

Lecture 5: Maximum Likelihood Estimation (MLE)

MoM examples

Ex. $X_n \stackrel{iid}{\sim} U(0, \theta)$

Let's get the MoM est. for θ .

$$\mu_1 = E[X_n] = m_1 = \frac{1}{N} \sum_{n=1}^N X_n = \bar{X}$$

$$\mu_1 = E[X_1] = \frac{0+\theta}{2} = \frac{\theta}{2} = \bar{X}$$

Solve for θ then $\hat{\theta}_{\text{mom}} = 2\bar{X}$

Ex. $X_n \stackrel{iid}{\sim} \text{Beta}(\alpha, 1)$

What is the MoM est. for α ?

$$\mu_1 = EX_n = m_1 = \bar{X}$$

$$\mu_1 = EX_n = \frac{\alpha}{\alpha+1} = \bar{X}$$

Solve for α : $\alpha = (\alpha+1)\bar{X}$

$$\Rightarrow \alpha = \alpha\bar{X} + \bar{X}$$

$$\Rightarrow \alpha - \alpha\bar{X} = \bar{X}$$

$$\Rightarrow \alpha(1-\bar{X}) = \bar{X}$$

$$\Rightarrow \hat{\alpha}_{\text{mom}} = \frac{\bar{X}}{1-\bar{X}}$$

Maximum Likelihood Estimator (MLE)

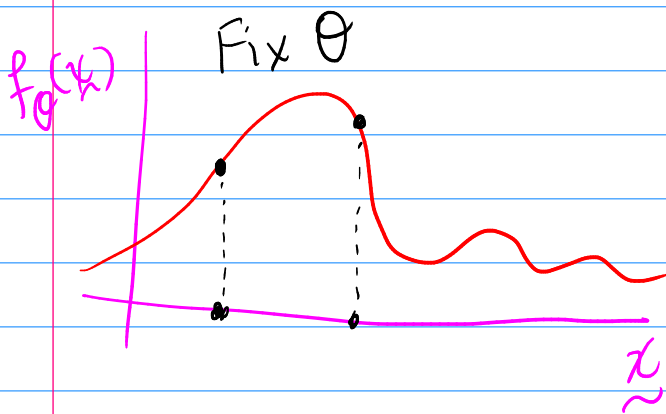
Assume $X_n \stackrel{iid}{\sim} f_\theta$ where $\theta \in \Theta$

recall: joint dist of my data

$$f_\theta(\underline{x}) = \prod_{n=1}^N f_\theta(x_n)$$

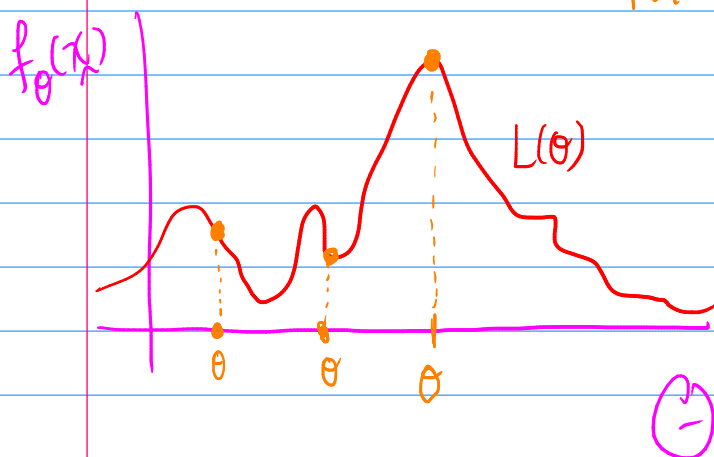
typically we look at this as a fn of \underline{x}

Way 1: $f_\theta: \mathbb{R}^N \rightarrow \mathbb{R}$



We can also fix \underline{x} and think of this as a fn of θ

This view is called the likelihood function



$$L: \Theta \rightarrow \mathbb{R}$$

$$L(\theta) = f_\theta(\underline{x})$$

MLE says choose $\hat{\theta}_{MLE}$ as the val. of θ that maximizes L

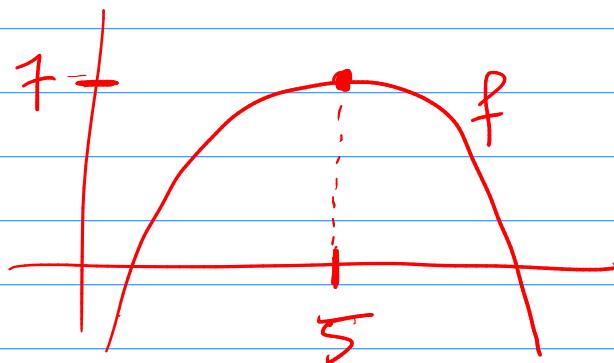
Defn: Maximum Likelihood estimator

$$\hat{\theta}_{MLE} = \underset{\theta \in (-)}{\operatorname{argmax}} L(\theta)$$

Ex,

$$\max_x f(x) = 7$$

$$\underset{x}{\operatorname{argmax}} f(x) = 5$$



Often we work with the log-likelihood fn

$$l(\theta) = \log L(\theta)$$

$\log = \text{natural log}$

Alt. defn of $\hat{\theta}_{MLE}$

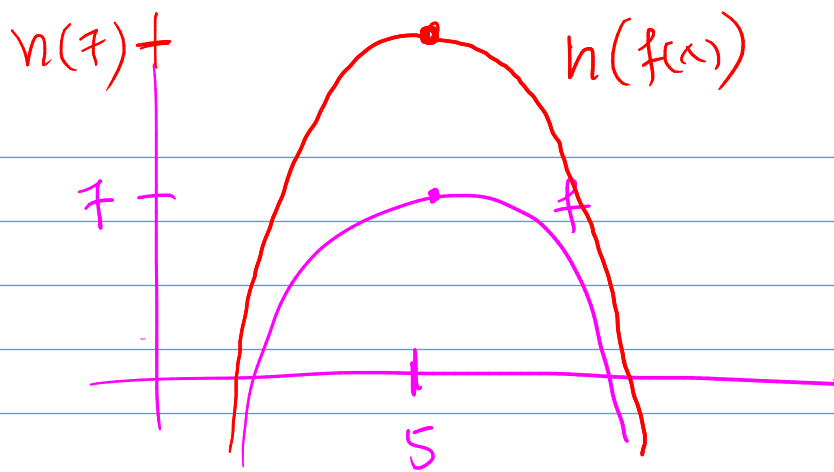
$$\hat{\theta}_{MLE} = \underset{\theta \in (-)}{\operatorname{argmax}} l(\theta)$$

\uparrow equivalent defn.

b/c \log is an increasing function

$$\underset{x}{\operatorname{argmax}} f(x) = \underset{x}{\operatorname{argmax}} h(f(x))$$

\uparrow inc. fn.



by defn an inc. fn says $x_1 < x_2$ then $h(x_1) < h(x_2)$

Ex. $X_n \stackrel{iid}{\sim} N(0, 1)$ where $\theta \in \mathbb{R}$

what's the MLE?

① let's get $l(\theta)$

$$L(\theta) = f_{\theta}(\underline{x}) = \prod_{n=1}^N f_{\theta}(x_n) = \prod_{n=1}^N \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}(x_n - \theta)^2\right)$$

$$= (2\pi)^{-N/2} \exp\left(\sum_{n=1}^N -\frac{1}{2}(x_n - \theta)^2\right)$$

$$L(\theta) = (2\pi)^{-N/2} \exp\left(-\frac{1}{2} \sum_{n=1}^N (x_n - \theta)^2\right) \propto e^{-\theta^2}$$

$$l(\theta) = \log L(\theta) = -\frac{N}{2} \log(2\pi) - \frac{1}{2} \sum_{n=1}^N (x_n - \theta)^2$$

② let's take a derivative $\frac{\partial l}{\partial \theta}$

$$\frac{\partial \ell}{\partial \theta} = 0 - \frac{1}{2} \sum_n \underbrace{\frac{\partial}{\partial \theta} (x_n - \theta)^2}_{\rightarrow 2(x_n - \theta)(-1) = -2(x_n - \theta)}$$

$$= -\frac{1}{2} \sum_n -2(x_n - \theta)$$

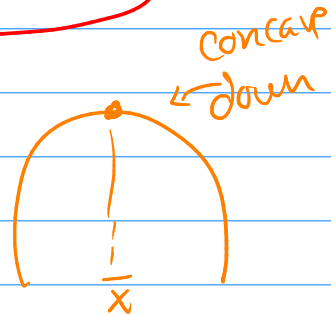
$$= \sum_n (x_n - \theta)$$

$$= \sum_n x_n - N\theta = 0$$

Calc says set $\frac{\partial \ell}{\partial \theta} = 0$ and solve for θ

$$\Rightarrow \sum_n x_n = N\theta$$

$$\Rightarrow \boxed{\hat{\theta}_{MLE} = \frac{1}{N} \sum_n x_n = \bar{x}}$$



Technically need to look at $\frac{\partial^2 \ell}{\partial \theta^2} < 0$

and also need to consider $\lim_{\theta \rightarrow \pm\infty} \ell(\theta) = -\infty$

$$\frac{\partial^2 \ell}{\partial \theta^2} = \frac{\partial}{\partial \theta} \left(\sum_n x_n - N\theta \right) = -N < 0$$

Theorem: MLE are based on sufficient stats

$$\hat{\theta}_{MLE} = \text{function}(T)$$

\uparrow sufficient for θ

Pf. By factorization theorem if T is sufficient then

$$L(\theta) = f_{\theta}(\underline{x}) = \underbrace{g(T, \theta)}_{\text{only shows up through } T} h(\underline{x})$$

and

$$\begin{aligned}\hat{\theta}_{MLE} &= \underset{\theta \in \Theta}{\operatorname{argmax}} L(\theta) = \underset{\theta \in \Theta}{\operatorname{argmax}} g(T, \theta) h(\underline{x}) \\ &= \underset{\theta \in \Theta}{\operatorname{argmax}} g(T, \theta) \\ &\rightarrow \text{function}(T)\end{aligned}$$

Ex. Let $X_n \stackrel{iid}{\sim} \text{Bern}(p)$, $p \in [0, 1]$

What is \hat{p}_{MLE} ?

① let's get $L(p)$ and $l(p)$

$$L(p) = f_p(\underline{x}) = \prod_{n=1}^N f_p(x_n)$$

$$= \prod_{n=1}^N p^{x_n} (1-p)^{1-x_n} \mathbb{1}(x_n = 0 \text{ or } 1)$$

$$= p^{\sum x_n} (1-p)^{\sum (1-x_n)} \prod_n \mathbb{1}(x_n = 0 \text{ or } 1)$$

$$= p^{N\bar{x}} (1-p)^{N-N\bar{x}} \prod_n \mathbb{1}(x_n = 0 \text{ or } 1)$$

Bern!

$$f_p(x) = p^x (1-p)^{1-x} \mathbb{1}(x=0 \text{ or } 1)$$

$$= \begin{cases} p & x=1 \\ 1-p & x=0 \end{cases}$$

$$\sum_{n=1}^N x_n = N\bar{x}$$

$$\ell(p) = \log L(p) = N\bar{x} \log p + (N - N\bar{x}) \log(1-p) + \log\left(\prod_n \mathbb{1}(x_n = 0 \text{ or } 1)\right)$$

(2) take a derivative wrt. p

$$\frac{\partial \ell}{\partial p} = \frac{N\bar{x}}{p} - \frac{N - N\bar{x}}{1-p}$$

always disappear if parameter doesn't change support

(3) Set $\frac{\partial \ell}{\partial p} = 0$ and solve for p

$$\frac{N\bar{x}}{p} - \frac{N - N\bar{x}}{1-p} = 0 \Rightarrow \frac{N\bar{x}}{p} = \frac{N - N\bar{x}}{1-p}$$

$$\Rightarrow N\bar{x}(1-p) = (N - N\bar{x})p$$

$$\Rightarrow N\bar{x} - \cancel{pN\bar{x}} = Np - \cancel{pN\bar{x}}$$

$$\Rightarrow N\bar{x} = Np$$

$$\Rightarrow \boxed{\hat{p}_{MLE} = \bar{x}} = \frac{\sum x_n}{N} = \frac{\# \text{ of } 1\text{'s}}{\text{total } \#}$$

= pct. of 1s in my data

Consider

$$\eta = \frac{p}{1-p} = \text{odds}$$

I might want $\hat{\eta}_{MLE}$.

$$\boxed{\eta = \frac{p}{1-p}} \Rightarrow (1-p)\eta = p$$

$$\Rightarrow \eta - p\eta = p$$

$$\Rightarrow \eta = p(1+\eta)$$

$$\Rightarrow \boxed{p = \frac{\eta}{1+\eta}}$$

lets get the likelihood in terms of η
(not p)

$$L(p) = p^{N\bar{x}} (1-p)^{N-N\bar{x}}$$

$$\boxed{L(\eta) = \left(\frac{\eta}{1+\eta}\right)^{N\bar{x}} \left(1 - \frac{\eta}{1+\eta}\right)^{N-N\bar{x}}}$$

$$\hat{\eta}_{MLE} = \arg \max_{\eta} L(\eta)$$