

Exp. fam.

Said:

$$f_{\theta}(\underline{x}) = h(\underline{x}) c(\theta) \exp(t(\underline{x}) w(\theta))$$

joint.

then this is an exp. fam.

Also said!

$$f_{\theta}(x_n) = h_0(x_n) c_0(\theta) \exp(t_0(x_n) w(\theta))$$

marginal

then the joint is an exp.

This is true:

$$\begin{aligned} f_{\theta}(\underline{x}) &= \prod_n f_{\theta}(x_n) = \prod_n h_0(x_n) c_0(\theta) \exp(t_0(x_n) w(\theta)) \\ &= \underbrace{\left(\prod_n h_0(x_n) \right)}_{h(\underline{x})} \underbrace{c_0(\theta)^N}_{c(\theta)} \exp\left(\overbrace{\left(\sum_n t_0(x_n) \right)}^{t(\underline{x})} w(\theta) \right) \end{aligned}$$

$\prod_n e^{a_n} = e^{\sum_n a_n}$

$$h(\underline{x}) = \prod_n h_0(x_n)$$

$$c(\theta) = c_0(\theta)^N \quad w(\theta) \text{ same}$$

$$t(\underline{X}) = \sum_n t_0(X_n)$$

last time:

$$X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$$

$$\textcircled{1} a^b e^b = e^{a+b}$$

$$\textcircled{2} e^{\log a} = a$$

$$\textcircled{3} \log a^b = b \log a$$

could do:

$$f_\lambda(X_n) = \frac{\lambda^{X_n} e^{-\lambda} \mathbb{1}(X_n \in \mathbb{N}_0)}{X_n!} = \underbrace{\left(\frac{1}{X_n!}\right)}_{h_0(X_n)} \underbrace{(e^{-\lambda})}_{f_0(\lambda)} e^{\overbrace{X_n \log \lambda}^{t_0(X_n)}} \omega(\lambda)$$

$$t(\underline{X}) = \sum_n X_n.$$

Notation:

dual life random and number

$$\bar{X} = \frac{1}{N} \sum_{n=1}^N X_n$$

↑
RV

$$v. \quad \bar{x} = \frac{1}{N} \sum_{n=1}^N x_n$$

↑
number

$$e.g. \quad X_{(1)} \quad v. \quad x_{(1)}$$

$$X_{(n)} \quad v. \quad x_{(n)}$$

$$S^2 \quad v. \quad s^2$$

generically T v. t

Exp. fams

Examples: Poisson, Exp, Normal, Gamma, Beta,

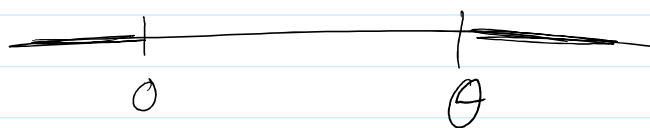
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Ex. $X_n \stackrel{iid}{\sim} U(0, \theta)$

$$\frac{1}{\theta} \text{-----} f_{\theta}(x)$$

Exp fam?

$$f_{\theta}(x) = \frac{1}{\theta} \mathbb{1}(0 < x < \theta)$$



no way to
separate into
 $(f_n x)(f_n \theta)$

So $U(0, \theta)$ isn't an exp. fam.

General fact: If the support of my dist
depends on θ ,
then it isn't an exp. fam.

Setup: $X_n \stackrel{iid}{\sim} f_{\theta}$, $\theta \in \Theta$

Defn: Sufficiency

A statistic $T = T(X)$ is sufficient for

A statistic $T = T(\underline{X})$ is sufficient for a parameter θ if

$f_{\underline{X}|T=t}(\underline{x})$ is "free" of θ .

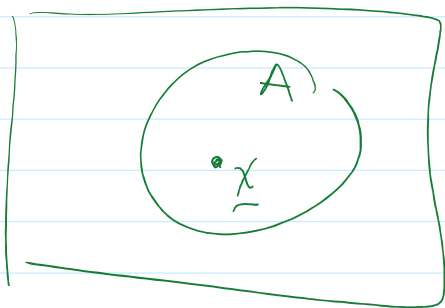
↑ θ doesn't show up in the formula

$\underline{x}, \underline{x} \in \mathbb{R}^N, A \subset \mathbb{R}^N$

$$P(\underline{X} = \underline{x} \text{ and } \underline{X} \in A)$$

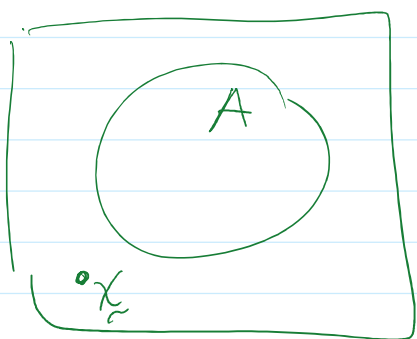
two options: $\underline{x} \in A$

①



$$P(\underline{X} = \underline{x} \text{ and } \underline{X} \in A) = P(\underline{X} = \underline{x})$$

②



$\underline{x} \notin A$

$$P(\underline{X} = \underline{x} \text{ and } \underline{X} \in A) = \underline{\underline{0}}$$

Clever! $P(\underline{X} = \underline{x} \text{ and Statement about } \underline{X})$

$$= P(\underline{X} = \underline{x}) \mathbb{I}(\text{Statement about } \underline{x})$$

Joint dist: $f_{\underline{x}, T}(\underline{x}, t) = f_{\underline{x}}(\underline{x}) \mathbb{I}(T(\underline{x}) = t)$

Ex. Let $X_1, X_2, X_3 \stackrel{iid}{\sim} \text{Bernall}(\theta)$
 $\theta \in [0, 1]$

Let $T = X_1 + X_2 + X_3 \sim \text{Bin}(3, \theta)$

is T sufficient for θ ,

$$f(\underline{x} | T=t) = \frac{f_{\underline{x}, T}(\underline{x}, t)}{f_T(t)}$$

$T \sim \text{Bin}(3, \theta)$

$P(\underline{X} = \underline{x})$

$= \prod_n \theta^{x_n} (1-\theta)^{1-x_n} \mathbb{I}(x_n = 0 \text{ or } 1)$

$= \theta^{\sum x_n} (1-\theta)^{3-\sum x_n}$

$$= \frac{P(\underline{X} = \underline{x}, T=t)}{P(T=t)}$$

$= \frac{P(\underline{X} = \underline{x}) \mathbb{I}(T(\underline{x}) = t)}{P(T=t)}$

$= \frac{\theta^{\sum x_n} (1-\theta)^{3-\sum x_n} \mathbb{I}(\sum x_n = t)}{\binom{3}{t} \theta^t (1-\theta)^{3-t}}$

$= \frac{\mathbb{I}(\sum x_n = t)}{\binom{3}{t}}$

\leftarrow No θ !

$$\binom{3}{t}$$

So T is sufficient for θ .

Ex. let $X_n \stackrel{iid}{\sim} f_\theta$

$$T = (X_1, X_2, \dots, X_N)$$

Q: is this sufficient for θ ?

$$f(\underline{x} | T=t) = \frac{f_{\underline{x}|T}(\underline{x}, t)}{f_T(t)}$$

$$= \frac{f_{\underline{x}, \underline{x}}(\underline{x}, \underline{x})}{f_{\underline{x}}(\underline{x})}$$

$$= f_{\underline{x}}(\underline{x}) / f_{\underline{x}}(\underline{x}) = 1 \quad \leftarrow \text{free of } \theta$$

Factorization Theorem

T is sufficient for θ iff

there is a fn $g(\theta, t)$ and $h(\underline{x})$
so that

$$f_{\theta}(\underline{x}) = g(\theta, t) h(\underline{x}).$$

only has \underline{x} 's through t

Ex. $X_n \stackrel{iid}{\sim} \text{Bern}(\theta)$

$$T = \sum_{n=1}^N X_n, \quad t = \sum_n X_n$$

is T sufficient for θ ?

$$\begin{aligned} f_{\theta}(\underline{x}) &= \prod_n \theta^{x_n} (1-\theta)^{1-x_n} \mathbb{1}(x_n = 0 \text{ or } 1) \\ &= \underbrace{\theta^{\sum_n x_n} (1-\theta)^{N - \sum_n x_n}}_{g(\theta, t)} \underbrace{\prod_n \mathbb{1}(x_n = 0 \text{ or } 1)}_{h(\underline{x})} \\ &= \theta^t (1-\theta)^{N-t} \end{aligned}$$

Then $f_{\theta}(\underline{x}) = g(\theta, t) h(\underline{x})$ so T is suff. for θ .

Ex. let $X_n \stackrel{iid}{\sim} U(0, \theta)$

Can I find a suff. stat. for θ ?

$$f(\underline{x}) = \prod_n \frac{1}{\theta} \mathbb{I}(0 < x_n < \theta)$$

$$= \theta^{-N} \prod_n \mathbb{I}(0 < x_n < \theta)$$

all x_n betw. 0 and θ
min > 0 and max $< \theta$

$$= \theta^{-N} \mathbb{I}(x_{(1)} > 0) \mathbb{I}(x_{(n)} < \theta)$$

$$= g(\theta, t) h(\underline{x})$$

$$\text{let } t = x_{(n)}$$

so $x_{(n)}$ is sufficient for θ

$$\text{and } g(\theta, t) = \theta^{-N} \mathbb{I}(x_{(n)} < \theta)$$

$$h(\underline{x}) = \mathbb{I}(x_{(1)} > 0)$$

Theorem: Exp. and Sufficiency

let $x_n \stackrel{\text{iid}}{\sim} f_\theta$ and

$$\text{joint } f_\theta(\underline{x}) = h(\underline{x}) c(\theta) \exp(t(\underline{x}) w(\theta))$$

so that it is an exp. fam.

Then $t(\underline{x})$ is sufficient for θ .

$$\text{let } x_n \stackrel{\text{iid}}{\sim} N(\mu, 1)$$

$$| \quad -a + b \quad a \quad b$$

Let $X_n \stackrel{iid}{\sim} N(\mu, 1)$

$$f(x_n) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}(x_n - \mu)^2\right)$$

$$e^{a+b} = e^a e^b$$

$$= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}x_n^2 + x_n\mu - \frac{1}{2}\mu^2\right)$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x_n^2} e^{x_n\mu} e^{-\frac{1}{2}\mu^2}$$

$$= \underbrace{\frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x_n^2}}_{h_0(x_n)} \underbrace{e^{-\frac{1}{2}\mu^2}}_{c_0(\mu)} \exp\left(\underbrace{x_n}_{t_0(x_n)} \underbrace{\mu}_{w(\mu)}\right)$$

$$t(\underline{x}) = \sum_n t_0(x_n) = \sum_{n=1}^N x_n$$

So by our theorem $\sum_n x_n$ is suff. for θ .

Ex. Let $X_n \stackrel{iid}{\sim} \text{Exp}(\lambda)$

$$f(x_n) = \underbrace{\lambda e^{-\lambda x_n}}_{c_0(\lambda)} \underbrace{\mathbb{I}(x_n > 0)}_{h_0(x_n)}$$

So $\text{Exp}(\lambda)$ is a exp. fam.

$$\text{and } t(\underline{x}) = \sum_n t_0(x_n) = \sum_n x_n.$$

So by our theorem $\sum_n X_n$ is suff. for λ .