The Econometrics of DSGE Models

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EIEF

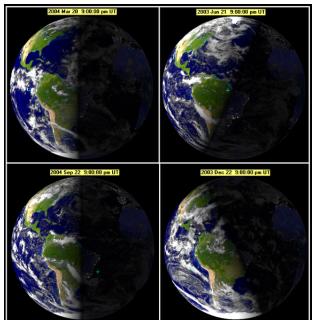
Lecture 9: Twilight zone of DSGE models

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Twilight Zone

The **twilight zone** or "grey line" is a moving line that separates the illuminated day side and the dark night side of a planetary body

Twilight Zone



Twilight Zone of DSGE Model

- Hybrid approach to DSGE estimation
- Forecasting performances of DSGE models
- Go beyond the linear approximation
- Identification

- The basic idea is to combine an unrestricted VAR and DSGE model
- ullet The combination is indexed by a scalar parameter λ
 - ▶ $\lambda \to \infty$ the model converges to the DSGE model
 - $\lambda \to 0$ the model converges to the unrestricted VAR
- The "best" hybrid model is the one associated with the $\hat{\lambda}$, the value of λ the correspond to the highest value of the marginal likelihood.

Details of the procedure

We start with a VAR, whose likelihood is given by

$$L(Y|\Phi,\Sigma)$$

where Φ and Σ are parameters, e.g.

$$Y_t = \Phi Y_{t-1} + u_t, \quad E[u_t u_t'] = \Sigma$$

• Given priors $p(\Phi, \Sigma)$ the posterior is formally given

$$p(\Phi, \Sigma|Y) \propto L(Y|\Phi, \Sigma)p(\Phi, \Sigma)$$

Suppose now that priors depends on "hyper-parameters"

$$p(\Phi, \Sigma | \theta, \lambda)$$

where θ are the structural parameters of a DSGE model



Details of the procedure

The joint posterior

$$p(\Phi, \Sigma, \theta | Y, \lambda) = \frac{L(Y | \Phi, \Sigma) p(\Phi, \Sigma | \theta, \lambda)}{\int_{\Phi, \Sigma} L(Y | \Phi, \Sigma) p(\Phi, \Sigma | \theta, \lambda) d(\Phi, \Sigma)}$$

Notice that

$$p(Y|\theta,\lambda) = \int_{\Phi,\Sigma} L(Y|\Phi,\Sigma)p(\Phi,\Sigma|\theta,\lambda)d(\Phi,\Sigma),$$

is the marginal likelihood, which can be evaluated (under "conditions") analytically for given values of θ and λ .

ullet We search for the "best" λ

$$\hat{\lambda} = rg \max_{\lambda \in \Lambda} p(Y|\theta,\lambda)$$



Details of the procedure

- Conditions:
 - Likelihood:

$$L(Y|\Phi,\Sigma) = \prod_{t=1}^{T} \rho(Y_t|Y_{t-1},\Phi,\Sigma),$$

where the $p(\cdot|\cdot,\Phi,\Sigma)'s$ are is normal densities

Priors

$$p(\Phi, \Sigma | \theta, \lambda) \propto \underbrace{p(\Phi | \Sigma, \theta, \lambda)}_{\text{Normal}} \underbrace{p(\Sigma | \theta, \lambda)}_{\text{Inverted Wishal}}$$

- ▶ The hyperparameters of the prior distribution of Φ and Σ are obtained from the (linearized DSGE model) with (structural) parameter θ
- Under these conditions
 - The posterior can be simulated using the Gibbs sampling (marginals are normal and inverted Wishart)

DSGE models: forecasting performances

Dynamic stochastic general equilibrium (DSGE) models use modern macroeconomic theory:

- to explain comovements of aggregate time series over the business cycle
- to forecast future values of economic aggregates
- to perform policy analysis

How well these models performs in terms of forecast performances?

DSGE models: forecasting performances

- Dynamic stochastic general equilibrium (DSGE) models have been trashed, bashed, and abused during the Great Recession and after
- One of the many reasons for the bashing was the models' alleged inability to forecast the recession itself.
- Oddly enough, there's little evidence on the forecasting performance of DSGE models during this turbulent period.

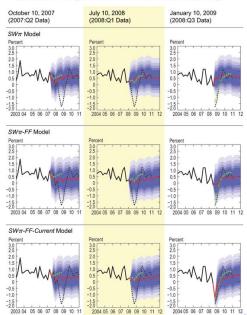
DSGE models: forecasting performances

"DSGE Model-Based Forecasting," prepared for Elsevier's Handbook of Economic Forecasting, Del Negro, Schorfheide, Herbst

Findings:

- what it really matters what information you feed into your model: Feed in the right information, and even a dingy DSGE model may not do so poorly at forecasting the recession.
- compared with the "Blue Chip Economic Consensus" forecasts, DSGE models do about the same, if not better, in fall 2007, in summer 2008, before the Lehman crisis, and at the beginning of 2009–provided one incorporates up-to-date financial data into the DSGE model.

DSGE models: forecasting performances Forecasts for Output Growth: DSGE versus Blue Chip



Nonlinearities

- DSGE models are essentially nonlinear representation of our economic reality
- When estimated a DSGE is not actually solved, but rather the solution is approximated by a log-linear function whose coefficients are nonlinear functions of model parameters.
- What passes for the DSGE model is actually the driving processes passed through a filter.
- How accurate is this approximation?

Consider the prototypical DSGE model

$$E_t[f(y_{t+1}, y_t, x_{t+1}, x_t)] = 0$$

where

• x_t : denotes predetermined (or *state*) variables

$$x_t = \begin{pmatrix} x_{1t} \\ x_{2t} \end{pmatrix}$$

where x_{1t} endogenous **predetermined** states and x_{2t} exogenous state

- y_t : denotes predetermined (or *control*) variables
- x_0 : the initial condition for the economy is given

Example: Neoclassical growth model

Households

$$\max E_0 \sum_{t=1}^{\infty} eta^t rac{c_t^{1-\gamma}}{1-\gamma}, \quad 0 < eta < 1, ext{ and } \gamma
eq 1$$

The period-by-period budget constraint

$$A_t k_t^{\alpha} = c_t + k_{t+1} - (1 - \delta)k_t$$

where k_t is the capital stock. In period t the capital is **predetermined.**

ullet The variable A_t denotes exogenous technological change

$$(A_{t+1}-1) = \rho(A_t-1) + \sigma \varepsilon_{t+1}$$



Example: Neoclassical growth model

The Lagrangian of the household's optimization problem

$$\mathscr{L} = E_0 \sum_{t=0}^{\infty} \beta^t \left\{ \frac{c_t^{1-\gamma}}{1-\gamma} + \lambda_t \left[A_t k_t^{\alpha} - c_t - k_{t+1} + (1-\delta) k_t \right] \right\}$$

The first order optimality conditions are

$$c_t^{-\gamma} = \beta E_t c_{t+1}^{-\gamma} [\alpha A_{t+1} k_{t+1}^{\alpha - 1} + 1 - \delta]$$
$$A_t k_t^{\alpha} = c_t + k_{t+1} - (1 - \delta) k_t$$

- k_t is endogenously **predetermined**, so it belongs to x_{1t}
- A_t is exogenous state, so it belongs to x_{2t}
- A_t denotes exogenous technological change

Example: Neoclassical growth model

Let

$$x_t = \begin{pmatrix} k_t \\ A_t \end{pmatrix}, \quad y_t = c_t$$

then

$$E_{t}f(y_{t+1}, y_{t}, x_{t+1}, x_{t}) = E_{t} \begin{bmatrix} y_{t}^{-\gamma} - \beta y_{t+1} \left(\alpha x_{2t+1} x_{1t+1}^{\alpha-1} + 1 - \delta \right) \\ y_{t} + x_{t+1} - x_{2t} x_{1t}^{\alpha} - (1 - \delta) x_{1t} \\ (x_{2t+1} - 1) - \rho(x_{2t} - 1) \end{bmatrix}$$

Policy functions

 The solution to models belonging to the class described previously is given by

$$y_t = g(x_t)$$

$$x_{t+1} = h(x_t) + \eta \sigma \varepsilon_{t+1}$$

• The matrix η is

$$oldsymbol{\eta} = egin{bmatrix} 0 \ ilde{oldsymbol{\eta}} \end{bmatrix}$$

with the dimension of 0 equal to the dimension of x_{1t} .

That is,

$$y_t = g(x_t)$$

$$x_{1t+1} = h_1(x_t)$$

$$x_{2t+1} = h_2(x_t) + \tilde{\eta} \sigma \varepsilon_{t+1}$$

• The key idea of perturbation methods is to express the solution as function of the state vector **and** of σ , the amount of uncertainty in the economy

$$y_t = g(x_t, \sigma)$$

$$x_{t+1} = h(x_t, \sigma) + \eta \sigma \varepsilon_{t+1}$$

- Given this interpretation, a perturbation methods finds a local approximation of the functions g and h
- By local approximation, we mean an approximation that is valid at a particular point $(\bar{x}, \bar{\sigma})$

Taking a series approximation of of the function g and h around this point we have

$$g(x,\sigma) = g(\bar{x},\bar{\sigma}) + g_x(\bar{x},\bar{\sigma})(x-\bar{x}) + g_\sigma(\bar{x},\bar{\sigma})(\sigma-\bar{\sigma})$$

$$+ \frac{1}{2}g_{xx}(\bar{x},\bar{\sigma})(x-\bar{x})^2 + \frac{1}{2}g_{\sigma\sigma}(\bar{x},\bar{\sigma})(\sigma-\bar{\sigma})^2$$

$$+ \frac{1}{2} + g_{x\sigma}(\bar{x},\bar{\sigma})(\sigma-\bar{\sigma})(x-\bar{x}) + \cdots$$

and

$$h(x,\sigma) = h(\bar{x},\bar{\sigma}) + h_x(\bar{x},\bar{\sigma})(x-\bar{x}) + h_\sigma(\bar{x},\bar{\sigma})(\sigma-\bar{\sigma})$$

$$+ \frac{1}{2}h_{xx}(\bar{x},\bar{\sigma})(x-\bar{x})^2 + \frac{1}{2}h_{\sigma\sigma}(\bar{x},\bar{\sigma})(\sigma-\bar{\sigma})^2$$

$$+ \frac{1}{2} + h_{x\sigma}(\bar{x},\bar{\sigma})(\sigma-\bar{\sigma})(x-\bar{x}) + \cdots$$

Let

$$F(x,\sigma) = E_t f(y_{t+1}, y_t, x_{t+1}, x_t)$$

= $E_t f(g(h(x,\sigma) + \eta \sigma \varepsilon'), g(x,\sigma), h(x,\sigma) + \eta \sigma \varepsilon', x)$

Notice that

$$F(x,\sigma)=0$$

for all x and σ

• This simple results, together with the fact that

$$\frac{\partial F(x,\sigma)}{\partial^i x \partial^j \sigma} = 0, \quad \forall x, \sigma, i, j$$

is used to identify $h_x, g_x, h_\sigma, g_\sigma, h_{\sigma x}, g_{\sigma x}, \dots$

- What is a good local point? The steady state of the economy, which coincides with the case $\sigma=0$
- ullet Notice that at the deterministic steady-state of the economy $(\sigma=0)$

$$g(\bar{x},0) = \bar{y}$$
$$h(\bar{x},0) = \bar{x}$$

- The reason why the steady state is particularly convenient is that in most cases is possible to solve for the steady state
- With the steady state values at hand, you can find the derivatives of F

 If we limit ourselves to the linear approximation of the policy functions, we have

$$g(x,0) = g(\bar{x},0) + g_x(\bar{x},0)(x-\bar{x}) + g_\sigma(\bar{x},0)\sigma$$

$$h(x,0) = h(\bar{x},0) + h_x(\bar{x},0)(x-\bar{x}) + h_\sigma(\bar{x},0)\sigma$$

We know that

$$g(\bar{x},0) = \bar{y}, \quad h(\bar{x},0) = \bar{x}$$

- The remaining unknown coefficient are the four first derivatives: $g_x(\bar{x},0), h_x(\bar{x},0), g_\sigma(\bar{x},0), h_\sigma(\bar{x},0)$
- We have

$$F_{\sigma}(\bar{x},0) = f_{y'}[g_{x}h_{\sigma} + g_{\sigma}] + f_{y}g_{\sigma} + f_{x'}h_{\sigma}$$

$$F_{x}(\bar{x},0) = f_{y'}g_{x}h_{x} + f_{y}g_{x} + f_{x'}h_{x} + f_{x}$$

Solving the linear system of equation we have

$$0 = F_{\sigma}(\bar{x}, 0) \implies [f_{y'}g_x + f_{x'} \quad f_{y'} + f_y] \begin{pmatrix} h_{\sigma} \\ g_{\sigma} \end{pmatrix} = 0$$

- This is *linear* and *homogenous* in g_{σ} and h_{σ} system of equations
- If a unique solution exists, we have that

$$h_{\sigma}=0, \quad g_{\sigma}=0$$

- Important theoretical results: in general, up to first order, one need not to correct the constant term of the approximation of the policy function for the size of the variance of the shocks
- ullet In the linear approximation, certainty equivalence holds the policy function is independent of of the variance covariance matrix of $arepsilon_t$

• To find g_x and h_x , observe that

$$\begin{bmatrix} f_{X'} & f_{y'} \end{bmatrix} \begin{pmatrix} I \\ g_X \end{pmatrix} h_X = -\begin{bmatrix} f_X & f_y \end{bmatrix} \begin{pmatrix} I \\ g_X \end{pmatrix}$$

- Let $A = [f_{x'} \quad f_{y'}]$ and $B = -[f_x \quad f_y]$. Both A and B are known
- Let P the matrix of eigenvector be such

$$h_{x}P=P\Lambda$$

where Λ is diagonal. Let also

$$Z = \begin{pmatrix} I \\ g_X \end{pmatrix} P$$

Then, from

$$[f_{x'} \quad f_{y'}] \begin{pmatrix} I \\ g_x \end{pmatrix} h_x P = -[f_x \quad f_y] \begin{pmatrix} I \\ g_x \end{pmatrix} P,$$

it follows

$$AZ\Lambda = BZ$$



• We can map the above problem into a generalized eigenvalue problem

$$AZ\Lambda = BZ$$

• A and B can be written as (in matlba [V,D] = eig(B,A))

$$AV\Lambda = BV$$

then

$$A[V_1 \quad V_2] \begin{bmatrix} D_{11} & 0 \\ 0 & D_{22} \end{bmatrix} = B[V_1 \quad V_2]$$

Comparing terms,

$$\Lambda = D_{11}$$

and

$$\begin{pmatrix} I \\ g_x \end{pmatrix} P = V_1 \equiv \begin{bmatrix} V_{11} \\ V_{12} \end{bmatrix}$$

Thus,

$$h_{x} = V_{11}D_{11}V_{11}^{-1}$$

 $g_{x} = V_{12}V_{11}^{-1}$



Problem with first order approximation
In many economic application we are interested in finding the effect of uncertainty on the economy
 e.g., up to first order, the mean of the rate of return of all assets must be the same; we cannot study risk premia with linear approximated models
* e.g., how uncertainty affect welfare cannot be studied with linear approximation; any two policies that give rise to the same steady state yield, up to first order, the same level of welfare

To go beyond the linear approximation

Given

$$y_t = g(x_t, \sigma)$$

 $x_{t+1} = h(x, \sigma) + \eta \sigma \varepsilon_{t+1}$

There are three main solutions to estimate θ

- Extended Kalman Filter
 - Unscented Kalman Filter
 - Particle Filter

Particle Filter

Needed:

- Particle methods assume x_k and the observations y_k can be modeled in this form:

$$x_k | x_{k-1} \sim p_{x_k | x_{k-1}}(x | x_{k-1})$$

and with an initial distribution $p(x_0)$.

3 The observations Y_0, Y_1, \ldots are conditionally independent provided that x_0, x_1, \ldots are known

Particle Filter

- Fernandez-Villaverde, Jesus, and Juan F. Rubio-Ramirez. "Estimating macroeconomic models: A likelihood approach." The Review of Economic Studies 74.4 (2007): 1059-1087.
- Flury, Thomas, and Neil Shephard. "Bayesian inference based only on simulated likelihood: particle filter analysis of dynamic economic models." Econometric Theory 27.5 (2011): 933.
- Fernandez-Villaverde, Jesus. "The econometrics of DSGE models." SERIEs 1.1-2 (2010): 3-49.

Alternative estimation

- Indirect Inference
- Approximate Bayesian Computation (ABC)
- Going back to drawing board (Gallant, Giacomini, Ragusa; 2013)