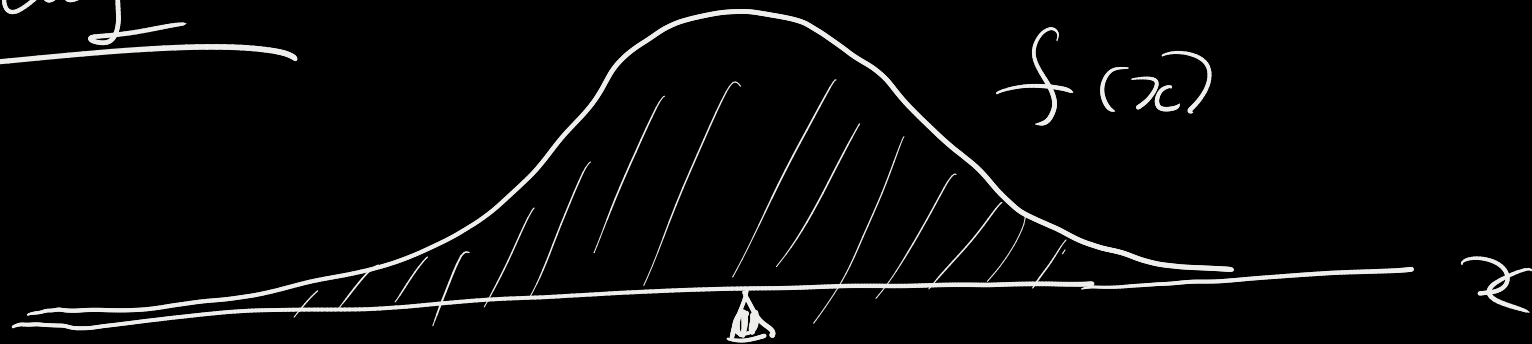


Chapter 4 Expected Value

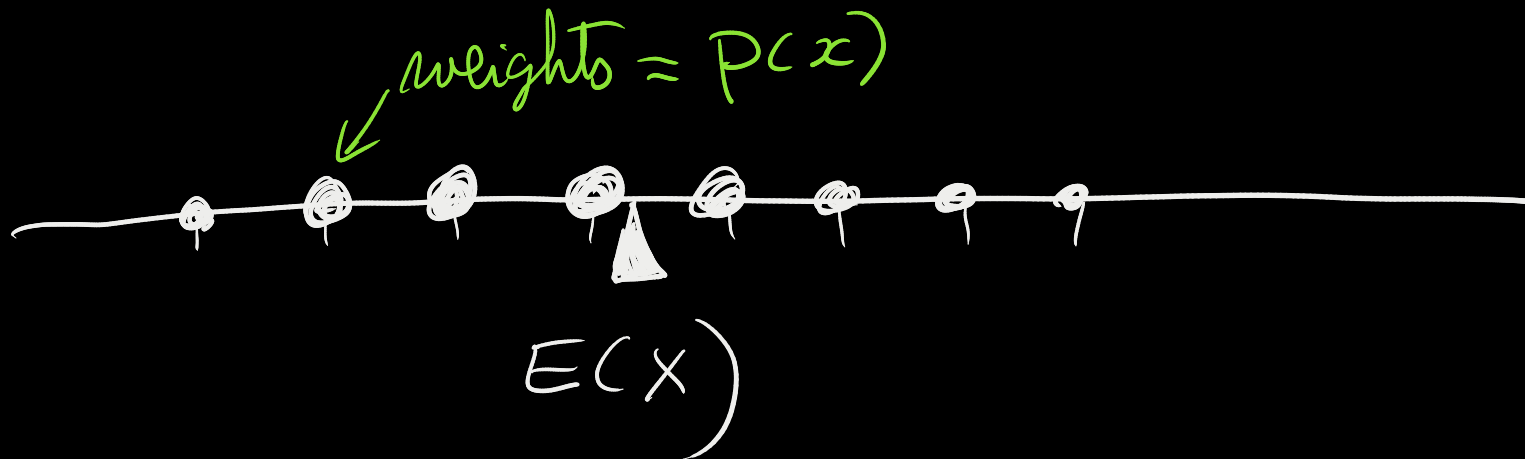
Geometrically

$$\underline{X \sim f}$$



$E(X)$ is point of balance

$$X \sim p$$



Defn $E(X) = \sum x p(x)$ provided $\sum |x| p(x) < \infty$
otherwise "not defined"

$$E(X) = \int x f(x) dx \text{ provided } \int |x| f(x) dx < \infty$$

Examples:

Exponential (λ)

$$f(x) = \lambda e^{-\lambda x}$$

$$x > 0, \lambda > 0$$

$$E(X) = \int_0^{\infty} x \lambda e^{-\lambda x} dx$$

$$= \Gamma(2) \frac{1}{\lambda} \int \underbrace{\frac{x^{2-1} \lambda^2}{\Gamma(2)} e^{-\lambda x}}_{\text{Gamma}(2, \lambda)} dx$$

$$= \frac{1}{\lambda}$$

Use Gamma distribution

$$\int_0^{\infty} \frac{1}{\Gamma(\alpha)} \lambda^{\alpha} x^{\alpha-1} e^{-\lambda x} dx = 1$$

Poisson: $E(x) = \sum_{k=0}^{\infty} \underbrace{\frac{k \lambda^k}{k!}}_{\text{first term is 0}} e^{-\lambda}$

$$= e^{-\lambda} \sum_{k=1}^{\infty} \frac{\lambda^k}{(k-1)!}$$

We know $\sum_{j=0}^{\infty} \frac{\lambda^j}{j!} = e^{\lambda}$

$$\begin{aligned}
 &= e^{-\lambda} \lambda \left(\sum_{k=1}^{\infty} \frac{\lambda^{k-1}}{(k-1)!} \right) \quad \text{but } 0 = \sum_{k=0}^{\infty} \frac{\lambda^k}{k!} \\
 &= e^{-\lambda} \lambda e^{\lambda} \\
 &= \lambda
 \end{aligned}$$

St. Petersburg Paradox

How to make a sure \$1

Find someone willing to make
fair bets on the toss of a coin.

i.e. if coin is H they give me \$N
" " " T I give them \$N

Let X = amount I win.

$$E(X) = \frac{1}{2}N + \frac{1}{2}(-N) = 0$$

"Fair bet"

My strategy:

- Bet \$1
- If I win stop
- If I lose bet double the amount
- Keep doubling until I win

If my first win is at the k th bet
my winnings are:

$$\boxed{-1 - 2 - 4 \dots - 2^j - \dots - 2^{k-1}} + 2^k$$

$= \$1$

~~I am sure to eventually win \$1~~

next-first

$$\sum_{i=0}^{n-1} a^i = \frac{a^n - 1}{a - 1}$$

if $a \neq 1$

ratio-1

Recall for
geometric series

— works for all finite n

— For " $n = \infty$ " need $|a| < 1$

Let w = amount won

It looks like $E(w) = 1$

not much, but sure!

Formally

$$E(w) = 1 \cdot \frac{1}{2} + 1 \cdot \frac{1}{4} + 1 \cdot \frac{1}{8} + \dots +$$

$$= 1 \left(\frac{1}{2} + \frac{1}{4} + \dots \right)$$

$$= 1$$

But let X be amount won on first winning bet:

$$E(X) = 1 \cdot \frac{1}{2} + 2 \cdot \frac{1}{4} + 4 \cdot \frac{1}{8} + 8 \cdot \frac{1}{16} + \dots$$

$$= \frac{1}{2} + \frac{1}{2} + \frac{1}{2} + \dots$$

$$= \infty$$

Let L be amount lost until first winning bet.

$$L = X - 1$$

$$E(L) = \infty.$$

$$\boxed{\infty - \infty = 1 \quad ?}$$

Problem: You need ∞ capital to play this game

- If you have finite capital you might go bankrupt before winning
- Suppose you only have enough money to lose K times, then

$$P(W = -\sum_{i=1}^K 2^{i-1}) = 1 - \sum_{i=1}^K \left(\frac{1}{2}\right)^i$$

$$P\left(W = -\frac{2^k - 1}{2 - 1}\right) = 1 - \frac{\left(\left(\frac{1}{2}\right)^{k+1} - \frac{1}{2}\right)}{\frac{1}{2} - 1}$$

$$P\{W = -(2^k - 1)\} = 1 - \left(1 - \left(\frac{1}{2}\right)^k\right) = \frac{1}{2^k}$$

$$E(W) = -(2^k - 1) \times \frac{1}{2^k} + 1 \left(1 - \frac{1}{2^k}\right)$$

$$= -1 + \frac{1}{2^k} + 1 - \frac{1}{2^k}$$

$$= 0$$

So a "fair" bet with a high probability of winning small at the cost of a small probability of losing big.

Scam strategy: Play the game with IOU's.

Gamma

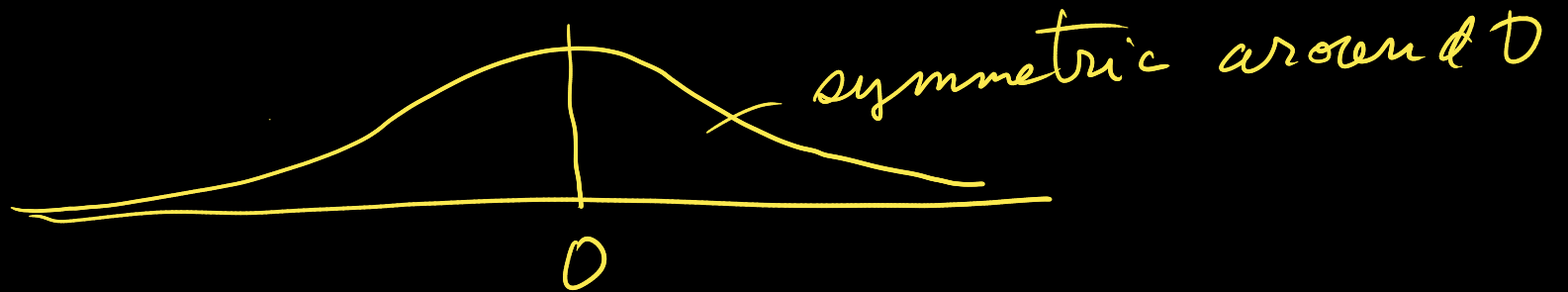
$$E(X) = \int_0^{\infty} x \frac{\lambda^{\alpha}}{\Gamma(\alpha)} x^{\alpha-1} e^{-\lambda x} dx$$

$$= \int_0^{\infty} \left(\frac{\lambda^{\alpha}}{\Gamma(\alpha)} \right) x^{\alpha} e^{-\lambda x} dx$$

$$= \frac{\lambda^{\alpha}}{\Gamma(\alpha)} \frac{\Gamma(\alpha+1)}{\lambda^{\alpha+1}} \int_0^{\infty} \underbrace{\frac{\lambda^{\alpha+1}}{\Gamma(\alpha+1)} x^{(\alpha+1)-1} e^{-\lambda x} dx}_{\text{Gamma}(\alpha+1) \text{ density}}$$

$$= \frac{\Gamma(\alpha+1)}{\Gamma(\alpha)} \frac{1}{\lambda} = \frac{\alpha}{\lambda} \quad \text{since } \Gamma(\alpha+1) = \alpha \Gamma(\alpha)$$

Cauchy: $f(x) = \frac{1}{\pi} \left(\frac{1}{1+x^2} \right), -\infty < x < +\infty$



But $E(|x|) = \int_{-\infty}^{\infty} |x| \frac{1}{\pi} \frac{1}{1+x^2} dx$

$$= 2 \int_0^{\infty} \frac{1}{\pi} \frac{x}{1+x^2} dx$$

$$= \frac{2}{\pi} \left[\frac{1}{2} \log(x^2 + 1) \right]_0^{\infty}$$

$$= \infty !$$

Cauchy CDF

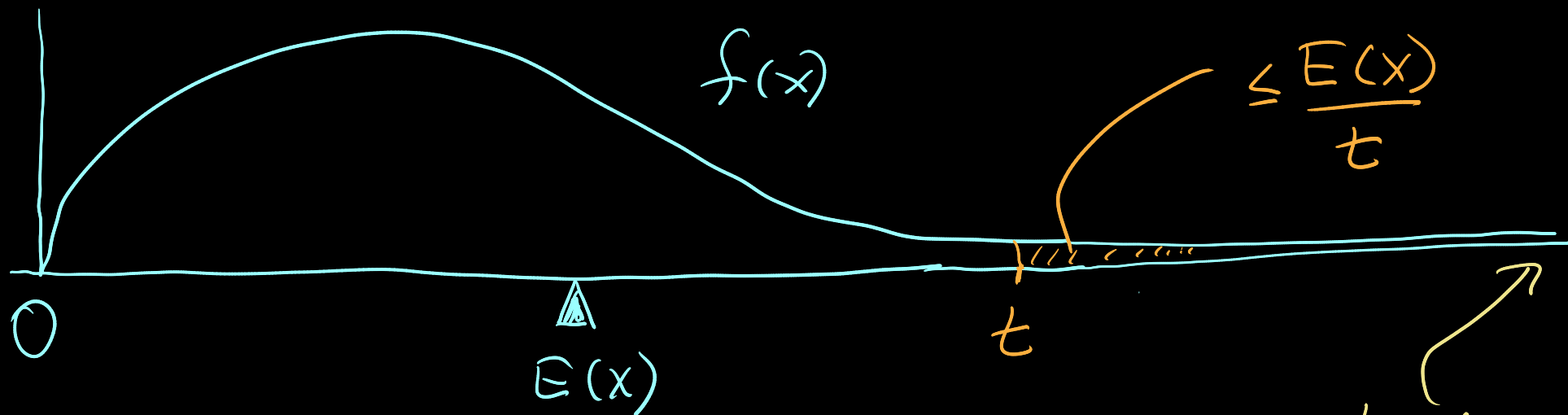
$$F^{-1}(u) = \tan^{-1}\left(\pi u - \frac{\pi}{2}\right)$$

easy to simulate.

Markov Inequality :

If 1) X is a RV with $P(X \geq 0) = 1$
2) $E(X) < \infty$

Then for any $t > 0$ $P(X \geq t) \leq \frac{E(X)}{t}$



Note $E(X)$ very affected by tail

Proof: $E(X) = \int_0^{\infty} x f(x) dx$

$$= \int_0^t x f(x) dx + \int_t^{\infty} x f(x) dx$$

$$\geq \int_t^{\infty} x f(x) dx$$

$$\geq \int_t^{\infty} t f(x) dx$$

$$= t P(X \geq t)$$

need this for
discrete case

