

Lecture 7: Evaluation

Defn: Mean-Squared Error (MSE)

If $X_n \stackrel{iid}{\sim} f_\theta$ where $\theta \in \Theta$ and let $\hat{\theta}$ be an est. of θ .

Define the MSE of $\hat{\theta}$ estimating θ as

$$MSE_\theta(\hat{\theta}) = E[(\hat{\theta} - \theta)^2]$$

If $\hat{\theta}$ is good, MSE is small, vice-versa.

Idea: If I have two ests, $\hat{\theta}_1$ and $\hat{\theta}_2$ I might prefer the est. w/ a smaller MSE.

Defn: Bias The bias of $\hat{\theta}$ est. θ is

$$B_\theta(\hat{\theta}) = E[\hat{\theta} - \theta] = E[\hat{\theta}] - \theta$$

If $B(\hat{\theta}) > 0$ I tend to over-estimate
 < 0 " under-estimate

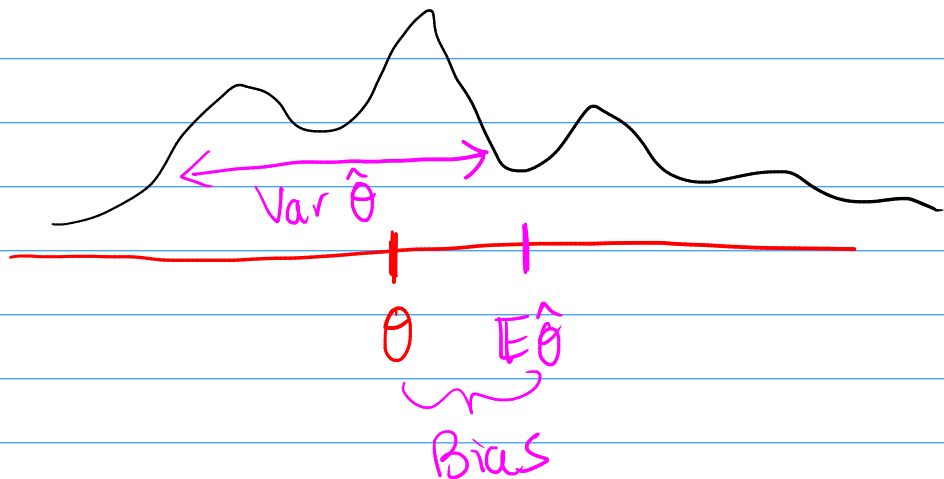
If $B(\hat{\theta}) = 0$ we say $\hat{\theta}$ is unbiased
i.e. gets pred. correct on avg.

Variance

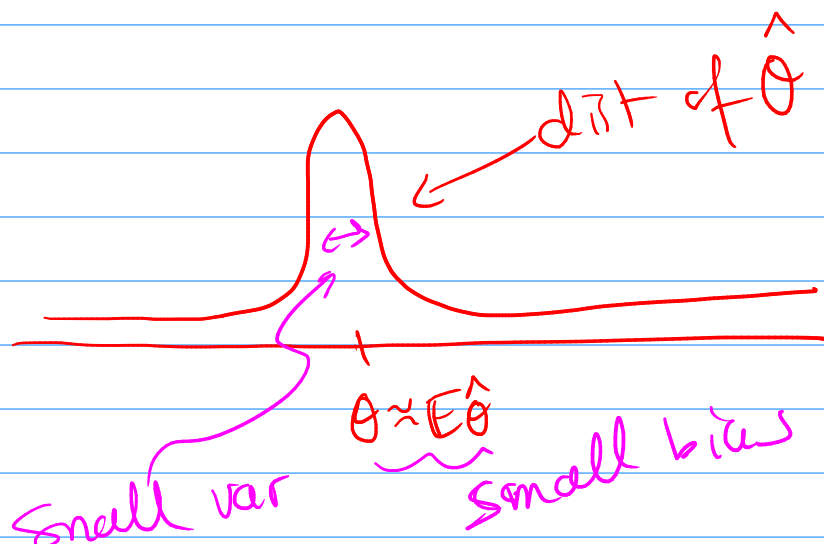
Recall that $\hat{\theta} = \hat{\theta}(X)$ so it is random.

So $\hat{\theta}$ has a variance: $\text{Var}_{\theta}(\hat{\theta})$

Ex.

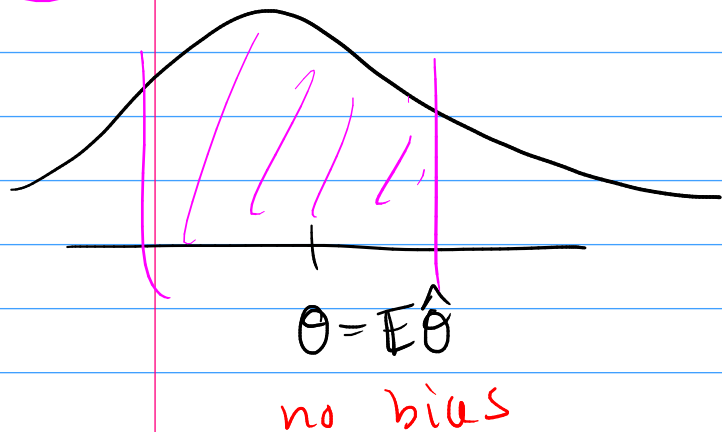


Ideally $B(\hat{\theta})$ is small and so is $\text{Var}(\hat{\theta})$

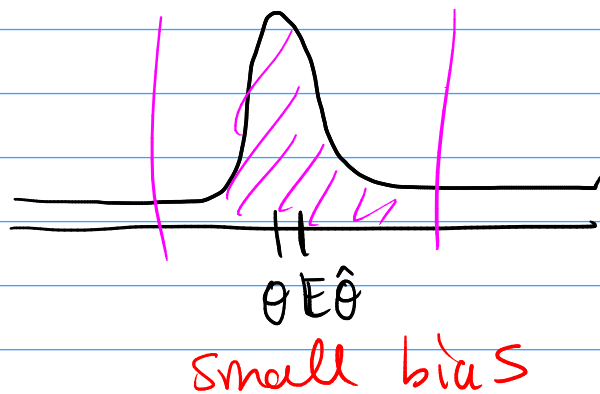


Sometimes, a (non-zero) biased est is better.

① high var



② low var



Theorem: $MSE = \text{bias}^2 + \text{Var}$

pf. $MSE(\hat{\theta}) = E[(\hat{\theta} - \theta)^2]$

$$= E[(\underbrace{\hat{\theta} - E\hat{\theta}}_a + \underbrace{E\hat{\theta} - \theta}_b)^2]$$

$$(a+b)^2 = a^2 + b^2 + 2ab$$

$$= E[a^2] + E[b^2] + 2E[ab]$$

$$= E[(\hat{\theta} - E\hat{\theta})^2] + E[(\theta - E\hat{\theta})^2]$$

$$+ 2E[(\hat{\theta} - E\hat{\theta})(E\hat{\theta} - \theta)]$$

$\text{Var}(z) = E[(z - E z)^2]$

①

③

$$(1) E[(\hat{\theta} - E\hat{\theta})^2] = \text{Var}(\hat{\theta})$$

$$(2) \cancel{E[(\theta - E\hat{\theta})^2]} = (E[\hat{\theta}] - \theta)^2 \\ = B(\hat{\theta})^2$$

$$(3) -E[(\theta - E\hat{\theta})(\hat{\theta} - E[\hat{\theta}])] \\ = (E[\hat{\theta}] - \theta) E(\hat{\theta} - E[\hat{\theta}]) \\ = (E[\hat{\theta}] - \theta) \underbrace{(E[\hat{\theta}] - E\hat{\theta})}_0 \\ = 0$$

$$E[aX] = aEX$$

Ex. $X_n \stackrel{iid}{\sim} f$ where $\mu = E(X_n)$
 $\sigma^2 = \text{Var}(X_n)$

We prev. showed generally if $\hat{\mu} = \bar{X}$
 then $E[\hat{\mu}] = \mu \rightsquigarrow$ unbiased
 $\text{Var}(\hat{\mu}) = \sigma^2/N$ $B(\hat{\mu}) = 0$

$$\text{So } \text{MSE}(\hat{\mu}) = B(\hat{\mu})^2 + \text{Var}(\hat{\mu}) \\ = 0^2 + \sigma^2/N = \sigma^2/N$$

Notice: If Bias = 0 then MSE = Var.

Ex, Consider $S_{N-1}^2 = \frac{1}{N-1} \sum_n (X_n - \bar{X})^2$.

Showed that $E[S_{N-1}^2] = \sigma^2$

and so S_{N-1}^2 is unbiased

Consider $X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ we also showed
that $\frac{N-1}{\sigma^2} S_{N-1}^2 \sim \chi^2(N-1)$

$$\text{So } \text{Var} \left[\frac{N-1}{\sigma^2} S_{N-1}^2 \right] = 2(N-1)$$

$$\text{So } \frac{(N-1)^2}{\sigma^4} \text{Var}(S_{N-1}^2) = 2(N-1)$$

$$\text{so } \boxed{\text{Var}(S_{N-1}^2) = \frac{2\sigma^4}{N-1}}$$

So since S_{N-1}^2 was unbiased for σ^2 then

$$\text{MSE} = \text{Var} = \frac{2\sigma^4}{N-1}$$

$$S_{N-1}^2 = \frac{1}{N-1} \sum_n (x_n - \bar{x})^2$$

The MLE of σ^2 in the $N(\mu, \sigma^2)$ case was

$$\hat{\sigma}^2 = \frac{1}{N} \sum_n (x_n - \bar{x})^2 = \frac{N-1}{N} S_{N-1}^2$$

Let's get MSE.

$$\mathbb{E}[\hat{\sigma}^2] = \mathbb{E}\left[\frac{N-1}{N} S_{N-1}^2\right] = \frac{N-1}{N} \mathbb{E}[S^2] = \frac{N-1}{N} \sigma^2$$

$$\begin{aligned} \text{So } B(\hat{\sigma}^2) &= \mathbb{E}[\hat{\sigma}^2] - \sigma^2 \\ &= \frac{N-1}{N} \sigma^2 - \sigma^2 \\ &= -\frac{1}{N} \sigma^2 \end{aligned}$$

$$\begin{aligned} \text{Var}(\hat{\sigma}^2) &= \text{Var}\left(\frac{N-1}{N} S_{N-1}^2\right) = \frac{(N-1)^2}{N^2} \text{Var}(S_{N-1}^2) \\ &= \frac{(N-1)^2}{N^2} \frac{2\sigma^4}{N-1} \\ &= \frac{2(N-1)\sigma^4}{N^2} \end{aligned}$$

$$\text{MSE}(\hat{\sigma}^2) = B^2 + \text{Var} = \left(-\frac{1}{N} \sigma^2\right)^2 + \frac{2(N-1)}{N^2} \sigma^4$$

$$= \frac{\sigma^4}{N^2} + \frac{2(N-1)\sigma^4}{N^2}$$

$$= \frac{(2N-1)\sigma^4}{N^2}$$

$$MSE(S_{N-1}^2) = \frac{2\sigma^4}{N-1}$$

Consider

$$MSE(\hat{\sigma}^2) = \frac{(2N-1)\sigma^4}{N^2} = \frac{(2N-1)}{N^2} \overbrace{\left(\frac{N-1}{2}\right) \left(\frac{2}{N-1}\right)}^1 \sigma^4$$

$MSE(S_{N-1}^2)$

$$\text{So } MSE(\hat{\sigma}^2) = \underbrace{\frac{(2N-1)(N-1)}{2N^2}}_{>1? <1?} MSE(S_{N-1}^2)$$

$$= \frac{2N^2 - 3N + 1}{2N^2} < \frac{2N^2}{2N^2} = 1$$

$$\text{So } MSE(\hat{\sigma}^2) < MSE(S_{N-1}^2).$$

More generally what do I multiply S^2 by to get the lowest MSE?

$$MSE(cS^2)$$

$$= B(cS^2)^2 + \text{Var}(cS^2)$$

$$= (E[cS^2] - \sigma^2)^2 + c^2 \text{Var}(S^2)$$

$$= (c E[S^2] - \sigma^2)^2 + c^2 \text{Var}(S^2)$$

$$= (c\sigma^2 - \sigma^2)^2 + c^2 \frac{2\sigma^4}{N-1}$$

$$= \sigma^4 (c-1)^2 + \frac{2c^2\sigma^4}{N-1}$$

take deriv wrt c set to zero, solve.

$$\frac{\partial}{\partial c} MSE = \cancel{\sigma^4} 2(c-1)(1) + \frac{4c\cancel{\sigma^4}}{N-1} = 0$$

$$\Rightarrow \cancel{2}(c-1)(N-1) + \frac{2}{1}c = 0$$

$$\Rightarrow cN - c - N + 1 + 2c = 0$$

$$\Rightarrow (N+1)c - (N-1) = 0$$

$$\Rightarrow \boxed{c^* = \frac{N-1}{N+1}}$$

So min MSE is at $c^*S^2 = \frac{N-1}{N+1} S^2 = \left[\frac{1}{N+1} \sum_n (X_n - \bar{X})^2 \right]$

I want to find a "best" estimator.

Problem: If I am too permissive in what I call an estimator, there is no all around best option.

Ex. $X_n \stackrel{iid}{\sim} N(\mu, 1)$

want to find some μ^* that is best so that

$$MSE_{\mu}(\mu^*) \leq MSE_{\mu}(\hat{\mu})$$

\forall possible $\hat{\mu}$
 $\forall \mu$

