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- · Want to know how network position of an agent affects economic outcome
- Network is high dimensional ⇒ summarize using centrality measures
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- · Suppose researchers observe agents in a network
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- Use OLS to study relationship between outcome and centrality
- When is linear regression with centrality measures statistically valid?

#### Examples

- Better-networked venture capital firms are more profitable (Hochberg, Ljungqvist, and Lu, 2007)
- Greater take-up when seeding microfinance to central villagers (Banerjee, Chandrasekhar, Duflo, and Jackson, 2013)
- Central families overrepresented in political offices in the Philippines (Cruz, Labonne, and Querubin, 2017)

Researchers study:

$$Y_i = C_i \beta + X'_i \gamma + \varepsilon_i$$
,  $E[X_i \varepsilon_i] = 0, E[C_i \varepsilon_i] = 0$  centrality controls

- $\beta$  is parameter of interest
- Different centrality measures capture different ways of being important
- Estimate  $\beta$  by OLS; conduct inference using t-test

# Statistical Challenges

- Networks may be sparse
  - · Many more agents than links per agent
  - · Not enough variation to identify eta
  - Sparsity is "stylized fact"

#### Statistical Challenges

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  - · Not enough variation to identify  $\beta$
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  - · Interaction of interest may not be observed
  - · Use a related network instead
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## Statistical Challenges

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- Networks may be constructed using proxies
  - · Interaction of interest may not be observed
  - · Use a related network instead
  - E.g. want risk sharing network, have friendship network
- ⇒ Weak signals harder to separate from noise

# When is linear regression on centrality measures statistically valid?

- · Degree, diffusion and eigenvector centralities
- Cross-sectional: one large network
- Novel asymptotic framework with sparse, proxy networks
- More similar to data  $\Rightarrow$  asymp. approx. more accurate in finite sample

#### Contributions

- 1. Show that OLS can become inconsistent with sparse, proxy networks
  - Characterize threshold at which inconsistency occurs
  - · Show that eigenvector is less robust than degree and diffusion
  - · Comparing significance involves both economic and statistical properties
  - Rule-of-Thumb for sparsity regime

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  - Rule-of-Thumb for sparsity regime
- 2. Distributional theory with sparse, proxy networks
  - Even when consistent, OLS estimators are asymptotically biased
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- 3. Novel bias correction and inference methods

#### Related Literature

- Linear Regression with Centrality Measures
  - Eigenvectors: Le and Li (2020); Cai, Yang, Zhu, Shen, and Zhao (2021)
- Estimation of Centrality Statistics with Proxy Networks
  - · Simulations: Costenbader and Valente (2003); Borgatti, Carley, and Krackhardt (2006)
  - Theory: Dasaratha (2020); Avella-Medina, Parise, Schaub, and Segarra (2020)
- Econometric Challenges from Sparse Networks
  - Network Formation: Jochmans (2018); Graham (2020b)

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- 6. Conclusion

- Network data is symmetric adjacency matrix A
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- ·  $A_{ij}$  records intensity of relationship between i and j
- · Centrality measures summarize network positions into "importance"
- · When A is known, centrality measures exactly computable
- Many different ways to measure importance ⇒ many centrality measures
- Focus on degree, diffusion and eigenvector

# Degree and Diffusion Centralities

• Degree is the total intensity of direct connections:

$$C^{(1)}(A) = A\iota$$

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• Degree is the total intensity of direct connections:

$$C^{(1)}(A) = A\iota$$

• Diffusion reflects ability to broadcast messages in a network:

$$C^{(T)}(A) = \left(\sum_{t=1}^{T} \delta^t A^t\right) \iota$$

 $\cdot$  *T*,  $\delta$  are chosen by researchers

## **Eigenvector Centrality**

• Eigenvector is the leading eigenvector of A, scaled

$$C^{(\infty)}(A) = a_n v_1(A)$$

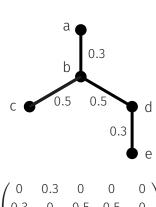
· Friends of important agents are themselves more important

# **Eigenvector Centrality**

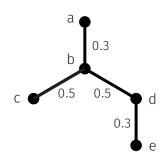
- Various choices of  $a_n$  in literature
- $a_n$  turns out to matter for statistical properties
- Today:  $a_n = \sqrt{\lambda_1(A)}$

	Applied Work	Econometrics
$a_n = 1$	Banerjee et al. (2013) Cruz et al. (2017)	Dasaratha (2020)
$a_n = \sqrt{n}$	Chandrasekhar et al. (2018) Banerjee et al. (2019)	Avella-Medina et al. (2020) Cai et al. (2021)

**Table 1:** Examples of  $a_n$  in econometric theory and empirical work



Degree  $C^{(1)}$ 

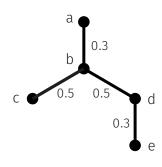


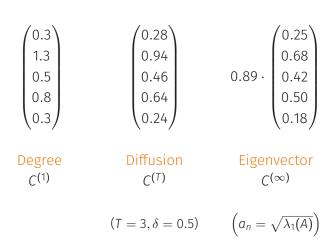
$$\begin{pmatrix}
0.3 \\
1.3 \\
0.5 \\
0.8 \\
0.3
\end{pmatrix}$$

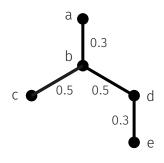
$$\begin{pmatrix}
0.28 \\
0.94 \\
0.46 \\
0.64 \\
0.24
\end{pmatrix}$$

# Diffusion $C^{(T)}$

$$(T = 3, \delta = 0.5)$$



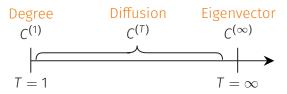




- When T = 1,  $C^{(1)}(A) \propto C^{(T)}(A)$
- Banerjee et al. (2019): If  $\delta$  is larger than the inverse of leading eigenvalue of A,

$$\lim_{T\to\infty}C^{(T)}(A)\propto C^{(\infty)}(A)$$

• The statistics we study can thus be represented on a line:



Model and Assumptions

#### Model and Assumptions

- For simplicity, consider regression without other covariates
- For  $d \in \{1, T, \infty\}$ :

$$Y_i = \beta^{(d)} C_i^{(d)} + \varepsilon_i^{(d)}$$

- $\cdot$  Will make enough assumptions so  $eta^{(d)}$  is slope of CEF
- DGP yields i.i.d. draws of  $\{(\varepsilon_i, U_i)\}$
- ·  $U_i \sim U[0,1]$  is unobserved latent type used to construct network

#### **True Network**

- Let A be the  $n \times n$  symmetric adjacency matrix
- For  $f:[0,1]^2 \to [0,1]$ ,  $p_n \in (0,1]$  and j > i, let

$$A_{ij} := p_n f(U_i, U_j)$$

Symmetry:  $A_{ij} = A_{ji}$ ; normalisation:  $A_{ii} = 0$ 

- When  $A_{ij}$  is large, agents i and j have a strong relationship
- $p_n \rightarrow 0$  reflects sparsity

#### Examples

#### Informal Insurance and Consumption Smoothing

- $Y_i$ : variance in consumption expenditure
- $A_{ij}$ : probability that i lends j money or vice versa
- *U<sub>i</sub>*: social class

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#### **Graphons** (*f*)

- "Full" Insurance:  $f(U_i, U_j) = 1$
- Assortative:  $f(U_i, U_j) = 1 (U_i U_j)^2$

#### **True Network**

- · When A is observed, can exactly compute centrality measures
- Form the estimator

$$\tilde{\beta}^{(d)} = \frac{\mathsf{Y}'\mathsf{C}^{(d)}}{\left(\mathsf{C}^{(d)}\right)'\mathsf{C}^{(d)}}$$

## **Proxy Network**

• When A is not observed, use  $\hat{A}$ : for j > i,

$$\hat{A}_{ij} \mid \mathbf{U} \stackrel{\text{iid}}{\sim} \text{Bernoulli}(A_{ij})$$

Below the diagonal,  $\hat{A}_{ij} = \hat{A}_{ji}$ 

Proxy error is "white noise"

# Proxy Network

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Below the diagonal,  $\hat{A}_{ij} = \hat{A}_{ji}$ 

- · Proxy error is "white noise"
- Use  $\hat{A}$  as plug-in for A to compute  $\hat{C}^{(d)}$
- Estimate

$$\hat{\beta}^{(d)} = \frac{Y'\hat{C}^{(d)}}{\left(\hat{C}^{(d)}\right)'\hat{C}^{(d)}}$$

#### Examples

#### Informal Insurance and Consumption Smoothing

- $Y_i$ : variance in consumption expenditure
- $A_{ij}$ : probability that i lends j money or vice versa
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#### **Proxy Networks**

- Social:  $\hat{A}_{ij} = 1$  if either *i* or *j* reports the other as a friend
- Financial:  $\hat{A}_{ij} = 1$  if *i* borrows from or lends to *j*

## Sparsity

- · Want to understand the properties of  $\tilde{\beta}^{(d)}$  and  $\hat{\beta}^{(d)}$  when networks are sparse
- Theoretical device:  $p_n \to 0$  as  $n \to \infty$ :

$$A_{ij} := \underset{p}{p}_{n} f(U_{i}, U_{j})$$
 ,  $\hat{A}_{ij} = \text{Bernoulli}\left(\underset{p}{p}_{n} f(U_{i}, U_{j})\right)$ 

- Weak interaction in true network:  $A_{ii} \rightarrow 0$
- Sparse proxy networks: many  $\hat{A}_{ij}$ 's are 0
- Rate at which  $p_n \to 0$  reflects different amounts of sparsity



#### Model and Assumptions

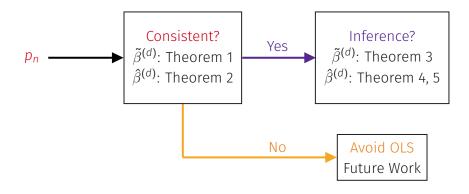
- · Proxy Networks: Â is observed but A economically meaningful
  - Node-Level Regressions: Auerbach (2022), Le and Li (2020), Cai et al. (2021)
  - Dyadic Models: see e.g. De Paula (2017) Section 3, Graham (2020a) Section 6
- Bickel-Chen Model of Sparsity (Bickel and Chen 2009)
  - Graham (2020a): "The Bickel-Chen model is the default one in the nonparametric statistics and machine learning literatures on random graphs."
  - Dyadic Models: Jochmans (2018), Graham (2020b)
- Standard models of proxy networks and sparsity to study linear regression

#### **Preview of Results**

• Main question: how do  $\tilde{\beta}^{(d)}$  and  $\hat{\beta}^{(d)}$  behave at as we vary  $p_n$ ?

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Theoretical Results

# Consistency with True Network A

#### Theorem 1

 $\tilde{\beta}^{(1)}$  and  $\tilde{\beta}^{(T)}$  are consistent if and only if

$$p_n\gg n^{-\frac{3}{2}}\;.$$

$$\tilde{\beta}^{(\infty)}$$
 is consistent if  $a_n = \sqrt{\lambda_1(A)}$ .

- $\cdot n \cdot n^{-3/2} \rightarrow 0$
- Inconsistent only under extreme sparsity
- Consistency requires  $\|C^{(d)}\|_2 \to \infty$  w.p.a. 1
- $\cdot$   $a_n$  inflates eigenvector, counters sparsity

# Consistency with Proxy Network $\hat{A}$

#### Theorem 2

 $\hat{\beta}^{(1)}$  and  $\hat{\beta}^{(T)}$  are consistent if and only if

$$p_n \gg n^{-1}$$
.

- $\cdot$   $\hat{\beta}^{(1)}$ ,  $\hat{\beta}^{(7)}$  less robust to sparsity than  $\tilde{\beta}^{(1)}$ ,  $\tilde{\beta}^{(7)}$
- Consistency with proxy networks in dense regime

# Consistency with Proxy Network $\hat{A}$

#### Theorem 2

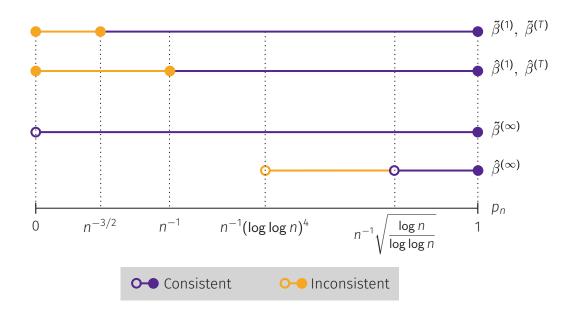
$$\hat{\beta}^{(\infty)}$$
 consistent if

$$p_n \gg n^{-1} \sqrt{\frac{\log n}{\log \log n}} \ .$$

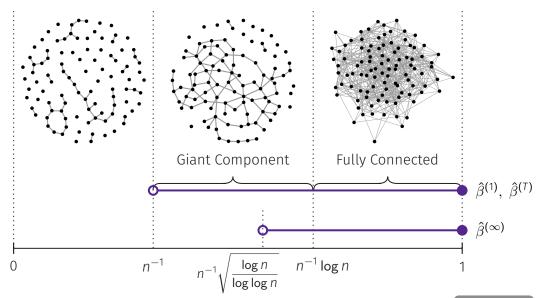
It is inconsistent if

$$n^{-1} (\log \log n)^4 \ll p_n \ll n^{-1} \sqrt{\frac{\log n}{\log \log n}}.$$

# **Consistency Thresholds**



# Rule of Thumb for Consistency



## Theoretical Results

Distributional Theory

## Distributional Theory with True Network A

#### Theorem 3

For  $d \in \{1, T, \infty\}$ , suppose  $\tilde{\beta}^{(d)}$  is consistent. Then,

$$\frac{\tilde{\beta}^{(d)} - \beta^{(d)}}{\sqrt{V_0^{(d)}}} \stackrel{d}{\rightarrow} N(0,1) ,$$

where 
$$V_0^{(d)} = E[(C_i^{(d)})^2]^{-2} E[(C_i^{(d)}\varepsilon_i)^2].$$

#### Theorem 4

Suppose for  $d \in \{1, T, \infty\}$  that  $\hat{\beta}^{(d)}$  is consistent. Then if  $\beta^{(d)} = 0$ ,

$$\frac{\hat{\beta}^{(d)}}{\sqrt{V_0^{(d)}}} \stackrel{d}{\to} N(0,1)$$

where 
$$V_0^{(d)} = E\left[\left(C_i^{(d)}\right)^2\right]^{-2} E\left[\left(C_i^{(d)}\varepsilon_i\right)^2\right].$$

- · Plug-in estimation works for  $V_0^{(d)}$  in the cases above
- Robust/hc t-statistic appropriate for  $\tilde{\beta}^{(d)}$  for any null
- Appropriate for  $\hat{\beta}^{(d)}$  if null is  $\beta^{(d)} = 0$ .

#### Theorem 4

For  $d \in \{1, T\}$ , suppose  $\hat{\beta}^{(d)}$  is consistent. If  $\beta^{(d)} \neq 0$ ,

$$\frac{\hat{\beta}^{(d)} - \beta^{(d)} \left(1 - B^{(d)}\right)}{\sqrt{V^{(d)}}} \stackrel{d}{\to} N(0, 1)$$

$$\underbrace{\sqrt{V_0^{(d)}}}_{O_p\left(n^{-3/2}p_n^{-1}\right)} \ll \underbrace{\sqrt{V^{(d)}}}_{O_p\left(n^{-1}p_n^{-1/2}\right)}^{p_n \to 0} \stackrel{\mathcal{B}^{(d)}}{\ll} O_p\left(n^{-1}p_n^{-1}\right)$$

- Using proxy network slows down rate of convergence
- Bias can be much larger than variance if  $p_n \to 0$
- · Bias correction necessary for obtaining non-degenerate limit distribution
- hc/robust t-statistic not appropriate when  $\beta^{(d)} \neq 0$
- t-statistic based confidence intervals are invalid

- $\hat{B}^{(d)}$  and  $\hat{V}^{(d)}$  in paper
- · Bias-corrected estimator:

$$\check{\beta}^{(d)} = \frac{\hat{\beta}^{(d)}}{1 - \hat{\beta}^{(d)}}$$

- Adjusted tests and confidence intervals in paper
- No additional data requirement;  $p_n$  need not be specified
- · Bias estimation is challenging since  $B^{(d)}$  much larger than  $\sqrt{V^{(d)}}$

#### Theorem 5

Suppose f has rank R  $< \infty$ . Suppose also for  $\eta > 0$  that

$$p_n \gg n^{-1} \left( \frac{\log n}{\log \log n} \right)^{\frac{1}{2} + \eta} \tag{1}$$

and that  $a_n = \sqrt{\lambda_1(A)}$ . Then,

$$\frac{\hat{\beta}^{(\infty)} - \beta^{(\infty)}}{\sqrt{V_0^{(\infty)}}} \stackrel{d}{\to} N(0,1) , \qquad (2)$$

where

$$V_0^{(\infty)} = E\left[\left(C_i^{(\infty)}\right)^2\right]^{-2} E\left[\left(C_i^{(\infty)}\varepsilon_i\right)^2\right]$$

## **Eigenvector Centrality**

- Chose  $a_n$  so that distribution is easy to characterize
- Usual hc/robust *t*-statistic valid for all null hypotheses
- Trade-off: choosing a model with slower, known rate of convergence for one with faster, unknown rate
- · Using  $\sqrt{\lambda_1(\hat{A})}$  does not change result

	Proxy Network		True Network	
	$\beta^{(1)}/\beta^{(T)}$	$eta^{(\infty)}$	$\beta^{(1)}/\beta^{(T)}/\beta^{(\infty)}$	
$H_0:\beta^{(d)}=0$		Dense: t-test*		
$H_0: \beta^{(d)} = b$				
Confidence Intervals				

**Table 2:** \*:  $p_n \gg n^{-1/2}$  (Le and Li, 2020). For  $\beta^{(\infty)}$ ,  $a_n = \sqrt{\lambda_1(A)}$ .

	Proxy Network		True Network
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			t-test
$H_0: \beta^{(d)} = b$			
Confidence Intervals			t-stat based

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	Proxy Network		True Network
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$H_0: \beta^{(d)} = 0$	t-test	Dense: <i>t</i> -test*	t toot
$H_0: \beta^{(d)} = b$	New Method		t-test
Confidence Intervals	New Method		t-stat based

Table 2: \*:  $p_n \gg n^{-1/2}$  (Le and Li, 2020). For  $\beta^{(\infty)}$ ,  $a_n = \sqrt{\lambda_1(A)}$ . Key: Theorem 3, Theorem 4,

	Proxy Network		True Network
	$\beta^{(1)}/\beta^{(T)}$	$\beta^{(\infty)}$	$\beta^{(1)}/\beta^{(T)}/\beta^{(\infty)}$
$H_0: \beta^{(d)} = 0$	t-test	Dense: <i>t</i> -test* Sparse: <i>t</i> -test	t-test
$H_0: \beta^{(d)} = b$	New Method	t-test	t-test
Confidence Intervals	New Method	t-stat based	t-stat based

Table 2: \*:  $p_n \gg n^{-1/2}$  (Le and Li, 2020). For  $\beta^{(\infty)}$ ,  $a_n = \sqrt{\lambda_1(A)}$ . Key: Theorem 3, Theorem 4, Theorem 5

# Simulations

#### **Simulations**

• Suppose f = 1 so that

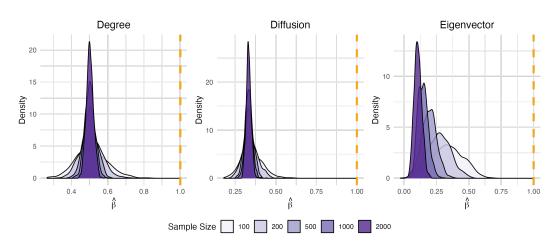
$$A_{ij} = \begin{cases} p_n & \text{if } i \neq j, \\ 0 & \text{otherwise.} \end{cases}$$

· Our regression model is:

$$Y_i = \beta C_i^{(d)} + \varepsilon_i^{(d)}$$

 $\varepsilon_i^{(d)} \overset{\text{i.i.d.}}{\sim} N(0,1) \text{ and } \varepsilon_i^{(d)} \perp \!\!\!\perp \hat{A}_{jk} \text{ for all } i,j,k \in [n].$ 

## Inconsistency with Proxy Networks



**Figure 1:** Distribution of  $\hat{\beta}^{(d)}$  for  $p_n = 1/n$ . For  $\hat{\beta}^{(\infty)}$ ,  $a_n = \sqrt{n}$ .  $\beta = 1$  (orange dashed line).

## **Consistency with True Networks**

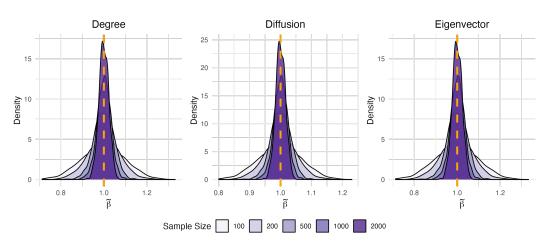


Figure 2: Distribution of  $\tilde{\beta}^{(d)}$  for  $p_n=1/n$ . For  $\tilde{\beta}^{(\infty)}$ ,  $a_n=\sqrt{n}$ .  $\beta=1$  (orange dashed line).

#### Bias correction is effective

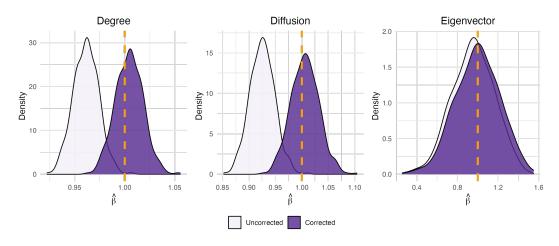
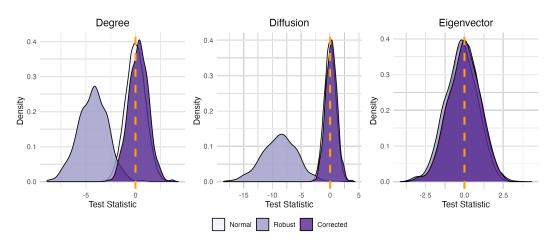


Figure 3: Distributions of  $\hat{\beta}^{(d)}$  and their bias corrected versions  $\check{\beta}^{(d)}$  for  $p_n = 1/\sqrt{n}$ , n = 500,  $a_n = \sqrt{\lambda_1(\hat{A})}$ .  $\beta = 1$  (orange dashed line).

## Distributional theory is accurate



**Figure 4:** Distribution of the centered and scaled test statistics. Robust refers to tests based on *t*-statistic with robust (hc) standard errors.  $p_n = 1/\sqrt{n}$ , n = 500,  $a_n = \sqrt{\lambda_1(\hat{A})}$ .

## Adjusted tests are better

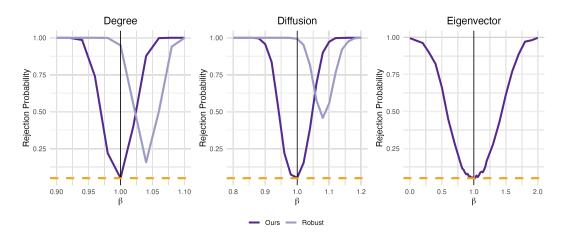
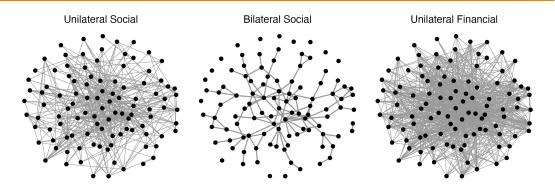


Figure 5: Power of the two-sided test of  $H_0$ :  $\beta=1$  under various alternatives. Test at 5% level of significance (orange dashed line).  $p_n=1/\sqrt{n}$ , n=500,  $a_n=\sqrt{\lambda_1(\hat{A})}$ .

- De Weerdt and Dercon (2006): want to know if informal insurance can help consumption smoothing
- · Village of 119 households in Nyakatoke, Tanzania
- · Regress variance in food expenditure on centrality in network



(n = 119)	Mean	Median	Min	Max
Unilateral Social	8.02	7	1	31
Bilateral Social	2.30	2	0	10
Unilateral Financial	16.53	14	3	79

**Table 3:** Degree distributions of various networks in Nyakatoke

		Estimate	p-value	Atten.	Bias Corr.
Unilateral Social	Degree	-1064	0.67	0.91	-1172
	Diffusion	-4274	0.77	1.00	-4290
	Eigenvector	-12353	0.86	0.91	-13548
Bilateral Social	Degree	-11604	0.06	0.74	-15592
	Diffusion	-23672	0.16	0.95	-24883
	Eigenvector	-10543	0.93	0.78	-13434
Unilateral Financial	Degree	-412	0.70	0.96	-429
	Diffusion	-4559	0.74	1.00	-4561
	Eigenvector	-15040	0.77	0.96	-15699

**Table 4:** Regression results. For diffusion,  $\delta=1/\sqrt{\lambda_1(\hat{A})}$ , T=2. For eigenvector,  $a_n=\sqrt{\lambda_1(\hat{A})}$ .



Conclusion

#### Conclusion '

- 1. Show that OLS can become inconsistent with sparse, proxy networks
  - Characterize threshold at which inconsistency occurs
  - · Show that eigenvector is less robust than degree and diffusion
  - · Comparing significance involves both economic and statistical properties
  - Rule-of-Thumb for sparsity regime
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  - Even when consistent, OLS estimators are asymptotically biased
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Regularized Eigenvectors

# Regularized Eigenvectors

• Suppose  $d = np_n$  is known. Define:

$$W_i := \min \left\{ \frac{2d}{C_i^{(1)}(\hat{A})}, 1 \right\}$$

- $w_i$  is the ratio by which the degree of i exceeds 2d.
- Let the regularized matrix  $\tilde{A}$  be defined as follows:

$$\tilde{A}_{ij} = \sqrt{w_i w_j} \, \cdot \, \hat{A}_{ij}$$

 $\cdot$   $\tilde{A}$  is the adjacency matrix in which we down-weight the links of high-degree agents so that degree is windsorized at 2d.

# Consistency with Proxy Networks

• Le et al. (2017) show that this regularized matrix concentrates to A in spectral norm even under sparsity.

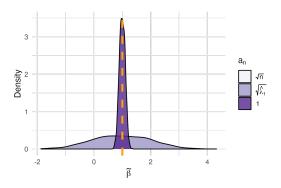
## Proposition 1 (Regularized Eigenvector)

Suppose  $a_n \to \infty$ . The linear regressions of Y on  $C^{(\infty)}(\tilde{A})$  is consistent if and only if

$$p_n\gg n^{-1}.$$

**Additional Simulations** 

## Scaling of Eigenvector Matters



**Figure 6:** Distribution of  $\tilde{\beta}^{(\infty)}$  for n=100,  $p_n=1/n$  under various  $a_n$ .  $\beta=1$  (orange dashed line).

# Eigenvector is more sensitive to sparsity than Degree and Diffusion

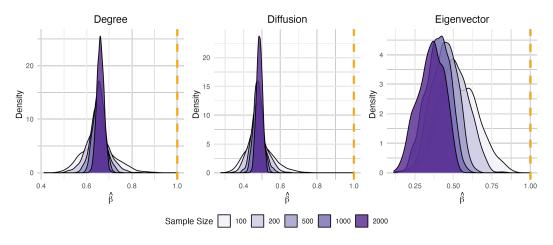


Figure 7: Distribution of  $\hat{\beta}^{(d)}$  for  $p_n = n^{-1} \sqrt{\log n / \log \log n}$ . For  $\tilde{\beta}^{(\infty)}$ ,  $a_n = \sqrt{n}$ .  $\beta = 1$  (orange dashed line).

# Size of $H_0: \beta = 1$

	Statistic		Sample Size					
<i>Pn</i>			100	200	500	1000	2000	
	Degree	Ours Robust	0.055 0.656	0.052 0.673	0.067 0.690	0.062 0.668	0.065 0.674	
0.1	Diffusion	Ours Robust	0.049 0.889	0.053 0.894	0.064 0.887	0.059 0.871	0.060 0.898	
	Eigenvector		0.045	0.043	0.037	0.056	0.044	
	Degree	Ours Robust	0.066 0.330	0.065 0.450	0.067 0.573	0.058 0.705	0.065 0.783	
$n^{-1/3}$	Diffusion	Ours Robust	0.080 0.645	0.070 0.734	0.074 0.813	0.057 0.888	0.064 0.934	
	Eigenvector		0.045	0.042	0.051	0.042	0.058	
$n^{-1/2}$	Degree	Ours Robust	0.072 0.659	0.049 0.801	0.051 0.949	0.037 0.993	0.062 0.999	
	Diffusion	Ours Robust	0.071 0.881	0.045 0.948	0.053 0.993	0.037 1.000	0.059 1.000	
	Eigenvector		0.077	0.045	0.050	0.050	0.047	

## Power of $H_0: \beta = 0$ under the alternative $H_1: \beta = 1$

	Charlania			Sample Size	ļ	
рn	Statistic	100	200	500	1000	2000
0.1	Degree - Robust	1.000	1.000	1.000	1.000	1.000
	Diffusion - Robust	1.000	1.000	1.000	1.000	1.000
	Eigenvector	0.845	0.995	1.000	1.000	1.000
$n^{-1/3}$	Degree - Robust	1.000	1.000	1.000	1.000	1.000
	Diffusion - Robust	1.000	1.000	1.000	1.000	1.000
	Eigenvector	0.998	1.000	1.000	1.000	1.000
n <sup>-1/2</sup>	Degree - Robust	1.000	1.000	1.000	1.000	1.000
	Diffusion - Robust	1.000	1.000	1.000	1.000	1.000
	Eigenvector	0.832	0.947	0.994	1.000	1.000

**Table 6:** Power of 5% level two-sided tests of  $H_0$ :  $\beta = 0$  when  $\beta = 1$ . Under this  $H_0$ , the our test statistics is the usual t-statistic with robust (heteroskedasticity-consistent) standard errors.

## **Local Asymptotics**

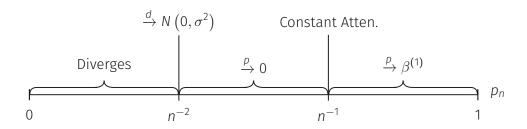
- · Similar in spirit to modeling:
  - Correlation of weak instruments and endogenous variables decaying to 0 (e.g. Staiger and Stock 1997).
  - Power of tests using local alternatives (Pitman drift, see e.g. Rothenberg 1984).
  - · Local to unity asymptotics for time series (e.g. Chan and Wei 1987)



## **Weak Ties Theory**

- Granovetter (1973): Weak ties which are more numerous are key drivers of outcomes
  - Weak ties: A<sub>ij</sub> is small
  - Numerous: most  $A_{ij}$  non-zero  $(O(n^2))$
- · Job referrals in Newton, MA:
  - Most recent job changers found jobs through friends "marginally included in the current network of contacts".
  - "It is remarkable that people receive crucial information from individuals whose very existence they have forgotten."
- Other examples: innovation (e.g. Reagans and Zuckerman 2001), economic development (e.g. Eagle et al. 2010), job referrals (e.g. Rajkumar et al. 2022).

# All Phase Transitions in $\hat{\beta}^{(1)}$





# Rule of Thumb for Consistency

#### Rule of Thumb

- (a) Treat  $\hat{\beta}^{(1)}$  and  $\hat{\beta}^{(T)}$  as consistent if there exists a giant component with at least N/2 nodes.
- (b) Treat  $\hat{\beta}^{(\infty)}$  as consistent if the network is fully connected.



## **Bias Correction**

· Bias correction reduces to the problem of

$$E\left[\iota'\left(\hat{A}-A\right)^{2}t\iota\mid U\right]$$

at a sufficiently fast rate.

- Given that  $E[(\hat{A}_{ij} A_{ij})^2 | U] \approx A_{ij}$ , an intuitive estimator is  $\iota' A^t \iota$
- · Consistent, but does not converge fast enough
- Difference between the two relates to the number of paths of a given length on a line in which each edge is traversed at least twice.



# Low Rank Assumption

## Assumption 1 (Rank R Graphon)

Suppose f has rank  $R < \infty$ :

$$f(u,v) = \sum_{r=1}^{R} \tilde{\lambda}_r \phi_r(u) \phi_r(v) \quad , \tag{3}$$

where  $\|\phi_r\| = 1$  for all  $r \in [R]$  and if  $r \neq s$ ,

$$\int_{[0,1]} \phi_r(u) \phi_s(u) du = 0 .$$

Furthermore, suppose that

$$\Delta_{\min} = \min_{1 \ge r \ge R-1} \left| \tilde{\lambda}_r - \tilde{\lambda}_{r+1} \right| > 0$$

## Low Rank Assumption

- The rank assumption means the networks have "structure" (Chatterjee 2015).
- Many popular network models are low rank
  - Stochastic Block Model (Holland et al. 1983)
  - · Random Dot Product Graphs (Young and Scheinerman 2007)
- Also common in the matrix completion literature (e.g. Candès and Tao 2010, Negahban and Wainwright 2012, Athey et al. 2021).

# **Empirical Demonstration**

	90%	95%	99%
Degree	(-19500, ∞) (-18800, ∞)		
Diffusion	$(-45000, \infty)$ $(-25200, \infty)$		

**Table 7:** One-sided confidence intervals for degree and diffusion in Bilateral Social network.

