

Single Sample: $f_{\theta}(x)$, $\log f_{\theta}(x)$

Multiple samples:

$$L(\theta) = f_{\theta}(\underline{x}) \quad , \quad \ell(\theta) = \log f_{\theta}(\underline{x})$$

$$\begin{array}{c} \swarrow \quad \searrow \\ \frac{\partial \ell}{\partial \theta} \quad \frac{\partial^2 \ell}{\partial \theta^2} \end{array}$$

Score: basically treat $\frac{\partial \ell}{\partial \theta}$ as random
(replace \underline{x} , $\underline{\underline{x}}$)

$$S_{\theta} = \frac{\partial}{\partial \theta} \log f_{\theta}(\underline{\underline{x}})$$

Fisher Info: one sample version

$$I(\theta) = - \mathbb{E} \left[\frac{\partial^2}{\partial \theta^2} \log f_{\theta}(\underline{\underline{x}}) \right]$$

$$I_N(\theta) = NI(\theta) = - \mathbb{E} \left[\frac{\partial^2}{\partial \theta^2} \log f_{\theta}(\underline{\underline{x}}) \right]$$

$$= - \mathbb{E} \left[\frac{\partial^2 \ell}{\partial \theta^2} \right]$$

Ex.

let $X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$ assume known

what is the Fisher Info $I_N(\mu)$?

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$$\textcircled{1} f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2}(x-\mu)^2\right)$$

$$\textcircled{2} \log f(x) = -\frac{1}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2}(x-\mu)^2$$

$$\begin{aligned}\textcircled{3} \frac{\partial \log f(x)}{\partial \mu} &= -\frac{1}{2\sigma^2} 2(x-\mu)(-1) \\ &= \frac{1}{\sigma^2}(x-\mu)\end{aligned}$$

$$\frac{\partial^2 \log f(x)}{\partial \mu^2} = -1/\sigma^2$$

$$\begin{aligned}\textcircled{4} I(\mu) &= -\mathbb{E}\left[\frac{\partial^2 \log f}{\partial \mu^2}\right] = -\mathbb{E}\left[-1/\sigma^2\right] \\ &= 1/\sigma^2\end{aligned}$$

$$I_N(\mu) = N/\sigma^2$$

$$\left[\text{Var}(\bar{X}) = \frac{\sigma^2}{N} \right]$$

Ex. $X_n \stackrel{\text{iid}}{\sim} \text{Pois}(\lambda)$

Recall $E X_n = \lambda = \text{Var } X_n$

$$\text{Sd}(X_n) = \sqrt{\lambda} = \psi \Leftrightarrow \psi^2 = \lambda$$

Reparameterize Pois in terms of ψ

$$f_{\lambda}(x) = \frac{\lambda^x e^{-\lambda}}{x!} = \frac{\psi^{2x} e^{-\psi^2}}{x!} = f_{\psi}(x)$$

Can get $I_N(\psi)$ using prev. procedure,

$$(1) \log f_{\psi} = 2x \log \psi - \psi^2 - \log(x!)$$

$$(2) \frac{\partial}{\partial \psi} \log f_{\psi} = \frac{2x}{\psi} - 2\psi$$

$$\frac{\partial^2}{\partial \psi^2} \log f_{\psi} = -\frac{2x}{\psi^2} - 2$$

$\lambda = \psi^2$

$$(3) I(\psi) = -E \left[\frac{\partial^2}{\partial \psi^2} \log f_{\psi} \right] = \frac{2EX}{\psi^2} + 2$$
$$= \frac{2\psi^2}{\psi^2} + 2 = 4$$

$$I_N(\psi) = 4N$$

$$\text{Recall: } I_N(\lambda) = \frac{N}{\lambda}$$

Derivative Review:

$$y = f(x) \Leftrightarrow x = f^{-1}(y)$$

$$\frac{dy}{dx} \xleftrightarrow{\text{rel. ?}} \frac{dx}{dy}$$

$$\text{Recall: } \frac{dy}{dx} = \frac{1}{\frac{dx}{dy}} \Leftrightarrow \frac{dx}{dy} = \frac{1}{\frac{dy}{dx}}$$

Theorem: If $\theta = \mathcal{I}(\psi)$
 \uparrow same fn

Then

$$I(\theta) = \left(\frac{\partial \psi}{\partial \theta} \right)^2 I(\psi)$$

$$I(\psi) = \left(\frac{\partial \theta}{\partial \psi} \right)^2 I(\theta)$$

Revisit Pois example,

$$I(\lambda) = 1/\lambda$$

$$\psi = \sqrt{\lambda} \Leftrightarrow \psi^2 = \lambda$$

$$\frac{d\lambda}{d\psi} = 2\psi$$

$$I(\psi) = \left(\frac{d\lambda}{d\psi}\right)^2 I(\lambda)$$

$$= (2\psi)^2 \frac{1}{\lambda}$$

$$= 4\psi^2 \frac{1}{\psi^2} = 4$$

Why do we care?

$T(\theta)$ could be θ

$\hat{\theta}^*$ is the UMVUE for $T(\theta)$,

if

$$(1) E\hat{\theta}^* = T(\theta)$$

$$(2) \text{Var}(\hat{\theta}^*) \leq \text{Var}(\hat{\theta})$$

↑ unbiased est
of $T(\theta)$

Theorem:

If $X_n \stackrel{\text{iid}}{\sim} f_\theta$ where $\theta \in \Theta$ and

$\hat{\theta}$ is unbiased for $T(\theta)$

(*) and f_θ is nice enough
read: 1-dim'l exp. family (*)

then

$$\text{Var}(\hat{\theta}) \geq \frac{\left(\frac{\partial T}{\partial \theta}\right)^2}{I_N(\theta)}$$

← called the Cramér-Rao lower bound (CRLB)

Note: If $T(\theta) = \theta$, $\left(\frac{\partial T}{\partial \theta}\right)^2 = 1^2 = 1$

So the CRLB is just $1/I_N(\theta)$

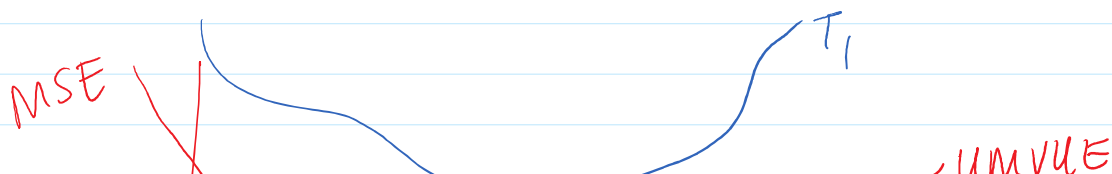
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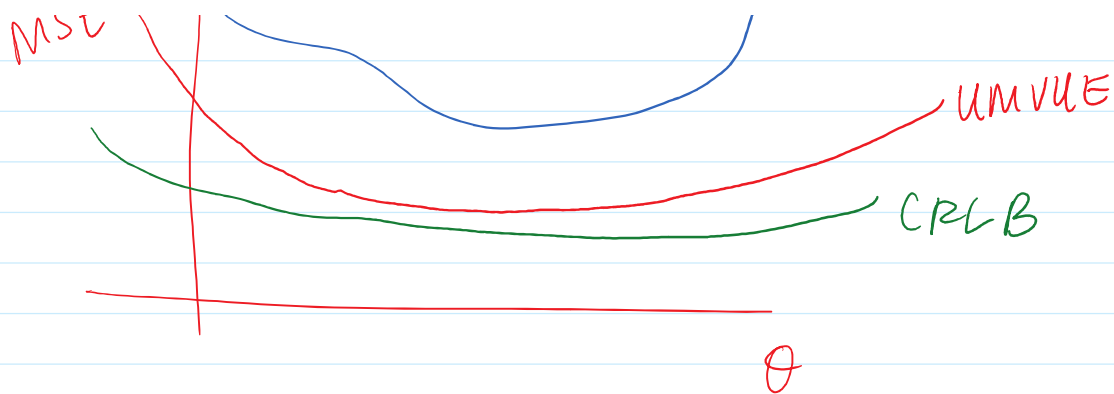
(1) If I have an unbiased est. $\hat{\theta}^*$ for $T(\theta)$ and I can show that

$$\text{Var}(\hat{\theta}^*) = \text{CRLB}$$

then $\hat{\theta}^*$ is the UMVUE.

(2) If an estimator doesn't achieve the CRLB I don't otherwise know it's not the UMVUE.





Ex. $X_n \stackrel{iid}{\sim} \text{Pois}(\lambda)$

let $\hat{\lambda} = \bar{X}$ Est λ , $\tau(\lambda) = \lambda$

$E\bar{X} = \lambda$ (its unbiased for λ)

$$\text{Var}(\bar{X}) = \frac{\text{Var}X_1}{N} = \left(\frac{\lambda}{N} \right)$$

CRLB

$$\downarrow$$

$$\beta = \frac{1}{I_N(\lambda)} = \frac{1}{N/\lambda} = \left(\frac{\lambda}{N} \right)$$

General Steps:

- ① propose unbiased est.
- ② calc its variance
- ③ calc. CRLB
- ④ show ② & ③ are equal \Rightarrow this is UMVUE

Ex. $X_n \stackrel{iid}{\sim} \text{Exp}(\lambda)$

Recall: $E X_n = \frac{1}{\lambda}$
 $\text{Var } X_n = \frac{1}{\lambda^2}$

Goal: Find UMVUE for $\frac{1}{\lambda}$. $[T(x) = \frac{1}{\lambda}]$

① Propose unbiased est. : \bar{X}

$$E \bar{X} = \frac{1}{\lambda} ;$$

② Calc. var

$$\text{Var}(\bar{X}) = \frac{1}{\lambda^2 N}$$

③ Calc. CRLB for estimating $\frac{1}{\lambda}$

$$T(x) = \frac{1}{\lambda}, \quad \frac{\partial T}{\partial \lambda} = -\frac{1}{\lambda^2} ; \quad \boxed{\left(\frac{\partial T}{\partial \lambda}\right)^2 = \frac{1}{\lambda^4}}$$

$$\frac{I(\lambda)}{I(\lambda)}$$

$$\rightarrow f_{\lambda}(x) = \lambda e^{-\lambda x}$$

$$\rightarrow \log f_{\lambda} = \log \lambda - \lambda x$$

$$\rightarrow \frac{\partial \log f_{\lambda}}{\partial \lambda} = \frac{1}{\lambda} - x$$

$$\rightarrow I(\lambda) = -E\left[-\frac{1}{\lambda^2}\right] = \frac{1}{\lambda^2}$$

$$\rightarrow \frac{\partial \log f_{\lambda}}{\partial \lambda} = \frac{1}{\lambda} - \lambda = \frac{1}{\lambda^2}$$

$$\rightarrow \frac{\partial^2 \log f_{\lambda}}{\partial \lambda^2} = -\frac{1}{\lambda^2}$$

$$\text{So } I_N(\lambda) = \frac{N}{\lambda^2}$$

$$B = \left(\frac{\partial T}{\partial \lambda} \right)^2 / I_N(\lambda) = \frac{\frac{1}{\lambda^4}}{N/\lambda^2} = \frac{1}{N\lambda^2}$$

Since \bar{X} unbiased for $\frac{1}{\lambda}$ and

$$\text{Var}(\bar{X}) = \frac{1}{N\lambda^2} = B$$

\bar{X} is the UMVUE.

Ex. $X_n \stackrel{\text{iid}}{\sim} N(\mu, \sigma^2)$ σ^2 is known

Saw earlier: $I_N(\mu) = \frac{N}{\sigma^2}$

If we want to estimate μ ,

(1) $E\bar{X} = \mu$ (\bar{X} is unbiased for μ)

(2) $\text{Var}\bar{X} = \frac{\sigma^2}{N}$

(3) $R = \frac{1}{\dots} = \frac{\sigma^2}{N}$

$$\textcircled{3} B = \frac{1}{I_N(\mu)} = \left(\frac{\sigma^2}{N} \right)$$

④ \bar{X} is the UMVUE.

Ex. Let $X_n \stackrel{iid}{\sim} U(0, \theta)$

Want the UMVUE for θ .

① Propose unbiased Est. for θ .

$$T = \frac{N+1}{N} X_{(N)}$$

then $E T = \theta$

Can show:
 $E[X_{(N)}] = \frac{N}{N+1} \theta$

② Calc. Variance

$$\text{Var}(T) = \frac{\theta^2}{N(N+2)}$$

③ CRLB

$$f_{\theta}(x) = \frac{1}{\theta} \mathbb{I}(0 < x < \theta)$$

$$\log f_{\theta}(x) = -\log \theta + \log(\mathbb{I}(0 < x < \theta)) \dots$$