

Review of Iterated ExpectationTheorem: Iterated Expectation

$$E[X] = E[E[X|Y]]$$

\uparrow wrt X \uparrow wrt Y \uparrow wrt $X|Y$
 $f(x)$ $f(y)$ $f(x|y)$

$$E_X[X] = \int x f(x) dx$$

$$E_Y[\sim] = \int \sim f(y) dy$$

$$E_{X|Y}[\sim] = \int \sim f(x|y) dx$$

pf. $f(x|y) = \frac{f(x,y)}{f(y)}$ and $f(y|x) = \frac{f(x,y)}{f(x)}$

(cts) so $f(x,y) = \underline{f(x|y)} f(y) = \underline{f(y|x)} f(x)$

$$\begin{aligned}
 E[X] &= \int x \underline{f(x)} dx = \int x \underbrace{f(x,y)}_{f(x)} dy dx \\
 &= \iint x f(x,y) dy dx \\
 &= \iint x \underline{f(x|y)} f(y) dx dy \\
 &= \int \left[\int x \underline{f(x|y)} dx \right] f(y) dy \\
 &= \int \underbrace{E[X|Y=y]}_{g(y)} \underline{f(y)} dy \\
 &= E[E[X|Y]]
 \end{aligned}$$

$$E[X|Y=y] = g(y)$$

$$E[X|Y] = g(Y)$$

$$= E \left[\underbrace{E[X|Y]}_{g(Y)} \right]$$

Theorem: Law of Total Variance

$$\text{Var}(X) = E \left[\underbrace{\text{Var}(X|Y)}_{\text{r.v.}} \right] + \text{Var} \left(\underbrace{E[X|Y]}_{\text{r.v.}} \right)$$

Ex. $X|Y=y \sim \text{Bin}(y, p)$, $p \in [0, 1]$
 $Y \sim \text{Pois}(\lambda)$, $\lambda \geq 0$

$E[X]$?

$$E[X] = E[E[X|Y]]$$

① $E[X|Y=y] = yp$
 $\text{Bin}(yp)$

② $E[X|Y] = Yp$

③ $E[E[X|Y]] = E[Yp] = p E[Y]$
 $= p\lambda$

$$\text{Var}(X) = \text{Var}(E[X|Y]) + E[\text{Var}(X|Y)]$$

$$\text{Var}(X) = \text{Var}(\mathbb{E}[X|Y]) + \mathbb{E}[\text{Var}(X|Y)]$$

$$\textcircled{1} \mathbb{E}[X|Y] = Y/p$$

$$\begin{aligned} \textcircled{2} \text{Var}(\mathbb{E}[X|Y]) &= \text{Var}(Y/p) \\ &= p^2 \text{Var}(Y) \\ &= p^2 \lambda \end{aligned}$$

$$\textcircled{1} \text{Var}(X|Y=y) = yp(1-p)$$

$$\textcircled{2} \text{Var}(X|Y) = Yp(1-p)$$

$$\begin{aligned} \textcircled{3} \mathbb{E}[Y] &= \mathbb{E}[Yp(1-p)] \\ &= p(1-p) \mathbb{E}[Y] \\ &= p(1-p) \lambda \end{aligned}$$

$$\text{Var}(X) = p^2 \lambda + p(1-p) \lambda = p \lambda$$

Back to Math Stats.

CRLB doesn't always work (non-exp fams).

We want a more general method for finding the UMVUE.

Fact: Let $\hat{\theta}$ be unbiased for $\tau(\theta)$

so that $\mathbb{E} \hat{\theta} = \tau(\theta)$

and let $W = W(X)$ be some function of \underline{X}

(not nec. a statistic
w could depend on θ)

and let

$$\underline{\varphi} = \varphi(w) = \mathbb{E}[\hat{\theta}|w].$$

depends on w

depends on X

notice

$$\mathbb{E}_w[\varphi] = \mathbb{E}_w[\mathbb{E}[\hat{\theta}|w]] = \mathbb{E}[\hat{\theta}] = \tau(\theta)$$

If φ is a statistic, then it is an unbiased est. of $\tau(\theta)$.

Fact:

$$\text{Var}(\varphi) = \text{Var}(\mathbb{E}[\hat{\theta}|w])$$

Law of total variance

$$\text{Var}(\hat{\theta}) = \text{Var} \mathbb{E}[\hat{\theta}|w] + \mathbb{E} \text{Var}(\hat{\theta}|w)$$

rearrange

$$\begin{aligned} \underline{\text{Var}(\varphi)} &= \text{Var} \mathbb{E}[\hat{\theta}|w] = \text{Var}(\hat{\theta}) - \mathbb{E} \text{Var}(\hat{\theta}|w) \\ &= \text{Var}(\hat{\theta}) - \text{something} \end{aligned}$$

positive

proof line:

$$\boxed{\text{Var}(\varphi) \leq \text{Var}(\hat{\theta})}$$

Summary: ① $\hat{\theta}$ unbiased stat. for $\tau(\theta)$
② w same quantity (for \mathbb{X})

$$\textcircled{3} \boxed{\varphi = E[\hat{\theta} | w]}$$

THEN ① $E[\varphi] = \tau(\theta)$

$$\textcircled{ii} \text{Var}(\varphi) \leq \text{Var}(\hat{\theta})$$

implication if φ is a statistic
(formula for φ doesn't depend on θ)
then φ is a better estimator than $\hat{\theta}$
(we have refuted $\hat{\theta}$)

Ex. let $X_n \stackrel{iid}{\sim} N(\theta, 1)$

define $\hat{\theta} = \frac{1}{2}(X_1 + X_2)$

notice: $E[\hat{\theta}] = \frac{1}{2}E[X_1] + \frac{1}{2}E[X_2] = \theta$

$\hat{\theta}$ unbiased for θ

$$\text{Var}(\hat{\theta}) = \frac{1}{4}\text{Var}(X_1) + \frac{1}{4}\text{Var}(X_2) = \frac{1}{2}$$

Let $W = X_1$

$$\begin{aligned}\varphi &= \mathbb{E}[\hat{\theta}|W] = \mathbb{E}\left[\frac{1}{2}(X_1 + X_2) \mid X_1\right] \\ &= \frac{1}{2}(\mathbb{E}[X_1|X_1] + \mathbb{E}[X_2|X_1])\end{aligned}$$

Aside: $\mathbb{E}[Z|Z] = ?$

$$\mathbb{E}[Z|Z=5] = 5$$

$$\mathbb{E}[Z|Z=z] = z \Rightarrow \mathbb{E}[Z|Z] = Z$$

$$= \frac{1}{2}(X_1 + \mathbb{E}[X_2])$$

$$\varphi = \frac{1}{2}(X_1 + \theta)$$

↑ not a statistic!

notice however

$$\mathbb{E}[\varphi] = \frac{1}{2}(\mathbb{E}[X_1] + \theta) = \theta \quad \checkmark$$

$$E[\varphi] = \frac{1}{2}(E[X_1] + \theta) = \theta \quad \checkmark$$

$$\text{Var}(\varphi) = \frac{1}{4} \text{Var}(X_1) = \frac{1}{4} < \frac{1}{2} = \text{Var}(\hat{\theta}). \quad \checkmark$$

Consider instead

$$\boxed{W = \bar{X}} \quad (\text{redo calculations})$$

$$\begin{aligned} \varphi &= E[\hat{\theta} | W] = E\left[\frac{1}{2}(X_1 + X_2) \mid \bar{X}\right] \\ &= \frac{1}{2}(E[X_1 | \bar{X}] + E[X_2 | \bar{X}]) \end{aligned}$$

$X_n \stackrel{iid}{\sim} f_\theta$ so $E[X_1 | \bar{X}] = E[X_2 | \bar{X}]$

$$= E[X_3 | \bar{X}] = \dots = E[X_N | \bar{X}]$$

$$\rightarrow = \frac{1}{2} 2 E[X_1 | \bar{X}]$$

$$= E[X_1 | \bar{X}]$$

$$= \frac{1}{N} N E[X_1 | \bar{X}]$$

$$= \frac{1}{N} \underbrace{(E[X_1 | \bar{X}] + E[X_1 | \bar{X}] + \dots + E[X_1 | \bar{X}])}_{N \text{ times}}$$

$$\begin{aligned}
&= \frac{1}{N} (E[X_1|\bar{X}] + E[X_2|\bar{X}] + E[X_3|\bar{X}] + \dots + E[X_N|\bar{X}]) \\
&= E\left[\frac{X_1 + X_2 + \dots + X_N}{N} \mid \bar{X}\right] \\
&= E[\bar{X} \mid \bar{X}] = \bar{X}.
\end{aligned}$$

So $\varphi = E[\hat{\theta} | W] = \bar{X}$

\uparrow \nwarrow this a statistic

we know

① $E(\varphi) = E(\bar{X}) = \theta$

② $\text{Var}(\varphi) = \text{Var}(\bar{X}) = \frac{1}{N} < \frac{1}{2} = \text{Var}(\hat{\theta})$

$X_n \stackrel{\text{iid}}{\sim} N(\theta, 1)$

③ φ is a stat.

φ is a "better" stat./est. of θ .

(in fact φ is the UMVUE)

Theorem: Rao-Blackwell Theorem

Let $\hat{\theta}$ be unbiased for $T(\theta)$ and W
a sufficient stat for θ . Then if

a sufficient stat for θ . Then if

$$\varphi = E[\hat{\theta}|w]$$

- (1) $E[\varphi] = T(\theta)$
- (2) $\text{Var}(\varphi) \leq \text{Var}(\hat{\theta})$
- (3) φ is a statistic.

i.e. φ is a better est. for $T(\theta)$ than $\hat{\theta}$.

pf. we already proved (1) and (2)

lets show (3): $\hat{\theta} = \theta(x)$ and $w = w(x)$

$$\begin{aligned} \varphi &= E[\hat{\theta}(x) | w(x)] \\ &= \int \hat{\theta}(x) f(x|w) dx \end{aligned} \quad \begin{aligned} E[X|Y=y] &= \int x f(x|y) dx \\ E[g(X)|Y=y] &= \int g(x) f(x|y) dy \end{aligned}$$

however w is sufficient
for θ

by defn $f(x|w)$ is free of θ

$\rightarrow \hat{\theta}$ free of θ , $f(x|w)$ free of θ

\Rightarrow so whole integral is free of \mathcal{O}

\Rightarrow so \mathcal{L} is free of \mathcal{O} .
i.e. \mathcal{L} is a stat!
