

1 Temperature and Economic Outcomes

1.1 Temperature Aggregation

See the do-file "1_1_temperature_aggregation.do".

1.2 US Climate Impacts: County-Year Damages

1.

Figure 1 displays the linear relationship between log transformed `emp_farm` and the vector of binned temperature controls, including county and year fixed effects. The omitted temperature bin is 16-20°C. Thus, the coefficient of roughly 0.0015 on the $> 32^{\circ}\text{C}$ bin can be interpreted as “Relative to an additional 16-20°C day during the year, an additional $> 32^{\circ}\text{C}$ day is associated with an increase in farm employment of roughly 0.15%. The relationship is statistically significant at the 5% level.”

2.

COME BACK Figure 2 displays the relationship between $\log(\text{per capita farm prop income})$ using the restricted cubic spline, including county and year fixed effects.

3.

Table 1 displays the results of using the binned temperature estimator to design a test for whether we observe treatment effect heterogeneity. Specifically, the number of $> 32^{\circ}\text{C}$ days is interacted with the respective temperature bins. As in Exercise 1.1, the omitted bin is 16-20°C. While the majority of the coefficients on the interaction terms are not statistically significant at the 5% level, the coefficients on “`tempB0 × tempA32`” and “`temp4to8 × tempA32`” are statistically significant at the 5% level. Conducting an F-test on the interaction terms jointly equalling zero yields an F-statistic of roughly 13; hence, we can reject the null hypothesis of no treatment effect heterogeneity.

2 Hedonic Air Quality Analysis

Note: to maintain the same sample throughout the questions in this section, I dropped the 47 observations that were missing any of the variables included in the dataset, leaving me with 966 observations. All

2.1 Questions

1.

- Table 2 displays the estimates of regressing housing prices on pollution levels with (Column 2) and without (Column 1) controls.
- Without controls, my estimates imply that a unit increase in the change in the annual geometric mean of total suspended particulates pollution (TSPs) from 1969-72 to 1977-80 is associated with roughly a 0.24% increase in housing values from 1970 to 1980 (statistically sig. at 5%).
- However, the point estimate is reduced by almost an order of magnitude after controlling for other housing price determinants and is no longer statistically significant. This implies that the regression without controls suffers from omitted variables bias, where the change in pollution is correlated with other housing price determinants.
- To confirm this, I regress `dgts` on the observable measures of economic shocks: `dincome`, `dunemp`, `dmnf` (Table 3) as well as compute simple correlations between `dgts` and these three variables. Indeed, increases in income and manufacturing employment are positively correlated with increases in TSPs while increases in unemployment are negatively correlated with increases in TSPs.
- From Table 2, these three covariates have a nonzero relationship with housing prices. Thus, the regression without controls suffers from omitted variables bias since `dgts` is correlated with omitted determinants of housing prices.

2.

For mid-decade regulatory status (`tsp7576`) to be a valid instrument for pollution changes when the outcome of interest is housing price changes, it must satisfy two criteria:

1. Relevance: `tsp7576` must be correlated with the determinant of interest, `dgts`. This can be seen in Column 1 of Table 4.
2. Exclusion: `tsp7576` is uncorrelated with the error term, or other variables that determine housing prices. Looking at the coefficients on the economic shock measures `dincome`, `dunemp`, `dmnf` in Column 2 of Table 4, `dincome` is statistically significantly related to `tsp7576` at the 5% level, while `dunemp` and `dmnf` are not. While we are able to control for `dincome` in our regression, this finding is not reassuring that there are no unobserved factors in the error term that are correlated with our instrument and the outcome of interest.

3.

Table 5 displays the results for this problem.

- Columns 1 and 2 display the first-stage relationship between regulation and air pollution changes without and with controls, respectively. Both with and without controls, the first stage relationship appears to be strong (i.e., EPA regulation appears to be a strong and statistically significant predictor of air pollution changes), though is slightly attenuated when including the controls. Interpreting our findings without (with) controls in Column 1 (Column 2), counties that were regulated in either 1975 or 1976 experienced a roughly 10.2 (8.14) unit reduction in the annual geometric mean of TSPs from 1969-72 to 1977-80 compared to counties that were not regulated (holding fixed other covariates in the regression).
- Columns 3 and 4 of Table 5 display the reduced form relationship between regulation and housing price changes without and with controls, respectively. Both coefficients are similar in magnitude and are statistically significant at the 5% level. Interpreting our findings without (with) controls in Column 3 (Column 4), counties that were regulated in either 1975 or 1976 experienced a roughly 3.46% (3.70%) increase in housing values from 1970 to 1980 compared to counties that were not regulated (holding fixed other covariates in the regression).
- Columns 5 and 6 of Table 5 display the 2SLS relationship between air pollution changes and housing price changes without and with controls, respectively. The 2SLS estimates can be obtained by dividing the reduced form coefficients by the first stage coefficients. Both coefficients are similar in magnitude and are statistically significant at the 5% level. Interpreting our findings without (with) controls in Column 5 (Column 6), a 1 unit increase in TSPs from 1969-72 to 1977-80 is associated with roughly a 0.34% (0.45%) decrease in housing values from 1970 to 1980 (holding fixed other covariates in the regression). Assuming the IV assumptions are met, these estimates reflect the local average treatment effect, or the average treatment effect for the subset of counties whose air pollution changed as a result of EPA regulation.

4.

Table 6 displays the results for this problem.

- Columns 1 and 2 display the first-stage relationship between regulation (as defined by the annual geometric mean of TSPs in 1974 being greater than 75 units) and air pollution changes without and with controls, respectively. Both with and without controls, the first stage relationship appears to be strong, though is slightly attenuated when including the controls. Interpreting our findings without (with) controls in Column 1 (Column 2), counties with annual geometric mean of TSPs above 75 units in 1974 experienced a roughly 13.15 (10.35) unit reduction in the annual geometric mean of TSPs from 1969-72 to 1977-80 compared to counties with annual geometric mean of TSPs below 75 units in 1974 (holding fixed other covariates in the regression).

- Columns 3 and 4 of Table 5 display the reduced form relationship between our new instrumental variable and housing price changes without and with controls, respectively. Both point estimates are significant at the 5% level, and the point estimate increases by over 50% after including controls. Interpreting our findings without (with) controls in Column 3 (Column 4), counties with annual geometric mean of TSPs above 75 units in 1974 experienced a roughly 3.12% (4.98%) increase in housing values from 1970 to 1980 compared to counties with annual geometric mean of TSPs below 75 units in 1974 (holding fixed other covariates in the regression).
- Columns 5 and 6 of Table 5 display the 2SLS relationship between air pollution changes and housing price changes without and with controls, respectively. As in the reduced form relationship, both point estimates are statistically significant at the 5% level, and the point estimate with controls is larger in magnitude (roughly double) than the point estimate without controls. Interpreting our findings without (with) controls in Column 5 (Column 6), a 1 unit increase in TSPs from 1969-72 to 1977-80 is associated with roughly a 0.24% (0.48%) decrease in housing values from 1970 to 1980 (holding fixed other covariates in the regression). These results are similar to those obtained in the previous part.

5.

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- Using this discontinuity in treatment assignment, we could derive an alternate estimator using a regression discontinuity design. The key assumption in this design is continuity, which means that the expected potential outcomes are continuous at the cutoff. In both cases, the RD design estimates the impacts of eligibility for EPA regulation at the 75 unit 1974 TSPs cutoff.
- Figure 3 plots the results of estimating the nonparametric bivariate relationship between pollution changes and 1974 TSPs levels and housing price changes and 1974 TSPs levels using a bandwidth of 2. The LATE estimate at the 75 unit TSPs cutoff is the vertical distance between the right-hand (red) and left-hand (blue) lines. Comparing this estimate to the first-stage estimate in part (4),
- Figure 4 plots the results of estimating the nonparametric bivariate relationship between housing price changes and 1974 TSPs levels using a bandwidth of 2. The LATE estimate at the 75 unit TSPs cutoff is the vertical distance between the right-hand (red) and left-hand (blue) lines.

6.

Figure 5 plots (i) the nonparametric bivariate relation between the single-index measure of the housing price changes predicted to occur due to other variables changing and 1974 TSPs levels against (ii) the smoothed housing price changes from (5). Since the former lowess smoother shows a relatively smooth and continuous relationship between the predicted house price changes and 1974 TSPs without any significant jumps or discontinuities around the 75 unit threshold, this suggests that the continuity condition required by RDD is likely met.

7.

COME BACK Figure 6 and 7

8.

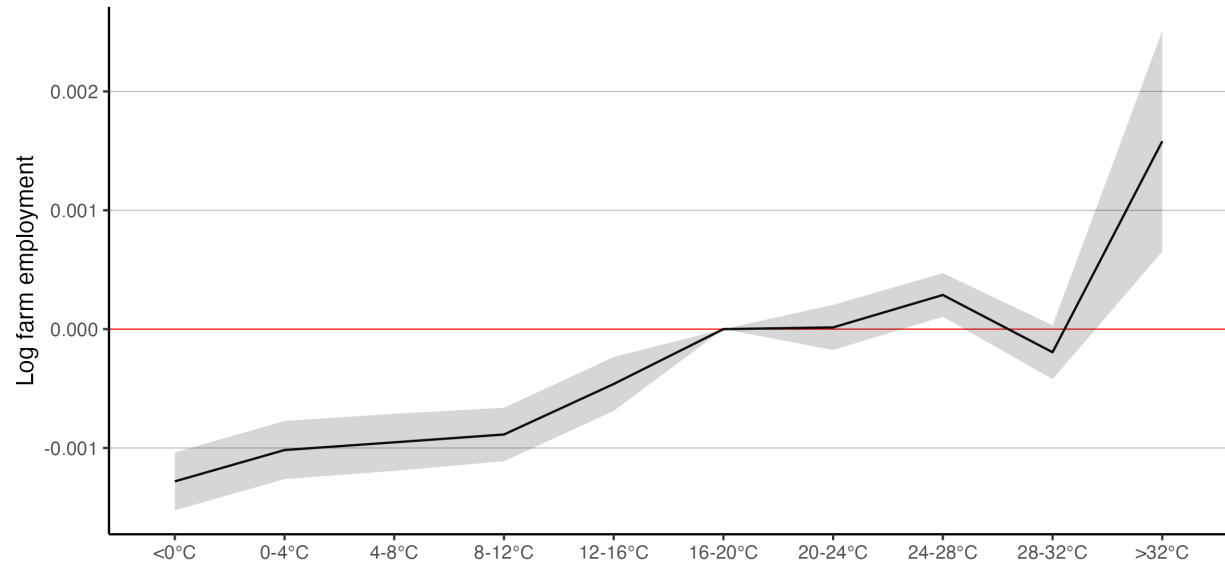
In general, two-stage least squares only recovers the LATE, or the ATE for the subpopulation whose treatment status is influenced by the instrumental variables (i.e., the compliers). This effect is identified under our usual IV assumptions of the exclusion restriction and a non-zero first stage. When one uses EPA regulation as an instrumental variable for air pollution, the LATE estimates how housing values are affected by changes in air pollution that occur only because of the changes in EPA regulation.

9.

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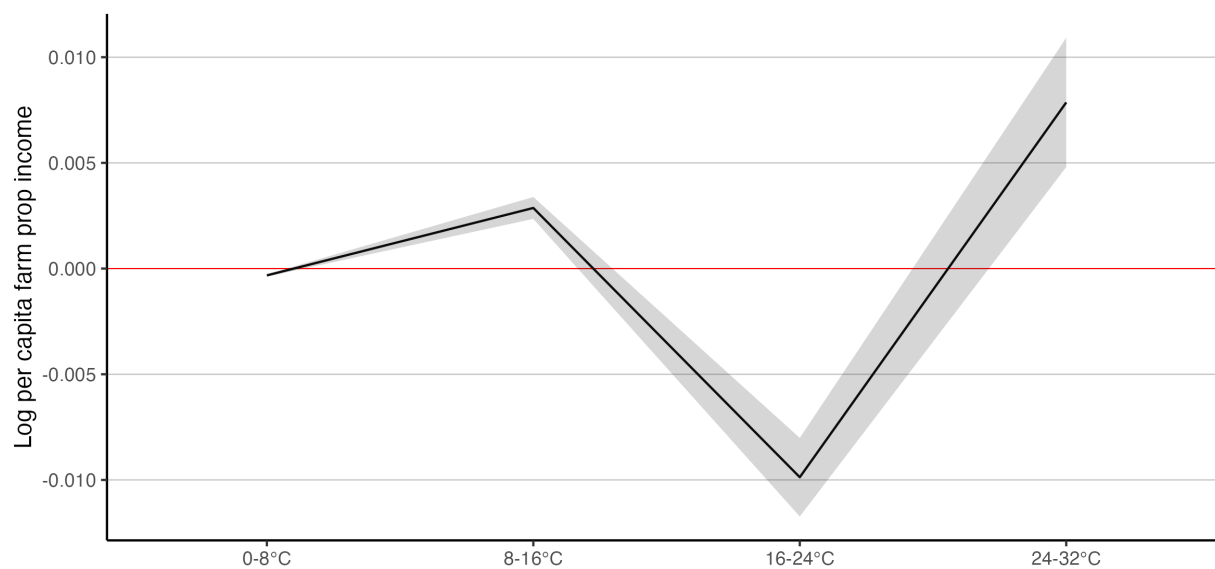
Figures

Figure 1: Exercise 1.2.1



The solid black line shows the point estimates of the regression specified in the problem. Gray shaded areas display 95% confidence intervals.

Figure 2: Exercise 1.2.1



The solid black line shows the point estimates of the regression specified in the problem. Gray shaded areas display 95% confidence intervals.

Figure 3: Exercise 2.5, pollution changes and 1974 TSPs levels

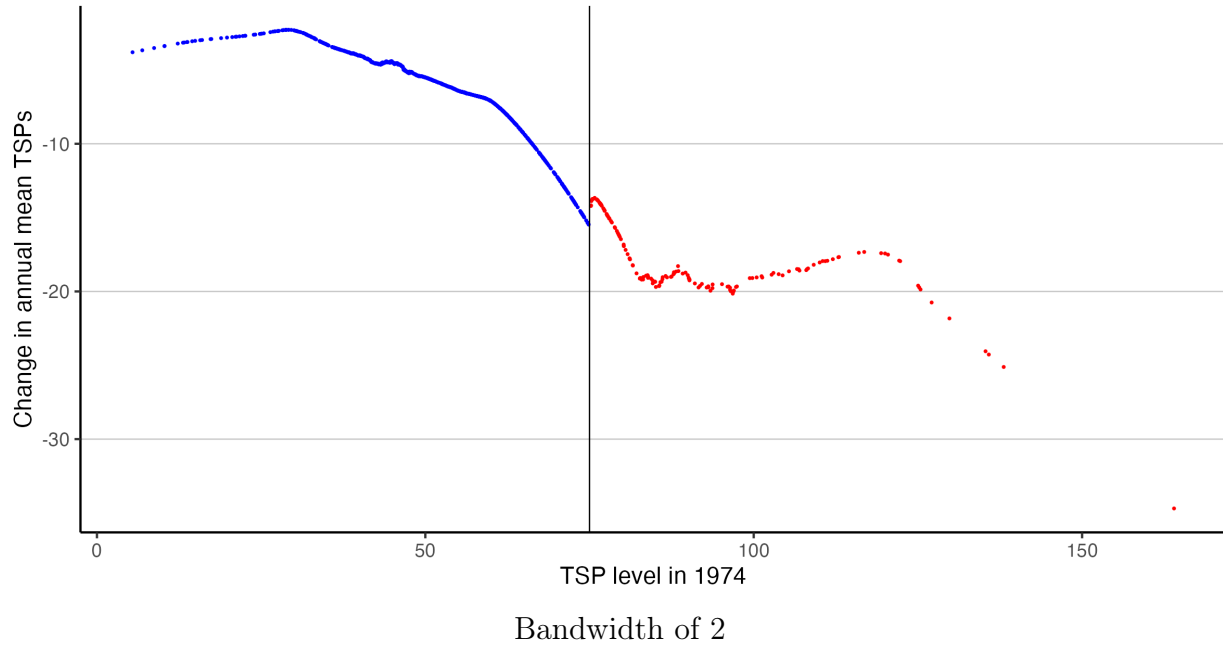


Figure 4: Exercise 2.5, housing price changes and 1974 TSPs levels

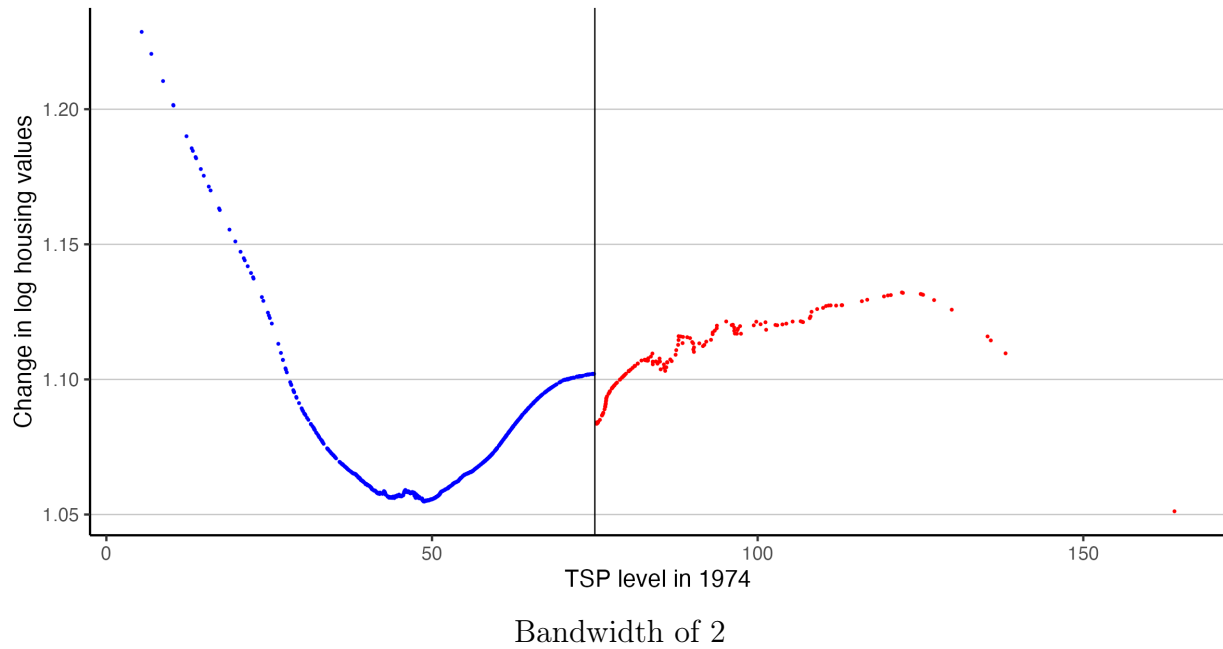
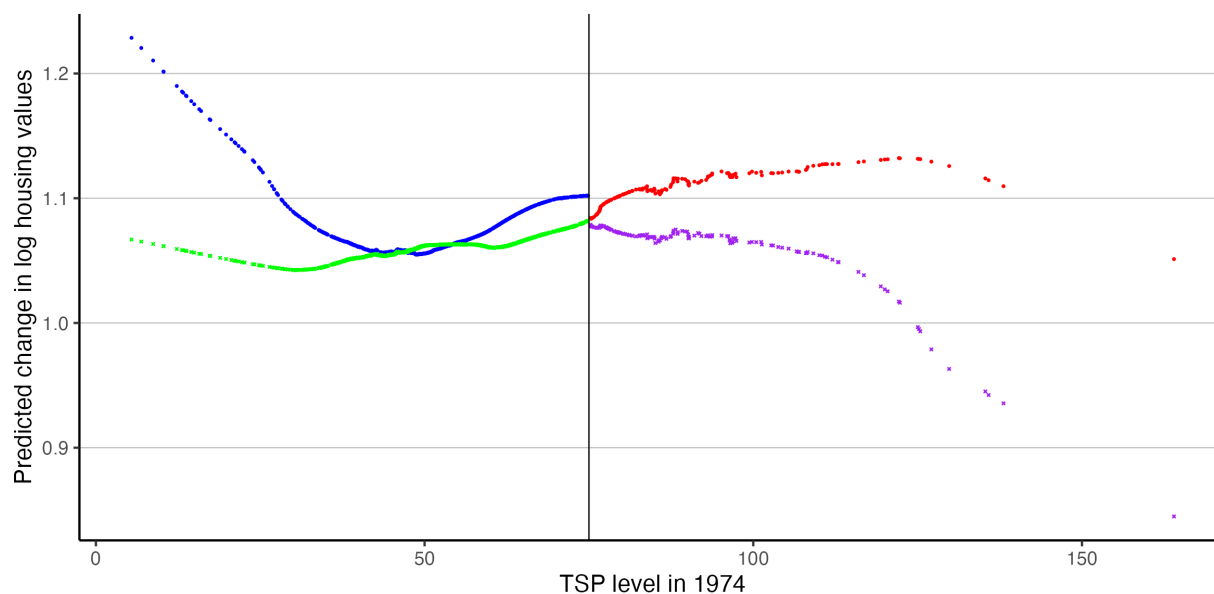
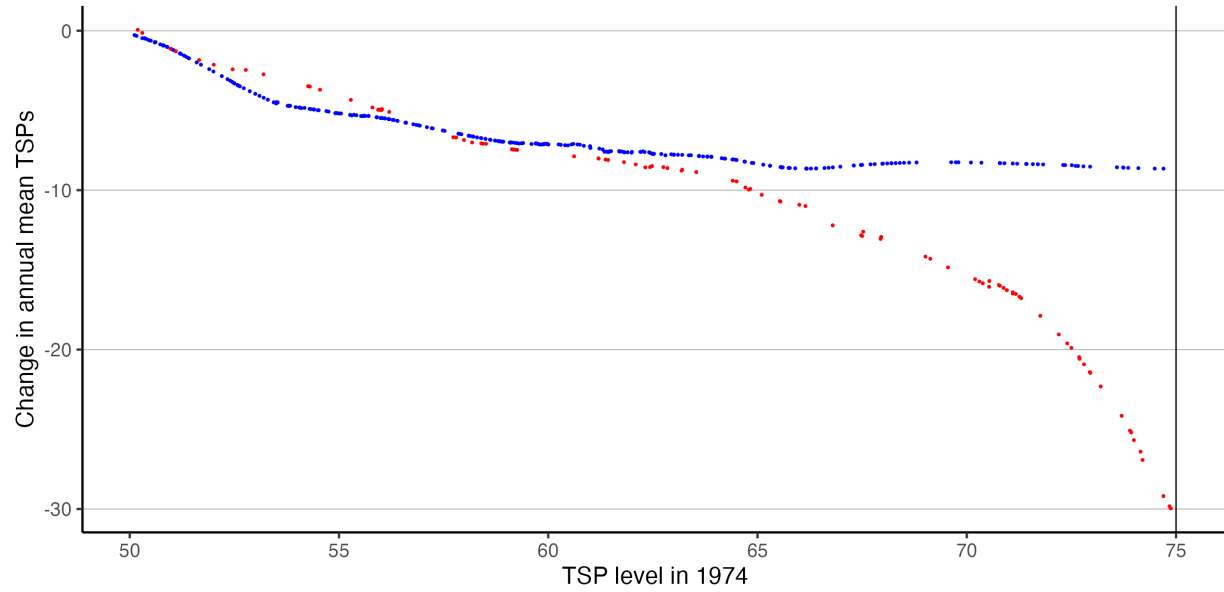


Figure 5: Exercise 2.6, predicted housing price changes and 1974 TSPs levels



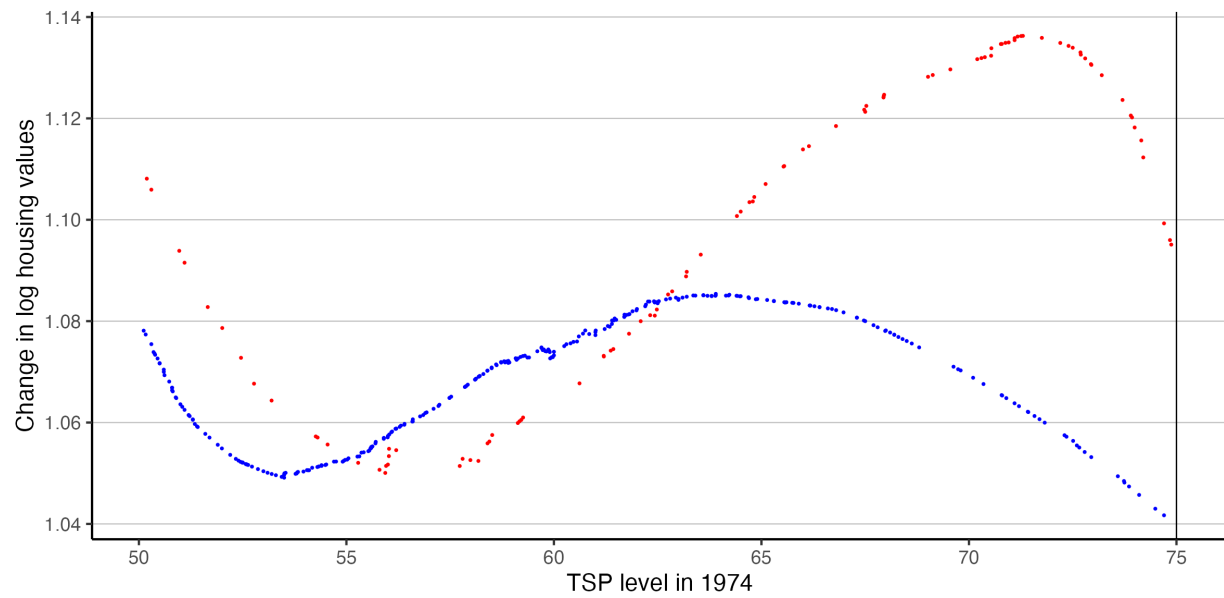
The figure plots (i) the nonparametric bivariate relation between the single-index measure of the housing price changes predicted to occur due to other variables changing and 1974 TSPs levels (green and purple) against (ii) the smoothed housing price changes from (5) (blue and red). A bandwidth of 2 is used in both cases.

Figure 6: Exercise 2.7, pollution



Bandwidth of 6

Figure 7: Exercise 2.7, house



Bandwidth of 6

Tables

Table 1: 1.2 - Climate Impacts, regression 3

	(1) log_emp_farm
tempB0	-0.001283 (0.000125)
temp0to4	-0.000981 (0.000125)
temp4to8	-0.001004 (0.000123)
temp8to12	-0.000929 (0.000115)
temp12to16	-0.000477 (0.000116)
temp16to20	
temp20to24	0.000017 (0.000097)
temp24to28	0.000288 (0.000094)
temp28to32	-0.000235 (0.000117)
tempA32	-0.001437 (0.008392)
tempB0 \times tempA32	-0.000186 (0.000058)
temp0to4 \times tempA32	-0.000118 (0.000088)
temp4to8 \times tempA32	0.000198 (0.000059)
temp8to12 \times tempA32	0.000033 (0.000032)
temp12to16 \times tempA32	0.000050 (0.000038)
temp20to24 \times tempA32	-0.000004 (0.000037)
temp24to28 \times tempA32	0.000049 (0.000034)
temp28to32 \times tempA32	-0.000032 (0.000030)
tempA32 \times tempA32	-0.000047 (0.000033)
constant	6.829056 (0.026407)
N	126619
R^2	0.948

Table 2: Hedonic analysis - Question 1

	(1) dlhouse	(2) dlhouse
dgtsp	0.00243 (0.00031)	0.00033 (0.00023)
ddens		-0.00000 (0.00001)
dmnfcg		-0.76163 (0.18422)
dwhite		-1.40777 (0.10401)
dfeml		-1.81695 (0.74477)
dage65		-2.36604 (0.47919)
dhs		-0.40267 (0.19850)
dcoll		-0.23226 (0.28428)
durban		-0.05510 (0.08695)
dunemp		-4.06593 (0.31417)
dincome		0.00010 (0.00001)
dpoverly		-0.77298 (0.27440)
dvacant		1.22289 (0.19539)
downer		0.25705 (0.14346)
dplumb		0.20047 (0.35452)
drevenue		-0.00024 (0.00007)
dtaxprop		-0.00026 (0.00007)
depend		0.00018 (0.00006)
constant	1.07764 (0.00767)	1.10880 (0.03698)
N	966	966
R^2	0.058	0.584

Table 3: Hedonic analysis - Question 1

	(1) dgtsp
dincome	0.0077 (0.0010)
dunemp	-32.3595 (34.9028)
dmnfcg	112.8088 (21.4736)
constant	-17.4984 (1.4480)
N	966
R^2	0.136

Table 4: Hedonic analysis - Question 2

	(1) dgtsp	(2) tsp7576
tsp7576	-8.14085 (1.19315)	
ddens	0.00260 (0.00171)	-0.00008 (0.00005)
dmnfcg	104.36659 (24.82599)	-0.77796 (0.67531)
dwhite	-3.73083 (14.31255)	-1.84777 (0.38495)
dfeml	208.90189 (101.63079)	-10.43262 (2.74565)
dage65	-85.69488 (65.11719)	0.01062 (1.77254)
dhs	19.46714 (27.02294)	-1.24735 (0.73447)
dcoll	66.56744 (38.67720)	-2.44999 (1.04981)
durban	-5.29556 (11.82536)	-0.11746 (0.32187)
dunemp	-58.67318 (42.68796)	0.61089 (1.16183)
dincome	0.00530 (0.00138)	0.00011 (0.00004)
dpoverly	38.15745 (37.62402)	3.99450 (1.01591)
dvacant	29.79659 (26.58129)	0.78969 (0.72311)
downer	-17.62769 (19.53872)	1.11996 (0.53061)
dplumb	25.01099 (48.23253)	-1.39844 (1.31214)
drevenue	-0.00569 (0.00888)	-0.00011 (0.00024)
dtaxprop	-0.02798 (0.00912)	-0.00091 (0.00025)
depend	0.00461 (0.00874)	0.00033 (0.00024)
constant	-13.45794 (5.05826)	0.67893 (0.13591)
N	966	966
R^2	0.220	0.164

Table 5: Hedonic analysis - Question 3

	(1) dgtsp	(2) dgtsp	(3) dlhouse	(4) dlhouse	(5) dlhouse	(6) dlhouse
tsp7576	-10.19936 (1.17933)	-8.14085 (1.19315)	0.03460 (0.01227)	0.03696 (0.00870)		
ddens		0.00260 (0.00171)		0.00000 (0.00001)		0.00001 (0.00002)
dmnfcg		104.36659 (24.82599)		-0.69669 (0.18100)		-0.22284 (0.26072)
dwhite		-3.73083 (14.31255)		-1.33577 (0.10435)		-1.35271 (0.12523)
dfeml		208.90189 (101.63079)		-1.33529 (0.74098)		-0.38684 (0.96463)
dage65		-85.69488 (65.11719)		-2.39448 (0.47476)		-2.78355 (0.58331)
dhs		19.46714 (27.02294)		-0.34689 (0.19702)		-0.25850 (0.24037)
dcoll		66.56744 (38.67720)		-0.11343 (0.28199)		0.18880 (0.35707)
durban		-5.29556 (11.82536)		-0.05217 (0.08622)		-0.07622 (0.10415)
dunemp		-58.67318 (42.68796)		-4.10931 (0.31123)		-4.37570 (0.38427)
dincome		0.00530 (0.00138)		0.00009 (0.00001)		0.00012 (0.00001)
dpoverly		38.15745 (37.62402)		-0.91878 (0.27431)		-0.74553 (0.32833)
dvacant		29.79659 (26.58129)		1.20134 (0.19380)		1.33662 (0.23558)
downer		-17.62769 (19.53872)		0.20691 (0.14246)		0.12688 (0.17488)
dplumb		25.01099 (48.23253)		0.26405 (0.35166)		0.37761 (0.42656)
drevenue		-0.00569 (0.00888)		-0.00024 (0.00006)		-0.00026 (0.00008)
dtaxprop		-0.02798 (0.00912)		-0.00023 (0.00007)		-0.00036 (0.00008)
depend		0.00461 (0.00874)		0.00017 (0.00006)		0.00019 (0.00008)
dgtsp					-0.00339 (0.00136)	-0.00454 (0.00129)
constant	-10.36680 (0.81884)	-13.45794 (5.05826)	1.02375 (0.00852)	1.07750 (0.03688)	0.98858 (0.02197)	1.01639 (0.05027)
N	966	966	966	966	966	966
R^2	0.072	0.220	0.008	0.591	-0.277	0.393

Table 6: Hedonic analysis - Question 4

	(1) dgtsp	(2) dgtsp	(3) dlhouse	(4) dlhouse	(5) dlhouse	(6) dlhouse
regulation_iv	-13.14764 (1.28796)	-10.35120 (1.44058)	0.03163 (0.01361)	0.04976 (0.01051)		
ddens		0.00165 (0.00171)		0.00001 (0.00001)		0.00002 (0.00002)
dmnfcg		82.81606 (25.04792)		-0.59139 (0.18267)		-0.19325 (0.26156)
dwhite		-16.09102 (14.61217)		-1.27233 (0.10657)		-1.34969 (0.12736)
dfeml		199.72153 (101.45782)		-1.26845 (0.73993)		-0.30829 (0.97451)
dage65		-125.71700 (65.18880)		-2.20210 (0.47542)		-2.80648 (0.59259)
dhs		-2.86815 (27.29043)		-0.23680 (0.19903)		-0.25058 (0.24432)
dcoll		41.53536 (38.97434)		0.01225 (0.28424)		0.21193 (0.36170)
durban		-5.33985 (11.79525)		-0.05171 (0.08602)		-0.07738 (0.10597)
dunemp		-44.58942 (42.65563)		-4.17835 (0.31109)		-4.39271 (0.39020)
dincome		0.00531 (0.00138)		0.00009 (0.00001)		0.00012 (0.00001)
dpoverly		15.05514 (37.24900)		-0.81640 (0.27166)		-0.74403 (0.33411)
dvacant		42.51517 (26.63062)		1.13848 (0.19422)		1.34287 (0.23955)
downer		-7.26578 (19.63143)		0.15466 (0.14317)		0.11973 (0.17764)
dplumb		20.36295 (48.13271)		0.28944 (0.35103)		0.38734 (0.43383)
drevenue		-0.00398 (0.00886)		-0.00025 (0.00006)		-0.00027 (0.00008)
dtaxprop		-0.02594 (0.00906)		-0.00024 (0.00007)		-0.00036 (0.00008)
depend		0.00046 (0.00871)		0.00019 (0.00006)		0.00019 (0.00008)
dgtsp					-0.00241 (0.00112)	-0.00481 (0.00125)
constant	-11.54289 (0.68702)	-9.77051 (5.14272)	1.03144 (0.00726)	1.05829 (0.03751)	1.00366 (0.01841)	1.01132 (0.05059)
N	966	966	966	966	966	966
R ²	0.098	0.224	0.006	0.593	-0.173	0.371