}{2}$

$W = X_1$

$\varphi = E[\hat{\theta}|W] = E\left[\frac{1}{2}(X_1 + X_2) | X_1\right]$

$E[z|z=5] = 5$

$E[z|z=6] = 6$

$E[z|z=7] = 7$

$E[z|z] = z$

$= \frac{1}{2}E[X_1|X_1] + \frac{1}{2}E[X_2|X_1]$

$= \frac{1}{2}X_1 + \frac{1}{2}E[X_2]$

$= \frac{1}{2}X_1 + \frac{1}{2}\theta$

$E[\varphi] = E\left[\frac{1}{2}X_1 + \frac{1}{2}\theta\right] = \frac{1}{2}EX_1 + \frac{1}{2}\theta = \frac{1}{2}\theta + \frac{1}{2}\theta = \theta$

$$\text{Var}(\varphi) = \text{Var}\left(\frac{1}{2}X_1 + \frac{1}{2}X_2\right) = \frac{1}{4} \text{Var}X_1 = \frac{1}{4}$$

$$\text{Let } W = \bar{X}$$

$$\begin{aligned}\varphi &= E[\hat{\theta} | W] = E\left[\frac{1}{2}(X_1 + X_2) \mid \bar{X}\right] \\&= \frac{1}{2} E[X_1 | \bar{X}] + \frac{1}{2} E[X_2 | \bar{X}] \\&\quad \swarrow \quad \searrow \text{same} \\&= \frac{1}{2} 2 E[X_n | \bar{X}] \\&= E[X_n | \bar{X}] \\&= \frac{1}{N} N E[X_n | \bar{X}] \\&= \frac{1}{N} \sum_{n=1}^N E[X_n | \bar{X}] \\&= E\left[\frac{1}{N} \sum_n X_n \mid \bar{X}\right] \\&= E[\bar{X} | \bar{X}]\end{aligned}$$

$$\varphi = \bar{X}$$

and by our theorem:

$$(1) E\varphi = E\bar{X} = \theta$$

$$(2) \text{Var} \varphi = \text{Var}(\bar{X}) \leq \text{Var}(\hat{\theta})$$

Theorem: Rao-Blackwell Theorem

If $\hat{\theta}$ is unbiased for $T(\theta)$ and W is sufficient for θ , then if

$$\varphi = E[\hat{\theta}|W]$$

① $E\varphi = T(\theta)$

② $\text{Var } \varphi \leq \text{Var } \hat{\theta}$

③ φ is a statistic (i.e. no θ in formula)

Pf of ③

$$\varphi = E[\hat{\theta}(x)|W]$$

$$= \int \underbrace{\hat{\theta}(x)}_{\text{no } \theta} \underbrace{f_{x|W}(x)}_{\text{no } \theta} dx$$

doesn't depend on θ

sufficient

$$Eg(x) = \int g(x)f(x)dx$$

Theorem: Lehmann - Scheffe Theorem

Let W be a (complete) sufficient statistic for θ and let $\hat{\theta}$ be an unbiased est. for $T(\theta)$ that depends on \underline{X} only through

$$W: \quad \hat{\theta} = \hat{\theta}(W) = \hat{\theta}(W(\underline{X}))$$

then $\hat{\theta}$ is the UMVUE for $T(\theta)$.

Basically: If I can find a unbiased est for $T(\theta)$ from a S.S. it is the UMVUE.

Ex. $X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

μ is unknown, σ^2 is known

Want: the UMVUE for μ .

Use Lehmann-Scheffe.

① Find a SS for μ , \bar{X} .

② Guess a fn of \bar{X} unbiased for μ .

$$\hat{\mu} = \bar{X}$$

Since $E\hat{\mu} = E\bar{X} = \mu$

and is a fn of the SS, then it is the UMVUE.

Ex. let $T(\mu) = \mu^2$.

① get SS for μ , \bar{X}

② Find a fn of \bar{X} that is unbiased for $T(\mu) = \mu^2$.

$$\begin{aligned} E[\bar{X}^2] &= \text{Var}(\bar{X}) + E[\bar{X}]^2 \\ &= \sigma^2/N + \mu^2 \end{aligned}$$

Consider $\hat{\mu}^2 = \bar{X}^2 - \sigma^2/N$

$$\text{Now } E[\hat{\mu}^2] = E[\bar{X}^2] - \sigma^2/N = \mu^2.$$

So $\bar{X}^2 - \sigma^2/N$ is the UMVUE for μ^2 .

Lehman-Schoffe Takeaway

How to find the UMVUE

- ① Find a SS for θ - call it W
- ② Find a fn of W that is unbiased for $T(\theta)$

i) Guess a fn $\hat{\theta}(w)$ so that
$$E[\hat{\theta}(w)] = T(\theta)$$

ii) Use Rao-Blackwell,
get some really simple unbiased est. V
and form
$$\hat{\theta} = E[V | W]$$

Ex. $X_n \stackrel{iid}{\sim} U(0, \theta)$

What is the UMVUE for $T(\theta) = \theta$.

- ① Find a SS for θ : $X_{(N)}$
- ② Find some fn of $X_{(N)}$ that is unbiased for θ

Last time: $E[X_{(N)}] = \frac{N}{N+1} \theta$

So if $\hat{\theta} = \frac{N+1}{N} X_{(N)}$

then $E\hat{\theta} = \theta$ and $\hat{\theta}$ is a fn of the SS

so $\hat{\theta}$ is the UMVUE.

Pf. of Lehmann-Scheffe

If $\hat{\theta} = \hat{\theta}(W)$ where W is (complete) sufficient
and unbiased for $T(\theta)$
then for any other unbiased est V of
 $T(\theta)$ want to show that

$$\text{Var}(V) \geq \text{Var}(\hat{\theta})$$

We'll do this by Rao-Blackwellizing V using W .

We'll show that $E[V|W] = \hat{\theta}$.

Rao-Blackwell says:

$$\text{if } \varphi = \varphi(W) = E[V|W]$$

then ① $E\varphi = T(\theta)$

② $\text{Var}\varphi \leq \text{Var}(V)$

③ φ is a stat

If I can show that $\varphi = \hat{\theta}$ then

$$\text{Var } \hat{\theta} = \text{Var } \varphi \leq \text{Var}(V).$$

This is what I want.

Consider $g(w) = \hat{\theta}(w) - \varphi(w)$

we'll show that $g \equiv 0 \quad \forall \theta$

Completeness of W

$$\mathbb{E}_{\theta}[h(w)] = 0 \quad \forall \theta \Leftrightarrow h \equiv 0$$

$$\begin{aligned} \text{Notice } \mathbb{E}_{\theta} g(w) &= \mathbb{E}[\hat{\theta}] - \mathbb{E}[\varphi] = T(\theta) - T(\theta) \\ &= 0 \quad \forall \theta \end{aligned}$$

If W is complete then $g \equiv 0$

$$\text{So } g(w) = 0 = \hat{\theta} - \varphi$$

$$\text{So } \hat{\theta} = \varphi.$$
