

Lecture 8 : UMVUEs

Defn: Uniformly Minimum Variance Unbiased Estimator (UMVUE)

Note: If $B(\hat{\theta}) = 0$ then $MSE(\hat{\theta}) = \text{Var } \hat{\theta}$

We call θ^* the UMVUE of $T(\theta)$

if

- ① unbiased for $T(\theta)$

$$E[\theta^*] = T(\theta)$$

- ② minimum variance — uniformly

$$\text{Var } \theta^* \leq \text{Var } \hat{\theta} \quad \forall \theta \in \Theta$$

for all $\hat{\theta}$ that are unbiased for $T(\theta)$

Defn: Score Basically $\frac{\partial \ell}{\partial \theta}$ but viewed as random

recall: \underline{x} deterministic, \underline{X} is random

If $\underline{X}_n \stackrel{\text{iid}}{\sim} f_{\theta}$ when $\theta \in \Theta$ then the score is

$$S = S_{\theta} = S_{\theta}(\underline{X}) = \frac{\partial \log f_{\theta}(\underline{X})}{\partial \theta} = \frac{\partial \ell}{\partial \theta}$$

$f_{\theta}(\underline{x})$

E_{X_1} $X_n \stackrel{iid}{\sim} \text{Exp}(\lambda)$ where $\lambda > 0$

$$\text{then } L(\lambda) = \prod_{n=1}^N \lambda e^{-\lambda x_n} \mathbb{1}(x_n > 0) \\ = \lambda^N e^{-\lambda \sum_n x_n} \mathbb{1}(x_{(1)} > 0)$$

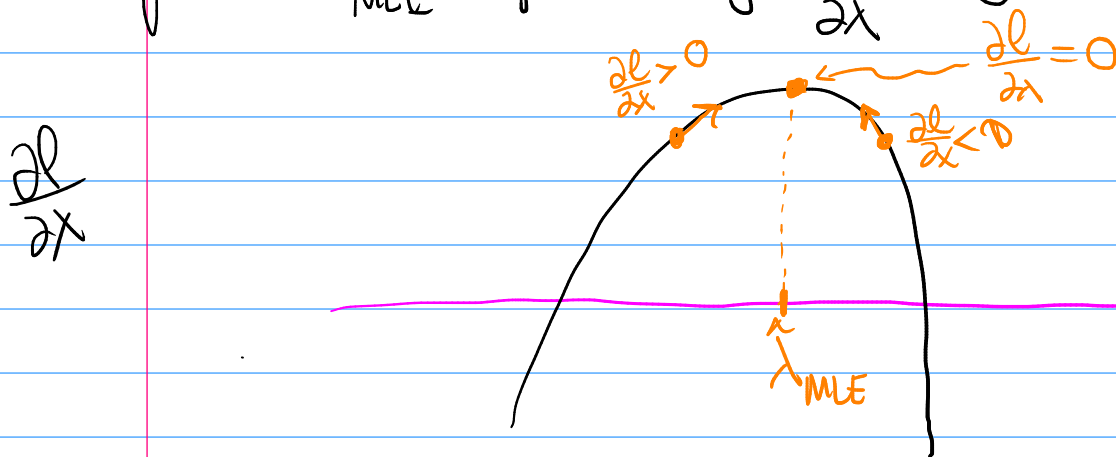
$$\ell(\lambda) = \log L(\lambda) = N \log \lambda - \lambda \sum_n x_n + \log \mathbb{1}(x_{(1)} > 0)$$

$$\frac{\partial \ell}{\partial \lambda} = \frac{N}{\lambda} - \sum_n x_n$$

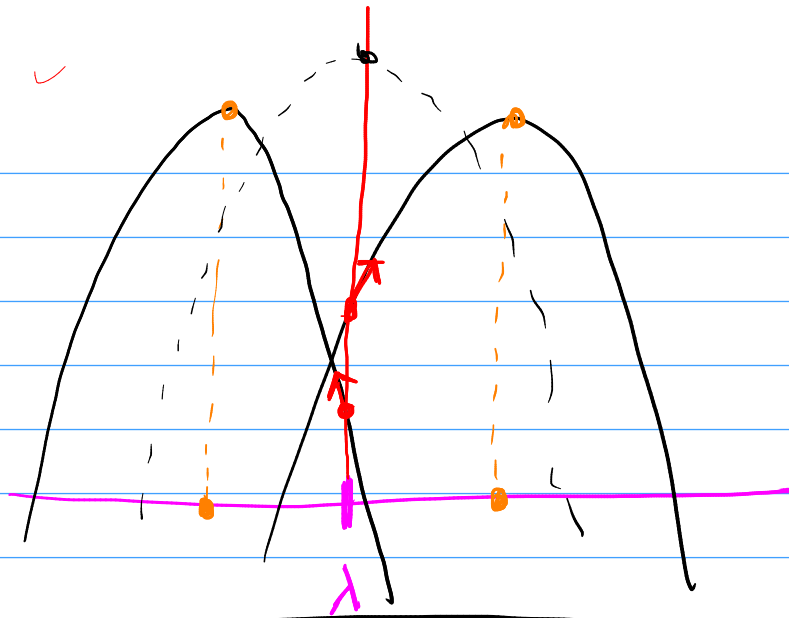
Score promotes x_n to X_n

$$S_\lambda \left| S_\lambda = \frac{N}{\lambda} - \sum_n X_n \right. \leftarrow \text{now random}$$

I get $\hat{\lambda}_{MLE}$ by setting $\frac{\partial \ell}{\partial \lambda} = 0$



Score \approx random ver. of $\frac{\partial \ell}{\partial \lambda}$



Theorem: $E[S_\theta] = 0$.

Back to ex.

$$S_\lambda = \frac{N}{\lambda} - \sum_n X_n.$$

$$E[S_\lambda] = \frac{N}{\lambda} - \sum_n E[X_n] = \frac{N}{\lambda} - \sum_n \frac{1}{\lambda} = \frac{N}{\lambda} - \frac{N}{\lambda} = 0$$

My MLE is $\hat{\lambda} = 1/\bar{X}$

$$\text{and } E\hat{\lambda} = E[1/\bar{X}] \neq \lambda$$

← MLE is not unbiased for λ

Thm: $E[S_\theta] = 0$.

pf.

$$E[\underline{S_\theta}] = \int \underline{S_\theta}(\underline{x}) f(\underline{x}) d\underline{x}$$

← lazy notation
... $\int \dots \int \dots dx_1 dx_2 \dots$

$$Eg(x) = \int g(x) f(x) dx$$

$$= \int \frac{\partial \ell}{\partial \theta} f_{\theta}(x) dx$$

Aside: $\frac{\partial \ell}{\partial \theta} = \frac{\partial}{\partial \theta} \log f_{\theta}(x)$

$$= \frac{\frac{\partial f}{\partial \theta}}{f_{\theta}(x)}$$

$$= \int \frac{\frac{\partial f}{\partial \theta} f_{\theta}(x)}{f_{\theta}(x)} dx$$

$$= \int \frac{\partial f}{\partial \theta} dx$$

(*)

$$= \frac{\partial}{\partial \theta} \int f(x) dx$$

$$= \frac{\partial}{\partial \theta} (1)$$

$$= 0$$

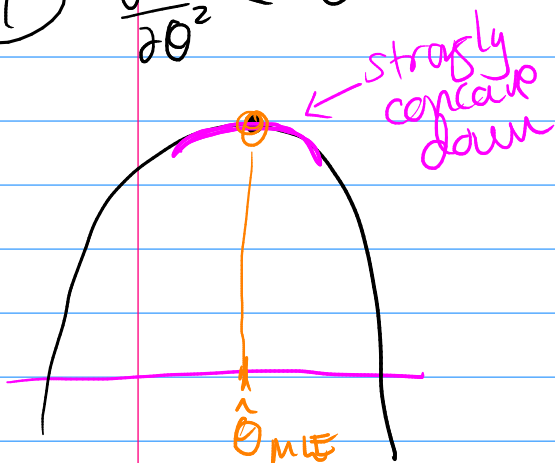
Need enough "regularity"
i.e. f_{θ} has to be nice
enough

This works for Exp. Fams.

What about $\frac{\partial^2 \ell}{\partial \theta^2}$?

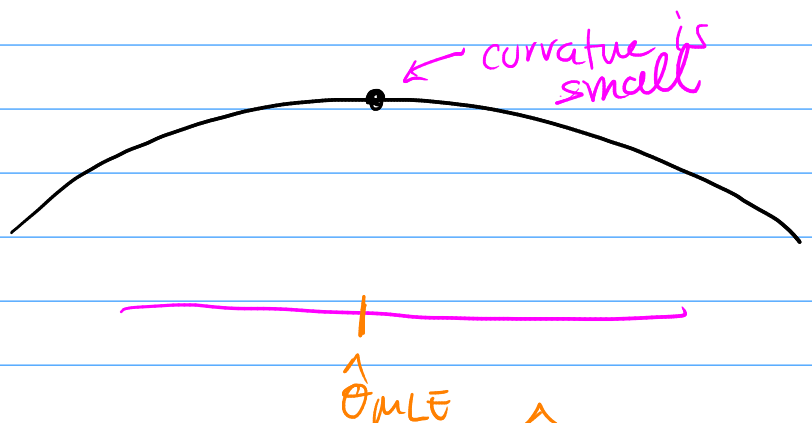
Two possibilities

(1) $\frac{\partial^2 \ell}{\partial \theta^2} \ll 0$



Strongly prefer $\hat{\theta}$ over other choices

(2) $\frac{\partial^2 \ell}{\partial \theta^2} < 0$ but close to zero



Weakly prefer $\hat{\theta}$ over other choices

Theorem: (*) also needs regularity of f_θ

$$\text{Var}(S_\theta) = \mathbb{E}[S_\theta^2] = -\mathbb{E}\left[\frac{\partial^2 \ell}{\partial \theta^2}\right]$$

b/c
 $\mathbb{E}S_\theta = 0$

viewing as
random
promote \mathcal{X} to \mathcal{X}_\sim

so if $S_\theta = \frac{\partial \ell}{\partial \theta}$ then

$$\mathbb{E}\left[\left(\frac{\partial \ell}{\partial \theta}\right)^2\right] = -\mathbb{E}\left[\frac{\partial^2 \ell}{\partial \theta^2}\right]$$

Defn: Fisher Information

We define the fisher info. for θ contained in
 \mathcal{X} ($N=1$) is defined as

$$I(\theta) = -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log f_\theta(\mathcal{X})\right]$$

$\frac{\partial \ell}{\partial \theta}$ w/ $N=1$

If I have N samples $\mathcal{X}_n \stackrel{\text{iid}}{\sim} f_\theta$ then the
Fisher info. contained about θ in \mathcal{X}_\sim is

$$I_N(\theta) = -\mathbb{E}\left[\frac{\partial^2 \ell}{\partial \theta^2}\right]$$

Theorem: $I_N(\theta) = N I(\theta)$

If I have N indep. samples - I have N times as much info.

pf- $I_N(\theta) = -\mathbb{E}\left[\frac{\partial^2 \ell}{\partial \theta^2}\right]$

$= -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log f(\mathbf{x})\right]$ ← joint

$= -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log\left(\prod_n f(x_n)\right)\right]$

$= -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \sum_n \log f(x_n)\right]$

$= \sum_n \underbrace{-\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log f(x_n)\right]}_{I(\theta)}$

$= N I(\theta)$

Ex. $X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$, $\lambda > 0$

Find $I_N(\lambda)$.

Let's find $I(\lambda)$ and multiply by N .

① Find $\log f_\lambda(x)$

$$f_\lambda(x) = \frac{\lambda^x e^{-\lambda}}{x!} \Rightarrow \log f_\lambda = x \log \lambda - \lambda - \log(x!)$$

② Take two derivs.

$$\rightarrow \frac{\partial}{\partial x} \log f_{\lambda} = \frac{x}{\lambda} - 1$$

$$\rightarrow \frac{\partial^2}{\partial x^2} \log f_{\lambda} = -\frac{x}{\lambda^2}$$

promote x to X to get $-\frac{X}{\lambda^2}$

need to calc

$$I(x) = -\mathbb{E}\left[\frac{\partial^2 \log f_{\lambda}}{\partial x^2}\right] = -\mathbb{E}\left[-\frac{X}{\lambda^2}\right] = \frac{1}{\lambda^2} \mathbb{E}X = \frac{\lambda}{\lambda^2} = \frac{1}{\lambda}$$

$$\text{So } I_N(x) = \frac{N}{\lambda}$$

$$\begin{aligned} \mathbb{E}\bar{X} &= \lambda \text{ (unbiased)} \\ \text{Var } \bar{X} &= \frac{\lambda}{N} \end{aligned}$$

Ex. Let $X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$ known

let's get $I_N(\mu)$.

$$\textcircled{1} f_{\mu}(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x-\mu)^2\right)$$

$$\textcircled{2} \log f_{\mu} = -\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2}(x-\mu)^2$$

$$\textcircled{3} \frac{\partial}{\partial \mu} \log f_{\mu} = -\frac{1}{2\sigma^2} (2)(x-\mu)(-1) = \frac{1}{\sigma^2}(x-\mu) = \frac{x}{\sigma^2} - \frac{\mu}{\sigma^2}$$

$$(4) \frac{\partial^2}{\partial \mu^2} \log f_{\mu} = -\frac{1}{\sigma^2}$$

$$(5) \text{Promote } x \text{ to } X \text{ ---- to get } -\frac{1}{\sigma^2}$$

$$(6) -\mathbb{E}[\dots] = -\mathbb{E}\left[-\frac{1}{\sigma^2}\right] = \frac{1}{\sigma^2}$$

$$I(\mu) = 1/\sigma^2$$

$$\text{So } \boxed{I_N(\mu) = N I(\mu) = N/\sigma^2}$$

$$\mathbb{E}\bar{X} = \mu \text{ (unbiased)}$$

$$\text{Var}\bar{X} = \sigma^2/N$$

$$\underline{\text{Ex.}} \quad X_n \stackrel{\text{iid}}{\sim} \text{Pois}(\lambda), \quad \mathbb{E}X_n = \text{Var}X_n = \lambda$$

$$\text{Sd}(X_n) = \sqrt{\lambda} = \psi$$

$$\text{i.e. } \psi^2 = \lambda$$

To get $I(\psi)$ I can re-parameterize Poiss in terms of ψ and apply my procedure

$$f_{\lambda}(x) = \frac{\lambda^x e^{-\lambda}}{x!} = \frac{(\psi^2)^x e^{-\psi^2}}{x!} = f_{\psi}(x)$$