

ESM 296
Individual Assignment 2

Answer Key

The questions below are from Stock and Watson and from Wooldridge

Question 1 (8 points):

Earnings functions attempt to find the determinants of earnings, using both continuous and binary variables. One of the central questions analyzed in this relationship is the returns to education.

(a) Collecting data from 253 individuals, you estimate the following relationship

$$\ln(\hat{E}arn) = 0.54 + 0.083 \times Educ, R^2 = 0.2, SER = 0.445$$

(0.14) (0.011)

Where *Earn* is average hourly earnings and *Educ* is years of education.

What is the effect of an additional year of schooling? If you had a strong belief that years of high school education were different from college education, how would you modify the equation? That if your theory suggested that there was a “diploma effect”?

ANSWER: *One additional year of education carries an 8.3% increase, or return, on earnings. You would need additional data to see if this coefficient was different for high school versus college education. Including both variables in the regression would then allow you to test for equality of the coefficients. A “diploma effect” could be studied by creating a binary variable for a high school diploma, a junior college diploma, a B.A. or B.Sc. diploma, etc.*

(b) In Labor Economics, we teach a model of human capital investments where there are returns to on-the-job training. To approximate on-the-job training, researchers often use a potential experience variable, which is defined as $Exper = Age - Educ - 6$.

You incorporate the potential experience variable into your original regression

$$\ln(\hat{E}arn) = -0.01 + 0.101 \times Educ + 0.033 \times Exper - 0.0005 \times Exper^2,$$

(0.16) (0.012) (0.006) (0.0001)

$$R^2 = 0.34$$

Test for the statistical significance of each of the coefficients. Why has the coefficient on education changed little compared to (a)?

ANSWER: The t-values for the coefficients on *Educ* (t-value: 8.42), *Exper* (t-value: 5.5), and $Exper^2$ (t-value: -5) are all significant (they are greater than 1.96, the appropriate value for testing significance at the 95% level).

In part (a), the coefficient was 0.083, while here it is 0.101. These are relatively close to each other. This can be explained by the nature of the experience variables. Generally, workers with many years of education do not work more years (i.e. have more experience) than those with fewer years of education, so education and experience may be roughly orthogonal (uncorrelated). So in the regression in part (a), perhaps experience was not a source of omitted variable bias.

There may be a little mechanical bias induced by the rule assigning experience. If anything, 1 additional year of education generally results from not participating in the workforce (gaining 1 year of experience). After including this, the coefficient on education increases marginally, so perhaps education and experience are slightly negatively correlated. However, this effect seems to be minor.

(c) You want to find the effect of introducing two variables, gender and marital status. Accordingly, you specify a binary variable that takes on the value of one for females and is zero otherwise (*Female*), and another binary variable that is one if the worker is married but is zero otherwise (*Married*). Adding these variables to the regression results in:

$$\ln(\hat{E}arn) = 0.21 + 0.093 \times Educ + 0.032 \times Exper - 0.0005 \times Exper^2$$

(0.16) (0.012) (0.006) (0.0001)

$$- 0.289 \times Female + 0.062 \times Married$$

(0.049) (0.056)

$$R^2 = 0.43, SER = 0.378$$

Are the coefficients of the two added binary variables individually statistically significant? Are they economically important? In percentage terms, how much less do females earn per hour, controlling

for education and experience? How much more do married people make? What is the percentage difference in earnings between a single male and a married female? What is the marriage differential between males and females?

ANSWER: *The coefficient for the female binary variable is statistically significant even at the 1% level. The coefficient for the married binary variable only has a t-statistic of 1.11 and is not statistically significant at the 10% level. Both coefficients indicate economic importance, since females make approximately 28.9% less than males and married people earn roughly 6.2 percent more (though this is insignificant in statistical terms). A married female earns 22.7% less $(-0.289 + 0.062)$ than a single male. Married females earn 28.9% less than married males, the same percentage that single females earn less than single males.*

(d) In your final specification, you allow for the binary variables to interact. The results are as follows:

$$\ln(\hat{E}arn) = 0.14 + 0.093 \times Educ + 0.032 \times Exper - 0.0005 \times Exper^2$$

$$\begin{array}{cccc} (0.16) & (0.011) & (0.006) & (0.0001) \end{array}$$

$$- 0.158 \times Female + 0.173 \times Married - 0.218 \times (Female \times Married)$$

$$\begin{array}{ccc} (0.075) & (0.080) & (0.097) \end{array}$$

$$R^2 = 0.44, SER = 0.375$$

Repeat the exercise in (c) of calculating the various percentage differences between gender and marital status.

ANSWER: *The default is the single male. Single females earn 15.8% less. Married males earn 17.3% more. Married females earn 20.3% less $(-0.158 + 0.173 - 0.218)$ than single males. Comparing married females with married males now results in a percentage differential of 37.6% in favor of the males $(-0.158 - 0.218)$.*

Question 2 (8 points):

The question below requires the STATA data file "GPA2.dta" is available on the class website. The same file also available in spreadsheet format "GPA2.csv". Now consider the following regression model:

$$COLLGPA = \beta_0 + \beta_1 hsize + \beta_2 hsize^2 + \beta_3 hspc + \beta_4 sat + \beta_5 female + \beta_6 athlete + u$$

Where *COLLGPA* is cumulative college grade point average, *hsize* is size of high school graduating class (in hundreds), *hspc* is academic percentile in graduating class, *sat* is combined SAT score, *female* is a binary gender variable, and *athlete* is a binary variable equal to one for student-athletes.

(a) Estimate the parameters of the regression model above by OLS. What is the estimated GPA differential between athletes and non-athletes?

ANSWER:

```
. regress colgpa hsize hsizesq hsperc sat female athlete, robust
```

Regression with robust standard errors

```
Number of obs = 4137
F( 6, 4130) = 301.41
Prob > F = 0.0000
R-squared = 0.2925
Root MSE = .5544
```

colgpa	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
hsize	-.0568543	.0169123	-3.36	0.001	-.0900115	-.0236971
hsizesq	.0046754	.0023388	2.00	0.046	.00009	.0092608
hsperc	-.0132126	.0005639	-23.43	0.000	-.0143182	-.012107
sat	.0016464	.0000666	24.72	0.000	.0015158	.001777
female	.1548814	.017923	8.64	0.000	.1197427	.1900201
athlete	.1693064	.0369629	4.58	0.000	.0968391	.2417736
_cons	1.241365	.0799464	15.53	0.000	1.084627	1.398103

The estimated GPA differential between athletes and non-athletes is 0.169 GPA points in favor of athletes conditional on SAT score.

(b) Drop *sat* from the model and re-estimate the parameters of the regression model. What is the estimated GPA differential between athletes and non-athletes? Explain why the estimate is different than the one in (a).

ANSWER:

```
. regress colgpa hsize hsizesq hsperc female athlete, robust
```

Regression with robust standard errors

```
Number of obs = 4137
F( 5, 4131) = 187.97
Prob > F = 0.0000
R-squared = 0.1885
Root MSE = .59368
```

colgpa	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
hsize	-.0534038	.0180048	-2.97	0.003	-.0887028	-.0181048
hsizesq	.0053228	.0024964	2.13	0.033	.0004285	.0102171
hsperc	-.0171365	.0005962	-28.74	0.000	-.0183054	-.0159675
female	.0581231	.0187994	3.09	0.002	.0212662	.0949801
athlete	.0054487	.0392878	0.14	0.890	-.0715765	.0824739
_cons	3.047698	.033987	89.67	0.000	2.981065	3.114331

When not controlling for SAT score, the estimated GPA differential between athletes and non-athletes is 0.005 points in favor of athletes although it is not significant. This suggests that the

variables *sat* and *athletes* are negatively correlated. (Recall the formula for omitted variable bias in Lecture 3).

As an alternative way to test for this, you could regress SAT on athlete and all the other variables:

```
. regress sat hsize hsize_sq hperc female athlete, robust
```

This gives a coefficient on *athlete* of roughly -100 SAT points and is very significant. This is evidence of strong correlation.

(c) Including the *sat* variable, re-estimate the model while allowing the effect of being an athlete to differ for males and females and test the null hypothesis that there is no difference in the GPA of female athletes and non-athletes. What about male athletes and non-athletes?

ANSWER:

```
. gen femaleXathlete=female*athlete
```

```
. regress colgpa hsize hsize_sq hperc sat female athlete femaleXathlete, robust
```

Regression with robust standard errors

```
Number of obs =    4137
F(   7,   4129) =   258.25
Prob > F       =    0.0000
R-squared      =    0.2925
Root MSE      =    .55446
```

colgpa	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
hsize	-.0568006	.016936	-3.35	0.001	-.0900043	-.0235969
hsize_sq	.0046699	.0023415	1.99	0.046	.0000793	.0092605
hperc	-.0132114	.0005636	-23.44	0.000	-.0143164	-.0121065
sat	.0016462	.0000667	24.70	0.000	.0015155	.0017769
female	.1546151	.018304	8.45	0.000	.1187294	.1905007
athlete	.1674185	.0411887	4.06	0.000	.0866665	.2481705
femaleXath~e	.0076921	.0862602	0.09	0.929	-.1614243	.1768086
_cons	1.241575	.0800111	15.52	0.000	1.08471	1.398439

```
. test female + athlete + femaleXathlete = female
```

```
( 1) athlete + femaleXathlete = 0
```

```
F(   1,   4129) =    5.14
Prob > F      =    0.0234
```

Holding the other variables in the regression constant, the GPA of female athletes is significantly higher than female non-athletes by 0.175 points (0.167+0.007). The coefficient on *athlete* is the estimated difference in the GPA of male athletes and non-athletes. The GPA of male athlete is 0.167 points higher than that of male non-athletes. The estimate is significant at the 1% level ($t\text{-stat} = 4.06 > 2.58$).

(d) Does the effect of *sat* on *COLLGPA* differ by gender? Justify your answer.

ANSWER:

```
. gen femaleXsat=female*sat
```

```
. regress colgpa hsize hsize^2 hperc sat female athlete femaleXathlete
femaleXsat, robust
```

Regression with robust standard errors

Number of obs = 4137
F(8, 4128) = 228.03
Prob > F = 0.0000
R-squared = 0.2925
Root MSE = .55452

colgpa	Coef.	Robust Std. Err.	t	P> t	[95% Conf. Interval]	
hsize	-.0568198	.0169421	-3.35	0.001	-.0900355	-.0236041
hsize^2	.0046773	.0023424	2.00	0.046	.0000849	.0092696
hperc	-.0132236	.0005634	-23.47	0.000	-.0143281	-.0121192
sat	.001624	.0000871	18.64	0.000	.0014532	.0017949
female	.0990198	.1328792	0.75	0.456	-.161495	.3595346
athlete	.1643156	.0420874	3.90	0.000	.0818016	.2468296
femaleXath~e	.0136833	.08806	0.16	0.877	-.1589618	.1863284
femaleXsat	.0000539	.0001271	0.42	0.671	-.0001952	.0003031
_cons	1.265315	.0994457	12.72	0.000	1.070347	1.460282

The estimated coefficient on the interaction between female and sat is not significant at conventional levels, implying that the effect of sat on COLLGPA does not differ by gender.