

Lecture 18: Likelihood Ratio Test

Recall: $L(\theta) = f_{\theta}(\underline{x})$

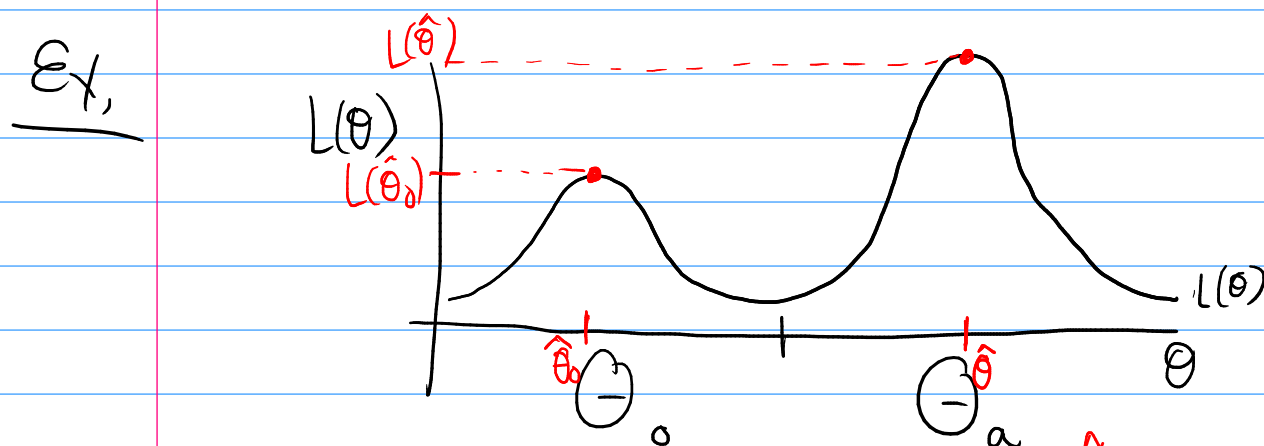
Want to test a hypothesis

$$H_0: \theta \in \Theta_0 \quad \text{v.} \quad H_a: \theta \in \Theta_a$$

the Likelihood Ratio Test Statistic is defined as

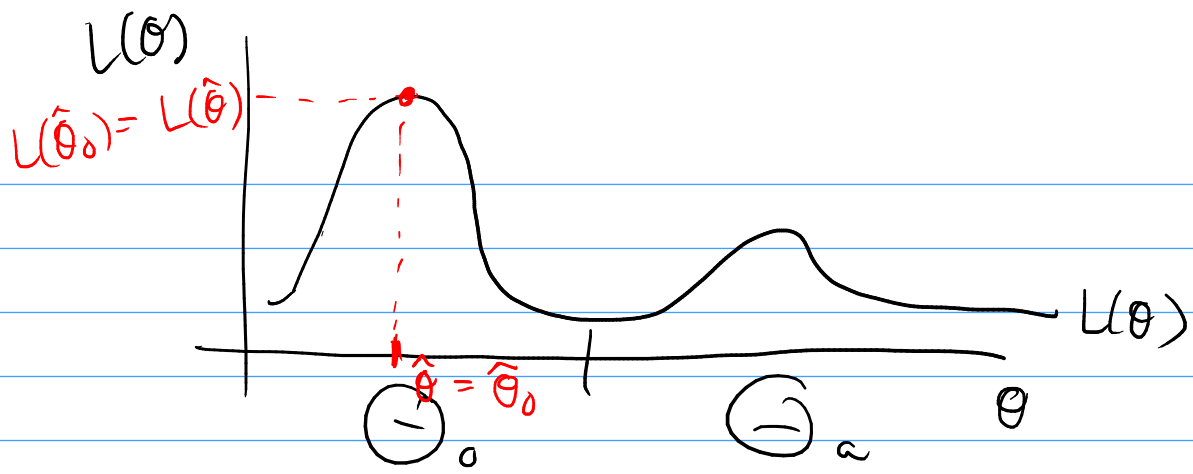
$$\lambda(\underline{x}) = \frac{\max_{\theta \in \Theta_0} L(\theta)}{\max_{\theta \in \Theta} L(\theta)} = \frac{\text{max. val. of } L(\theta) \text{ in } \Theta_0}{\text{max val. of } L(\theta) \text{ in } \Theta}$$

$$\hat{\theta} = \text{MLE} \quad \hat{\theta}_0 = \text{MLE restricted to } \Theta_0 \quad = \frac{L(\hat{\theta}_0)}{L(\hat{\theta})} \leq 1$$

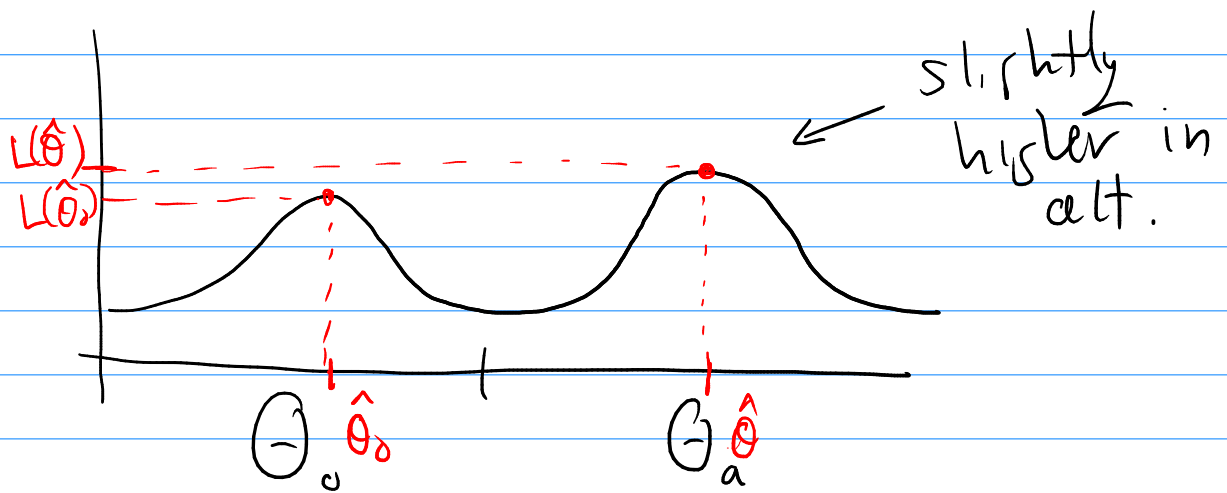


$$\text{So } \lambda(\underline{x}) = \frac{L(\hat{\theta}_0)}{L(\hat{\theta})} < 1$$

↑ probably reject H_0



$S_0 \quad \lambda = \frac{L(\hat{\theta}_0)}{L(\hat{\theta})} = 1$ ↑ don't reject H_0



$\lambda \approx 1$ ↑ maybe reject H_0 ?

Overall: $\lambda \approx 1$ H_a not much more likely than H_0

$\lambda \approx 0$ H_a much more likely than H_0

The likelihood ratio test says reject when

$$\lambda = \frac{L(\hat{\theta}_0)}{L(\hat{\theta})} \leq c$$

↑ some threshold we can choose to balance Type I and II errors

Typically, c small \Rightarrow less type I
more type II
 c large \Rightarrow reverse

Equiv. $R = \{x \in \mathcal{X} : \lambda(x) \leq c\}$

Ex. $X_n \stackrel{iid}{\sim} N(\theta, \sigma^2)$

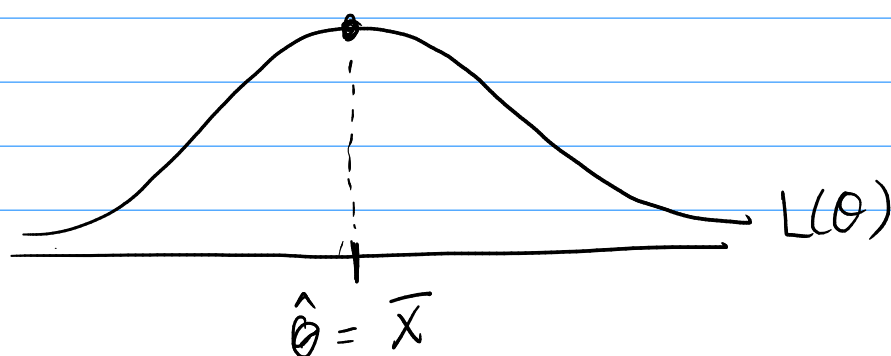
$H_0: \theta \leq a$ v. $H_a: \theta > a$

Let's form the LRT.

$\Theta = \mathbb{R}$, $\Theta_0 = (-\infty, a]$, $\Theta_a = (a, \infty)$

$$L(\theta) = \prod_n \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x_n - \theta)^2\right)$$
$$= \dots = (2\pi)^{-N/2} (\sigma^2)^{-N/2} \exp\left(-\frac{1}{2\sigma^2} \sum_n (x_n - \theta)^2\right)$$

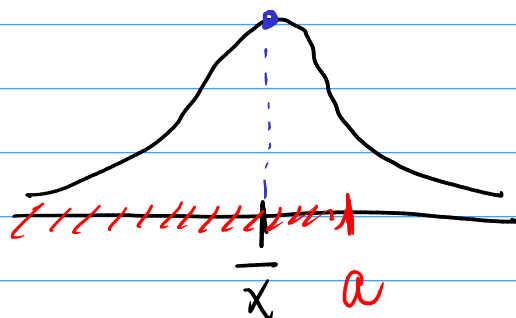
looks quadratic in θ
Kinda like $e^{-\theta^2}$



$$\lambda = \frac{L(\hat{\theta}_0)}{L(\hat{\theta})} = \frac{L(\hat{\theta}_0)}{L(\bar{x})}$$

To find $\hat{\theta}_0$, two cases $\bar{x} \leq a$ and $\bar{x} > a$

$\bar{x} \leq a$

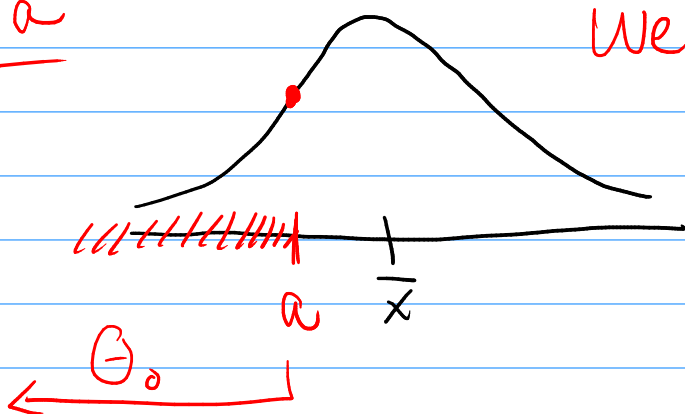


$$\hat{\theta}_0 = \operatorname{argmax}_{\theta \in \Theta_0} L(\theta)$$

$$\Theta_0 = (-\infty, a]$$

we get $\hat{\theta}_0 = \bar{x}$

$\bar{x} > a$



we find $\hat{\theta}_0 = a$

So
$$\hat{\theta}_0 = \begin{cases} \bar{x} & \text{if } \bar{x} \leq a \\ a & \text{if } \bar{x} > a \end{cases}$$

So
$$\lambda(\bar{x}) = \frac{L(\hat{\theta}_0)}{L(\hat{\theta})} = \begin{cases} \frac{L(\bar{x})}{L(\bar{x})} = 1 & \bar{x} \leq a \\ \frac{L(a)}{L(\bar{x})} & \bar{x} > a \end{cases}$$

$H_0: \theta \leq a$

$H_a: \theta > a$

So LRT says to reject when

$$\frac{L(a)}{L(\bar{X})} \leq c \quad \text{where } c \in (0,1)$$

$$\frac{L(a)}{L(\bar{X})} = \frac{(\cancel{2\pi})^{-N/2} (\cancel{\sigma^2})^{-N/2} \exp\left(-\frac{1}{2\sigma^2} \sum_n (\cancel{x_n} - a)^2\right)}{(\cancel{2\pi})^{-N/2} (\cancel{\sigma^2})^{-N/2} \exp\left(-\frac{1}{2\sigma^2} \sum_n (\cancel{x_n} - \bar{X})^2\right)}$$

$$= \frac{\exp\left(-\frac{1}{2\sigma^2} \sum_n (\cancel{x_n}^2 - 2a\cancel{x_n} + a^2)\right)}{\exp\left(-\frac{1}{2\sigma^2} \sum_n (\cancel{x_n}^2 - 2\bar{X}\cancel{x_n} + \bar{X}^2)\right)}$$

$$\frac{e^{a+b}}{e^{a+c}}$$

$$= \exp\left(-\frac{1}{2\sigma^2} (-2aN\bar{X} + Na^2 + 2N\bar{X}^2 - N\bar{X}^2)\right)$$

$$= \exp\left(-\frac{1}{2\sigma^2} (N\bar{X}^2 - 2aN\bar{X} + Na^2)\right)$$

$$= \exp\left(-\frac{N}{2\sigma^2} (\bar{X} - a)^2\right)$$

So the LRT says to reject when

$$\lambda = \exp\left(-\frac{N}{2\sigma^2} (\bar{X} - a)^2\right) \leq c$$

$$\Leftrightarrow -\frac{N}{2\sigma^2} (\bar{X} - a)^2 \leq \log c$$

$$\Leftrightarrow \frac{N}{\sigma^2} (\bar{X} - a)^2 \geq -2 \log c$$

$$\sqrt{a^2} = |a|$$

$$\Leftrightarrow \frac{\sqrt{N}}{\sigma} (\bar{X} - a) \geq \sqrt{-2 \log c}$$

we assume $\bar{X} > a$
 $\bar{X} - a > 0$

$$\Leftrightarrow \frac{\bar{X} - a}{\sigma/\sqrt{N}} \geq \sqrt{-2 \log c}$$

c^*
 some number > 0

So LRT says to reject when

$$\boxed{\frac{\bar{X} - a}{\sigma/\sqrt{N}} \geq c^*} \Leftrightarrow \bar{X} \geq a + c^* \frac{\sigma}{\sqrt{N}}$$

How should we choose c^* ?

Maybe want the LRT to be level α test?
 (size α test)

$$\text{i.e. } \max_{\theta \in \Theta_0} \underbrace{P(\text{reject})}_{\beta(\theta)} \leq \alpha \quad [\text{level}]$$

$$\beta(\theta) = P_{\theta}(\lambda \leq c) = P_{\theta}\left(\frac{\bar{X} - a}{\sigma/\sqrt{N}} \geq c^*\right)$$

$$= P_{\theta}\left(\frac{\bar{X} - a}{\sigma/\sqrt{N}} + \frac{a - \theta}{\sigma/\sqrt{N}} \geq c^* + \frac{a - \theta}{\sigma/\sqrt{N}}\right)$$

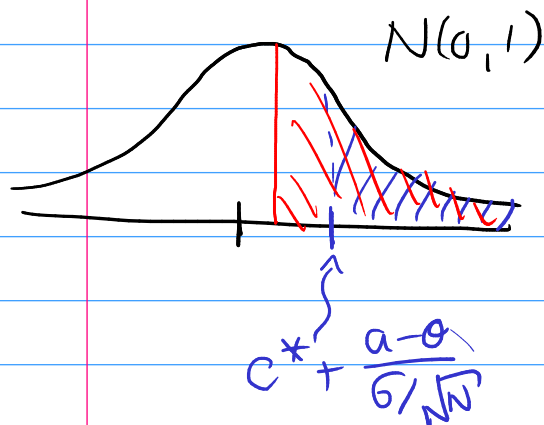
$$X_n \stackrel{iid}{\sim} N(\theta, \sigma^2)$$

$$\bar{X} \sim N(\theta, \sigma^2/n)$$

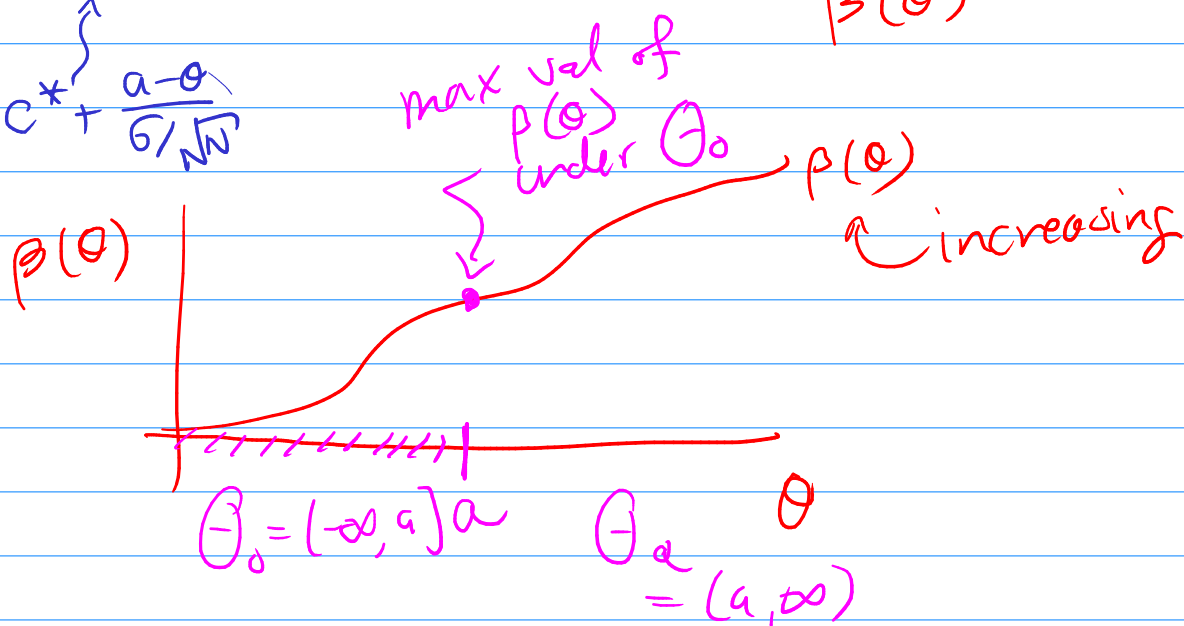
$$= P_{\theta} \left(\frac{\bar{X} - \theta}{\sigma/\sqrt{n}} \geq c^* + \frac{a - \theta}{\sigma/\sqrt{n}} \right)$$

$$\frac{\bar{X} - \theta}{\sigma/\sqrt{n}} \sim N(0, 1)$$

$$= \underbrace{P(Z \geq c^* + \frac{a - \theta}{\sigma/\sqrt{n}})}_{Z \sim N(0, 1)} \beta(\theta)$$



as $\theta \uparrow$, $P(\dots) \uparrow$
 $\bar{\beta}(\theta)$



So $\max \beta(\theta)$ under H_0 is at $\theta = a$

So if I want

$$\max_{\theta \in \theta_0} \beta(\theta) = \beta(a) \leq \alpha$$

So choose c^* so that $\beta(a) \leq \alpha$

$$\beta(a) = P\left(Z \geq c^* + \frac{a - a}{\sigma/\sqrt{n}}\right) = P(Z \geq c^*)$$

So choose c^* so that $\overbrace{1 - F_Z(c^*)} = \alpha$

If F_Z is the CDF of a $N(0,1)$

then want

$$1 - F_Z(c^*) = \alpha$$

$$\text{or } \boxed{c^* = F_Z^{-1}(1 - \alpha)} = z_\alpha$$

