

STAT 110 Practice Examination

Annotated Solutions

1. The sample mean \bar{y} is closest to
 - A. 41.3
Just one observation, not the average of all five.
 - B. 39.7
Again only a single value.
 - C. **42.0**
 - D. 52.5
Mean can't be larger than every data point.
 - E. 209.9
That is the sum of the numbers, not their mean.
2. The sample standard deviation s is closest to
 - A. 1.0
Would imply the five numbers are almost identical.
 - B. 8.2
Needs values about 8 away from the mean—none are.
 - C. 94.0
Mixes up variance and standard deviation.
 - D. **2.9**
 - E. 1.7
Used n in the denominator instead of $n - 1$; too small.
3. Watching one T20 cricket match that totals 325 runs, choose the best statement:
 - A. It is unlikely totals ever exceed 500.
One match can't tell us that.
 - B. 325 is “typical” because the game was random.
Need a league-wide distribution.
 - C. **We can't tell if 325 is unusual without knowing likely score ranges.**
 - D. More than 300 runs is normal.
No supporting evidence.

- E. Every game has exactly 325 runs.
Plainly impossible.
4. Are variation and uncertainty important in statistics?
- A. **Yes—models describe variation and CIs/SEs quantify uncertainty.**
- B. No; parameters are known exactly.
Population parameters are unknown.
- C. Yes; it makes analysis look sophisticated.
Not a scientific reason.
- D. No; big data eliminate uncertainty.
Large n reduces but never removes it.
- E. No; once data are collected there's no uncertainty.
Randomness persists.
5. $\Pr(Y = 3 \text{ or } 4)$ for the gene-copy study
- A. 0.30
Omits the 4-copy probability.
- B. **0.80**
- C. 0.20
Only the missing 0.20 probability.
- D. 0.50
Counts 3-copies but not 4-copies.
- E. 1.00
Would mean everyone has 3 or 4 copies.
6. $E[Y]$
- A. 2.25
Dropped one probability weight.
- B. **2.85**
- C. 2.50
Used wrong weights.
- D. 3.00
Added instead of weighting.
- E. 2.80
Rounded from a wrong calculation.
7. Best description of $E[Y]$
- A. A person will have exactly $E[Y]$ copies.
It's a long-run average, not a guarantee.
- B. **It's the expected (average) number of copies for one randomly chosen person.**

- C. Half the people have fewer than $E[Y]$.
That defines the median.
- D. Exactly $E[Y]$ copies means higher risk.
Mean alone says nothing about risk.
- E. Copy number is normal with mean $E[Y]$.
Normality wasn't given.
8. Meaning of $\Pr(B \mid V^c)$
- A. **Probability a customer chooses a plant-based burger *given they are not vegetarian*.**
- B. Probability a customer is not vegetarian, given they chose the burger.
Condition reversed.
- C. Probability a vegetarian chooses the burger.
Condition is the opposite group.
- D. Joint probability of non-vegetarian & burger.
Not conditional.
- E. Probability of choosing a burger regardless of diet.
Ignores the condition.
9. $\Pr(V \mid B)$
- A. **0.632**
- B. 0.600
That's $\Pr(B \mid V)$.
- C. 0.180
Only one piece of Bayes' rule.
- D. 0.285
Another single piece.
- E. 0.105
Yet another single piece.
10. $\Pr(V \cup B)$
- A. **0.405**
- B. 0.180
Just vegetarians who choose burgers.
- C. 0.300
 $\Pr(V)$ only.
- D. 0.285
 $\Pr(B)$ only.
- E. 1.000
Would imply every customer is vegetarian or buys a burger.
11. $\Pr(B)$

- A. 0.180
Only vegetarians choosing burgers.
 - B. 0.300
 $\Pr(V)$, not $\Pr(B)$.
 - C. **0.285**
 - D. 0.105
Only non-vegetarians choosing burgers.
 - E. 1.000
Would mean everyone buys a burger.
12. Best definition of a random variable
- A. Summarises both population and sample.
It's defined before any sample is taken.
 - B. Normally distributed variable.
Distribution can be anything.
 - C. **A random process with a numerical outcome.**
 - D. Fixed but unknown value.
The value is random before observing.
 - E. Depends on the observed sample.
Definition doesn't rely on realised data.
13. Which is best modelled as continuous?
- A. Eggs in a nest.
Count data, discrete.
 - B. Purchasing visitors.
Count data, discrete.
 - C. **pH value of seawater**
 - D. Voters supporting a candidate.
Discrete count.
 - E. Tasks completed.
Discrete count.
14. $E[2X - 3Y]$ with $E[X] = 10$, $E[Y] = 4$
- A. 4
Used $X - 3Y$ instead of $2X - 3Y$.
 - B. 20
Multiplied rather than added/subtracted.
 - C. 12
Dropped one coefficient.
 - D. **8**

- E. 0
Incorrectly cancelled the terms.
15. $\text{sd}(2X - 3Y)$ (independent X, Y)
- A. 10.50
That's the variance, not the s.d.
- B. 1.22
Square-rooted the wrong total.
- C. **3.24**
- D. 2.12
Left out a variance component.
- E. 1.50
Left out two variance components.
16. A healthy adult is 1.5 s.d. below the mean ($\mu = 40, \sigma = 8$). Score?
- A. -1.5
z-score itself, not the raw score.
- B. 38.5
Moved in the wrong direction.
- C. 52.0
Moved above the mean.
- D. **28.0**
- E. 40.0
That's the mean ($z = 0$).
17. Code to find $P(\text{score} > 48)$
- A. **1-pnorm(1.0)**
- B. 1-pnorm(48)
Treats raw score as a z-score.
- C. 1-pnorm(-1.0)
Upper tail of the wrong z.
- D. pnorm(1.0)
Gives lower tail.
- E. pnorm(48)
Again uses raw score as z.
18. Sampling distribution of \bar{y} for $n = 64$
- A. Normal(40, 8)
Forgot to divide by $\sqrt{64}$.
- B. **Normal(40, 1)**

- C. Normal(0, 1)
Standardised already.
- D. t_{63}
 σ is known, so we use z .
- E. Unknown.
Central Limit Theorem gives it.
19. Same set-up as 18 but population not normal (still $n = 64$)
- A. Normal(40, 8)
Same issue—must divide by \sqrt{n} .
- B. **Approximately normal**(40, 1)
- C. Normal(0, 1)
Already standardised.
- D. t_{63}
 σ known.
- E. Unknown.
CLT still applies for large n .
20. Which formula gives a 95 % confidence interval for a *population mean* when the population standard deviation is *unknown*?
- A. $\bar{y} \pm z^* \sqrt{\frac{p(1-p)}{n}}\%$
This is the CI for a proportion, not a mean.
- B. $\bar{y} \pm z^* \frac{\sigma}{\sqrt{n}}\%$
Uses the unknown σ instead of the sample s .
- C. $\bar{y} \pm t_{n-1}^* \frac{\sigma}{\sqrt{n}}\%$
 σ is unknown; we should plug in s , not σ .
- D. $\bar{y} \pm z^* s\%$
Forgot to divide by \sqrt{n} and uses z instead of t .
- E. $\bar{y} \pm t_{n-1}^* \frac{s}{\sqrt{n}}$
21. If we increase the confidence level from 90 % to 99 % while keeping the same sample,
- A. The point estimate changes.
The sample mean stays the same.
- B. The standard error changes.
SE depends only on s and n .
- C. The sample size must change.
Sample size is fixed.
- D. **The interval widens to guarantee the higher confidence.**

- E. A higher confidence level always narrows the interval.
It actually widens.
22. A 99 % confidence interval for the mean content of coffee jars is (201.2 g, 203.0 g). Which interpretation is correct?
- A. 99 % of individual jars weigh between 201.2 g and 203.0 g.
CI is about the mean, not single jars.
- B. 99 % of *future* jars will fall in this range.
Again, CI targets the mean.
- C. There is a 99 % chance the true mean lies in this very interval.
The probability refers to the method, not to μ once the interval is drawn.
- D. If we took many samples, 99 % of the *sample means* would be in this range.
CI's vary from sample to sample; means do not cluster in one fixed interval.
- E. **In the long run, 99 % of such intervals built this way would contain the true mean μ .**
23. Which statement about a p -value is correct?
- A. A large p -value proves the null hypothesis is true.
"Prove" is too strong.
- B. $p = 0.03$ means there is a 3 % chance the alternative is false.
 p -value is calculated under H_0 , not H_A .
- C. **$p = 0.03$ means that, if H_0 were true, we would see data this extreme only 3 % of the time.**
- D. p -value is the probability calculations were wrong.
It measures data extremeness, not mistakes.
- E. A small p -value guarantees practical importance.
Statistical \neq practical significance.
24. Which step is genuine *statistical estimation* of a population mean?
- A. Just replace μ by \bar{y} in a sentence.
Pure description, no inference.
- B. Compute the standard error only.
SE alone isn't an estimate of μ .
- C. Use the full population data.
Then it's no longer estimation.
- D. Guess a number that "looks right."
Not statistical.
- E. **Build a confidence interval or carry out a test based on the sample.**
25. In the penguin data, the first line of R code calculates `mean(bill_length_mm)`. What does this value summarise?

- A. **The average bill length for the sampled penguins.**
 - B. The average flipper length.
Wrong variable.
 - C. The standard deviation of bill length.
Mean, not spread.
 - D. The median bill length.
Mean \neq median.
 - E. The sample size.
Mean is not a count.
26. The next line of code asks for `sd(flipper_length_mm)`. What statistic is that?
- A. Mean flipper length.
It's the spread, not the centre.
 - B. Median flipper length.
Again, centre vs spread.
 - C. Range of flipper length.
Range is max-min.
 - D. **The standard deviation of flipper length (how much they vary).**
 - E. Sample size.
Standard deviation isn't a count.
27. A one-sample t -test compares pollution *before* and *after* a policy. Which null hypothesis is correct?
- A. $\mu_{\text{after}} - \mu_{\text{before}} = 0$
 - B. $\mu_{\text{after}} = 0$
Should compare difference, not absolute level.
 - C. A small p -value means results are “not unusual.”
It actually means they are unusual under H_0 .
 - D. The test measures correlation between before & after.
It tests mean difference.
 - E. Significance alone decides policy.
Statistical vs practical issues.
28. You have paired “before/after” data. Which test is appropriate?
- A. **Paired-sample t -test on the differences.**
 - B. Two-sample independent t -test.
Ignores pairing.
 - C. Two-proportion z -test.
Data are numeric, not proportions.
 - D. Chi-square test.
Counts, not means.

- E. ANOVA.
More than two groups required.
29. Statistical *power* is
- A. The chosen α level.
That's Type I error rate.
 - B. Probability of rejecting H_0 when H_0 is true.
That's Type I error again.
 - C. **Probability of rejecting H_0 when a specified alternative is true.**
 - D. The observed p -value.
p-value and power differ.
 - E. Guaranteed by $p \leq 0.05$.
Low p doesn't ensure high power.
30. Which change *increases* the power of a test (all else fixed)?
- A. Call the data “paired” when they are not.
Inflates Type I error, not power properly.
 - B. Reduce the sample size.
Smaller n lowers power.
 - C. Lower α from 0.05 to 0.01.
Harder to reject, so power drops.
 - D. Replicate the same small study many times with no changes.
Each run keeps low power.
 - E. **Increase the sample size.**
31. Standard error for the difference of two independent means: which value is correct?
($s_1 = 4.5$, $n_1 = 25$; $s_2 = 6.0$, $n_2 = 36$)
- A. **2.11**
 - B. 4.44
Forgot to divide by sample sizes.
 - C. 0.61
Only used one group's variance.
 - D. 2.92
Miscalculated the square-root step.
 - E. 16.32
Added variances, then squared again.
32. Which R command correctly performs an *independent* two-sample t -test?
- A. `t.test(groupA, groupB, var.equal = FALSE)`
 - B. `t.test(groupA, groupB, paired = TRUE)`
Sets a paired test you don't have.

- C. `chisq.test(groupA, groupB)`
Chi-square uses counts.
 - D. `prop.test(groupA, groupB)`
Proportions, not means.
 - E. `wilcox.test(groupA, groupB)`
Non-parametric alternative.
33. Interpreting a 95 % CI for the difference (drug A minus placebo) of $(-4.3, 1.2)$:
- A. CI gives the mean reduction for drug A alone.
It describes the difference.
 - B. Drug A is definitely better.
CI includes zero—no guarantee.
 - C. A single patient will improve by 1.2 units.
CI is about the mean, not an individual.
 - D. The CI proves no reduction is possible.
Includes zero but doesn't prove equality.
 - E. **Because the CI crosses 0, we lack strong evidence that drug A outperforms placebo.**
34. A regression-output question asked, “Which statement is *not* correct?” Only one option is wrong.
- A. The errors are normally distributed.
This is a valid model assumption.
 - B. Homoscedasticity was assessed.
Also true.
 - C. Linearity appears reasonable.
Also true.
 - D. **There is no random error term in the fitted model.**
Every regression has ε .
 - E. Independence of observations was checked.
Also true.
35. With $\hat{y} = 150 - 0.75x$, the fitted mean at $x = 20$ is
- A. **135**
 - B. 148.5
Added instead of subtracted 0.75×20 .
 - C. 135 for every bird.
Mean model; individuals vary around it.
 - D. 150
Forgot the slope term.
 - E. The model gives no spread.
Residual SD gives the spread.

36. Interpret the slope $\hat{\beta}_1 = -0.75$ in $\hat{y} = 150 - 0.75x$.
- A. **On average, y decreases by 0.75 units for each 1-unit increase in x .**
 - B. The line rises by 0.75 units per x .
Sign is negative.
 - C. One bird's y always drops 0.75.
Slope describes the mean, not individuals.
 - D. When x changes 10 units, y always drops 7.5.
Again, average not guarantee.
 - E. -0.75 is the intercept.
Intercept is 150.
37. A scatterplot gives $r = -0.64$ between age and aptitude.
- A. **The variables have a moderate negative linear relationship.**
 - B. Positive correlation.
The sign is negative.
 - C. No linear trend.
 $|r|$ is well away from 0.
 - D. Relationship is curved.
Nothing suggests curvature.
 - E. Outliers dominate.
No evidence shown.
38. Least-squares line-fitting uses which criterion?
- A. Draw a line by eye.
Subjective, not least squares.
 - B. Minimise the *sum* of residuals.
They cancel to 0 anyway.
 - C. Maximise correlation.
Not the fitting rule.
 - D. **Minimise the sum of *squared* residuals.**
 - E. Find a perfect fit.
Rarely possible with real data.
39. Which equation gives the *predicted* aptitude from age?
- A. $\hat{x} = 150 - 0.75y$
Swapped variables.
 - B. $\hat{y} = 150 - 0.75x + \varepsilon$
 ε is not part of the prediction.
 - C. $\hat{y} = 150 - 0.75x$
 - D. $y = 150 - 0.75x$
Drops the "hat"—not clearly prediction.

- E. $x = 150 - 0.75y$
Predicts age from aptitude.
40. The regression output gives $p = 0.00177$ for the slope. Which option matches this?
- A. 0.020
Too large.
- B. 0.15
Too large.
- C. 0.051
Too large.
- D. 0.50
Far too large.
- E. **0.0018 (rounded).**
41. The standard error for $\hat{\beta}_1$ is 0.310. Which description is correct?
- A. The standard error tells us the average error made when predicting \hat{y} %
That is the residual standard error, not $SE(\hat{\beta}_1)$.
- B. It is the mean of the sampling distribution of $\hat{\beta}_1$ %
The mean of that distribution is the true β_1 , not the SE.
- C. It measures variability of the y -values around the regression line%
That quantity is σ_ε , not $SE(\hat{\beta}_1)$.
- D. It is the difference between observed and true slope%
That difference is one realisation, not the spread.
- E. **It describes the variability in the sampling distribution of $\hat{\beta}_1$.**
42. With multiplier 2.093, the 95 % CI for β_1 is closest to
- A. $(-1.44, -0.82)$
Half-width too small (used wrong SE).
- B. $(99.26, 120.48)$
Uses the intercept scale.
- C. $(-1.78, -0.48)$
- D. $(-11.74, 9.48)$
Uses residual SD, not slope SE.
- E. $(-2.43, 0.17)$
Wrong half-width and sign range includes 0.
43. Should we interpret $\hat{\beta}_0$ in this study?
- A. Yes; it gives aptitude when age = 0 months.
Age = 0 never occurs and is outside data range.
- B. Yes; intercept is always meaningful.
Not if $x = 0$ is implausible.

- C. No; R^2 is too low.
 R^2 doesn't govern interpretability of β_0 .
- D. No; SE is large.
Uncertainty \neq scientific meaning.
- E. **No; age = 0 is impossible here, so the intercept lacks scientific meaning.**
44. Two intervals from `predict(..., interval = "|A|")` and `"|B|"`:
- A. Interval A: CI for mean, Interval B: PI
Their widths are opposite.
- B. Both are prediction intervals.
One is clearly narrower.
- C. Interval A for low aptitude, B for high.
Aptitude not in code.
- D. **Interval A is a 95 % *prediction* interval; Interval B a 95 % CI for the mean response.**
- E. Interval B is narrower due to bigger n .
Same n ; difference is PI vs CI.
45. Predicted aptitude for a child who first speaks at 60 months:
- A. 177
Plugged 60 into a wrong formula.
- B. 6591
Multiplied instead of subtracting.
- C. 104
Used +1.13 rather than -1.13.
- D. **42**
- E. 60
Simply echoed the age.
46. Can the prediction in Question 45 be trusted?
- A. Yes; $R^2 > 0.3$ guarantees reliability.
 R^2 alone never guarantees it.
- B. Yes; residual SD is small.
Still outside data range.
- C. No; negative correlation forbids prediction.
Sign isn't the issue.
- D. **No; age = 60 is an extrapolation beyond the observed 7–42 months.**
- E. No; linear models are never useful.
Too extreme.
47. Multiple linear regression allows us to

- A. Have at most two predictors.
Any number is allowed.
 - B. Exclude categorical variables.
They enter via dummies.
 - C. Abandon `lm` for `mlm`.
lm fits multiple predictors.
 - D. **Assess each predictor's effect while *controlling* for the others.**
 - E. Interpret only one predictor's effect.
Each coefficient has meaning.
48. Interpreting $\hat{\beta}_0$ from the model with temperature only
- A. **Temperature = 0 °C lies within the data (-4.5 – 24.8 °C), so $\hat{\beta}_0 = 1873.9$ is the estimated mean metabolic rate at 0 °C.**
 - B. We mustn't interpret it; 0 °C is biologically impossible.
0 °C is observed.
 - C. 0 °C is outside data – so extrapolation.
It's inside.
 - D. Intercept gives change per degree.
Slope does that.
 - E. Intercepts are never interpreted.
Context determines interpretability.
49. From the model with temperature + activity, which pair is $(\hat{\beta}_2, s_{\hat{\beta}_2})$?
- A. 1623.69, 19.38
That's the intercept.
 - B. **52.64, 2.95**
 - C. -18.63, 1.05
Those are for temperature.
 - D. 52.64, 19.38
SE mismatched to estimate.
 - E. -18.63, 2.95
Estimate and SE from different predictors.
50. Best wording of the 95 % CI for β_2 (activity)
- A. 46.8–58.5 change; temp ignored.
Must say "holding temperature fixed."
 - B. 46.8–58.5 *level* at 1 h.
CI is for change.
 - C. -20.7—16.6 for temperature.
That's the other predictor.

- D. We are 95 % confident the *mean* metabolic rate increases by 46.8–58.5 Kcal/day for each extra hour of activity, *holding temperature constant*.
- E. Mixed effects of both predictors.
CI isolates activity's slope.
51. $R^2 = 84.8\%$. Which statement is *not* correct?
- A. R^2 is the squared correlation between y and \hat{y} .
True.
- B. R^2 ranges from 0 to 1.
True.
- C. R^2 is the proportion of variance explained.
True.
- D. The label “Multiple R-squared” shows R^2 in R output.
True.
- E. **A model is useless unless $R^2 > 0.5$.**
52. Residual-plot shows a funnel shape. Which assumption fails?
- A. Linearity
Trend looks straight.
- B. Independence
Plot gives no info on independence.
- C. Outliers
None obvious.
- D. **Constant variance (homoscedasticity).**
- E. No assumptions violated
We see heteroscedasticity.
53. ANOVA: hypotheses being tested
- A. Difference of two means.
Only two groups.
- B. Test of independence.
Different procedure.
- C. All four means unequal.
Too strict.
- D. $H_0 : \mu_1 = \mu_2 = \mu_3 = \mu_4$; H_A : **at least one mean differs.**
- E. Equality of full distributions.
Focus is on means.
54. F-value for that ANOVA
- A. **27890**

- B. 355 921
That is SS_{type} .
- C. 118 640
That is MS_{type} .
- D. 4
Residual MS .
- E. 14 455
 SS_{res}

55. Which statement is *not* correct about ANOVA?

- A. ANOVA compares between-group to within-group variance.
Correct.
- B. Under H_0 , the F-value follows an F-distribution.
Correct.
- C. Large F suggests group means explain much variation.
Correct.
- D. df in the table choose the reference F-distribution.
Correct.
- E. **The ANOVA table itself tells us which group means differ.**

56. Which reference distribution gives the p-value?

- A. Standard normal.
Wrong family.
- B. χ^2_3 .
Not used here.
- C. $F_{(3, 3398)}$
- D. χ^2_{3398}
Wrong.
- E. t_{3398}
Wrong family.

57. Interpreting the tiny p-value from ANOVA ($\alpha = 0.01$)

- A. **Because $p \leq \alpha$, the data are very unlikely if all pitch-type means were equal.**
- B. $p \leq \alpha$ proves slider = change-up speed.
ANOVA is omnibus.
- C. All four means differ.
Post-hoc needed to know.
- D. Shows Kershaw is “hard to face.”
Beyond the scope of speed.

- E. Identifies slowest and fastest pitch.
Need pairwise tests.
58. TukeyHSD(a.baseball) does what?
- A. Fits separate regressions.
No, it's post-hoc on the ANOVA.
- B. Tests equality of variances.
Levene/Bartlett do that.
- C. **Compares every pair of pitch-type means with a multiple-comparison adjustment.**
- D. Gives an updated ANOVA table.
Already have one.
- E. Tests each pitch vs overall mean.
Not Tukey's HSD.
59. Interpreting the row FF-CU in Tukey output
- A. More fastballs than curveballs per 100 pitches.
Counts, not speeds.
- B. Faster than “non-fastballs.”
CU only.
- C. More strikes per 100 pitches.
Not strike data.
- D. **Fastballs average 30.04 – 30.60 km/h *faster* than curveballs (95 % confidence).**
- E. 95 % probability statement.
Confidence, not probability about a fixed quantity.
60. Which is *not* an example of binary data?
- A. **The *number* of eggs in a nest**
Takes many integer values – a count, not yes/no.
- B. Presence / absence of a gene.
Two outcomes.
- C. Dolphin breeding status (breeder / non-breeder).
Two outcomes.
- D. Task completed within time? (yes / no)
Two outcomes.
- E. Penguin age status (adult / juvenile).
Two outcomes.
61. Which *binomial* assumption is *violated* in this bird-nest study?
- A. Two possible outcomes (success / failure) for each trial hold.
This one is satisfied.

- B. The probability of success p is the same for every trial.
Also satisfied.
- C. Successive trials are independent.
No evidence of dependence was given.
- D. Each trial outcome is recorded accurately.
No hint of mis-classification.
- E. **The number of trials n is fixed in advance.**
Here the count of nests keeps growing, so n is not predetermined.
62. What proportion of visits (6 430 of 7 729) resulted in a sale?
- A. **0.83**
- B. 0.53
6 430 \div 12 200 was used—wrong denominator.
- C. 0.94
Divided 7 729 by 8 200.
- D. 0.75
Rounded from 5 800 \div 7 729, which is not the right count.
- E. 0.78
Used an intermediate rounded figure, not the exact ratio.
63. What fraction of tagged birds later returned? (419 of 791)
- A. 0.33
419 \div 1 270—wrong denominator.
- B. **0.53**
- C. 0.91
791 \div 870—swapped numerator and denominator.
- D. 0.41
327 \div 791—wrong numerator.
- E. 0.80
Rounded 634 \div 791—again wrong numerator.
64. In R, `prop.test(x1, x2)` reports a 95 % CI labelled “ $p_1 - p_2$ ”. What parameter does that interval estimate?
- A. p_1
Only the first group’s proportion.
- B. p_2
Only the second group’s proportion.
- C. $p_1 - p_2$ (the **difference** between the two proportions).
- D. $\frac{p_1}{p_2}$
That would be a risk ratio.

- E. The pooled overall proportion.
Not what the CI targets.
65. By default, `prop.test` for a 2×2 table tests the null hypothesis
- A. $p_1 = 0.5$
Only one group.
 - B. $p_2 = 0.5$
Only one group.
 - C. $p_1 + p_2 = 1$
Adds instead of compares.
 - D. $p_1 = p_2$ (equivalently, $p_1 - p_2 = 0$).
 - E. $p_1 p_2 = 0$
Product is irrelevant here.
66. For a 2×2 table, `prop.test` and `chisq.test` (with Yates correction) will always give
- A. **The same χ^2 statistic and therefore the same p -value.**
 - B. Different χ^2 but same p -value.
Statistic itself is identical.
 - C. Same statistic but a different reference distribution.
Both use χ^2_1 .
 - D. An F-test in one case.
Neither uses the F-family here.
 - E. Completely unrelated results.
They address the same hypothesis.
67. Expected count (cancer cases, “light smokers” cell) is calculated as (row total \times column total)/grand total. Which option matches that formula?
- A. Adds row to column totals.
Should multiply.
 - B. **Multiplies the row total by the column total, then divides by the grand total.**
 - C. Uses the larger of the two totals.
Not the expected-count rule.
 - D. Divides column by row total.
Reverses the formula.
 - E. Grand total divided by both subtotals.
Upside-down.
68. Using that rule, the expected number of cancer cases in the “never-smoked” group is approximately
- A. 4.6
Halved the correct value.

- B. 18.4
Doubled the correct value.
- C. 12.0
Rounded a mid-step mistakenly.
- D. **9.2**
- E. 25.0
Used column total instead of grand total.
69. The χ^2 test on that 4×2 table checks
- A. Whether the four row totals equal the two column totals.
Totals are fixed.
- B. Goodness-of-fit against a theoretical Poisson.
Wrong framework.
- C. **Independence between smoking status (4 levels) and cancer (yes/no).**
- D. Equal variances across rows.
Different concept.
- E. Equality of four separate proportions at once.
Independence phrasing is preferred.
70. The degrees of freedom for a 4×2 contingency table are
- $$(r - 1)(c - 1) = (4 - 1)(2 - 1) =$$
- A. 1
Forgot $(r - 1)$.
- B. 2
Computed $(c - 1)$ only.
- C. **3**
- D. 5
Added instead of multiplied.
- E. 6
Used $r \times (c - 1)$.
71. The test yielded $p \approx 0.30$. At $\alpha = 0.05$ we therefore
- A. Declare smoking *causes* cancer.
Causality not established and $p > 0.05$.
- B. Prove H_0 is true.
Failing to reject \neq proof.
- C. Need a bigger sample: $n < 500$.
 $n > 6\,000$ already.
- D. **Fail to reject H_0 ; evidence of association is weak.**

- E. Automatically switch to Fisher's exact test.
Expected counts are large enough.
72. What does the Mann–Whitney (Wilcoxon rank-sum) test compare?
- A. Equality of *means*.
It compares distributions via ranks.
 - B. Variances of two groups.
That's Levene's/Bartlett's test.
 - C. **Whether one distribution tends to give higher (or lower) values than the other.**
 - D. Medians *only*.
Shift in distribution, not just the median.
 - E. Counts of observations above 0.
Not its statistic.
73. Its test statistic is the
- A. Sum of the *signed* residuals.
Signed ranks belong to the Wilcoxon signed-rank test for paired data.
 - B. Mean of raw values in group A.
Raw data not used directly.
 - C. Variance of the pooled sample.
Not relevant.
 - D. Largest observation's rank.
Uses all ranks, not just one.
 - E. **Sum of the ranks for one of the two groups.**
74. Which statement about a non-parametric rank test is *true*?
- A. They are always *more* powerful than *t*-tests.
Often less powerful.
 - B. They provide easy CIs for a difference of means.
Mean-based CIs are awkward.
 - C. **They require no normality assumption for the underlying data.**
 - D. Significance implies causation.
Design, not p-values, yields causality.
 - E. Their p-value is guaranteed to be larger than in a parametric test.
Could be smaller or larger.
75. A p-value of 0.004 from a Mann–Whitney test means
- A. The average of the ranks equals the average of raw data.
Ranks are not raw values.
 - B. 0.4 % probability the null is true.
 $p \neq \Pr(H_0 \text{ true})$.

- C. If the two distributions were identical, we would see a rank-sum this extreme only 0.4 % of the time.
 - D. There is no evidence of a difference.
 $p = 0.004$ is strong evidence against H_0 .
 - E. Effect size must be large.
Significance \neq effect magnitude.
76. A stratified random sample
- A. Might ignore one stratum entirely.
By definition, each stratum is sampled.
 - B. Picks exactly the same # units from each stratum.
Proportional or unequal allocation is common.
 - C. **Draws separate random subsamples within every stratum.**
 - D. Always uses systematic sampling inside strata.
Can be simple random or other methods.
 - E. Guarantees lower variance than an SRS.
Often but not always.
77. In the cannabis-use survey, the biggest threat to validity was
- A. Selection bias from non-response.
High response rate reported.
 - B. Confounding by age.
Age was adjusted for.
 - C. **Information (reporting) bias: people may under-report cannabis use.**
 - D. Instrument calibration error.
Self-report, not instruments.
 - E. Over-powered sample inflating small effects.
Not the main concern.
78. The MAO-A gene paper comparing Māori and non-Māori was criticised mainly because
- A. It involved no ethics approval.
Approval was obtained.
 - B. **Researchers failed to involve Māori sufficiently in study design and interpretation.**
 - C. The statistical model assumed independence.
Standard assumption, not unique flaw.
 - D. Results were not peer-reviewed.
They were published in a journal.
 - E. The gene was sequenced inaccurately.
Genotyping method wasn't the issue.
79. Which practice *violates* the principle of genuine co-design with communities?

- A. Researchers present preliminary ideas at a community hui.
Involves community.
 - B. Community representatives sit on the advisory board.
Also involvement.
 - C. Community veto over data sharing is respected.
Aligns with co-design.
 - D. **Investigators finalise aims and methods *before* any community consultation.**
 - E. Draft results are fed back for comment.
Again involves community.
80. In the CARE principles for Indigenous data, the “C” stands for
- A. **Collective benefit**
 - B. Consent
Important, but not the “C” in CARE.
 - C. Custodianship
Covered by the “A” and “R” elements.
 - D. Confidentiality
Not the first pillar here.
 - E. Collaboration
Embedded in the overall framework, but not the C-word.
81. In a clinical trial, what best describes a *placebo*?
- A. Some control groups receive no treatment at all.
“No-treatment” ≠ placebo if nothing is given.
 - B. A placebo pill lets participants know they are in the control arm.
Blinding means they shouldn’t know.
 - C. **An inert treatment that looks the same as the active drug, helping blind participants and researchers.**
 - D. Placebos are unethical because they deny therapy.
Ethics boards permit them when no proven therapy exists.
 - E. Placebos make the study single-blind only.
They enable single or double blinding.
82. Which variable qualifies as a *confounder* in an observational study?
- A. Related only to the *predictor* but not the outcome.
Must relate to both.
 - B. Related only to the *outcome* but not the predictor.
Must relate to both.
 - C. Completely independent of both predictor and outcome.
Cannot confound.

- D. **Associated with both predictor and outcome, potentially distorting the observed relationship.**
 - E. Any variable measured after the outcome.
Timing is wrong for confounding.
83. The “replication crisis” in science refers to
- A. Journals publishing too *few* papers.
The issue is about reproducibility, not volume.
 - B. Choosing Bayesian methods over frequentist ones.
Both paradigms face replication issues.
 - C. **Many published findings failing to reproduce when the studies are repeated independently.**
 - D. A shortage of funding for statistics.
Not the core of the crisis.
 - E. Peer reviewers demanding larger samples.
Better power aids replication.
84. Testing dozens of hypotheses on one data set without adjustment mainly
- A. Controls the Type I error at 5 % overall.
It inflates the family-wise error.
 - B. Leaves the false-positive rate unchanged.
Actual rate increases.
 - C. **Inflates the chance of at least one false positive (Type I error).**
 - D. Guarantees smaller p -values are true discoveries.
They may be false positives.
 - E. Eliminates the need for follow-up studies.
Replication is still needed.
85. “HARKing” (Hypothesising After Results are Known) is the practice of
- A. **Formulating or changing hypotheses *after* seeing the data, then presenting them as if specified beforehand.**
 - B. Pre-registering study aims.
That is the opposite practice.
 - C. Randomising subjects after consent.
Unrelated to hypotheses.
 - D. Adding more predictors during peer review.
Could be data dredging, but not the definition of HARKing.
 - E. Publishing only significant outcomes.
That is publication bias.
86. Which statement about *maximum likelihood estimation* (MLE) is correct?

- A. MLE requires a Bayesian prior.
MLE is purely likelihood based.
- B. MLE only works for very large samples.
It is defined for any sample size.
- C. Computes p -values directly.
MLE gives point estimates; tests use additional steps.
- D. **MLE chooses parameter values that maximise the data's likelihood under the assumed model.**
- E. Seldom used in real research.
It is standard in many fields.

87. In Bayesian analysis, which statement is *false*?

- A. Bayesian inference combines prior information with the data likelihood.
Core principle—true.
- B. The posterior distribution expresses updated beliefs about parameters.
True.
- C. **Bayesian methods ignore the likelihood and rely only on the prior.**
- D. Bayesian credible intervals are probability statements about parameters.
True.
- E. Bayesian methods are widely applied in modern statistics.
True.