

Appendix A: model diagnostics

We used Stan (Carpenter et al., 2017) as the probabilistic language behind the estimation of our Bayesian models in this study, with brms as its R interface (Bürkner, 2017). This language implements the Markov Chain Monte-Carlo (MCMC) algorithm to explore the posterior distribution of the model. Broadly, this algorithm is used to iteratively sample the joint sampling space of the parameters to be estimated in the model, and compute, for each value sampled, its likelihood under some probability distribution previously defined. We run 4 MCMC chains, each 1000 iterations long each. The correct performance of this algorithm is critical to the quality of the statistical evidence to which the outcomes of the model lead.

One way to diagnose the behaviour of MCMC is to inspect whether the different MCMC chains (if more than one) have converged to a similar region of the posterior. The Gelman-Rubin diagnostic (\hat{R} or R-hat Gelman & Rubin, 1992) provides a measure of chain convergence by comparing the variance within each chain *versus* the variance between each chain. Both are expected to be identical when chains have perfectly converged, so that $\hat{R} = 1$. Values lower than 1.01 are recommended, while values higher than 1.05 indicate that chains might have trouble converging and therefore the estimated parameters must be taken with caution. Figure 1 (A) shows the distribution of \hat{R} values for the coefficients of the fixed effect of our models, which we used for statistical inference. Most values are lower than 1.01, and never higher than 1.05, which provides evidence of successful MCMC convergence.

Another diagnostic of good MCMC converge is the ratio of effective sample size to total sample size (N_{eff}/N), which indicates the proportion of samples in the chain that resulted from a non-divergent transition. Values closer to 1 are ideal, as they indicate that all posterior samples from the MCMC were used to estimate the posterior distribution of the parameter. Values larger than 0.1 are recommended. Figure 1 (B) shows the distribution of the effective sample sizes of the coefficients of the fixed effects in our models. Most values are larger than 0.1, although model 0 (\mathcal{M}_0) accumulates most effective sample sizes close to 0.1.

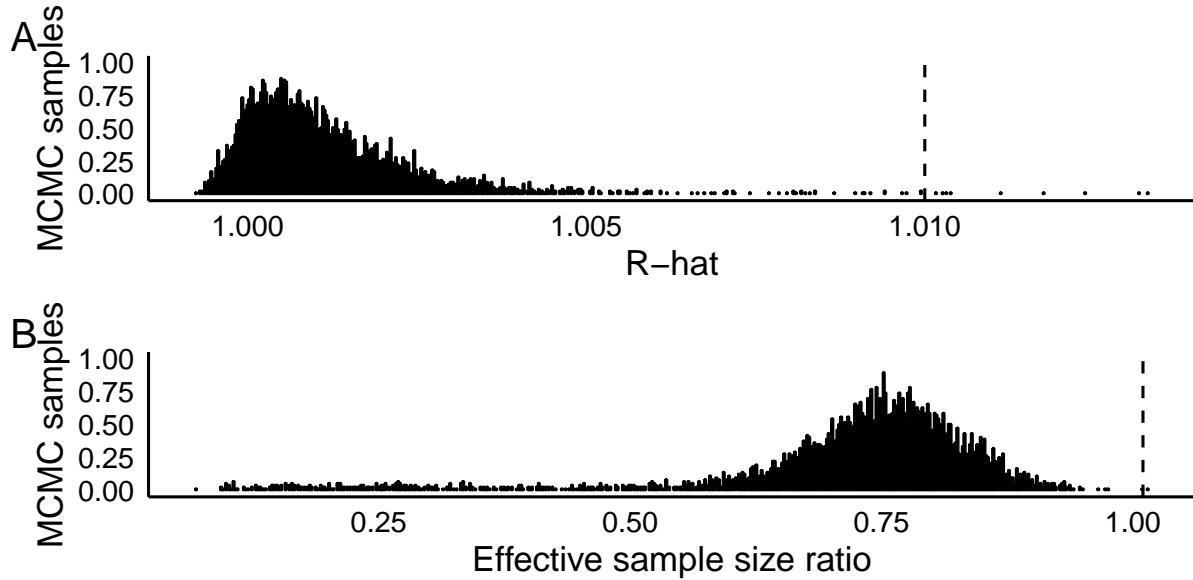


Figure 1: MCMC convergence diagnostic of all parameters in the model. Each dot represents the score of one parameter. (A) Distribution of the Gelman-Rubin (R-hat) scores. (B) Distribution of the ratio of effective sample size.

23 **Posterior draws and bi-variate scatterplots**

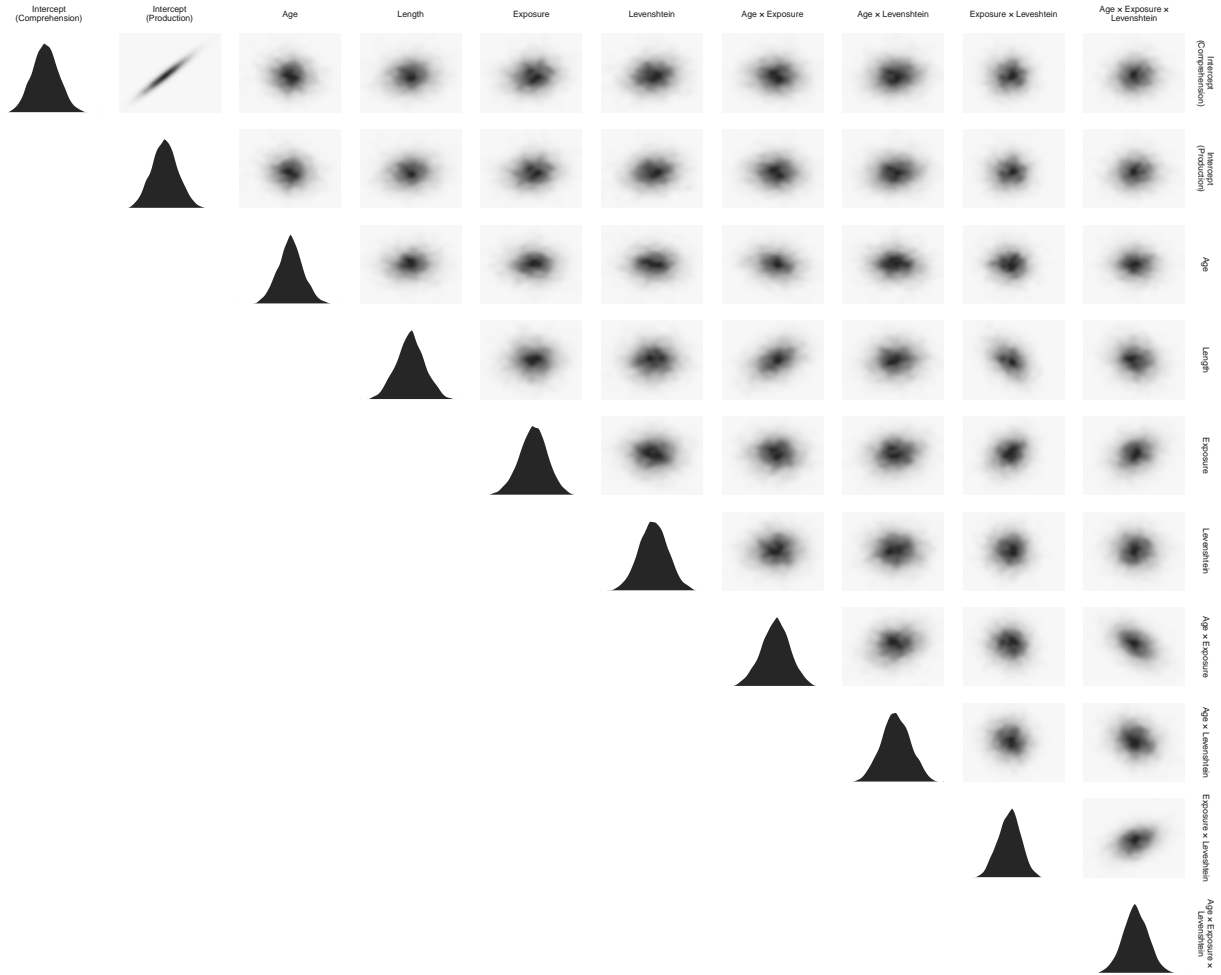


Figure 2: Marginal distribution and bi-variate scatterplot of posterior samples for the fixed regression coefficients in Model 3.

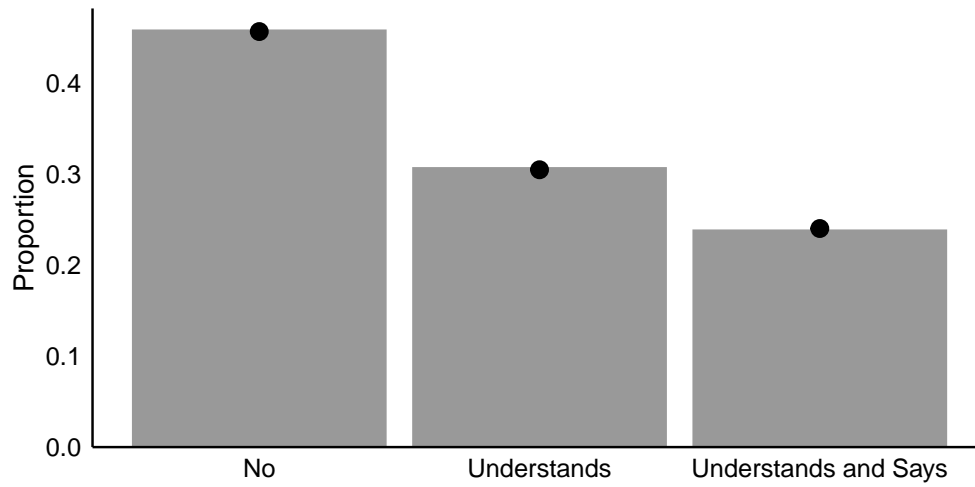
24 **Posterior-predictive checks**

Figure 3: Model posterior predictive checks (PPC). Bars indicate the observed proportion of responses to each category (No, Understands, and Understands and Says). Blue dots and error bars represent the mean proportion of responses simulated from the posterior for each category, and its 95% interval.

Table 1: Posterior distribution of regression coefficients of the model including the *Exposure* composite predictor, and of the model including lexical frequency (*Frequency*) and degree of exposure (*DoE*) separately. Median: median of the posterior distribution in the probability scale. 95% HDI: 95% highest density interval of the distribution. $p(\text{ROPE})$: overlap between the 95% HDI and the ROPE, indicating the posterior probability that the true value of the coefficient is equivalent to zero. In the right column, we show a graphical representation of the median and 95% HDI of each coefficient, with a reference dotted line indicating the location of zero.

	β	95% HDI	$p(\text{ROPE})$
Model: Exposure			
Length (+1 SD, 1.56 phonemes)	0.485	[0.478, 0.49]	.000
Age (+1 SD, 4.86 months)	0.600	[0.589, 0.612]	.000
Exposure (+1 SD, 1.81)	0.558	[0.549, 0.566]	.000
Cognateness (+1 SD, 0.26)	0.514	[0.502, 0.525]	.194
Exposure \times Cognateness	0.486	[0.483, 0.488]	.000
Age \times Exposure	0.518	[0.509, 0.526]	.000
Age \times Cognateness	0.503	[0.5, 0.507]	.936
Age \times Exposure \times Cognateness	0.495	[0.493, 0.497]	.874
Model: Frequency & DoE			
Age (+1 SD, 4.86, months)	0.600	[0.588, 0.612]	.000
Phonemes (+1 SD, 1.56 phonemes)	0.486	[0.48, 0.492]	.000
Frequency (+1 SD, 0.19)	0.527	[0.515, 0.538]	.000
DoE (+1 SD, 0.3)	0.557	[0.549, 0.565]	.000
Cognateness (+1 SD, 0.26)	0.516	[0.505, 0.527]	.059
Age \times DoE	0.518	[0.51, 0.527]	.000
DoE \times Cognateness	0.486	[0.483, 0.489]	.000
Age \times Cognateness	0.504	[0.501, 0.507]	.865
Age \times DoE \times Cognateness	0.495	[0.493, 0.497]	.884

Appendix B: frequency and language exposure as separate predictors

As a robustness check, we fit a model similar to the one described in the main manuscript, but including lexical frequency and language degree of exposure as separate predictors, instead of the composite measure *Exposure*. Language degree of exposure (*DoE*) was included in interaction with *Age* and *Cognateness*, while lexical frequency (*Frequency*) was included as a main effect. Table 1 shows a comparison between the posterior distribution of the regression coefficients of both models. Overall, results are equivalent.

Table 2: Sample of items included in the BVQ questionnaire and their syllabified SAMPA transcriptions in Catalan and Spanish

Translation	Item	X-SAMPA	Syllables	Item	X-SAMPA	Syllables
bib	pitet	pi"tEt	2	babero	ba"Be.4o	3
man	home	"O.m@	2	hombre	"om.b4e	2
butterfly	papallona	p@.p@"Lo.n@	4	mariposa	ma.4i"po.sa	4
book	conte	"kon.t@	2	cuento	"kwen.to	2
paper	paper	p@"pe	2	papel	pa"pel	2
eye	ulls	uLs	1	ojos	"o.xos	2
red	vermell	b@r"mEL	2	rojo	"ro.xo	2
potty	orinal	u.4i"na5	3	orinal	o.4i"nal	3
skate (object)	patins	p@"tins	2	patines	pa"ti.nes	3
crocodile	cocodril	ku.ku"D4i5	3	cocodrilo	ko.ko"D4i.lo	4
plate	plat	p5at	1	plato	"pla.to	2
pear	pera	"pE.4@	2	pera	"pe.4a	2
dinosaur	dinosaure	di.nu"saw.4@	4	dinosaurio	di.no"saw.4jo	4
witch	bruixa	"b4u.S@	2	bruja	"b4u.xa	2
coin	cèntim	"sEn.tim	2	céntimo	Ten.ti.mo	3

Appendix C: frequency and language exposure as separate predictors

We define syllable frequency as the rate of appearance of individual syllables in the word-forms included in the Barcelona Vocabulary Questionnaire (BVQ) (Garcia-Castro et al., 2023). Each item corresponds to a Catalan or Spanish word, and has an associated phonological transcription in X-SAMPA format (Wells, 1995). These transcriptions are syllabified. Some examples:

Most Catalan and Spanish words had two syllables, with Spanish words having three and four syllables more often than Catalan words. Less than 1% of the words included in the analyses presented in the main body of the manuscripts had five syllables. No words had more than five syllables (see Figure 4). We extracted lexical frequencies from the English corpora in the CHILDES database (MacWhinney, 2000; Sanchez et al., 2019). Using the Catalan and Spanish corpora was not possible due to the low number of children and tokens included in the corpora.

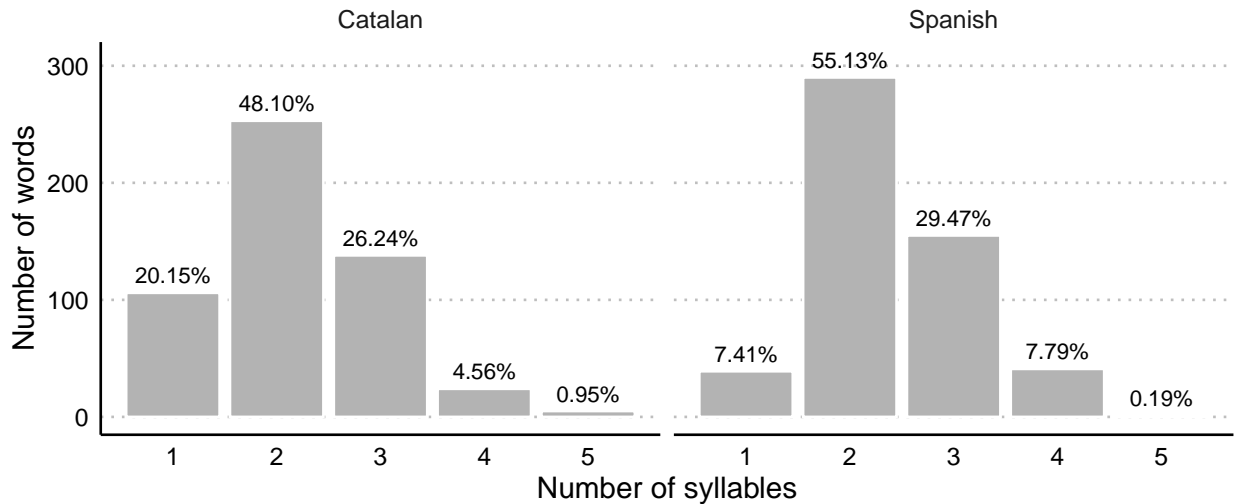


Figure 4: Distribution of the number of syllables in Catalan and Spanish

We now present how syllable frequencies were calculated. Every exposure to a word-form also counts as a exposure to each of the syllables that make up such word. Every time a child hears the word *casa* [house], they are exposed to the syllables *ca* and *sa*. Syllables that appear embedded in words with higher frequency will also be more frequent. To compute the relative frequency of each syllable in Catalan and Spanish (i.e., how many times the syllables appears

in every million words in Catalan or Spanish speech), we summed the relative lexical frequency in CHILDES of every word that contains such syllable in the corresponding language. Figure 5 shows the distribution of frequencies across syllables in Catalan and Spanish. In the log10 scale, syllable frequencies in Catalan and Spanish followed a slightly asymmetric distribution, with most syllables scoring around 1,000 counts per million, and a longer tail to the right of the distribution.

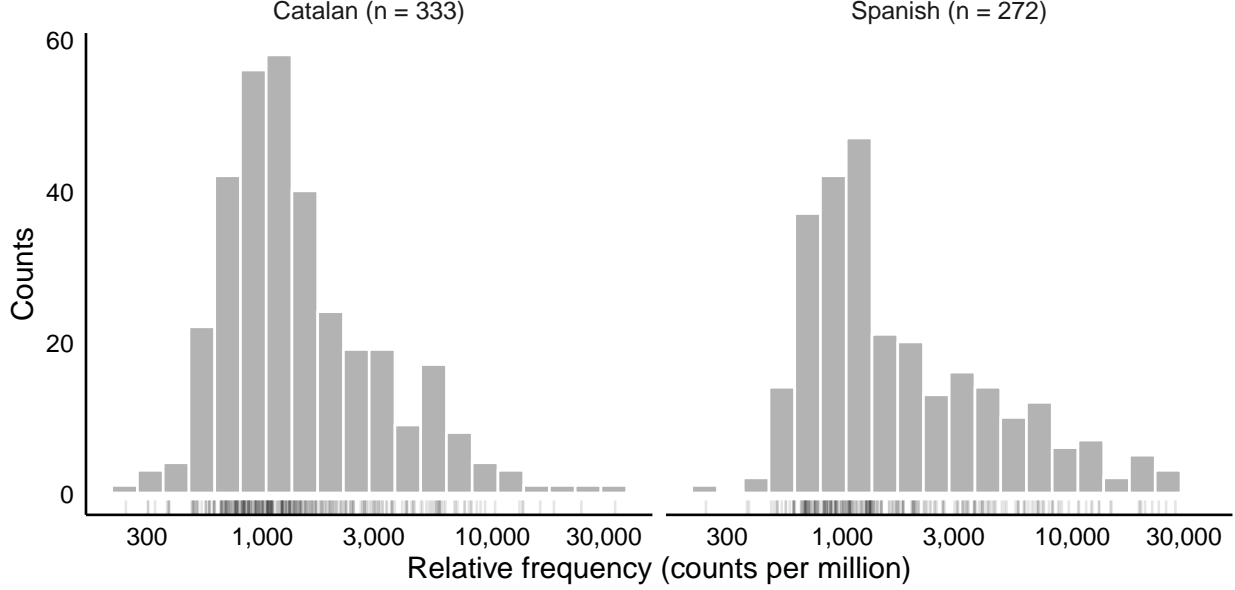


Figure 5: Distribution of apositional syllable frequencies in Spanish and Catalan

To estimate the association between word-level syllabic frequency and cognateness, while controlling for the number of syllables in the word, as words are expected to necessarily increase the syllabic frequency of the word), we fit a multilevel, Bayesian linear regression model with syllabic frequency (the sum of the syllabic frequency of the syllables in a word) as response variable, and the main effect of the number of syllables (*Syllables*) and *Cognateness* (Levenshtein similarity between a word and its translation equivalent, [Levenshtein, 1966](#)) as predictors. We added translation equivalent-level random effects for the intercept and the main effect of *Syllables* (some translation pairs had a different number of syllables in each language). We used a Gaussian distribution to model syllabic frequency scores after standardising this variable and the predictors. We used a weakly informative prior for all parameters involved in the model (see Equation 1 for a formal equation of this model and its prior). We conducted statistical inference by evaluating the proportion of the 95% highest density interval (HDI) of the posterior distribution of each coefficient that fell into the region of practical equivalence (ROPE, see the main manuscript for a more detailed explanation, [Kruschke & Liddell, 2018](#)).

$$\begin{aligned}
 \text{Syllable frequency} &\sim \mathcal{N}(\mu, \sigma) \\
 \mu &= (\beta_0 + u_{0_i}) + (\beta_1 + u_{1_i})\text{Syllables} + \beta_2\text{Cognateness} \\
 \beta_{0-3} &\sim \mathcal{N}(0, 10) \\
 u_{0-1_i} &\sim \mathcal{N}(0, \sigma_{u_i}) \\
 \sigma_y &\sim \text{Exponential}(2) \\
 \begin{pmatrix} u_{0_i} \\ u_{1_i} \end{pmatrix} &\sim \mathcal{N}\left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma_u\right) \\
 \Sigma_u &= \begin{pmatrix} \sigma_{u_0} & \rho_{u_0} \sigma_{u_0} \sigma_{u_1} \\ \rho_{u_1} \sigma_{u_1} \sigma_{u_0} & \sigma_{u_1} \end{pmatrix} \\
 \sigma_{u_{0-1}} &\sim \mathcal{N}_+(1, 0.1) \\
 \rho_u &\sim \text{LKJcorr}(2)
 \end{aligned} \tag{1}$$

We fit this model running 4 sampling chains with 1,000 iterations each. Table 3 shows a summary of the posterior distribution of the fixed effects in the model. As expected, words with more syllables scored higher in syllabic fre-

Table 3: Posterior distribution of regression coefficients.

	β	95% HDI	$p(\text{ROPE})$
Intercept	16.090	[16.022, 16.162]	NA
Syllables (+1 SD, 0.802)	5.644	[5.575, 5.713]	.000
Cognateness (+1 SD, 0.24)	0.009	[-0.056, 0.081]	1.000

quency: all posterior draws for the regression coefficient of the main effect of this predictor fell outside the ROPE defined between -0.5 and +0.5 ($\beta = 5.64$, 95% HDI = [5.58, 5.71]). Keeping the number of syllables constant, the effect of cognateness was negligible: all of the posterior distributions of this predictor fell within the ROPE, providing evidence that the true value of the increment in syllabic frequency for every increase in cognateness is equivalent to zero ($\beta = 0.01$, 95% HDI = [-0.06, 0.08]).

Figure 6 shows the median posterior-predicted syllabic frequencies for words with one to four syllables, for the whole range of cognateness values. Overall, cognate words' syllabic frequency is equivalent to that of non-cognates. This suggests that the cognate facilitation effect in word acquisition reported in the present study is not the result from an association between cognateness and higher syllabic frequencies.

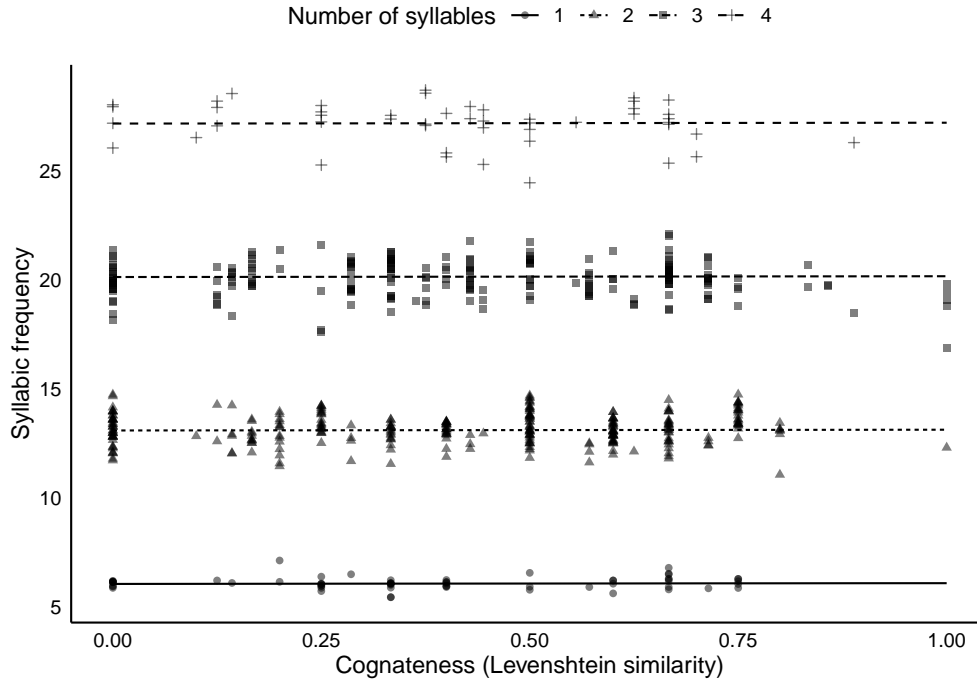


Figure 6: Posterior-predictions of the syllabic frequency model. Thicker lines indicate the median of the posterior predictions, and thinner lines indicate individual posterior predictions.

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