

Gyujin Park

CSE 6740

HOMEWORK #3

1. QUESTIONS

Question 1. (a)

$$\begin{aligned}\hat{\theta} &= (X^T X)^{-1} X^T Y \\ E(\hat{\theta}) &= E((X^T X)^{-1} X^T Y) \\ &= (X^T X)^{-1} X^T E(Y) \\ &= (X^T X)^{-1} X^T X \theta \\ &= \theta\end{aligned}$$

$\therefore \hat{\theta}$ is unbiased estimator for θ .

(b)

$$\begin{aligned}Cov(\hat{\theta}) &= Cov((X^T X)^{-1} X^T Y) \\ &= (X^T X)^{-1} X^T Cov(Y) ((X^T X)^{-1} X^T)^T \\ &= \sigma^2 (X^T X)^{-1} X^T X (X^T X)^{-1} \\ &= \sigma^2 (X^T X)^{-1}\end{aligned}$$

(c)

I agree. $\hat{\theta}$ follows Gaussian distribution because $\hat{\theta}$ is a linear transformation of Y .

Question 2. (a)

$$\begin{aligned}\hat{\theta} &= (X^T X + \lambda I)^{-1} X^T Y \\ E(\hat{\theta}) &= (X^T X + \lambda I)^{-1} X^T X \theta\end{aligned}$$

$$\begin{aligned}Cov(\hat{\theta}) &= Cov((X^T X + \lambda I)^{-1} X^T Y) \\ &= (X^T X + \lambda I)^{-1} X^T Cov(y) ((X^T X + \lambda I)^{-1} X^T)^T \\ &= \sigma^2 (X^T X + \lambda I)^{-1} X^T X (X^T X + \lambda I)^{-1}\end{aligned}$$

Question 3.2. (a)

Let's assume the class distribution is Bernoulli with parameter θ .

$$p(x|y = 1) = (2\pi)^{-\frac{d}{2}} |\Sigma_1|^{-\frac{1}{2}} e^{-\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1} (x-\mu_1)}$$

$$p(x|y = -1) = (2\pi)^{-\frac{d}{2}} |\Sigma_2|^{-\frac{1}{2}} e^{-\frac{1}{2}(x-\mu_2)^T \Sigma_2^{-1} (x-\mu_2)}$$

$$\begin{aligned} h(x) &= \log \left(\frac{p(x|y = 1)\theta}{p(x|y = -1)(1 - \theta)} \right) \\ &= \log \left(\sqrt{\frac{|\Sigma_2|}{|\Sigma_1|}} \frac{\theta}{1 - \theta} e^{-\frac{1}{2}(x-\mu_1)^T \Sigma_1^{-1} (x-\mu_1) + \frac{1}{2}(x-\mu_2)^T \Sigma_2^{-1} (x-\mu_2)} \right) \\ &= \frac{1}{2} \log \frac{|\Sigma_2|}{|\Sigma_1|} + \log \frac{\theta}{1 - \theta} - \frac{1}{2}(x - \mu_1)^T \Sigma_1^{-1} (x - \mu_1) + \frac{1}{2}(x - \mu_2)^T \Sigma_2^{-1} (x - \mu_2) \end{aligned}$$

Decision boundary ($h(x) = 0$): A quadratic curve

(b)

If $\Sigma_1 = \Sigma_2$, Quadratic terms in $h(x)$ cancel out and $h(x)$ becomes linear.

\therefore Decision boundary: A line

(c)

If $\Sigma_1 = \Sigma_2 = I_d$,

$$(x - \mu_1)^T (x - \mu_1) - (x - \mu_2)^T (x - \mu_2) = 2 \log \frac{\theta}{1 - \theta}$$

$$\|x - \mu_1\|^2 - \|x - \mu_2\|^2 = 2 \log \frac{\theta}{1 - \theta}$$

Decision boundary: A line perpendicular to the line connecting μ_1 and μ_2 .

Question 4. (a)

$$\begin{aligned} L(z) &= \log(1 + e^{-z}) \\ L'(z) &= \frac{-e^{-z}}{1 + e^{-z}} \\ L''(z) &= \frac{e^z}{(1 + e^z)^2} > 0 \end{aligned}$$

$\therefore L(z)$ is a convex function.

(b)

We can prove it using mathematical induction. By definition, it's true when $m = 2$.

Let's assume the inequality holds when $m = n$.

If $m = n + 1$ and $\sum_{i=1}^{n+1} \alpha_i = 1$,

$$\begin{aligned} f\left(\sum_{i=1}^{n+1} \alpha_i x_i\right) &= f\left((\alpha_1 + \alpha_2) \left(\frac{\alpha_1}{\alpha_1 + \alpha_2} x_1 + \frac{\alpha_2}{\alpha_1 + \alpha_2} x_2\right) + \sum_{i=3}^{n+1} \alpha_i x_i\right) \\ &\geq (\alpha_1 + \alpha_2) f\left(\frac{\alpha_1}{\alpha_1 + \alpha_2} x_1 + \frac{\alpha_2}{\alpha_1 + \alpha_2} x_2\right) + \sum_{i=3}^{n+1} \alpha_i f(x_i) \\ &\geq (\alpha_1 + \alpha_2) \left(\frac{\alpha_1}{\alpha_1 + \alpha_2} f(x_1) + \frac{\alpha_2}{\alpha_1 + \alpha_2} f(x_2)\right) + \sum_{i=3}^{n+1} \alpha_i f(x_i) \\ &= \sum_{i=1}^{n+1} \alpha_i f(x_i) \end{aligned}$$

\therefore We are done.

Question 5. (a)

$$\begin{aligned}
 P(Y = 1|X = x) &= \frac{e^{w_0 + w^T x}}{1 + e^{w_0 + w^T x}} \\
 P(Y = 0|X = x) &= \frac{1}{1 + e^{w_0 + w^T x}} \\
 \log \left(\frac{P(Y = 1|X = x)}{P(Y = 0|X = x)} \right) &= \log \left(e^{w_0 + w^T x} \right) \\
 &= w_0 + w^T x
 \end{aligned}$$

\therefore log-odds of success is a linear function.

(b) [Extra Credit]

$$\begin{aligned}
 P(y^i|x^i) &= \frac{e^{\theta_{y^i} \cdot x^i}}{\sum_c e^{\theta_c \cdot x^i}} \\
 L = \prod_i P(y^i|x^i) &= \prod_i \frac{e^{\theta_{y^i} \cdot x^i}}{\sum_c e^{\theta_c \cdot x^i}} \\
 l = \log L &= \sum_i \log \frac{e^{\theta_{y^i} \cdot x^i}}{\sum_c e^{\theta_c \cdot x^i}} \\
 &= \sum_i \left(\theta_{y^i} \cdot x^i - \log \sum_c e^{\theta_c \cdot x^i} \right) \\
 \therefore \frac{dl}{d\theta_c} &= \sum_i \frac{d}{d\theta_c} (\theta_{y^i} \cdot x^i) - \frac{d}{d\theta_c} \left(\sum_i \log \sum_c e^{\theta_c \cdot x^i} \right) \\
 &= \sum_i I(y_i = c) x^i - \sum_i \frac{e^{\theta_c \cdot x^i}}{\sum_c e^{\theta_c \cdot x^i}} x^i
 \end{aligned}$$

Question 6. (a)

$$\begin{aligned}
 E &= \sum_{u,i} \left(M_{ui} - \sum_{k=1}^r U_{uk} V_{ik} \right)^2 \\
 \frac{\partial E}{\partial U_{uk}} &= \sum_i 2 \left(M_{ui} - \sum_{k=1}^r U_{uk} V_{ik} \right) (-V_{ik}) \\
 \frac{\partial E}{\partial V_{ik}} &= \sum_u 2 \left(M_{ui} - \sum_{k=1}^r U_{uk} V_{ik} \right) (-U_{uk})
 \end{aligned}$$

(b)

$$\begin{aligned}
 E &= \sum_{u,i} \left(M_{ui} - \sum_{k=1}^r U_{uk} V_{ik} \right)^2 + \lambda \sum_{u,k} U_{uk}^2 + \lambda \sum_{i,k} V_{ik}^2 \\
 \frac{\partial E}{\partial U_{uk}} &= \sum_i 2 \left(M_{ui} - \sum_{k=1}^r U_{uk} V_{ik} \right) (-V_{ik}) + \lambda \cdot 2 \cdot U_{uk} \\
 \frac{\partial E}{\partial V_{ik}} &= \sum_u 2 \left(M_{ui} - \sum_{k=1}^r U_{uk} V_{ik} \right) (-U_{uk}) + \lambda \cdot 2 \cdot V_{ik}
 \end{aligned}$$