

WE CONSIDER THE EDGE RECONNECTING  
MODEL WITH  $\boxed{\kappa=1}$ , STARTED FROM  
THE COMPLETE GRAPH  $K_n$  ON  $n$  VERTICES.  
THUS THE EDGE DENSITY  $\boxed{\rho=1}$

DENOTE BY  $X_n(t)$  THE ADJACENCY MATRIX  
OF THE GRAPH AFTER  $t$  STEPS.

DENOTE BY 
$$\nu^n(t) = \frac{1}{\binom{n}{2}} \sum_{i=1}^n \sum_{j=1}^{i-1} \mathbb{1}[X_{ij}(t) = 0]$$

IN WORDS:  $\nu^n(t)$  IS THE DENSITY OF  
UNORDERED PAIRS OF VERTICES WITH NO  
EDGE BETWEEN THEM.

WE HAVE  $\nu^n(0) = 0$ , BECAUSE IN THE COMPLETE  
GRAPH THERE IS NO PAIR OF VERTICES WITH NO  
EDGE BETWEEN THEM.

CLAIM: IF  $n$  IS LARGE AND  $\boxed{T = \lfloor t \cdot \frac{n^2}{2} \rfloor}$

$$\nu^n\left(\left\lfloor t \cdot \frac{n^2}{2} \right\rfloor\right) \approx (1 - e^{-t}) \cdot e^{-(1 - e^{-t})}$$

$$\nu^n(15 \cdot n^2) \approx (1 - e^{-25}) \cdot e^{-(1 - e^{-25})}$$

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THUS FOR  $[s=3]$  WE ALREADY HAVE

$$N^n(3 \cdot n^2) \approx (1 - e^{-s}) \cdot e^{-(1 - e^{-s})} \approx \frac{1}{e}$$

SIMILARLY FOR  $[s=4], [s=5] \dots$

PROOF OF CLAIM: BY THEOREM 1,

$$N^n\left(\left\lfloor t \cdot \frac{n^2}{2} \right\rfloor\right) = N^n\left(\left\lfloor t \cdot \frac{s(w_0) \cdot n^2}{2} \right\rfloor\right) \approx$$

$$\approx \int_0^1 \int_0^1 w_t(x, y, 0) dx dy \stackrel{(32)}{=}$$

$$= \int_0^1 \int_0^1 \sum_{h=0}^{\infty} w_0(x, y, h) \cdot q(t, h, 0, \frac{D(w_0, x) \cdot D(w_0, y)}{s(w_0)}) = (\star)$$

$w_0(x, y, h) = \mathbb{I}[h=1] \in$  MULTIGRAPH LIMIT OF COMPLETE GRAPHS

$$\text{THUS } s(w_0) = 1, D(w_0, x) \stackrel{(12)}{=} 1$$

$$(\star) = \int_0^1 \int_0^1 q(t, 1, 0, 1) dx dy = q(t, 1, 0, 1) \stackrel{(24)}{=}$$

$$\sum_{l=0}^{\infty} b(l, 1, \bar{e}^t) \cdot p(0-l, (1-\bar{e}^t)) = b(0, 1, \bar{e}^t) \cdot p(0, 1-\bar{e}^t)$$

$$\stackrel{(12), (13)}{=} \binom{1}{0} \cdot 1 \cdot (1-\bar{e}^t)^1 \cdot e^{-(1-\bar{e}^t)} = (1-\bar{e}^t) \cdot \exp(-(1-\bar{e}^t))$$

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BUT  $\frac{1}{e}$  IS NOT THE STATIONARY VALUE OF

$N^m(T)$ ,  $T \rightarrow \infty$ !

IF  $S=1$  AND  $R=1$ , THEN  $\hat{W}_\infty(x, y, z)$  IS  
DEFINED BY (41), THEREFORE

$$N^m(\infty) \approx \int_0^1 \int_0^1 P(0, F^{-1}(x), F^{-1}(y)) dx dy, \text{ WHERE}$$

$$F(x) = \int_0^x g(y, 1, 1) dy = 1 - e^{-x}, \text{ THAT IS}$$

$N^m(\infty) = E(\exp(-X \cdot Y))$ , WHERE  $X$  AND  $Y$   
ARE INDEPENDENT RANDOM VARIABLES WITH  
 $\text{EXP}(1)$  DISTRIBUTION.

$$\begin{aligned} N^m(\infty) &= \int_0^\infty \int_0^\infty \exp(-x \cdot y) \cdot e^{-x} \cdot e^{-y} dx dy = \\ &= \int_0^\infty \frac{1}{1+x} \cdot e^{-x} dx \approx 0.5963... \end{aligned}$$

NOTE THAT  $0.5963 \neq \frac{1}{e} = 0.3678...$

THUS  $\frac{1}{e}$  APPEARS TO BE THE STATIONARY  
VALUE OF  $N^m(T)$  ON THE  $m^2$  TIMESCALE,  
BUT IT IS NOT!

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ABOUT YOUR FIGURE 3:  
EVOLUTION OF DEGREES ON THE  $n^2$  TIMESCALE.

WE FOLLOW THE NOTATION OF (38), (39):

$$D_n(T) = \frac{1}{n} \cdot d(X_n(T), i)$$

NOW  $D_n(0) = 1$   $m = \binom{n}{2}$   $S = 1$   $K = 1$

THUS

$$E(D_n(T+1) - D_n(T) | X_n(T)) \approx \frac{1}{n^3} \cdot (1 - D_n(T)) \quad (38)$$

$$\text{Var}(D_n(T+1) - D_n(T) | X_n(T)) \approx \frac{1}{n^3} \cdot 2 \cdot D_n(T) \quad (39)$$

HEURISTICS: IF  $T \propto n^2$ , THEN  $D_n(T)$  IS NOT TOO DIFFERENT FROM  $D_n(0) = 1$ , SO LET'S SAY

$$E(D_n(T+1) - D_n(T)) \approx 0$$

$$\text{Var}(D_n(T+1) - D_n(T)) \approx \frac{2}{n^3}$$

THUS IF  $T = S \cdot n^2$

THEN

$$E(D_n(T)) \approx 1$$

$$\text{Var}(D_n(T)) \approx T \cdot \frac{2}{n^3} = \frac{2 \cdot S}{n}$$

"INDEPENDENT" INCREMENTS  $\Rightarrow$

$\Rightarrow$  IF  $T = S \cdot n^2$ , THEN

$D_n(S \cdot n^2)$  IS NORMALLY DISTRIBUTED:  $N(\mu=1, \sigma=\sqrt{\frac{2S}{n}})$

THUS THE EMPIRICAL DEGREE DISTRIBUTION  
AT TIME  $T = S \cdot n^2$  IS NORMAL WITH

MEAN  $\approx n$

STANDARD  
DEVIATION  $\approx \sqrt{2 \cdot S \cdot n}$

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