# A biometric analysis of infant mortality and temperature, northern Sweden 1895-1950

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## Abstract

The effect of extreme temperatures on infant mortality in the Umeå and Skellefteå regions 1895-1950 is studied in a biometric analysis setting. More precisely, the effect of climate and weather, measured by temperature, average and extremes, on infant mortality is investigated. It turns out that climate (average) is more important than weather (extremes), low average temperatures are more important than temporary dips in temperature, but effects are different in neonatal and postneonatal settings.

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## 1 Introduction

The impact of ambient temperature variations on infant mortality is studied for two northern Sweden areas, the Umeå coastal region and Skellefteå coastal and inland regions, during the first half of the twentieth century. Two recent papers (Junkka et al., 2021; Karlsson et al., 2021) studied neonatal mortality and temperature variations in this geographical area during the years 1880–1950. Climate and mortality in general is a research area that has generated great interest over the last years, see Bengtsson and Broström (2010) for an example of our own efforts.

The effect of seasonal variation and the occurrence of extreme monthly temperatures is studied and interacted with social class and time period. Studies are performed separately for neonatal and postneonatal mortality, and for winter and summer seasons.

Figure 1 shows the study area within Sweden, with the weather stations marked. The map is taken (with permission) from the paper by Junkka et al. (2021).

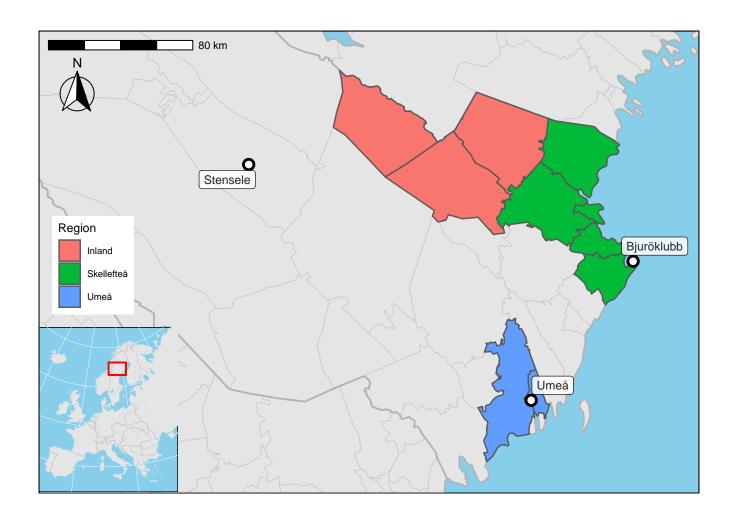


Figure 1: Umeå, Skellefteå (orange) and its inland (green).

## 2 Data

We have two sources of data which we combine into one data set suitable for our purpose. The first is demographic data obtained from the *Centre for Demographic and Ageing Research* (CEDAR, https://cedar.umu.se), the second is daily temperature measurements obtained

from the Swedish Meteorological and Hydrological Institute (SMHI, https://www.smhi.se).

## 2.1 Infant mortality

Individual data with all births between 1 January 1895 and 31 December 1950 in two coastal and one inland areas of north Sweden, Skellefteå (51560 births) and Umeå (31213 births). They were followed until death or age one year, whichever came first. The following *static* characteristics were observed on each child:

birthdate Date of birth.

sex Girl or boy.

exit Number of days under observation.

event Logical, TRUE if a death is observed.

socBranch Working branch of father (if any).

**socStatus** Social status *within* the given working branch.

illeg Mother unmarried?

parity Order among siblings.

Some crude statistics about infant, neonatal, and postneonatal mortality are shown in Figures.

Figure 2 shows the average weekly crude infant mortality, and a clear seasonal pattern is visible.

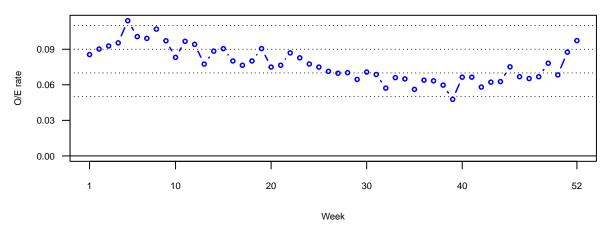


Figure 2: Crude infant mortality by week of year, Umeå/Skellefteå 1895–1950.

The average weekly neonatal mortality is shown in Figure 3.

The seasonal pattern is similar to the one we found above for infant mortality.

The average weekly postneonatal mortality is shown in Figure 4.

The seasonal pattern is once again similar to the one we found for infant mortality. Next, the decline over the years in Figures 5 and 6.

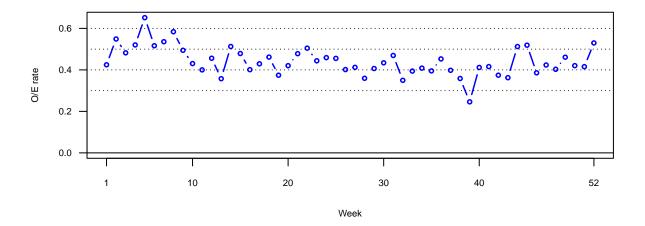


Figure 3: Crude neonatal mortality by week of year, Umeå/Skellefteå 1895–1950.

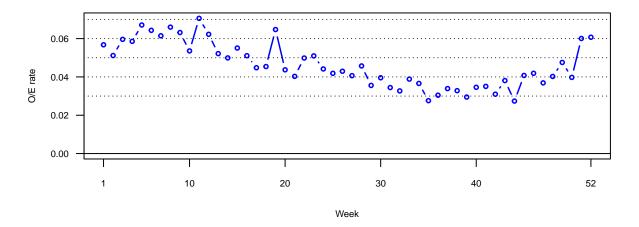


Figure 4: Crude postneonatal mortality by week of year, Umeå/Skellefteå 1895–1950.

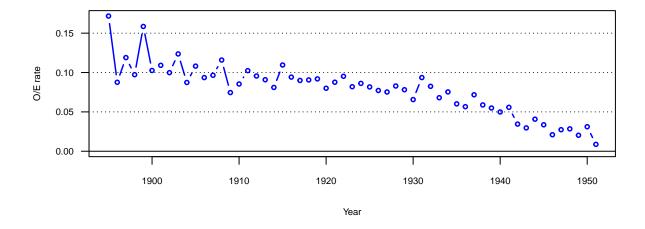


Figure 5: Crude IMR by year, Umeå-Skellefteå 1895–1950.

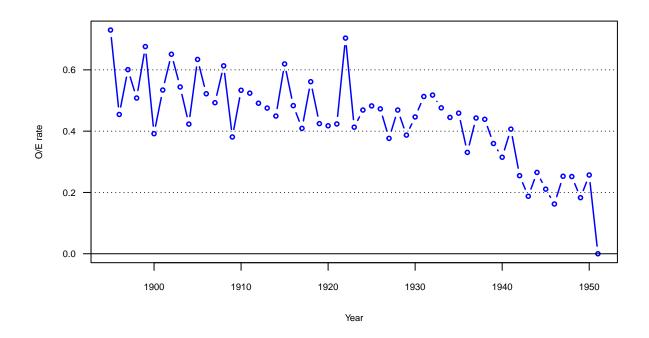


Figure 6: Crude NMR by year, Umeå-Skellefteå 1895–1950.

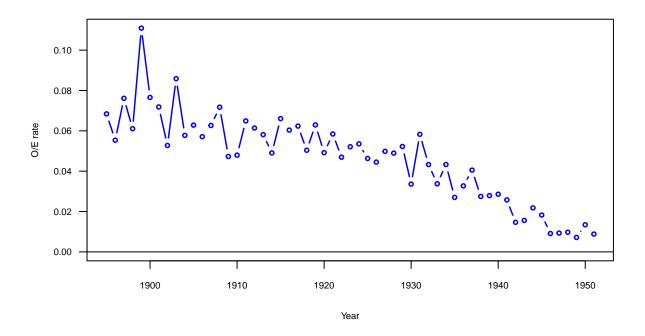


Figure 7: Crude postneonatal mortality by year, Umeå-Skellefteå 1895–1950.

Note that the estimates of neonatal and postneonatal mortality are non-standard here: They are calculated as *occurrence-exposure rates*, that is, the number of deaths divided by *exposure time*. They are thus crude *hazard rates*, not probabilities, which is more common in definitions of mortality.

## 2.2 Temperature

Temperature data are collected from three weather stations, *Umeå*, *Bjuröklubb* (used with population data from Skellefteå coastal area), and *Stensele* (Inland). All stations deliver daily temperature data covering our time period, usually three measures per day, morning, noon, and evening. In Table 1, the Umeå data from the week 1–7 January, 1923 is shown.

#### [Table 1 about here.]

There are three measurements per day, or 21 per week. In the forthcoming analyses, the weekly data are summarized in a few measurements, see Table 2. Our rule for week numbering is that week No. 1 always start on January 1. Week No. 52 always ends at December 31, and so will be eight days long, except for leap years, when it will be nine days long.

## [Table 2 about here.]

Weekly averages (mintemp, maxtemp, meantemp) are calculated by week and year, and deviations from the averages (emintemp, emaxtemp, emeantemp) of the weekly averages are used as time-varying *communal covariates*. As an example, see Figure 8, where the variation around the average minimum temperature (emintemp) week 1 is shown.

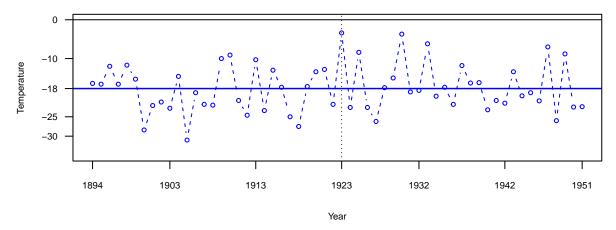


Figure 8: Minimum temperature the first week of each year.

Curiously, our randomly selected year 1923 turns out contain the warmest first week of all years, see Figure 8.

Figure 9 shows the average monthly distribution over all years. The subregional patterns and levels are very similar.

Time trends of yearly average temperatures, see Figure 10.

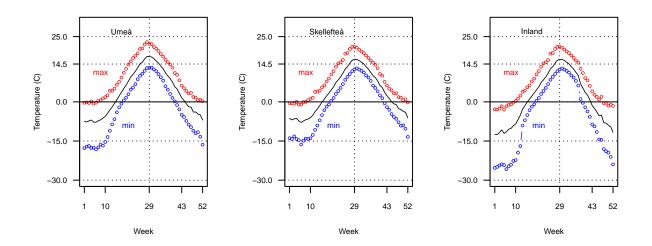


Figure 9: Weekly max, mean, and min temperature averages, 1895–1950.

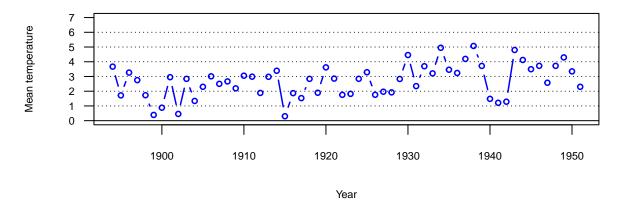


Figure 10: Yearly average temperatures, Umeå and Skellefteå.

## 2.3 Temperature as communal covariates

The two data sets, mortality and weather, are combined into one by treating temperature data as a communal covariate and incorporate it as such in the mortality data set. The function make.communal in the  $\mathbf{R}$  (R Core Team, 2021) package eha (Broström, 2021a,b) is used for that purpose. Resulting data drame is partly shown in Table 3.

[Table 3 about here.]

## 3 Statistical modelling

It turns out that extremely low temperature (lowTemp) is bad during all seasons except summer, and extremely high temperature (highTemp) is bad during summer, but good otherwise. So we group season into two categories, *summer* and *not summer*. In each case separate analyses for neonatal and postneonatal mortality are performed.

The *summer* half-year consists of the weeks 14–39, about 1 April to 30 September, and the *winter* half-year is the rest, weeks 1-13 and 40–52, 1 January to 31 March and 1 October to 31 December. This is the division made in Karlsson et al. (2021), and we keep it for comparability reasons.

## 4 Results

The results regarding neonatal mortality is much in accordance with the results found by Junkka et al. (2021). However, they used temperature in a "hockey-stick" regression with a breakpoint at 14.5 degrees Celsius and a negative slope (decreasing risk) to the left and a positive slope (increasing risk) to the right. Instead, we are using the average weekly temperature for the 52 weeks of a year, for each week averaging over all the years in the study, as our "reference points" ("climate"), adding deviances up and down ("weather") as "short-term temperature stress". This is similar to the way prices and mortality were related in for instance Bengtsson and Broström (2011), that is, a time series split into long time trend and short term variation.

Comments on other candidate covariates:

Birth month is left out in the analyses despite that fact that it is an important factor in neonatal mortality. However, we include time of year in terms of winter and summer as a time-varying covariate, and in the neonatal case it will coincide to a great extent with birth period. For the postneonal case the situation is different, but it turns out that for those infants who have survived the first month of life, birth month does not matter much. We separate the investigation into two parts, neonatal and postneonatal mortality. But first, a joint analysis.

Socioeconomic status is divided into two factor covariates: socBranch and socStatus. The latter should be seen relative to actual socBranch. We have information on whether the infant was illegitimate (mother unmarried), but we incorporate those cases in the category none of socBranch and unknown as category of socStatus.

Covariates may affect neonatal and postneonatal mortality differently, and one way to investigate that is to stratify the infant mortality data into two age intervals, one from birth to one month of age and the other from one month to one year of age. Then the interaction between the stratum variable and other variables of interest is investigated, with this result:

#### Single term deletions

```
Model:
oe(event, exposure) ~ strata(ageIvl) * (extemp + extemp.1 + emeantemp +
    socst + period + urban)
                                          Pr(>Chi)
                               AIC
                                      LRT
<none>
                            41696
strata(ageIvl):extemp
                          1 41694 0.026
                                             0.8730
strata(ageIvl):extemp.1
                           1 41696 2.279
                                             0.1312
strata(ageIvl):emeantemp
                          1 41713 19.018 1.295e-05 ***
                          5 41721 34.937 1.549e-06 ***
strata(ageIvl):socst
strata(ageIvl):period
                          2 41764 72.221 < 2.2e-16 ***
strata(ageIvl):urban
                           1 41695 1.269
                                             0.2599
```

We can see that the variables *emeantemp*, *socStatus* and *period* all have different effects on mortality in the neonatal case compared to the postneonatal case. In terms of *infant mortality*, we can say that these caovariates have non-proportional effects.

0 '\*\*\* 0.001 '\*\* 0.01 '\* 0.05 '.' 0.1 ' 1

## 4.1 Neonatal mortality

The analyses are split into two parts by season, winter is one, and summer the other. Summer covers the months April to September, and winter the rest of the months.

#### 4.1.1 Winter

Signif. codes:

This period refers to the months October to March. A Cox regression involves as interesting variables high Temp, an indicator of temperature at least four degrees above the expected for at least two weeks in a row, emeantemp the expected temperature the actual week, and extemp the excess temperature the actual week.

[Table 4 about here.]

**4.1.1.1 Temperature and sex** We check the interaction between temperature and sex and period.

Single term deletions

#### Model: oe(event, exposure) ~ sex \* (extemp + extemp.1 + emeantemp) Df AIC LRT Pr(>Chi) <none> 4507.1 1 4507.6 2.47699 sex:extemp 0.1155 sex:extemp.1 1 4505.6 0.50456 0.4775 1 4505.3 0.11737 sex:emeantemp 0.7319

Obviously no temperature and sex interactions

#### 4.1.2 Summer

This period refers to the months *April to September*. A Cox regression involves as interesting variables *highTemp*, an indicator of temperature at least four degrees above the expected for at least two weeks in a row, *emeantemp* the expected temperature the actual week, and *extemp* the *excess temperature* the actual week.

[Table 5 about here.]

## 4.2 Postneonatal mortality

## **4.2.0.1** By birth month See Figure 11.

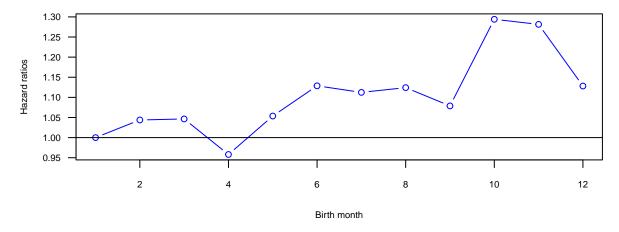


Figure 11: Crude postneonatal relative (to January = 1) mortality by birth month.

## 4.2.1 Winter

Interaction tests with temperature and period, postneonatal mortality:

## Single term deletions

#### Model: oe(event, exposure) ~ period \* (extemp + extemp.1 + emeantemp) LRT Pr(>Chi) Df AIC <none> 17721 2 17722 4.2212 period:extemp 0.1212 period:extemp.1 2 17719 1.4788 0.4774 period:emeantemp 2 17720 2.3455 0.3095

This shows that the temperature effects do not vary much with time period, but we continue with separate analyses for the time periods (1895,1914], (1914,1935], (1935,1951]. Just to make sure, will be joined when appropriate.

**4.2.1.1** Sex, urban, and temperature Interaction tests with temperature and sex, postneonatal mortality:

Single term deletions

#### Model:

```
oe(event, exposure) ~ (sex + urban) * (extemp + extemp.1 + emeantemp)
                Df
                             LRT Pr(>Chi)
                      AIC
<none>
                    18095
sex:extemp
                 1 18093 0.0714
                                   0.7894
sex:extemp.1
                 1 18093 0.0116
                                   0.9143
                 1 18093 0.0518
                                   0.8200
sex:emeantemp
urban:extemp
                 1 18094 0.6809
                                   0.4093
urban:extemp.1
                 1 18094 0.6013
                                   0.4381
urban: emeantemp
                 1 18094 0.2018
                                   0.6533
```

Obviously no interactions between temperature and sex, and urban.

#### 4.2.1.2 The time period 1895–1913

[Table 6 about here.]

Table 6 shows that the effects of extreme temperatures have the expected directions, extra high temperatures is *positive* (remember that this concerns winter conditions), but not statistically significant. *Climate* (emeantemp), on the other hand shows to be very important, where increasing temperature by one degree lower mortality by 0.7 per cent. The normal short term fluctuations (extemp) are not so important, but in the expected direction.

The importance of social class (socBranch) in explaining the effects of climate and weather is negligible, but there is variation between the regions regarding *climate*:

[Table 7 about here.]

## [Table 8 about here.]

We can conclude from Tables 7 and 8 that the main difference between Umeå and Skellefteå is that *climate* has more severe effect in Skellefteå (with inland), while *weather* is less important.

The effects of temperature in different social classes are best shown graphically, see Figure 13.

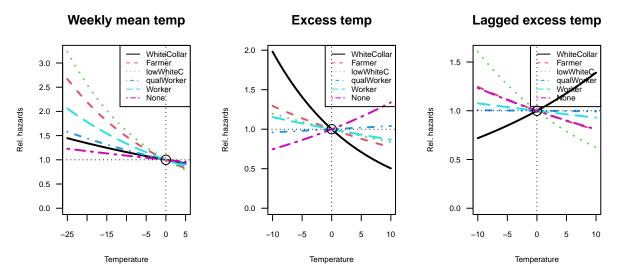


Figure 12: Interactions between social status and temperature, winter 1895–1913.

## 4.2.1.3 The time period 1914–1934

[Table 9 about here.]

Table 9 shows that the weather effect is lagged one week (*extemp.1*), and that climate continues to be very important. Let's look at regional differences.

[Table 10 about here.]

[Table 11 about here.]

The effects of temperature in different social classes are best shown graphically, see Figure 13.

## 4.2.1.4 The time period 1935–1950

[Table 12 about here.]

Table 12 shows that

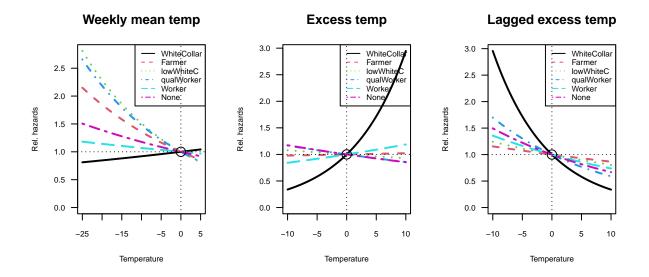


Figure 13: Interactions between social status and temperature, winter 1914–1934.

[Table 13 about here.]

[Table 14 about here.]

The effects of temperature in different social classes are best shown graphically, see Figure 13.

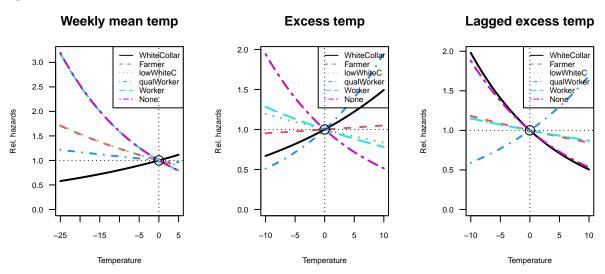


Figure 14: Interactions between social status and temperature, winter 1935–1950.

## **4.2.2** Summer

[Table 15 about here.]

The result in Table 15 shows that *climate* (emeantemp) is more important than *weather* (exTemp). Moreover, no signs of interaction between weather or climate and the rest of covariates (not shown).

## 5 Conclusion

Remains to be written, especially the discussion about temperature and mortality.

Used processing time (seconds):

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user system elapsed
1897.628 1603.543 414.167
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Done: 2022-02-15 18:43:01.

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Table 1: Raw temperature data from first week of 1923, Umeå.

Date	Time	Temperature	Quality
1923-01-01	07:00:00	0.4	G
1923-01-01	13:00:00	0.6	G
1923-01-01	20:00:00	0.0	G
1923-01-02	07:00:00	-1.4	G
1923-01-02	13:00:00	-1.4	G
1923-01-02	20:00:00	-1.2	G
1923-01-03	07:00:00	0.4	G
1923-01-03	13:00:00	0.8	G
1923-01-03	20:00:00	1.2	G
1923-01-04	07:00:00	1.4	G
1923-01-04	13:00:00	1.2	G
1923-01-04	20:00:00	1.0	G
1923-01-05	07:00:00	-1.4	G
1923-01-05	13:00:00	-3.2	G
1923-01-05	20:00:00	-3.4	G
1923-01-06	07:00:00	1.0	G
1923-01-06	13:00:00	0.4	G
1923-01-06	20:00:00	0.4	G
1923-01-07	07:00:00	0.6	G
1923-01-07	13:00:00	0.4	G
1923-01-07	20:00:00	0.4	G

Table 2: Weekly summarized temperature data: Umeå 1923, first week.

week	year	mintemp	maxtemp	meantemp	${\bf emintemp}$	$\mathbf{e}\mathbf{m}\mathbf{a}\mathbf{x}\mathbf{t}\mathbf{e}\mathbf{m}\mathbf{p}$	emeantemp
1	1923	-3.4	1.4	-0.1	-17.73	-0.36	-7.54

Table 3: Data with communal covariates.

enter	$\mathbf{exit}$	event	extemp	extemp.1	emeantemp	week	year
0.0000000	0.0180327	0	0	1	12.584483	35	1900
0.0180327	0.0372634	0	-4	0	11.286207	36	1900
0.0372634	0.0564942	0	-1	-4	10.086207	37	1900
0.0564942	0.0757250	0	1	-1	6.910345	38	1900
0.0757250	0.0949557	0	-2	1	5.315517	39	1900
0.0949557	0.1141865	0	0	-2	3.610345	40	1900

Table 4: Neonatal mortality, October to March.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	-0.024	-0.005	0.995	0.006	0.448
extemp.1	-0.026	-0.012	0.988	0.006	0.053
emeantemp	-4.140	-0.020	0.980	0.006	0.001
socst					0.348
high  White  C	0.029	0	1	(refe	rence)
farmer	0.471	-0.095	0.909	0.158	
lowWhiteC	0.076	-0.191	0.826	0.177	
qual Worker	0.130	0.050	1.051	0.162	
worker	0.270	-0.063	0.939	0.159	
none	0.024	-0.039	0.962	0.210	
period					0.000
(1895, 1914]	0.354	0	1	(refe	rence)
(1914, 1935]	0.372	-0.087	0.917	0.053	
(1935, 1951]	0.275	-0.624	0.536	0.071	
urban					0.530
FALSE	0.863	0	1	(refe	rence)
TRUE	0.137	-0.055	0.947	0.087	
subreg					0.273
ume	0.296	0	1	(refe	rence)
ske	0.487	0.057	1.059	0.059	
in land	0.217	-0.041	0.960	0.074	
sex					0.000
boy	0.516	0	1	(refe	rence)
girl	0.484	-0.228	0.796	0.048	
parity					0.000
1	0.260	0	1	(refe	rence)
2-4	0.455	-0.333	0.717	0.059	
5+	0.285	-0.062	0.940	0.063	
Events	1806	TTR	3899		
Max. logLik.	-2169				

Table 5: Neonatal mortality, April to September.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	0.022	0.001	1.001	0.012	0.911
extemp.1	0.022	0.001	1.001 $1.007$	0.012	0.530
emeantemp	9.817	-0.006	0.994	0.005	0.162
socst	0.011	0.000	0.001	0.000	0.102
highWhiteC	0.029	0	1	(refe	rence)
farmer	0.460	0.315	1.371	0.192	roncej
lowWhiteC	0.078	0.201	1.223	0.207	
qualWorker	0.135	0.316	1.372	0.196	
worker	0.272	0.246	1.280	0.192	
none	0.026	0.572	1.772	0.227	
period	0.0_0	0.01=		·	0.000
(1895, 1914]	0.333	0	1	(refe	rence)
(1914, 1935]	0.370	-0.127	0.881	0.056	,
(1935, 1951)	0.297	-0.537	0.584	0.072	
urban					0.038
FALSE	0.860	0	1	(refe	rence)
TRUE	0.140	-0.190	0.827	0.092	,
subreg					0.578
ume	0.295	0	1	(refe	rence)
ske	0.485	0.051	1.053	0.061	,
in land	0.220	-0.001	0.999	0.074	
sex					0.000
boy	0.511	0	1	(refe	rence)
girl	0.489	-0.217	0.805	0.049	
parity					0.000
1	0.269	0	1	(refe	rence)
2-4	0.451	-0.267	0.765	0.060	
5+	0.280	-0.105	0.900	0.066	
Events	1689	TTR	4026		
Max. logLik.	-2024				

Table 6: Postneonatal mortality, October to March 1895-1914.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
socst					0.004
high  White  C	0.016	0	1	(refe	rence)
farmer	0.557	0.750	2.117	0.359	
lowWhiteC	0.047	0.641	1.898	0.384	
qual Worker	0.094	0.976	2.653	0.365	
worker	0.264	0.920	2.510	0.360	
none	0.022	0.862	2.367	0.410	
extemp	-0.273	-0.018	0.982	0.007	0.016
extemp.1	-0.237	-0.016	0.984	0.008	0.036
emeantemp	-4.154	-0.050	0.951	0.008	0.000
urban					0.514
FALSE	0.938	0	1	(refe	rence)
TRUE	0.062	-0.086	0.918	0.132	
subreg					0.000
ume	0.305	0	1	(refe	rence)
ske	0.476	-0.390	0.677	0.070	
in land	0.218	-0.529	0.589	0.091	
sex					0.000
boy	0.514	0	1	(refe	rence)
girl	0.486	-0.269	0.764	0.061	
parity					0.000
1	0.187	0	1	(refe	rence)
2-4	0.423	-0.027	0.974	0.087	
5+	0.390	0.229	1.258	0.086	
Events	1111	TTR	15642		
Max. logLik.	-3882				

Table 7: Postneonatal mortality, Umeå, October to March 1895-1913.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	-0.223	-0.014	0.986	0.013	0.265
extemp.1	-0.183	-0.016	0.984	0.013	0.210
emeantemp	-3.621	-0.013	0.987	0.012	0.292
socst					0.003
highWhiteC	0.028	0	1	(refe	rence)
farmer	0.425	1.296	3.656	0.587	
lowWhiteC	0.071	0.930	2.535	0.621	
qual Worker	0.148	1.435	4.200	0.590	
worker	0.304	1.515	4.548	0.586	
none	0.024	1.550	4.712	0.647	
urban					0.581
FALSE	0.827	0	1	(reference)	
TRUE	0.173	-0.082	0.922	0.149	
sex					0.009
boy	0.513	0	1	(refe	rence)
girl	0.487	-0.252	0.777	0.097	
parity					0.034
1	0.204	0	1	(refe	rence)
2-4	0.452	-0.119	0.888	0.131	
5+	0.344	0.163	1.177	0.132	
Events	436	TTR	4777		
Max. logLik.	-1428				

Table 8: Postneonatal mortality, Skellefteå with inland, October to March 1895-1913.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	-0.294	-0.019	0.981	0.009	0.037
extemp.1	-0.260	-0.015	0.985	0.009	0.100
emeantemp	-4.388	-0.063	0.939	0.009	0.000
socst					0.267
highWhiteC	0.011	0	1	(refe	rence)
farmer	0.616	0.225	1.252	0.453	
lowWhiteC	0.036	0.322	1.381	0.490	
qual Worker	0.071	0.523	1.687	0.466	
worker	0.246	0.374	1.453	0.456	
none	0.021	0.208	1.231	0.536	
urban					0.397
FALSE	0.987	0	1	(reference)	
TRUE	0.013	0.277	1.320	0.315	
sex					0.000
boy	0.514	0	1	(refe	rence)
girl	0.486	-0.278	0.757	0.078	
parity					0.006
1	0.179	0	1	(refe	rence)
2-4	0.411	0.040	1.041	0.117	
5+	0.410	0.277	1.319	0.114	
Events	675	TTR	10865		
Max. logLik.	-2443				

Table 9: Postneonatal mortality, October to March 1914-1934. Adjusted for sex, parity, and subregion.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
socst					0.000
highWhiteC	0.023	0	1	(refe	rence)
farmer	0.497	0.771	2.163	0.327	
lowWhiteC	0.075	0.768	2.156	0.339	
qual Worker	0.114	0.763	2.146	0.332	
worker	0.266	1.039	2.825	0.327	
none	0.026	0.876	2.400	0.379	
extemp	-0.020	0.004	1.004	0.008	0.579
extemp.1	-0.044	-0.025	0.975	0.008	0.001
emeantemp	-4.170	-0.056	0.946	0.009	0.000
urban					0.459
FALSE	0.863	0	1	(refe	rence)
TRUE	0.137	-0.085	0.919	0.115	
subreg					0.000
ume	0.265	0	1	(refe	rence)
ske	0.508	-0.348	0.706	0.080	
in land	0.227	-0.436	0.647	0.100	
sex					0.000
boy	0.506	0	1	(refe	rence)
girl	0.494	-0.232	0.793	0.065	
parity					0.000
1	0.254	0	1	(refe	rence)
2-4	0.445	0.220	1.246	0.087	
5+	0.301	0.425	1.530	0.091	
Events	967	TTR	17087		
Max. logLik.	-3587				

Table 10: Postneonatal mortality, Umeå, October to March 1914-1934.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	-0.118	0.019	1.019	0.015	0.215
extemp.1	-0.138	-0.022	0.978	0.015	0.131
emeantemp	-3.619	-0.061	0.941	0.016	0.000
socst					0.012
highWhiteC	0.047	0	1	(refe	rence)
farmer	0.388	0.849	2.338	0.435	
lowWhiteC	0.138	0.425	1.529	0.452	
qual Worker	0.197	0.806	2.240	0.431	
worker	0.188	1.105	3.019	0.430	
none	0.041	0.677	1.969	0.521	
urban					0.729
FALSE	0.637	0	1	(reference)	
TRUE	0.363	-0.053	0.949	0.152	
sex					0.001
boy	0.513	0	1	(refe	rence)
girl	0.487	-0.391	0.677	0.118	
parity					0.007
1	0.304	0	1	(refe	rence)
2-4	0.458	0.081	1.084	0.145	
5+	0.238	0.447	1.564	0.156	
Events	303	TTR	4522		
Max. logLik.	-1067				

Table 11: Postneonatal mortality, Skellefteå with inland, October to March 1914-1934.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	0.015	-0.001	0.999	0.010	0.910
extemp.1	-0.011	-0.026	0.974	0.009	0.005
emeantemp	-4.369	-0.050	0.951	0.009	0.000
socst					0.004
highWhiteC	0.015	0	1	(reference)	
farmer	0.536	0.762	2.142	0.508	
lowWhiteC	0.052	1.077	2.935	0.524	
qual Worker	0.084	0.719	2.052	0.523	
worker	0.294	1.041	2.831	0.508	
none	0.020	1.052	2.865	0.570	
urban					0.895
FALSE	0.945	0	1	(reference)	
TRUE	0.055	0.025	1.025	0.187	
sex					0.039
boy	0.503	0	1	(reference)	
girl	0.497	-0.160	0.852	0.078	
parity					0.000
1	0.236	0	1	(refe	rence)
2-4	0.441	0.296	1.344	0.109	
5+	0.324	0.433	1.542	0.114	
Events	664	TTR	12566		
Max. logLik.	-2508				

Table 12: Postneonatal mortality, October to March 1935-1950. Adjusted for sex, parity, and subregion.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
	0.394	-0.007	0.993	0.013	0.590
extemp.1					
emeantemp	-4.129	-0.057	0.944	0.016	0.000
socst					0.081
high  White  C	0.053	0	1	(reference)	
farmer	0.313	1.158	3.183	0.520	
lowWhiteC	0.118	0.857	2.356	0.544	
qual Worker	0.204	1.037	2.820	0.523	
worker	0.283	1.194	3.301	0.517	
none	0.028	1.399	4.050	0.603	
urban					0.060
FALSE	0.764	0	1	(reference)	
TRUE	0.236	-0.366	0.693	0.199	
subreg					0.853
ume	0.318	0	1	(refe	rence)
ske	0.473	-0.090	0.914	0.159	
in land	0.210	-0.052	0.949	0.189	
sex					0.136
boy	0.516	0	1	(refe	rence)
girl	0.484	-0.183	0.833	0.123	,
parity					0.014
1	0.375	0	1	(refe	rence)
2-4	0.502	0.003	1.003	0.140	
5+	0.123	0.481	1.617	0.179	
Events	268	TTR	13484		
Max. logLik.	-1266				

Table 13: Postneonatal mortality, Umeå, October to March 1935-1950.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	0.376	0.014	1.014	0.031	0.652
extemp.1	0.364	0.002	1.002	0.031	0.939
emeantemp	-3.615	-0.036	0.964	0.031	0.228
socst					0.490
highWhiteC	0.086	0	1	(refe	rence)
farmer	0.226	0.557	1.746	0.651	
lowWhiteC	0.191	0.402	1.495	0.654	
qual Worker	0.262	0.319	1.376	0.640	
worker	0.187	0.892	2.441	0.629	
none	0.047	0.129	1.137	0.916	
urban					0.058
FALSE	0.510	0	1	(reference)	
TRUE	0.490	-0.527	0.590	0.281	
sex					0.728
boy	0.519	0	1	(refe	rence)
girl	0.481	0.082	1.085	0.234	
parity					0.315
1	0.417	0	1	(refe	rence)
2-4	0.503	-0.009	0.991	0.257	
5+	0.080	0.534	1.705	0.365	
Events	73	TTR	4281		
Max. logLik.	-354				

Table 14: Postneonatal mortality, Skellefteå with inland, October to March 1895-1913.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	0.407	-0.003	0.997	0.017	0.876
extemp.1	0.408	-0.011	0.989	0.017	0.515
emeantemp	-4.368	-0.065	0.937	0.017	0.000
socst					0.042
high  White  C	0.038	0	1	(reference)	
farmer	0.354	1.958	7.088	1.010	
lowWhiteC	0.084	1.580	4.854	1.045	
qual Worker	0.176	1.913	6.770	1.016	
worker	0.328	1.911	6.761	1.010	
none	0.019	2.591	13.345	1.072	
urban					0.480
FALSE	0.883	0	1	(reference)	
TRUE	0.117	-0.194	0.824	0.281	
sex					0.046
boy	0.515	0	1	(reference)	
girl	0.485	-0.288	0.750	0.145	
parity					0.045
1	0.355	0	1	(reference)	
2-4	0.501	0.017	1.017	0.167	
5+	0.143	0.472	1.604	0.206	
Events	195	TTR	9203		
Max. logLik.	-899				

Table 15: Postneonatal mortality, April to September. Adjusted for sex, parity, time period, and subregion.

Covariate	Mean	Coef	H.R.	S.E.	L-R p
extemp	0.006	0.007	1.007	0.011	0.507
extemp.1	0.003	-0.012	0.988	0.011	0.287
emeantemp	9.963	-0.020	0.980	0.004	0.000
socst					0.000
high  White  C	0.030	0	1	(reference)	
farmer	0.464	0.755	2.128	0.231	
lowWhiteC	0.078	0.576	1.779	0.245	
qualWorker	0.133	0.653	1.921	0.236	
worker	0.270	0.996	2.707	0.230	
none	0.025	0.870	2.386	0.268	
urban					0.016
FALSE	0.860	0	1	(refe	rence)
TRUE	0.140	-0.211	0.810	0.089	
period					0.000
(1895, 1914]	0.340	0	1	(refe	rence)
(1914, 1935]	0.370	-0.206	0.814	0.050	
(1935, 1951]	0.291	-1.001	0.367	0.077	
subreg					0.000
ume	0.294	0	1	(refe	rence)
ske	0.487	-0.414	0.661	0.054	
in land	0.219	-0.426	0.653	0.067	
parity					0.000
1	0.266	0	1	(refe	rence)
2-4	0.455	0.031	1.032	0.062	
5+	0.279	0.302	1.352	0.065	
sex					0.000
boy	0.512	0	1	(refe	rence)
girl	0.488	-0.168	0.845	0.046	
Events	1914	TTR	45958		
Max. logLik.	-7639				