## Choosing the weights for the logarithmic pooling of probability distributions.

<u>Luiz Max F. de Carvalho</u> [lmax.procc@gmail.com], Daniel Villela, Flavio Coelho & Leonardo S. Bastos [lsbastos@fiocruz.br]

Scientific Computing Programme (PROCC), Oswaldo Cruz Foundation, Fiocruz, Brazil. 60th World Statistics Congress ISI2015, Rio de Janeiro, Brazil.

July 29, 2015

- Obtain important insights on consensus belief formation and group decision making (Genest and Zidek, 1986);
- Applications in a range of fields, from infectious disease modelling (Coelho and Codeço, 2009) and wildlife conservation (Poole and Raftery, 2000) to engineering (Savchuk and Martz, 1994);
- One important question remains, however: how to give each expert/information source a weight without being arbitrary?



Let  $\mathbf{F}(\theta) = \{f_0(\theta), f_1(\theta), f_2(\theta), \dots, f_K(\theta)\}$  be the set of prior distributions representing the opinions of K+1 experts and let  $\alpha = \{\alpha_0, \alpha_1, \alpha_2, \dots, \alpha_K\}$  be the vector of weights, such that  $\alpha_i > 0 \ \forall i$  and  $\sum_{i=0}^K \alpha_i = 1$ . Then the log-pooled prior is

$$\pi(\theta) = t(\alpha) \prod_{i=0}^{K} f_i(\theta)^{\alpha_i}$$
 (1)

with  $t(\alpha) = \int_{\Theta} \prod_{i=0}^{K} f_i(\theta)^{\alpha_i} d\theta$ .

- Enjoys rather desirable properties, such as external Bayesianity (Genest and Zidek, 1986);
- Poole and Raftery (2000) prove that t(α) is always finite for the case K = 1 (2 experts/priors), which we extend for any finite K. See Theorem 1 in http://arxiv.org/pdf/1502.04206v1.pdf for a simple proof.



• If there is no information about the reliabilities of the experts one might want to assign  $\alpha$  so as to maximise entropy of the resulting distribution:

$$H_{\pi}(\theta) = -\int_{\Theta} \pi(\theta) \ln \pi(\theta) d\theta$$

$$H_{\pi}(\theta; \alpha) = \sum_{i=0}^{K} \alpha_{i} E_{\pi}[-\ln f_{i}(\theta)] - \ln t(\alpha)$$

ullet Formally, we want to find  $\hat{lpha}$  such that

$$\hat{\boldsymbol{\alpha}} := \operatorname{arg\,max} H_{\pi}(\theta; \boldsymbol{\alpha})$$

• Caveat: is not guaranteed to yield an unique solution.



- What if we want to minimise conflict between the consensus and each individual opinion?
- Let  $d_i = \mathsf{KL}(f_i||\pi)$  and let  $L(\alpha)$  be a loss function such that

$$L(\alpha) = \sum_{i=0}^{K} d_i \tag{2}$$

$$= -K \ln t(\alpha) + \sum_{i=0}^{K} \sum_{j\neq i}^{K} \alpha_j \mathsf{KL}(f_i||f_j) \tag{3}$$

$$\hat{\alpha} := \arg\min L(\alpha) \tag{4}$$

 Contrary to the above, the loss function is convex, thus there is a unique solution (Rufo et al., 2012).



- An appealing alternative is to place a (hyper) prior on the weights  $(\alpha)$ ;
- Two options:
  - (a) Dirichlet prior:

$$\pi(oldsymbol{lpha}) = rac{1}{\mathcal{B}(oldsymbol{X})} \prod_{i=0}^{K} lpha_i^{\mathsf{x}_i-1}$$

(b) logistic-normal:

$$\alpha_i = \frac{e^{m_i}}{\sum_{i=0}^K e^{m_i}}, \ m_i \sim N(\mu_i, \sigma_i^2)$$

- Advantage: accomodates uncertainty in natural way, and is very flexible;
- Caveat(s): may yield inconsistent results and hardly ever allows for analytical solutions.



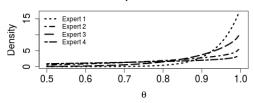
ullet  $Y \sim Bernoulli( heta)$  and

$$f_i(\theta; a_i, b_i) = \frac{\Gamma(a_i + b_i)}{\Gamma(a_i b_i)} \theta^{a_i - 1} (1 - \theta)^{b_i - 1}$$

- Allows for simple expressions for the entropy and KL divergence  $[\pi(\theta; \alpha)]$  is also Beta], and efficient sampling from the hyperpriors;
- Savchuk and Martz (1994) consider an example in which four experts are required supply prior information about the survival probability of a certain unit for which there have been y=9 successes out of n=10 trials;
- We propose to evaluate performance using integrated (marginal) likelihoods, a.k.a., prior evidence.



## **Expert Priors**



## **Pooled Priors and Posteriors**

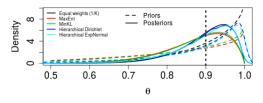




Table : Weights obtained using the three methods for the proportion estimation problem.  $^1$  – Kullback-Liebler  $^2$  – Posterior mean for  $\alpha$ .

Method	$\alpha_0$	$\alpha_1$	$\alpha_2$	$\alpha_3$
Maximum entropy	0.00	1.00	0.00	0.00
Minimum KL <sup>1</sup> divergence	0.04	0.96	0.00	0.00
Hierarchical prior <sup>2</sup>	0.26	0.24	0.26	0.23

Table : Integrated likelihoods for the priors of each expert as well as the combined priors.  $^1$  Calculated using the posterior mean of  $\alpha$ 

Expert priors		Pooled priors		
Expert 0	0.237	Equal weights	0.254	
Expert 1	0.211	Maximum entropy	0.211	
Expert 2	0.256	Minimum KL	0.223	
Expert 3	0.163	Hierarchical <sup>1</sup>	0.255	



- Our results are not yet decisive regarding which method is better;
- The Dirichlet approach seems the most natural from a Bayesian perspective, but prior sensitivity is currently unknown;
- Generalise the hyperprior approach to the exponential family;
- The relationship between the integrated likelihood of the combined (pooled) distribution in relation to that of each individual distribution.
- Applications to dynamic models (Poole and Raftery, 2000).



- Thank you very much for your attention!
- The authors would like to thank Professor Adrian Raftery (University of Washington) for helpful suggestions. DAMV was supported in part by Capes under Capes/Cofecub project (N. 833/15). FCC is grateful to Fundação Getulio Vargas for funding during this project.



- Coelho, F. C. and Codeço, C. T. (2009). Dynamic modeling of vaccinating behavior as a function of individual beliefs. *PLoS Comput. Biol.*, 5(7):e1000425.
- Genest, C. and Zidek, J. V. (1986). Combining probability distributions: A critique and an annotated bibliography. *Statistical Science*, pages 114–135.
- Poole, D. and Raftery, A. E. (2000). Inference for deterministic simulation models: the bayesian melding approach. *Journal of the American Statistical Association*, 95(452):1244–1255.
- Rufo, M., Martin, J., Pérez, C., et al. (2012). Log-linear pool to combine prior distributions: A suggestion for a calibration-based approach. *Bayesian Analysis*, 7(2):411–438.
- Savchuk, V. P. and Martz, H. F. (1994). Bayes reliability estimation using multiple sources of prior information: binomial sampling. *Reliability, IEEE Transactions on*, 43(1):138–144.