CS6190: Probabilistic Modeling Homework 0

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1. Let X and Y be continuous real-valued random variables. Prove the following:

(a)
$$E[E[X|Y]] = E[X]$$

$$\begin{split} E\left[E\left[X|Y\right]\right] &= \int_{-\infty}^{\infty} E\left[X|Y=y\right] p_Y(y) dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \; p_{X|Y=y}(x) dx \; p_Y(y) dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \; p_{X|Y=y}(x) p_Y(y) dx dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x \; p_{XY}(x,y) dx dy \\ &= \int_{-\infty}^{\infty} x \int_{-\infty}^{\infty} p_{XY}(x,y) dy \; dx \\ &= \int_{-\infty}^{\infty} x \; p_X(x) dx \\ &= E\left[X\right] \end{split}$$

Here $p_X(x)$, $p_Y(y)$ and $p_{X|Y=y}(x)$ are the probability density functions for X, Y, and the conditional probability of X given the value of Y. $p_{XY}(x,y)$ is the joint probability distribution function of X and Y. I needed to look up the proof [1] as I was not aware that the first step shown above could be done and due to this, I was getting an extra y in the final result. The result I got before I looked up the proof was yE[X].

(b)
$$Var(X) = E[Var(X|Y)] + Var(E[X|Y])$$

$$\begin{split} E\left[Var(X|Y)\right] + Var(E\left[X|Y\right]) &= E\left[E\left[X^{2}|Y\right] - E\left[X|Y\right]E\left[X|Y\right]\right] \\ &+ E\left[E\left[X|Y\right]E\left[X|Y\right]\right] - E\left[E\left[X|Y\right]\right]E\left[E\left[X|Y\right]\right] \\ &= E\left[E\left[X^{2}|Y\right]\right] - E\left[E\left[X|Y\right]E\left[X|Y\right]\right] \\ &+ E\left[E\left[X|Y\right]E\left[X|Y\right]\right] - E\left[X\right]E\left[X\right] \\ &= E\left[X^{2}\right] - E\left[X\right]E\left[X\right] \\ &= Var(X) \end{split}$$

Before I worked out the solution above, I was using the definition $Var(X) = E[X^2] - \mu_X^2$ and that did not help. After I realized that $\mu_X = E[X]$, the solution fell in place.

2. Consider random variables X and Y with joint pdf

$$p(x,y) = \begin{cases} x+y, & \text{for } 0 \le x \le 1, \text{and } 0 \le y \le 1\\ 0, & \text{otherwise} \end{cases}$$

(a) What is the marginal pdf p(x)?

$$\begin{split} p(x) &= \int_{-\infty}^{\infty} p(x,y) dy \\ &= \begin{cases} \int_{0}^{1} (x+y) dy, \text{for } 0 \leq x \leq 1 \\ 0, \text{otherwise} \end{cases} \\ &= \begin{cases} \left(xy + \frac{1}{2}y^2 \right) \Big|_{0}^{1}, \text{for } 0 \leq x \leq 1 \\ 0, \text{otherwise} \end{cases} \\ &= \begin{cases} x + \frac{1}{2}, \text{for } 0 \leq x \leq 1 \\ 0, \text{otherwise} \end{cases} \end{split}$$

(b) What is the conditional pdf p(y|x)?

$$\begin{aligned} p(y|x) &= \frac{p(x,y)}{p(x)} \\ &= \begin{cases} \frac{x+y}{x+\frac{1}{2}}, \text{for } 0 \leq x \leq 1, \text{and } 0 \leq y \leq 1 \\ 0, \text{otherwise} \end{cases} \end{aligned}$$

(c) What is the conditional expectation E[Y|X]?

$$E[Y|X] = \int_{-\infty}^{\infty} p_{Y|X}(y)dy$$

$$= \begin{cases} \int_{0}^{1} \frac{x+y}{x+\frac{1}{2}} dy, \text{ for } 0 \leq x \leq 1\\ 0, \text{ otherwise} \end{cases}$$

$$= \begin{cases} \frac{xy+\frac{1}{2}y^{2}}{x+\frac{1}{2}} \Big|_{0}^{1}, \text{ for } 0 \leq x \leq 1\\ 0, \text{ otherwise} \end{cases}$$

$$= \begin{cases} \frac{x+\frac{1}{2}}{x+\frac{1}{2}}, \text{ for } 0 \leq x \leq 1\\ 0, \text{ otherwise} \end{cases}$$

$$= \begin{cases} 1, \text{ for } 0 \leq x \leq 1\\ 0, \text{ otherwise} \end{cases}$$

(d) What is the covariance Cov(X, Y)?

$$E[XY] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xyp(x, y)dxdy$$
$$= \int_{-\infty}^{\infty} x \int_{-\infty}^{\infty} yp(x, y)dydx$$

Cov(X,Y) = E[XY] - E[X]E[Y]

$$\int_{-\infty}^{\infty} yp(x,y)dy = \begin{cases} \int_{0}^{1} y(x+y)dy, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \int_{0}^{1} (xy+y^{2})dy, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{1}{2}xy^{2} + \frac{1}{3}y^{3} \Big|_{0}^{1}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{1}{2}x + \frac{1}{3}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \int_{0}^{1} x \left(\frac{1}{2}x + \frac{1}{3} \right) dx, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \int_{0}^{1} \left(\frac{1}{2}x^{2} + \frac{1}{3}x \right) dx, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \left(\frac{1}{2} \frac{1}{3}x^{3} + \frac{1}{3} \frac{1}{2}x^{2} \right) \Big|_{0}^{1}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \left(\frac{1}{6} + \frac{1}{6} \right), & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \left(\frac{1}{3}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \left(\frac{1}{3}x + \frac{1}{2} \frac{1}{2}x^{2} \right) dx, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

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$$= \begin{cases} \left(\frac{1}{3} + \frac{1}{4}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

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$$= \begin{cases} \frac{1}{3} + \frac$$

$$p(y) = \int_{-\infty}^{\infty} p(x,y)dx$$

$$= \begin{cases} \int_{0}^{1} (x+y)dx, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \left(\frac{1}{2}x^{2} + yx\right)\Big|_{0}^{1}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} y + \frac{1}{2}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$E[Y] = \int_{-\infty}^{\infty} yp(y)dy$$

$$= \begin{cases} \int_{-\infty}^{\infty} y\left(y + \frac{1}{2}\right)dy, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \int_{-\infty}^{\infty} \left(y^{2} + \frac{1}{2}y\right)dy, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \left(\frac{1}{3}y^{3} + \frac{1}{2}\frac{1}{2}y^{2}\right)\Big|_{0}^{1}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{1}{3} + \frac{1}{4}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{7}{12}, & \text{for } 0 \leq x \leq 1, \text{ and } 0 \leq y \leq 1 \\ 0, & \text{otherwise} \end{cases}$$

$$Cov(X,Y) = E[XY] - E[X]E[Y]$$

$$= \begin{cases} \frac{1}{3} - \frac{7}{12} \frac{7}{12}, & \text{for } 0 \le x \le 1, \text{ and } 0 \le y \le 1\\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} \frac{48}{144} - \frac{49}{144}, & \text{for } 0 \le x \le 1, \text{ and } 0 \le y \le 1\\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} -\frac{1}{144}, & \text{for } 0 \le x \le 1, \text{ and } 0 \le y \le 1\\ 0, & \text{otherwise} \end{cases}$$

$$= \begin{cases} 0, & \text{otherwise} \end{cases}$$

- 3. Let $X \sim Exp(\lambda)$, i.e., the exponential distribution with pdf $p(x) = \lambda exp(-\lambda x)$. Let $Y = \sqrt{X}$.
 - (a) What is the density function p(y)?

$$Y = \sqrt{X}$$

$$f(x) = \sqrt{x}$$

$$f^{-1}(y) = y^{2}$$

$$p(y) = \left| \frac{d}{dy} (f^{-1}(y)) \right| p(f^{-1}(y))$$

$$= \left| \frac{d}{dy} y^{2} \right| p(y^{2})$$

$$= \begin{cases} 2y \lambda exp(-\lambda y^{2}), \text{ for } y \geq 0 \\ 0, \text{ otherwise} \end{cases}$$

(b) What is the cdf, $F(y) = P(Y \le y)$? Verify that F(0) = 0 and $F(\infty) = 1$. The cdf can be found by integrating the pdf.

$$F(y) = P(Y \le y) = \int_0^y 2y\lambda exp(-\lambda y^2)dy$$
Substituting $u = -\lambda y^2$, we get $dy = -\frac{1}{2\lambda y}du$

$$\int 2y\lambda exp(-\lambda y^2)dy = \int 2y\lambda exp(u) \times -\frac{1}{2\lambda y}du$$

$$= -\int exp(u)du$$

$$= -exp(u)$$

$$= -exp(-\lambda y^2)$$

$$F(y) = -exp(-\lambda y^2)\Big|_0^y$$

$$= -exp(-\lambda y^2) - (-exp(-\lambda 0^2))$$

$$= -exp(-\lambda y^2) + 1$$

$$= 1 - exp(-\lambda y^2)$$

$$F(0) = 1 - exp(-\lambda 0^2)$$

$$= 1 - 1$$

$$= 0$$

$$F(\infty) = 1 - exp(-\infty)$$

$$= 1 - 0$$

$$= 1$$

(c) What is the quantile function F^{-1} ?

If q_p is the quantile value given by the quantile function $F^{-1}(q_p)$, based on the cdf above:

$$q_{p} = 1 - exp(-\lambda(F^{-1}(q_{p}))^{2})$$

$$exp(-\lambda(F^{-1}(q_{p}))^{2}) = 1 - q_{p}$$

$$-\lambda(F^{-1}(q_{p}))^{2} = \ln(1 - q_{p})$$

$$(F^{-1}(q_{p}))^{2} = \frac{-1}{\lambda}\ln(1 - q_{p})$$

$$F^{-1}(q_{p}) = \sqrt{\frac{-1}{\lambda}\ln(1 - q_{p})}$$

(d) Compute the mean, E[Y], and variance, Var(Y). Hint: Use integration by parts.

$$\begin{split} E\left[Y\right] &= \int_{-\infty}^{\infty} y p(y) dy \\ &= \int_{0}^{\infty} 2y^{2} \lambda exp(-\lambda y^{2}) dy = \int_{0}^{\infty} x^{\frac{1}{2}} \lambda exp(-\lambda x) \\ Var(Y) &= E\left[Y^{2}\right] - E\left[Y\right] E\left[Y\right] \end{split}$$

4. Given a realization y_1, y_2, \dots, y_n from your random variable Y in the previous problem, what is the maximum likelihood estimate for λ ?

R Coding Part

4. Recall that if F is a cdf and $U \sim \text{Unif}(0,1)$, then $Y = F^{-1}(U)$ will be a random variable with cdf F. Write a function that uses this method to generate n random numbers from the distribution of Y in Problem 3 above (n and λ should be parameters to the function).

Hint: Look at the runif function in R for simulating uniform random varibles.

- (a) Generate 10,000 realizations of the random variable Y with $\lambda = 2$.
- (b) Plot a histogram of these numbers, using the hist function with option freq = FALSE.
- (c) Use the lines command to plot the pdf of Y on top.
- (d) Compute the sample mean and variance of your 10,000 realizations. Do they roughly match the values for E[Y] and Var(Y) you calculated above?
- 5. Now generate 20 realizations from Y with $\lambda = 2$. Plot the log likelihood function for this data. Plot a vertical line at the maximum likelihood estimate for lambda (using the equation you got in Problem 3).

References

[1] Conditional Expectation, www.statlect.com/fundamentals-of-probability/conditional-expectation.