

3.5 Functions of random vectors

Sums: Given a joint distribution for (X, Y)
find the distribution $Z = X + Y$

Example

Y	2	1	0
X	0	1	2
2	0	.2	.2
1	.1	.1	.1
0	.1	.2	0

What are
possible values
for $Z = X + Y$

$$P(Z=0) = \sum_{x,y: x+y=0} P(x,y) = \sum_{x,y=0-x} P(x,y)$$

$$= \sum_x p(x, 0-x)$$

$$P(Z=1) = \sum_x p(x, 1-x)$$

$$P(Z=2) = \sum_x p(x, 2-x)$$

$$P(Z=3) = \dots$$

$$P(Z=4) = \dots$$

Y			
2	0	.2	.2
1	.1	.1	.1
0	.1	.2	0
	0	1	2 X

General formula:

$$\text{of } Z = X + Y$$

Discrete
case

$$P_Z(z) = \sum_x p(x, z-x)$$

$$\stackrel{\text{or}}{=} \sum_y p(z-y, y)$$

$$\text{Continuous: } f_z(z) = \int_{-\infty}^{\infty} f(x, z-x) dx$$

$$= \int_{-\infty}^{\infty} f(z-y, y) dy$$

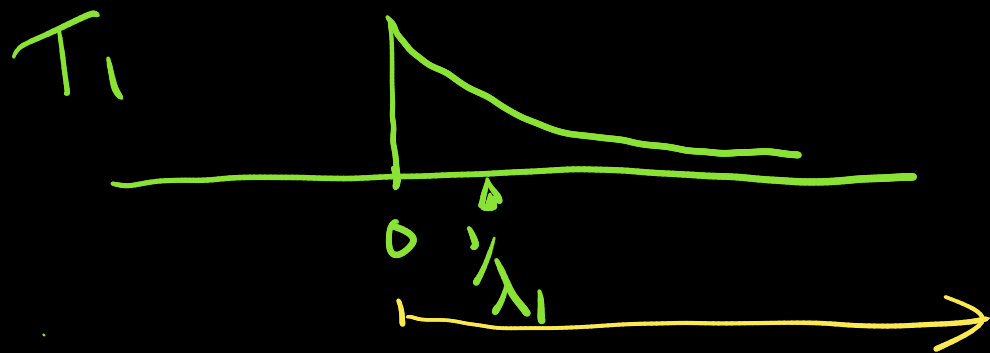
If X & Y are independent: $f(x, y) = f_x(x) f_y(y)$

and $f_2(z) = \int_{-\infty}^{\infty} f(x, z-x) dx$

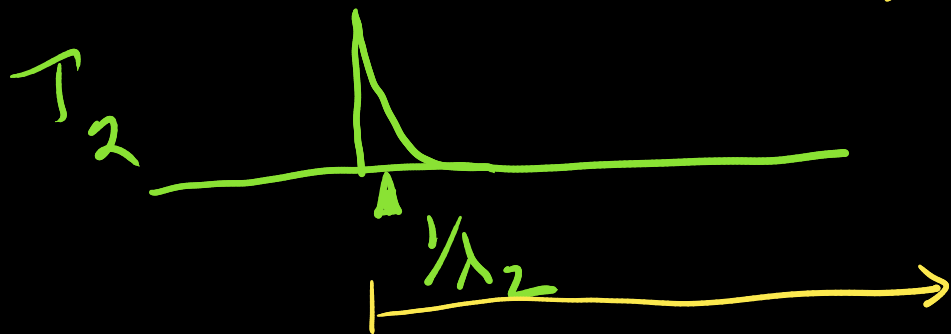
$$= \int_{-\infty}^{\infty} f_x(x) f_y(z-x) dx$$

"convolution of f_x & f_y "

Example 1: Computer has lifetime $T_1 \sim \text{Exponential}(\lambda_1)$
Automatic backup $T_2 \sim \text{Exponential}(\lambda_2)$



$$f_1(t_1) = \lambda_1 e^{-\lambda_1 t_1}, t_1 > 0$$

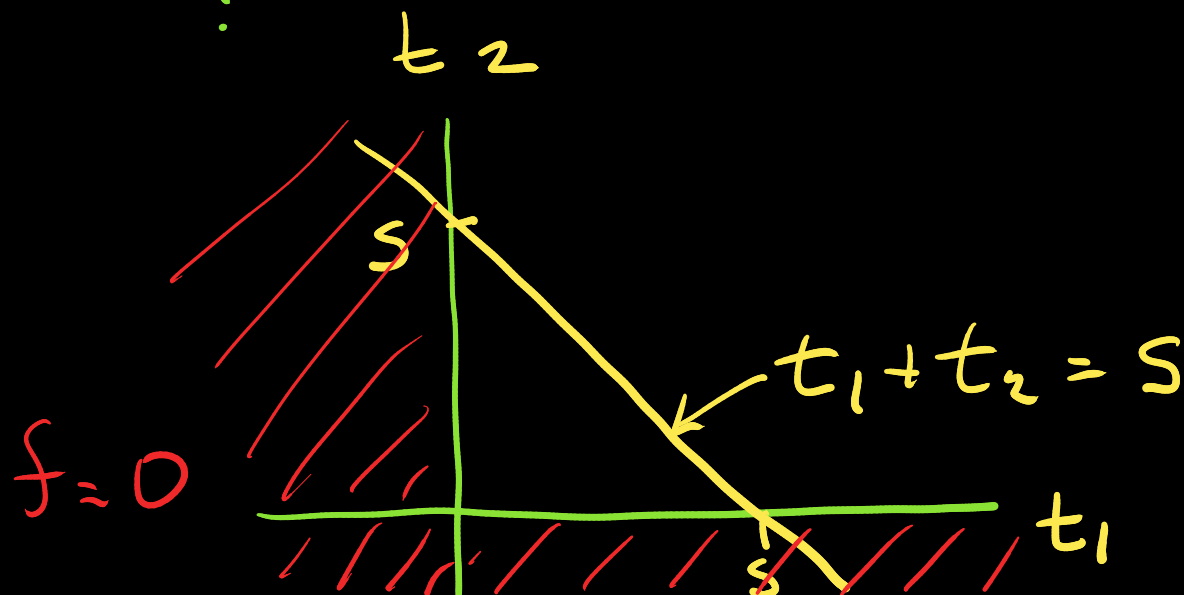


$$f_2(t_2) = \lambda_2 e^{-\lambda_2 t_2}, t_2 > 0$$

$$S = T_1 + T_2 \quad ; \quad T_1, T_2 \text{ independent}$$

$$f_s(s) = \int_{-\infty}^{\infty} f_1(t_1) f_2(s-t_1) dt_1$$

$$= \int_{?}^{?} \lambda_1 e^{-\lambda_1 t_1} \lambda_2 e^{-\lambda_2 (s-t_1)} dt_1$$



s
 $? < t_1 < ?$

$$= \int_{t_1=0}^{t_1=s} \lambda_1 e^{-\lambda_1 t_1} \lambda_2 e^{-\lambda_2 (s-t_1)} dt_1$$

"Simple" if $\lambda_1 = \lambda_2 = \lambda$ i.e. Same expected lifetime.

$$f_s(s) = \int_0^s \lambda e^{-\lambda t} \lambda e^{-\lambda (s-t)} dt \quad \lambda^2 e^{-\lambda s}$$

$$= \int_0^s \lambda^2 e^{-\lambda s} dt$$

$$= \left[t \lambda^2 e^{-\lambda s} \right]_0^s$$

$$= s \lambda^2 e^{-\lambda s}, \quad 0 < s$$

$$\sim \text{Gamma}(2, \lambda) \text{ i.e. } \frac{1}{\Gamma(\alpha)} s^{\alpha-1} \lambda^\alpha e^{-\lambda s},$$

$0 < s, \alpha = 2$

General Case : Using Jacobians

Transformation $\mathbb{R}^1 \rightarrow \mathbb{R}^d$

of $y = h(x)$

h is differentiable and monotone

"strictly" increasing on an interval I

or "decreasing" " " " "

$$f_y(y) = f_x(g^{-1}(y)) \times \left| \frac{dx}{dy} \right|$$

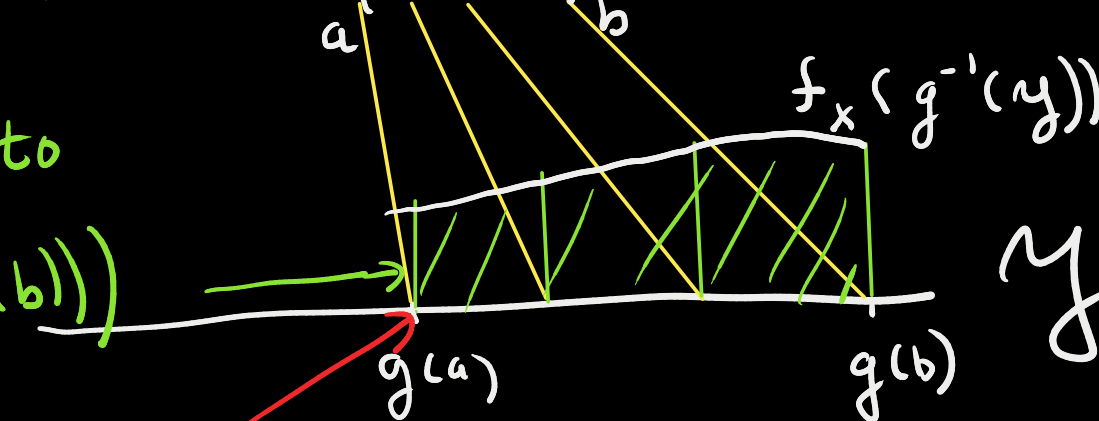
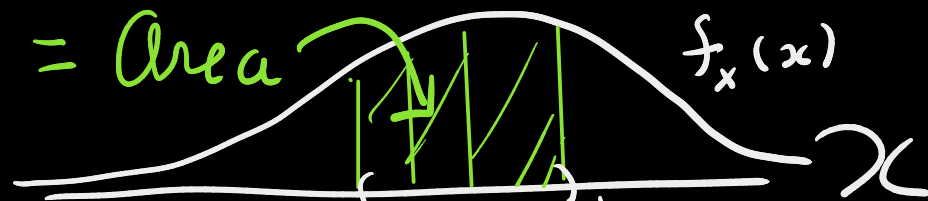
$|\dots|$ so it works
whether increasing or decreasing

$$P(X \in (a, b)) = \text{Area}$$

||

needs to
be equal to

$$P(Y \in (g(a), g(b)))$$



$g \downarrow$

But this area is too big because $g(x)$ stretches the base

What to do?

Undo the effect of stretching
by dividing by $\left| \frac{dy}{dx} \right|$
i.e. the amount of stretching

Now $\left| \frac{dy}{dx} \right| = \left| \frac{dx}{dy} \right|^{-1}$

So we can multiply by $\left| \frac{dx}{dy} \right|$

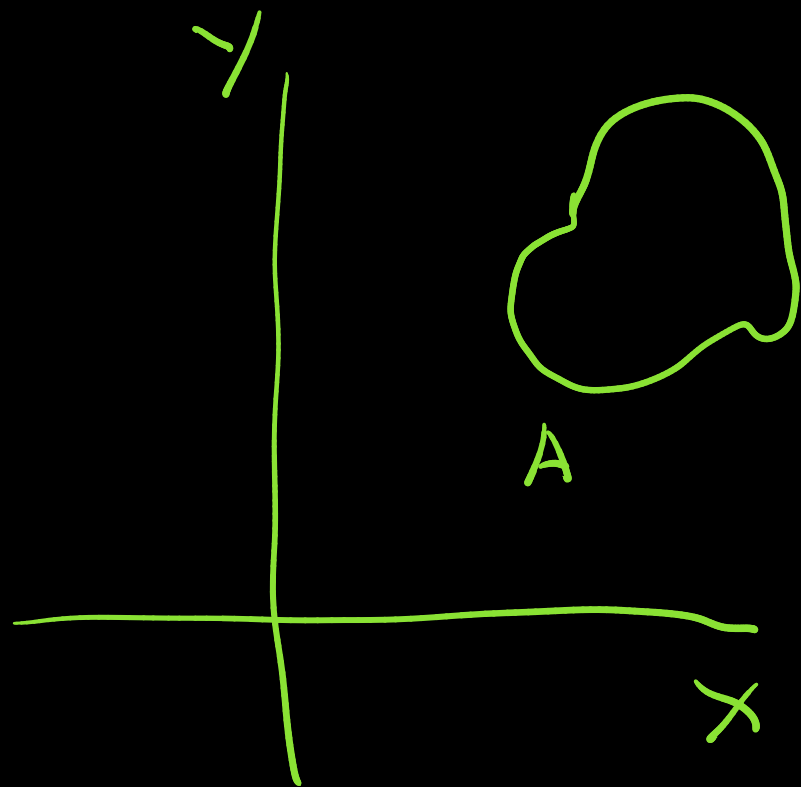
Formula:

$$f_y(y) = f_x(g^{-1}(y)) \times \left| \frac{dx}{dy} \right|$$

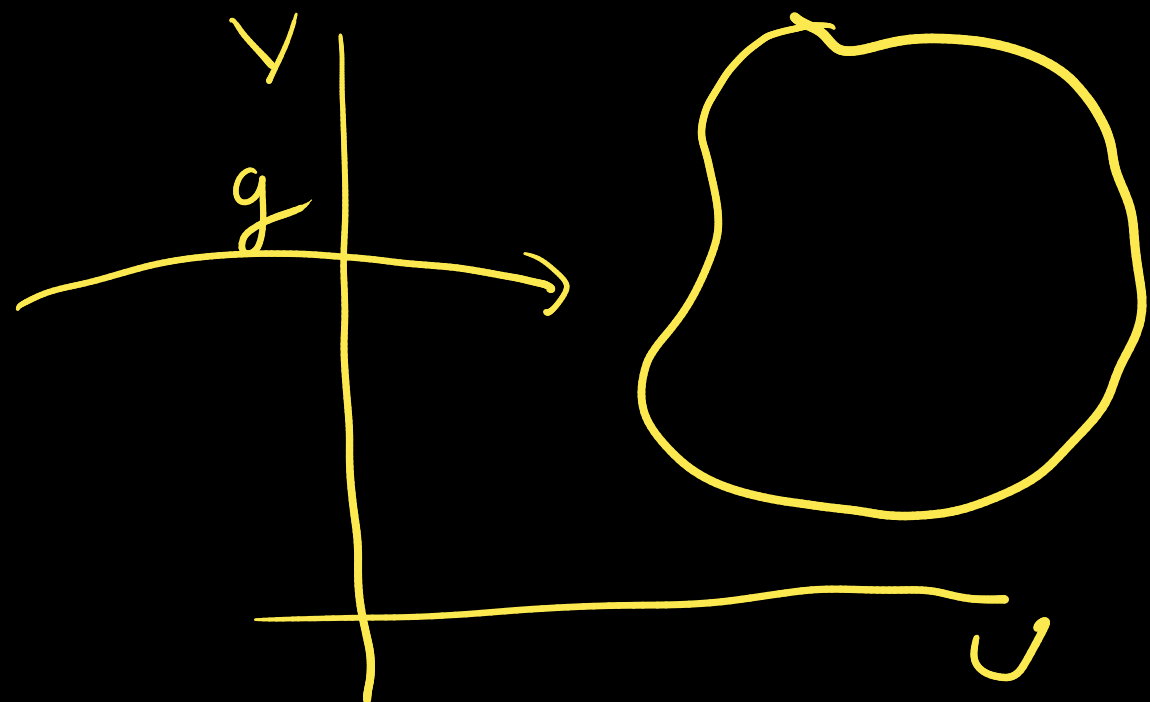
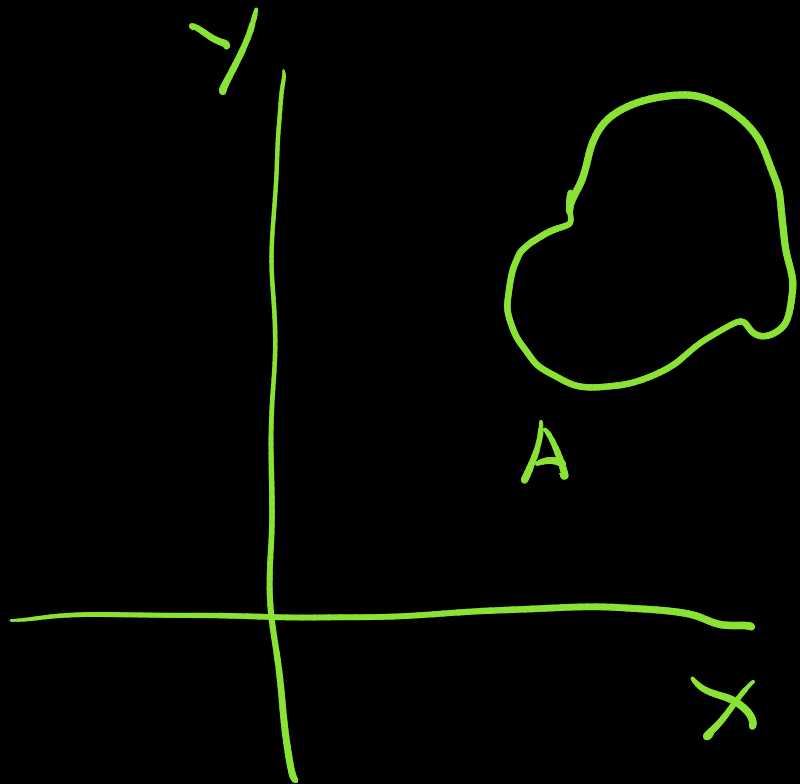
on $g(I)$

$$= f_x(g^{-1}(y)) \times \left| \frac{d}{dy} g^{-1}(y) \right|$$

How does this work in \mathbb{R}^2 ?



$$P((x, y) \in A) = \iint_A f_{xy}(x, y) dx dy$$



$$P\{(x, y) \in A\} = \iint_A f_{xy}(x, y) dx dy$$

too big if g stretched area

$$P\{(u, v) \in g(A)\} = \iint \underline{f_{xy}[x'(u, v), y'(u, v)]} du dv$$

need to undo stretching $g(A)$

stretching by g

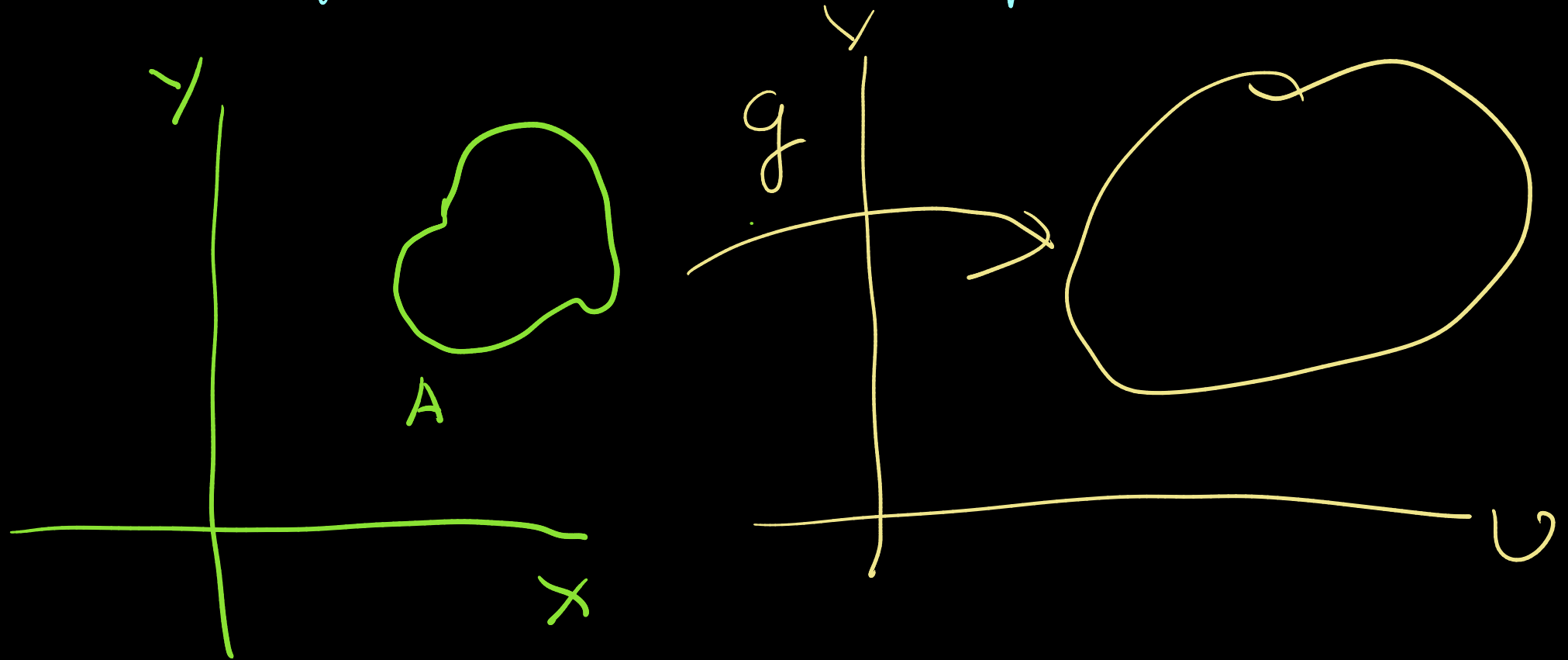
$$= \iint_{g(A)} \frac{f_{xy}(g^{-1}(u, v))}{\left| \frac{d(u, v)}{d(x, y)} \right|} du dv \quad \left. \vphantom{\frac{f_{xy}(g^{-1}(u, v))}{\left| \frac{d(u, v)}{d(x, y)} \right|}} \right\} \begin{array}{l} \text{stretching} \\ \text{by } g \end{array}$$

$$= \iint_{g(A)} f_{xy}(g^{-1}(u, v)) \underbrace{\|J_{g^{-1}}(u, v)\|}_{\text{shrinking by } g^{-1}} du dv$$

So: $f_{uv}(u, v) = f_{xy}(g^{-1}(u, v)) \underbrace{\|J_{g^{-1}}(u, v)\|}_{\substack{\text{Absolute value of} \\ \text{Jacobian determinant}}}$

Jacobian matrix
Jacobian determinant

What is $|J|$: Determinant of Jacobian Matrix



How much does g (stretch or shrink) area?

$$J_g(x, y) = \begin{bmatrix} \frac{du}{dx} & \frac{du}{dy} \\ \frac{dv}{dx} & \frac{dv}{dy} \end{bmatrix}$$

$$\underbrace{\left| \overbrace{J_g(x, y)}^{\text{determinant}} \right|}_{\text{absolute value}}$$

$$\underbrace{\left| \overbrace{\frac{du}{dx} \cdot \frac{dv}{dy} - \frac{dv}{dx} \frac{du}{dy}}^{\text{determinant}} \right|}_{\text{absolute value.}}$$

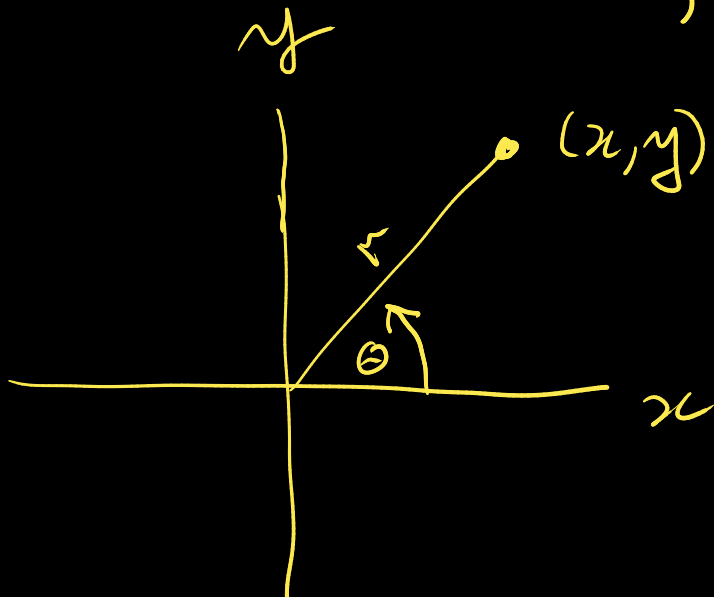
Example: Polar coordinates

Let (X, Y) have a bivariate standard normal distribution:

$$f(x, y) = \frac{1}{2\pi} e^{-\frac{1}{2}(x^2 + y^2)}$$

$$-\infty < x < \infty$$

$$-\infty < y < \infty$$



Let g be transformation
to polar co-ordinates.

$$r = \sqrt{x^2 + y^2}$$

$$\theta = \tan^{-1}(y/x)$$

Inverse easier

$$x = r \cos \theta$$

$$y = r \sin \theta$$

$$r \geq 0, \quad 0 \leq \theta < 2\pi$$

$$\begin{pmatrix} x \\ y \end{pmatrix} = g^{-1} \begin{pmatrix} r \\ \theta \end{pmatrix}$$

The arc-tangent of two arguments $\text{atan2}(y, x)$ returns the angle between the x-axis and the vector from the origin to (x, y) , i.e., for positive arguments $\text{atan2}(y, x) == \text{atan}(y/x)$.

$$\theta = \begin{cases} \tan^{-1}(y/x) & \text{if } x > 0 \\ \tan^{-1}(y/x) + \pi & \text{if } x < 0 \\ \frac{\pi}{2} \text{sgn}(y) & \text{if } x = 0, y \neq 0 \\ 0 & \text{if } x = y = 0 \end{cases}$$

$$J_{g^{-1}}(r, \theta) = \begin{bmatrix} \frac{dx}{dr} & \frac{dx}{d\theta} \\ \frac{dy}{dr} & \frac{dy}{d\theta} \end{bmatrix}$$

$$= \begin{bmatrix} \cos \theta & r \times (-\sin \theta) \\ \sin \theta & r \cos \theta \end{bmatrix}$$

$$\begin{aligned} x &= r \cos \theta \\ y &= r \sin \theta \end{aligned}$$

$$|J_{g^{-1}}| = r \cos^2 \theta - r(-\sin^2 \theta) = r(\cos^2 \theta + \sin^2 \theta) = r$$

$\|J_{g^{-1}}\| = |r| = r$ since $r > 0$ anyways.

$$f_{x,y}(x,y) = \frac{1}{2\pi} e^{-\frac{1}{2}(x^2+y^2)}$$

$$f_{R,\theta}(r,\theta) = \left(\frac{1}{2\pi} e^{-\frac{1}{2}r^2} \right) \times r, \quad \begin{array}{l} r > 0 \\ \underline{0 < \theta < 2\pi} \end{array}$$

See where $\frac{1}{2\pi}$ came from!

Now let $Z = R^2$

If one variable stays the same
you don't need to work out the Jacobian

$$\begin{pmatrix} z \\ \theta \end{pmatrix} = g \begin{pmatrix} R \\ \theta \end{pmatrix} = \begin{pmatrix} R^2 \\ \theta \end{pmatrix}$$

Note: Even if this is not 0
we still have $J_g = \frac{dz}{dr}$

$$J_g = \begin{bmatrix} \frac{dz}{dr} & \frac{dz}{d\theta} \\ \frac{d\theta}{dr} & \frac{d\theta}{d\theta} \end{bmatrix} = \begin{bmatrix} 2R & 0 \\ 0 & 1 \end{bmatrix} = 2R$$

$$f_{z,\theta}(z,\theta) = \left(\text{Substitute } z,\theta \text{ in } f_{r,\theta} \right) \times \left| \frac{dr}{dz} \right|$$

$$= \frac{1}{2\pi} e^{-\frac{1}{2}r^2} \times r \times \left| \frac{dr}{dz} \right|$$

$$\begin{aligned} z &= r^2 \\ \theta &= \theta \end{aligned}$$

$$= \frac{1}{2\pi} e^{-\frac{1}{2}z} \times \sqrt{z} \times \frac{1}{2} z^{-1/2}$$

$$r = \sqrt{z}$$

$$= \frac{1}{2\pi} \frac{1}{2} e^{-\frac{1}{2}z}, \quad \begin{aligned} z &> 0 \\ 0 &< \theta < 2\pi \end{aligned}$$

$$\frac{dr}{dz} = \frac{1}{2} z^{-1/2}$$

$$= \underbrace{f_{\theta}(\theta)}_{U(0, 2\pi)} \times \underbrace{f_z(z)}_{\text{Exponential}(1/2)}$$

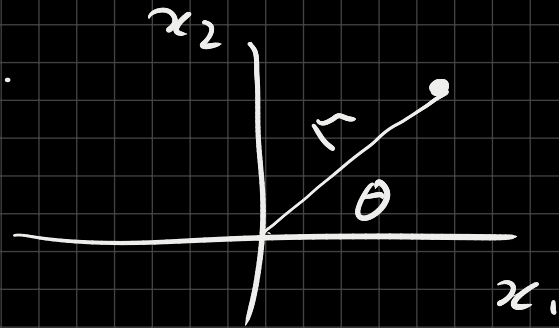
Independent.

Note: This gives us a way of
 generating random standard normals
 in pairs even though it is very difficult
 to generate them one at a time.

How?

Example for Wednesday Quiz

Let X_1, X_2 be independent with Exponential(λ) distribution.



Find the distribution of R and Θ

$$\text{where } X_1 = R \cos \Theta$$

$$X_2 = R \sin \Theta$$

$$f(x_1, x_2) = \lambda^2 e^{-\lambda x_1} e^{-\lambda x_2}, \quad x_1, x_2 > 0$$

Find $f_{R, \Theta}(r, \theta)$

How: 1) Substitute r, θ in $f(x_1, x_2)$

and 2) multiply by:

$$\|J_{g^{-1}}\| = \left\| \frac{\partial (x_1, x_2)}{\partial (r, \theta)} \right\| = \begin{vmatrix} \frac{\partial x_1}{\partial r} & \frac{\partial x_1}{\partial \theta} \\ \frac{\partial x_2}{\partial r} & \frac{\partial x_2}{\partial \theta} \end{vmatrix}$$

(OR)

divide by:

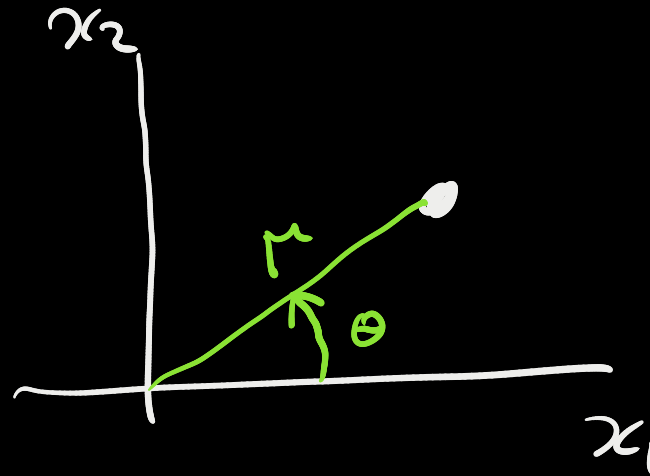
$$\|J_g\| = \left\| \frac{\partial (r, \theta)}{\partial (x_1, x_2)} \right\| = \begin{vmatrix} \frac{\partial r}{\partial x_1} & \frac{\partial r}{\partial x_2} \\ \frac{\partial \theta}{\partial x_1} & \frac{\partial \theta}{\partial x_2} \end{vmatrix}$$

$$f(x_1, x_2) = \lambda^2 e^{-\lambda x_1} e^{-\lambda x_2}, \quad x_1, x_2 > 0$$

$$x_1 = r \cos \theta$$

$$x_2 = r \sin \theta$$

Find $f_{R,\theta}(r, \theta)$



$$\begin{aligned} \frac{2(x_1, x_2)}{2(r, \theta)} &= \begin{vmatrix} \cos \theta & -r \sin \theta \\ \sin \theta & r \cos \theta \end{vmatrix} \\ &= r(\cos^2 \theta + \sin^2 \theta) = r \end{aligned}$$

$$f(x_1, x_2) = \lambda^2 e^{-\lambda x_1} e^{-\lambda x_2}, \quad x_1, x_2 > 0$$

$$x_1 = r \cos \theta$$

$$x_2 = r \sin \theta$$

$$f_{R,\theta}(r, \theta) = \lambda^2 e^{-\lambda r \cos \theta} e^{-\lambda r \sin \theta} \times r$$

$$= r \lambda^2 e^{-\lambda r (\cos \theta + \sin \theta)}$$

$$r > 0, \quad 0 < \theta < \frac{\pi}{2}$$

2) divide by

$$\|J_g\| = \left\| \frac{\partial (r, \theta)}{\partial (x_1, x_2)} \right\| = \left\| \begin{array}{cc} \frac{\partial r}{\partial x_1} & \frac{\partial r}{\partial x_2} \\ \frac{\partial \theta}{\partial x_1} & \frac{\partial \theta}{\partial x_2} \end{array} \right\|$$

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