

On the Relationship Between Equivariant Predictive Models and Structural Causal Model Identification

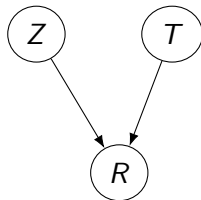
Grace Yin
Department of Statistics
University of British Columbia
MSC Presentation

April 19, 2022

Introduction

Structural Causal Model (SCM)

stone size treatment

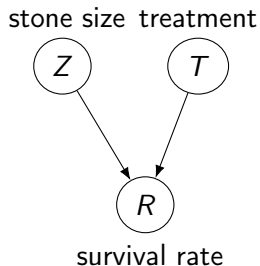


survival rate

$$\begin{cases} Z := f_1(N_Z) \\ T := f_2(N_T) \\ R := f_3(Z, T, N_R) \\ N_Z, N_T, N_R \text{ are jointly independent} \\ \text{noise terms} \end{cases}$$

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Introduction

Structural Causal Model (SCM) [2]

Definition

A structural causal model (SCM) $\mathcal{C} := (G, S, P_N)$ consists of a collection S of d (structural) assignments

$$X_j := f_j(PA_j, N_j), \quad j = 1, \dots, d$$

where $PA_j \subset \{X_1, \dots, X_d\} \setminus \{X_j\}$ are parents of X_j and $P_N = P_{N_1, \dots, N_d}$ is the joint (product) distribution over the noise variables which are assumed to be jointly independent. $G = (V, E)$ is the graph contains vertices and edges.

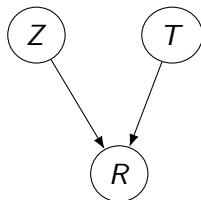
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Structural Causal Models (SCM)

intervention: $do(\dots)$: sets the variable value, without changing other nodes

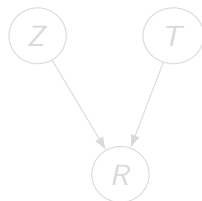
treatment: $T = A$ or $T = B$

stone size treatment



survival rate

stone size $do(T = A)$



survival rate

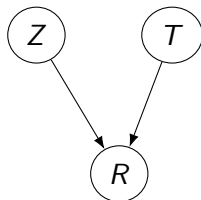
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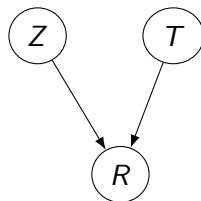
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Introduction

Question: Suppose there are two candidate SCMs, how can we identify which one is correct?

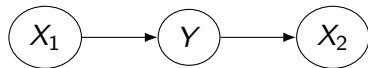


Figure: SCM 1

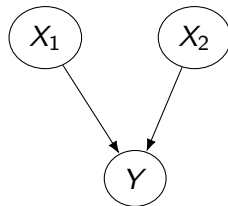


Figure: SCM 2

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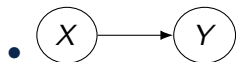
- Potential solution: conditional independence test
 - High computational costs [1]



- the functional assignment: $Y := f(X, E) = \beta_0 + \beta_1 X + E$
 - has constant risk for $X' \leftarrow X + a, a \in \mathbb{R}(+)$
- Our approach: apply constant risk theorem
 - across interventions described by the action of a group

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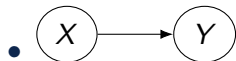
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Background and Notation

joint distribution of (X, Y) :

$$\tilde{P} = P \otimes Q_x,$$

P : marginal distribution on $(\mathbf{X}, \mathcal{X})$

Q_x : conditional distribution (Markov kernel) from $(\mathbf{X}, \mathcal{X})$ into $(\mathbf{Y}, \mathcal{Y})$

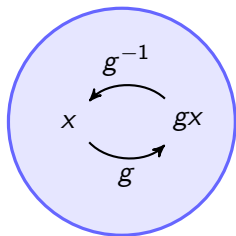
risk function is defined as

$$R(\tilde{P}, \rho) = \int_{\mathcal{X} \times \mathcal{Y} \times \mathcal{Z}} P(dx) Q_x(dy) \rho_x(dz) L(y, z).$$

ρ : decision procedure $\mathbf{X} \times \mathcal{Z} \rightarrow [0, 1]$ where $(\mathbf{Z}, \mathcal{Z})$ is the decision space

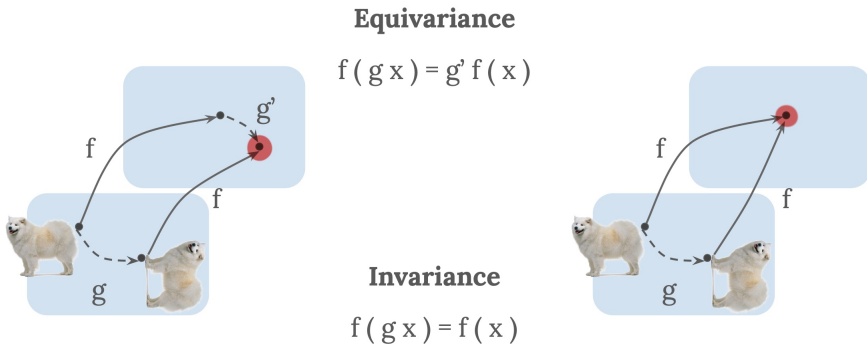
L : loss function $\mathcal{Y} \times \mathcal{Z} \rightarrow [-\infty, \infty)$.

Background and Notation



- \mathcal{G} : a group acting measurably on \mathbf{X} and \mathbf{Y}
- $g \in \mathcal{G}$: a group action, $\Phi(g, x) = gx$
- **conditional shift**: $g_x \tilde{P} = (P \circ g_x^{-1}) \otimes Q_x$
for $g_x \in \mathcal{G}$

Background and Notation



Figures adapted from Daniel E. Worrall

Figure: Equivariance and Invariance map [3]

Constant Risk Theorem

- equivariant markov kernel:

$$Q_{g_x} = Q_x \circ g^{-1}$$

- invariant loss function:

$$L(gy, gz) = L(y, z)$$

- equivariant decision procedure ρ :

$$\rho_{g_x} = \rho \circ g^{-1}$$

Theorem

For an invariant loss function L , the risk of a decision procedure ρ is constant under the conditional shift $g_x \tilde{P}$ for any group action $g_x \in \mathcal{G}$, if ρ is equivariant and the kernel Q_x is equivariant. i.e.,

$$\forall g_x \in \mathcal{G}, R(\rho, g_x \tilde{P}) = R(\rho, \tilde{P})$$

Identify the Structure of SCM

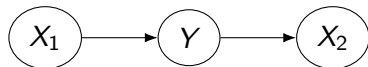


Figure: SCM 1

$$\begin{cases} X_1 \sim \mathcal{N}(0, \sigma^2) \\ Y := \beta_1 X_1 + \varepsilon_y \\ X_2 := \beta_2 Y + \varepsilon_2 \end{cases}$$

with $\varepsilon_2 \sim \mathcal{N}(0, 1)$, $\varepsilon_y \sim \mathcal{N}(0, \sigma^2)$

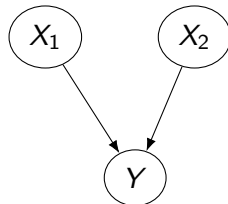


Figure: SCM 2

$$\begin{cases} X_1 \sim \mathcal{N}(0, \sigma^2) \\ X_2 := \varepsilon_2 \\ Y := \beta_1 X_1 + \beta_2 X_2 + \varepsilon_y \end{cases}$$

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Simulation Experiments and Discussion

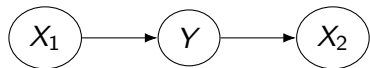


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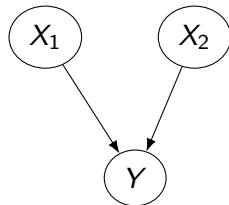


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intervention : $\begin{cases} X'_1 \leftarrow X_1 + g_1 \\ X'_2 \leftarrow X_2 + g_2 \end{cases}$

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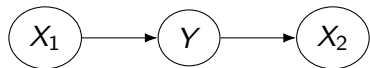


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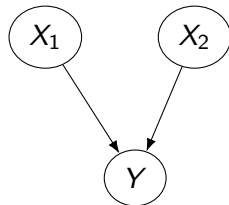


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Simulation Experiments and Discussion

Simulation experiments:

- simulate data from SCM 1
- simulate estimated coefficients in shifting environment 0
- compute risk for three models in different shifting environments

$$\left\{ \begin{array}{l} \ell_1 : \hat{y} = \underbrace{\hat{\beta}_0 + \hat{\beta}_1}_{\text{depend on } X_1} x_1 \\ \ell_2 : \hat{y} = \underbrace{\hat{\alpha}_0 + \hat{\alpha}_1}_{\text{depend on } X_2} x_2 \\ \ell_3 : \hat{y} = \underbrace{\hat{\gamma}_0 + \hat{\gamma}_1 x_1 + \hat{\gamma}_2 x_2}_{\text{depend on } X_1 \& X_2} \end{array} \right.$$

loss function: least square loss function

risk function:

$$R(\rho, P) = \int_{\mathcal{X} \times \mathcal{Y}} L(y, f(x)) dP(x, y), \text{ where } \rho = \delta_{f(x)}$$

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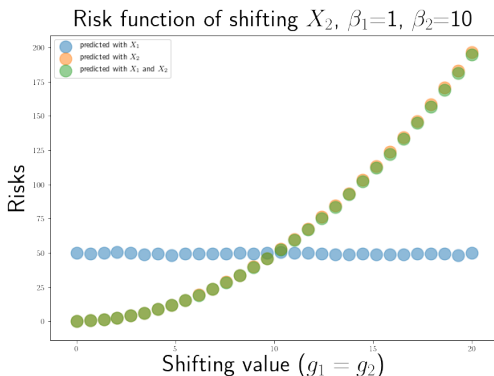
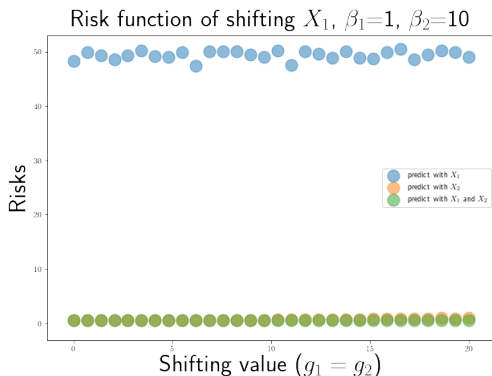
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Simulation Experiments and Discussion

The simulation results verified our constant risk theorem.

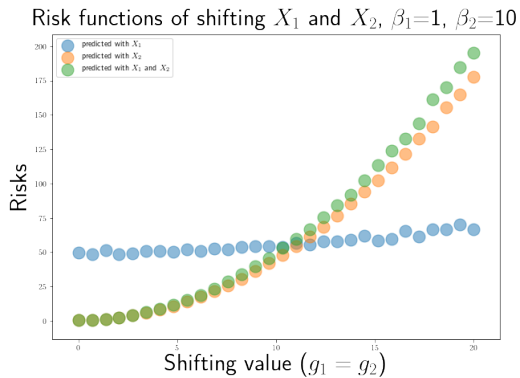
- shifting X_1 : constant risk for all three predictive models
- shifting X_2 : only ℓ_1 has constant risk



Simulation Experiments and Discussion

Differences between constant risk approach and risk minimization approach

- a crossover among three predictive models when shifting X_1 and X_2

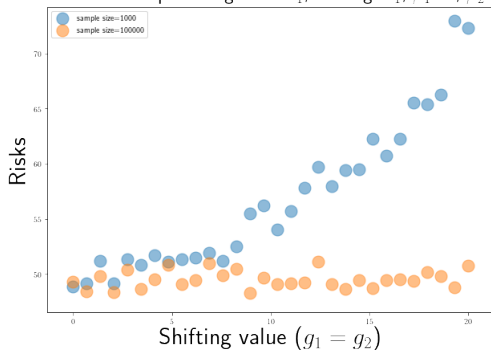


Simulation Experiments and Discussion

Sample size can influence the constant risk results:

- simulated $\hat{\beta}_0$ and $\hat{\beta}_1$ for ℓ_1
- sample size $n_1 = 1000$ vs sample size $n_2 = 100000$
- empirical risks: $\hat{R}(g) \propto (\beta_0 - \hat{\beta}_0)^2 + (x_{1,i} + g)^2(\beta_1 - \hat{\beta}_1)^2$

Risk functions of predicting with X_1 , shifting X_1 , $\beta_1=1$, $\beta_2=10$



Potential Extension

Potential directions:

- Apply on other examples of SCMs
- Identify the theoretical interconnection between the risk function among linear regression models
- Implement algorithms
- Explore non-linear predictive models

- [1] J. Peters, P. Bühlmann, and N. Meinshausen.
Causal inference using invariant prediction: identification and confidence intervals.
2015.
- [2] J. Peters, D. Janzing, and B. Schölkopf.
Elements of Causal Inference: Foundations and Learning Algorithms.
The MIT Press, 2017.
- [3] E. van der Pol.
Geometric deep learning and reinforcement learning, February 2021.

Acknowledgement

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- my supervisor: Prof. Ben Bloem-Reddy
- the various members of the UBC Department of Statistics

Github QR code:



Thank You for Listening!