

Goal: Find a UMVUE. (Best estimator)

Deterministic: $\underline{x} \in \mathbb{R}^N$ and $\theta \in \mathbb{R}$

<u>Fns of \underline{x}</u>	$f_{\theta}(\underline{x})$	$\log f_{\theta}(\underline{x})$	$\frac{\partial}{\partial \theta} \log f_{\theta}(\underline{x})$ or $\frac{\partial^2}{\partial \theta^2} \log f_{\theta}(\underline{x})$
<u>Fn of θ</u>	$L(\theta) = f_{\theta}(\underline{x})$	$\ell(\theta) = \log L(\theta)$	$\frac{\partial \ell}{\partial \theta}$ or $\frac{\partial^2 \ell}{\partial \theta^2}$

Random: Replace \underline{x} w/ \underline{X}

$f_{\theta}(\underline{X})$ or $\log f_{\theta}(\underline{X})$ or $\frac{\partial}{\partial \theta} \log f_{\theta}(\underline{X})$ or $\frac{\partial^2}{\partial \theta^2} \log f_{\theta}(\underline{X})$

Score: $S_{\theta} = S_{\theta}(\underline{X}) = \frac{\partial}{\partial \theta} \log f_{\theta}(\underline{X}) = \frac{\partial \ell}{\partial \theta}$

Theorem: If we have nice enough f_{θ}
(Exponential families work)

then

$$\textcircled{1} \quad \mathbb{E}[S_{\theta}] = 0 \quad \left(\mathbb{E}\left[\frac{\partial}{\partial \theta} \log f_{\theta}(\underline{X})\right] = 0 \right)$$

$$\left(\mathbb{E}\left[\frac{\partial \ell}{\partial \theta}\right] = 0 \right)$$

$$\textcircled{2} \quad \text{Var}(S_{\theta}) = \mathbb{E}[S_{\theta}^2] = -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log f_{\theta}(\underline{X})\right]$$

$$\hookrightarrow \text{Var}(S_\theta) = \mathbb{E}[S_\theta^2] = -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log f_\theta(X)\right]$$

If $\mathbb{E}[Z] = 0$
 $\text{Var}(Z) = \mathbb{E}[Z^2]$

$$\left(\mathbb{E}\left[\left(\frac{\partial \ell}{\partial \theta}\right)^2\right] = -\mathbb{E}\left[\frac{\partial^2 \ell}{\partial \theta^2}\right] \right)$$

↑ my preference.

Defn: Fisher Info.

$N=1$ $X \sim f_\theta$ then $I(\theta) = -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log f_\theta(X)\right]$

$N > 1$ $X_n \stackrel{\text{iid}}{\sim} f_\theta$ then $I_N(\theta) = -\mathbb{E}\left[\frac{\partial^2}{\partial \theta^2} \log f_\theta(X)\right]$

$$\boxed{I_N(\theta) = -\mathbb{E}\left[\frac{\partial^2 \ell}{\partial \theta^2}\right] = NI(\theta)}$$

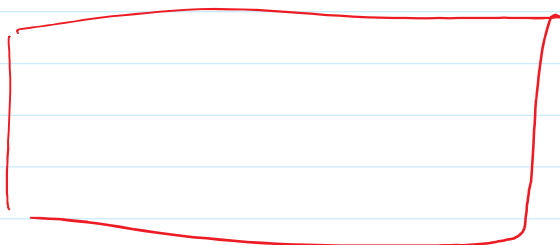
Ex. $X_n \stackrel{\text{iid}}{\sim} \text{Pois}(\lambda)$
 what is $I_N(\lambda)$?

① $I_N(\lambda) = NI(\lambda)$ so I only need to derive for $N=1$ sample

② $X \sim \text{Pois}(\lambda)$

③ $f_\lambda(x) = \frac{\lambda^x e^{-\lambda}}{x!}$

④ $\log f_\lambda(x) = x \log \lambda - \lambda - \log(x!)$



$$= \frac{1}{\lambda}$$

Claim 1 (Redo) $I_N(\lambda) = N I(\lambda)$

by defn $I_N(\lambda) = -\mathbb{E}\left[\frac{\partial^2}{\partial \lambda^2} \log f_\lambda(\underline{x})\right]$

$$\Rightarrow f_\lambda(\underline{x}) = \prod_n \frac{e^{-\lambda} \lambda^{x_n}}{(x_n)!} = \frac{e^{-N\lambda} \lambda^{\sum x_n}}{\prod_n x_n!}$$

$$\Rightarrow \log f_\lambda(\underline{x}) = -N\lambda + \sum x_n \log \lambda - \log(\prod x_n!)$$

$$\Rightarrow \frac{\partial}{\partial \lambda} \log f_{\lambda}(x) = -N + \frac{\sum_n x_n}{\lambda}$$

$$\Rightarrow \frac{\partial^2}{\partial \lambda^2} \log f_{\lambda}(x) = - \frac{\sum x_n}{\lambda^2}$$

$$\begin{aligned} \Rightarrow -E\left[\frac{\partial^2 \ell}{\partial \lambda^2}\right] &= -E\left[\frac{\partial^2}{\partial \lambda^2} \log f_{\lambda}(x)\right] = -E\left[-\frac{\sum x_n}{\lambda^2}\right] \\ &= \frac{\sum E[x_n]}{\lambda^2} \\ &= \frac{N \lambda}{\lambda^2} = \frac{N}{\lambda} \end{aligned}$$

Claim 2: $E\left[\left(\frac{\partial \ell}{\partial \theta}\right)^2\right] = -E\left[\frac{\partial^2 \ell}{\partial \theta^2}\right]$

$X \sim \text{Pois}(\lambda)$

$$\Rightarrow \frac{\partial \ell}{\partial \lambda} \frac{\partial}{\partial \lambda} \log f_{\lambda}(x) = \frac{x}{\lambda} - 1$$

$$\begin{aligned} \Rightarrow E\left[\left(\frac{\partial \ell}{\partial \lambda}\right)^2\right] &= E\left[\left(\frac{x}{\lambda} - 1\right)^2\right] \\ &= E\left[\frac{x^2}{\lambda^2} - \frac{2x}{\lambda} + 1\right] \\ &= \frac{E[x^2]}{\lambda^2} - \frac{2E[x]}{\lambda} + 1 \\ &= \frac{\lambda + \lambda^2}{\lambda^2} - \frac{2\lambda}{\lambda} + 1 \end{aligned}$$

$\text{Var}(X) = \lambda$
 $= E[x^2] - E[x]^2$
 $\text{so } E[x^2] = \lambda + \lambda^2$

$$\frac{1}{\lambda^2} - \frac{1}{\lambda} + 1 = \frac{1}{\lambda} + 1 - 2 + 1 = \frac{1}{\lambda} = I(\lambda)$$

Can I do this w/ $I_N(\lambda)$?
When working w/ $N > 1$

$$\frac{\partial \ell}{\partial \lambda} = -N + \frac{\sum X_n}{\lambda}$$

$$\begin{aligned} \text{So calculate } E\left[\left(\frac{\partial \ell}{\partial \lambda}\right)^2\right] &= E\left[\left(-N + \frac{\sum X_n}{\lambda}\right)^2\right] \\ &= E\left[\left(\frac{\sum X_n}{\lambda}\right)^2\right] - \frac{2N}{\lambda} E[\sum X_n] + N^2 \end{aligned}$$

ex. let $X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$
 σ^2 known
 μ unknown

What is $I_N(\mu)$?

$$(1) f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$$

$$(2) \ell = \log f(x) = -\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2}(x-\mu)^2$$

$$(3) \frac{\partial \ell}{\partial \mu} = \frac{1}{\sigma^2}(x-\mu)$$

$$\frac{\partial \mu}{\partial \sigma^2} = -\frac{1}{2\sigma^2} (x - \mu) (-1)$$

$$= \frac{1}{\sigma^2} (x - \mu)$$

$$(4) \frac{\partial^2 \ell}{\partial \mu^2} = -\frac{1}{\sigma^2}$$

$$(5) -E\left[\frac{\partial^2 \ell}{\partial \mu^2}\right] = 1/\sigma^2 = I(\mu).$$

$$(6) \boxed{I_N(\mu) = N/\sigma^2}$$

$$\left[\begin{array}{l} \text{Var}(\bar{x}) = \frac{\sigma^2}{N} \\ \text{suspicious ???} \end{array} \right]$$

Ex. $X_n \stackrel{\text{iid}}{\sim} \text{Pois}(\lambda)$

Recall $\text{Var}(X_n) = \lambda$ so $\text{sd}(X_n) = \sqrt{\lambda} = \psi$

Reparameterize poisson in terms of $\psi = \sqrt{\lambda}$ ($\lambda = \psi^2$)

$$f_{\lambda}(x) = \frac{\lambda^x e^{-\lambda}}{x!} = \frac{(\psi^2)^x e^{-\psi^2}}{x!} = \frac{\psi^{2x} e^{-\psi^2}}{x!}$$

$$= f_{\psi}(x)$$

Use prev. procedure:

$$(1) \ell = \log f_{\psi} = 2x \log \psi - \psi^2 - \log(x!)$$

$$(2) \frac{\partial \ell}{\partial \psi} = \frac{2x}{\psi} - 2\psi$$

$$(3) \frac{\partial^2 \mathcal{L}}{\partial \psi^2} = -\frac{2X}{\psi^2} - 2$$

$$(4) I(\psi) = -E\left[\frac{\partial^2 \mathcal{L}}{\partial \psi^2}\right] = -E\left[-\frac{2X}{\psi^2} - 2\right] \\ = \frac{2E[X]}{\psi^2} + 2 \\ = \frac{2\psi^2}{\psi^2} + 2 = 2 + 2 = 4.$$

$$(5) I_N(\psi) = 4N.$$

Derivative Review

$$y = f(x) \Leftrightarrow x = f^{-1}(y)$$

$$\frac{dy}{dx} \xleftrightarrow{\text{rel?}} \frac{dx}{dy}$$

Recall: $\frac{dy}{dx} = \frac{1}{\frac{dx}{dy}}$ or $\frac{dx}{dy} = \frac{1}{\frac{dy}{dx}} \leftarrow$

Theorem:

If $\theta = \tau(\psi)$ then

$$I(\theta) = \left(\frac{\partial \psi}{\partial \theta}\right)^2 I(\psi)$$

Transformation
Theorem for
Fisher Info.

A.H. $\left[\frac{\partial \psi}{\partial \theta} \right]^2$

Alt.

$$I(\psi) = \left(\frac{\partial \theta}{\partial \psi} \right)^2 I(\theta)$$

1/c If $I(\theta) = \left(\frac{\partial \psi}{\partial \theta} \right)^2 I(\psi)$ then $I(\psi) = \left(\frac{1}{\frac{\partial \psi}{\partial \theta}} \right)^2 I(\theta)$
 $= \left(\frac{\partial \theta}{\partial \psi} \right)^2 I(\theta)$

Revisit above example

$X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$ then $I(\lambda) = 1/\lambda$

then if $\psi = \sqrt{\lambda}$ then

$$I(\psi) = \left(\frac{d\lambda}{d\psi} \right)^2 I(\lambda)$$

$$= (2\psi)^2 \frac{1}{\lambda}$$

$$= (2\psi)^2 \frac{1}{\psi^2} = 4.$$

$$\psi = \sqrt{\lambda}$$
$$\lambda = \psi^2$$

$$\text{so } \frac{d\lambda}{d\psi} = 2\psi$$

Why we care?

Recall: $\hat{\theta}^*$ is the UMVUE for $\tau(\theta)$

$$(1) E[\hat{\theta}^*] = \tau(\theta)$$

$$(2) \text{Var}(\hat{\theta}^*) \leq \text{Var}(\hat{\theta}) \quad \forall \theta$$

$$(2) \text{Var}(\hat{\theta}^*) \leq \text{Var}(\hat{\theta}) \quad \forall \hat{\theta}$$

↑ any other unbiased est.

Theorem: Cramer - Rao Lower Bound

If $X_n \sim f_\theta$ when $\theta \in \Theta$ and $\hat{\theta}$ is unbiased for $T(\theta)$

(*) as long as f_θ is regular enough (i.e. exponential families) (*)

then

$$\text{Var}(\hat{\theta}) \geq \frac{\left(\frac{\partial T}{\partial \theta}\right)^2}{I_N(\theta)}.$$
