UNPACKING THE EDUCATION PRODUCTION FUNCTION:

ESSAYS ON TEACHING, PARENTING, AND MEASUREMENT IN EMERGING

CONTEXTS

A DISSERTATION

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FOR THE DEGREE OF

DOCTOR OF PHILOSOPHY

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June 2021

I certify that I have read this dissertation and that, in my opinion, it is fully adequate in scope and quality as a dissertation for the degree of Doctor of Philosophy.

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(Susan Athey)

Approved for the Stanford University Committee on Graduate Studies

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Preface

This dissertation consists of three essays that investigate topics concerning the three main characters in the education production function – the teacher, the parent, and the student. All three essays speak to modes of learning that have a broad reach. Mobile education in the cases of papers one and three, higher education in developing countries in the case of paper two. All three papers leverage novel sources of information to inform policymaking, design, and future research. Papers one and two are empirical inquiries powered by experimental and quasi-experimental evidence respectively, while paper three expands the methodological toolkit available to researchers and practitioners.

The first chapter of this dissertation brings experimental evidence to the problem of improving parental follow through in mobile education. The paper is authored by me while its contents belong to a part of a larger project that will be published in co-authorship with Susan Athey. Mobile learning apps offer us the opportunity to improve parenting at an unprecedented scale around the world. Effective parenting on these platforms is constrained by follow through. I investigate whether features on a recommendation page can increase follow through. I find that Cognitive overload is a major factor affecting parental follow through. Providing a top-of-the-page reason significantly decreased the completion rate by 70%. We also find null effects on completion from adding 1) a commitment question, 2) a guide to navigate back to the page, and 3) a link to the child's learning report. The hypothesized positive effects from these features are likely negated by the cognitive load penalties. However, increasing the number of recommended modules increased the number of modules each complier completed without increasing cognitive overload.

The second chapter of this dissertation presents quasi-experimental evidence for an age-old question – how does faculty research affect student learning? The paper is co-authored with Prashant

Loyalka. Other co-authors who contributed to data collection and discussions include Guirong Li, Elena Kardanova, Igor Chirikov, Ningning Yu, Shangfeng Hu, Huan Wang, Liping Ma, Fei Guo, Ou Lydia Liu, Ashutosh Bhuradia, Saurabh Khanna, and Yanyan Li. Whether faculty research affects college student achievement has long been the subject of debate. Previous studies use subjective measures of student achievement, focus on correlation rather than causation, and typically focus on one college or department, thus lacking generalizability. Using unique, large-scale survey and assessment data that we collected from nationally representative samples of four-year STEM undergraduates and faculty in China, India, and Russia as well as a student fixed effects identification strategy that accounts for differential sorting of students to faculty, we present generalizable estimates of the effect of faculty research on objective, standardized measures of student achievement. Results show that faculty research has a negative effect on student achievement, suggesting direct tradeoffs between the university's dual mission of producing research and learning.

The third chapter of this dissertation concerns the estimation of student models using large sparse data from online learning. The paper is authored by me while its contents belong to a part of a larger project that will be published in co-authorship with Susan Athey. The rise of popular mobile education applications produced data where a large number of students each answers a small subset of questions from a large question bank. Traditional approaches from the education measurement literature face important limitations in this context where data is large but sparse. We propose models based on latent factorization and Bayesian variational inference to address these challenges. Our models retrieve true parameters with greater fidelity than traditional models in simulations. They also scale well computationally to industrial-size datasets. Compared to traditional specifications, latent factorization models can make more accurate predictions on the hold-out test set in general. More latent factors and adding hierarchical dependence on question attributes contribute to better predictive performance in lower-frequency content areas. Our models also compare favorably in both computational performance and predictive accuracy against similar models from recent literature. We conclude by describing a real-world application of our models in personalizing homework assignments. In a future study, we plan to run experiments with this application to quantify the impact of personalization.

Acknowledgments

I have received support from many people as I worked towards this dissertation. I am particularly indebted to Susan Athey whose mentorship made these papers possible. Susan supported me through the initial phases of my research project even when output was slow and by no means guaranteed. Working as part of Susan's team, I was able to pursue research questions and apply methodologies that truly excites me. This would not have been possible otherwise. I am grateful to have borne witness to Susan's commitment to her students while multi-billion dollar firms are competing for the same slots on her calendar. Her approachability despite all her achievements is a constant lesson in humility for me as I move on to the next stage of my career.

I want to thank my co-advisors Susanna Loeb and Eric Bettinger, whose advice has been instrumental in helping me navigate the rich academic resources at Stanford. They encouraged me to chart my own path through the myriad of courses and research opportunities available here. My research and career direction would have looked very different had they not encouraged me to pursue these interests.

As an engineering student in my undergraduate years, my path towards a Ph.D. in the Economics of Education was shaped by the years I spent working with the Rural Education Action Program (REAP). I want to thank Prashant Loyalka for helping me as a mentor and a friend supporting me over my years working at REAP and my subsequent Ph.D. years. I am extremely grateful to Scott Rozelle for welcoming me into REAP and introducing me to economics as a way to help people in need. I am constantly inspired by Scott's dedication to helping people who lead very different lives than our own.

I would also like to extend my gratitude to the graduate program at Stanford. My graduate

Contents

List of Tables

List of Figures

Chapter 1

Improving Parental Follow Through in Mobile Education

1.1 Introduction

Large-scale mobile learning apps have risen in popularity in recent years. Byju's in India, 17Zuoye in China, and Khan Academy in the US are just some examples of companies that offer significant mobile options. Together, these three companies account for more than 180 million users worldwide (Singh, 2020; Khan Academy, 2020a; Sunny Education Inc., 2018). These apps often offer features for parents to enhance their engagement with their children's learning. Features for parents include those used to monitoring learning, increase practice and exposure, as well as personalized recommendations for shoring up weak links.

Mobile learning apps offer us the opportunity to improve parenting at an unprecedented scale around the world. This holds great promise for education in general, since parents are integral in their children's life successes. However, several key obstacles stand in the way of realizing this vision. Unlike traditional venues of parental engagement (e.g. teacher-parent meetings), effective parental engagement require three key components (Figure 1.1).

Firstly, parents need to use the App. With the explosion of number in the number of apps we install on our phones, this is not an easy task. Secondly, education apps are typically overflowing
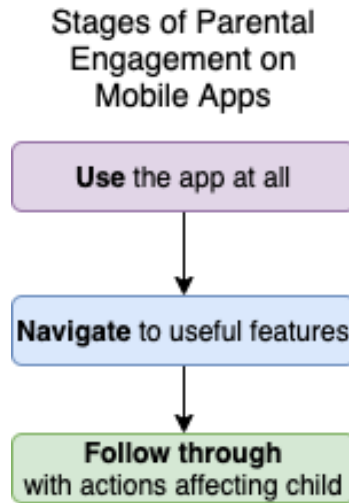
Figure 1.1: Effective parental engagement in mobile Apps

with complex features, it is hard for the parent to navigate to useful features most useful for their child's learning. Thirdly, even when a useful course of action is identified, the parent still needs to follow through and take action with their child to generate any meaningful impact on her.

In reality, mobile platforms face challenges in engaging parents in each of the three stages named in Figure 1.1. Even with a large user base, the average parental usage rate can be very low. Apps have many complex features and parents have a hard time navigating to useful content. Finally, parents often do not follow through in taking action with their children.

This study is a part of a series of studies examining each of these issues through experimentation. In this paper, we seek to address the issue of parental follow through with a large-scale experiment leveraging insights from behavioral economics. Behavioral interventions in parenting have been increasingly studied in recent years (see Bergman (2019) for a review).

In this study, we randomized the features shown to parents on a page recommending customized Learn-Practice-Explain (LPE) exercise modules for their children. We find that cognitive overload is a major factor in our experiment with features shown on a recommendation page. Contrary to our hypotheses, we find that providing a reason for recommendation at the top of the page significantly decreased the completion rate of LPE modules by 70%. Consistent with a cognitive overload interpretation, users who were shown a reason were 60% less likely to answer the survey question, but those who responded provided statistically indistinguishable answers from the control

group. We also find null effects on completion from adding 1) a commitment question, 2) a guide to navigate back to the page, and 3) a link to the child's learning report. The hypothesized positive effects from these features were likely muted or negated by the penalties they induce as cognitive overload.

While we find strong evidence of cognitive overload from the above-mentioned features, we find no evidence that increasing the number of modules increases cognitive overload as users shown more modules were equally likely to answer the survey question. Increasing the number of recommended modules did not increase the number of users completing any modules, but it increased the number of modules each complier completed.

Our experiments advance the existing behavioral literature in parenting through observing fine-grained behavioral data generated in mobile apps. We also reveal a set of issues in parental follow through that can inform future researchers and practitioners. Finally, my results contribute to the understanding and improvement of an important mode of education that serves a large number of the world's learners. The paper is organized as follows. Section 2 samples the related literature, section 3 describes our partnering company's platform, section 4 details the experimental design, section 5 lays out our results, and section 6 concludes.

1.2 Related work

Our study builds on an expanding literature of behavioral experiments in parenting (see Bergman (2019) for a review). Many randomized studies have revealed the salience of behavioral barriers may prevent parents from investing optimally in their children's education. Among the barriers identified are biased beliefs about their children's performance, limited cognitive bandwidth, and the cost of monitoring. Interventions ranging from text messaging about student's behavior (Bergman, 2015) to tax assistants (Bettinger et al., 2012) have been shown to be effective at addressing one or multiple barriers.

Apart from their primary findings, these studies often also uncover non-obvious pitfalls. Cunha et al. (2017) found that interactivity in a message program decreased learning outcomes. In addition, Gallego et al. (2017) found that the lack of certain cues was responsible for a loss of half of the effect of the program. Bergman et al. (2019) found that opt-in programs fail to deliver an impact due to

very low uptake rates (less than 11%) compared to an opt-out program (95% uptake).

Our study is related to the growing literature on personalized learning. Education technology companies have long made claims about the imperative of personalization. Efficacious programs, such as those examined by Muralidharan et al. (2019) and Banerjee et al. (2007), have personalized components. However, other technology-based learning programs without personalization have shown to be similarly efficacious (e.g. Mo et al. (2014); Lai et al. (2012)). We contribute to the literature on personalization by providing experimental evidence on a recommendation page for parents filled with personalized content for their children.

Since our experiment takes place on a recommendation page, we are also inspired by the literature around the design of recommendation systems. For example, in the literature on features of recommendation systems, user experience studies found that a larger number of recommendations can introduce choice fatigue (Bollen et al., 2010; Konstan and Riedl, 2012). Pu and Chen (2006) documents how explanatory text increase trust in recommendation system and intention to return to the page. We run multiple interventions in our experiment and one of our interventions adds an explanation to the page explaining to the parent why the exercise modules were recommended to her.

Our study is also related to the behavioral literature on commitment devices. Commitment devices have been studied extensively in behavioral economics. Bryan et al. (2010) discusses the various theories on commitment devices. They highlight the difference between hard and soft commitment. The latter incurs psychological costs instead of real economic penalties. Soft commitment devices have been applied to financial discipline (e.g. Thaler and Benartzi (2004)) and charitable giving (e.g. Breman (2011)). Recently, Mayer et al. (2015) reports on using commitment devices on parents to increase their usage of a reading App. One of our interventions involves varying the presence of a soft commitment question that parents can choose to check.

Finally, our study is closely linked to the literature on technology-induced cognitive overload. Originally proposed to explain the missing link between IT investment and productivity gain (Karr-Wisniewski and Lu, 2010), technology-induced cognitive overload has been observed in many settings beyond the workplace. In educational settings, teachers who taught classes that had more complicated reports generated by an AI tutor used the reports less and, in turn, had much smaller impacts

on their student achievements (Kim et al., 2019). Information overload patterns have also been observed with parents. For example, (Fricke et al., 2018) finds complicated messages increased drop out from the text messaging program. Cortes et al. (2018) found that compared to 3 messages per week, 5 messages per week decreased the effect of a parental messaging experiment.

1.3 Context

Our study takes place on the platform of our partner company, 17Zuoye. 17Zuoye runs 3 separate apps for elementary school teachers, students, and parents in China. They enjoy a high level of adoption in the country with 60 million users in 2018 (Sunny Education Inc., 2018). The app is designed around the functionality that allows teachers to post homework to students. The draw for teachers is that the app provides them with ready-to-use homework questions and grades their students' responses automatically, reducing their workload in day-to-day teaching. The app is free to use for teachers, students, and parents. For their revenue stream, 17Zuoye runs a "freemium" model for parents who want to purchase education products in the App.
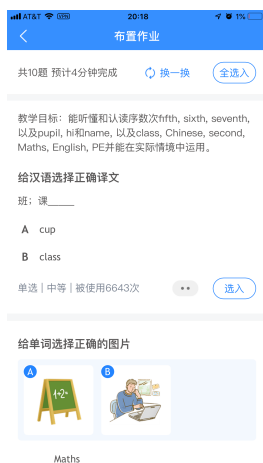
Figure 1.2: Assigning homework in the teacher app
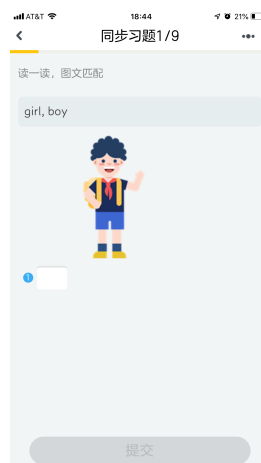
Figure 1.3: Completing homework in the student app

Figure 1.2 shows the teacher app's interface for assigning homework. Teachers get to choose from a bank of curriculum-relevant homework questions to assign to their class. Students complete the assigned homework on the student app (see Figure 1.3). Their responses are automatically graded

and recorded in the system. Teachers have access to student records through a data console in the teacher app.

Figure 1.4: HW monitoring page in the parent app   Figure 1.5: Purchase page in the parent app

In the parent app, parents have access to their child's homework records. Parents get to monitor their child's homework performance on particular pages where this information is displayed. Parents can also purchase education products for their children. Examples of products include educational games for math designed as a jungle adventure, e-books designed as a recap of course material, and online oral English courses taught in real-time. The app also carries a wide selection of free study resources such as English videos, ebooks, and practice questions. Parents may spend time in the app browsing these free items and assigning them to their children if they choose to.

1.4 Experimental designs

In the recommendation page features experiment, we want to investigate how features impact parental satisfaction and follow-through. To do this, we randomized the features shown on a page recommending customized Learn-Practice-Explain (LPE) exercise modules for their children. While the platform has content for math, English, and Chinese, LPE modules were only available for math. The recommendation page was only available to parents whose child has completed a math exam on the platform. The recommendation page is designed to prompt parents to get their child to complete the LPE exercise using the parent's cell phone.

Recommended modules LPE modules are generated based on the questions students missed on the math exam. Questions on the math exam are associated with knowledge points. Each LPE module is specially designed to teach one knowledge point (e.g. "interpret a figure to write down a multiplication expression"). For each student, her exam scores would be aggregated into the correct rate by each knowledge point. The LPE modules would then be ranked in order from the knowledge point with the lowest correct rate to the knowledge point with the highest correct rate. If the student received a perfect score on the math exam, then no LPE module will be recommended to the student. In this case, the parent does not have access to the recommendation page.

Our experiment population is composed of 51860 parents who entered the recommendation page between Jan 14th, 2021, and Jan 29th, 2021. Parents entered the page through either the entry point on the exam report page in the App or by clicking a system-generated push message reminding them to open the recommendation page. Our treatment conditions are variations of the features shown on the recommendation page. Assignment to treatment arms is randomly assigned and orthogonal to how the parent entered the page.

1.4.1 Sample recruitment and data collection

Sample recruitment took place on 17Zuoye's platform in their parent app. The recommendation page was available to any parent whose child had taken a commissioned exam. The parents who chose to open the recommendation page when the experiment was ongoing were all assigned to one of the treatment conditions. While our interventions were directed at parents, they call on parents to take action with their children by getting them to complete exercises on the parents' phones. Thus,

the assignment to treatment conditions may affect children's outcomes as well. Data collection is done automatically through 17Zuoye's data infrastructure.

1.4.2 Learn-Practice-Explain Module

The LPE modules integrate animation with practice questions and explanatory voice-over and text. The student is guided through a series of questions and animations that explains one knowledge point (see Figure 3.17). The typical time to complete an LPE module is between 2 and 4 minutes. As the student answers the questions, only a correct answer would allow the student to move forward. A wrong answer would trigger a hint and the student is directed to answer the question again. Importantly, the modules are designed to target granular knowledge points. For example, "rewriting numbers in units of Wan (ten thousand)" is a knowledge point and correspondingly, an LPE module, in 4th-grade math.

1.4.3 Sampling and treatment assignment

Our experimental population comes from elementary school parents whose children partook in math subject exams on the 17Zuoye platform. These exams are created by teachers, schools, or the local education authority and delivered to students via 17Zuoye's app. Exam questions are labeled according to a knowledge framework. As such, our recommendation page generates recommended LPE modules that tackle the same node in the knowledge framework as the students' most error-prone exam questions.

In total, the recommendation page was available to 1.08 million parents. Parents were prompted to open the page through in-App alerts and push messages. Our population consists of the 52K parents who entered the recommendation page in the period of data collection lasting from Jan 11th to Jan 29th, 2021.

The experimental arms are summarized in Table 1.1. The treatment assignment of each feature is independent of that of another feature. As such, the experiment is a fully-crossed design. Due to the experimental data yielding clear results for two features (recommendation reason, and the number of modules) in the first week, starting from Jan 23rd, these two features were set to values that optimized completion rates while other features' treatment assignment probabilities remained

Figure 1.6: Question in an LPE module
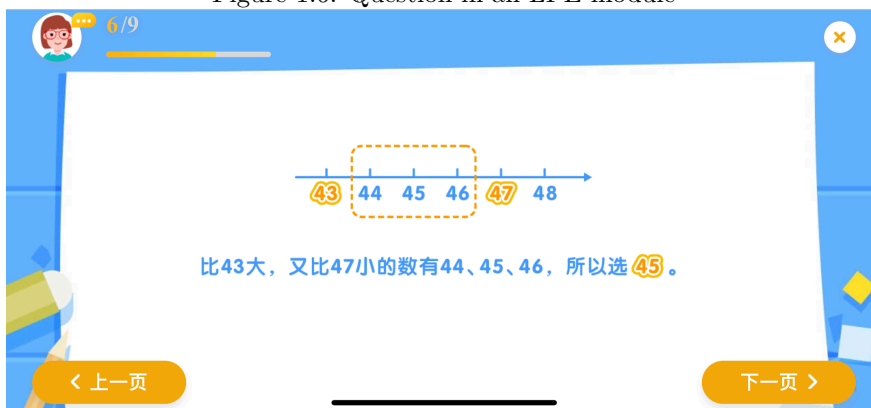
Figure 1.7: Explanatory text/audio in an LPE module

unchanged. Due to the need to experiment with various wordings of the commit question, the commit question feature was only finalized starting Jan 23rd. Hence, unlike the other features, the analysis of the commitment question feature uses only data from the 23rd onward.

Assignment value	Probability of assignment	Intervention
{Excluded, Customized version, Generic version}	$\{0.52, 0.24, 0.24\}$	Text to explain reason for recommendation
{1, 3, 5, 9}	$\{0.25, 0.25, 0.25, 0.25\}$	Maximum number of recommended modules
{Excluded, Base version, Reminder version}	$\{0.3, 0.35, 0.35\}$	Commitment question with checkbox
{Excluded, Included}	$\{0.5, 0.5\}$	Link to a guide to help the parent navigate back to the page
{Excluded, Included}	$\{0.5, 0.5\}$	Link to a report on the child's learning

Table 1.1: Recommendation page features experiment: Treatment assignment

Figure 1.8 shows how the different features in the way they appear in the parent App. On the left of the figure, we have mapped treatment conditions to sections of the screen they affect. On the right, we document the fixed features of the page that are shown to all users regardless of their treatment assignment.
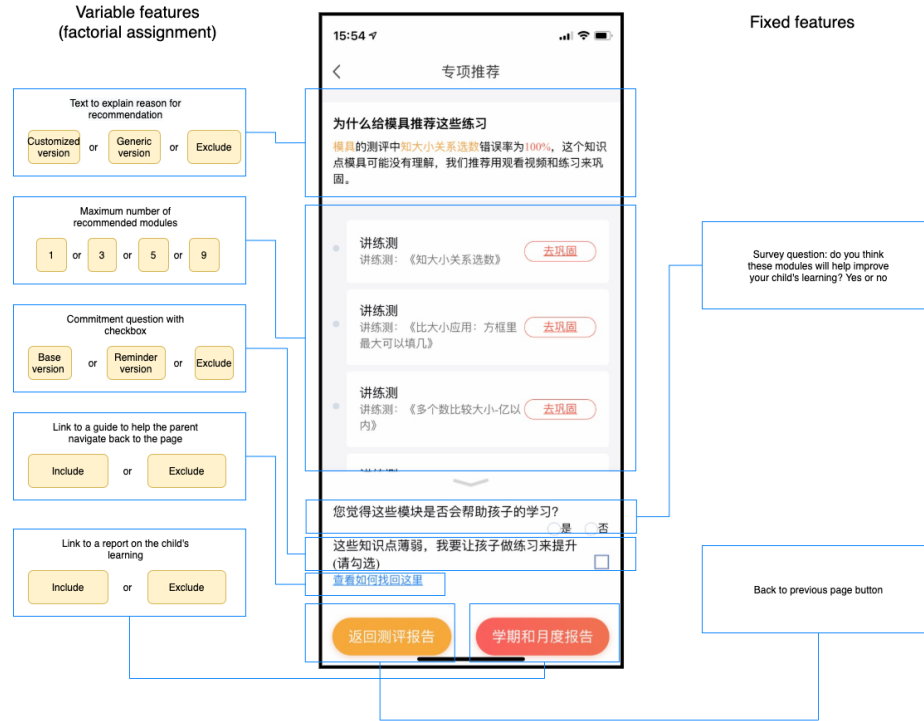
Figure 1.8: Recommendation page layout

Table 1.2 describes the outcomes and covariates used in this analysis. We use the extrapolated completion count as our primary outcome in this paper, which likely captures completion rates more accurately than the other measures. Nevertheless, when we use actual completion count and start counts, our results mirror the current results.

We do not have access to an indicator for whether the parent or the student actually completed the LPE modules. However, an LPE module takes more than two minutes to complete on average. Hence, although it is possible that parents were completing these modules, the vast majority of the completes were likely by students, not their parents. Nevertheless, the LPE module started metric is likely to have captured parents "shopping around" by clicking on the modules to reveal their contents.

Variable type	Name	Description
Outcome	useful	Whether a user answered "Yes" to the survey question (vs answering "No") conditional on responding.
Outcome	respond_survey	Whether a user responded to the survey question.
Outcome	n_complete_extrp	Extrapolated number of completed LPE modules. Since not all users who complete a course click on the "complete" button. As such, this measure not only counts the "complete" button events but also includes any start within 10 minutes of clicking a complete button.
Outcome	any_complete	Whether the user has completed any LPE modules.
Secondary outcome	n_complete	Number of completed LPE modules by counting of click on the "complete" button at the end of each module. This serves as a robustness check for n_complete_extrp.
Secondary outcome	n_started	Number of started LPE modules, this measure will capture parents' "shopping" behavior. This serves as a robustness check for n_complete_extrp.
Covariate	is_non_metro	Family lives in a non-metro county in China, typically an indicator for lower affluence.
Covariate	low_achievement	Child belongs to the bottom 30% of homework score distribution (on past exams in the current semester).
Covariate	high_achievement	Child belongs to the top 30% of homework score distribution (on past exams in the current semester).
Covariate	high_activity	Parents belonging to top 30% in number of days with any parent App activity or the parent has opened the learning report within the past two months.

Table 1.2: Recommendation page features experiment: Outcomes and covariates

Every feature in this experiment was designed to help improve parental satisfaction and follow through. Variations in the maximum number of modules displayed allow us to investigate the optimal number of recommended modules. Figure 1.9 explains the hypothesized causal pathways for each feature in the experiment.
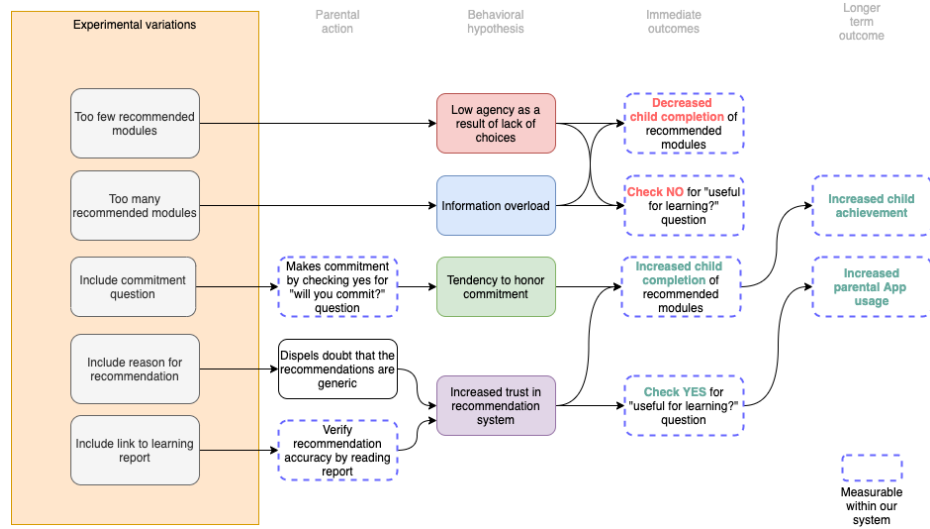
Figure 1.9: Hypotheses for the function of recommendation page features

1.5 Results

In this section, we document the results of our experiments. Throughout this paper, error bars in figures correspond to 95% confidence intervals unless otherwise noted. Regression tables show standard errors in parentheses.

My research question is how do features on a page recommending custom exercise modules (LPE) affect parents' assessment of the usefulness of the recommended items and completion rates? The features we evaluate include:

1. Existence of text to explain the reason for recommending the modules.

2. Existence of a checkbox with a piece of commitment text.

3. The maximum number of recommended modules the page displays.

4. Existence of a link to a report on the child's learning.

5. Existence of a link to a guide to help the parent navigate back to the page.

Table 1.3 summarizes the main metrics for this experiment. We see that the response rate to the survey question is generally low, at 7% of all those who entered the page. But out of those responding, the vast majority 93% answer "YES" to the usefulness question. A little more than 4% of students actually complete any LPE module after their parents entered the page. We also note that there are more starts of LPE modules than completion, indicating potential "shopping" behavior on the part of the parents.

Event	N	Proportion	Notes
Page entry (unique parent)	51860	1.00	
Unique parent answer useful question	3600	0.07	
Unique parent answer "YES" to useful question	3342	0.93	Within parents who responded to survey question.
Unique parent check commit question	1087	0.02	
Unique student that completed any LPE module	2320	0.04	
N clicks on learning report	3043	0.06	
N clicks on how-to-find-back link	535	0.01	
N starts for LPE module	13739	0.26	
N completes for LPE module	7799	0.15	

Table 1.3: Recommendation page features experiment: Summary table

We document the sample sizes for all the treatment conditions of this experiment in Table 1.4. As all the treatment conditions were assigned orthogonal to each other, also known as a factorial design, a parent may constitute a unit in multiple treatment arms.

Feature	Treatment arm	Sample size
Recommendation reason	Exclude, Generic reason, Customized reason	13,280, 5,775, 5,955
Maximum number of modules	1, 3, 5, 9	6,238, 6,315, 6,306, 6,151
Guide to navigate back	Exclude, Include	25,945, 25,915
Link to learning report	Exclude, Include	26,058, 25,802
Commitment question	Exclude, Question without reminder, Question with reminder	5,426, 6,498, 6,510

Table 1.4: Recommendation page features experiment: Sample sizes

1.5.1 Main effects

We find a strong negative effect (70% drop) of providing a reason for recommendation at the top of the page on the number of LPE modules completed. We document this in Figure 1.10[1]. Why might this be the case? Consistent with the interpretation that adding these features induces cognitive overload, users who were shown a reason were also 60% less likely to answer the survey question (see Figure 1.11). Nevertheless, those who did answer provided statistically indistinguishable answers from the group not shown the reason (see Figure 1.12).

In Figure 1.10, we can also see that we find null effects on completion from adding 1) a commitment question, 2) a guide to navigate back to the page, and 3) a link to the child's learning report. Using our estimated standard errors, we were able to rule out effect sizes larger than 0.016 ($\pm 13\%$), 0.016 ($\pm 13\%$), and 0.032 ($\pm 18\%$) for these treatment conditions respectively.

Also consistent with the cognitive overload interpretation, we see significantly lower rates in answering the survey question from adding these features (see Figure 1.11). Similar to the previous result, we find statistically indistinguishable answers from the control group for all of these treatments (see Figure 1.12). As such, our hypothesized positive effects from these features are likely muted or even negated by the penalties they induced through cognitive load.

Unlike the above-mentioned features, we find no evidence that increasing the number of modules

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[1]Effect sizes and standard errors estimates as OLS coefficient on the treatment variable (with feature = 1, no feature = 0) while including variables in the OLS regression for the other treatments in the factorial design.
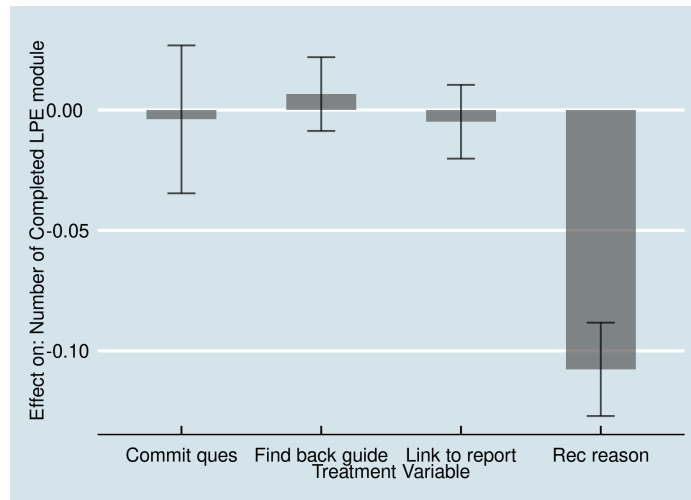
Figure 1.10: Effect of features on LPE module completion
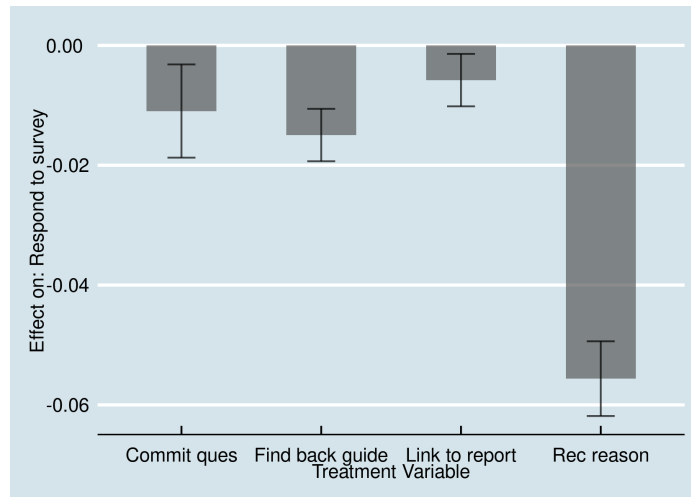
Figure 1.11: Effect of features on respond to survey

causes cognitive overload as they have statistically indistinguishable survey response rates (Figure 1.13) and affirmative answer rates (Figure 1.14)[2]. This suggests that increasing the number of homogeneous items such as LPE modules does not induce cognitive overload in the same way that the addition of other features does. It may take more cognitive capacity to read and understand features in different places on a page than to digest a greater number of similar items.

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[2]Means are computed as the OLS coefficient on the binary indicators for the number of modules (1, 3, 5, 9) without intercept while controlling for orthogonal assignments of other pieces of the factorial design.
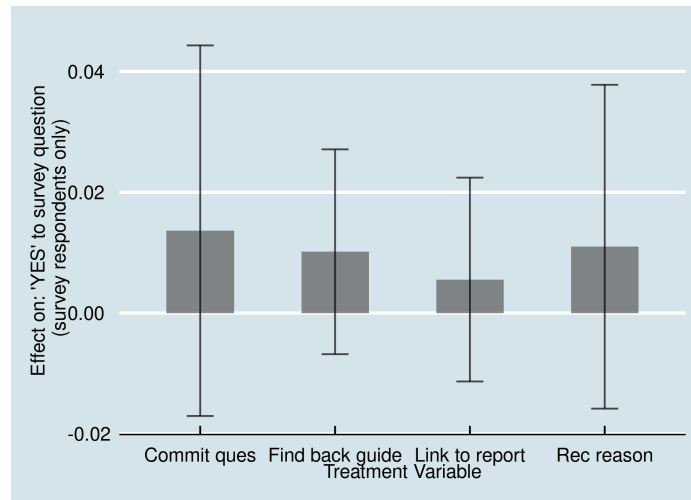
Figure 1.12: Effect of features on respond "YES" to survey (vs respond "No")
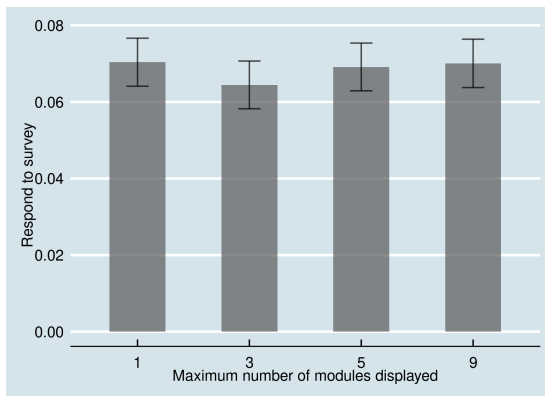
Figure 1.13: Respond to survey by number of modules
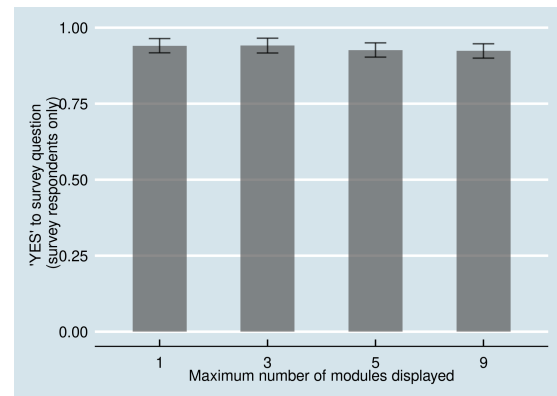
Figure 1.14: Respond "YES" to survey by number of modules

While increasing the number of recommended modules did not increase the number of users completing any modules (Figure 1.15), but increased the number of completions by compliers (Figure 1.16). This effect is likely the mechanical consequence of giving users a larger set of choices.

1.5.2 Completion patterns by number of modules

Since we have shown that increasing the number of available modules does not increase the number of users completing any modules, it must have increased completion by increasing the number completed per user. Figure 1.17 shows how the distribution of completions per user in the different
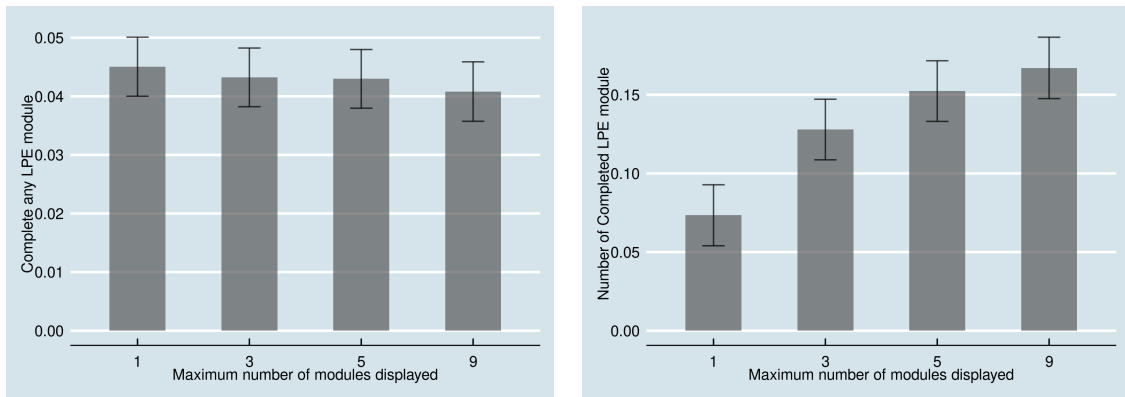
Figure 1.15: Completion of any LPE module by number of modules

Figure 1.16: LPE module completion rate by number of modules

treatment groups.

These histograms reveal two prominent patterns. Firstly, the number of users declines almost monotonically as the number of completions increases. Secondly. there is a group of users who complete the maximum number of allotted modules as evident by the spike in the left-most bin of the histogram for users assigned to the 3, 5, and 9 groups. Taken together, these patterns suggest that the distribution of user preferences seems to be continuous and decreasing with the number of completions. Increasing the number of available modules increases completions by allowing those who want to complete more to do so.
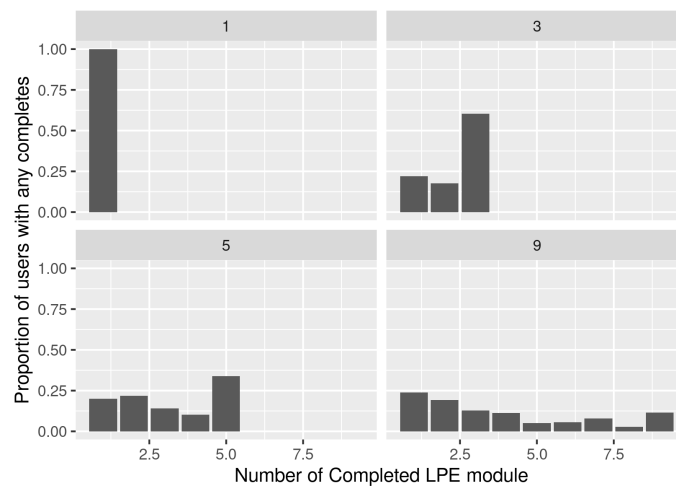
Figure 1.17: Number of modules completed by maximum number of modules available

1.5.3 Variations of features

In our experiment, the recommendation reason had two versions. The generic recommendation reason versions simply stated that the recommended LPE modules were based on the child's exam performance. The customized recommendation reason named the knowledge point that the student performed worse in and showed the percentage of questions within the knowledge point that the child had answered wrong. We see from Figure 1.18 that the effect on completion was statistically indistinguishable between the two versions of the recommendation reason.

The commitment question also had two variations. A base version asked the parent to commit to getting her child to complete the LPE modules. A reminder version also told the parent that she would be reminded. A reminder push message would be automatically sent to any parent who received the reminder version and checked the check-box. The reminder push message would be sent out at 8:30 PM on the same day the parent checked the check-box. We see from Figure 1.19 that the effect on completion was also indistinguishable between the two versions of the commitment question.
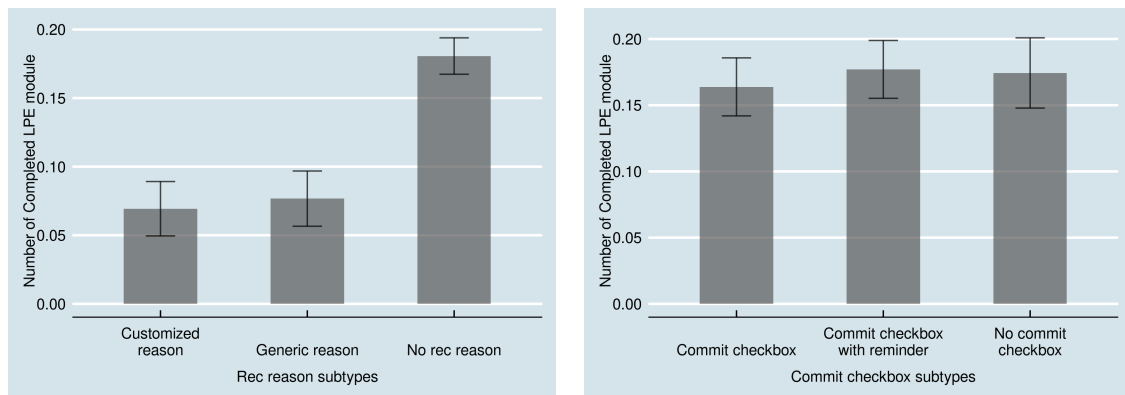
Figure 1.18: LPE module completion rate by rec-Figure 1.19: LPE module completion rate by commendation reason type   mitment question type

1.5.4 Modules vs features

We dig deeper into the differential effect on cognitive overload caused by two types of information present on the page. Having laid out the evidence in favor of a cognitive overload interpretation to our results, we seek to further examine how different types of information lead to cognitive overload.

Figure 1.20, using red boxes, highlights the two sections of the page in question. The first is the number of modules present in the middle of the page, the second is the section at the bottom of the page containing various features.

In this exercise, we want to examine how the survey response rate differs by the number of modules displayed and by the number of features available. The number of features is a count of the features available in the bottom section of the page for each user. We count the commitment question, the find back guide, and the link to the learning report. If all three features are present, then the count is 3; if only two out of the three are present, then the count is 2, and so on. Users who were assigned none of these features had a count of 0.

Figure 1.21 and Figure 1.22 shows the effects of the two types of information side by side. While increasing the number of modules shown did not decrease the survey response rate, increasing the number of features did. Low survey response rate can be a proxy for cognitive overload. As such, these results show that modules and features have different effects on cognitive overload. Perhaps because modules are homogeneous, increases in modules seem to require less cognitive energy to process than the increases in features.

Figure 1.20: Features and modules
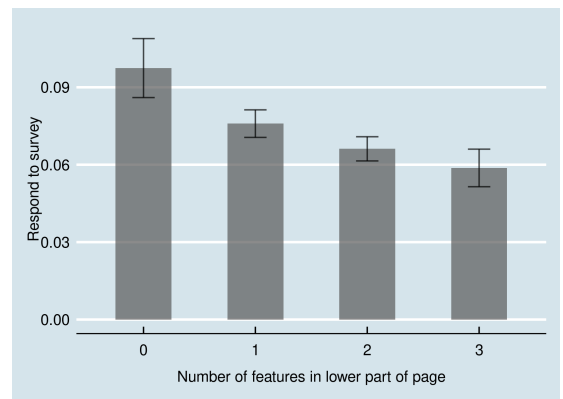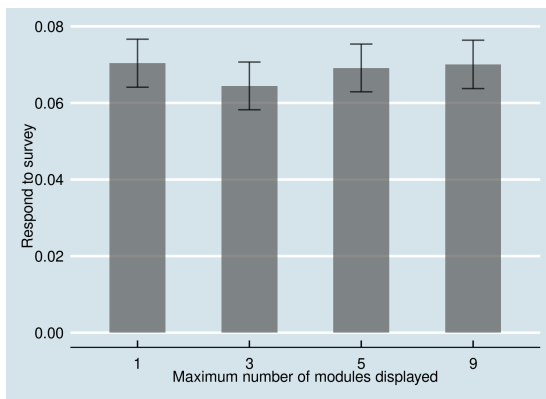
Figure 1.21: Respond to survey by number of modules

Figure 1.22: Respond to survey by number of features

1.6 Conclusion

To improve parenting through mobile apps, the parent needs to follow through and take action with their child to generate any meaningful impact on her. Our findings advance the literature on engaging parents through mobile devices. Our major contribution is demonstrating empirically that cognitive overload plays an important role in determining how much parents follow through. We also reveal that homogeneous modules do not incur cognitive overload in the same way that additional features do. As such, providing more homogeneous choices allow those who prefer to complete more to do so without introducing additional penalties on the rest of the users.

We showed that while many features may have hypothesized positive effects on parent follow through, the cognitive overload is an ever-present headwind. The hypothesized positive effects from any feature shown to the parent can be muted, or even negated, by the penalties they induce as cognitive load. Our findings should alert researchers and practitioners to think hard about the hypothesized benefits of a feature and weigh them against their potential cognitive load penalties. Above all, our results point to the importance of experimentation as it is almost impossible to establish the cognitive penalties of any feature a priori.

We also found that increasing the number of choices for homogeneous modules did not increase cognitive overload. But it increased follow through due to more choices being available. As we would expect with actions involving a cost in effort, the distribution of user preferences seems to be continuous and decreasing with the number of completions. Increasing the number of available modules increased completions by allowing those who want to complete more to do so. The implication for researchers and practitioners is that increasing the size of the choice set could induce more of the desired behavior. In our case, we did not find indications that doing so with homogeneous modules induced cognitive overload. However, our experiment capped the maximum number of modules shown at 9. These results may not hold in a setting where the choices are much more numerous than this.

Our results also revealed the modules and features seem to have a different effect on parents. Increasing the number of modules shown did not decrease the survey response rate but increasing the number of features did. Features seem to be more distracting than modules even though they occupy a smaller portion of the screen. This suggests that not everything on the screen is distracting. In

our case, features may be more distracting than modules because parents are not familiar with their functionalities and need to devote more effort to understand them. Researchers and practitioners should be aware of these differences and critically evaluate features based on the mental effort needed to understand them.

Chapter 2

The Effect of Faculty Research on Student Learning in College

2.1 Introduction

The effect of faculty research on the quality of college teaching and student learning has been the subject of intense debate for decades if not centuries (Humboldt, 1970; Newman, 2008; Feldman, 1987; Hattie and Marsh, 1996; Prince et al., 2007). Conventional wisdom argues that research complements teaching (Neumann, 1992; Smeby, 1998; Robertson and Bond, 2001). By engaging in research, faculty ostensibly provide students with up-to-date, advanced knowledge and encourage effort towards learning (Neumann, 1992). Others posit that, given trade-offs in the allocation of time between research and teaching activities, faculty research may have a negative or negligible effect on teaching quality and student learning (Fox, 1992; Linsky and Straus, 1975). Critics further argue that the higher pecuniary benefits and prestige tied with research as opposed to teaching may lead faculty to overinvest in research activities (Tuckman and Hagemann, 1976; Cech, 2003). They contend that universities need to prioritize teaching directly and not assume that a system that primarily rewards research will naturally result in high-quality teaching (Nature Editorial, 2010).[1]

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[1] These popular arguments typically do not consider complexities associated with incentivizing agents to accomplish multiple goals which may vary in terms of risk and performance measurement error.

Five decades of empirical investigation have not settled the debate (Feldman, 1987; Hattie and Marsh, 1996; Prince et al., 2007; Figlio and Schapiro, 2017). Meta-analytic reviews of the quantitative literature find a "zero correlation" between faculty research and teaching quality or student learning (Hattie and Marsh, 1996). Past studies are limited, however, in that they examine the relationship using subjective or non-standardized measures of teaching quality and student learning (Hattie and Marsh, 1996; Galbraith and Merrill, 2012). Most measures of student learning are self-reported, subject to bias (Galbraith and Merrill, 2012; Shevlin et al., 2000; Felton* et al., 2004), and uncorrelated with more distal, long-run measures of learning (Carrell and West, 2010). Another fundamental problem is that studies do not use causal identification strategies to estimate the impact of faculty research on student learning. In particular, most studies do not account for a classic selection bias problem in the economics and education literatures: the non-random sorting of students across different instructors (Hanushek and Rivkin, 2010). A third problem is that prior studies usually focus on one institution and lack generalizability (Figlio and Schapiro, 2017).

The purpose of this study is to provide more compelling and generalizable causal evidence on this long-standing debate. We examine the effect of faculty research on college student learning using matched student-faculty data that we collected from nationally representative samples of undergraduate students and faculty in STEM majors in China, India, and Russia. Not only do China, India, and Russia offer three separate national contexts in which to estimate effects, but they together also train approximately half of the STEM graduates in the world. Universities in these three countries are furthermore increasingly emphasizing and investing in faculty research in addition to teaching (Carnoy et al., 2013; Loyalka et al., 2014).

To conduct the study, we collected and analyzed nationally representative, standardized assessment data on the academic skills of 4,785 3rd-year STEM undergraduates (broadly defined computer science and electrical engineering related majors which comprise approximately 25-30% of all STEM majors) from China (802), India (3,403), and Russia (580). We used strict sampling procedures to randomly select institutions and students in China, India, and Russia. By paying close attention to survey implementation, we also achieved high total response rates among institutions (98%), faculty (85%) and enrolled students (94%). Our academic skills tests were specifically created to assess math and science (physics) skill levels among the computer science (CS) and electrical engineering (EE)

related majors across the three countries. The math and science tests each contained 35 items and lasted for 40 minutes and were given in random order to account for test fatigue. These items were piloted and selected for their reliability, uni-dimensionality, validity, scalability, and cross-national comparability (see Kardanova et al. (2016)).

Similar to a number of studies in the economics and education literatures, we utilize a cross-subject student fixed effects strategy to identify effects (Dee, 2005, 2007; Clotfelter et al., 2010; Kingdon and Teal, 2010; Lavy, 2015; Rivkin and Schiman, 2015; Zakharov et al., 2016). Using this strategy, we find that, in contrast to prior studies, faculty research tends to negatively impact student achievement. One additional faculty publication per year decreases student achievement by 0.044, 0.020, and 0.067 standard deviations (SDs) in China, India, and Russia. Other measures of faculty research such as time spent on research also suggest negative impacts. When compared with instructor value-added studies which typically find small effects from faculty quality (Bettinger and Long, 2010; Carrell and West, 2010; Hoffmann and Oreopoulos, 2009), the results suggest meaningful tradeoffs between faculty research and the achievement of STEM undergraduate students.

Besides having policy implications, our study contributes to the education and broader social sciences literature in several ways. Methodologically, the study is perhaps the first to use standardized assessments to describe the relationship between faculty research and student achievement. The vast majority of studies which use subjective measures of achievement likely have inflated standard errors (and potential bias) due to excessive (and perhaps non-random) measurement error in the dependent variable. In addition, the study is one of the first to use a quasi-experimental causal identification strategy to estimate effects. Perhaps equally as important, the study uniquely shows consistent, generalizable results across a large number of institutions—specifically, three major higher education systems that differ greatly in their institutional features (Carnoy et al., 2013). This gives us more confidence that our findings speak to a common phenomenon in higher education globally. Finally, our results speak to important tradeoffs in the organization of higher education in nations that train a large proportion of the world's STEM undergraduates.

2.2 Data and Methods

2.2.1 Sampling and Data Collection

Data were collected from nationally representative samples of Electrical Engineering (EE) related programs in China, India, and Russia and nationally representative samples of Computer Science (CS) related programs in India and Russia. In China, we first chose five provinces that represented China's northern (Beijing and Shandong), central (Henan and Shaanxi), and southern (Sichuan) regions as well as its economically more developed (Beijing ranked 2nd in GDP per capita out of 31 provincial-level administrative divisions), mid-developed (Shandong and Shaanxi ranked 10th and 14th) and less developed (Sichuan and Henan ranked 22nd and 24th) regions (National Bureau of Statistics of China, 2015). From each of the five provinces, we took a simple random sample of 6 universities from a list of all universities in each province that offered bachelor's degree programs in CS and EE. In India, we took a simple random sample of 8 elite colleges and 42 non-elite colleges that offered bachelor's degree programs in CS and EE. In Russia, we took a simple random sample of 34 universities from a list of all universities nationwide that offered bachelor's degree programs in CS and EE. We further limited the sample to universities with CS and EE programs that taught math and science courses to students.[2] Altogether, we sampled 802 EE students from 28 public four-year universities in China[3], 3,403 CS and EE students from 50 four-year colleges in India, and 580 CS and EE students from 29 public four-year universities in Russia. In addition, we surveyed the faculty that taught students math and science (physics) courses in the first two years of university.

Students were surveyed and tested in late November and early December 2015 in Russia and China and in October 2017 in India. These dates correspond to the beginning of the third year of their undergraduate studies. Trained enumerators administered valid and reliable math and science (physics) exams to students in a timed and controlled environment (Kardanova et al., 2016).[4] We standardized exam scores using the mean and standard deviation (SD) of scores within each country and subject separately. After the exams were completed, students provided information about their age, gender, major, and socioeconomic background (Appendix Table 2.6).

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[2]This excluded 5 universities in Russia that taught informatics to students in lieu of science courses.

[3]CS students were not universally taught science (physics) courses in China. As such, we exclude them from this analysis.

[4]Math and science (physics) courses are considered foundational courses for computer science and electrical engineering students in the three countries and are taken during the first two years.

Math and science (physics) faculty were surveyed near the same time as students. We asked faculty to report the number of publications (defined as the total number of published academic international journal articles, domestic journal articles, monographs, and edited volumes) from the last three years. Following the literature Hattie and Marsh (1996), we used the average number of publications per year as our primary measure of faculty research. We also asked the faculty to report the number of hours they spend, in a typical work week, on teaching classes, advising students, research, consulting, and administrative tasks. We construct a measure of faculty's time spent on research by dividing their reported research hours on the sum of hours from all the categories of work. Details and summary statistics for the indicator are located in the Appendix (Appendix Table 2.7).

A range of faculty professional background characteristics were collected to serve as controls in our analyses. Faculty reported their age, gender, and educational background. They also reported their academic title, whether they were full or part time faculty, and years of teaching experience at the university level. Based on faculty responses, we created dummy variables for whether faculty attended elite universities for their terminal degree program, and whether they had a terminal degree major in the subject (math or science) they were teaching. We further broke down the years of teaching variable into categorical dummies to allow for potential nonlinearities in the relationship between teaching experience and student achievement (Clotfelter et al. (2007), see Appendix Table 2.7).

We matched students to their respective math and science instructors. Following the economics literature, we created a composite faculty vector for each student that averaged all of the student's instructor characteristics weighted by the respective credit hours allocated to each instructor's course (Bettinger and Long, 2005, 2010). Each student was thus observed twice in our dataset: once with a math score and composite math faculty vector and once with a science score and composite science faculty vector.

2.2.2 Statistical Approach

Our main specification was a cross-subject student fixed effects model that exploits variation across math and science subjects within students. Cross-subject student fixed effects models minimize

potential bias induced by the non-random assignment of students across classes and schools. We are thus able to better identify causal effect estimates by only utilizing the variation that exists across different subjects (e.g. math and science) within the same student (Angrist and Pischke, 2008). Cross-subject student fixed effects models are frequently used in the econometric literature (e.g. Dee (2005, 2007); Clotfelter et al. (2010); Kingdon and Teal (2010); Lavy (2015); Rivkin and Schiman (2015)).

In our model, we compare outcomes within the same student in different subjects, controlling for other factors that may differ across these subjects. By utilizing this variation, the model controls for all factors which are constant across subjects and which may confound the relationship between faculty research and student learning outcomes. This includes unobserved student-level factors such as student innate ability and motivation. It also includes all administrative class and school level factors that do not vary across subjects, thus accounting for potential non-random assignment of students to administrative classes and schools. To further eliminate potential sources of selection bias, we controlled for a number of faculty characteristics that differed across math and science within each student. We also accounted for preexisting variation in cross-subject achievement by controlling for the average math and science achievement levels of cohorts of entering freshman in the same department.[5]

In our specific context, the cross-subject student fixed effects model takes the first difference between math and science subjects (scores and corresponding determinants) within each student. By doing so, the specification eliminates the confounding influence of unobserved subject-invariant factors. The specification is written as follows:

$$(y_{pi} - y_{mi}) = \beta(R_{pi} - R_{mi}) + \lambda(Z_{pi} - Z_{mi}) + \nu(X_i - X_i) + \gamma(A_{pi} - A_{mi}) + (\mu_i - \mu_i) + (\epsilon_{pi} - \epsilon_{mi}) \quad (2.1)$$

where y_{mi} and y_{pi} are a student's standardized scores in math and science. Similarly, R_{mi} and R_{pi} are the indicators for faculty research for math and science instructors. Z_{mi} and Z_{pi} consists of

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[5]Another option to account for pre-existing variation in cross-subject achievement is to use the self-reported college entrance scores in math and science we collected through our student questionnaires. These entrance exams were only taken by a subset of the students in our sample. As such, we added these controls to our specification and ran the estimator on this subset of students. The results from these robustness tests were substantively the same in magnitude and significance as those reported here.

observed faculty-level determinants of student scores for math and science. X_i is a vector of student characteristics that vary across students but not across subjects. A_{mi} and A_{pi} represent baseline achievement levels in math and science. μ_i is an unobserved student error term that does not vary between subjects. Finally, ϵ_{pi} and ϵ_{mi} are unobserved student-level error terms that vary by subject.

Conditional on a few assumptions, the student fixed effects model produces unbiased estimates of the effect of faculty research engagement (β). The first assumption is that the way in which faculty research affects student performance is similar across both math and science subjects. Under this assumption, non-subject specific student, administrative class, and school level characteristics (both observed and unobserved as captured by X_i and μ_i) are canceled out and do not induce bias in the estimation of β. We also add a subject (science or math) dummy variable to equation (1) to control for the situation in which faculty research influences student achievement more in one subject than in the other. The second assumption is that the remaining error term $\epsilon_{pi} - \epsilon_{mi}$ in equation (1) is uncorrelated with the faculty research($R_{pi} - R_{mi}$). This assumption implies that unobserved student or faculty characteristics that vary across the two subjects should not be jointly correlated with faculty research and student achievement (Schwerdt and Wuppermann, 2011). To account for potential violations of this assumption, we control for subject-specific mean scores of year 1 students from the same administrative class (in the case of China and Russia) or department (in the case of India). These serve as proxies for the preexisting variation in achievement levels across subjects ($A_{pi} - A_{mi}$). We also control for a number of faculty-level cross-subject factors($Z_{pi} - Z_{mi}$). These factors are documented in the faculty level summary statistics in Appendix Table 2.5.

As a robustness check, and for the sake of comparability with the existing literature, we also estimate the following ordinary least squares (OLS) model:

$$y_{si} = \beta R_{si} + \lambda Z_{si} + \nu X_{si} + \gamma A_{si} + \mu_i + \epsilon_{si} \tag{2.2}$$

Equation 2 yields causal estimates of the effect of faculty research on student achievement under the assumption of strong ignorability (Imbens and Rubin, 2015). Namely, if y_{si} and R_{si}, conditional on a rich set of student, class and faculty background characteristics including baseline test scores (as captured by A_{si} and Z_{si}, which includes the subject (science or math) dummy variable), are uncorrelated with the combined error term $\mu_i + \epsilon_{si}$, estimates of β capture the causal effect of

faculty research on student achievement. Unlike in the cross-subject fixed effects specification, in this value-added specification, unobserved factors such as student motivation may still confound the relationship between faculty research and student achievement.

2.3 Results

Our results show that faculty publication rate has a negative and statistically significant effect on student achievement. One additional faculty publication per year is associated with 0.044, 0.020, and 0.067 standard deviation decrease in their students' achievement in China, India, and Russia respectively (Table 2.1).[6] Not all publications are of the same quality and may demand different amount of effort on the part of the faculty. One concern, specific to our context, is that the countries we examine have diverse domestic venues for academic publication that may be of varying quality. We attempt to mitigate the noise from publication quality by repeating our analyses using only publications in international academic journals. We find substantively the same results, the point estimate shows an additional international publication per year by the faculty leads to a 0.025, 0.013, and 0.133 standard deviation decrease in students' achievement in China, India, and Russia respectively. However, in this case the point estimate for China is no longer statistically significant (Table 2.2).

When we examine use the much noisier measure of self-reported time spent on research, the effects are still negative but not always statistically significant (Table 2.3). A 10 percentage point increase in time spent on research leads to a 0.069 SD decrease in student achievement in Russia (statistically significant at the 5% level). The estimates are negative for China and India but not statistically significant. Given the strong likelihood of substantial measurement error in self-reporting time spent on research as opposed to the publications measure, estimates from the former are not only attenuated towards zero but also have larger standard errors (Zimmerman, 1992).

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[6]The reported R2 statistic only reflects the proportion of within-student variance explained. It does not report the variance reduction induced by adding the student fixed effects.

2.3.1 Robustness Checks

We run two important robustness checks. First, to account for noise from a long tail in faculty publication rates (as many faculty produce few publications and some produce many), we reran the results in Table 2.1 using a log-transform of the faculty publication rate. We find substantively the same results for faculty publications with and without a log transform (Appendix Table 2.4). In addition to our preferred student fixed effects specification which largely accounts for student-teacher sorting, we run OLS models that control for both student and faculty characteristics. The estimates from the OLS model are smaller in magnitude and not always statistically significant but are largely consistent with the results from the student fixed effect model (Appendix Table 2.5). Prior studies which relied on OLS specifications may have failed to detect negative effects due to unresolved omitted variables bias at the student level.

2.4 Conclusion

A strong foundation in STEM is believed to be important for national economic productivity and innovation (on Prospering in the Global Economy of the 21st Century, 2007). Policymakers and researchers from around the world have therefore made teaching and research in STEM majors a top priority (on Prospering in the Global Economy of the 21st Century, 2007; Lucena et al., 2008). This is especially true in countries such as China, India, and Russia where the sheer numbers of undergraduate enrolments in STEM majors are large compared to the United States (Carnoy et al., 2013). However, STEM higher education institutions are increasingly tasked with the dual objectives of producing research and learning. Since both of these outputs have important consequences for economic productivity and competitiveness, understanding whether they represent competing or complimentary objectives has major policy implications for student learning and the direction of higher education around the world.

Although smaller than the significant effect sizes typically reported in studies of pre-tertiary education, the estimated effect sizes reported in this study are non-trivial in the context of higher education. Multiple studies show that a one standard deviation increase in faculty value-added increases student achievement by 0.05 to 0.06 SDs (Carrell and West, 2010; Hoffmann and Oreopoulos,

2009). Furthermore, even our results may be underestimates as one disadvantage of student fixed effects specifications is that they remove substantial variation in the dependent variable (Angrist and Pischke, 2008).

On the whole, our results provide the first systematic evidence that, in STEM disciplines at a wide range of universities in three major countries, there is a significant and non-trivial tradeoff between research and teaching. Policymakers and university administrators that are concerned about the tradeoff may wish to more closely monitor how to balance research and teaching activities within their institutions. On the one hand, they may wish to readdress the distribution of faculty incentives between teaching and research in an effort to find an optimal tradeoff. On the other hand, they may wish to continue the trend of hiring faculty that specialize in teaching or research, with the caveat that the presence of faculty that engage in both areas may serve other purposes such as attracting prospective students and faculty (Wood, 1973).

	Student Fixed Effects Model (China)	Student Fixed Effects Model (Russia)	Student Fixed Effects Model (India)
Number of publications per year	-0.044*	-0.067***	-0.020***
	(0.024)	(0.019)	(0.006)
Science subject dummy (opposed to math)	0.040	-0.199***	-0.000
	(0.048)	(0.064)	(0.019)
Achievement levels of entering freshman (math and science)	0.259*	0.630***	0.216***
	(0.142)	(0.142)	(0.081)
Faculty level variables			
Female	-0.147*	-0.132	0.020
	(0.078)	(0.089)	(0.045)
Fulltime faculty	-0.062	-0.016	0.525***
	(0.287)	(0.127)	(0.106)
Faculty has PhD degree	0.075	0.482	-0.025
	(0.102)	(0.313)	(0.048)
Full Professor (compared to Lecturer or RA)	-0.090	-0.171	-0.103
	(0.143)	(0.332)	(0.074)
Associate Professor (compared to Lecturer or RA)	-0.025	-0.645*	-0.044
	(0.121)	(0.331)	(0.058)
Went to elite university for highest degree obtained	-0.038	0.000	0.072
	(0.119)	(.)	(0.050)
Majored in math or science for highest degree	-0.053	0.059	-0.046
	(0.097)	(0.123)	(0.046)
Number of years teaching (middle tercile in country)	0.079	-0.515***	0.034
	(0.123)	(0.127)	(0.044)
Number of years teaching (upper tercile in country)	0.113	-0.162	-0.019
	(0.133)	(0.123)	(0.054)
Observations	1570	1098	6414
R2	0.019	0.114	0.019

Notes:
1. Standard errors in parentheses.
2. The R2 reported denotes the proportion of within-student variance explained. The total proportion of (within- and across-student) variance explained is above 0.75 for all 3 models.

3. Statistical significance: * p <0.1, ** p <0.05, *** p <0.01

Table 2.1: The effect of faculty publications (per year) on student achievement

	Student Fixed Effects Model (China)	Student Fixed Effects Model (Russia)	Student Fixed Effects Model (India)
Number of international journal publications per year	-0.025 (0.034)	-0.133*** (0.034)	-0.013* (0.007)
Science subject dummy (opposed to math)	0.043 (0.049)	-0.208*** (0.068)	-0.009 (0.020)
Achievement levels of entering freshman (math and science)	0.260* (0.143)	0.609*** (0.141)	0.212** (0.084)
Faculty level variables			
Female	-0.148* (0.079)	-0.038 (0.092)	0.044 (0.046)
Fulltime faculty	-0.064 (0.312)	0.053 (0.132)	0.466*** (0.107)
Faculty has PhD degree	0.097 (0.110)	0.648** (0.312)	-0.031 (0.049)
Full Professor (compared to Lecturer or RA)	-0.142 (0.147)	-0.438 (0.324)	-0.005 (0.076)
Associate Professor (compared to Lecturer or RA)	-0.054 (0.127)	-0.833** (0.333)	-0.065 (0.060)
Went to elite university for highest degree obtained	-0.039 (0.122)	0.000 (.)	0.099* (0.051)
Majored in math or science for highest degree	-0.032 (0.097)	0.083 (0.129)	-0.015 (0.047)
Number of years teaching (middle tercile in country)	0.081 (0.128)	-0.506*** (0.125)	0.051 (0.045)
Number of years teaching (upper tercile in country)	0.124 (0.150)	-0.080 (0.124)	-0.025 (0.054)
Constant	0.039 (0.329)	1.030*** (0.220)	-0.427*** (0.103)
Observations	1518	1016	6078
R2	0.015	0.135	0.016

Notes:
1. Standard errors in parentheses.
2. The R2 reported denotes the proportion of within-student variance explained. The total proportion of (within- and across-student) variance explained is above 0.75 for all 3 models.

3. Statistical significance: * p <0.1, ** p <0.05, *** p <0.01

Table 2.2: The effect of faculty publications in international journals (per year) on student achievement

	Student Fixed Effects Model (China)	Student Fixed Effects Model (Russia)	Student Fixed Effects Model (India)
Number of international journal publications per year	-0.025 (0.034)	-0.133*** (0.034)	-0.013* (0.007)
Science subject dummy (opposed to math)	0.043 (0.049)	-0.208*** (0.068)	-0.009 (0.020)
Achievement levels of entering freshman (math and science)	0.260* (0.143)	0.609*** (0.141)	0.212** (0.084)
Faculty level variables			
Female	-0.148* (0.079)	-0.038 (0.092)	0.044 (0.046)
Fulltime faculty	-0.064 (0.312)	0.053 (0.132)	0.466*** (0.107)
Faculty has PhD degree	0.097 (0.110)	0.648** (0.312)	-0.031 (0.049)
Full Professor (compared to Lecturer or RA)	-0.142 (0.147)	-0.438 (0.324)	-0.005 (0.076)
Associate Professor (compared to Lecturer or RA)	-0.054 (0.127)	-0.833** (0.333)	-0.065 (0.060)
Went to elite university for highest degree obtained	-0.039 (0.122)	0.000 (.)	0.099* (0.051)
Majored in math or science for highest degree	-0.032 (0.097)	0.083 (0.129)	-0.015 (0.047)
Number of years teaching (middle tercile in country)	0.081 (0.128)	-0.506*** (0.125)	0.051 (0.045)
Number of years teaching (upper tercile in country)	0.124 (0.150)	-0.080 (0.124)	-0.025 (0.054)
Constant	0.039 (0.329)	1.030*** (0.220)	-0.427*** (0.103)
Observations	1518	1016	6078
R2	0.015	0.135	0.016

Notes:
1. Standard errors in parentheses.
2. The R2 reported denotes the proportion of within-student variance explained. The total proportion of (within- and across-student) variance explained is above 0.75 for all 3 models.

3. Statistical significance: * p <0.1, ** p <0.05, *** p <0.01

Table 2.3: The effect of research time spent by faculty on student achievement

2.5   Appendix

2.5.1   Sampling and Implementation Details

The national samples of universities represented the range of elite and non-elite institutions in each country. In China, "elite" universities were defined on the basis of their designation as Project 985 or 211 universities. Project 985 universities are the top 39 universities in China (Ministry of Education of the People's Republic of China, b) whereas Project 211 universities are the top 112 universities in China (Ministry of Education of the People's Republic of China, a), and both have received preferential national government funding. In India, "elite" institutions were defined as the India Institutes of Technology (IITs), the Indian Institutes of Information Technology (IIITs), the National Institutes of Technology (NITs), and other institutions that ranked in the top 100 of the National Institutional Ranking Framework (NIRF) rankings for 2016 developed by the Ministry of Human Resource Development, Government of India. In Russia, "elite" universities were defined as National Research Universities, "5-100" universities, and Federal universities. Altogether, we sampled 6 elite and 22 non-elite universities in China, 8 elite and 42 non-elite colleges in India, as well as 6 elite and 23 non-elite universities in Russia.

We next randomly sampled students within CS and EE departments within the sample universities. We first identified EE departments in China.[7] After identifying the EE departments, we randomly sampled 2 EE departments from each sample university.[8] We then randomly sampled 1 administrative group or "class" of year 3 (junior year) students within each department. All students within the sampled classes were selected for participation in the study.

In India, we randomly sampled 1 CS department and 1 EE department at each sampled college.[9] All year 3 (junior year) students within the sampled departments were selected for participation in the study.

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[7]The EE departments in China included Electrical Engineering, Electrical Engineering and Automation, Measurement and Control Technology and Instrumentation, Electronic Science and Technology, Electronic and Information Engineering, Electronic and Information Science and Technology, Communication Engineering, Optoelectronic Information Science and Engineering, Microelectronics Science and Engineering), and Automation.

[8]If there was only one EE department in the university, we sampled only that department.

[9]In India, the CS departments included Computer Engineering, Computer Science Engineering, Information Science and Engineering, and Information Technology. The EE departments in India included Electrical Engineering, Electronics and Communication Engineering, Electronics and Electrical Engineering, Electronics and Instrumentation Engineering, and Electronics and Telecommunications Engineering.

Similarly, we identified CS departments and EE departments in Russia.[10] After identifying the departments, we randomly sampled 3 CS departments and 3 EE departments from each university.[11] We then randomly sampled 1 administrative group or class of year 3 (junior year) students within each department.[12] All students within the sampled classes were selected for participation in the study.

The survey was conducted in late November and early December of 2015 in China and Russia, and in late October and early November of 2016 in India. Within each administrative class (or department in the case of India), one half of the sampled students were randomly chosen to take two closely proctored exams: math and science. The math and science exams each contained 35 items and lasted for 40 minutes. The different subject exams were given to examinees in random order to avoid fatigue-induced bias in exam results. Details about the construction and suitability of the math and science exams can be found in Kardanova et al. (2016).

After the exams were completed, students responded to a short questionnaire. Students were asked about their age, gender, major, parental education attainment, and several items of value in the home. We used the information on household and polychoric principal components analysis to create a proxy for wealth or socioeconomic status (Kolenikov and Angeles, 2009). Student participation rates for both the questionnaire and the exams were extremely high (95% for China, 95% for India, and 87% for Russia). Summary statistics for the student level variables are presented in Appendix Table 2.4.

2.5.2 Faculty Data

We collected faculty data in a follow up online survey after the main survey of students. All EE students in China and Russia, a subset of CS students in Russia, and all CS and EE students in India are expected to take mandatory courses in math and science in their first two years of

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[10]The CS departments in Russia include IT and Computer Facilities, Informational Systems and Technologies, Applied IT, Program Engineering, Mathematics and Computer Science, Fundamental IT and Information Technologies, Software and Administration of Information Systems, Information Security. The EE departments in Russia include Information and Communication Technology and Communication Systems, Design and Technology of Electronic Instrumentation, Radio Engineering, Electronics and Nanoelectronics, Electrical Power and Electric Engineering, Laser Equipment and Laser Technologies, Optics Engineering, Instrument Construction, Photonics and Optoinformatics.

[11]If there were less than three EE departments in the university, we sampled all departments.

[12]In China and Russia, we also sampled one class of year 1 (freshman) students from each department. In India, we also sampled all year 1 (freshman) students from each department. The test scores for these freshman students served as control variables in our analyses.

university (or college for India). By cross-referencing administrative records and a section of our student questionnaire where students wrote down the names of their math and science instructors, we compiled a list of all faculty that had taught math and science courses to the sampled students in the previous years of college. We subsequently asked faculty members to answer a questionnaire. Faculty reported their gender, academic title, education background, whether they worked full time or part time, whether they had a PhD degree, and how much they published in the past 3 years. Summary statistics for these faculty level variables are in Appendix Table 2.7.

We then matched students to their respective math and science instructors. In almost all cases, students took multiple courses in math and science and were therefore matched to multiple math and science instructors. Following the economics literature, we created one synthetic faculty member for each student for each subject (Bettinger and Long, 2005, 2010). The synthetic faculty member for math averaged all of the student's math instructor characteristics weighted by the respective credit hours allocated to each instructor's course. Similarly, the synthetic faculty member for science averaged all of the student's science instructor characteristics weighted by the respective credit hours allocated to each instructor's course. These operations resulted in a dataset where each student was observed twice: once for math (with variables for math achievement and math faculty characteristics) and once for science (with variables for science achievement and science faculty characteristics).

	Student Fixed Effects Model (China)	Student Fixed Effects Model (Russia)	Student Fixed Effects Model (India)
Number of publications per year (log)	-0.132**	-0.167**	-0.026
	(0.066)	(0.072)	(0.030)
Science subject dummy (opposed to math)	0.045	-0.199***	0.004
	(0.048)	(0.064)	(0.019)
Achievement levels of entering freshman (math and science)	0.236*	0.596***	0.244***
	(0.142)	(0.144)	(0.081)
Faculty level variables			
Female	-0.139*	-0.118	0.028
	(0.078)	(0.090)	(0.045)
Fulltime faculty	-0.080	-0.022	0.509***
	(0.287)	(0.128)	(0.106)
Faculty has PhD degree	0.064	0.445	-0.035
	(0.102)	(0.315)	(0.051)
Full Professor (compared to Lecturer or RA)	-0.088	-0.189	-0.047
	(0.142)	(0.337)	(0.073)
Associate Professor (compared to Lecturer or RA)	-0.026	-0.583*	-0.032
	(0.121)	(0.333)	(0.058)
Went to elite university for highest degree obtained	-0.027	0.000	0.077
	(0.118)	(.)	(0.050)
Majored in math or science for highest degree	-0.050	0.088	-0.039
	(0.097)	(0.125)	(0.046)
Number of years teaching (middle tercile in country)	0.087	-0.507***	0.036
	(0.122)	(0.128)	(0.044)
Number of years teaching (upper tercile in country)	0.133	-0.146	-0.045
	(0.130)	(0.124)	(0.054)
Constant	0.115	1.178***	-0.444***
	(0.324)	(0.223)	(0.103)
Observations	1570	1098	6414
R2	0.019	0.104	0.016

Notes:

1. Standard errors in parentheses.

2. The R2 reported denotes the proportion of within-student variance explained. The total proportion of (within- and across-student) variance explained is above 0.75 for all 3 models.

3. Statistical significance: * p <0.1, ** p <0.05, *** p <0.01

Table 2.4: The effect of log-transformed faculty publications (per year) on student achievement

	OLS	OLS	OLS
	(China)	(Russia)	(India)
Number of publications per year	-0.019	-0.067***	-0.012*
	(0.027)	(0.023)	(0.006)
Science subject dummy (opposed to math)	0.041	-0.199***	0.003
	(0.060)	(0.077)	(0.021)
Achievement levels of entering freshman (math and science)	0.233	0.630***	0.284***
	(0.179)	(0.172)	(0.087)
Student level variables			
Age	-0.056***	-0.026	-0.012***
	(0.021)	(0.021)	(0.004)
Female	0.028	-0.163**	-0.062***
	(0.049)	(0.065)	(0.023)
Father had gone to college	-0.012	0.087	-0.028
	(0.078)	(0.059)	(0.025)
Mother had gone to college	0.064	0.051	0.122***
	(0.088)	(0.068)	(0.026)
Household assets index	-0.019	-0.034	0.029***
	(0.020)	(0.033)	(0.008)
Faculty level variables			
Female	-0.102	-0.132	0.004
	(0.093)	(0.108)	(0.038)
Fulltime faculty	0.160	-0.016	0.324***
	(0.347)	(0.153)	(0.098)
Faculty has PhD degree	0.027	0.482	0.024
	(0.124)	(0.378)	(0.046)
Full Professor (compared to Lecturer or RA)	-0.090	-0.171	-0.090
	(0.171)	(0.400)	(0.070)
Associate Professor (compared to Lecturer or RA)	0.030	-0.645	-0.023
	(0.144)	(0.399)	(0.052)
Went to elite university for highest degree obtained	0.003	0.000	0.050
	(0.139)	(.)	(0.046)
Majored in math or science for highest degree	-0.027	0.059	-0.003
	(0.111)	(0.148)	(0.041)
Number of years teaching (middle tercile in country)	0.132	-0.515***	0.009
	(0.144)	(0.153)	(0.041)
Number of years teaching (upper tercile in country)	0.210	-0.162	-0.000
	(0.159)	(0.148)	(0.048)
Constant	0.928	1.770***	-0.077
	(0.610)	(0.540)	(0.128)
Observations	1544	1098	6732
R2	0.380	0.392	0.359

Notes:
1. Standard errors in parentheses.
2. Results control for class fixed effects.
3. Statistical significance: * p <0.1, ** p <0.05, *** p <0.01

Table 2.5: The effect of faculty publications per year on student achievement (OLS regression controlling for the full set of covariates)

	China			Russia			India		
	N	Mean	St. Dev.	N	Mean	St. Dev.	N	Mean	St. Dev.
Math subject test score	802	0.001	1.000	580	0.000	1.000	3403	0.000	1.000
Science subject test score	802	0.000	1.000	580	0.000	1.001	3403	0.000	1.000
Mean math score of entering freshman student from the same admin class / department	785	0.429	0.450	549	-1.082	0.304	3402	-0.012	0.650
Mean science score of entering freshman student from the same admin class / department	785	0.207	0.503	549	-0.993	0.384	3402	-0.012	0.715
Age	789	22.336	1.099	580	21.572	1.276	3392	19.508	2.375
Female	802	0.296	0.457	580	0.233	0.423	3403	0.393	0.489
Father has college education	802	0.182	0.386	580	0.681	0.466	3403	0.520	0.500
Mother has college education	802	0.138	0.346	580	0.793	0.405	3403	0.399	0.490
Household asset index	800	-0.608	1.251	579	0.823	0.821	3371	-0.026	1.590

Table 2.6: Student level summary statistics

	China			Russia			India		
	N	Mean	Std. Dev.	N	Mean	Std. Dev.	N	Mean	Std. Dev.
Number of publications per year	277	1.721	2.675	342	2.141	3.205	752	1.469	2.857
Number of international journal articles per year	271	1.050	2.254	335	0.797	2.125	701	1.381	3.615
Research time (units of 10% of working time per week)	280	2.470	1.839	341	1.551	1.632	720	2.468	1.754
Female	283	0.834	0.373	351	0.795	0.404	777	0.605	0.489
Full-time faculty	283	0.466	0.500	351	0.422	0.495	776	0.372	0.484
Faculty has PhD degree	283	0.965	0.185	345	0.864	0.344	769	0.982	0.134
Full Professor	283	0.371	0.484	350	0.854	0.353	768	0.337	0.473
Associate Professor	283	0.141	0.349	349	0.195	0.397	768	0.074	0.262
Assistant Professor, Lecturer, or RA	283	0.399	0.491	349	0.650	0.478	768	0.178	0.383
Went to elite university for highest degree obtained	283	0.459	0.499	349	0.155	0.362	768	0.747	0.435
Majored in math or science for highest degree	283	0.700	0.459	351	0.000	0.000	767	0.300	0.458
Number of years teaching (bottom tercile in country)	282	0.635	0.482	309	0.689	0.464	767	0.432	0.496
Number of years teaching (middle tercile in country)	282	0.262	0.441	349	0.309	0.463	768	0.316	0.465
Number of years teaching (upper tercile in country)	282	0.422	0.495	349	0.332	0.472	768	0.362	0.481

Table 2.7: Faculty level summary statistics

Chapter 3

Modeling Student Learning in Large-Scale Online Settings

3.1 Introduction

Online education has been transformed in recent years by the rise of mobile learning. Around the world, popular mobile learning apps offer students personalized paths of engagement. In India, the tutoring app Byju's personalizes students' learning journeys using a large knowledge graph (Bhatia, 2017). The US-based Khan Academy offers teachers the option of personalizing student assignments (Khan Academy, 2020b) and also allows students to choose their own pathway through practice exercises (Khan Academy, 2020c). Similar to Khan Academy, 17Zuoye offers choices to teachers and students in China (Sunny Education Inc., 2018). Together, these three apps account for more than 180 million users (Singh, 2020; Khan Academy, 2020a; Sunny Education Inc., 2018), and many other apps offer similar features to their users.

Choice over one's learning path is heralded as an essential part of a broader effort to improve learning through personalization. While this movement towards personalized learning is received with great fanfare in the online education industry, they also created important challenges for documenting student progress. Compared to the traditional setting of standardized tests, students are

not exposed to the same items in personalized learning paths. Typically, students are only exposed to a small fraction of items from a large question bank. Furthermore, new interaction data is generated in real-time and new students and new items are frequently added to the system.

As an example, in 17Zuoye's database, there are 1.5 million unique questions exposed to students in December of 2017. However, among active users, the median student is only exposed to 398 questions over the same month. The high-frequency user at 95 percentile only logs 1115 questions in the same period.

Traditionally, education measurement tools based on Item response theory (IRT) are designed for standardized tests with dense data. They are unable to scale to this setting where the student-item matrix is large and sparse. The literature on linking has procedures for accommodating data sparsity. But existing methods are ad hoc, heavily dependent on model specification, and do not computationally scale to large data.

In this paper, we propose new models based on latent factorization and Bayesian variational inference to address these challenges. We find that our models retrieve true parameters with greater fidelity than traditional models in small data settings (section 3.5). In addition, our models also scale well computationally to industrial-size datasets (section 3.6). Compared to the two-parameter model, our factorization models are able to make more accurate predictions on the hold-out test set in general. More latent factors and hierarchical dependence on question attributes contribute to better predictive performance in lower-frequency content areas (section 3.7). Benchmarking against similar models from the recent literature, our models are much faster to run and produce better predictions (section 3.8).

In section 3.9 of our paper, we describe a real-world application of our models. *Smart Homework* uses predictions from our models to make personalized question recommendations to students that are not too hard nor too easy. We plan to experimentally test the effect of personalization on student outcomes in a future study. Section 3.10 concludes the paper by discussing other potential applications of our models.

3.2 Related Work

Our methods are inspired by recent advances in inference techniques and machine learning models that employ user/item factorization. One central idea behind our proposed models is that students and items have latent vector representations whose dot product influences the probability of answering a question correctly. This approach draws from recent work on latent factorization that have been applied to providing online recommendations (Gopalan et al., 2015; Donnelly et al., 2020), analyzing complementarity and substitutability in consumer choice (Ruiz et al., 2017; Donnelly et al., 2019), and geographical preferences of restaurant-goers (Athey et al., 2018).

These works draw from the large literature on recommender systems where the standard approach is to find a try to find an approximation of the full matrix of user-item interactions using the product of two lower-rank matrices. Nonetheless, these recent works extend the latent vector representation approach to allow latents to depend on observed characteristics and to account for time-varying effects. We will also incorporate these innovations into our approach.

Advances in Bayesian variational inference make Bayesian inference computationally feasible on massive datasets. Variational inference recasts Bayesian inference as an optimization problem, lending it to stochastic optimization techniques (e.g. stochastic gradient descent) which allows the algorithm to scale to large datasets. See Blei et al. (2017) for a review. Recent engineering work allows models using stochastic variational inference to be implemented through off-the-shelf machine learning packages (Tran et al., 2016; Bingham et al., 2019).

Our work is also related to the large literature on linking in educational measurement. Linking refers to the practice that compares student performance across different tests (see Kolen and Brennan (2004) for a review). Similar to our goal, linking can be interpreted as a way to overcome sparsity in the combined student-item matrix of different tests that share common persons or common items (Reardon et al., 2019). A major difference, however, is that the linking literature is generally focused on the design of tests for particular settings where researchers control the recruiting of examinees and the administration of exams. As such, procedures for accommodating data sparsity are mostly ad hoc (e.g. estimating parameters from test A first, then keeping these parameters fixed when estimating test B) Kolen and Brennan (2004). There is no prevailing consensus on which methods

should be used and typical works in this literature are also only applicable to certain model specifications. Our work differs from this literature in the scale of data and the highly sparse nature of our context and the generality of our approach.

Our work is also related to the literature that casts the model of student learning as solutions to a predictive problem (see Pardos (2017) for a review). A prominent line of research in this realm centers around Bayesian Knowledge Tracing and builds temporal models of student learning (Corbett and Anderson, 1994). Recent advances along this line of research have employed deep and recurrent representation to this task (Piech et al., 2015), incorporating prior knowledge of the learners into the model (Yudelson et al., 2013), and modeling question difficulty (Pardos and Heffernan, 2011).

Many researchers used the temporality of student responses and leveraged recent developments in training deep neural networks to predict student-question responses (see Hernández-Blanco et al. (2019) for an extensive review). However, recent papers find that simpler models with psychological interpretations can behave just as well as deep learning approaches when structured to fit a few regularities (Khajah et al., 2016; Wilson et al., 2016a,b).

Our approach takes the middle ground where we produce interpretable parameters while using a factorization approach to flexibly model latent regularities. Our work essentially reduces the dimensionality of the temporal student-question sequence into a student-question matrix and attempt to solve a matrix completion problem (Candes and Plan, 2010; Mazumder et al., 2010). One of the major hurdles in recent work in this line of literature is that computational costs may be prohibitively high for large datasets using existing approaches (e.g. (Poole et al., 2008; Bailey, 2007; Shor and McCarty, 2011)). To remedy this, Imai et al. (2016) proposed a series of EM algorithms that significantly improved the computational performance in the estimation of ideal points (estimating latent traits for legislators and bills) by as much as 1000 times over existing MCMC approaches. Most recently, computer scientists Wu et al. (2020) proposed optimizing a specialized loss function that lower bounds the marginal likelihood over a student's responses. This loss is named Variational approach for Item response theory based on a novel lower BOund (VIBO). The authors showed that VIBO scales well to predictive exercises in large real-world data such as the PISA science dataset.

3.3 Model specification

We use Bayesian variational inference for all parameters in our proposed models. See section 3.11.1 for an summary of Bayesian variational inference. We place Gaussian posteriors on every parameter with flexible mean and scale. We also initialize most parameters with Gaussian priors with mean 0 and a standard deviation of 1. In practice, our optimization methods also use stochastic gradient descent and various computational optimizations implemented by Bingham et al. (2019).

3.3.1 Model 1: Two-Parameter model

Our first model takes the form of the classic Two-Parameter model from the item response literature. The only difference to traditional models is that we estimate the parameters through Bayesian variational inference.

In this model θ_i is the student parameter and α_j, β_j are question parameters. For the probability of student i answering question j correctly ($Y_{ij} = 1$), the two-parameter model is defined as:

$$P(Y_{ij} = 1 | \theta_i, \alpha_j, \beta_j) = \frac{1}{1 + \exp(-\alpha_j(\theta_i - \beta_j))}$$

3.3.2 Model 2: Latent factorization model

In this model, we map students and questions into latent vectors (θ_i, α_j) and allow their inner product to influence the probability of correctness. The model is given by:

$$P(Y_{ij} = 1 | \theta_i, \alpha_j, \beta_j) = \frac{1}{1 + \exp(-(\theta_i^\top \alpha_j - \beta_j))}$$

The main benefit of latent factorization in this model is analogous to that of methods commonly found in recommendation systems. These types of models allow the parameters to learn from the structure of the student-question matrix. The richness of the latent factors allows us to quantify student ability in multi-dimensional ways as the predictions for correctness will be different for items with different latent factors.

3.3.3 Model 3: Hierarchical factorization model

In this model, we retain the structure of the latent factorization model. However, we replace a simple latent vector α_j with a concatenation of the latent α_j and a latent transformation of observed question covariates X_j through the transformation matrix H_α. We denote this concatenated vector as $\alpha_j \oplus H_\alpha X_j$. Accordingly, we make θ_i to be the same length as the result of the concatenation.

The trainable parameters of this model are $\theta_i, \alpha_j, H_\alpha, \beta_j$. The model is given by:

$$P(Y_{ij} = 1|\theta_i, \alpha_j, H_\alpha, \beta_j, X_j) = \frac{1}{1 + \exp(-(\theta_i^\top(\alpha_j \oplus H_\alpha X_j) - \beta_j))}$$

By adding a flexible dependency on the observed characteristics of questions, we are allowing question attributes to influence the probability of correctness directly. Similar to Athey et al. (2018), this hierarchical structure may allow the model to perform better, especially for questions that appear in the data with low frequency.

3.3.4 Software implementation

Our code base[1] uses the probabilistic programming package Pyro (Bingham et al., 2019) for Bayesian stochastic variational inference. Pyro is built on top of Pytorch (Paszke et al., 2017) and uses data structure, automatic differentiation, and optimizers from the latter.

3.4 Data collection

We apply our models to data generated on the 17Zuoye platform. The data comes from homework assignments and exams in three subject areas–English, math, and Chinese. Records are logged at the students-question level.

Exam data on the 17Zuoye platform are also tagged with question attributes. Attributes include the appropriate grade level of the question. They also include two types of domain knowledge tags. One system maps questions to competencies. The other maps questions to skills. These tags are manually labeled by content specialists.

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[1]Code access is available at `https://github.com/henrishi/bm_model`.

3.5 Parameter retrieval

In this section, our goal is to compare how well our proposed models and estimation strategies retrieve true data generating parameters. In addition, we also compare the parameters retrieved by our approach to those retrieved by a widely-used traditional item response model package *ltm* (Rizopoulos, 2006). *ltm* results are labeled as *traditional_2param* in the following figures.

We focus on the two-parameter models for this exercise because the factorization models are under determined systems and there are multiple parameter arrangements that can yield the same prediction. *ltm* produces frequentist point estimates and standard errors. For the sake of comparison, with our Bayesian models, we take the mean of the posterior distribution as our estimator and the standard deviation of the posterior distribution as our standard error.

3.5.1 Simulated dense data

We fit *ltm* and *Bayesian two-parameter model* on a small simulated data set. Our data generating process samples θ and β from a normal distribution with mean 0 and standard deviation of 1. α is sampled from a normal distribution with mean 1.2 and standard deviation of 0.5 while constrained to be positive. The student-question response data is then drawn from a Bernoulli distribution where the probability of for a correct answer from student i and question j is $P(Y_{ij} = 1) = \frac{1}{1+\exp(-\alpha_j(\theta_i - \beta_j))}$. We have 30 questions and 50 students in the simulated data and every student has a response for every question.

We focus on the estimate for the β parameter since it is the easiest among the three parameters to estimate. Figure 3.1 compare the estimates from *ltm* and *Bayesian two-parameter model* alongside the true data-generating parameter. Table 3.1 presents the correlation (both linear and ranking correlations) between the estimates and the true parameter values.

We see that the *Bayesian two-parameter model* is recovering parameters better than *ltm* . *ltm* has large devious estimates for the parameter with the lowest value and large error ranges for certain parameters with mid-range values. Not only are the point estimates more devious than *Bayesian two-parameter model* , but the standard errors are also bigger for *ltm* in general. This is especially pronounced for the devious estimates. These deviations hurt the correlations of the estimates from *ltm* with the true parameter across all the correlation metrics we report (Pearson, Kendall, and
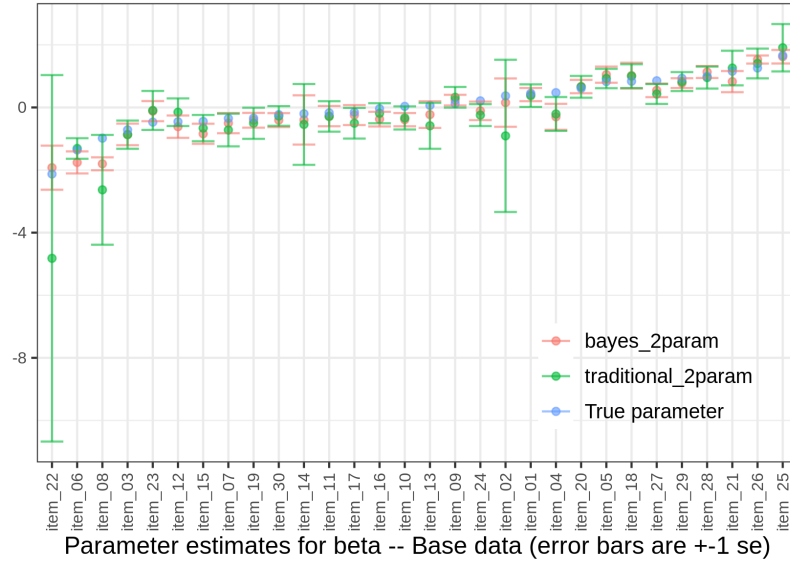
Spearman).

Figure 3.1: Small dense data: true and estimated beta

Stats	Pearson correlation (linear)	Kendall correlation (ranking)	Spearman correlation (ranking)
Bayesian two-parameter model - True parameters	0.95	0.83	0.93
ltm - True parameters	0.91	0.69	0.83

Table 3.1: Small dense data: correlation stats for beta

3.5.2 Simulated overlap data with missings

In the current exercise, we generate a simulated dataset where two groups of students share some overlapping questions but not others. The way missing data is structured is shown in Figure 3.2. Our data generating process samples θ from a normal distribution with mean 0 and standard deviation of 1. α is sampled from a normal distribution with mean 1.2 and a standard deviation of 0.5 while constrained to be positive.

Different from the section 3.5.1, we sample the β for overlap questions and those only available for students 1 - 50 from a normal distribution with mean 0 and standard deviation of 1. However, we sample the β for questions only available for students 51 - 100 from a normal distribution with mean 0.5 and standard deviation of 0.7. We sampled the β parameters this way to mimic real-world settings where question difficulty is usually different for different groups of students. As in 3.5.1 the student-question response data is then drawn from a Bernoulli distribution where the probability of a correct answer is a logistic function of the parameters.
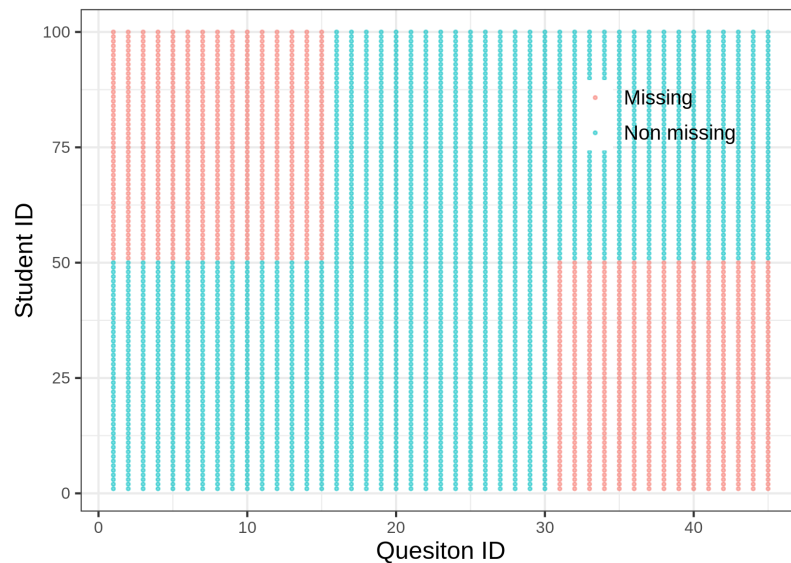

Figure 3.2: Overlap data missing pattern

In this setting, we again compare the performance of the *ltm* model to the *Bayesian two-parameter model* model in estimates of β. We find that the problem of devious estimates is greatly exacerbated for *ltm* . Figure 3.3 shows that the error ranges are excessively large for some estimates from *ltm* .

For a more informative figure, we take out the standard error bars of the outliers from the *ltm* model and re-plot in Figure 3.4. We see that while the majority of parameters have similar under both models, *ltm* is yielding very devious estimates for items 11, 31, and 37. These items are not overlapping items (see point map above) and are only taken by a single group of students.

Bayesian two-parameter model produces point estimates that are much closer to the true parameters for the parameters that *ltm* failed to estimate accurately. This corroborates the Bayesian inference property that the prior distribution serves as a regularizer and allows the model to be more numerically stable. Finally, we see that *Bayesian two-parameter model* again dominates *ltm* in correlation stats with the true parameters by a large margin in this data setting.
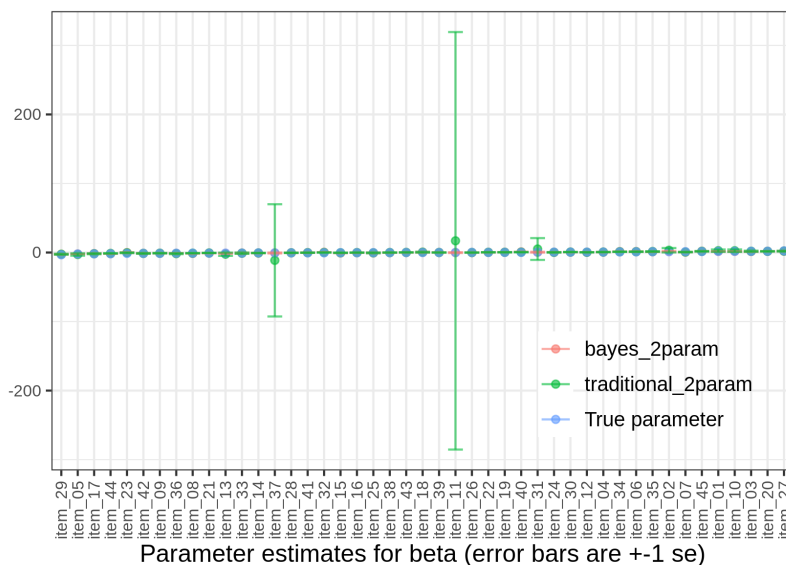

Figure 3.3: Overlap data: true and estimated beta

Stats	Pearson correlation (linear)	Kendall correlation (ranking)	Spearman correlation (ranking)
Bayesian two-parameter model - True parameters	0.96	0.84	0.96
ltm - True parameters	0.41	0.76	0.91

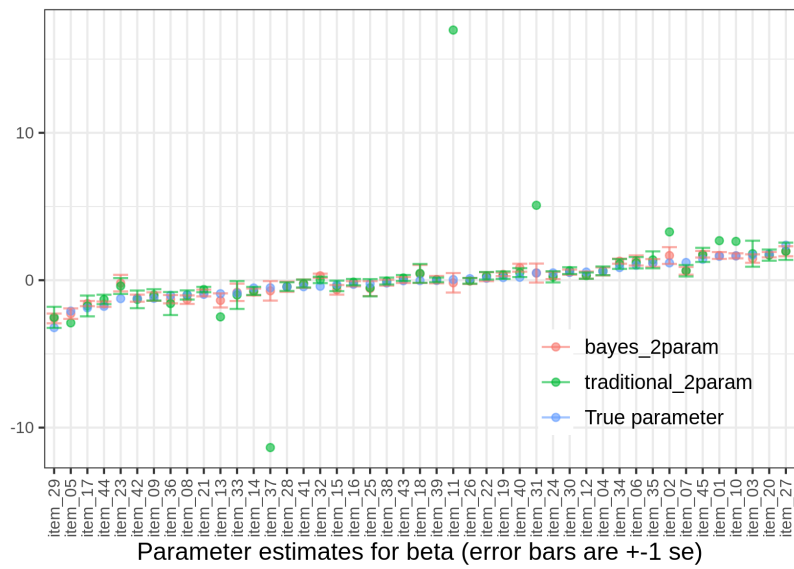Table 3.2: Overlap data: correlation stats for beta

Figure 3.4: Overlap data: true and estimated beta (zoomed in)

3.6   Computation performance

In this section, we report the computation performance of our proposed models using an industrial-scale dataset with 88 million responses. The purpose of this section is to show that our proposed Bayesian models *Bayesian two-parameter model* and *Bayesian factorization model* are computationally tractable for real-world applications. Traditional methods such as *ltm* cannot be compared here because they could not handle data of this size. Specifically, *ltm* is unable to produce any useful estimates for even very small datasets when sparsity is at this level. [2]

This dataset comes from homework records of 1000 schools over a period of 2 months. In this dataset, we have 334K questions and 162K students. The overall density of the student-question matrix is 0.16%. The training of our models terminates when convergence has been achieved. We define convergence as the change in loss averaged over the last 5 iterations falls below 0.1% of the change in loss from the first to the second iterations.

In Figure 3.5 we show the training loss for different model specifications over time. Even though we adopt stochastic gradient descent for optimization, the overall loss curve is smooth due to the large amount of data used for training. The more complex models, factorization models with longer latent vectors, tend to take marginally longer to train. Most models converge within 30 minutes, the longest model to train took less than 40 minutes (see Figure 3.6).

3.7   Predictive performance

We compare the prediction performance of our proposed models using a large exam dataset where we have access to question attributes. We want to compare the performance of all three models, *Bayesian two-parameter model* , *Bayesian factorization model* , and *Bayesian hierarchical factorization model* . Since *Bayesian hierarchical factorization model* needs question attributes as inputs, we needed to test our predictive performance on a dataset that has question attributes. The exams

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[2] *ltm* produces numerical errors for even very small samples from this dataset (e.g. a sample of 1000 records). For example, a random sample of 1000 records translated into a matrix of 639 students by 717 questions. *ltm* returns numerical errors for the resulting matrix. One can get *ltm* to run if the missing entries are replaced values (e.g. change all missing values to 0), but running *ltm* on the resulting data takes 16 minutes to converge. My testing shows that *ltm* convergence time is roughly $O(n^2)$ meaning that doubling the amount of data takes 4 times as long to converge. This is an exorbitant amount of time considering that 1000 records are a mere 0.00011% of the full dataset with 88 million records.
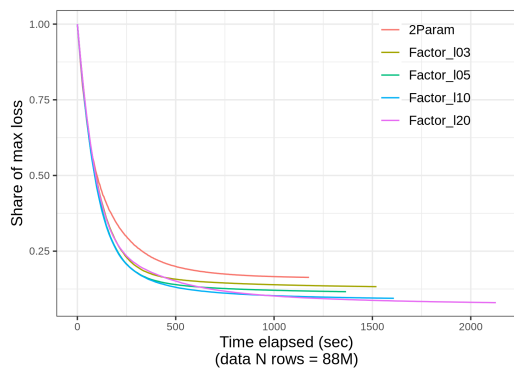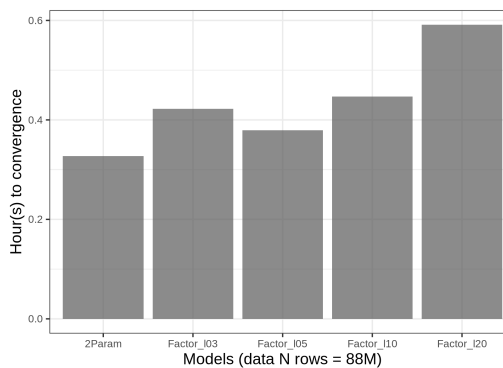
Figure 3.5: Training loss

Figure 3.6: Convergence time

are also higher stakes than typical homework assignments, as such the data may be more reflective of actual competencies and less noisy as a result.

In this dataset, we have 8.6 million student-question records. There are 3.6k questions and 261k students in total. The overall density of the student-question matrix is 0.92%.

The data were randomly divided into 80% training, 10% validation, and 10% test sets at the level of student-question interactions. This means data for a single student may appear in any of the three sets. The same goes for data from a single question. Model training would stop automatically once convergence is achieved. To get a better metric of the models' capabilities, we define a stricter convergence criterion – as average per-iteration changes in loss becoming 0.05% the initial change.

We compare *Bayesian factorization model* against *Bayesian two-parameter model* in sections 3.7.1 and 3.7.2. Having established the superiority of factorization models, we move on to quantify the gains from adding hierarchical dependency in *Bayesian hierarchical factorization model* in section 3.7.3.

3.7.1 Overall predictive performance

We first document the overall predictive accuracy for *Bayesian two-parameter model* and *Bayesian factorization model* in Figures 3.7 and 3.8. For *Bayesian factorization model* , we show results from three models where the length of the latent vectors θ, α are taken to be 3, 5, 10, and 20 respectively. We see that *Bayesian factorization model* models enjoy higher AUC in the test set than *Bayesian two-parameter model* . The F1 statistic tells the same story where *Bayesian factorization model*

models are superior predictive performance.[3] We note that in all models, training set accuracy is higher than the test set. This is partially attributable to the sparse nature of the student-question matrix. Some students or questions may only show up in the training set or the test set, making test set predictions less accurate than the training set.
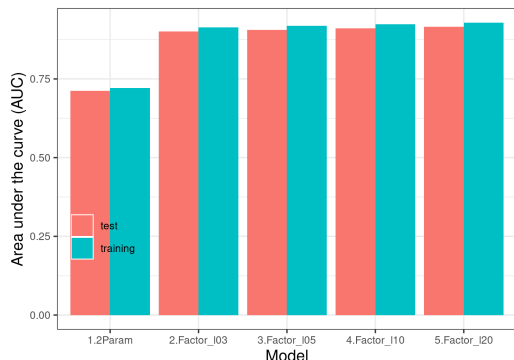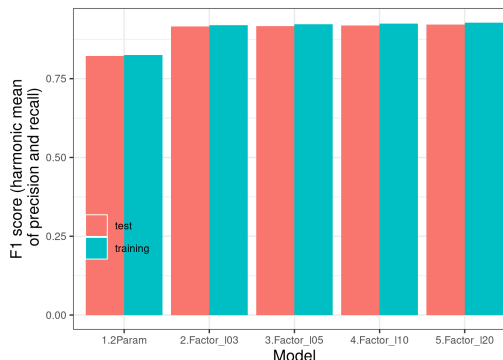

Figure 3.7: Area under the curve (AUC) by model

Figure 3.8: F1 statistic by model

3.7.2 Predictive performance by content area

Do factorization models with more latent factors perform better than simpler models on predictive accuracy in less frequent content areas. In this section, we answer this question by examining the predictive performance of candidate models by question knowledge labels. We have 22 knowledge labels in our data. For example, a knowledge label for a math question may be "arithmetic" or "geometry", one for an English question may be "English spelling" or "English pronunciation".

Figure 3.9 shows the distribution of student-question records by knowledge labels. From left to right, we see knowledge labels in descending popularity. We note that some knowledge labels have significantly lesser data than the most popular knowledge labels.

We document the predictive performance of different models across knowledge labels in 3.10 and 3.11. We see that confirm that *Bayesian factorization model* dominates *Bayesian two-parameter model* in both AUC and F1 across the knowledge labels. Interestingly, the more complicated factorization models (the 10-factor and 20-factor models) have better performance than simpler factorization models (the 3-factor and 5-factor models) in predictive performance for the less frequent

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[3] In calculating the F1 statistic, we set the predictive threshold at $P > 0.5$ for a positive prediction.
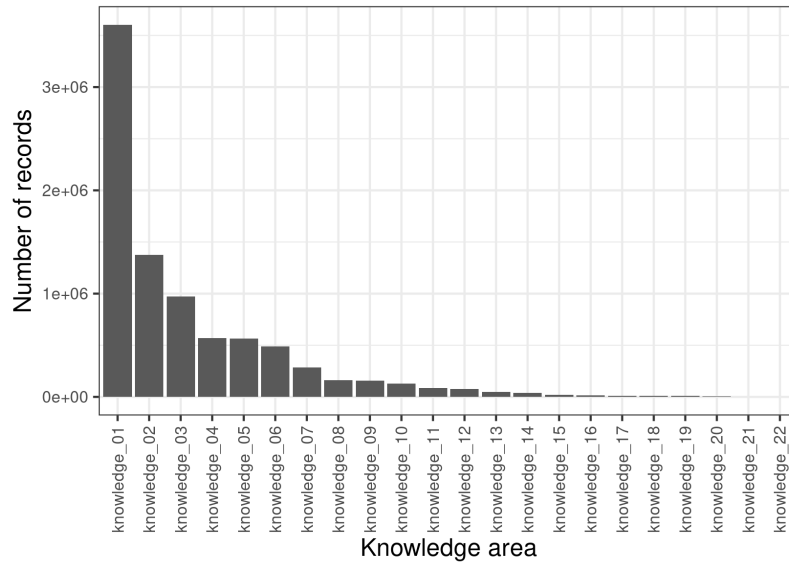
Figure 3.9: Count of student-question records by knowledge label

knowledge labels. This suggests that additional latent factors may have picked up additional heterogeneity useful in predicting less frequent knowledge labels.
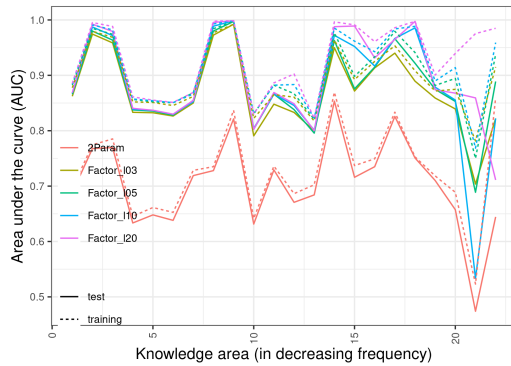

Figure 3.10: Area under the curve (AUC) by knowledge label
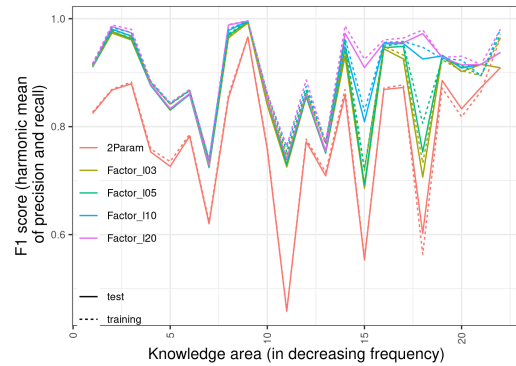

Figure 3.11: F1 statistic by knowledge label

3.7.3 Performance gains from adding hierarchical dependency

Having established the superiority of factorization models over *Bayesian two-parameter model* , we move on to quantify the gains from adding a hierarchical dependency in *Bayesian hierarchical*

factorization model . To make models more comparable, we hold the size of latent vector constant and compare the *Bayesian factorization model* model with the same model that adds different hierarchical dependency.

We use two types of question attributes, one of which is the one-hot vector of knowledge label described in section 3.7.2. The other attribute is a multi-hot vector of skills involved in the question. The attribute is encoded as a multi-hot vector because a single question can be associated with multiple skills. For our candidate models with hierarchical dependency, we have freedom in choosing the size of the latent matrix H_α. Following Athey et al. (2018) we pick H_α such that some entries are 0 so that certain types of question attributes can only contribute to certain parts of the resulting latent vector $H_\alpha X_j$. As specified in 3.3.3, the latent vector $H_\alpha X_j$ is a linear combination of these representations. We choose two model specifications, the first structures H_α such that the knowledge labels and skills each map to one latent scalar so the resulting $H_\alpha X_j$ is a length 2 vector. The second is slightly more complex in that the knowledge labels and skills each map to two latent scalars so the resulting $H_\alpha X_j$ is a length 4 vector.

In Figure 3.12 we compare the factorization model with a length 3 latent vector to a hierarchical model with the same setup. We see that the hierarchical models outperform the factorization models slightly overall, but the improvement is more significant for certain low-frequency knowledge labels. The same is true when we look at factorization model with a length 5 latent vector and hierarchical models with the same set up (see Figure 3.13).[4]

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[4]For the sake of brevity, we focus on AUC for our comparison, but the F1 statistics results are substantively the same as the AUC results.
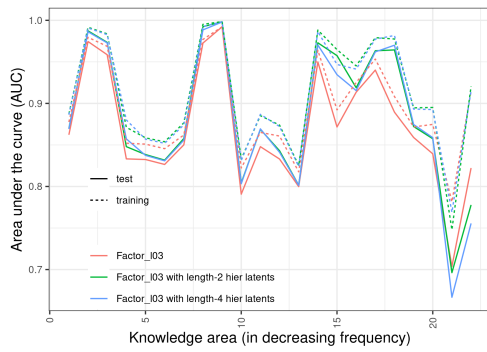
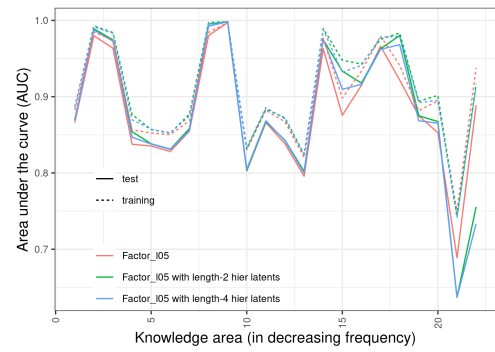Figure 3.12: Area under the curve (AUC) by knowledge label (models where latent vector size = 3)

Figure 3.13: Area under the curve (AUC) by knowledge label (models where latent vector size = 5)

3.8 Benchmark comparisons

In this section, we compare our model against prominent contenders in recent literature on fast, accurate, and scalable IRT model estimations. Imai et al. (2016) proposed a series of EM algorithms that significantly improved the computational performance in the estimation of ideal points (estimating latent traits for legislators and bills) by as much as 1000 times over existing MCMC approaches. The work is highly cited in the political science literature (hereafter *emIRT*). Most recently, Wu et al. (2020) proposed coupling a specialized loss function named VIBO with variational inference to estimate IRT models (hereafter *varIRT*). They have shown that their model works well in parameter retrieval exercises with simulation data and scales well to predictive exercises with real-world data.
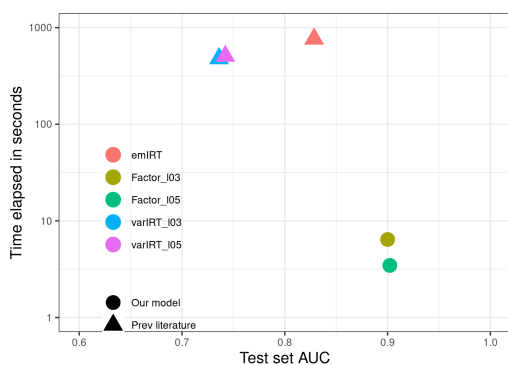

Figure 3.14: Computational and predictive performance: proposed models vs literature benchmarks
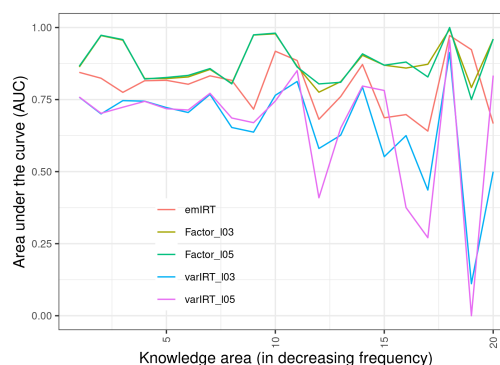
Figure 3.15: Test set AUC by knowledge label: proposed models vs literature benchmarks

In this exercise, we want to run both *emIRT* and *varIRT* on the same dataset with the same computer as we used for our previous predictive exercises (8.6M student-question records) and compare the models' computational and predictive performance. While we attempted to run all models on the whole dataset, *emIRT* and *varIRT* could not digest the entire dataset as it was too massive and demanded resources beyond the physical resources of the computer.[5]

As a result, we trained all our models using a smaller subset of the data. The subset was

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[5]Both *emIRT* and *varIRT* resulted in out-of-memory errors on our computer, which had 16G of memory and was able to perform estimations with our proposed models with ease. One possible contributor to the disparity is that *emIRT* and *varIRT* both require data to be fed to it in wide form where each row represents a student and each column a question. This greatly increases the size of the data in memory compared to the long format feed data that our implementation adopts.

created by sampling 10K students from the original dataset at random and keeping all the questions associated with each student. The resulting data has $327,082$ student-question records, covering $10,000$ student and 3201 questions. The student-question matrix has a density of 1.02%. We randomly split the data into 90% training set and 10% test set.

Similar to our model, *varIRT* allows for more than one latent dimension. As such, we ran *Bayesian factorization model* and *varIRT* each with 3 and 5 latent dimensions. The official implementation of *emIRT* only allows for one latent dimension so we did not produce variations of the model. Note that we use *Bayesian factorization model* and not *Bayesian hierarchical factorization model* in this exercise because *varIRT* does not support hierarchical dependency on observables and using such would result in an unfair comparison.

Figure 3.14 shows how our model compares to *emIRT* and *varIRT* in both computational performance and test set predictive performance (AUC). We see that *Bayesian factorization model* dominates both alternatives with faster computational performance and better predictive accuracy. While the lead in predictive performance over *emIRT* and *varIRT* is considerable, both versions of the *Bayesian factorization model* also represent a substantial gain in computational performance. They ran on the data in under 10 seconds, while *varIRT* and *emIRT* took around 10 times the amount of time to run – on the order of hundreds of seconds (the y axis of Figure 3.14 is log scaled).

Figure 3.15 shows how our model compares to *emIRT* and *varIRT* by predictive performance across knowledge labels. Barring a few knowledge labels where *Bayesian factorization model* had similar performance with *emIRT* , the former largely dominated both *emIRT* and *varIRT* . Notably, despite being more complex in model specification, *varIRT* underperformed *emIRT* in high-frequency knowledge labels and most of the low-frequency knowledge labels.

3.9 Model application: recommender system for personalized questions

This section describes a practical application of our proposed models. Our models were put into use at our partner company to power a recommender system producing personalized questions for students. The recommender system is one piece of a broader product, *Smart Homework* , which

combines personalized questions with immediate feedback when the student incorrectly answers a question. Figure 3.16 shows the relationship between the different components of the system.

The recommender system aims to produce questions that are not too difficult nor too easy. This aligns with a long line of research in educational psychology stemming from the Zone of Proximal Development proposed by Vygotsky (1980). The predicted probabilities for student i getting question j correct is computed using our proposed models from historical data.

Once the probabilities are calculated, the question pool is filtered by taking out questions that the student has already attempted in the past. Then the pool goes through a ranking algorithm for each student. The algorithm ranking questions based on a weighted sum of different factors. The most prominent factor is how close the predicted correctness probability is to 0.7. Other factors include how recent was the question created, whether the question includes a picture or a table. The top 10 questions are selected to form the *Smart Homework* .

As the student works through her *Smart Homework* she moves from one question to the next if she answers correctly. When she answers a question incorrectly, she is taken through a series of explanations and exercises also known as the learn-practice-explain (LPE) module. The student is guided through a series of questions and animations that explains one knowledge point for the triggering question (see Figure 3.17). The typical time to complete this module is between 2 and 4 minutes. As the student answers the questions, only a correct answer would allow the student to move forward. A wrong answer would trigger a hint and the student is directed to answer the question again.

After completing the module, the student is asked to answer another question similar to the original question she missed. Regardless of whether she answered this question correctly, she is taken back to other questions in the *Smart Homework* . *Smart Homework* exits when the student has completed every question.

3.9.1 Batch updates and parameter freeze

We added two features to the implementation of our models to address practical challenges for building *Smart Homework* . Firstly, we implemented batch updating of parameters. Since our models are based on Bayesian inference, making updates to parameter estimates is very straightforward.
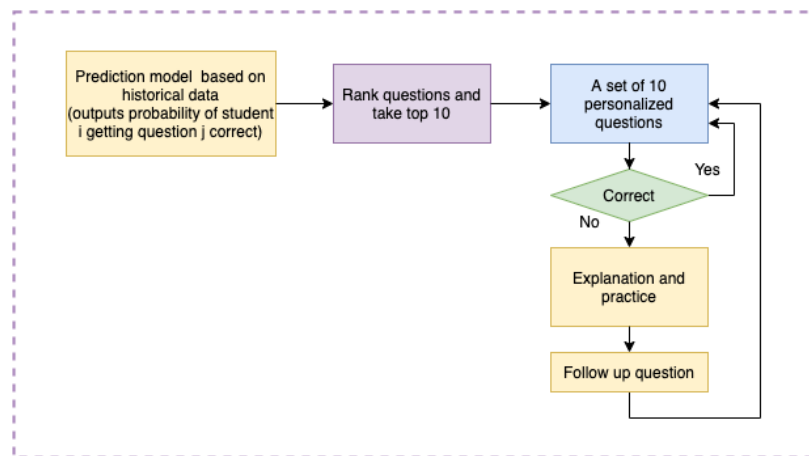
Figure 3.16: *Smart Homework* is powered by a recommeder system and immediate feedback
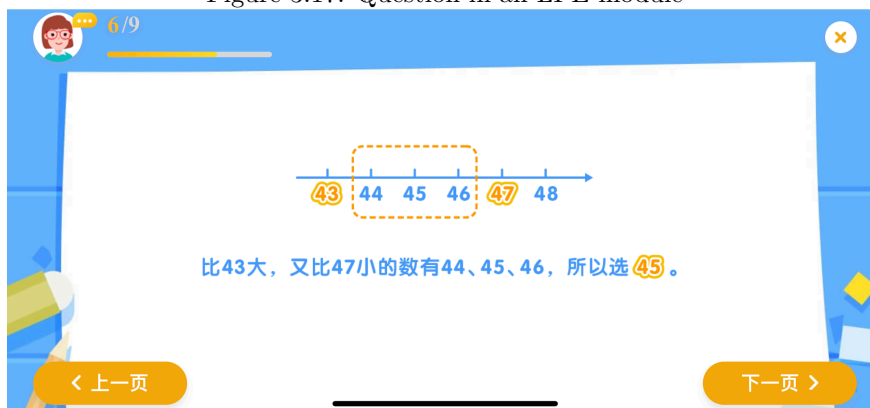

Figure 3.17: Question in an LPE module

Figure 3.18: Explanatory text/audio in an LPE module

Our features load pre-existing parameters into the model in the form of prior distributions. The prior distributions encapsulate all the information the model learned from previous data. Thus, instead of re-fitting the model every time new data comes in, the model needs only to use incremental data to update itself using the prior distributions as its starting point. This helps reduce the computational costs of keeping models up-to-date. In practice, our model updates itself with incremental data once a week.

Secondly, we added a feature to allow parameters to be fixed. The fitting of the model will only alter trainable parameters but not fixed parameters. This allows question parameters, typically estimated over a large amount of historical data, to remain unchanged in the model fitting process. This lessens the burden of optimization and can further reduce computational costs when a large number of questions are involved.

3.9.2 Future experimentation

In a future study, we plan to experimentally test whether *Smart Homework* improves student outcomes. Namely, we will assign three versions of *Smart Homework* as the treatment condition and compare its effects on student engagement and achievement against a control group that received a curated (but non-personalized) homework. The three versions of *Smart Homework* will be 1) the full version with both personalization and feedback, 2) a partial version with only personalization, and 3) a partial version with only feedback.

Our partner company is currently (as of May 2021) conducting a pilot of *Smart Homework* with 30K student participants. The pilot stage aims to reveal bugs and inefficiencies in the product's engineering. The pilot stage also aims to establish an understanding of the baseline student behaviors including open rate, completion rate, and utilization of feedback. The main study will proceed after the issues revealed in the pilot stage are addressed.

3.10 Conclusion

We have shown that latent factorization and hierarchical modeling powered by Bayesian variational inference can make important gains in modeling student-question interactions using large datasets

from online education. This set of modern approaches perform well in parameter retrieval and predictive accuracy. They are also numerically stable and can be applied to industrial-scale datasets with computational ease. Our models compare favorably in both computational performance and predictive accuracy against similar models from recent literature. Finally, they also allow for desirable features such as batch updates and parameter freeze to be implemented with ease.

Smart Homework is only one of the many applications that our models can power. Our models can be used to identify learning gaps. In particular, our factorization approach can identify content domains where a student is weaker than average. Early warning systems and remedial content recommendations can be generated based on such information.

Our models can also improve the informational landscape of parents and teachers. One example can be building learning progress dashboards. For this application, the predictions from our models will be used to derive statistics about student progress. An example of this is using trained models to make predictions about the average rate at which the student will get a particular set of pre-selected domain-representative questions correct. This statistic can be used as an indicator for proficiency.

Finally, applications can make use of the latent parameters the models learned from data. One such application is the automatic labeling of questions based on the latent parameters produced by our models. Another application is discovering the preferences of teachers in terms of which types of question a particular teacher likes to assign.

3.11 Appendix

3.11.1 Bayesian Variational Inference

The main method of inference for parameters in our proposed models is Bayesian variational inference (see (Blei et al., 2017) for a review). Bayesian variational inference is increasingly popular in estimation tasks involving large amounts of data. It has superior runtime performance compared to traditional Bayesian estimation techniques such as Markov Chain Monte Carlo. Variational inference has been shown to work well in mean-fields approximation tasks, but existing methods are less well suited to recover correlations between parameters.

Variational inference proposes a parameterized class of posterior distributions and reduces the Bayesian inference task to one of optimizing a divergence value between the proposed posterior and the true posterior. For this inference task, I use the Kullback-Leibler divergence defined by

$$\lambda^* = \arg\min_{\lambda} KL(q(z; \lambda) || p(z|x))$$

$$= \arg\min_{\lambda} E_{q(z;\lambda)}[\log(q(z; \lambda)) - \log(p(z|x))]$$

where z is the parameter vector of interest, x is the data, and λ is the parameter vector characterizing the proposed posterior distribution. However, directly optimizing this expression is not possible since it involves the posterior distribution $p(z|x)$. However, since we can reformulate the divergence as

$$KL(q(z; \lambda) || p(z|x)) = \log(p(x)) - \text{ELBO}(\lambda)$$

where $\text{ELBO}(\lambda)$ is the evidence lower bound defined by

$$\text{ELBO}(\lambda) = E_{q(z;\lambda)}[\log(p(x, z)) - \log(q(z; \lambda))]$$

we can derive gradients to perform optimization on the ELBO.

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