Chi-square test

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1 Definition [?, ?]

The χ^2 -test is a statistical method to test hypothesis, where random variable follows multinomial distribution. It tests a null hypothesis stating that the frequency distribution of certain events observed in an observer sample is consistent with a particular theoretical distribution.

The most common χ^2 -test is a Pearson's chi-square test in which the test statistic is of following kind

$$\chi^2 = \sum_{i=1}^n \frac{(O_i - E_i)^2}{E_i} = N \sum_{i=1}^n \frac{(O_i/N - p_i)^2}{p_i}$$
 (1)

where

- χ^2 Pearson's cumulative test statistic, which asymptotically approaches a χ^2 distribution.
- O_i number of observation of type i
- N total number of observations
- $E_i = Np_i$ the expected (theoretical) count of type i
- n the number of cells in the table $(rows \cdot columns)$

If the test fails, then the appropriate test statistic has approximately a noncentral χ^2 distribution with the same degrees of freedom (df) and a noncentrality parameter λ , which depends on alternative considered.

2 Test types [?]

Pearson's chi-squared test is used to assess three types of hypothesis testing: goodness of fit, homogeneity, and independence.

- A test of **goodness of fit** establishes whether an observed frequency distribution differs from a theoretical distribution. The χ^2 test statistic follows the χ^2 distribution with df = n - 1.
- A test of **homogeneity** compares the distribution of counts for two or more groups using the same categorical variable. The χ^2 test statistic follows the χ^2 distribution with $df = (rows - 1) \cdot (columns - 1)$.
- A test of **independence** assesses whether observations consisting of measures on two variables, expressed in a contingency table, are independent of each other.

The χ^2 test statistic follows the χ^2 distribution with $df = (rows - 1) \cdot (columns - 1)$.

2.1 Test of goodness of fit

Let $X = (X_1, ..., X_k)$ be a multinomial ranom variable with parameters $n, p_1, ..., p_k$. Suppose we wish to test

$$H_0: p_i = p_i^0 i = 1, 2, ..., k$$
 (2)

against

$$H_a$$
: not all p's are as given by H_0 (3)

where the p'_i are given expected numbers. The value of the chi-square test-statistic is

$$\chi_{H_0}^2 = \sum_{i=1}^k \frac{\left(x_i - np_i^0\right)^2}{np_i^0} = \sum_{i=1}^k \frac{\left(\hat{p}_i - p_i^0\right)^2}{p_i^0} \tag{4}$$

The chi-square test reject H_0 if

$$\chi_{H_0}^2 > \chi_{k-1,1-\alpha}^2,\tag{5}$$

where α is the significance level, and $\chi^2_{k-1,1-\alpha}$ is the quantile of order $1-\alpha$ of the χ^2 distribution with k-1 degrees of freedom. The p-value of the test is

$$p - value = P\left(X < \chi_{H_0}^2\right) \tag{6}$$

To evaluate the power of the test let's precisely define the alternative H_a as follow.

$$H_a: p_i = p_i^a \qquad i = 1, 2, ..., k$$
 (7)

Thus the power o the test is

$$Power = P\left(X_a > \chi^2_{k-1,1-\alpha}\right) \tag{8}$$

where X_a is a ranom variable that follows the noncentral χ^2 distribution with the noncentrality parameter

$$\lambda = n \sum_{i=1}^{k} \frac{\left(p_i^a - p_i^0\right)^2}{p_i^0} \tag{9}$$

2.2 Test of independence

Let $X = (X_{ij}) \in \mathbb{R}^{r \times c}$ be a multinomial random variable with parameters n, p_{ij} where i = 1, 2, ..., r, j = 1, 2, ..., c and $\sum_{i=1}^{r} \sum_{j=1}^{c} p_{ij} = 1$. Suppopse we wish to test independence

$$H_0: p_{ij} = p_i.p._j i = 1, 2, ..., r j = 1, 2, ..., c$$
 (10)

against

$$H_a$$
: not all the equatons given under H_0 are satisfied (11)

where $p_{i.} = \sum_{j=1}^{c} p_{ij}$ and $p_{.j} = \sum_{i=1}^{r} p_{ij}$. The value of the chi-square test-statistic is

$$\chi_{H_0}^2 = \sum_{i=1}^r \sum_{j=1}^c \frac{(x_{ij} - x_{i.} x_{.j} / n)^2}{x_{i.} x_{.j} / n}$$
(12)

where $x_{i.} = \sum_{j=1}^{c} x_{ij}$ and $x_{.j} = \sum_{i=1}^{r} x_{ij}$. We observe that

The chi-square test reject H_0 if

$$\chi_{H_0}^2 > \chi_{(r-1)\cdot(c-1),1-\alpha}^2,\tag{13}$$

To evaluate the power of the test let's precisely define the alternative H_a as follow.

$$H_a: p_{ij} = p_{i.}p_{.j} + \frac{c_{ij}}{\sqrt{n}}, \quad i = 1, 2, ..., r \quad j = 1, 2, ..., c, \quad where \quad \sum_{i=1}^{r} \sum_{j=1}^{c} c_{ij} = 0,$$
 (14)

Thus the power o the test is

$$Power = P\left(X_a > \chi^2_{(r-1)\cdot(c-1),1-\alpha}\right) \tag{15}$$

where X_a is a ranom variable that follows the noncentral χ^2 distribution with the noncentrality parameter

$$\lambda = \sum_{i=1}^{r} \sum_{j=1}^{c} \frac{c_{ij}^{2}}{p_{i} \cdot p_{\cdot j}} - \sum_{i=1}^{r} \frac{c_{i \cdot}^{2}}{p_{i}} - \sum_{j=1}^{c} \frac{c_{\cdot j}^{2}}{p_{\cdot j}}, \tag{16}$$

where $c_{i\cdot} = \sum_{j=1}^{c} c_{ij}$ and $c_{\cdot j} = \sum_{i=1}^{r} c_{ij}$. If $\Delta_{ij} = c_{ij}/\sqrt{n}$, then

$$\lambda = \frac{1}{n} \left[\sum_{i=1}^{r} \sum_{j=1}^{c} \frac{\Delta_{ij}^{2}}{p_{i} \cdot p_{\cdot j}} - \sum_{i=1}^{r} \frac{\Delta_{i\cdot}^{2}}{p_{i\cdot}} - \sum_{j=1}^{c} \frac{\Delta_{\cdot j}^{2}}{p_{\cdot j}} \right], \tag{17}$$

2.3 Test of homogenity

Let $X_i = (X_{ij}) \in \mathbb{R}^c$ be a multinomial random variable with parameters n_i, p_{ij} for i = 1, 2, ..., r and $\sum_{j=1}^c p_{ij} = 1$. Suppopse we wish to test homogenity

$$H_0: p_{1j} = p_{2j} = \dots = p_{rj} = p_j \quad j = 1, 2, \dots, c$$
 (18)

against

$$H_a$$
: not all the equators given under H_0 are satisfied (19)

The value of the chi-square test-statistic is

$$\chi_{H_0}^2 = \sum_{i=1}^r \sum_{j=1}^c \frac{(x_{ij} - x_i \cdot x_{\cdot j}/n)^2}{x_i \cdot x_{\cdot j}/n}$$
(20)

where $x_{i.} = \sum_{j=1}^{c} x_{ij} = n_i$ and $n = \sum_{i=1}^{r} n_i$. The chi-square test reject H_0 if

$$\chi_{H_0}^2 > \chi_{(r-1)\cdot(c-1),1-\alpha}^2,\tag{21}$$

To evaluate the power of the test let's precisely define the alternative H_a as follow.

$$H_a: p_{ij} = p_j + \frac{c_{ij}}{\sqrt{n}}, \qquad j = 1, 2, ..., c \quad where \quad \sum_{j=1}^{c} c_{ij} = 0,$$
 (22)

Thus the power o the test is

$$Power = P\left(X_a > \chi^2_{(r-1)\cdot(c-1),1-\alpha}\right) \tag{23}$$

where X_a is a ranom variable that follows the noncentral χ^2 distribution with the noncentrality parameter

$$\lambda = \sum_{j=1}^{c} \frac{1}{p_j} \left[\sum_{i=1}^{r} c_{ij}^2 \frac{n_i}{n} - \left(\sum_{i=1}^{r} c_{ij} \frac{n_i}{n} \right)^2 \right], \tag{24}$$

If $\Delta_{ij} = c_{ij}/\sqrt{n}$, then

$$\lambda = \frac{1}{n} \sum_{j=1}^{c} \frac{1}{p_j} \left[\sum_{i=1}^{r} \Delta_{ij}^2 \frac{n_i}{n} - \left(\sum_{i=1}^{r} \Delta_{ij} \frac{n_i}{n} \right)^2 \right]$$
 (25)

3 Sample size

References

- [1] Chi-squared-test https://en.wikipedia.org/wiki/Pearson%27s_chi-squared_test.
- $[2] \ \ Guenther, W.\ (1977).\ Power\ and\ Sample\ Size\ for\ Approximate\ Chi-Square\ Tests.\ The\ American\ Statistician,\ 31(2),\ 83-85.$