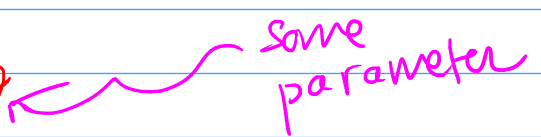


## Lecture 3: Exponential Families and Sufficiency

Probability: Given  $X_n \stackrel{iid}{\sim} f_\theta$   some parameter

Ex: If  $\theta = 5$  calculate  $P(X_n = \dots)$

Statistics: Given  $X_n \stackrel{iid}{\sim} f_\theta$

Ex: I don't know  $\theta$ , how can I estimate it?

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Generally we'll work w/ parameterized fams of dists:

e.g. (\*)  $N(\mu, \sigma^2)$  when  $\mu \in \mathbb{R}, \sigma^2 > 0$

(\*)  $\text{Exp}(\lambda)$  where  $\lambda > 0$

(\*)  $U(0, \theta)$  where  $\theta > 0$

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### Exponential Families

Assume we have a fam of dists parameterized by  $\theta \in \Theta \subset \mathbb{R}$

so that  $X_n \stackrel{iid}{\sim} f_\theta$

where

$$f_{\theta}(\underline{x}) = h(\underline{x}) c(\theta) \exp(T(\underline{x}) w(\theta))$$

Joint dist fns of  $\theta$  not  $\underline{x}$   
fns of  $\underline{x}$  not  $\theta$

then we say the  $X_n$ s are from an exponential family.

Ex. Poisson, Exp, Normal, Gamma, Beta, ....

Ex.  $X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$  for  $\lambda > 0$  |  $f(x) = \frac{1}{x!} \lambda^x e^{-\lambda} \mathbb{1}(x \in \mathbb{N}_0)$

$$f(\underline{x}) = \prod_{n=1}^N f(x_n) = \prod_{n=1}^N \frac{1}{x_n!} \lambda^{x_n} e^{-\lambda} \mathbb{1}(x_n \in \mathbb{N}_0)$$

$$= \prod_n \left( \frac{1}{x_n!} \right) \prod_n (\lambda^{x_n}) \prod_n e^{-\lambda} \prod_n \mathbb{1}(x_n \in \mathbb{N}_0)$$

$$= \prod_n \left( \frac{1}{x_n!} \right) \prod_n \mathbb{1}(x_n \in \mathbb{N}_0) \underbrace{\lambda^{\sum_n x_n}}_{\substack{\text{color: red} \\ e^{\log(\lambda^{\sum x_n})} = e^{(\sum x_n) \log \lambda}}} e^{-N\lambda}$$

$$a = e^{\log a}$$

$$\log(a^b) = b \log a$$

$$= \underbrace{\prod_n \left( \frac{1}{x_n!} \right) \prod_n \mathbb{1}(x_n \in \mathbb{N}_0)}_{h(\underline{x})} \underbrace{e^{-N\lambda}}_{c(\underline{x})} \exp \left( \underbrace{\left( \sum_n x_n \right)}_{T(\underline{x})} \underbrace{\log \lambda}_{w(\underline{x})} \right)$$

So  $\text{Pois}(\lambda)$  is an exp. fam.

Short-cut: can just check the marginal.

If

$$f_{\theta}(x) = h_{\theta}(x) c_{\theta}(\theta) \exp(T_{\theta}(x) \omega_{\theta}(\theta))$$

then

$$f_{\theta}(\underline{x}) = \prod_n f_{\theta}(x_n) = \prod_n h_{\theta}(x_n) c_{\theta}(\theta) \exp(T_{\theta}(x_n) \omega_{\theta}(\theta))$$

$$= \prod_n h_{\theta}(x_n) \prod_n c_{\theta}(\theta) \exp(\sum_n T_{\theta}(x_n) \omega_{\theta}(\theta))$$

$$\exp(\omega_{\theta}(\theta) \sum_n T_{\theta}(x_n))$$

$$f_{\theta}(\underline{x}) = \underbrace{\prod_n h_{\theta}(x_n)}_{h(\underline{x})} \underbrace{\prod_n c_{\theta}(\theta)}_{c(\theta)} \exp(\underbrace{(\sum_n T_{\theta}(x_n))}_{T(\underline{x})} \underbrace{\omega_{\theta}(\theta)}_{\omega(\theta)})$$

punchline:  $h(\underline{x}) = \prod_n h_{\theta}(x_n)$

$$c(\theta) = \prod_n c_{\theta}(\theta)$$

$$T(\underline{x}) = \sum_n T_{\theta}(x_n)$$

$$\omega(\theta) = \omega_{\theta}(\theta)$$

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Ex.  $X_n \stackrel{\text{iid}}{\sim} \text{Pois}(\lambda)$

$$f(x_n) = \frac{1}{x_n!} \lambda^{x_n} e^{-\lambda} \mathbb{1}(x_n \in \mathbb{N}_0)$$

$$= \underbrace{\frac{1}{X_n!}}_{h_0(X_n)} \underbrace{\mathbb{1}(X_n \in \mathbb{N}_0)}_{c_0(\lambda)} \underbrace{e^{-\lambda}}_{T_0(X_n)} \underbrace{\exp(X_n \log \lambda)}_{w(X)}$$

$$T(\underline{X}) = \sum_n T_0(X_n) = \sum_n X_n.$$

Ex.  $X_n \stackrel{iid}{\sim} U(0, \theta)$ ,  $\theta > 0$ .

Exp fam?

$$f_\theta(X_n) = \frac{1}{\theta} \text{ for } 0 < X_n < \theta$$

$$= \frac{1}{\theta} \mathbb{1}(0 < X_n < \theta)$$

no way to write as  
 $(f_n \text{ of } X) \cdot (f_n \theta)$

No, not an exp. fam.

Short-cut 2: general fact, if the support of the dist depends on the parameter, it's not an exp. family.

Defn: Sufficiency  $[X_n \stackrel{iid}{\sim} f_\theta, \theta \text{ unknown}]$

We say a statistic  $T = T(\underline{X})$  is sufficient for a parameter  $\theta$  if

$f_{\underline{X}|T=t}(\underline{X})$  is "free" of  $\theta$ .

free = formal for dist doesn't depend on  $\theta$ .

Ex. If I have data  $X_1, \dots, X_N$  is  $X_1 = T$  sufficient for  $\theta$ ?

No.

Ex. Is  $T = \underline{X}$  sufficient for  $\theta$ ?

Yes.

$$f(\underline{x} | T=t) = \frac{f_{\underline{x}, T}(\underline{x}, t)}{f_T(t)} = \frac{f_{\underline{x}, \underline{x}}(\underline{x}, \underline{x})}{f_{\underline{x}}(\underline{x})}$$

$$= \frac{f_{\underline{x}}(\underline{x})}{f_{\underline{x}}(\underline{x})} = 1$$

no  $\theta$ .

Theorem: Factorization Theorem

$T$  is sufficient for  $\theta$  iff

there is a fn  $g(\theta, T)$  and  $h(\underline{x})$  so that

joint  $\rightarrow f_{\theta}(\underline{x}) = g(\theta, T) h(\underline{x})$   $\leftarrow$  no  $\theta$   
fn of  $\theta$   
and data only through  $T$

Ex. let  $X_n \stackrel{iid}{\sim} \text{Bern}(\theta)$ ,  $\theta \in [0, 1]$

$$T = \frac{1}{N} \sum_{n=1}^N X_n = \bar{X}$$

Is  $T$  sufficient for  $\theta$ ?

$$f(x) = \begin{cases} \theta & x=1 \\ 1-\theta & x=0 \end{cases} \\ = \theta^x (1-\theta)^{1-x} \mathbb{1}_{(x=0,1)}$$

$$f_{\theta}(\underline{x}) = \prod_n \theta^{x_n} (1-\theta)^{1-x_n} \mathbb{1}_{(x_n=0 \text{ or } 1)}$$

$$\bar{X} = \frac{1}{N} \sum_n X_n$$

$$\sum_n X_n = N\bar{X}$$

$$a^{b+c} = a^b a^c$$

$$= \theta^{\sum_n X_n} (1-\theta)^{\sum_n (1-X_n)} \prod_n \mathbb{1}_{(x_n=0 \text{ or } 1)}$$

$$= \theta^{\sum_n X_n} (1-\theta)^{N - \sum_n X_n} \prod_n \mathbb{1}_{(x_n=0 \text{ or } 1)}$$

$$= \theta^{N\bar{X}} (1-\theta)^{N - N\bar{X}} \prod_n \mathbb{1}_{(x_n=0 \text{ or } 1)}$$

$$= \theta^{N\bar{X}} (1-\theta)^N (1-\theta)^{-N\bar{X}} \prod_n \mathbb{1}_{\dots}$$

$$= \frac{\theta^{N\bar{X}}}{(1-\theta)^{N\bar{X}}} (1-\theta)^N \prod_n \mathbb{1}_{\dots}$$

$$= \left( \frac{\theta}{1-\theta} \right)^{N\bar{X}} (1-\theta)^N \prod_n \mathbb{1}_{(x_n=0 \text{ or } 1)}$$

$$g(T, \theta)$$

$$h(\underline{X})$$

$$g(T, \theta) = \left( \frac{\theta}{1-\theta} \right)^{NT} (1-\theta)^N$$

So  $\bar{X}$  is sufficient for  $\theta$ .

Ex. let  $X_n \stackrel{iid}{\sim} U(0, \theta)$

Can I find a SS?

↑ one-dim'l

$$f_{\theta}(X_n) = \frac{1}{\theta} \mathbb{I}(0 < X_n < \theta)$$

$$= \frac{1}{\theta} \mathbb{I}(X_n > 0) \mathbb{I}(X_n < \theta)$$

$$f_{\theta}(\underline{x}) = \prod_n \frac{1}{\theta} \mathbb{I}(X_n > 0) \mathbb{I}(X_n < \theta)$$

$$= \theta^{-N} \prod_n \mathbb{I}(X_n < \theta) \prod_n \mathbb{I}(X_n > 0)$$

$$= \theta^{-N} \mathbb{I}(\text{all } X_n < \theta) \mathbb{I}(\text{all } X_n > 0)$$

$$\prod \mathbb{I}(A_n) = \mathbb{I}(\text{all } A_n)$$

$$= \theta^{-N} \mathbb{I}(X_{(n)} < \theta) \mathbb{I}(X_{(1)} > 0)$$

$$T = X_{(n)} \quad g(T, \theta)$$

$$h(\underline{x})$$

So  $X_{(n)}$  is sufficient for  $\theta$ .

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