Session 1

A practical introduction to MCMC

Matt Denwood 2021-06-22

Course Outline and Practicalities

Overview

Date/time:

- 28th June to 1st July 2021
- 09.00 12.30 daily
- Use the same Zoom link all week!

Teachers:

- Matt Denwood (University Of Copenhagen)
- Nils Toft (IQinAbox)
- Søren Saxmose Nielsen (University Of Copenhagen)
- Maj Beldring Henningsen (University of Copenhagen)

Motivation

Diagnostic test evaluation: with gold standard

• Evaluation of a new test against a gold standard is simple!

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```
se <- c(1, 0.6)
    sp < -c(1, 0.9)
    N <- 1000; prevalence <- 0.25
4
    data <- tibble(Status = rbinom(N, 1, prevalence)) %>%
5
      mutate(Test 1 = Status) %>%
6
      mutate(Test 2 = rbinom(N, 1, se[2]*Status + (1-sp[2])*(1-Status)))
    twoXtwo <- with(data, table(Test 1, Test 2))
9
10
    (sensitivity <- twoXtwo[2,2] / sum(twoXtwo[2,1:2]))
11
    ## [1] 0.6015625
    (specificity <- twoXtwo[1,1] / sum(twoXtwo[1,1:2]))
    ## [1] 0.9018817
```

Diagnostic test evaluation: no gold standard

What happens when the reference test is imperfect?

```
1  se <- c(0.9, 0.6)
2  sp <- c(0.95, 0.9)
3  N <- 1000
4  prevalence <- 0.25
5
6  library("tidyverse")
7
8  data <- tibble(Status = rbinom(N, 1, prevalence)) %>%
9  mutate(Test_1 = rbinom(N, 1, se[1]*Status + (1-sp[1])*(1-Status))) %>%
10  mutate(Test_2 = rbinom(N, 1, se[2]*Status + (1-sp[2])*(1-Status))) %>%
11  print()
```

The solution

We can

Revision

Bayes Rule

Bayes' theorem is at the heart of Bayesian statistics:

$$P(\theta|Y) = \frac{P(\theta) \times P(Y|\theta)}{P(Y)}$$

6

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Where: θ is our parameter value(s);

Y is the data that we have observed;

 $P(\theta|Y)$ is the posterior probability of the parameter value(s);

 $P(\theta)$ is the prior probability of the parameters;

 $P(Y|\theta)$ is the likelihood of the data given the parameters value(s);

P(Y) is the probability of the data, integrated over parameter space.

• In practice we usually work with the following:

$$P(\theta|Y) \propto P(\theta) \times P(Y|\theta)$$

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$$P(\theta|Y) \propto P(\theta) \times P(Y|\theta)$$

- Our Bayesian posterior is therefore always a combination of the likelihood of the data, and the parameter priors
- But for more complex models the distinction between what is 'data' and 'parameters' can get blurred!

MCMC

- A way of obtaining a numerical approximation of the posterior
- Highly flexible
- Not inherently Bayesian but most widely used in this context
- Assessing convergence is essential, otherwise we may not be summarising the true posterior
- Our chains are correlated so we need to consider the effective sample size

Preparation

- Any questions so far? Anything unclear?
- Do we all have R and JAGS installed?
- Can we all access the teaching material from GitHub?

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Any problems: see us during the first practical session!

Session 1a: Theory and application

of MCMC

MCMC in Practice

- We can write a Metropolis algorithm but this is complex and inefficient
- There are a number of general purpose languages that allow us to define the problem and leave the details to the software:
 - WinBUGS/OpenBUGS
 - Bayesian inference Using Gibbs Sampling
 - JAGS
 - Just another Gibbs Sampler
 - Stan
 - Named in honour of Stanislaw Ulam, pioneer of the Monte Carlo method

JAGS

- JAGS uses the BUGS language
 - This is a declarative (non-procedural) language
 - The order of statements does not matter
 - The compiler converts our model syntax into an MCMC algorithm with appropriately defined likelihood and prior
 - You can only define each variable once!!!

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 - The compiler converts our model syntax into an MCMC algorithm with appropriately defined likelihood and prior
 - You can only define each variable once!!!
- Different ways to run JAGS from R:
 - rjags, runjags, R2jags, jagsUI
- See http://runjags.sourceforge.net/quickjags.html
 - This is also in the GitHub folder

A simple JAGS model might look like this:

```
model{
1
      # Likelihood part:
      Positives ~ dbinom(prevalence, TotalTests)
5
      # Prior part:
      prevalence ~ dbeta(2, 2)
6
7
      # Hooks for automatic integration with R:
8
      #data# Positives, TotalTests
9
      #monitor# prevalence
10
      #inits# prevalence
11
12
```

There are two model statements:

- 1 Positives ~ dbinom(prevalence, TotalTests)
 - states that the number of *Positive* test samples is Binomially distributed with probability parameter *prevalence* and total trials *TotalTests*

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These are very similar to the likelihood and prior functions defined in the preparatory exercise

The other lines in this model:

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Compared to our Metropolis algorithm, this JAGS model is:

- Eaiser to write and understand
- More efficient (lower autocorrelation)
- Faster to run

To run this model, copy/paste the code above into a new text file called "basicjags.bug" in the same folder as your current working directory. Then run:

```
library('runjags')
   ## Attaching runjags (version 2.2.0-2) and setting user-specified

→ options

   ##
   ## Attaching package: 'runjags'
   ## The following object is masked from 'package:tidyr':
   ##
   ##
          extract
3
   # data to be retrieved by runjags:
   Positives <- 7
3
   TotalTests <- 10
4
5
   # initial values to be retrieved by runjags:
   prevalence <- list(chain1=0.05, chain2=0.95)
```

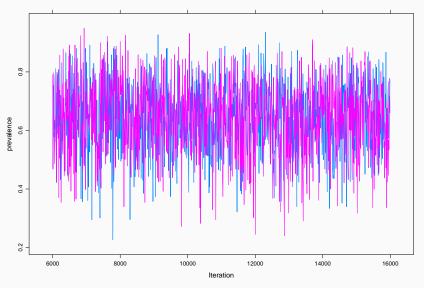
```
results <- run.jags('basicjags.bug', n.chains=2, burnin=5000,

⇔ sample=10000)
```

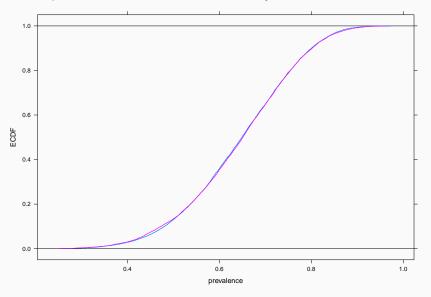
First check the plots for convergence:

plot(results)

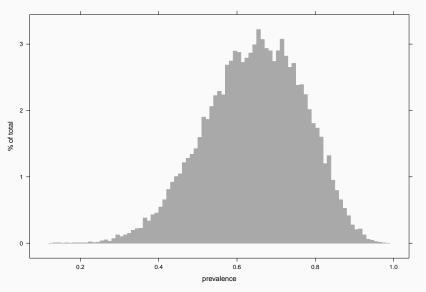
Trace plots: the two chains should be stationary:



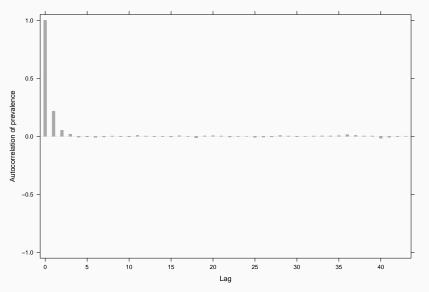
ECDF plots: the two chains should be very close to each other:



Histogram of the combined chains should appear smooth:



Autocorrelation plot tells you how well behaved the model is:



Then check the effective sample size (SSeff) and Gelman-Rubin statistic (psrf):

```
results
   ##
1
     JAGS model summary statistics from 20000 samples (chains = 2;
      adapt+burnin = 6000):
   ##
   ##
              Lower95 Median Upper95 Mean SD Mode
   ## prevalence 0.4063 0.65005 0.87294 0.64338 0.12382 -
6
   ##
                  MCerr MC%ofSD SSeff AC.10
   ##
                                            psrf
   ##
9
   ## Total time taken: 0.2 seconds
10
```

Reminder: we want psrf < 1.05 and SSeff > 1000

Exercise

- Run this model yourself in JAGS
- Change the initial values for the two chains and make sure it doesn't affect the results
- Reduce the burnin length does this make a difference?
- Change the sample length does this make a difference?

Optional Exercise

- Change the number of chains to 1 and 4
 - Remember that you will also need to change the initial values
 - What affect does having different numbers of chains have?
- Try using the run.jags argument method='parallel' what affect does this have?

Session 1b: Working with basic

models (apparent prevalence)

Other runjags options

There are a large number of other options to runjags. Some highlights:

- The method can be parallel or background or bgparallel
- You can use extend.jags to continue running an existing model (e.g. to increase the sample size)
- You can use coda::as.mcmc.list to extract the underlying MCMC chains
- Use the summary() method to extract summary statistics
 - See ?summary.runjags and ?runjagsclass for more information

Using embedded character strings

 For simple models we might not want to bother with an external text file. Then we can do:

```
mt. <- "
1
    model{
      Positives ~ dbinom(prevalence, TotalTests)
      prevalence ~ dbeta(2, 2)
4
 5
      #data# Positives, TotalTests
6
      #monitor# prevalence
      #inits# prevalence
     п
10
11
    results <- run.jags(mt, n.chains=2)
12
```

But I would advise that you stick to using a separate text file!

Setting the RNG seed

• If we want to get numerically replicable results we need to add .RNG.name and .RNG.seed to the initial values, and an additional #modules# lecuyer hook to our basicjags.bug file:

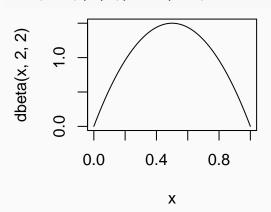
```
1
   model{
     Positives ~ dbinom(prevalence, TotalTests)
     prevalence ~ dbeta(2, 2)
3
4
     #data# Positives, TotalTests
5
      #monitor# prevalence
6
      #inits# prevalence, .RNG.name, .RNG.seed
      #modules# lecuyer
8
    .RNG.name <- "lecuver::RngStream"
1
    .RNG.seed <- list(chain1=1, chain2=2)
   results <- run.jags('basicjags.bug', n.chains=2)
```

Every time this model is run the results will now be identical

A different prior

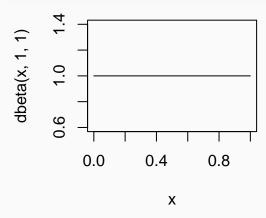
• A quick way to see the distribution of a prior:

curve(dbeta(x, 2, 2), from=0, to=1)



• A minimally informative prior might be:

curve(dbeta(x, 1, 1), from=0, to=1)



• Let's change the prior we are using to dbeta(1,1):

```
model{
model{
model{
positives ~ dbinom(prevalence, TotalTests)
prevalence ~ dbeta(1, 1)

# Hooks for automatic integration with R:
# data# Positives, TotalTests
#monitor# prevalence
#inits# prevalence
}
```

An Equivalent Model

We could equivalently specify an observation-level model:

```
model{
      # Likelihood part:
      for(i in 1:TotalTests){
         Status[i] ~ dbern(prevalence)
6
      # Prior part:
8
      prevalence ~ dbeta(1, 1)
10
      # Hooks for automatic integration with R:
11
      #data# Status, TotalTests
      #monitor# prevalence
12
      #inits# prevalence
13
14
```

But we need the data in a different format: a vector of 0/1 rather than total positives!

```
Status <- c(rep(0, TotalTests-Positives), rep(1, Positives))
```

A GLM Model

```
1
    model{
       # Likelihood part:
       for(i in 1:TotalTests){
3
4
         Status[i] ~ dbern(predicted[i])
         logit(predicted[i]) <- intercept</pre>
5
       }
6
7
       # Prior part:
8
       intercept ~ dnorm(0, 10^-6)
9
10
11
       # Derived parameter:
       prevalence <- ilogit(intercept)</pre>
12
13
14
       # Hooks for automatic integration with R:
       #data# Status, TotalTests
15
16
       #monitor# intercept, prevalence
       #inits# intercept
17
     }
18
```

- This is the start of a generalised linear model, where we could add covariates at individual animal level.
- We introduce a new distribution dnorm() notice this is mean and precision, not mean and sd!
- For a complete list of the distributions available see:
 - https://sourceforge.net/projects/mcmcjags/files/Manuals/4.x/
 - This document is also provided on the GitHub repository
- However, notice that the prior is specified differently...

Exercise

- Run the original version, the observation-level version, and the GLM version of the model and compare results with the same data
- Now try a larger sample size: e.g. 70 positives out of 100 tests are the posteriors from the two models more or less similar than before?
- Now try running the GLM model with a prior of dnorm(0, 0.33) (and the original data) - does this make a difference?

Optional Exercise

Another way of comparing different priors is to run different models with no data - as there is no influence of a likelihood, the posterior will then be identical to the priors (and the model will run faster).

One way to do this is to make all of the response data (i.e. either Positives or Status) missing. Try doing this for the following three models, and compare the priors for prevalence:

- The original model with prior prevalence ~ dbeta(1,1)
- The GLM model with prior intercept ~ dnorm(0, 10^-6)
- The GLM model with prior intercept ~ dnorm(0, 0.33)