# Data Analysis Report - Holy Crop!

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## Abstract

This report explores the effects of land and fertilizer on total crop yield across many different regions. There were many variables considered in the analysis of crop yield. We chose to look at the Entity, amount of fertilizer used, and the amount of arable land available to grow crops and to observe how each of these affected individual crops as well as the total yield of all the cereal crops. Some of the main regions we analyze the effects of are the Americas (North and South), Asia (Southern, Western, and South Eastern), Australia and New Zealand, the Caribbean, Central America, Africa (Eastern, Western, and Southern), Europe (Eastern, Western, and Southern), Oceania, To study these effects, we utilize logistic regression analysis to quantify the effects of these variables on the crop yield. By using our logistic regression model, we are able to compare the changes in yield based on the different factors for any given year as well as predict a potential cereal crop yield for future years based on our data. Through these findings, we can quantify the importance of fertilization management and strategies to increase productivity and yield of crops across many different countries and regions.

After testing many different logistic models, we came to find out that arable land provided the least effect on the crop yield which was very interesting as it was speculated to have been one of the bigger contributing factors to total yield. Fertilizer had the highest positive correlation with respect to the crop yield and had a more dramatic effect on the total amount of crops produced. The crops that had the most significant increase in yield (with respect to the p-value) due to the amount of fertilizer used were rice, potatoes, cassava, and beans.

## Introduction

Crop yield is our main area of concern, and the problem that arises out of this is what conditions and factors can affect it given data recorded from previous years. What regions around the world are affected differently than others, naturally was one of the first areas that we thought to look at. Which crops have the largest impact on the yield overall, also

appeared as we were analyzing the data. How has nitrogen fertilizer affected crop outputs and changed over the years.

To answer these questions, we created multiple logistic regression models to study the strength of each variable's effects on both individual crops and the total crop yield. With these models, we compared the AIC and p-values of each covariate and its effects on the yield. After careful analysis and visualizations of the different variables and their effects, we noticed that our original assumptions about region/Entity having a relatively great effect on crop yield was not as significant as we had first thought. However, fertilizer did in fact have a larger impact on each individual crop yield, as well as the overall cereal yield which confirmed our second assumption. We chose to look at many different factors and covariates for why our original assumption about the region having an impact on crop yield/ overall cereal yield wasn't confirmed and what reasons there could be behind that. We investigated many different models to attempt to see the effects, and we found that the quasi-poisson model most accurately depicts the interactions of multiple variables on the total crop yield.

### Data and Methods

The data that we used was the Global Crop Yields from Our World In Data, through the tidytuesday github. There were five datasets in total and all share the categorical variables of Entity (Country/Region Name), Code (Country Code) and Year. The first dataset was focused on yields of key crops in tonnes per hectare such as wheat, rice, maize, etc. The second dataset included the arable land needed to produce a fixed quantity of crops normalized to 1961. The third dataset included nitrogen fertilizer usage in kilograms per hectare and cereal yields in tonnes per hectare. The fourth dataset focused on cereal yield index with a variable for the change to land area for cereal production since 1961, as well as population. The fifth data set looked at tractors per 100 sq km of arable land, as well as cereal yield in kg per hectare and the total population as well.

We decided to only focus on the first three datasets to cut down on the number of variables, then to clean the data further we filtered out the data to only include larger Entities, reducing our data to large regions/continents so we could use Entity as a categorical variable with a much smaller number of categories. Then after merging the filtered crop yields, arable land and fertilizer datasets we started looking at the year variable. Some Entities did not have values for certain years, so we had to limit our models to only use a certain range of years, where we settled on 1995-2014. There were still some Entities that did not have one specific type of crop, but these were ignored when we made the models.

After cleaning the data, we classify our outcomes as the total yield of crops in kilograms per hectare and denote the observed values as  $y_i$  where i = 1, ..., n, n = 440. We then started searching for models using stepwise regression starting with the full model, removing some variables to minimize the AIC. At this point we split our focus to creating two models, one that focuses on the effects of arable land, Entity and specific crops over time on total yield, and another that looks at the effect of fertilizer on total yield. We also found that a quasi-poisson model was able to show us better models when looking at the residuals

For the Entity model we assumed that each logit( $p_i$ ) is a realization of a random variables  $Y_i \sim \text{Quasi-Poisson}(\lambda_i)$  independently, and model the count of total yield using the following logistic regression with these variables:

$$logit(p_i) = \beta_0 + \beta_1(Entity)_i + \beta_2(Barley)_i + \beta_3(Cocoa)_i$$
 Where  $i = 1, ..., n$  and  $n = 440$ .

For the Year model we assumed that each logit( $p_i$ ) is a realization of a random variables  $Y_i$  ~ Quasi-Poisson( $\lambda_i$ ) independently, and model the count of total yield using the following logistic regression with these variables:

$$\begin{aligned} & \operatorname{logit}(p_i) = \beta_0 + \beta_1 \operatorname{I}(\operatorname{Year})_i + \beta_2 (\operatorname{Wheat})_i + \beta_3 (\operatorname{Maize})_i + \beta_4 (\operatorname{Potatoes})_i + \beta_5 (\operatorname{Cassava})_i + \beta_6 (\operatorname{Barley})_i + \beta_7 (\operatorname{Cocoa})_i \end{aligned}$$

Where i = 1, ..., n, n = 440 and I() in the equation indicates a dummy variable, taking value 1 when the condition is true and zero otherwise.

For the Arable Land model we assumed that each  $logit(p_i)$  is a realization of a random variables  $Y_i \sim Quasi\text{-Poisson}(\lambda_i)$  independently, and model the count of total yield using the following logistic regression with these variables:

$$\begin{aligned} & \operatorname{logit}(p_i) = \beta_0 + \beta_1(\operatorname{Land})_i + \beta_2(\operatorname{Wheat})_i + \beta_3(\operatorname{Maize})_i + \beta_4(\operatorname{Potatoes})_i + \beta_5(\operatorname{Cassava})_i + \\ & \beta_6(\operatorname{Barley})_i + \beta_7(\operatorname{Cocoa})_i \text{ Where } i = 1, ..., n \text{ and } n = 440. \end{aligned}$$

For the Crop model we assumed that each logit( $p_i$ ) is a realization of a random variables  $Y_i$  ~ Quasi-Poisson( $\lambda_i$ ) independently, and model the count of total yield using the following logistic regression with these variables:

$$\begin{array}{lll} \operatorname{logit}(p_i) &=& \beta_0 \,+\, beta_1(\operatorname{Wheat})_i \,+\, \beta_2 \,+\, \left(\operatorname{Maize}\right)_i \,+\, \beta_3(\operatorname{Soybeans})_i \,+\, \beta_4(\operatorname{Potatoes})_i \,+\, \beta_5(\operatorname{Cassava})_i \,+\, \beta_6(\operatorname{Barley})_i \,+\, \beta_7(\operatorname{Cocoa})_i \,\, \operatorname{Where}\,\, i = 1, ..., n \,\, \text{and} \,\, n = 440. \end{array}$$

For the Fertilizer model we assumed that each logit( $p_i$ ) is a realization of a random variables  $Y_i \sim \text{Quasi-Poisson}(\lambda_i)$  independently, and model the count of total yield using the following logistic regression with these variables:

$$\begin{aligned} & \operatorname{logit}(p_i) = \beta_0 + \beta_1(\operatorname{Fertilizer})_i + \beta_2(\operatorname{Wheat})_i + \beta_3(\operatorname{Rice})_i + \beta_4(\operatorname{Maize})_i + \beta_5(\operatorname{Soybeans})_i + \\ & \beta_6(\operatorname{Potatoes})_i + \beta_7(\operatorname{Beans})_i + \beta_8(\operatorname{Cassava})_i + \beta_9(\operatorname{Barley})_i + \beta_{10}(\operatorname{Cocoa})_i + \beta_{11}(\operatorname{Bananas})_i \end{aligned}$$
 Where  $i = 1, ..., n$  and  $n = 440$ .

We also took a look at correlation between each of our variables and the yield. The resulting graphs and values showed an extreme amount of correlation between them all. There were a few instances of insignificant correlation, primarily cocoa beans when compared to all other variables, except against arable land, wheat, and bananas.

## Results

#Individual Modeling

Overall, we modeled yield as our response variable. We first started with individual variables to determine which might have the largest effect on crop yield. (Findings were excluded if they didn't provide significant effects, i.e. p < 0.05.)

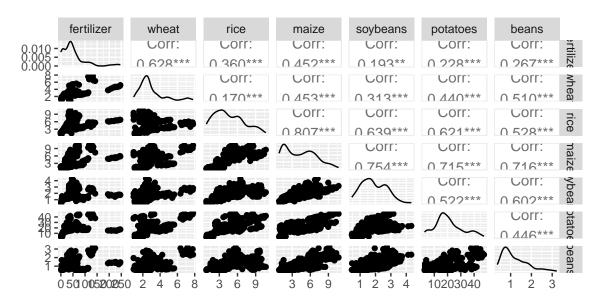


Figure 1: This graph shows and example of the correlation plot values, this particular one has the variables Fertilizer, Wheat, Rice, Maize, Soybeans, Potatoes, and Beans. Basically, all other coordinate plots look similar to this one.

### Year

Modeling year as a factor variable, we found that a majority of years had significant effects on yield in the quasi-poisson model (negated 1998, 2000, 2005, 2006, 2013, 2014). We did find that the biggest increase in yield came in 2007 on a logistic scale — a change from 1995 to 2007 corresponds to a 1.108-multiplicative odds increase in crop yield. 2001 and 2003 follow closely behind, with 1.099 and 1.043-multiplicative odds increases in yield, respectively. There were no significant multiplicative-odds decreases. Entity

# Entity

We decided to not have a baseline for the model consisting solely of entity as the covariate — Africa would have been the baseline, but any other continent with respect to this simply didn't make sense, and we were not getting any significant effects.

Removing Africa as a baseline yielded significant effects for every region studied — this was one of just two models overall which we chose to stick with poisson models instead of quasi-poisson because the dispersion parameter was close enough to 1 that it didn't really make a difference on estimates nor p-values.

Micronesia yielded a 1.147-multiplicative-odds increase in crop yield (remember, there's no baseline so it's not exactly with respect to any other entity). How we interpreted this was it was simply increasing from zero.

## Arable land

With arable land, we also removed the baseline because it is difficult to determine what the baseline would even be — maybe non-arable land? When we removed this, we found that there was a significant effect and a 1.409-multiplicative-odds increase in crop yield.

## Crops

With each crop, we originally had a baseline, but, similarly to arable land, it was difficult to determine what that was. In addition, all the crops were giving us insignificant effects, leading us to determine removing the baseline was the best way to go. Again, these are just individual crops' effects on yield, these are by no means our best models.

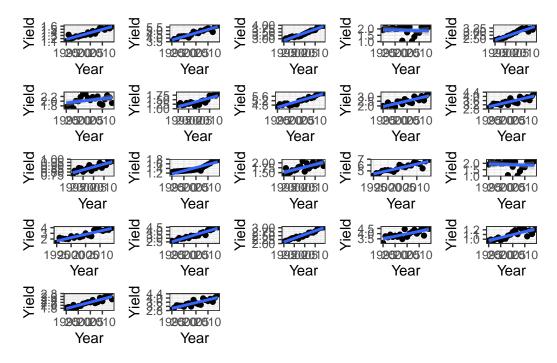
We found that cocoa beans had the highest multiplicative-odds increase of 2.18. This makes sense because, as you will see in other models, this was part of our best crops model. Bananas, on the other hand, had the least positive multiplicative-odds increase, with just 0.036, on yield, which also makes sense because this fruit rarely appeared in any of our good models.

## **Interaction Effects**

# Entity\*Year

With Africa as a baseline, there was not a single significant effect. Removing the baseline of Africa did not change much for the interaction effects. There were still no significant effects among interactions — all the significant effects with the regular variables were described above.

Year vs. Yield plots by Entity



Most of these plots are fairly similar in their slope and correlation between yield and year over time. We found that the greatest correlation in the logistic regression between year and yield came with Western Europe by value, but Northern America by increasing slope. The one with the smallest correlation came with Australia and New Zealand, perhaps indicating that not a lot of farming goes on here with respect to the rest of the world.

Summary of interaction effects with entity and crops:

Entity interaction with wheat: A one-unit change in wheat produced in Western Africa with respect to wheat produced in Africa corresponds to a 0.709-multiplicative-odds decrease in crop yield.

Entity interaction with rice: A one-unit change in rice produced in the Caribbean with respect to rice produced in Africa corresponds to a 0.9745-multiplicative-odds decrease in crop yield.

Entity interaction with maize: A one-unit change in maize produced in Australia & New Zealand with respect to maize produced in Africa corresponds to a 1.58-multiplicative-odds decrease in crop yield.

Entity interaction with soybeans: A one-unit change in soybeans produced in Asia with respect to soybeans produced in Africa corresponds to a 1.58-multiplicative-odds decrease in crop yield.

Entity interaction with potatoes: A one-unit change in potatoes produced in Oceania with respect to potatoes produced in Africa corresponds to a 0.184-multiplicative-odds decrease in crop yield.

Entity interaction with beans: A one-unit change in beans produced in Australia &

New Zealand or Oceania with respect to beans produced in Africa corresponds to a 3.26-multiplicative-odds decrease in crop yield.

Entity interaction with peas: A one-unit change in peas produced in Europe with respect to peas produced in Africa corresponds to a 2.35-multiplicative-odds decrease in crop yield.

Entity interaction with cassava: A one-unit change in cassava produced in Northern Africa with respect to cassava produced in Africa corresponds to a 2.17-multiplicative-odds decrease in crop yield.

Entity interaction with barley: A one-unit change in barley produced in Western Africa with respect to barley produced in Africa corresponds to a 0.81-multiplicative-odds decrease in crop yield.

Entity interaction with cocoa\_beans: A one-unit change in cocoa beans produced in Microne-sia with respect to cocoa beans produced in Africa corresponds to a 15.56-multiplicative-odds decrease in crop yield.

Entity interaction with bananas: A one-unit change in bananas produced in Europe with respect to bananas produced in Africa corresponds to a 0.198-multiplicative-odds decrease in crop yield.

# Year\*Crop

There were no significant effects found when the year and crop were acting together, interestingly enough. Intuitively, this may make sense in a way, just because if you're only looking at the year, perhaps the total yield isn't changing that much with each crop on an annual basis.

The residual plot shows a few things about our model. The first thing to notice is that the values seem to have a constant dispersion meaning that we do have homoscedasticity. There are a few areas where the distribution is a little off of constant, but as a whole. There are, however, a few problems that need to be addressed with the residual plots. The first being a clear pattern in the residual plot, this shows that the relationship between variables and predictors are not linear, it can also mean that the data is not well modeled. Since this exact same pattern repeats over every graph, it seems that the relationship between variables and predictors is more likely than an unfit model. The final problem again occurs because of the bands seen in the plots, it also means that we have not effectively found the underlying cause of the relationship. While these do not bode well for our models, it is imperative for us to state that these graphs are an improvement from previous models we used. We used the Poisson distribution at first and found all the same problems as previously mentioned, but also the dispersion was not constant as well. In looking at these problems we changed to the Quasi-Poisson distribution which fixed our dispersion issue. The QQ-Plots for each model all can be safely assumed to have normal distributions. There are a couple of spots in the plot that trail off the main line a little bit, but generally they stay on the line.

This is the best fitting model using Entity as one of the variables.

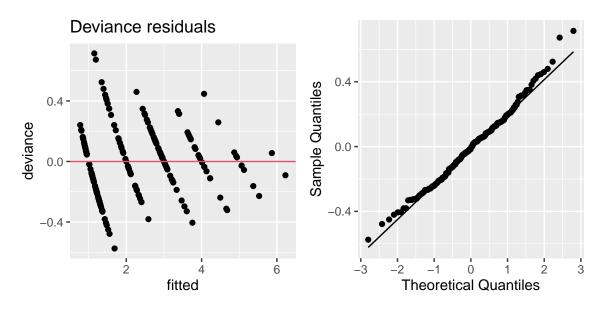


Figure 2: The left graph shows the deviance residuals for Year vs. fitted values and the right shows the Q-Q Normal Plot of the deviance of Year.

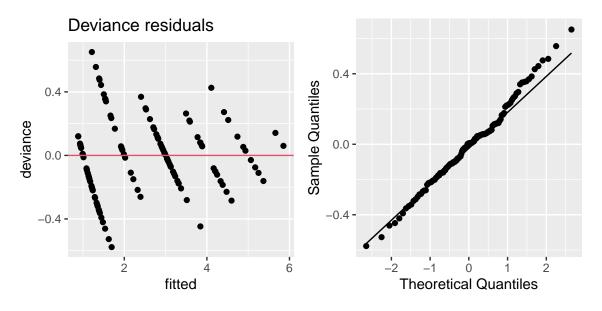


Figure 3: The left graph shows the deviance residuals for Fertilizer vs. fitted values and the right shows the Q-Q Normal Plot of the deviance of Fertilizer.

### Standard errors: MLE

Est.	S.E.	t val.	p
0.08	0.08	1.03	0.30
0.64	0.08	8.00	0.00
0.88	0.06	14.89	0.00
0.45	0.07	6.62	0.00
-0.07	0.06	-1.08	0.28
-0.13	0.07	-1.90	0.06
0.09	0.06	1.39	0.17
0.53	0.07	7.63	0.00
0.45	0.06	7.61	0.00
-0.52	0.08	-6.38	0.00
0.33	0.03	10.57	0.00
-0.47	0.12	-3.82	0.00
	0.08 0.64 0.88 0.45 -0.07 -0.13 0.09 0.53 0.45 -0.52 0.33	0.08 0.08 0.64 0.08 0.88 0.06 0.45 0.07 -0.07 0.06 -0.13 0.07 0.09 0.06 0.53 0.07 0.45 0.06 -0.52 0.08 0.33 0.03	0.08 0.08 1.03 0.64 0.08 8.00 0.88 0.06 14.89 0.45 0.07 6.62 -0.07 0.06 -1.08 -0.13 0.07 -1.90 0.09 0.06 1.39 0.53 0.07 7.63 0.45 0.06 7.61 -0.52 0.08 -6.38 0.33 0.03 10.57

Estimated dispersion parameter = 0.05

This model uses Year in the model with the best fitting variables to go with it.

## Standard errors: MLE

	Est.	S.E.	t val.	p
(Intercept)	-0.94	0.09	-10.13	0.00
as.factor(Year)1996	-0.09	0.08	-1.15	0.25
as.factor(Year)1997	-0.07	0.08	-0.88	0.38
as.factor(Year)1998	-0.18	0.08	-2.33	0.02
as.factor(Year)1999	-0.13	0.08	-1.73	0.09
as.factor(Year)2000	-0.09	0.08	-1.15	0.25
as.factor(Year)2001	-0.13	0.08	-1.66	0.10
as.factor(Year)2002	-0.11	0.08	-1.49	0.14
as.factor(Year)2003	-0.21	0.08	-2.76	0.01
as.factor(Year)2004	-0.14	0.08	-1.79	0.08
as.factor(Year)2005	-0.11	0.07	-1.49	0.14
as.factor(Year)2006	-0.15	0.08	-2.04	0.04
as.factor(Year)2007	-0.18	0.07	-2.36	0.02
as.factor(Year)2008	-0.20	0.07	-2.69	0.01
as.factor(Year)2009	-0.10	0.07	-1.35	0.18
as.factor(Year)2010	-0.15	0.07	-2.02	0.05
as.factor(Year)2011	-0.19	0.07	-2.59	0.01
as.factor(Year)2012	-0.13	0.07	-1.84	0.07

as.factor(Year)2013	-0.08	0.07	-1.16	0.25
as.factor(Year)2014	-0.17	0.07	-2.34	0.02
wheat	0.28	0.02	14.75	0.00
maize	0.18	0.01	12.18	0.00
soybeans	0.32	0.04	8.26	0.00
potatoes	-0.03	0.00	-11.13	0.00
cassava	0.03	0.00	16.98	0.00
barley	-0.02	0.03	-0.62	0.54
cocoa	0.27	0.11	2.53	0.01

## Estimated dispersion parameter = 0.06

This is the best fitting model using Arable Land as one of the variables.

#### Standard errors: MLE

	Est.	S.E.	t val.	р
(Intercept)	-1.92	0.16	-12.08	0.00
land	1.58	0.19	8.36	0.00
wheat	0.36	0.02	17.99	0.00
maize	0.26	0.01	19.57	0.00
potatoes	-0.03	0.00	-11.34	0.00
cassava	0.04	0.00	16.41	0.00
barley	0.06	0.03	2.05	0.04
cocoa	0.34	0.11	3.17	0.00

Estimated dispersion parameter = 0.06

This model uses Crops as the variables without any others, here is what fitetd the data the best.

#### Standard errors: MLE

	Est.	S.E.	t val.	р
(Intercept)	-1.02	0.08	-13.52	0.00
wheat	0.28	0.02	15.07	0.00
maize	0.18	0.01	12.43	0.00
soybeans	0.31	0.04	8.59	0.00
potatoes	-0.03	0.00	-11.16	0.00
cassava	0.03	0.00	16.91	0.00

barley	-0.02	0.03	-0.61	0.54
cocoa	0.25	0.10	2.45	0.02

Estimated dispersion parameter = 0.06

This is the best fitting model using Fertilizer as one of the variables.

#### Standard errors: MLE

	Est.	S.E.	t val.	р
(Intercept)	-0.58	0.17	-3.33	0.00
fertilizer	0.00	0.00	1.97	0.05
wheat	0.24	0.04	5.46	0.00
rice	-0.01	0.02	-0.56	0.58
maize	0.15	0.03	4.87	0.00
barley	0.05	0.05	0.93	0.35
soybeans	0.27	0.05	5.54	0.00
potatoes	-0.02	0.01	-3.98	0.00
beans	-0.14	0.10	-1.42	0.16
cassava	0.02	0.01	2.61	0.01
cocoa	-0.10	0.19	-0.54	0.59
bananas	-0.00	0.00	-0.50	0.62

Estimated dispersion parameter = 0.06

In this section, discuss and interpret your results. Explain the implications and significance of your findings, and relate them to your research question and existing literature or theories.

## Conclusion

```
Analysis of Deviance Table
```

```
Model 1: round(Yield) ~ wheat + maize + soybeans + potatoes + cassava +
    barley + cocoa
Model 2: round(Yield) ~ land + wheat + maize + potatoes + cassava + barley +
    cocoa
Model 3: round(Yield) ~ as.factor(Year) + wheat + maize + soybeans + potatoes +
    cassava + barley + cocoa
Model 4: round(Yield) ~ Entity + barley + cocoa
    Resid. Df Resid. Dev Df Deviance
```

```
1
         186
                10.5148
2
                 10.5235
         186
                            0 - 0.00873
3
         167
                 9.4761
                          19
                               1.04738
4
         182
                 9.0465 -15
                               0.42959
```

Throughout our analysis of the effects of different variables on the crop yield, we found that there is a high correlation between the interactions of entity, year, land, fertilizer, and the region. Year and fertilizer we found have the lowest correlation. Fertilizer having one of the lower correlations was surprising because our original thoughts were that land and fertilizer would have the greatest effects on the crop yield (both individually and for total cereal yield). However, our correlation plots did confirm our original question which was does the region and amount of arable land have a large impact on crop yield?

Although we may have previously thought that all of our covariates would have a significantly high effect on crop yield, we were somewhat mistaken. The quasi-poisson distribution does not yield any AIC values because the distribution doesn't depend on the same likelihood function as other distributions do. Therefore, before realizing that the quasi-poisson model yielded better results to us than the poisson model, we were using AIC values from the poisson model as a way to judge the best ones.

We did find that based on our best models (which have been noted in our R chunks), arable land had the weakest relationship and effect on the yield — this somewhat makes sense relative to all the other covariates because we would intuitively imagine that year or region or crop helps give the best correlation to the yield.

We can, in fact, see in our year vs. yield plots by entity, that there is a positive correlation in the logistic regression fitted line for almost every entity. Annual change over time notably distinguishes a positive correlation between the two variables, and this was consistent throughout almost all entities except a couple.

Based on our best models, since we cannot use AIC to compare our models, we use the residual deviance as the next-best alternative to determine the best models. We find that our best entity model — with entity, barley and cocoa beans as covariates — was our best overall model to estimate crop yield, with a residual deviance of 9.0465. This fits perfectly with what we previously said about cocoa beans yielding the highest multiplicative-odds increase when doing individual modeling. Barley was also determined to be a good factor in modeling the highest change in yield — both of these crop covariates were in all of our best models.

The next-best model we had was our year model, which had a residual deviance of 9.4761; followed by our best crops model: 10.5148 residual deviance; and followed by our best arable land model: 10.5235 residual deviance.

Further deductions reveal that arable land, as we had concluded, was our worst predictor, followed by year. We can also deduce that no interaction effects is, a lot of the time, better than having interaction effects because this makes more parameters, increasing both the AIC and residual deviance.

Some factors that could also have an effect on the crop yield would be the weather of specific regions, which is something our data did not provide. In future studies it would be beneficial to get weather data to look at that covariate alongside the arable land and region of where the crops are being grown. We believe that would have a more significant impact on each specific crop's yield. Based on our findings, we can say that while fertilizer can be beneficial to crop yield, the biggest factors when considering where to grow certain crops heavily depends on the region/entity and the types of crops produced.

#### References

jthomasmack. "Tidytuesday/Data/2020/2020-09-01/Readme.Md at Master · Rfordata-science/Tidytuesday." GitHub, 1 Sept. 2020, github.com/rfordatascience/tidytuesday/blob/master/data/209-01/readme.md.

## **Appendix**

```
(bananas.mod <- summary(glm(round(Yield) ~ bananas -1, data = merged_data, family = "d
(best rice <- summary(glm(round(Yield) ~ wheat + rice + maize + soybeans + potatoes +
(best_beans <- summary(glm(round(Yield) ~ wheat + maize + soybeans + potatoes + beans
(best peas <- summary(glm(round(Yield) ~ peas + wheat + maize + soybeans + potatoes +
(best_bananas <- summary(glm(round(Yield) ~ wheat + rice + maize + soybeans + potatoes
summary(glm(round(Yield) ~ Entity*as.factor(Year) - 1, data = merged data, family = "]
summary(glm(round(Yield) ~ Entity*wheat, data = merged_data, family = "quasipoisson"))
summary(glm(round(Yield) ~ Entity*rice, data = merged data, family = "quasipoisson"))
summary(glm(round(Yield) ~ Entity*maize, data = merged_data, family = "quasipoisson"))
summary(glm(round(Yield) ~ Entity*soybeans, data = merged data, family = "quasipoisson
summary(glm(round(Yield) ~ Entity*potatoes, data = merged_data, family = "quasipoisson")
summary(glm(round(Yield) ~ Entity*beans, data = merged data, family = "quasipoisson"))
summary(glm(round(Yield) ~ Entity*peas, data = merged data, family = "quasipoisson"))
summary(glm(round(Yield) ~ Entity*cassava, data = merged data, family = "quasipoisson")
summary(glm(round(Yield) ~ Entity*barley, data = merged_data, family = "quasipoisson")
summary(glm(round(Yield) ~ Entity*cocoa, data = merged data, family = "quasipoisson"))
summary(glm(round(Yield) ~ Entity*bananas, data = merged data, family = "quasipoisson"
summary(glm(round(Yield) ~ as.factor(Year)*wheat, data = merged_data, family = "quasi
summary(glm(round(Yield) ~ as.factor(Year)*rice, data = merged data, family = "quasipe")
summary(glm(round(Yield) ~ as.factor(Year)*maize, data = merged_data, family = "quasi
summary(glm(round(Yield) ~ as.factor(Year)*soybeans, data = merged data, family = "quantum data")
summary(glm(round(Yield) ~ as.factor(Year)*potatoes, data = merged_data, family = "quantum data")
```

```
summary(glm(round(Yield) ~ as.factor(Year)*beans, data = merged_data, family = "quasi
summary(glm(round(Yield) ~ as.factor(Year)*peas, data = merged data, family = "quasipe
summary(glm(round(Yield) ~ as.factor(Year)*cassava, data = merged_data, family = "quas
summary(glm(round(Yield) ~ as.factor(Year)*barley, data = merged data, family = "quas;
summary(glm(round(Yield) ~ as.factor(Year)*cocoa, data = merged_data, family = "quasi
summary(glm(round(Yield) ~ as.factor(Year)*bananas, data = merged data, family = "quas
entity1 = merged_data %>% summarise(entity = Entity, wheat = wheat, rice = rice, maize
plot1 = ggpairs(entity1, cardinality threshold = 24, progress = F)
entity2 = merged_data %>% summarise(entity = Entity, peas = peas, cassava = cassava, )
ggpairs(entity2, cardinality_threshold = 24, progress = F)
year1 = merged data %>% summarise(year = Year, wheat = wheat, rice = rice, maize = mai
ggpairs(year1, cardinality_threshold = 24, progress = F)
year2 = merged data %>% summarise(year = Year, peas = peas, cassava = cassava, barley
ggpairs(year2, cardinality threshold = 24, progress = F)
fer1 = merged_data %>% summarise(fertilizer = fertilizer, wheat = wheat, rice = rice,
ggpairs(fer1, cardinality threshold = 24, progress = F)
fer2 = merged_data %>% summarise(fertilizer = fertilizer, peas = peas, cassava = cassava
ggpairs(fer2, cardinality threshold = 24, progress = F)
yield1 = merged data %>% summarise(yield = Yield, wheat = wheat, rice = rice, maize =
ggpairs(yield1, cardinality_threshold = 24, progress = F)
yield2 = merged data %>% summarise(yield = Yield, peas = peas, cassava = cassava, bar
ggpairs(yield2, cardinality_threshold = 24, progress = F)
land1 = merged data %>% summarise(land = land, wheat = wheat, rice = rice, maize = max
ggpairs(land1, cardinality threshold = 24, progress = F)
land2 = merged_data %>% summarise(land = land, peas = peas, cassava = cassava, barley
ggpairs(land2, cardinality threshold = 24, progress = F)
main1 = merged_data %>% summarise(land = land, entity = Entity, year = Year, fertilize
```

```
ggpairs(main1, cardinality_threshold = 24, progress = F)
crop1 = merged data %>% summarise( peas = peas, cassava = cassava, barley = barley, se
ggpairs(crop1, cardinality threshold = 24, progress = F)
crop2 = merged_data %>% summarise( wheat = wheat, rice = rice, maize = maize, cocoa =
ggpairs(crop2, cardinality threshold = 24, progress = F)
crop3 = merged_data %>% summarise( peas = peas, cassava = cassava, barley = barley, wl
ggpairs(crop3, cardinality threshold = 24, progress = F)
crop4 = merged_data %>% summarise(soybeans = soybeans, potatoes = potatoes, beans = be
ggpairs(crop4, cardinality_threshold = 24, progress = F)
(top crops <- (glm(round(Yield) ~ wheat + maize + soybeans + potatoes + cassava + bar
fitted = top_crops$fitted.values
d = resid(top crops, type = "deviance")
resid df6 = data.frame(fitted = fitted, deviance = d)
ggplot(resid_df6, aes(x = fitted, y = deviance)) +
       geom point() +
       geom hline(yintercept = 0, col = 2) +
       labs(title = "Deviance residuals")
ggplot(resid df6, aes(sample = deviance)) + stat qq() + stat qq line() + labs(x = "Theorem 1) + stat qq() + stat
(best_air <- glm(round(Yield) ~ land + wheat + maize + potatoes + cassava + barley + c
fitted = best_air$fitted.values
d = resid(best air, type = "deviance")
resid_df5 = data.frame(fitted = fitted, deviance = d)
ggplot(resid df5, aes(x = fitted, y = deviance)) +
       geom_point() +
       geom_hline(yintercept = 0, col = 2) +
       labs(title = "Deviance residuals")
ggplot(resid_df5, aes(sample = deviance)) + stat_qq() + stat_qq_line() + labs(x = "Theorem 1) + stat_qq() + stat_qq_line() + labs(x = "Theorem 2) + stat_qq() + stat_qq_line() + labs(x = "Theorem 3) + stat_qq_line() + labs(x = "Theorem 3) + stat_qq_line() + labs(x = "Theorem 3) + stat_qq_line() + stat_qq_line() + labs(x = "Theorem 3) + stat_qq_line() + stat_qq_line() + labs(x = "Theorem 3) + stat_qq_line() + s
(best entity <- glm(round(Yield) ~ Entity + barley + cocoa, data = merged data, family
fitted = best_entity$fitted.values
```

```
d = resid(best_entity, type = "deviance")
resid_df3 = data.frame(fitted = fitted, deviance = d)

ggplot(resid_df3, aes(x = fitted, y = deviance)) +
    geom_point() +
    geom_hline(yintercept = 0, col = 2) +
    labs(title = "Deviance residuals")

ggplot(resid_df3, aes(sample = deviance)) +stat_qq() + stat_qq_line() + labs(x = "Thee
```