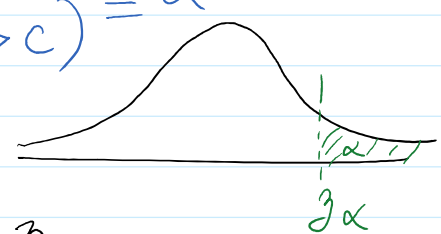


Ex. $H_0: \theta \leq \theta_0$ v. $H_a: \theta > \theta_0$

$$X_n \stackrel{iid}{\sim} N(\theta, \sigma^2)$$

Let $T = \frac{\bar{X} - \theta_0}{\sigma/\sqrt{N}}$ Known so that $P_{\theta_0}(T > c) = \alpha$



The test $T > c$ when $c = z_\alpha$

KR says this is the UMP level α test

So long as T has the MLR

Need to check that T has the MLR property.

MLR: if $\theta_1 < \theta_2$ then $\frac{f_{\theta_2}(t)}{f_{\theta_1}(t)}$ inc. fn of t .

$$\bar{X} \sim N(\theta, \sigma^2/N)$$

$$T = \frac{\bar{X} - \theta_0}{\sigma/\sqrt{N}} \sim N(\underbrace{\theta - \theta_0}_{\mu}, 1)$$

$$T \sim N(\mu, 1)$$

$$f_{\mu}(t) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}(t-\mu)^2\right)$$

$$= \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}t^2 - \frac{1}{2}\mu^2 + t\mu\right)$$

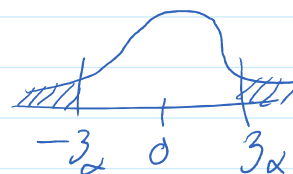
$$= \underbrace{\frac{1}{\sqrt{2\pi}} \exp\left(-\frac{1}{2}t^2\right)}_{h(t)} \underbrace{\exp\left(-\frac{1}{2}\mu^2\right)}_{c(\mu)} \exp\left(\overset{t}{\uparrow} \overset{w(\mu)=\mu}{\uparrow} t\mu\right)$$

Had a theorem said if w inc. as a fn of μ
then T has the MLR property. ✓

let's again consider $X_n \stackrel{\text{iid}}{\sim} N(0, \sigma^2)$

Test 1: rej. if $\frac{\bar{X} - \theta_0}{\sigma/\sqrt{n}} > z_\alpha$

Test 2: rej. if $\frac{\bar{X} - \theta_0}{\sigma/\sqrt{n}} < -z_\alpha$



Neyman-Pearson

$$H_0: \theta = \theta_0 \quad \text{v.} \quad H_a: \theta = \theta_a \quad \Rightarrow \quad \text{UMP } \alpha \text{ is } T1 \quad \text{--- } \theta_a > \theta_0$$

$$H_0: \theta = \theta_0 \quad \text{v.} \quad H_a: \theta = \theta_a \quad \Rightarrow \quad \text{UMP } \alpha \text{ is } T2$$

\uparrow
 $\theta_a < \theta_0$

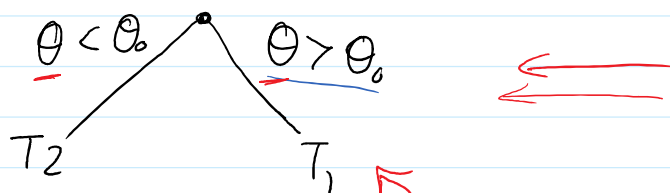
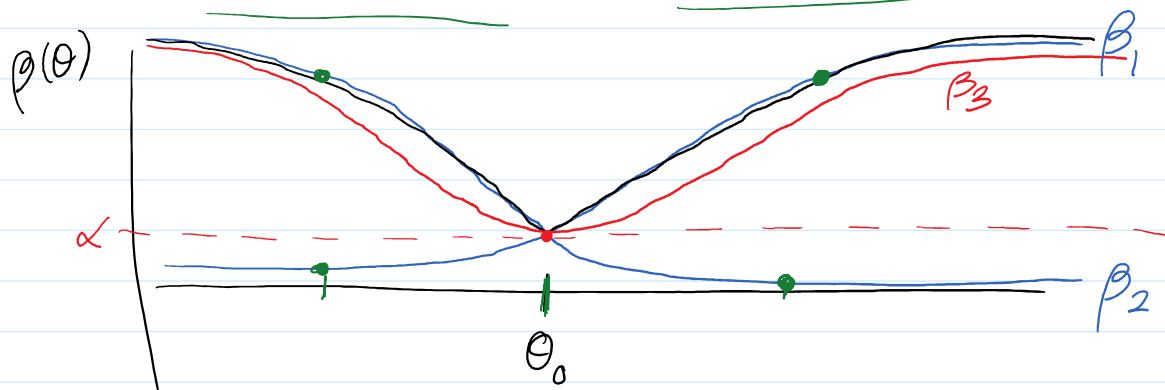
Karlin-Rubin

$$H_0: \theta \leq \theta_0 \quad \text{v.} \quad H_a: \theta > \theta_0 \quad \Rightarrow \quad \text{UMP } \alpha \text{ is } T1$$

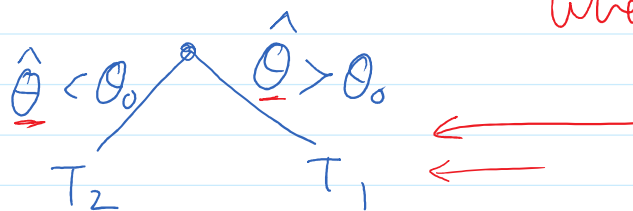
$$H_0: \theta \geq \theta_0 \quad \text{v.} \quad H_a: \theta < \theta_0 \quad \Rightarrow \quad \text{UMP } \alpha \text{ is } T2$$

Q: What about

$$H_0: \theta = \theta_0 \quad \text{v.} \quad H_a: \theta \neq \theta_0 \quad ?$$



Test 3:
$$\frac{|\bar{X} - \theta_0|}{\sigma/\sqrt{n}} > z_{\alpha/2}$$



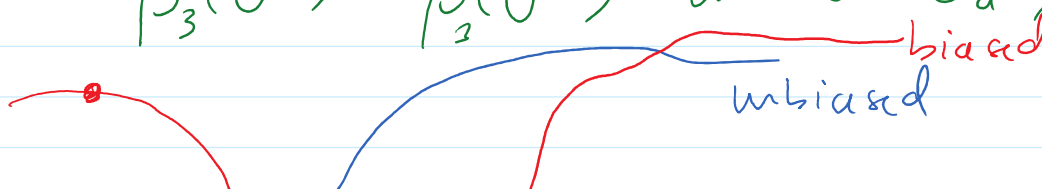
when $\theta \approx \theta_0$

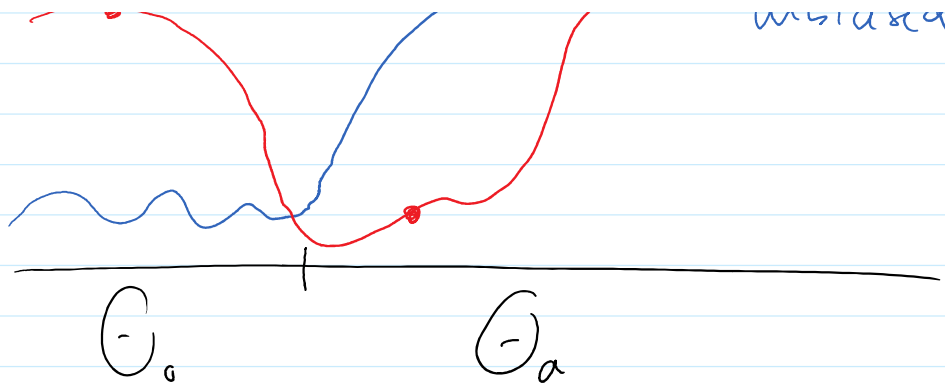
level α

Test 3 the Uniformly Most Powerful Unbiased Test

Unbiased:

$$\beta_3(\theta') \geq \beta_2(\theta'') \quad \text{for } \theta' \in \Theta_a, \theta'' \in \Theta_0$$





Interval Estimation

Point Estimation: $\hat{\theta} \approx \theta$

New: want to say $\theta \in C$ (approx.) when

$$C = C(\underline{x}) \subset (-\infty, \infty)$$

preferably, C is some interval.

Defn: Interval Estimator

An interval est. of $\theta \in (-\infty, \infty) \subset \mathbb{R}$ is
any pair of fns

$$L = L(\underline{x}), \quad U = U(\underline{x}) \text{ that}$$

satisfy $L \leq U$.

We say: $L \leq \theta \leq U$ (at least approx.)

Sometimes: might want a "one-sided" interval

i.e. $L = -\infty$ or $U = \infty$

so that our interval is

$(-\infty, u]$ or $[L, \infty)$.

Say:

$\theta \leq u$

$\theta \geq L$

Ex. let $X_1, \dots, X_n \stackrel{iid}{\sim} N(\mu, 1)$ then an interval est. of μ is

$$[\underbrace{\bar{X} - 1}_L, \underbrace{\bar{X} + 1}_U]$$

might say: $\bar{X} - 1 \leq \theta \leq \bar{X} + 1$.

why? just use \bar{X} ?

$$P(\bar{X} = \mu) = 0.$$

Need to attach some meas. of error. ($sd(\bar{X})$)

Alt.

$$P(\underbrace{\bar{X} - 1}_\uparrow \leq \underbrace{\mu}_{\text{fixed}} \leq \underbrace{\bar{X} + 1}_\uparrow) > 0$$

random

Ex. cont.

$$\mathbb{P}(\mu \in [\bar{x} - 1, \bar{x} + 1])$$

$$= \mathbb{P}(\bar{x} - 1 \leq \mu \leq \bar{x} + 1)$$

$$= \mathbb{P}(\bar{x} - 1 \leq \mu, \bar{x} + 1 \geq \mu)$$

$$X_1, \dots, X_n \stackrel{\text{iid}}{\sim} N(\mu, 1)$$

$$= \mathbb{P}(\bar{x} - \mu \leq 1, \bar{x} - \mu \geq -1) \quad \bar{x} - \mu \sim N(0, 1/4)$$

$$= \mathbb{P}(-1 \leq \bar{x} - \mu \leq 1)$$

$$= \mathbb{P}\left(-2 \leq \underbrace{\frac{\bar{x} - \mu}{1/2}}_{N(0,1)} \leq 2\right)$$

$$Z \sim N(0, 1)$$

$$= \mathbb{P}(|Z| \leq 2)$$

$$\approx .95$$



So the chance that $[\bar{x} - 1, \bar{x} + 1]$ covers μ is 95%.

Defn: For an interval estimator $[L, u]$ of a parameter θ we define the coverage probability to be

$$P_{\theta}(L \leq \theta \leq u)$$

Defn: Confidence Coefficient

Worst-case coverage prob.

For an interval est. $[L, u]$ the conf. coef. is

$$1 - \alpha = \min_{\theta \in \Theta} P_{\theta}(L \leq \theta \leq u)$$

Defn: Confidence Interval

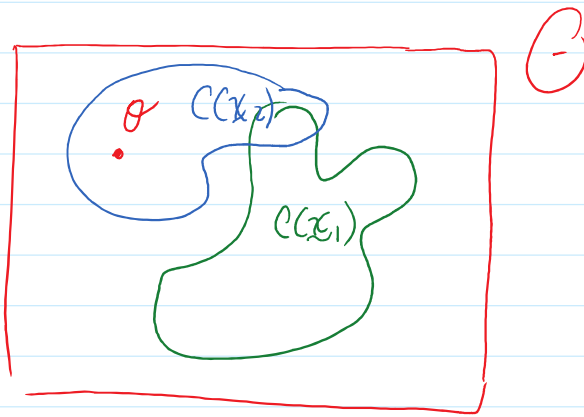
An interval est. along w/ its conf coef

Defn: Confidence Set

A set $C(\underline{x}) \subset \Theta$ and its assoc. conf. coef

$$1 - \alpha = \min_{\theta \in \Theta} P(\theta \in C)$$





How do we build a confidence set / interval.

Basically one way: invert a hypothesis test

$$HT \Leftrightarrow \text{Conf. Set.}$$

Ex. $X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ known

Consider HT for

$$H_0: \mu = \mu_0 \quad v. \quad H_a: \mu \neq \mu_0$$

For some fixed sig. level α the UMPU level α test is to reject when

$$\frac{|\bar{X} - \mu_0|}{\sigma/\sqrt{n}} > z_{\alpha/2}$$

$$R(\mu_0) = \left\{ \underline{x} \in \mathcal{X} \mid \frac{|\bar{x} - \mu_0|}{\sigma/\sqrt{N}} > z_{\alpha/2} \right\}$$

↑ reject when data not in agreement w/ μ_0

$$A(\mu_0) = \mathcal{X} \setminus R(\mu_0)$$

$$= \left\{ \underline{x} \in \mathcal{X} \mid \frac{|\bar{x} - \mu_0|}{\sigma/\sqrt{N}} \leq z_{\alpha/2} \right\}$$

↑ data is in good agreement w/ μ_0

$$\Leftrightarrow -z_{\alpha/2} \leq \frac{\bar{x} - \mu_0}{\sigma/\sqrt{N}} \leq z_{\alpha/2}$$

$$\Leftrightarrow \mu_0 - \frac{\sigma}{\sqrt{N}} z_{\alpha/2} \leq \bar{x} \leq \mu_0 + \frac{\sigma}{\sqrt{N}} z_{\alpha/2}$$

$$\Leftrightarrow \underbrace{\bar{x} - \frac{\sigma}{\sqrt{N}} z_{\alpha/2}}_L \leq \mu_0 \leq \underbrace{\bar{x} + \frac{\sigma}{\sqrt{N}} z_{\alpha/2}}_U$$

Since my test is size α then

$$P_{\mu_0}(\text{rej}) = \alpha$$

$$\Leftrightarrow P_{\mu_0}(\text{accept}) = 1 - \alpha$$

and so

$$(*) P_{\mu_0} \left(\bar{X} - \frac{\sigma}{\sqrt{N}} z_{\alpha/2} \leq \mu \leq \bar{X} + \frac{\sigma}{\sqrt{N}} z_{\alpha/2} \right) = 1 - \alpha$$

↑ true $\forall \mu$

$$\text{So if } \left. \begin{aligned} L &= \bar{X} - \frac{\sigma}{\sqrt{N}} z_{\alpha/2} \\ U &= \bar{X} + \frac{\sigma}{\sqrt{N}} z_{\alpha/2} \end{aligned} \right\}$$

$$\text{then } \min_{\mu \in \mathbb{R}} P(L \leq \mu \leq U) = 1 - \alpha \quad]$$

and so ... $[L, U]$ is a $1 - \alpha$ conf. interval.

Test Inversion

For any $\theta_0 \in \Theta$ let $A(\theta_0)$ be the accept. region for the α -level test

$$H_0: \theta = \theta_0 \quad \text{v.} \quad H_a: \theta \neq \theta_0$$

and let

$$C(\underline{x}) = \{ \theta \mid \underline{x} \in A(\theta) \} \subset \Theta$$

Then C is a $1-\alpha$ confidence set.

