Switching Linear Dynamical Systems

Information Theory and Inference



Intuition

Switching state space models

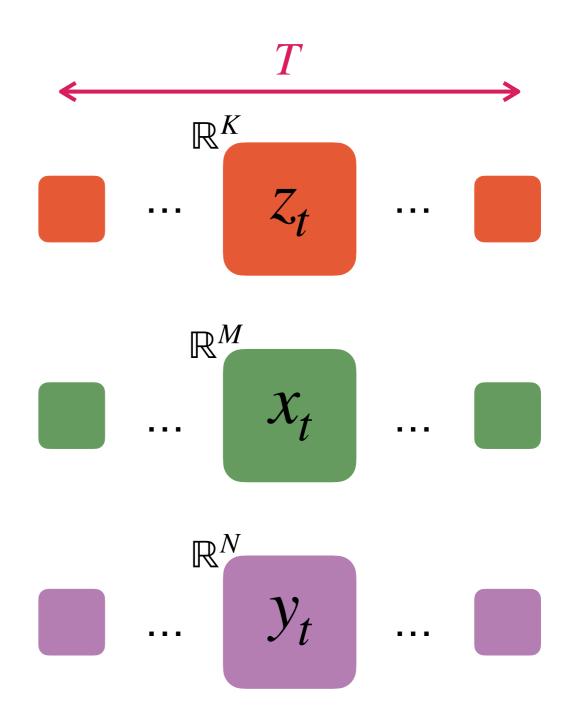
Imagine driving a vehicle, its macroscopic motion can be quite difficult to understand but it can be seen as the consequence of four basic actions performed by the driver (accelerate, brake, turn left, turn right)

Switching state space models provide a framework to understand and predict such dynamic behaviors by allowing for transitions between different underlying states in systems that can switch between different regimes over time

The layers

The model is composed of: K discrete latent states M continuous latent states N visible states

Finally, the time step is denoted with t = 1,...,T



The dynamics

The latent discrete dynamic is Markovian

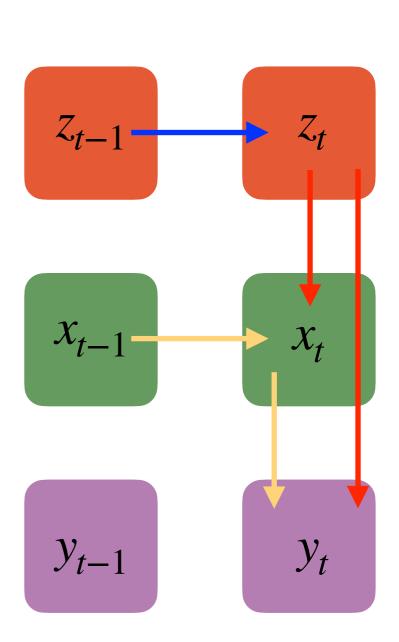
$$z_t \sim \pi_{z_{t-1}} \in \mathbb{R}^{K \times K}$$

And it is propagated through the other layers as

$$x_{t} = A_{z_{t}} x_{t-1} + b_{z_{t}} + v_{t}$$

$$y_t = C_{z_t} x_t + d_{z_t} + w_t$$

Where ν_t , w_t are Gaussian noise terms



Hidden continuous layer

Remember how

$$x_t = A_{z_t} x_{t-1} + b_{z_t} + v_t$$
 with $v_t \sim \mathcal{N}(0, Q_{z_t})$

It follows that

$$A_k \in \mathbb{R}^{M \times M}$$

$$b_k \in \mathbb{R}^M$$

$$Q_k \in \mathbb{R}^{M \times M}$$

In principle different for any latent state k = 1,...,K

Visible layer

In a similar way

$$y_t = C_{z_t} x_t + d_{z_t} + w_t$$
 with $w_t \sim \mathcal{N}(0, S_{z_t})$

It follows that

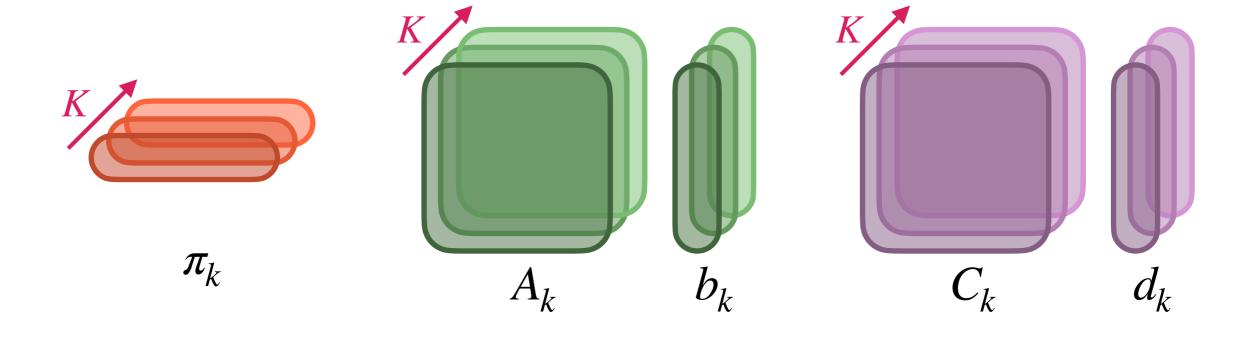
$$C_k \in \mathbb{R}^{N \times M}$$

$$d_k \in \mathbb{R}^N$$

$$Q_k \in \mathbb{R}^{N \times N}$$

Also different for any latent state k = 1,...,K

Recap



$$\theta = [(\underline{\pi_k}, A_k, Q_k, b_k, C_k, S_k, d_k)]_{k=1}^K$$

Simplified Model

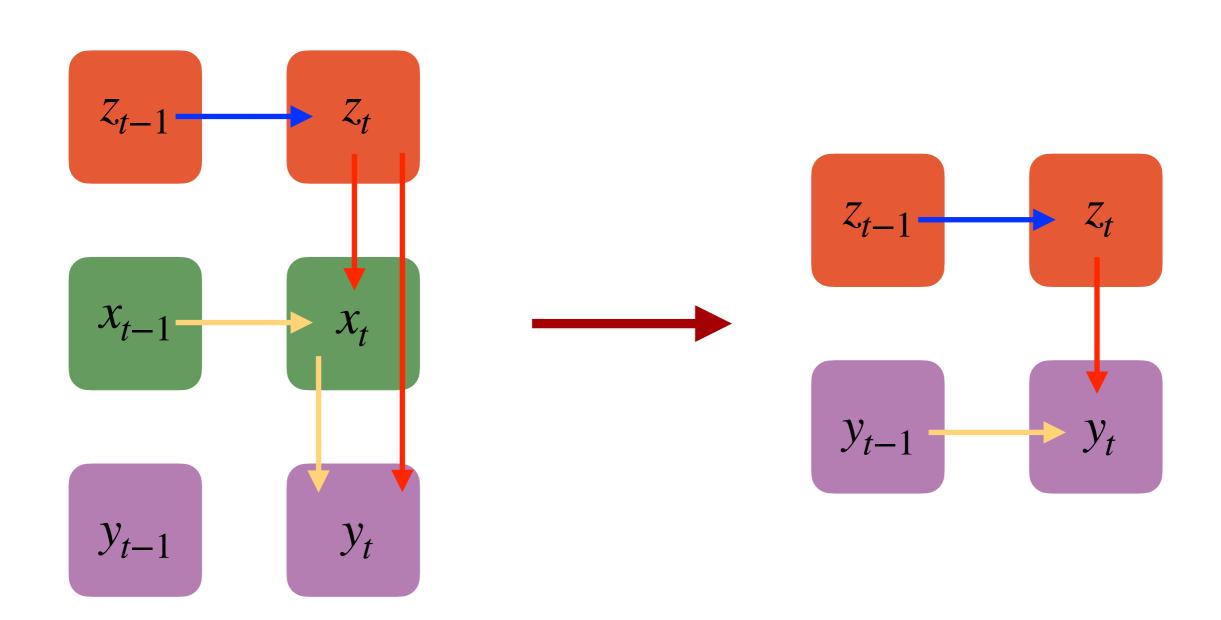
Despite having developed the model from a theoretical standpoint with three layers in practice we only used two of them, dropping the intermediate continuous layer

The two main reasons are linked to:

- 1 Explainability
- 2 Problem complexity

Simplified Model

Recap

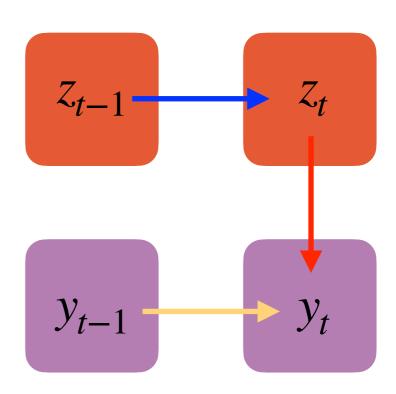


Simplified Model

Recap

$$z_{t} \sim \pi_{z_{t-1}}$$

$$y_{t} = A_{z_{t}} y_{t-1} + b_{z_{t}} + v_{t}$$



JAGS vs Stan



Stan - The inference problem

The first approach was setting up an appropriate Hamiltonian Monte Carlo model with Rstan.

Once the model was set, we needed a way to actually retrieve the latent states z. This was performed in two steps

- 1 Likelihood estimation with the Forward Algorithm
- 2 MAP z sequence retrieval with the Viterbi Algorithm

Stan - Forward algorithm

The forward algorithm has been implemented for both sampling efficiency and a Stan requirement, since discrete parameters are not allowed

The final goal is to compute the joint likelihood

$$p(x_{1:T}, y_{1:T})$$

We can start from

$$\gamma_t(k) = p(z_t = k, x_{1:t}, y_{1:t})$$

Stan - Forward algorithm

From

$$\gamma_t(k) = p(z_t = k, x_{1:t}, y_{1:t})$$

We can sum over the hidden states z at t-1 and then apply the chain rule to retrieve a recurrent expression

$$\gamma_t(k) = p(y_t | z_t = k, x_t) p(x_t | z_t = k, x_{t-1}) \sum_{j=1}^{K} \pi_{jk} \gamma_{t-1}(j)$$

Stan - Forward algorithm

Finally

$$\mathcal{L}_{k}(y_{t}) = p(y_{t} | z_{t} = k, x_{t}) = \mathcal{N}(C_{k}x_{t} + d_{k}, S_{k})$$

$$\mathcal{L}_{k}(x_{t}) = p(x_{t} | z_{t} = k, x_{t-1}) = \mathcal{N}(A_{k}x_{t-1} + b_{k}, Q_{k})$$

And so

$$\gamma_t(k) = \mathcal{L}_k(y_t) \mathcal{L}_k(x_t) \sum_j \pi_{jk} \gamma_{t-1}(j)$$

Stan - Forward algorithm

$$\gamma_t(k) = p(z_t = k, x_{1:t}, y_{1:t}) = \mathcal{L}_k(y_t) \mathcal{L}_k(x_t) \sum_j \pi_{jk} \gamma_{t-1}(j)$$

For the first step:

$$\gamma_1(k) = p(z_t = k, x_1, y_1) = \mathcal{L}_k(y_1) p(x_1 | z_1 = k) p(z_1 = k)$$

Where

 $p(x_1 | z_1 = k)$ is a multivariate Gaussian $p(z_1 = k)$ is a uniform distribution

Stan - Viterbi algorithm

Most of the steps are similar to the forward algorithm so we won't get in too much detail

$$\eta_t(k) = \operatorname{argmax}_{z_{1:t-1}} p(z_{1:t-1}, z_t = k, x_{1:t}, y_{1:t})$$

We can retrieve the recursive relation

$$\eta_t(k) = \mathcal{L}_k(y_t) \mathcal{L}_k(x_t) \operatorname{argmax}_j [\pi_{jk} \eta_{t-1}(j)]$$

With initialization

$$\eta_1(k) = p(z_1 = k, x_1, y_1)$$

Stan - Extracting information

Putting everything together we have

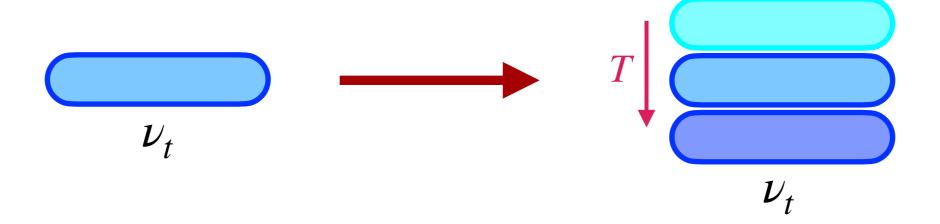
$$p(x_{1:T}, y_{1:T}) = \sum_{k=1}^{K} p(z_T = k, x_{1:T}, y_{1:T}) = \sum_{k=1}^{K} \gamma_T(k)$$

$$\hat{z}_{1:T} = \arg\max_{z_{1:T}} p(z_{1:T}, x_{1:T}, y_{1:T}) = \arg\max_{k} \eta_T(k)$$

In JAGS most of the issues that plagued the Stan implementation were, fortunately, absent and as a result the model was easier to code since:

- 1 It is enough to specify the stochastic dependance between variables to directly retrieve posterior distributions for all parameters
- 2 The Viterbi algorithm wasn't needed thanks to the native support for discrete parameters

On the other hand, JAGS does not allow redefining variables, resulting in the need for much more complex, multi-dimensional arrays.



Moreover, some functions are not vectorized

Dummy Random Walk

The first dataset we generated to test the system was a polar random walk with two states, also known in the literature as "Markov windshield wiper".

state 1: "move clockwise with step θ "

state 2: "move counter clockwise with step θ "

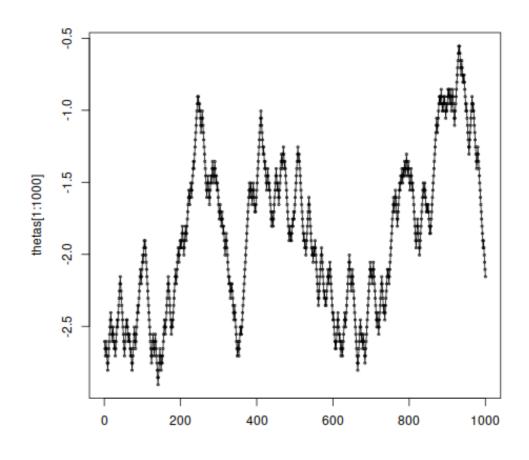
As for the transition matrix it simply was

$$\pi = \begin{pmatrix} 0.7 & 0.3 \\ 0.3 & 0.7 \end{pmatrix}$$

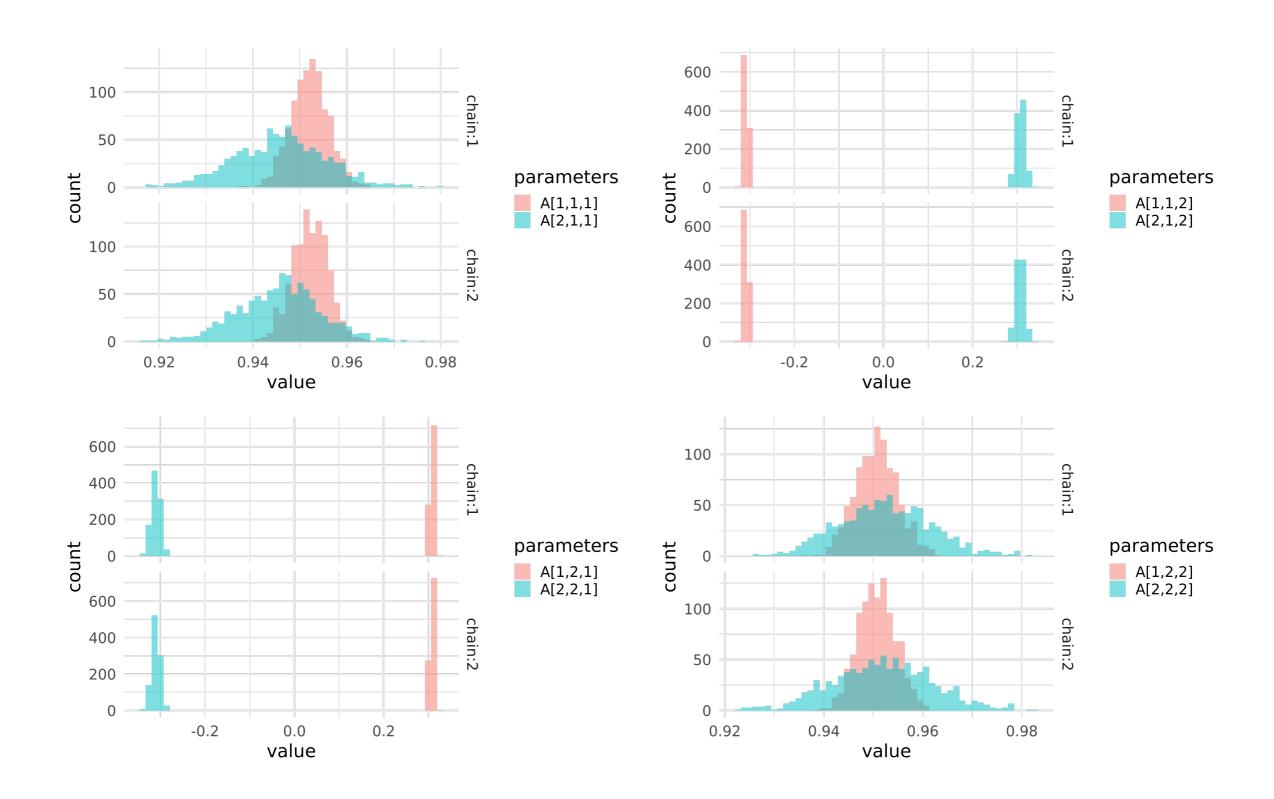
Dummy Random Walk

The first dataset we generated to test the system was a polar random walk with two states, also known in the literature as "Markov windshield wiper".

The dataset was then saved as (x, y) coordinates

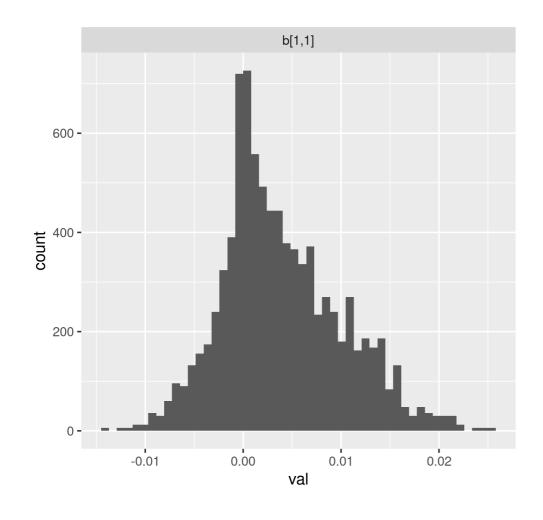


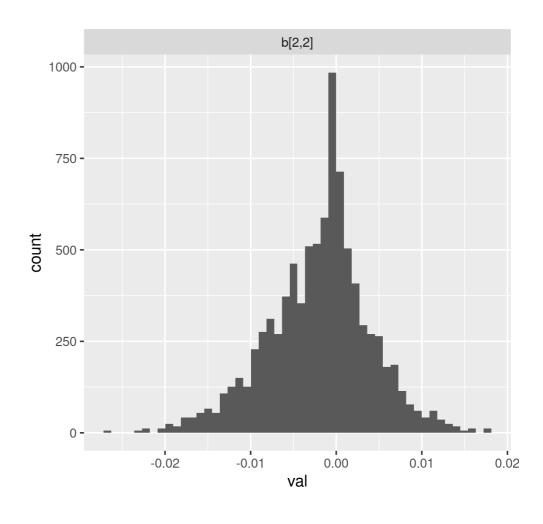
Results



Results

Distributions of the b vector entries for the first hidden state





Recursive Model

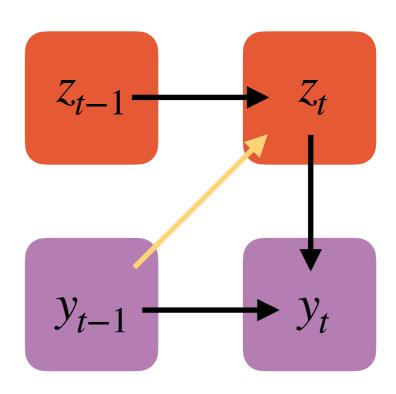
Refinement

One way to make the model more flexible is to modify the Markov process by introducing a conditional dependence on the previous state

$$z_t \sim \pi_{SB}(\nu_t)$$
 with $\nu_t = R_{z_{t-1}} y_{t-1} + r_{z_{t-1}}$

To have consistent dimensionality

$$R_k \in \mathbb{R}^{K-1 \times M}$$
$$r_k \in \mathbb{R}^{K-1}$$



Recursive Model

Stick Breaking

The stick breaking distribution is defined as

$$\pi_{SB}(\nu) = [\pi_{SB}^{(1)}(\nu), \dots, \pi_{SB}^{(K)}(\nu)]$$

$$\pi_{SB}^{(K)}(\nu) = \begin{cases} \sigma(\nu_k) \prod_{j=1}^k (1 - \sigma(\nu_j)) & \text{if } k \leq K - 1, \\ \prod_{j=1}^{K-1} (1 - \sigma(\nu_j)) & \text{if } k = K, \end{cases}$$

Where $\sigma(x)$ is the logistic function

$$\sigma(x) = \frac{e^x}{1 + e^x}$$

Recursive Model

Stick Breaking

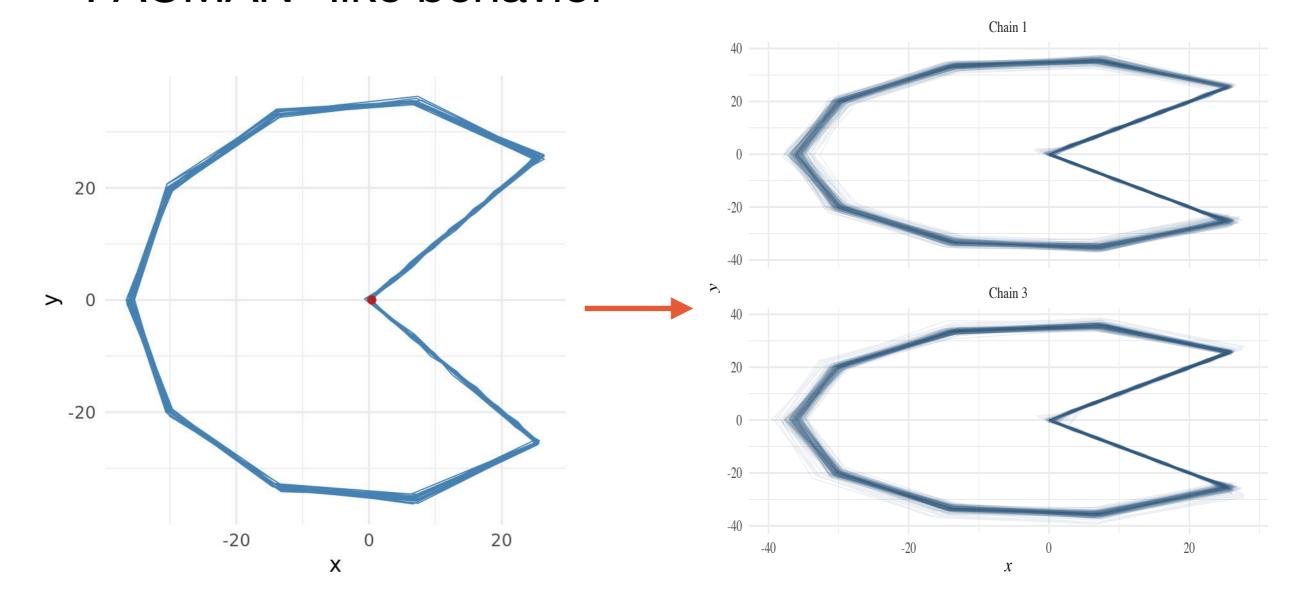
Both the Forward and Viterbi algorithms can be adapted to accomodate the Stick Breaking update rule, we will also need additional priors.

$$(R_k, r_k) \sim \text{MN}(M_r, \Sigma_r, \Omega_r)$$

PACMAN Dataset

Data generation

We moved to a slightly more complex dataset with a "PACMAN" like behavior

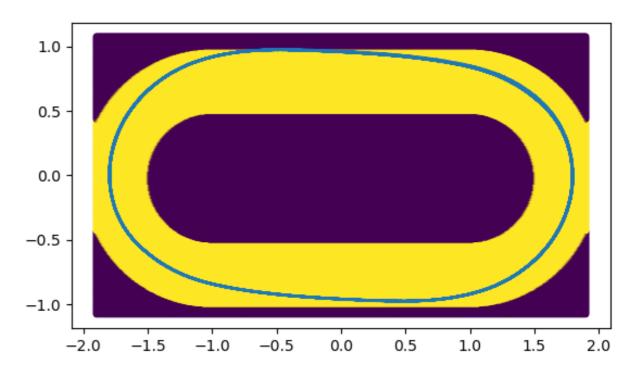


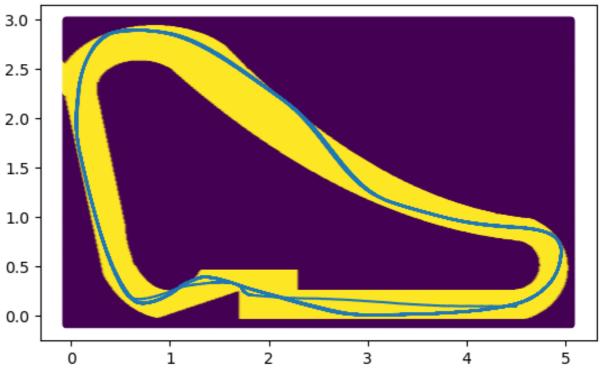
Racing circuit datasets

NASCAR and Monza circuits

A simple code has been implemented to simulate the driving behavior around a racing track.

First a simpler NASCAR style circuit, then a heavily romanticized rendition of the Monza circuit*

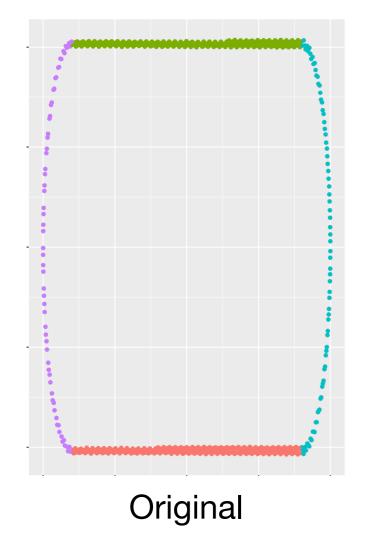


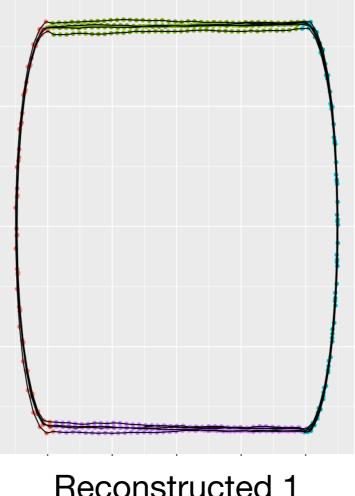


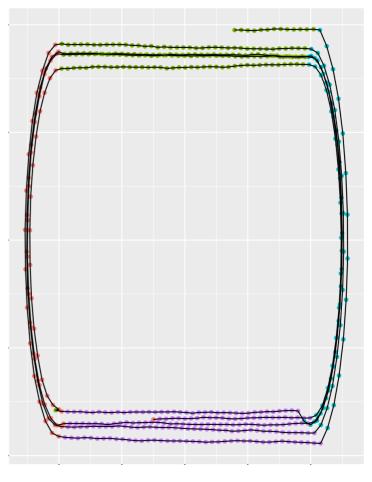
^{*}which notably lacks its only decent corner, the Ascari chicane. [editor note]

Results **NASCAR - JAGS**

Final simulations have been run with 20000 burn-in steps and then generated 10 sequences of reconstructed data







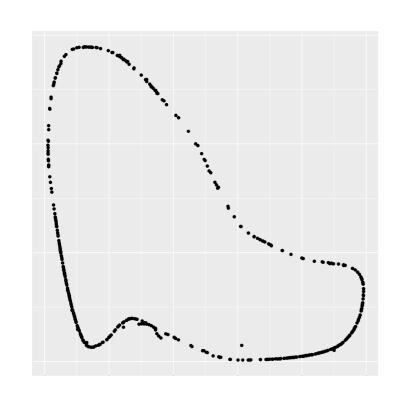
Reconstructed 1

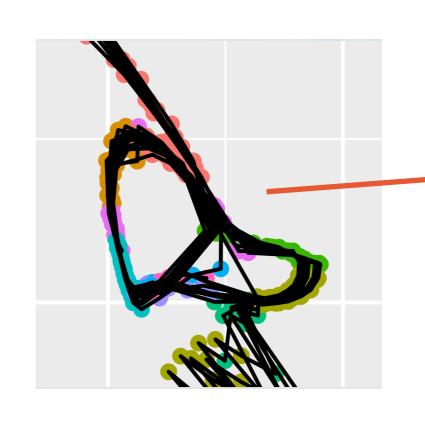
Reconstructed 2

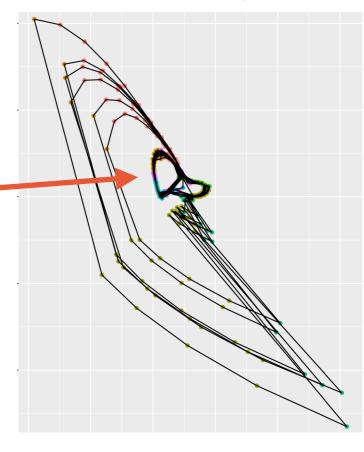
Results

Monza - JAGS

In this case the added complexity of the model tends to push the boundaries of the system, as a matter of fact stability is hard to achieve (10 hidden states)







Original

Reconstructed first iterations

Reconstructed later iterations

Results

Monza - JAGS

While distributions are reasonably centered the spread is much more pronounced compared to the NASCAR dataset, even with a 100'000 iterations burn in period

