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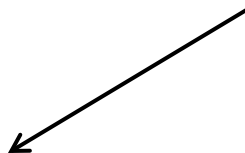
## Lecture 4: Regression specification

- Dummy / indicator variables
- Interactions with indicator variables
- Interactions with “continuous” variables
- Functional form: log-linear regressions
- Applications in Assignment #2
- Chapter 8 in S&W

## Regression with indicator variables

- Suppose we are interested in calculating the percent difference in earnings between males and females
- $Y_i = \ln$  weekly earnings of person  $i$
- $D_{1i} = 1$  (if person  $i$  is female)
- $Y_i = \beta_0 + \beta_1 D_{1i} + u_i$
- $\beta_0 =$  Average log weekly earnings of males
- $\beta_1 =$  Difference in average log weekly earnings between females and males ( $\approx$  percent difference in weekly wages between females and males – see slides at end)

Note: in regressions with only indicator variables, the intercept has a clear interpretation



## Algebra:

- $\beta_0$  = Average log weekly earnings of males
- Since  $\beta_0 = E[Y_i | D_{1i}=0]$
- $\beta_1$  = Difference in average log weekly earnings between females and males
- Since  $E[Y_i | D_{1i}=1] = \beta_0 + \beta_1$
- $\Rightarrow \beta_1 = E[Y_i | D_{1i}=1] - E[Y_i | D_{1i}=0]$

# STATA application (CPS data)

```
. regress lwkearn female, robust;
```

Linear regression

```
Number of obs =      8454
F(   1,   8452) =   487.02
Prob > F       =    0.0000
R-squared      =    0.0542
Root MSE      =    .54535
```

-----							
		Robust					
lwkearn		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
female		-.2613968	.0118448	-22.07	0.000	-.2846155	-.2381782
_cons		6.799713	.008433	806.32	0.000	6.783182	6.816244
-----							

This means that on average, women earn about 26% less than males (not conditional on other attributes, like job type, hours worked)

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## Current Population Survey (CPS)

- ❑ Monthly survey of about 60,000 households
- ❑ Administered by U.S. Census Bureau for the Bureau of Labor Statistics (BLS)
- ❑ Sample represents the civilian noninstitutional U.S. population
- ❑ The survey asks about the employment status of each member of the household 15 years of age or older in the reference week
- ❑ ⇒ Used to construct official unemployment rate series
- ❑ Also ask about demographics, education, wages and income (some months), industry, etc

## Adding more indicator variables to the model

- $D_{2i} = 1$  (if person  $i$  has a college degree)
- $Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + u_i$
- $\beta_0$  = Average log weekly earnings of males without a college degree
- $\beta_1$  = Percent difference in average weekly earnings between females and males, for those without a college degree
- $\beta_2$  = Percent difference in average weekly earnings between college graduates and non-college graduates, irrespective of gender

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□ Algebra:

□  $\mu_{00} = E[Y|D_1=0, D_2=0] = \beta_0$  [omit i subscript]

□  $\mu_{10} = E[Y|D_1=1, D_2=0] = \beta_0 + \beta_1$

□  $\mu_{01} = E[Y|D_1=0, D_2=1] = \beta_0 + \beta_2$

□  $\mu_{11} = E[Y|D_1=1, D_2=1] = \beta_0 + \beta_1 + \beta_2$

□ So:

□  $\beta_0 = \mu_{00}$

□  $\beta_1 = \mu_{10} - \mu_{00}$

□  $\beta_2 = \mu_{01} - \mu_{00} = \mu_{11} - \mu_{10}$

# STATA application

```
. regress lwkearn female college, robust;
```

Linear regression

Number of obs = 8454

F( 2, 8451) = 1251.87

Prob > F = 0.0000

R-squared = 0.2318

Root MSE = .49154

-----							
		Robust					
lwkearn		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
-----+-----							
female		-.2794053	.0106736	-26.18	0.000	-.3003283	-.2584824
college		.4965453	.0115014	43.17	0.000	.4739998	.5190909
_cons		6.635888	.0083724	792.59	0.000	6.619476	6.6523
-----							

This means that on average, non-college graduate women earn about 28% less than non-college graduate males, and the average return to college is about 50%



## Adding an interaction to this model

- There is no reason to restrict the “return to college” is the same for males and females
- Add an interaction between  $D_{1i}$  and  $D_{2i}$  to the model:
- $Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 D_{2i} + \beta_3 (D_{1i} * D_{2i}) + u_i$
- Called “saturated” or “fully-interacted” model (no other functions of  $D_1$  and  $D_2$  can be included in model)
  - Otherwise perfect multicollinearity
- Example of the “difference-in-difference” estimator (to come later)

## Interpretation

- $\beta_0$  = Average log weekly earnings of males without a college degree
- $\beta_1$  = Percent difference in average weekly earnings between females and males without a college degree
- $\beta_2$  = Percent difference in average weekly earnings of males with and without a college degree (i.e., “male college wage premium”)
- $\beta_3$  = Female-male difference in the return to college

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- Algebra:

- $\mu_{00} = E[Y|D_1=0, D_2=0] = \beta_0$  [omit i subscript]

- $\mu_{10} = E[Y|D_1=1, D_2=0] = \beta_0 + \beta_1$

- $\mu_{01} = E[Y|D_1=0, D_2=1] = \beta_0 + \beta_2$

- $\mu_{11} = E[Y|D_1=1, D_2=1] = \beta_0 + \beta_1 + \beta_2 + \beta_3$

- So:

- $\beta_0 = \mu_{00}$

- $\beta_1 = \mu_{10} - \mu_{00}$

- $\beta_2 = \mu_{01} - \mu_{00}$

- $\beta_3 = (\mu_{11} - \mu_{10}) - (\mu_{01} - \mu_{00})$

# STATA application

```
. regress lwkearn female college fcollege, robust;
```

Linear regression

```
Number of obs =      8454
F(   3,   8450) =    838.19
Prob > F       =    0.0000
R-squared      =    0.2319
Root MSE      =    .49153
```

-----						
		Robust				
lwkearn		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
female		-.2702025	.012817	-21.08	0.000	-.295327    -.245078
college		.5095113	.0165479	30.79	0.000	.4770735    .5419492
fcollege		-.0264149	.0229823	-1.15	0.250	-.0714659    .0186361
_cons		6.63161	.0091209	727.08	0.000	6.613731    6.649489
-----						

This means that on average, non-college graduate women earn about 27% less than non-college graduate males, that the average return to college for males is about 51%, and that there is no statistically significance F-M difference in the return to college

## Interactions between indicator variables and 'continuous' variables

- Consider a slightly different model:
- $Y_i = \log$  weekly earnings of person  $i$
- $D_{1i} = \mathbf{1}$  (if person  $i$  is female)
- $X_{1i} =$  years of education of person  $i$
- $Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 X_{1i} + u_i$
- $\beta_0 = E[Y_i | D_{1i}=0, X_{1i}]$
- $\beta_1 = E[Y_i | D_{1i}=1, X_{1i}] - E[Y_i | D_{1i}=0, X_{1i}]$
- $\beta_2 =$  percent increase in weekly earnings associated with an 1 additional year of education ("return to education"), restricted to be the same for males and females

# STATA application

```
. regress lwkearn female yrseduc, robust;
```

Linear regression

Number of obs = 8454  
F( 2, 8451) = 1460.34  
Prob > F = 0.0000  
R-squared = 0.2731  
Root MSE = .47814

-----						
		Robust				
lwkearn		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
female		-.2953517	.0104344	-28.31	0.000	-.3158056    -.2748978
yrseduc		.1032773	.0021292	48.51	0.000	.0991036    .107451
_cons		5.390107	.0293252	183.80	0.000	5.332623    5.447592
-----						

This means that on average, women earn about 30% less than males, and that the average return to an additional year of education is about 10%

## Allowing the return to education to vary by gender

- $Y_i = \beta_0 + \beta_1 D_{1i} + \beta_2 X_{1i} + \beta_3 (D_{1i} * X_{1i}) + u_i$
- $\beta_0 = E[Y_i | D_{1i}=0, X_{1i}]$
- $\beta_1 = E[Y | D_{1i}=1, X_{1i}] - E[Y | D_{1i}=0, X_{1i}]$
- $\beta_2 = \text{male return to education}$
- $\beta_3 = \text{female-male difference in the return to education}$

# STATA application

```
. regress lwkearn female yrseduc fyrseduc, robust;
```

Linear regression

Number of obs = 8454  
F( 3, 8450) = 982.28  
Prob > F = 0.0000  
R-squared = 0.2732  
Root MSE = .47814

-----						
		Robust				
lwkearn		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
-----+-----						
female		-.2415737	.0597732	-4.04	0.000	-.3587438 -.1244036
yrseduc		.1050051	.0027211	38.59	0.000	.099671 .1103392
fyrseduc		-.0038881	.0043246	-0.90	0.369	-.0123653 .0045891
_cons		5.366525	.0369454	145.26	0.000	5.294103 5.438947
-----						



## 8.4

Through the use of the interaction term  $X_i \times D_i$ , the population regression line relating  $Y_i$  and the continuous variable  $X_i$  can have a slope that depends on the binary variable  $D_i$ . There are three possibilities:

1. Different intercept, same slope (Figure 8.8a):

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + u_i;$$

2. Different intercept and slope (Figure 8.8b):

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 D_i + \beta_3 (X_i \times D_i) + u_i;$$

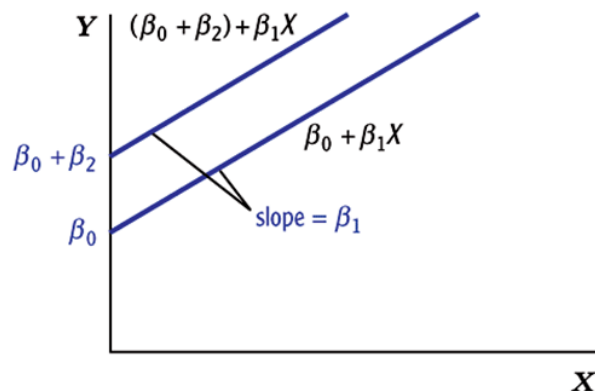
3. Same intercept, different slope (Figure 8.8c):

$$Y_i = \beta_0 + \beta_1 X_i + \beta_2 (X_i \times D_i) + u_i.$$

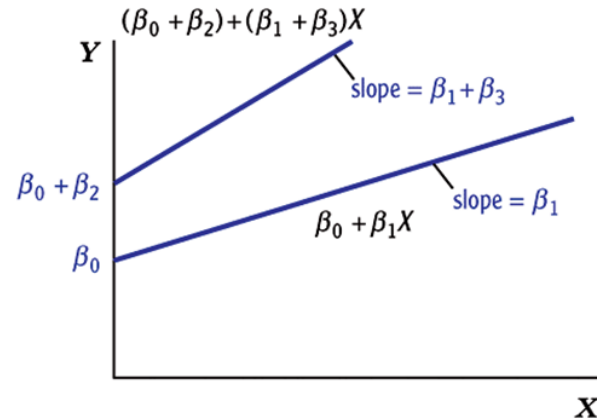
Note: “intercept” in this slide does not refer to the regression intercept ( $\beta_0$ ) per se, but to level difference in the population regression function that comes from the group indicator  $D_i$

The intercept for the group where  $D_i=0$  is  $\beta_0$ , while the intercept for the group where  $D_i=1$  is  $\beta_0 + \beta_2$

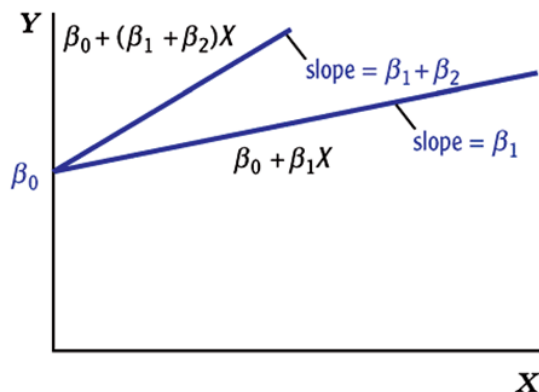
**FIGURE 8.8** Regression Functions Using Binary and Continuous Variables



(a) Different intercepts, same slope



(b) Different intercepts, different slopes



(c) Same intercept, different slopes

Interactions of binary variables and continuous variables can produce three different population regression functions: (a)  $\beta_0 + \beta_1 X + \beta_2 D$  allows for different intercepts but has the same slope; (b)  $\beta_0 + \beta_1 X + \beta_2 D + \beta_3 (X \times D)$  allows for different intercepts and different slopes; and (c)  $\beta_0 + \beta_1 X + \beta_2 (X \times D)$  has the same intercept but allows for different slopes.

## Interactions between two 'continuous' variables

- $Y_i$  = log weekly earnings of person  $i$
- $X_{1i}$  = years of education of person  $i$
- $X_{2i}$  = age of person  $i$
- $Y_i = \beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i} + \beta_3 (X_{1i} * X_{2i}) + u_i$
- The interaction allows the return to education to vary by age:

$$\frac{\partial Y_i}{\partial X_{1i}} = \beta_1 + \beta_3 X_{2i}$$

- This allows marginal effects to vary by value of  $X_{2i}$

# STATA application

```
. regress lwkearn age educ age yrseduc, robust;
```

## Linear regression

Number of obs = 8454

$$F(3, 8450) = 709.70$$

Prob > F = 0.0000

R-squared = 0.2068

Root MSE = .49951

		Robust					
lwkearn		Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age		.0053146	.0034072	1.56	0.119	-.0013643	.0119935
educ		.1058259	.0113045	9.36	0.000	.0836664	.1279854
age_yrseduc		-.0001394	.0002455	-0.57	0.570	-.0006207	.0003419
_cons		5.061251	.1574626	32.14	0.000	4.752585	5.369916

```
. lincom educ + age_yrs*44.92 <
```

```
( 1) educ + 44.92*age_yrseduc = 0
```

For linear marginal effects in STATA use “lincom”. Here 44.92 is the average age in the sample, so this gives the return to education evaluated at the average age

lwkearn	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]
(1)	.0995637	.0021612	46.07	0.000	.0953273 .1038002

## (Natural) Logarithmic regression

- By far, this is the most frequently used “nonlinear” regression model
- Why:
- Logs convert changes into percentage change
- A log-log regression yields estimates of elasticities
- Often fits data better
- Always use the natural log:  $x = \ln(\exp(x))$ , where  $e = 2.71828$
- I will write “log” but mean “ln” (also true with Stata)

## Review: Properties of 'ln' function

- (a)  $\log(1/x) = -\log(x)$
- (b)  $\log(a^*x) = a^*\log(x)$
- (c)  $\log(x/a) = \log(x) - \log(a)$
- (d)  $\partial \log(x) / \partial x = 1/x$
  
- $\Rightarrow$  Natural Log transformation models relations in “percentage” terms, rather than in natural units (linearly)
  
- *Here's why:*  $\ln(x+\Delta x) - \ln(x) = \ln\left(1 + \frac{\Delta x}{x}\right) \approx \frac{\Delta x}{x}$

# I. Linear-log regression

- $Y_i = \beta_0 + \beta_1 \log(X_{1i}) + u_i$
- $\beta_1 \equiv \partial Y_i / \partial \log(X_{1i})$
- $\Rightarrow \beta_1$  measures the unit change in  $Y$  arising from a proportionate change in  $X_{1i}$
- Why:

$$\frac{\partial Y_i}{\partial \log(X_{1i})} = \beta_1 = \frac{\partial Y_i}{\partial X_{1i}} \frac{\partial X_{1i}}{\partial \log(X_{1i})} = X_{1i} \frac{\partial Y_i}{\partial X_{1i}} \approx \frac{\Delta Y_i}{\Delta X_{1i} / X_{1i}}$$

## II. Log-linear regression

- $\log(Y_i) = \beta_0 + \beta_1 X_{1i} + u_i$
- $\beta_1 \equiv \partial \log(Y_i) / \partial X_{1i}$
- Measures the proportionate change in Y arising from a 1-unit change in  $X_{1i}$

- Why:

$$\frac{\partial \log(Y_i)}{\partial X_{1i}} = \beta_1 = \frac{\partial \log(Y_i)}{\partial Y_i} \frac{\partial Y_i}{\partial X_{1i}} = \frac{1}{Y_i} \frac{\partial Y_i}{\partial X_{1i}} \approx \frac{\Delta Y_i / Y_i}{\Delta X_{1i}}$$

- In other words  $\beta_1$  measures the percent effect on Y of a 1-unit change in  $X_1$
- Classic example: Estimating the return to schooling



### III. Log-log regression

- $\log(Y_i) = \beta_0 + \beta_1 \log(X_{1i}) + u_i$
- $\beta_1 \equiv \partial \log(Y_i) / \partial \log(X_{1i})$
- $\beta_1$  = proportionate change in  $Y_i$  arising from a proportionate change in  $X_{1i}$
- Measures the elasticity of  $Y_i$  with respect to  $X_{1i}$
- Why:

$$\frac{\partial \log(Y_i)}{\partial \log(X_{1i})} = \beta_1 = \frac{\partial \log(Y_i)}{\partial Y_i} \partial Y_i \frac{1}{\partial \log(X_{1i})} \frac{\partial X_{1i}}{\partial X_{1i}} = \frac{\partial Y_i}{Y_i} \frac{X_{1i}}{\partial X_{1i}} \approx \frac{\Delta Y_i / Y_i}{\Delta X_{1i} / X_{1i}}$$