

# Optimal parsing of LZ78 for memoryless sources

## Corrections

- The formal definition of  $D_n^{\text{LZ}}$  in (3) is misleading because it contradicts the previous verbal definition. For  $w$  a word of size  $n$ , the formula

$$\frac{1}{M_n(w)} \sum_{j=0}^{M_n(w)-1} |u_j^{\text{LZ}}|$$

would rather be the empirical average length of a phrase in the Lempel-Ziv parsing of a word  $\bar{D}_n^{\text{LZ}}$ , whereas  $D_n^{\text{LZ}}$  is used in the rest of the paper as the length of a randomly selected phrase. These two aren't equal: if we build a Lempel-Ziv DST from a word, then  $D_n^{\text{LZ}}$  can be seen as the depth of a random node, which is different from the average path length computed on all the nodes.

- In *Remark 2*, I think the definition of  $v$  and  $t$  should rather be  $v = a_{i'} \dots a_i$  and  $t = a_{i+1} \dots a_n$ .
- The result (14) should be an equality, and it is one indeed because of the flexible parsing algorithm. A proof by contradiction can show this. However, since we upperbound  $g(j)$  by  $+\infty$  in (23) there might be no purpose to using (14) instead of just (13).
- The proof around *Theorem 1* has several flaws. The notation  $X$  for a sequence depending on  $n$  and not a random variable is misleading. On the other hand, it should appear that both  $g$  and  $j$  are random variables, as the randomness of  $j$  is used in the end of the proof. I wrote some possible definitions [here](#), and applied them to make some computations that seemed otherwise incorrect because of their use of randomness outside of a probability measure ([here](#)).
- As for the arguments that link  $|L_{g(j)-k}|$  to  $D_n^{\text{LZ}}$ , I have indicated how I think they could be developed in [this part](#). These arguments are the most controversial part right now I think.
- *Theorem 2* is false as stated: we proved *Theorem 1* using a random  $j$ . The randomness remains, so the quantifier 'for any  $j < M_n$ ' should be removed. This would be true for *Theorem 1* too.
- The proof of *Theorem 2* may stop at (26) since we can directly prove this upperbound goes to 0. This yields a tighter upperbound for *Theorem 2*. I detailed this analysis in the [last part of this report](#).
- In that same proof, the step between (25) and (26) relies heavily on a result from [6]. A bit more context on this result (and why it does apply here) would make things clearer.
- The conclusion claims to use *Cramer's* theorem to link

$$\max_{0 \leq k \leq g_w(J)} \{L_{g_w(J)-k} - k\}$$

to  $D_n^{\text{FLEX}}$ , which is a sum of random variables. Since *Cramer* applies to independent random variables and the lengths of successive phrases are not independent, something must be missing there.

## Notations

These are definitions and notations in order to write the proof of *Theorem 1*.

**Definition 1** For all  $n \in \mathbb{N}$ , calling  $\Omega_n$  the set of words of length  $n$ .

**Definition 2** Defining  $W \in \Omega_n$  to be a random variable which outputs words of length  $n$  from a memoryless source.

**Definition 3** Considering  $J \in \mathbb{N}^*$  to be a random variable which, in the event  $\{W = w\}$ , uniformly randomly picks the index of one of the phrases of  $w$ . The joint law of  $J$  with  $W$  being:

$$PW = w, J = j = \begin{cases} 1/M_n(w) & \text{if } j \leq M_n(w) \\ 0 & \text{else} \end{cases}$$

**Remark 1** We might choose another randomness for  $J$ , but this one seems more natural.

**Definition 4** For a given word  $w \in \Omega_n$ , and for all  $i \in \mathbb{N}^*$ , we consider  $g_w(i)$  defined by

$$g_w(i) = f_w(i) + |L_{f(i)}|$$

where  $f_w(i)$  is the starting index of the  $i^{\text{th}}$  phrase of the flexible parsing of  $w$ , and  $|L_{f_w(i)}|$  is the length of the longest greedy phrase given by the Lempel-Ziv parsing of this same word.

**Definition 5** We define  $g_w(J)$  to be the random variable which outputs  $g_w(j)$  during the events  $\{W = w\}$  and  $\{J = j\}$ .

**Definition 6** For all  $k \in \mathbb{N}$ , let  $L_{g_w(J)-k}$  be the random variables which gives the  $(k+1)^{\text{th}}$  possible phrase for the flexible parsing at index  $g_w(J)-k$ . Its only randomness comes from  $W$  and  $J$ . If  $i \leq 0$ , we might assume that  $L_i$  will be the empty word of size 0.

**Notation 1** Denoting by

$$B_{J,W}^k = |L_{g_w(J)-k}|$$

the length of this  $(k+1)^{\text{th}}$  candidate.

We can now study the random variable

$$\max_{0 \leq k \leq g_w(J)} \{B_{J,W}^k - k\}$$

Given any  $(j, w) \in \mathbb{N}^* \times \Omega_n$ , under the events  $\{J = j\}$ ,  $\{W = w\}$  and  $\{J \leq M_n(W)\}$ , this random variable is the length of the  $j^{\text{th}}$  phrase of the flexible parsing of the word  $w$ .

**Definition 7** Defining the random variable  $D_n^{\text{FLEX}}$  as

$$\begin{aligned} D_n^{\text{FLEX}}(w) &= \frac{1}{M_n(w)} \sum_{j=0}^{M_n(w)-1} |u_j^{\text{FLEX}}(w)| \\ &= \frac{1}{M_n(w)} \sum_{j=0}^{M_n(w)-1} \max_{0 \leq k \leq g_w(j)} \{B_{j,w}^k - k\} \end{aligned}$$

where  $D_n^{\text{FLEX}}(w)$  is the empirical average of the lengths of the phrases of the flexible parsing of a word  $w$ , contrary to  $\max_{0 \leq k \leq g_w(J)} \{B_{J,w}^k - k\}$ , with  $J$  a random variable and  $w$  fixed, which is the length of a uniformly randomly selected phrase of the flexible parsing of  $w$ .

**Definition 8** We denote  $x_n$  the average value of  $D_n^{LZ}$  :

$$x_n = \frac{1}{h} \log_2 \left( \frac{nh}{\log_2(n)} \right)$$

**Remark 2** The random variable  $D_n^{LZ}$  is the length of a phrase randomly taken from the Lempel-Ziv parsing of a memoryless-generated word. It is not the same as the empirical average length of a phrase, which we can denote, for  $w \in \Omega_n$  :

$$\overline{D}_n^{LZ}(w) = \frac{1}{M_n(w)} \sum_{j=0}^{M_n(w)-1} |u_j^{FLEX}|$$

**Notation 2** We denote  $b_n^\delta$  as

$$b_n^\delta = x_n + (c_3 x_n)^\delta$$

We will study

$$P \max_{0 \leq k \leq g_W(J)} \{B_{J,W}^k - k\} > b_n^\delta$$

**Computations**

With these definitions, we can do the formal computations at the beginning of the proof of *Theorem 1*. By conditioning on  $W$  and  $J$ , we obtain

$$\begin{aligned} P \max_{0 \leq k \leq g_W(J)} \{B_{J,W}^k - k\} > b_n^\delta &= \sum_{w \in \Omega_n} \sum_{j \in \mathbb{N}^*} P \max_{0 \leq k \leq g_W(J)} \{B_{J,W}^k - k\} > b_n^\delta, W = w, J = j \\ &= \sum_{w \in \Omega_n} \sum_{j \in \mathbb{N}^*} P \bigcup_{k=0}^{g_w(j)} \{B_{w,j}^k > k + b_n^\delta\}, W = w, J = j \\ &\leq \sum_{w \in \Omega_n} \sum_{j \in \mathbb{N}^*} \sum_{k=0}^{g_w(j)} P B_{w,j}^k > k + b_n^\delta, W = w, J = j \\ &\leq \sum_{w \in \Omega_n} \sum_{j \in \mathbb{N}^*} \sum_{k=0}^{+\infty} P B_{w,j}^k > k + b_n^\delta, W = w, J = j \\ &= \sum_{k=0}^{+\infty} \sum_{w \in \Omega_n} \sum_{j \in \mathbb{N}^*} P B_{w,j}^k > k + b_n^\delta, W = w, J = j \\ &= \sum_{k=0}^{+\infty} P B_{W,J}^k > k + b_n^\delta \end{aligned}$$

For all  $k \in \mathbb{N}$ , we may now prove that

$$P B_{J,W}^k > k + b_n^\delta \leq P D_n^{LZ} > k + b_n^\delta$$

Let  $k \in \mathbb{N}$ . Currently, the proof to show this stands on three arguments :

- (1) The first is that  $L_{g_W(J)-k}$  is a random phrase from the Lempel-Ziv parsing of a word of length  $N$ , where  $N \leq g_W(J) \leq n$ . In the event  $\{N = n'\}$ , we consider  $D_{n'}^{LZ}$ .
- (2) The second, is that  $|L_{g_W(J)-k}|$  can therefore be considered the same as  $D_N^{LZ}$  i.e at least equal in law.
- (3) The third is that, for all  $n' \leq n$ ,  $D_{n'}^{LZ} \leq D_n^{LZ}$ .

Although they seem generally true, there are different problems with each of these arguments :

- N isn't clearly established, so  $D_N^{\text{LZ}}$  isn't really known.
- If we can identify N and, let's say, condition our probability with  $\{N = n'\}$ , it is not obvious that the choice of a phrase at position  $g_W(J) - k$  knowing  $\{N = n'\}$  is the same as the uniform choice that operates when choosing a random phrase from a word of size  $n'$ , in  $D_{n'}^{\text{LZ}}$ .
- As for (3), this result is true on average, but not in all cases. Indeed, since  $D_n^{\text{LZ}}$  (resp.  $D_{n'}^{\text{LZ}}$ ) is concentrated around  $x_n$  (resp.  $x_{n'}$ ), and  $x_{n'} < x_n$  since  $n' < n$ , we can show that this result holds with high probability. To write the proof, we may condition using events of the type

$$\{|D_n^{\text{LZ}} - x_n| \leq k_n^{(1)} v_n\}$$

and

$$\{|D_{n'}^{\text{LZ}} - x_{n'}| \leq k_n^{(2)} v_{n'}\}$$

where

$$v_n = \sqrt{\log(nh/\log(n))}$$

A sketch of the proof is to apply concentration inequalities to these events while picking  $(k_n^{(1)}, k_n^{(2)})$  such that

$$k_n^{(1)} v_n + k_n^{(2)} v_{n'} < x_n - x_{n'}$$

and having  $k_n^{(1)}$  and  $k_n^{(2)}$  go to  $+\infty$  for  $n$  going to  $+\infty$  in order for the upper-bound probability to converge to zero.

**Upperbound proof**

This is a proof that the upperbound in (26) goes to 0. Assuming

$$\sum_{k=0}^{+\infty} \text{PD}_n^{\text{LZ}}(W) > k + b_n^\delta \leq A \alpha^{(c_3 x_n)^{\delta-1/2}} \sum_{i=0}^{+\infty} \alpha^{i/\sqrt{c_3 x_n}} \quad (26)$$

We can prove directly that the upperbound term goes to 0 for  $n$  going to  $+\infty$ , without resorting to another majoration. Since  $\sqrt{c_3 x_n}$  goes to infinity for  $n \rightarrow +\infty$ , we can pick  $n$  such that  $\sqrt{c_3 x_n} > 1$ . Therefore  $\alpha^{1/\sqrt{c_3 x_n}} < 1$  and the geometric sum gives

$$\sum_{i=0}^{+\infty} \alpha^{i/\sqrt{c_3 x_n}} = \frac{1}{1 - \alpha^{1/\sqrt{c_3 x_n}}}$$

It remains to prove that

$$\lim_{n \rightarrow +\infty} \frac{\alpha^{(c_3 x_n)^{\delta-1/2}}}{1 - \alpha^{1/\sqrt{c_3 x_n}}} = 0$$

which is done by using L'Hospital's rule. Define :

$$f: \begin{cases} \mathbb{R}_+^* \longrightarrow \mathbb{R} \\ x \longmapsto \alpha^{x^{\delta-1/2}} \end{cases} \quad \text{and} \quad g: \begin{cases} \mathbb{R}_+^* \longrightarrow \mathbb{R} \\ x \longmapsto 1 - \alpha^{\frac{1}{\sqrt{x}}} \end{cases}$$

Let  $x \in ]0; +\infty[$ . Derivating yields :

$$f'(x) = \ln \alpha \left( \delta - \frac{1}{2} \right) x^{\delta-3/2} f(x) \quad \text{and} \quad g'(x) = \ln \alpha \frac{1}{2x\sqrt{x}} \alpha^{\frac{1}{\sqrt{x}}}$$

We proceed to show that  $\frac{f'(x)}{g'(x)}$  goes to 0 as  $x$  goes to  $+\infty$ . We have

$$\frac{f'(x)}{g'(x)} = \frac{\left(\delta - \frac{1}{2}\right) x^{\delta-3/2} \alpha^{x^{\delta-1/2}}}{\frac{1}{2x\sqrt{x}} \alpha^{\frac{1}{\sqrt{x}}}} = 2 \left(\delta - \frac{1}{2}\right) x^{\delta} \cdot \frac{\alpha^{x^{\delta-1/2}}}{\alpha^{\frac{1}{\sqrt{x}}}}$$

Since  $\alpha^{\frac{1}{\sqrt{x}}} \xrightarrow{x \rightarrow +\infty} 1$ , we are left to study  $x^{\delta} \alpha^{x^{\delta-1/2}}$ .

Writing  $x^{\delta} \alpha^{x^{\delta-1/2}} = e^{\delta \log x + \log \alpha \cdot x^{\delta-1/2}}$

and taking the log, since  $\delta > 1/2$  and  $\log \alpha < 0$  we see that

$$\delta \log x + \log \alpha \cdot x^{\delta-1/2} \xrightarrow{x \rightarrow +\infty} -\infty$$

Therefore  $x^{\delta} \alpha^{x^{\delta-1/2}} \xrightarrow{x \rightarrow +\infty} 0$ , and given that  $f(0) = g(0) = 0$ , L'Hospital's rule applies, proving that  $\frac{f(x)}{g(x)} \xrightarrow{x \rightarrow +\infty} 0$ .