

Lecture 19: 08 November 2023

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19.1 Notions of Convergence

Definition 19.1 (Almost Sure Convergence). A sequence of random variables X_1, X_2, \dots, X_n converges **almost surely** to a random variable X if

$$\Pr \left(\lim_{n \rightarrow \infty} |X_n - X| > \epsilon \right) = 0, \forall \epsilon > 0. \quad (19.1)$$

and we denote this as $X_n \xrightarrow{\text{as}} X$.

Equivalently, we can write (19.1) as

$$\Pr (|X_n - X| > \epsilon) \rightarrow 0, n \geq N, N \rightarrow \infty. \quad (19.2)$$

Definition 19.2 (Convergence in Probability). We say that X_n converges to X in probability and write $X_n \xrightarrow{\text{P}} X$ if

$$\lim_{n \rightarrow \infty} \Pr (|X_n - X| > \epsilon) = 0 \quad (19.3)$$

or equivalently

$$\Pr (|X_n - X| > \epsilon) \rightarrow 0, n \rightarrow \infty. \quad (19.4)$$

Definition 19.3 (Convergence in Distribution). We say that a sequence of random variables $\{X_n\}$ converges in distribution to another random variables X if for all x where F_X is continuous,

$$X_n \xrightarrow{\text{d}} X. \quad (19.5)$$

As an example, consider X_1, X_2, \dots, X_n which are iid samples drawn from an independent distribution f_X . Then, $F_{X_n}(x) = F_X(x)$ for all x and n , but X_n does not converge to X almost surely and in probability.

Theorem 19.4 (Strong Law of Large Numbers). *If X_1, X_2, \dots, X_n are iid with $\mathbb{E}[X_1] = \mu < \infty$, then*

$$\frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{\text{as}} \mu \quad (19.6)$$

We state two important results for these notions of convergence.

Theorem 19.5 (Weak Law of Large Numbers). *If X_1, X_2, \dots, X_n are iid with $\mathbb{E}[X_1] = \mu < \infty$, then*

$$\frac{1}{n} \sum_{i=1}^n X_i \xrightarrow{\text{P}} \mu. \quad (19.7)$$

19.2 Mean Squared Error of DME With Location Families

In the DME problem,

$$\mathbb{E}[Y_i] = \Pr(X_i \leq \theta) \quad (19.8)$$

$$= F_X(\theta - \mu) \quad (19.9)$$

$$\implies \frac{1}{n} \sum_{i=1}^n Y_i \xrightarrow{\text{as}} F_X(\theta - \mu). \quad (19.10)$$

We quote another result for F_X^{-1} .

Theorem 19.6 (Continuous Mapping Theorem). *If $g : \mathbb{R} \rightarrow \mathbb{R}$ is continuous and differentiable, then the following identities hold.*

$$1. X_n \xrightarrow{\text{as}} X \implies g(X_n) \xrightarrow{\text{as}} g(X).$$

$$2. X_n \xrightarrow{\text{P}} X \implies g(X_n) \xrightarrow{\text{P}} g(X).$$

$$3. X_n \xrightarrow{\text{d}} X \implies g(X_n) \xrightarrow{\text{d}} g(X).$$

Using Theorem 19.6, we see that

$$F_X^{-1} \left(\frac{1}{n} \sum_{i=1}^n Y_i \right) \xrightarrow{\text{as}} F_X^{-1}(F_X(\theta - \mu)) \quad (19.11)$$

$$\implies \hat{\mu} \xrightarrow{\text{as}} \mu. \quad (19.12)$$

We also make use of the central limit theorem.

Theorem 19.7 (Central Limit Theorem). *If X_1, X_2, \dots, X_n are iid samples with $\mathbb{E}[X_i] = \mu$ and $\text{Var}X_i = \sigma^2$, then*

$$\frac{\sum_{i=1}^n (X_i - \mu)}{\sqrt{n\sigma^2}} \xrightarrow{\text{d}} \mathcal{N}(0, 1). \quad (19.13)$$

From (19.13), we have

$$Z \triangleq \sqrt{n} \left(\left(\frac{1}{n} \sum_{i=1}^n X_i \right) - \mu \right) \xrightarrow{\text{d}} \mathcal{N}(0, \sigma^2) \quad (19.14)$$

$$\implies \mathbb{E}[Z^2] = n(\text{MSE}) \rightarrow \sigma^2 \quad (19.15)$$

$$\implies \text{MSE} \rightarrow \frac{\sigma^2}{n}. \quad (19.16)$$

However, we still cannot conclude anything about the MSE of $\hat{\mu}$. We require another tool.

19.3 Delta Method

Suppose $\phi : \mathbb{R} \rightarrow \mathbb{R}$ is differentiable $\{X_i\}_{i=1}^n$ is a sequence of random variables and $\{r_i\}_{i=1}^n$ is a sequence of real numbers such that $r_n \rightarrow \infty$ as $n \rightarrow \infty$. Further, suppose that

$$r_n (X_n - \theta) \xrightarrow{d} X. \quad (19.17)$$

Then,

$$r_n (\phi(X_n) - \phi(\theta)) \xrightarrow{d} \left(\frac{d\phi}{dt} \Big|_{t=\theta} \right) X \quad (19.18)$$

Using the Delta Method for our problem, define

$$Y_n \triangleq \frac{1}{n} \sum_{i=1}^n X_i. \quad (19.19)$$

To estimate $\phi(\mu)$, we use the *plugin estimator*

$$\phi(\hat{\mu}) = \phi\left(\frac{1}{n} \sum_{i=1}^n Y_i\right). \quad (19.20)$$

Hence,

$$\sqrt{n} (Y_n - \mu) \xrightarrow{d} Y \sim \mathcal{N}(0, \sigma^2) \quad (19.21)$$

$$\implies \sqrt{n} (\phi Y_n - \phi(\mu)) \xrightarrow{d} (\phi'(\mu)) Y \quad (19.22)$$

$$n (\text{MSE}) \xrightarrow{d} 2\mu Y \quad (19.23)$$

$$n \text{MSE}(\mu^2) = 4\mu^2 \sigma^2. \quad (19.24)$$