

Theorem: Central Limit Theorem

If I have  $X_n$  that are iid w/  $EX_n = \mu$   
and  $\text{Var}(X_n) = \sigma^2 < \infty$  then

$$\sqrt{N} \left( \frac{\bar{X} - \mu}{\sigma} \right) \xrightarrow{d} N(0, 1).$$

Other forms:

$$\sqrt{N}(\bar{X} - \mu) \xrightarrow{d} N(0, \sigma^2)$$

or

$$\bar{X} \sim AN(\mu, \sigma^2/N)$$

Ex.  $X_n \stackrel{\text{iid}}{\sim} \text{Exp}(\lambda)$

then  $EX_n = \frac{1}{\lambda} = \mu$ ,  $\text{Var}(X_n) = \frac{1}{\lambda^2} = \sigma^2$

and so

$$\sqrt{N} \left( \frac{\bar{X} - \frac{1}{\lambda}}{\sqrt{\frac{1}{\lambda^2}}} \right) \xrightarrow{d} N(0, 1)$$

Q: what about  $g(\bar{X})$ ?  $\frac{1}{\bar{X}}$ ,  $\log \bar{X}$ , ...

Theorem: First-Order Delta Method

If  $Y_n$  is a seq of RVs where

$$\sqrt{N}(Y_n - \theta) \xrightarrow{d} N(0, \psi^2) \quad \psi = \psi(\theta)$$

Obvious example is  $Y_n = \bar{X}$ ,  $\theta = \mu$ ,  $\psi = \sigma^2$  then we have such a seq of  $Y_n$ s.

then if  $g$  is a differentiable function and  $g'(\theta) \neq 0$ , then

$$\sqrt{N}(g(Y_n) - g(\theta)) \xrightarrow{d} N(0, [g'(\theta)]^2 \psi^2)$$

Another way:

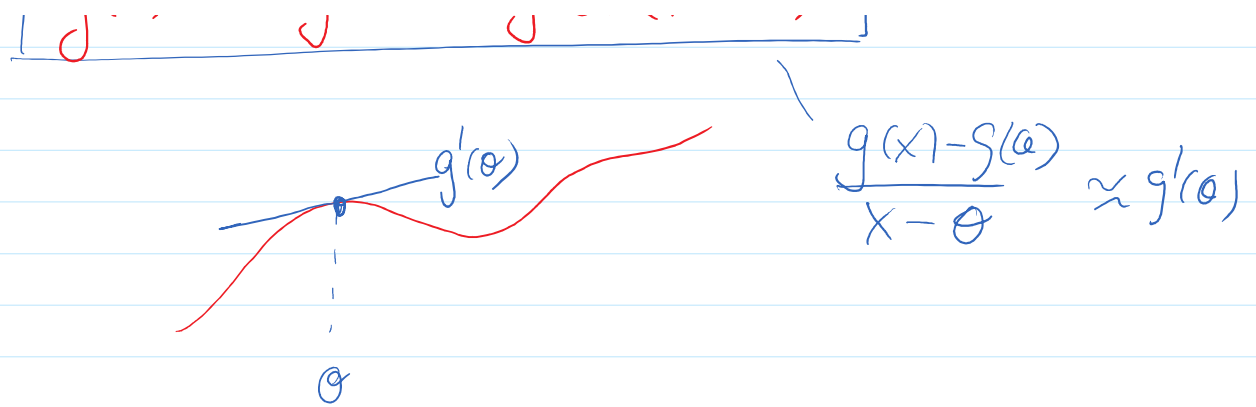
$$Y_n \sim AN(\theta, \psi^2/N)$$

then  $g(Y_n) \sim AN(g(\theta), [g'(\theta)]^2 \psi^2/N)$

Note:  $\text{Var}(cX) = c^2 \text{Var}(X)$

pf. <sup>First-Order</sup> Taylor approx of  $g$  around  $\theta$  is

$$g(x) \approx g(\theta) + g'(\theta)(x - \theta)$$



then

$$g(x) - g(\theta) \approx g'(\theta)(x - \theta)$$

so

$$\sqrt{n}(g(x) - g(\theta)) \approx g'(\theta) \sqrt{n}(x - \theta)$$

so

$$\underbrace{\sqrt{n}(g(Y_n) - g(\theta))}_{\xrightarrow{d} N(0, [g'(\theta)]^2 \psi^2)} \approx \underbrace{g'(\theta)}_{\xrightarrow{d} N(0, \psi^2)} \underbrace{\sqrt{n}(Y_n - \theta)}_{\xrightarrow{d} N(0, \psi^2)}$$

$$sd(g(x)) \approx |g'(x)| sd(x)$$

Ex. CLT says

$$\sqrt{n}(\bar{X} - \mu) \xrightarrow{d} N(0, \sigma^2)$$

If  $g(x) = \log(x)$  then  $g'(x) = \frac{1}{x}$

$$\Rightarrow [g'(x)]^2 = \frac{1}{x^2}$$

and so the  $\Delta$ -method says

$$[g'(\mu)]^2$$

$$\sqrt{N}(g(\bar{X}) - g(\mu)) \xrightarrow{d} N(0, \frac{1}{\mu^2} \sigma^2)$$

i.e.  $\boxed{\bar{X} \sim AN(\mu, \sigma^2/N) \text{ then } \log \bar{X} \sim AN(\log \mu, \frac{\sigma^2}{N\mu^2})}$

Ex: Let  $X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$  then the CLT says

$$\sqrt{N}(\bar{X} - \lambda) \xrightarrow{d} N(0, \lambda)$$

Consider  $g(x) = 1/x$  then  $g'(x) = -1/x^2$

$$\text{so } [g'(x)]^2 = 1/x^4$$

and thus the  $\Delta$ -method says

$$\sqrt{N}(g(\bar{X}) - g(\mu)) \xrightarrow{d} N(0, [g'(\mu)]^2 \sigma^2)$$

i.e.

$$\begin{aligned} \sqrt{N}\left(\frac{1}{\bar{X}} - \frac{1}{\lambda}\right) &\xrightarrow{d} N\left(0, \left(\frac{1}{\lambda}\right)^4 \lambda\right) \\ &= N\left(0, \frac{1}{\lambda^3}\right) \end{aligned}$$

i.e.  $\bar{X} \sim \text{AN}(\mu, \sigma^2/N)$  then  $\frac{1}{\bar{X}} \sim \text{AN}(\frac{1}{\mu}, \frac{1}{N\mu^3})$

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## Ex. Variance-Stabilizing Transformation

Generically:  $Y \sim \text{AN}(\theta, \psi^2/N)$   $\psi^2$  may depend on  $\theta$

Q: is there some transformation  $g$  so that

$$g(Y) \sim \text{AN}(g(\theta), \text{---})$$

$\uparrow$  doesn't depend on  $\theta$

Solu: use  $\Delta$ -method

b/c it says

$$g(Y) \sim \text{AN}(g(\theta), [g'(\theta)]^2 \psi^2)$$

Could do: choose  $g$  so that  $[g'(\theta)]^2 \psi^2 \stackrel{!!}{=} \text{constant}$  doesn't depend on  $\theta$ .

So our condition is

$$[g'(\theta)]^2 \psi(\theta)^2 = c \quad ] \text{ ODE}$$

Ex.  $X_n \stackrel{\text{iid}}{\sim} \text{Pois}(\lambda)$

CLT:

CLT:  $\sqrt{n}(\bar{X} - \mu) \xrightarrow{d} N(0, \sigma^2)$   $\Leftrightarrow \bar{X} \sim N(\mu, \sigma^2/n)$

in this case our ODE

$$g'(\lambda) \frac{\lambda}{N} = c$$

$$\Rightarrow \left(\frac{dg}{dx}\right)^2 \frac{\lambda}{N} = C$$

$$\Rightarrow \frac{dg}{dx} = \sqrt{\frac{cN}{\lambda}}$$

$$\Rightarrow dy = \sqrt{\frac{CN}{x}} dx$$

$$\Rightarrow g = \int dg = \int \frac{\sqrt{\epsilon N}}{\sqrt{x}} dx \propto \int \frac{1}{\sqrt{x}} dx \propto \sqrt{x}$$

so  $g(x) = \sqrt{x}$

By  $\Delta$ -method:

$$r(r-1)^2 \dots$$

$$g(\bar{X}) \sim AN(g(\lambda), [g'(\lambda)]^{-1} \Psi^c)$$

i.e.

$$\sqrt{\bar{X}} \sim AN(\sqrt{\lambda}, \underbrace{\left(\frac{1}{2\sqrt{\lambda}}\right)^2 \lambda}_{\frac{1}{4} \frac{1}{\lambda} \lambda = \frac{1}{4}})$$

so  $\sqrt{\bar{X}} \sim AN(\sqrt{\lambda}, \frac{1}{4})$ .

Theorem: Second Order  $\Delta$ -method

If

$$\sqrt{N}(\bar{Y}_n - \theta) \xrightarrow{d} N(0, \Psi^2)$$

and if  $g$  is twice-differentiable

but  $g'(\theta) = 0$

multiply  $\chi^2(1)$   
by  $\Psi^2 g''(\theta)$

then

$$\boxed{N(g(\bar{Y}_n) - g(\theta)) \xrightarrow{d} \frac{\Psi^2 g''(\theta)}{2} \chi^2(1)}$$

Ex. let  $X_n \stackrel{iid}{\sim} \text{Bernoulli}(p)$  and let

$$g(t) = t \log(t/p) - (1-t) \log\left(\frac{1-t}{1-p}\right)$$

↪ KL-divergence  
(dist. measure between  $\text{Bern}(p)$  and  $\text{Bern}(t)$ )

Q: what can I say about  $g(\bar{X})$ ?

CLT:  $\sqrt{N}(\bar{X} - p) \xrightarrow{d} N(0, p(1-p))$

$$\mathbb{E}X_n = p \text{ and } \text{Var}(X_n) = p(1-p)$$

notice though that

$$g'(t) = \log\left(\frac{t}{1-t}\right) - \log\left(\frac{p}{1-p}\right)$$

$$\text{and so } g'(p) = \log\left(\frac{p}{1-p}\right) - \log\left(\frac{p}{1-p}\right) = 0$$

↑  
problem for  
First-Order  
 $\Delta$  method.

let's apply the second order  $\Delta$ -method

$$g''(t) = \frac{1}{t} + \frac{1}{1-t} = \frac{1}{t(1-t)}$$

and so

$$N(g(\bar{X}) - g(p)) \xrightarrow{d} \frac{\psi^2 g''(p)}{2} \chi^2(1)$$



1.2.

$$N(g(\bar{x}) - g(p)) \xrightarrow{d} \underbrace{\frac{p(1-p)}{2} \frac{1}{p(1-p)}}_{\frac{1}{2}} \chi^2(1)$$

"proof" of Second Order Δ-method

Assumption:  $\sqrt{N}(\bar{Y}_n - \theta) \xrightarrow{d} N(0, \psi^2)$

and  $g'(\theta) = 0$

Second order

Taylor Expansion:

$$g(x) \approx g(\theta) + \cancel{g'(\theta)(x-\theta)} + \frac{g''(\theta)}{2}(x-\theta)^2$$

if  $g'(\theta) = 0$  then this is

$$g(x) \approx g(\theta) + \frac{g''(\theta)}{2}(x-\theta)^2$$

so

$$g(\bar{Y}_n) \approx g(\theta) + \frac{g''(\theta)}{2}(\bar{Y}_n - \theta)^2$$

so

$$g(\bar{Y}_n) - g(\theta) \approx \frac{g''(\theta)}{2}(\bar{Y}_n - \theta)^2$$

multiply by  $N$

"

$$\begin{aligned}
 N(g(\bar{Y}_n) - g(\theta)) &\approx \frac{g''(\theta)}{2} N(\bar{Y}_n - \theta)^2 \\
 &= \frac{g''(\theta)}{2} \underbrace{(\sqrt{n}(\bar{Y}_n - \theta))^2}_{\xrightarrow{d} N(0, \psi^2)} \\
 &\quad \xrightarrow{d} \psi N(0, 1) \\
 &\xrightarrow{d} \frac{g''(\theta)}{2} (\psi N(0, 1))^2 \\
 &\quad \underbrace{\hspace{10em}}_{\frac{g''(\theta) \psi^2}{2} \chi^2(1)}
 \end{aligned}$$


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Back to estimation:

For a finite sample we looked for estimators that are unbiased and have a low variance

Asymptotically, we want estimators that are

① asymptotically unbiased (consistency)

$$\hat{\theta} \xrightarrow{P} \theta$$

(2) asymptotic variance to be small.

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Theorem: MLEs are consistent (\*) Some conditions needed (works for Exp. fams)

If  $\hat{\theta}$  is the MLE of  $\tau(\theta)$  then

$$\hat{\theta} \xrightarrow{P} \tau(\theta)$$

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