Want to find "best" estimaters.

Fact: In general if I'm too permissive in what I allow to be an estimator—

there is no "best" estimator.

Ex. Xn ild N(u, 1)

Want to estimate u.

1. P. want some u* so flut

MSE(u*) < MSE(û) Yu

fer any other is.

let û3 = 5

$$MSE(\hat{\mu}_3) = E[(\hat{\mu}_3 - 5)^2] = 0$$

Need to restrict class of allowable estimators otherwise there is no "best" estimator.

One way: restrict "allamable" estimaters fo unbiased ests.

Defn: Uniformly Minimu-Variance Unbiased Estimator (UMVUE)

Note: $B_{\theta}(\hat{\theta}) = 0$ then $MSE(\hat{\theta}) = Var(\hat{\theta})$

We call \hat{O}^* the UMVIIE of T(0) some for of O, if Unbiased.

E[\hat{O}^*] = T(0)

2) minimum variance

 $Var(\hat{o}^*) \leq Var(\hat{o}) \quad \forall o \in C$

nd all whiased estimators of of T(0)

Defn: Score:

Basically 30 but viewed es random.

Recall: & deterministic, & random

If Xn iid for where O∈ O ther the score

$$S_{\theta} = S_{\theta}(X) = \frac{\partial \log f_{\theta}(X)}{\partial \theta} = \frac{d \partial Q}{\partial \theta}$$

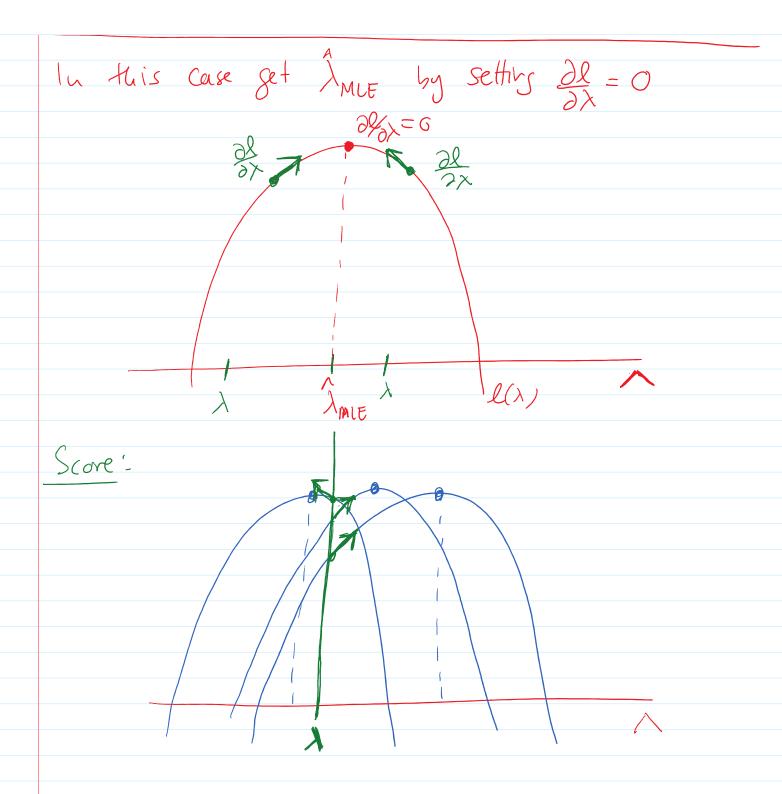
Ex. Xn i'd Exp(x)

Her $L(\chi) = \prod_{n=1}^{N} \lambda e^{-\lambda \chi_n} I(\chi_n > 0)$ = $\lambda e^{-\lambda \chi_n} I(\chi_{(1)} > 0)$

 $l(\lambda) = log L = N(g\lambda - \lambda \sum_{n} \chi_{n} + log \mathbf{1}(\chi_{co} > 0)$

 $\frac{\partial l}{\partial \lambda} = \frac{N}{\lambda} - \sum_{n} \chi_{n} \leftarrow \frac{\text{deriv of } los - lik}{n}$

Score: $S_{\lambda} = \frac{N}{\lambda} - \sum_{n} X_{n}$ now random



(n prev. example:

$$E[S_{\lambda}] = 0.$$

$$E[N - 2x_{h}] = 0$$

$$\text{true:} \frac{N}{\lambda} - 2x_{h} = 0$$

$$My MUE: S N - 2x_{h} = 0$$

$$ad get \lambda = \frac{1}{x}$$

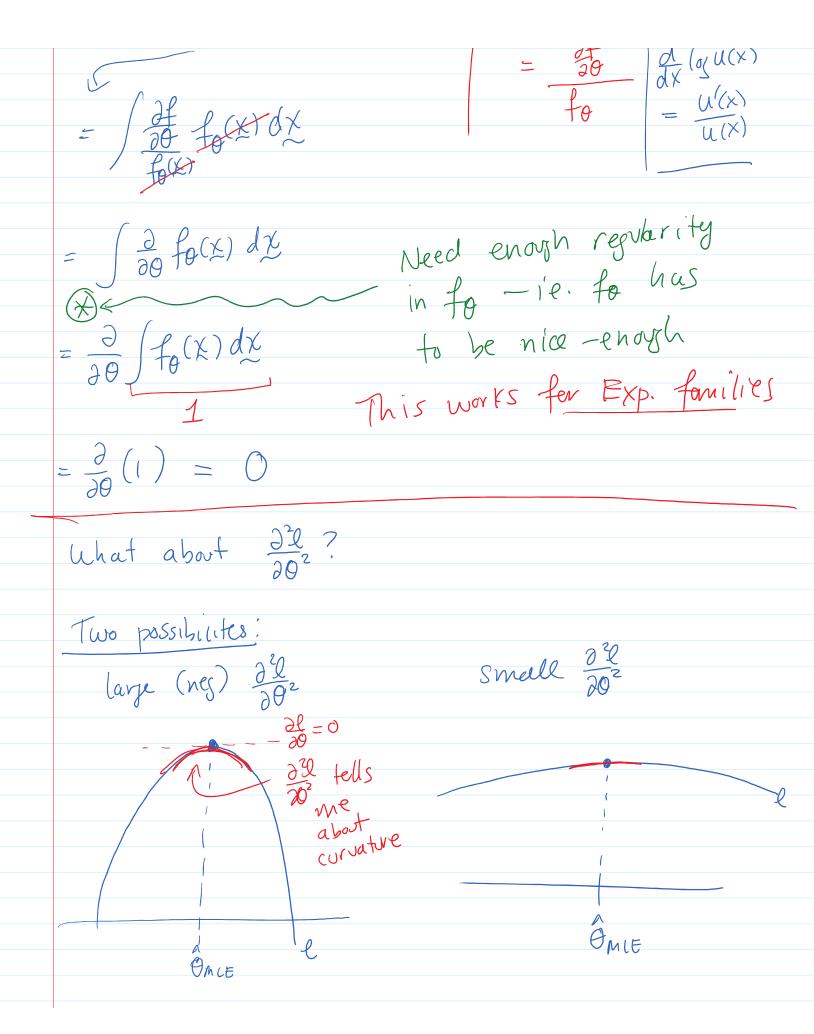
$$E[\frac{1}{x}] \neq \lambda$$

Pf.
$$E[S_0] = 0$$
 | $azy notation:$

$$S[S_0] = E[S_0(x)] = S_0(x) f_0(x) dx$$

$$E[g(x)] = S_0(x) f(x) dx$$

$$= \int \frac{\partial l}{\partial \theta} f_0(x) dx$$



$$Var(S_0) = \mathbb{E}[S_0^2] = -\mathbb{E}\left[\frac{\partial^2 \ell}{\partial O^2}\right]$$

Thinking of as random

$$\mathbb{E}\left[\left(\frac{\partial l}{\partial \theta}\right)^2\right] = -\mathbb{E}\left[\frac{\partial^2 l}{\partial \theta^2}\right].$$

Defn: Fisher Information

We define the fisher info. for θ contained in χ (N=1)

$$I(\theta) = -E\left[\frac{\partial^2}{\partial \theta^2} \log f_{\theta}(x)\right]$$

If I have N samples Xn iid then
the Fisher info in X about 0 is

$$I_N(\theta) = -E\left[\frac{\partial^2 \ell}{\partial \theta^2}\right]$$

Pf:
$$I_{N}(\theta) = -\mathbb{E}\left[\frac{\partial^{2} \ell}{\partial \theta^{2}}\right]$$

$$= -\mathbb{E}\left[\frac{\partial^{2} \ell}{\partial \theta^{2}}(\sigma_{S}f_{\theta}(X))\right]$$

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Find In(x).

(1) Find
$$\log f_{\chi}(x)$$

$$f_{\chi}(x) = \frac{\lambda e^{-\lambda}}{\chi!} \Rightarrow \log f_{\chi}(x) = \chi \log \lambda - \lambda - (og(\chi!))$$

(2) Find
$$\frac{2^2}{3\lambda^2} \left(\sigma f_{\lambda}(x) \right)$$

$$\frac{2}{3\lambda} \longrightarrow \frac{\chi}{\lambda} - 1$$

$$\frac{2^2}{3\lambda^2} \longrightarrow -\frac{\chi}{\lambda^2}$$

(3) Form
$$-\mathbb{E}\left[\frac{\partial^2(gf_{\lambda})}{\partial \lambda^2}\right] = \mathbb{I}(\lambda)$$

$$-\mathbb{E}\left[-\frac{\chi}{\lambda^2}\right] = \frac{\mathbb{E}\chi}{\lambda^2} = \frac{\lambda}{\lambda^2} = \frac{1}{\lambda}$$

$$\begin{array}{ccc}
4 & I_N(x) = NI(x) \\
& I_N(x) = N_{\lambda}.
\end{array}$$