

Time-Series-Assessment-2.R

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2022-04-03

#Code is in deep blue

#Loading of library packages to enable the functionality of my functions
library(tseries)

Warning: package 'tseries' was built under R version 4.0.5

Registered S3 method overwritten by 'quantmod':
method from
as.zoo.data.frame zoo

library(forecast)

Warning: package 'forecast' was built under R version 4.0.5

Q1

#.....#

#(a)

#Setting random number generator at value 639

set.seed(639)

#Running ARIMA Simulation

#n= number of values, 350 values in this case

#model list defines parameters of AR and MA series

#alpha 1=0.4, alpha 2=-0.1 for AR series

#beta1 =0.57

x=arima.sim(n=350,model=list(ar=c(0.4,-0.1),ma=0.57))

..... •

Explanation:

We are simulating a time series of length 350, and ARMA(2,1) with parameters $\alpha_1 = 0.4, \alpha_2 = -0.1$ and $\beta_1 = 0.57$ respectively

Also, the set.seed, ensure we are able to replicate our time series each time, without randomly generating different values each time.

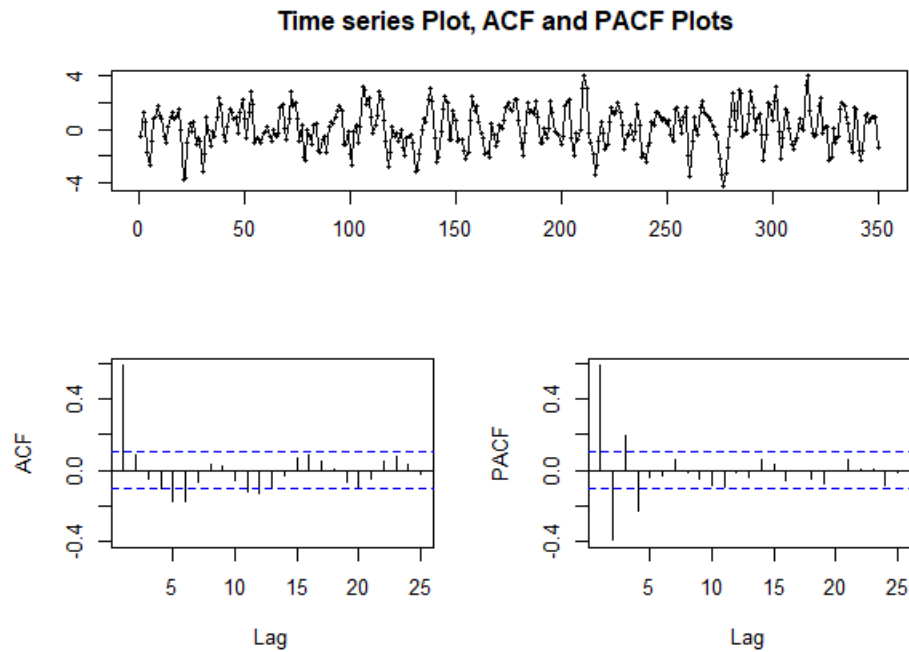
..... •

Q1(b)

#We use the `tsdisplay` function to plot the ARIMA time series simulation of `x`
#This produces a graph of the Time series, its respective ACF and PACF graphs

#code

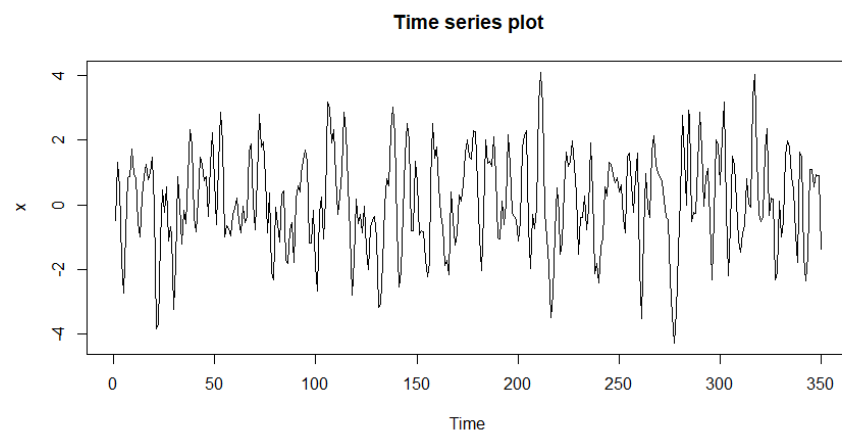
```
tsdisplay(x,main="Time series Plot, ACF and PACF Plots ")
```



Q1b)a #Can also be produced using `ts.plot()`, `acf()` and `pacf()`

```
ts.plot(x)
```

Output



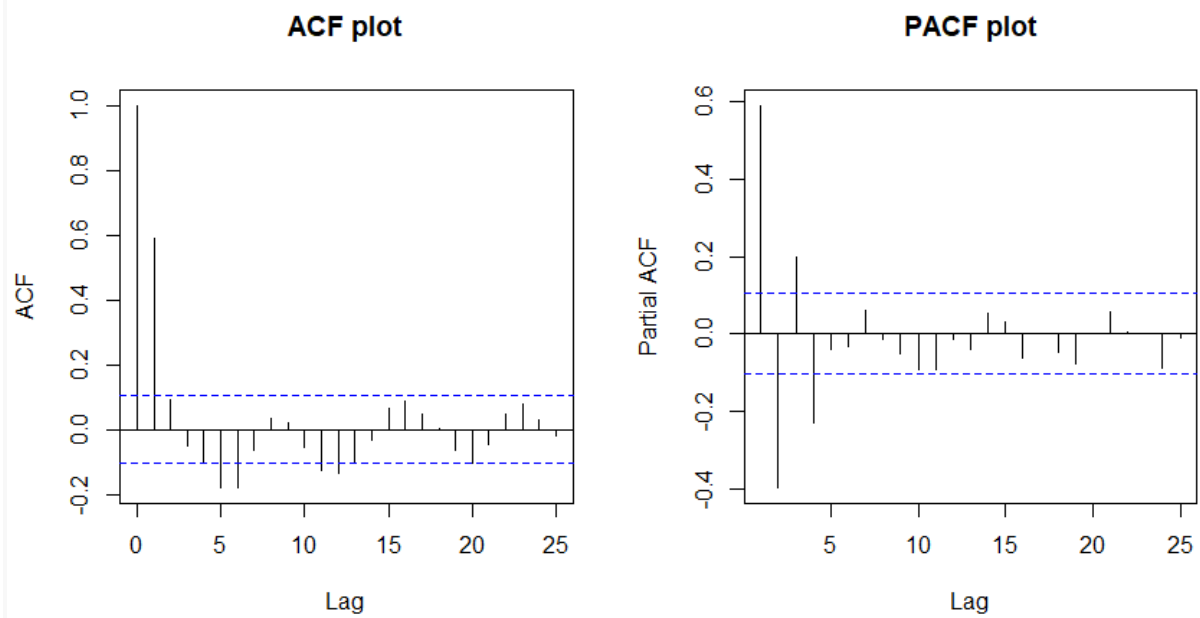
Q1b)b

#code:

```
par(mfrow=c(1,2))
```

```
acf(x)
```

```
pacf(x)
```



Q1(c)

*Our simulation of Time series is stationary, but I am checking for clarity
Running Dicky-Fuller Test to prove stationarity*

#code:

```
adf.test(x)
```

Output

```
## Warning in adf.test(x): p-value smaller than printed p-value
```

```
##
```

```
## Augmented Dickey-Fuller Test
```

```
##
```

```
## data: x
```

```
## Dickey-Fuller = -6.695, Lag order = 7, p-value = 0.01
```

```
## alternative hypothesis: stationary
```

Q2c)a

#Fitting the ARIMA time series to get ARMA(2,2)

#ARIMA(2,0,2) or ARMA(2,2)

#I have assigned ARMA(2,2) as fit 1

```
fit1=arima(x,order=c(2,0,2))
```

#Calling fit1

Output

```
fit1
```

```
##
```

```
## Call:
```

```
## arima(x = x, order = c(2, 0, 2))
```

```
##
```

```
## Coefficients:
```

```
##          ar1          ar2          ma1          ma2  intercept
```

```
##      -0.0655   -0.0207   1.0440   0.3108         0.0982
```

```
## s.e.    0.6073    0.2067   0.6063   0.4021         0.1229
```

```
##
```

```
## sigma^2 estimated as 1.129:  log likelihood = -518.38,  aic = 1048.77
```

#(b)

#ARIMA(2,0,1) or ARMA(2,1)

#Assigning ARMA(2,1) as fit2

```
fit2=arima(x,order = c(2,0,1))
```

#Calling fit2

Output

```
fit2
```

```
##
```

```
## Call:
```

```
## arima(x = x, order = c(2, 0, 1))
```

```
##
```

```
## Coefficients:
##          ar1      ar2      ma1  intercept
##      0.3775 -0.1402  0.5969    0.0986
## s.e.  0.0832   0.0726  0.0694    0.1188
##
## sigma^2 estimated as 1.131:  log likelihood = -518.64,  aic = 1047.28
```

Q1(c)

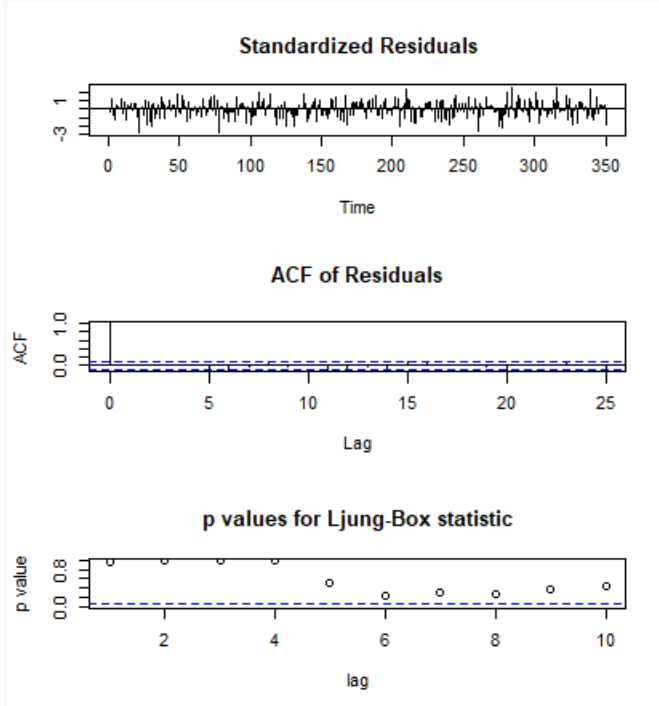
I choose model 2, ARMA(2,1) as it has the lowest AIC, Maximum likelihood value. Hence ARMA(2,1) is the best model.

ARMA(2,2) --→ log likelihood = -518.38, aic = 1048.77

ARMA(2,1) --→ log likelihood = -518.64, aic = 1047.28

```
Q1(d)
#Running a diagnostic test of fit2
tsdiag(fit2)
```

Output



Equation of model chosen:

ARMA(2,1)

$$y_t = 0.4y_1 - 0.1y_2 - 0.57\varepsilon_1$$

The output suggests some forecasting can be done, also, the ACF plots suggest white noise as well as the p-values of Ljung-Box statistic being greater than the significance value of 5%.

Also, since our p-values are greater than 5%, we can suggest, our null hypothesis is not true, hence no evidence of serial correlation amongst the fitted model (fit2).

Looking at the 2nd plot, we can observe the lags of the ACF residuals do not exceed the 95% level, hence model is quite a good fit!

We will look into this further in part(e)

Q1(e)

```
# I assumed my x value to be my chosen fitted ARIMA series
#fitdf= number of degrees of freedom to be subtracted if x is a series of
#residuals.
#In our case its 4, as our p=2 from AR terms and q=1 from MA terms and -1
#We normally don't tend to use it, but as I am using my fitted time series to
#check for serial correlation
#The question demanded a lag of length 10
#Type of Box.test is Ljung-Box
Box.test(residuals(fit2),fitdf =4,lag = 10,type = "Ljung-Box")
```

Output

```
##
## Box-Ljung test
##
## data: residuals(fit2)
## X-squared = 9.8832, df = 6, p-value = 0.1297
```

Explanation:

$$Q = T(T+2) \sum_{k=1}^s r_k^2 / (T-k) \rightarrow \chi_{s-p-q}^2 \quad (4.11)$$

X-squared represents the Q value for Ljung-Box test (Portmanteau statistic), so our Q value is 9.8832, total degree of freedom = 10-3-1=6, p-value = is probability of our Q value occurring.

So in this context:

Our p-value > 5%, hence adequacy of fitted ARMA(p,q) should be re-analyzed.

H_0 = Residuals of white noise observed

H_1 = No residuals of white noise observed

T = 350

```
#using the checkresiduals() to see if my answer is right
checkresiduals(fit2)
```

Output

```
##
## Ljung-Box test
##
## data: Residuals from ARIMA(2,0,1) with non-zero mean
```

```
## Q* = 9.8832, df = 6, p-value = 0.1297
```

```
##
```

```
## Model df: 4.    Total lags used: 10
```

```
#Since p-value is greater than 5%, we need to use the GARCH model
```

```
#.....#
```


Question 2

Generate a random sample of y !

a) $g(y) = \frac{1}{2} \sin(y) \quad ; 0 < y < \pi$

$$G_Y(u) = \int_{-\infty}^{\infty} g_Y(y) dy$$

$$= \int_0^u \frac{1}{2} \sin(y) dy$$

$$= \frac{1}{2} \int_0^u \sin(y) dy$$

$$= \frac{1}{2} \left[-\cos(y) \right]_0^u$$

$$= \frac{1}{2} \left[\cos(y) \right]_u^0$$

$$= \frac{1}{2} \left[1 - \cos(u) \right]$$

$$G_Y(u) = \frac{1}{2} (1 - \cos(u))$$

Making u the subject:

$$G_Y(u) = y$$

$$y = \frac{1}{2} (1 - \cos(u))$$

$$2y = 1 - \cos(u)$$

$$1 - 2y = -\cos(u)$$

$$2y - 1 = \cos(u)$$

$$\arccos(2y - 1) = u$$

$$\therefore u = \arccos(2y - 1) =$$

b) Method of rejection sampling

$$Y \sim g(y) = \frac{1}{2} \sin(y)$$

$$X \sim f(x) = \frac{1}{\pi} \quad \} \quad X \sim U(0, \pi)$$

We know the $G_Y^{-1}(u) = \arccos(2u - 1)$
 $X = Y$

$$\sup_y \frac{g(y)}{f(y)} = M$$

$$M = \sup_y \frac{g(y)}{f(y)}$$

$$= \frac{\frac{1}{2} \sin(y)}{\frac{1}{\pi}}$$

$$= \frac{\pi}{2} \sin(y) \Rightarrow \ln\left(\frac{\pi}{2} \sin(y)\right)$$

$$\frac{d}{dy} \left[\ln\left(\frac{\pi}{2} \sin(y)\right) \right] = \frac{\cos(y)}{\sin(y)}$$

$$\frac{\cos(y)}{\sin(y)} = 0$$

$$\cos(y) = 0$$

$$y = \arccos(0)$$

$$y = \frac{\pi}{2} = n\frac{\pi}{2} \text{ for } n=1, 2, 3, \dots$$

$$\frac{d^2y}{dy^2} = \frac{-\cos^2(y)}{\sin^2(y)} - 1 < 0$$

$\frac{d^2y}{dy^2} < 0$, hence $y = \frac{\pi}{2}$ is a maximum turning point.

$$m = \sup_y \left(\frac{g(y)}{f(y)} \right)$$

$$= \left(\frac{1}{2} \sin\left(\frac{\pi}{2}\right) * \pi \right)$$

$$= \frac{\pi}{2} \sin\left(\frac{\pi}{2}\right)$$

$$c_1 \frac{g(y)}{m f(y)} = \frac{\frac{1}{2} \sin(y)}{\frac{\pi}{2} \sin\left(\frac{\pi}{2}\right) * \pi} = \frac{\sin(y)}{2} \cdot \frac{1}{\frac{1}{2} * \frac{\pi}{2} \sin\left(\frac{\pi}{2}\right)}$$

$$= \frac{\sin(y)}{\frac{1}{2} \sin\left(\frac{\pi}{2}\right)} = \frac{\sin(y)}{2} * \frac{1}{\sin\left(\frac{\pi}{2}\right)}$$

$$\frac{f(y)}{n(y)} = \frac{f_n(y)}{f_n(\pi/2)} = f_n(y)$$

$$\text{Since } f_n(\pi/2) = 1$$

Algorithm

- ① Generate a random number U_1
- ② Set $Y = \arccos(2U_1 - 1)$ [Solved in (a) by inversion]
- ③ Generate another random number U_2
- ④ If $U_2 \leq f_n(Y)$, then set $X=Y$,
or otherwise start all over from ①

Question 3

$$f(x) = \frac{e^{-x/2}}{(\phi(3) - 0.05)\sqrt{2\pi}}, \quad 0 < x < 3$$

$$g(y) = 2\alpha e^{-2y}, \quad 0 < y < 3$$

$$(a) \int_{-\infty}^{\infty} g_p(y) dy = 1$$

$$\int_0^3 2\alpha e^{-2y} dy = 1$$

$$\left[\frac{2\alpha e^{-2y}}{-2} \right]_0^3 = 1$$

$$\left[-\alpha e^{-2y} \right]_0^3 = 1$$

$$\left[\alpha e^{-2y} \right]_3^0 = 1$$

$$\alpha e^{-0} - \alpha e^{-6} = 1$$

$$\alpha - \alpha e^{-6} = 1$$

$$\alpha(1 - e^{-6}) = 1$$

$$\alpha = \frac{1}{1 - e^{-6}}$$

$$\therefore \alpha = (1 - e^{-6})^{-1}$$

Hence distribution function for Y is $\text{Exp}(2)$

$$\text{Since } (1 - e^{-6})^{-1} \approx 1$$

$$\text{p.d.f} = \begin{cases} 2(1 - e^{-6})^{-1} e^{-2y}, & 0 < y < 3 \\ 0, & \text{otherwise} \end{cases}$$

b)

Algorithm for generating samples of random variable Y using standard uniform $U(0,1)$
 $Y \sim \text{Exp}(2)$

$$G_Y(u) = \int_{-\infty}^u g_Y(y) dy$$

$$= \int_0^u g_Y(y) dy$$

$$= \int_0^u 2(1 - e^{-6})^{-1} e^{-2y} dy$$

using integral by substitution!

$$\frac{2}{1 - e^{-6}} \int_0^u e^{-2y} dy$$

$$\text{let } u = -2y$$

$$\frac{du}{dy} = -2$$

$$-\frac{du}{2} = dy$$

$$\frac{2}{1 - e^{-6}} \int_0^u e^u \frac{du}{-2} \Rightarrow \frac{-1}{1 - e^{-6}} \int_0^u e^u du$$

$$\frac{-1}{1-e^{-t}} \int_0^u e^u du$$

$$\frac{-1}{1-e^{-t}} \left[e^u \right]_0^u$$

$$\frac{-1}{1-e^{-t}} (e^u - e^0) = \frac{-e^u + e^0}{1-e^{-t}} = \frac{1-e^u}{1-e^{-t}}$$

$$= \frac{1-e^{-2y}}{1-e^{-t}} \quad \text{since } u=2y$$

$$\therefore G_y(u) = \frac{1-e^{-2y}}{1-e^{-t}} \quad \text{let } G_y(u) = u$$

$$\text{Setting: } \frac{1-e^{-2y}}{1-e^{-t}} = u$$

$$1-e^{-2y} = u(1-e^{-t})$$

$$1-u(1-e^{-t}) = e^{-2y}$$

$$\ln[1-u(1-e^{-t})] = -2y$$

$$y = - \frac{\ln[1-u(1-e^{-t})]}{2}$$

$$y = - \frac{\ln[1-u(1-e^{-t})]}{2}$$

(c)

X has p.d.f of $f(x)$, and the inversion is hard.

Y has p.d.f of $g(x)$, and the inversion is easy.

Let's assume:

$$\frac{f(x)}{g(x)} \leq M < \infty \quad \text{for all values of } x.$$

Generate a random variable Y with p.d.f $g(y)$.

$$f(x) = \frac{e^{-x^2/2}}{(\Phi(1) - 0.5)\sqrt{2\pi}}$$

$$g(x) = 2(1-e^{-x})^{-1}e^{-2x}$$

let $x = Y$ ↗

$$M = \sup_x \frac{f(x)}{g(x)}$$

We will use MLE to get a proper estimate of the maximum value of x :

$$= \frac{e^{-x^2/2}}{(\Phi(1) - 0.5)\sqrt{2\pi}} \rightarrow \frac{(1-e^{-x})}{2e^{-2x}}$$

$$= \frac{e^{-x^2/2 + 2x} (1-e^{-x})}{2(\Phi(1) - 0.5)\sqrt{2\pi}}$$

$$\ln \left[\frac{e^{-x^2/2 + 2x} (1-e^{-6})}{(\Phi(3) - 0.5) \sqrt{2\pi} (2)} \right]$$

$$-\frac{x^2}{2} + 2x - \ln [(\Phi(3) - 0.5) \sqrt{2\pi}] + \ln [1-e^{-6}]$$

$$\frac{d}{dx} \left[-\frac{x^2}{2} + 2x \right] - \frac{d}{dx} [(\Phi(3) - 0.5) \sqrt{2\pi}] + \frac{d}{dx} (1-e^{-6})$$

$$\frac{d}{dx} = \frac{-2x}{2} + 2$$

Set to 0.

$$\frac{d}{dx} = 0; \quad -\frac{x}{1} + 2 = 0, \quad -x + 2 = 0$$

$$\therefore x = 2 \quad (\text{Turning point})$$

Maximum of x occurs at 2

$$\frac{d^2}{dx^2} = -1 \quad (\text{Hence maximum occurs at } x=2)$$

$$\frac{f(x)}{Mg(x)} = \frac{e^{-x^2/2 + 2x} (1-e^{-6})}{2(\Phi(3) - 0.5) \sqrt{2\pi} \cdot 2}$$

$$= \frac{e^{-x^2/2 + 2x} (1-e^{-6})}{4\sqrt{2\pi}(\Phi(3) - 0.5)} = e^{-x^2/2 + 2x + 2} \cdot \frac{1-e^{-6}}{4\sqrt{2\pi}(\Phi(3) - 0.5)}$$

$$M = \frac{e^{-x^2/2 + 2x} (1 - e^{-6})}{2(\Phi(3) - 0.5)\sqrt{2\pi}}, \text{ where } x=2$$

$$M = \frac{e^{-2^2/2 + 2(2)} (1 - e^{-6})}{2(\Phi(3) - 0.5)(\sqrt{2\pi})}$$

$$M = \frac{e^{-2+4} (1 - e^{-6})}{2(\Phi(3) - 0.5)(\sqrt{2\pi})}$$

$$= \frac{e^{+2} (1 - e^{-6})}{2(\Phi(3) - 0.5)\sqrt{2\pi}}$$

$$\frac{f(x)}{Mg(x)} = \frac{e^{-x^2/2} (1 - e^{-6}) e^{2x}}{2\sqrt{2\pi} (\Phi(3) - 0.5) e^2 (1 - e^{-6})}$$

$$= e^{-x^2/2 + 2x - 2}$$

$$= e^{-x^2/2 + 2x - 2}$$

Algorithm!

① Generate a random number u_1

② Set $y = -\frac{1}{2} \ln[1 - u_1(1 - e^{-6})]$

③ Generate another random number u_2

④ If $u_2 \leq e^{-y^2/2 + 2y - 2}$, then set $x = y$. Otherwise
Go to step ①.

c) Algorithm to perform run n(c)

① Generate a random number u_1

$$u_1 = 0.86$$

② Set $y = -\frac{1}{2} \ln [1 - u_1(1 - e^{-6})]$

$$y = -\frac{1}{2} \ln [1 - 0.86(1 - e^{-6})]$$

$$y = 0.9755$$

③ Generate another random number u_2

$$u_2 = 0.34, \text{ where } h(y) = e^{-y/2 + 2y - 2}$$

④ If $u_2 \leq e^{-y/2 + 2y - 2}$

$$u_2 \leq e^{-0.9755/2 + 2(0.9755) - 2}$$

$$u_2 \leq 0.59167$$

Since $u_2 \leq 0.59167$, we ~~go to step 1~~ ^{end here!}.

~~and start the process over!~~

$$\text{So } u \leq h(y) \Rightarrow M g(y) u \leq f(y)$$

where $M g(y) u$ is a ^{random} point below $M g(y)$,

Hence we accept this point below $f(y)$.

(d) Z = number of attempts until we accept x
 p = probability (accept x at each attempt)

$$p = P(\text{Accept } x \text{ at each attempt}) = \int_{-\infty}^{\infty} h(y) g(y) dy$$

$$= \int_{-\infty}^{\infty} \frac{f(y)}{mg(y)} g(y) dy$$

$$= \frac{1}{M} \int_{-\infty}^{\infty} f(y) dy$$

$$= \frac{1}{M}$$

$$M = \frac{e^2 (1 - e^{-6})}{2(\sqrt{2\pi}) (\Phi(3) - 0.5)}$$

$$\frac{2\sqrt{2\pi} (\Phi(3) - 0.5)}{e^2 (1 - e^{-6})}$$

$$f = \frac{2\sqrt{2\pi} (\Phi(3) - 0.5)}{e^2 - e^{-4}}$$

$$\Phi(3) = 0.9986501$$

$$f = 0.3392$$

• $E(z) = \text{number of pairs of random numbers.}$

$$= \frac{1}{p} \approx M$$

$$= 2.948$$

$\approx 2.95 \approx 3$ pairs of Random numbers.