

Lecture 20 Interval Estimation

Point estimation!

$$\hat{\theta} = \hat{\theta}(\underline{x}) \in \Theta \quad \text{idea } "\hat{\theta} \approx \theta"$$

Interval Estimation!

$$C = C(\underline{x}) \subset \Theta \quad \text{idea } "\theta \in C"$$

prefer if C is an interval.

Defn: Interval Estimator

An interval est. of $\theta \in \Theta \subset \mathbb{R}$ is a pair of fns

$$L = L(\underline{x}) \quad \text{and} \quad U = U(\underline{x})$$

that satisfy $L \leq U \quad \forall \underline{x}$

idea: want to say $"L \leq \theta \leq U"$

Ex let $X_n \stackrel{iid}{\sim} N(\mu, 1)$. let $N=4$.

Then an interval est. of μ is

$$\left[\underbrace{\bar{X} - 1}_L, \underbrace{\bar{X} + 1}_U \right]$$

might say " $\bar{x} - 1 \leq \mu \leq \bar{x} + 1$ ".

Why use an interval est? Just use \bar{x} ?

notice! $P(\bar{x} = \mu) = 0$

So we typically attach some uncertainty to \bar{x}
i.e. $sd(\bar{x}) = 1/\sqrt{4}$

Alt: use an interval est. b/c

$$P(\bar{x} - 1 \leq \mu \leq \bar{x} + 1) > 0$$

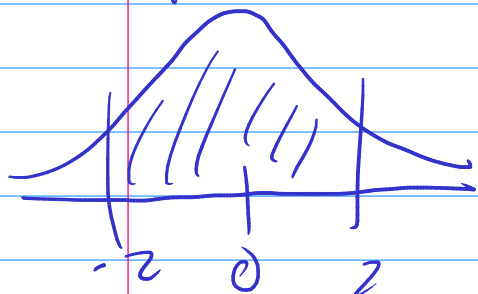
in this case

$$\hookrightarrow = P(\bar{x} - \mu \leq 1 \text{ and } \bar{x} - \mu \geq -1)$$

$$= P(-1 \leq \bar{x} - \mu \leq 1)$$

$$N(0,1) = P(-2 \leq \frac{\bar{x} - \mu}{\sqrt{1/4}} \leq 2) \approx 95\%$$

$N(0,1)$



Defn: Coverage Prob.

For int. est. $[L, U]$ of a param θ
the coverage prob is

$$P_{\theta}(L \leq \theta \leq U)$$

\uparrow depends on θ .

Defn: Confidence Coef.

Worst-case coverage prob.

$$1 - \alpha = \min_{\theta \in \Theta} P_{\theta}(L \leq \theta \leq U)$$

more generally if I have a set $C(\underline{x}) \subset \Theta$
its assoc. conf. coef is

$$1 - \alpha = \min_{\theta \in \Theta} P_{\theta}(\theta \in C(\underline{x}))$$

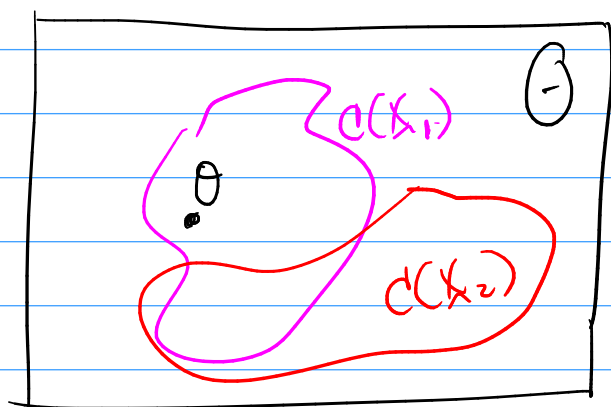
conf interval = interval est. + conf. coef

conf. set = set est. + conf. coef.

When we look at $P_{\theta}(\theta \in C(X))$

random

fixed but unknown



How do I build a conf set / interval?

Basically one way ... invert a hypothesis test

HT \longleftrightarrow Conf. set.

Ex. $X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$

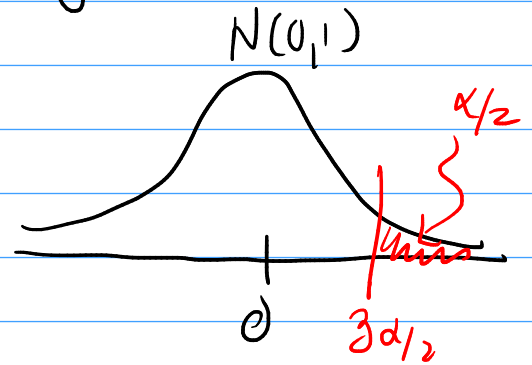
known

Let's build a HT for the hypothesis

$H_0: \mu = \mu_0$ v. $H_a: \mu \neq \mu_0$

A level α test for this hypothesis is to reject when

$$\left| \frac{\bar{X} - \mu_0}{\sigma/\sqrt{N}} \right| > z_{\alpha/2}$$



$$R(\mu_0) = \left\{ \underline{x} : \frac{|\bar{X} - \mu_0|}{\sigma/\sqrt{N}} > z_{\alpha/2} \right\}$$

↑ set of \underline{x} that don't agree w/ $H_0: \mu = \mu_0$

$$A(\mu_0) = \mathcal{X} \setminus R(\mu_0) = \left\{ \underline{x} : \frac{|\bar{X} - \mu_0|}{\sigma/\sqrt{N}} \leq z_{\alpha/2} \right\}$$

↑ set of \underline{x} that agree w/ H_0 .

"invert"
i.e. solve for
 μ_0 in the middle

$$-z_{\alpha/2} \leq \frac{\bar{X} - \mu_0}{\sigma/\sqrt{N}} \leq z_{\alpha/2}$$

$$\Leftrightarrow -z_{\alpha/2} \sigma/\sqrt{N} \leq \bar{X} - \mu_0 \leq \sigma/\sqrt{N} z_{\alpha/2}$$

$$\Leftrightarrow \underbrace{\bar{X} - \sigma/\sqrt{N} z_{\alpha/2}}_L \leq \mu_0 \leq \underbrace{\bar{X} + \sigma/\sqrt{N} z_{\alpha/2}}_U$$

Claim: $[L, U]$ is a $1-\alpha$ CI for μ .

why? Need to show $P_{\mu}(L \leq \mu \leq U) \geq 1-\alpha$

Test is level α so

$$\max_{\theta \in \Theta_0} P_{\theta}(\text{reject}) \leq \alpha$$

$$\Leftrightarrow P_{\mu_0}(\text{reject}) \leq \alpha$$

$$\Leftrightarrow P_{\mu_0}(\underline{X} \in R(\mu_0)) \leq \alpha$$

$$\Leftrightarrow P_{\mu_0}(\underline{X} \in A(\mu_0)) \geq 1-\alpha$$

$$\Leftrightarrow P_{\mu_0}(\bar{X} - \sigma/\sqrt{N} z_{\alpha/2} \leq \mu_0 \leq \bar{X} + \sigma/\sqrt{N} z_{\alpha/2}) \geq 1-\alpha$$

$$\Leftrightarrow P_{\mu_0}(L \leq \mu_0 \leq U) \geq 1-\alpha$$

\uparrow arbitrary

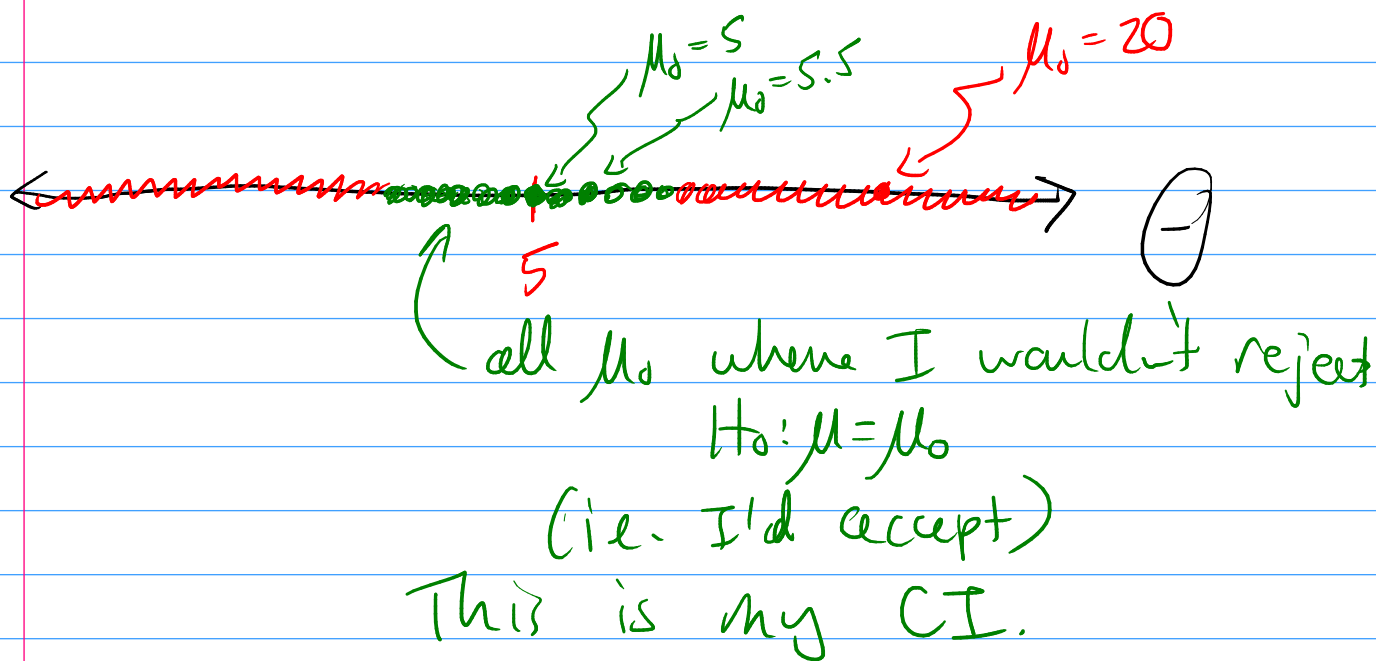
so

$$\forall \mu \quad P_{\mu}(L \leq \mu \leq U) \geq 1-\alpha$$

$$\min_{\mu} P_{\mu}(L \leq \mu \leq U) \geq 1-\alpha$$

so L, U is a $1-\alpha$ CI for μ .

$$H_0: \mu = \mu_0, \quad \bar{X} = 5$$



Test Inversion

For $\theta_0 \in \Theta$ let $A(\theta_0)$ be the acceptance region of an α -level test for

$$H_0: \theta = \theta_0 \quad \text{v.} \quad H_a: \theta \neq \theta_0$$

then if I let

$$C(\underline{X}) = \{ \theta : \underline{X} \in A(\theta) \}$$

this is a $1-\alpha$ confidence set for θ .

Two worlds!

HT: fix θ_0 and test $H_0: \theta = \theta_0$

by determining some rule

$$A(\theta_0) = \{ \text{set of } \underline{x} \text{ consistent w/ } H_0 \}$$

$\subset \mathcal{X}$

CI: Fix \underline{x} and want to determine which θ are consistent w/ \underline{x}
do this as

$$C(\underline{x}) = \{ \text{set of } \theta \text{ consistent w/ } \underline{x} \} \subset \Theta$$

Ex. Let $X_n \stackrel{\text{iid}}{\sim} \text{Exp}(\beta)$

$$E X_n = \beta$$

$$f(x) = \frac{1}{\beta} e^{-x/\beta} \quad \text{for } x \geq 0$$

Let's make a CI for β by inverting the LRT.

$$H_0: \beta = \beta_0 \quad \text{v.} \quad H_a: \beta \neq \beta_0$$

$$\lambda = \frac{L(\hat{\beta}_0)}{L(\hat{\beta})} = \frac{L(\beta_0)}{L(\bar{x})} = \frac{\frac{1}{\beta_0^N} \exp(-N\bar{x}/\beta_0)}{\frac{1}{\bar{x}^N} \exp(-N)}$$

$$= \left(\frac{\bar{X}}{\beta_0} \right)^N e^N e^{-N\bar{X}/\beta_0}$$

LRT rejects when $\lambda \leq C$

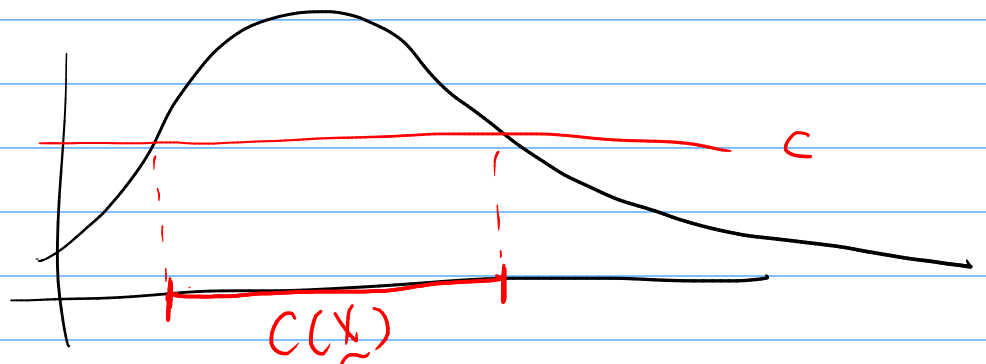
↑ set to make
some α level
test

$$R(\beta_0) = \left\{ \underline{x} : \left(\frac{\bar{X}}{\beta_0} \right)^N e^N e^{-N\bar{X}/\beta_0} \leq C \right\}$$

$$A(\beta_0) = \left\{ \underline{x} : \left(\frac{\bar{X}}{\beta_0} \right)^N e^N e^{-N\bar{X}/\beta_0} > C \right\}$$

$$C(\underline{x}) = \left\{ \beta : \underline{x} \in A(\beta) \right\}$$

$$= \left\{ \beta : \left(\frac{\bar{X}}{\beta} \right)^N e^N e^{-N\bar{X}/\beta} > C \right\}$$



Often trying to invert LRT is nasty.

What is a HT?

level α

A HT is simply a region $R \subset \mathcal{X}$ so that

$$\max_{\theta \in \Theta_0} P_{\theta}(\underline{X} \in R) \leq \alpha$$

So if I consider

$$H_0: \theta = \theta_0 \quad \text{v.} \quad H_a: \theta \neq \theta_0$$

then a level α HT is simply a region R
so that

$$P_{\theta_0}(X \in R) \leq \alpha$$

or equiv. a region A where

$$P_{\theta_0}(X \in A) \geq 1 - \alpha$$

same statement about \tilde{X}

So a level α HT is just a statement
about \tilde{X} so that

$$P_{\theta_0}(\text{statement}) \geq 1 - \alpha$$

↑ implicitly defines A

Generalize:

$$P_{\theta_0}(Q(\tilde{X}, \theta_0) \in A) \geq 1 - \alpha$$

"statement"

then this implicitly defines accept region of
a HT as

$$\{\underline{x} : Q(\underline{x}, \theta_0) \in A\}$$

So to build a conf. set I can invert this!

$$C(\underline{x}) = \{\theta : Q(\underline{x}, \theta) \in A\}.$$

This will work so long as

$$\min_{\theta} P_{\theta}(Q(\underline{x}, \theta) \in A) \geq 1 - \alpha$$

i.e. this statement has to have prob $\geq 1 - \alpha$ for all possible θ .

One way to ensure this is if

① $\text{dist } Q(\underline{x}, \theta)$ doesn't depend on θ

② A doesn't depend on θ

b/c then

$$P_{\theta}(Q(\underline{x}, \theta) \in A) \geq 1 - \alpha$$

↑
dist doesn't
depend on θ

↑
no θ