

Setup:  $\underline{X}_n \stackrel{iid}{\sim} f_\theta$  and  $\theta \in \Theta$ .

$T$  captures all the info about  $\theta$  in the sample  $\underline{x}$

Defn: A statistic  $T = T(\underline{X})$  is sufficient for  $\theta$  if

$f_{\underline{X} | T=t}(\underline{x})$  is "free" of  $\theta$

conditional on knowing  $T=t$

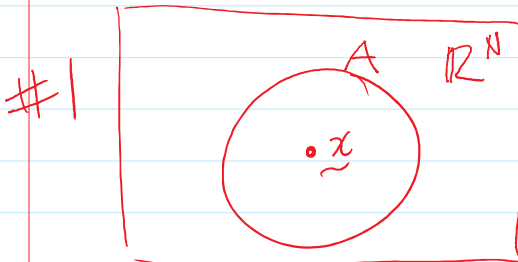
Ex.  $\underline{x} \in \mathbb{R}^N$ ;  $A \subset \mathbb{R}^N$

$$P(\underline{X} = \underline{x} \text{ and } \underline{x} \in A)$$

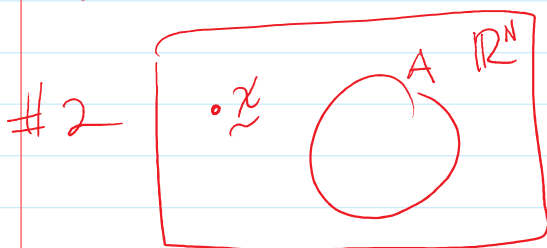
dist. formula doesn't depend on  $\theta$

Two options: statement about  $\underline{x}$

$$\underline{x} \in A$$



$$P(\underline{X} = \underline{x} \text{ and } \underline{x} \in A) = P(\underline{X} = \underline{x}) \quad \text{if } \underline{x} \in A$$



$$P(\underline{X} = \underline{x} \text{ and } \underline{x} \in A) = 0$$

intersection is empty if  $\underline{x} \notin A$

we could simply summarize as

$$P(\underline{X} = \underline{x}, \underline{x} \in A) = P(\underline{X} = \underline{x}) \mathbb{I}(\underline{x} \in A)$$

or

$$P(X = \underline{x}, \underbrace{T(X) = t}_{X \in A}) = P(X = \underline{x}) \mathbb{I}(T(X) = t)$$

Ex. let  $X_1, X_2, X_3 \stackrel{iid}{\sim} \text{Bernoulli}(\theta)$   $\theta \in [0, 1]$

let  $T = X_1 + X_2 + X_3 \sim \text{Bin}(3, \theta)$   
 IS  $T$  sufficient for  $\theta$ ?

$$f(\underline{x} | T=t) = \frac{f_{X,T}(\underline{x}, t)}{f_T(t)}$$

$$= \frac{P(X = \underline{x}, T=t)}{P(T=t)}$$

$$= \frac{P(X = \underline{x}) \mathbb{I}(T=t)}{P(T=t)}$$

$$= \frac{\theta^{\sum x_n} (1-\theta)^{3-\sum x_n} \mathbb{I}(T=t)}{\binom{3}{t} \theta^t (1-\theta)^{3-t}}$$

$$= \frac{1}{\binom{3}{t}} \mathbb{I}(\sum x_n = t)$$

no  $\theta$ !

$$f_{X|Y=y} = \frac{f(x,y)}{f(y)}$$

$$\begin{aligned} f(\underline{x}) &= P(X = \underline{x}) \\ &= \prod_{n=1}^3 \theta^{x_n} (1-\theta)^{1-x_n} \\ &= \theta^{\sum x_n} (1-\theta)^{3-\sum x_n} \end{aligned}$$

$$T \sim \text{Bin}(3, \theta)$$

$$\begin{aligned} f(t) &= P(T=t) \\ &= \binom{3}{t} \theta^t (1-\theta)^{3-t} \end{aligned}$$

$$t = \sum x_n$$

So  $T$  is sufficient for  $\theta$ .

Ex. let  $X_n \stackrel{iid}{\sim} f_\theta$

$$T = (X_{(1)}, X_{(2)}, X_{(3)}, \dots, X_{(N)})$$

$\swarrow$   $N$ -dimensional statistic

recall

$$f_T(t) = n! \prod_{n=1}^N f(t_n) \quad t = (t_1, \dots, t_N)$$

Q: is  $T$  sufficient for  $\theta$ ?

$$f_{X|T=t}(x) = \frac{f_{X,T}(x, t)}{f_T(t)}$$

$\swarrow$  only need to consider where  $T(X) = t$

$$= \frac{f_X(x)}{f_T(t)} = \frac{\prod_{n=1}^N f(x_n)}{n! \prod_{n=1}^N f(t_n)} = \frac{1}{n!}$$

recall:  $t_n$  are a re-ordering of  $x_n$

So since this is free of  $\theta$ ,  $T$  is sufficient.

Ex.  $X_n \stackrel{iid}{\sim} N(\mu, 1)$

let  $T = \bar{X}$ . We know  $\bar{X} \sim N(\mu, 1/N)$ .

Q: Is  $X$  sufficient for  $\mu$ ?

need to consider  $\bar{X} = t$

$$f(X|T=t) = \dots = \frac{f_X(X)}{f_T(t)}$$

$$T = \bar{X} \sim N(\mu, 1/N)$$

$$\frac{\prod_{n=1}^N \frac{1}{\sqrt{2\pi}} \exp(-\frac{1}{2}(x_n - \mu)^2)}{\frac{1}{\sqrt{2\pi/N}} \exp(-\frac{N}{2}(\bar{X} - \mu)^2)}$$

$$\prod_n e^{a_n} = e^{\sum a_n}$$

$$= (2\pi)^{-N/2} \exp\left(-\frac{1}{2} \sum_{n=1}^N (x_n - \mu)^2\right)$$

$$(2\pi/N)^{-1/2} \exp\left(-\frac{N}{2}(\bar{X} - \mu)^2\right)$$

$$\rightarrow \sum_n (x_n - \mu)^2 = \sum_n (x_n^2 - 2\mu x_n + \mu^2)$$

$$\bar{X} = \frac{1}{N} \sum x_n$$

$$= \sum x_n^2 - 2\mu \underbrace{\sum x_n}_{N\bar{X}} + N\mu^2$$

$$\rightarrow N(\bar{X} - \mu)^2 = N\bar{X}^2 - 2\mu N\bar{X} + N\mu^2$$

$$\propto \frac{\exp(-\frac{1}{2}(\sum x_n^2 - 2\mu N\bar{X} + N\mu^2))}{\exp(-\frac{1}{2}(N\bar{X}^2 - 2\mu N\bar{X} + N\mu^2))}$$

$$\exp(-\frac{1}{2}(N\bar{X}^2 - 2\mu N\bar{X} + N\mu^2))$$

↑ doesn't depend on  $\mu$ .

So  $T = \bar{X}$  is sufficient for  $\mu$ .

so  $I = X$  is sufficient for  $\mu$ .

### Factorization Theorem:

$T$  is sufficient for  $\theta$

iff

there is a function  $g(\theta, T)$  and  $h(X)$

so that

$$f(X) = g(\theta, T) h(X)$$

and  $\underset{\sim}{X}$  depends on  $\theta$  only through  $T = T(X)$

depends on  $\underset{\sim}{X}$  not  $\theta$

Ex.  $X_n \stackrel{\text{iid}}{\sim} N(\mu, 1)$

Show  $\bar{X}$  sufficient for  $\mu$ .

$$\rightarrow f(\underline{X}) = \prod_{n=1}^N \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}(\chi_n - \mu)^2\right)$$

$$= (2\pi)^{-N/2} \exp\left(-\frac{1}{2} \sum_n (\chi_n - \mu)^2\right)$$

$$= (2\pi)^{-N/2} \exp\left(-\frac{1}{2} \left(\sum_n \chi_n^2 - 2\mu N\bar{X} + N\mu^2\right)\right)$$

$$= \underbrace{(2\pi)^{-N/2} \exp\left(-\frac{1}{2} \sum_n \chi_n^2\right)}_{h(X)} \underbrace{\exp\left(-\frac{1}{2} (-2\mu N\bar{X} + N\mu^2)\right)}_{g(\mu, \bar{X})}$$

$$h(\underline{x}) = \exp\left(-\frac{1}{2}\sum_{i=1}^n (x_i - \mu)^2\right)$$

$$g(\mu, \underline{x}) = \exp\left(-\frac{1}{2}(-2\mu \sum_{i=1}^n x_i + N\mu^2)\right)$$

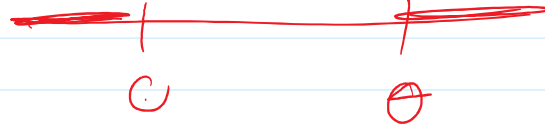
$$g(\mu, t) = \exp\left(-\frac{1}{2}(-2\mu Nt + N\mu^2)\right)$$

So since I can factor in this way,  
 $\bar{X}$  is sufficient for  $\theta$ .

Ex. let  $X_n \stackrel{iid}{\sim} U(0, \theta)$

$\frac{1}{\theta}$

Can I find a  
 sufficient stat  
 for  $\theta$ ?



$$f(\underline{x}) = \prod_{n=1}^N \frac{1}{\theta} \mathbb{1}(0 < x_n < \theta)$$

$$\mathbb{1}(A)\mathbb{1}(B) = \mathbb{1}(A \text{ and } B)$$

$$= \theta^{-N} \prod_{n=1}^N \mathbb{1}(0 < x_n < \theta)$$

all  $x_n$  between 0 and  $\theta$

$$= \theta^{-N} \mathbb{1}(x_{(1)} > 0) \mathbb{1}(x_{(n)} < \theta)$$

all have to be true  $\min > 0$   
 $\max < \theta$

$$= g(\theta, T) h(\underline{x})$$

$h(\underline{x}) = \mathbb{1}(x_{(1)} > 0)$

$$g(\theta, T) = \theta^{-N} \mathbb{1}(x_{(n)} < \theta)$$

$$g(\theta, T) = \theta^T \mathbb{1}(\underbrace{X_{(n)}}_T < \theta)$$

$$= \theta^{-N} \mathbb{1}(T < \theta) \text{ when } T = X_{(n)}$$

So  $T = X_{(n)}$  is sufficient for  $\theta$ .

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Ex.  $X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$ .  $\lambda > 0$

$$f(x) = \frac{\lambda^x e^{-\lambda}}{x!}$$

Sufficient stat for  $\lambda$ .

$$f(\underline{x}) = \prod_n f(x_n) = \prod_n \frac{\lambda^{x_n} e^{-\lambda}}{x_n!} = \left( \prod_n \frac{1}{x_n!} \right) \lambda^{\sum x_n} e^{-N\lambda}$$

$$= h(\underline{x}) g(\lambda, T)$$

$$h(\underline{x}) = \left( \prod_n \frac{1}{x_n!} \right); \quad g(\lambda, t) = \lambda^t e^{-N\lambda}$$

$$T = \frac{\sum x_n}{N} = \bar{X}$$

So  $\bar{X}$  is sufficient for  $\lambda$ .

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$$\text{Notice } E[X_i] = \lambda$$

$$\text{and so } E[\bar{X}] = \lambda$$

$$E[\bar{X}] = E\left[\frac{1}{N} \sum_n X_n\right] = \frac{1}{N} \sum_n E[X_n] = \frac{1}{N} \sum_n \lambda = \lambda$$


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Theorem: Exponential Families and Sufficiency

## Theorem: Exponential Families and Sufficiency

If we have  $X_n \stackrel{iid}{\sim} f_\theta$  and  $f_\theta$  is an exponential family, then

$$f_\theta(\underline{x}) = h(\underline{x}) c(\theta) \exp(T(\underline{x}) w(\theta)).$$

then  $T$  is sufficient for  $\theta$ .

pf. let  $g(\theta, t) = c(\theta) \exp(t w(\theta))$

then

$$f_\theta(\underline{x}) = h(\underline{x}) g(\theta, T)$$

so  $T$  is sufficient for  $\theta$ .

Ex.  $X_n \stackrel{iid}{\sim} \text{Exp}(\lambda)$   $f(x) = \lambda e^{-\lambda x} \mathbb{I}(x > 0)$

then we saw

$$\begin{aligned} f_\theta(\underline{x}) &= \prod_n \lambda e^{-\lambda x_n} \mathbb{I}(x_n > 0) \\ &= \lambda^N e^{-\lambda \sum_n x_n} \prod_n \mathbb{I}(x_n > 0) \end{aligned}$$