

Teoria degli Algoritmi

Corso di Laurea Magistrale in Matematica Applicata

a.a. 2020-21



SAPIENZA
UNIVERSITÀ DI ROMA

Gabriele Tolomei

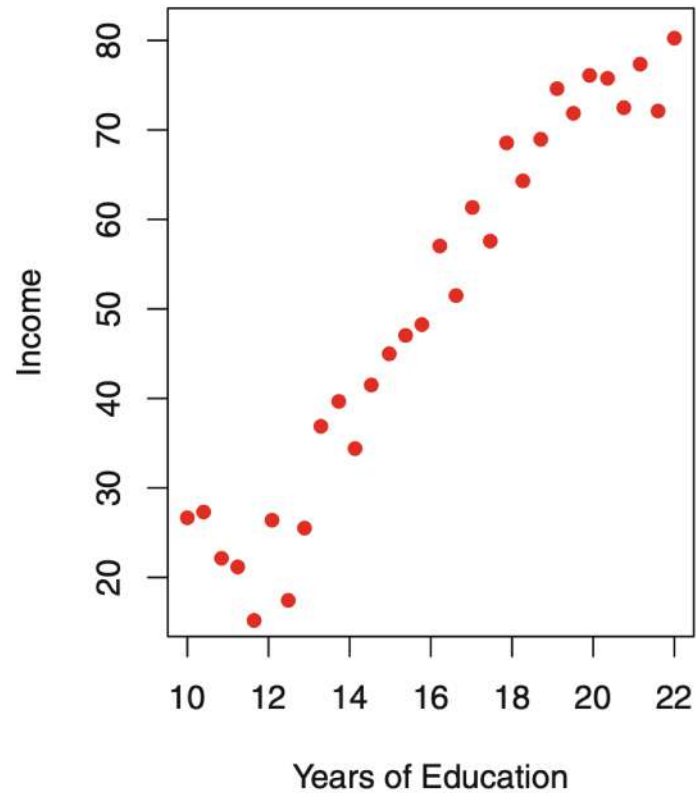
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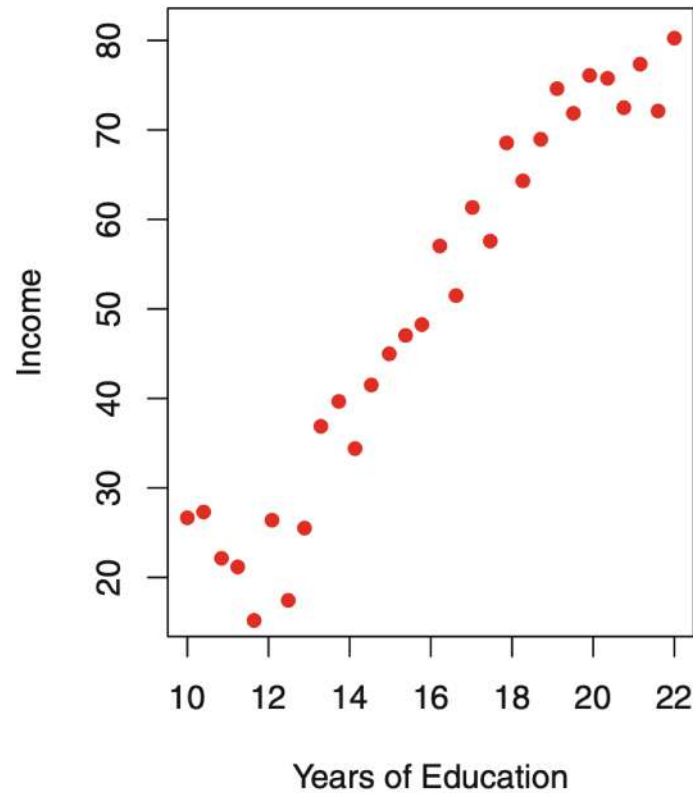
LINEAR REGRESSION

Example: Y =Income vs. X =Education

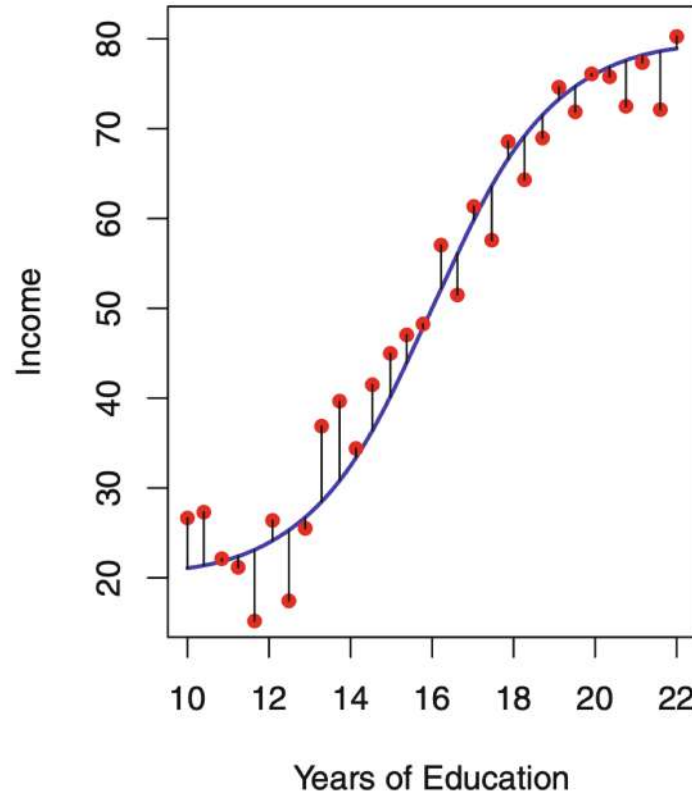


Observations
(simulated)

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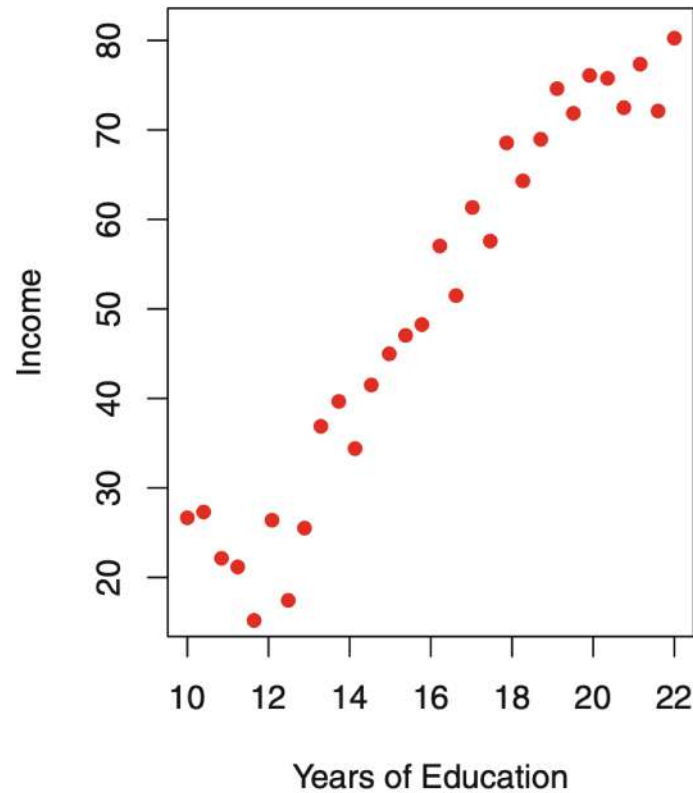
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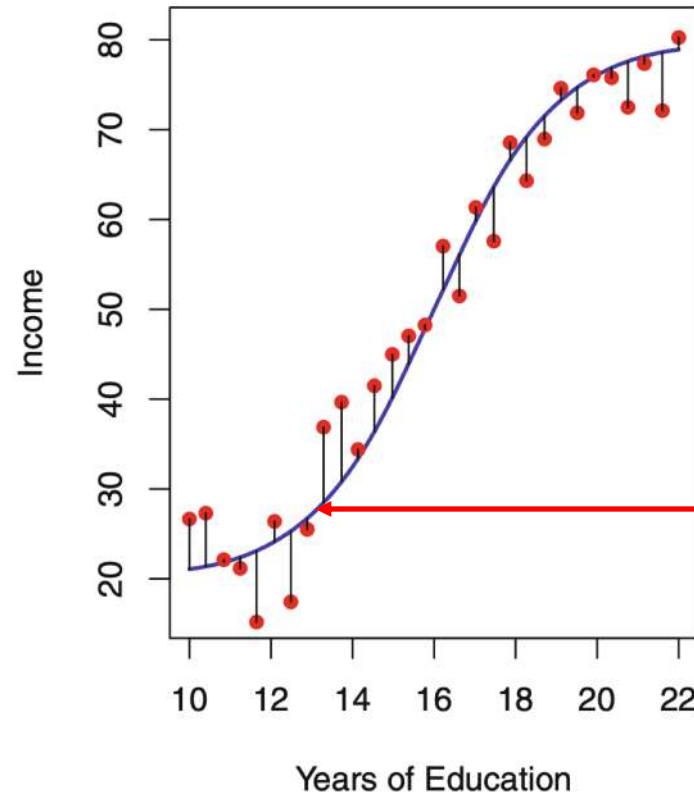
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between X and Y

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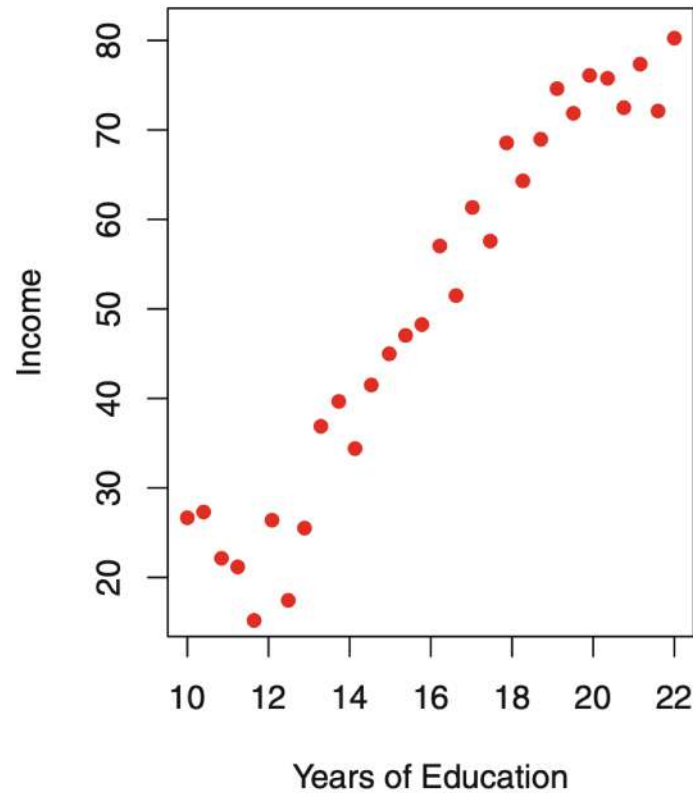
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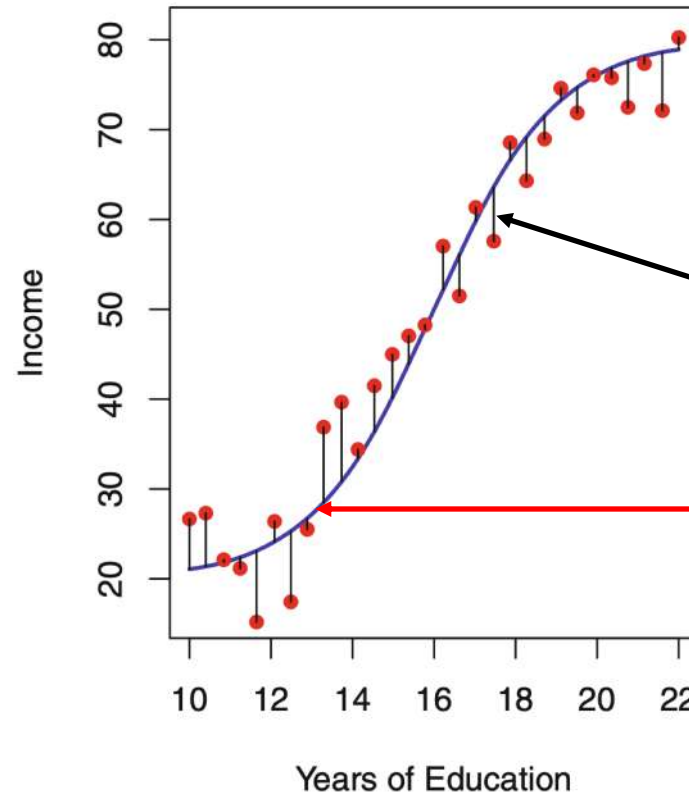
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True yet unknown relationship
between X and Y

Gaussian error term

$$Y = f(X) + \varepsilon$$

Assumptions

- There exists a relationship between \mathcal{X} (features) and \mathcal{Y} (values)

$$\mathcal{Y} = f(\mathcal{X}) + \epsilon$$

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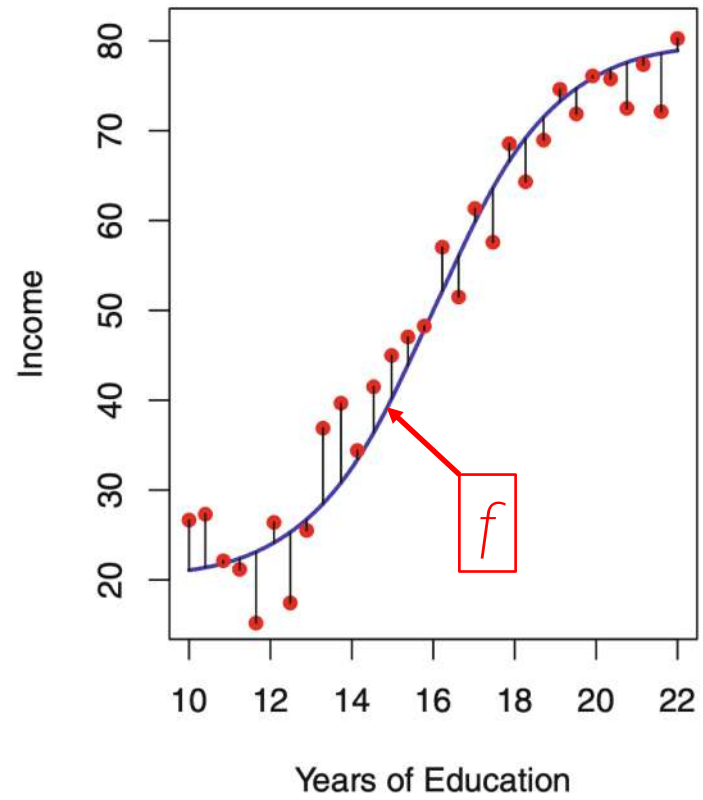
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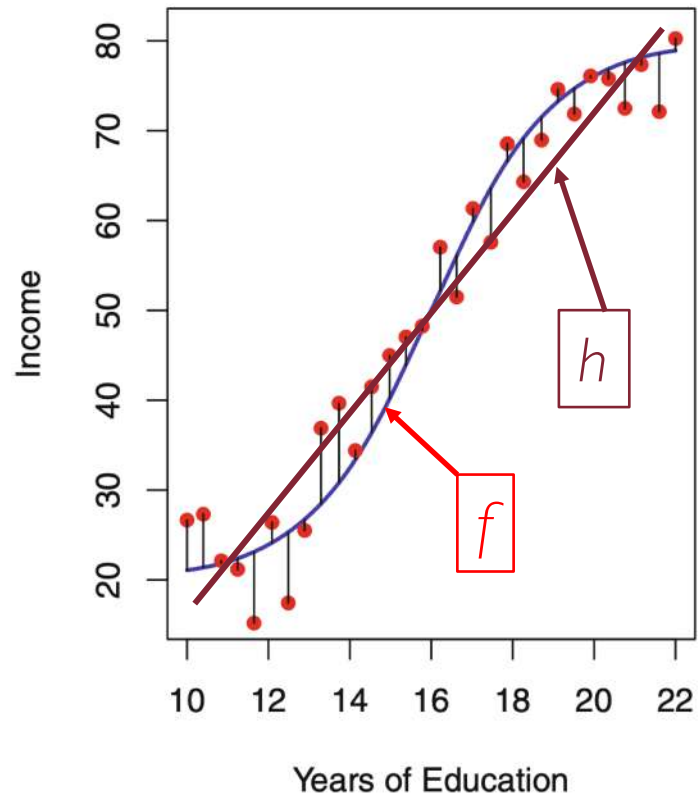
$$Y = f(X) + \epsilon$$

- f is some fixed but unknown function of X
- ϵ is a random error term, which is independent of X and has 0-mean
- In this formulation, f represents the systematic information that X provides about Y

Goal

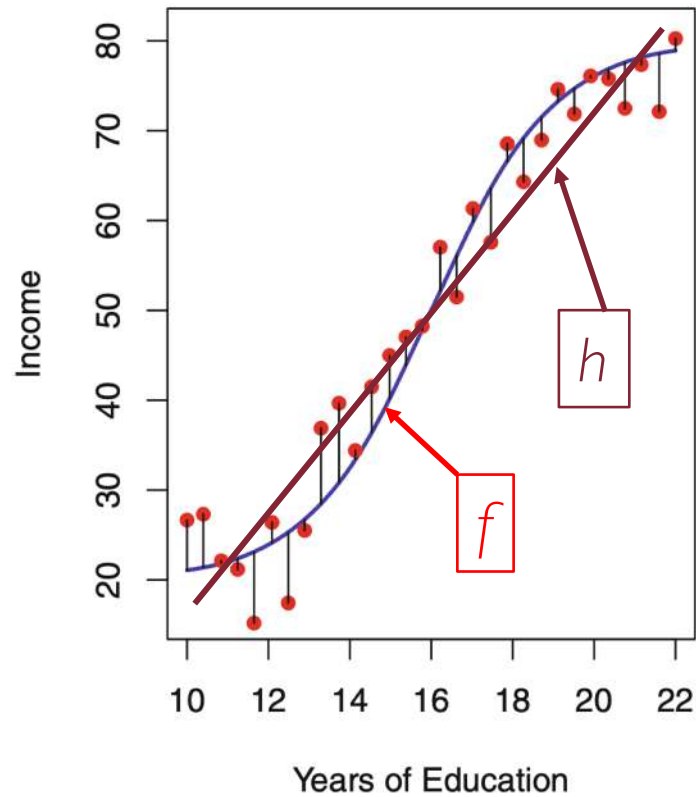


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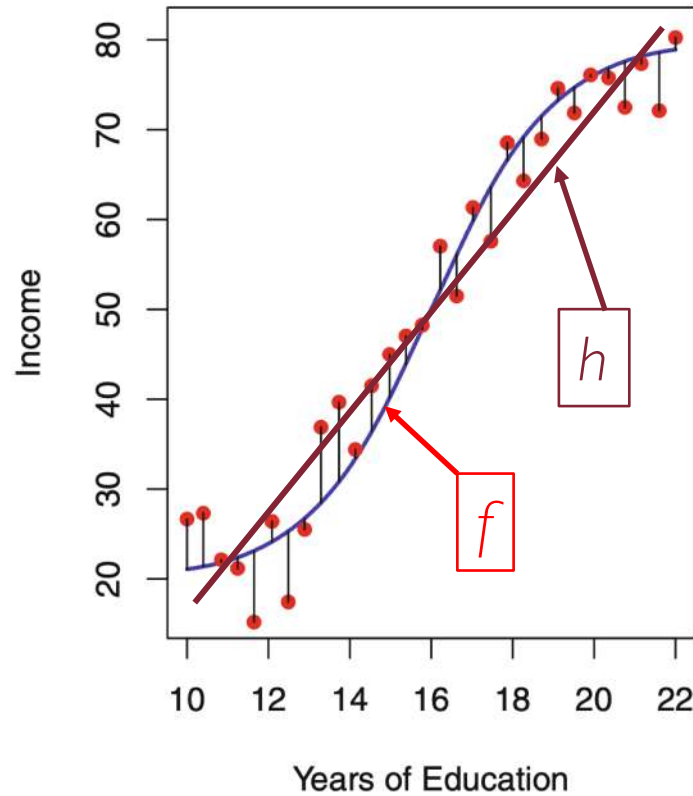
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- Choose h from a specific hypothesis space H (i.e., linear functions)
- Use a dataset D of observations to learn h

$$h(X) \sim f(X)$$

Recap of Notation

$$\mathcal{X} \subseteq \mathbb{R}^n$$

input feature space

$$\mathcal{Y}$$

output space

$$\mathcal{Y} \subseteq \mathbb{R}$$

real-value label(**regression**)

$$(\mathbf{x}_i, y_i)$$

i -th labeled instance

$$\mathbf{x}_i = (x_{i,1}, \dots, x_{i,n}) \in \mathcal{X}$$

n -dimensional feature vector of the i -th instance

$$y_i \in \mathcal{Y}$$

label of the i -th instance

$$\mathcal{D} = \{(\mathbf{x}_1, y_1), \dots, (\mathbf{x}_m, y_m)\}$$

dataset of m i.i.d. labeled instances

The Hypothesis Space H

The hypothesis space is defined as follows:

$$\mathcal{H} = \{h_{\boldsymbol{\theta}} : \mathcal{X} \mapsto \mathcal{Y} \mid h_{\boldsymbol{\theta}}(\mathbf{x}) = \theta_0 x_0 + \theta_1 x_1 + \dots + \theta_n x_n\}$$

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$\boldsymbol{\theta}$ $n+1$ -dimensional vector of model parameters

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Among all the possible instantiations of $\boldsymbol{\theta}$ the learning algorithm selects $\boldsymbol{\theta}^*$ as the one which minimizes a **loss function** measured on D

Residuals of Sum Squares (RSS)

$$y_i = f(\mathbf{x}_i) + \epsilon_i \quad i\text{-th observation}$$

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$$\text{RSS}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \sum_{i=1}^m e_i^2 = \sum_{i=1}^m (\hat{y}_i - y_i)^2 = \sum_{i=1}^m (h_{\boldsymbol{\theta}}(\mathbf{x}_i) - y_i)^2$$

Ordinary Least Squares (OLS)

- Remember that the supervised learning problem can be generally defined as the following optimization problem

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$$\begin{aligned}\text{MSE}(h_{\theta}, \mathcal{D}) &= \frac{1}{m} \text{RSS}(h_{\theta}, \mathcal{D}) = \\ &= \frac{1}{m} \sum_{i=1}^m (h_{\theta}(\mathbf{x}_i) - y_i)^2\end{aligned}$$

The OLS Learning Algorithm

OLS aims at solving the following optimization problem:

$$\begin{aligned} h^* = h_{\theta^*} &= \operatorname{argmin}_{\theta} \operatorname{MSE}(h_{\theta}, \mathcal{D}) = \\ &= \operatorname{argmin}_{\theta} \left[\frac{1}{m} \sum_{i=1}^m (h_{\theta}(\mathbf{x}_i) - y_i)^2 \right] \end{aligned}$$

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How do we solve that?

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The function to minimize can be proven **convex**

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The Hessian
matrix is positive
semi-definite

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- If the function is convex (concave) finding the minimum (maximum) can be done just by computing the first derivative and set it to 0
- In the case of a multivariate function, this generalizes to compute the gradient (∇) of the function and set it to 0

The Gradient ∇

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Solving $\nabla f = \mathbf{0}$ means finding the n -dimensional vector \mathbf{x} such that:

$$\nabla f = \left(\frac{\partial f}{\partial x_1}, \frac{\partial f}{\partial x_2}, \dots, \frac{\partial f}{\partial x_n} \right) = \underbrace{(0, 0, \dots, 0)}_n = \mathbf{0}$$

Solving the Optimization Problem

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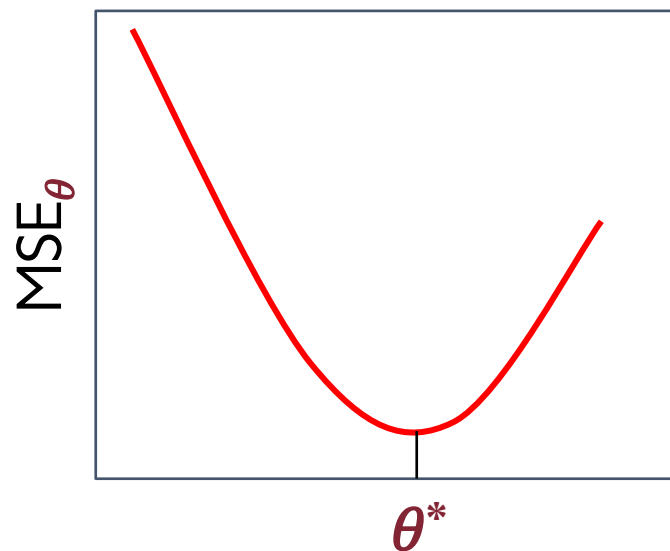
Each term of the summation is a multivariate linear function of the model parameters $\boldsymbol{\theta}$

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Convex functions have a **unique local minimum**, which therefore happens to be the **global minimum**

Computing the Gradient of MSE

$$\nabla \text{MSE}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \nabla \left[\frac{1}{m} \sum_{i=1}^m (h_{\boldsymbol{\theta}}(\mathbf{x}_i) - y_i)^2 \right]$$

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$$\frac{\partial f}{\partial t}(\alpha t) = \alpha \frac{\partial f}{\partial t}(t) \quad \alpha \in \mathbb{R}, \text{ constant}$$

scalar multiple rule

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$$\nabla \text{MSE}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \frac{1}{m} \left[\sum_{i=1}^m \nabla (h_{\boldsymbol{\theta}}(\mathbf{x}_i) - y_i)^2 \right]$$

Computing the Gradient of MSE (1 instance)

To make things easier, let's assume the dataset D contains a single instance (\mathbf{x}, y)

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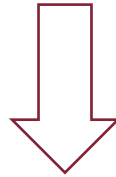
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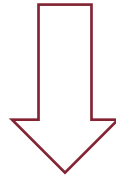
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Vectorized Notation

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The resulting gradient is an $(n+1)$ -dimensional vector as expected!

Setting the Gradient Equal to Zero

$$\nabla \text{MSE}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \begin{bmatrix} 2(\boldsymbol{\theta}^T \cdot \mathbf{x} - y) \\ 2(\boldsymbol{\theta}^T \cdot \mathbf{x} - y)x_1 \\ \vdots \\ 2(\boldsymbol{\theta}^T \cdot \mathbf{x} - y)x_n \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} = \mathbf{0}$$

Setting the Gradient Equal to Zero

$$\nabla \text{MSE}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \begin{bmatrix} 2(\boldsymbol{\theta}^T \cdot \mathbf{x} - y) \\ 2(\boldsymbol{\theta}^T \cdot \mathbf{x} - y)x_1 \\ \vdots \\ 2(\boldsymbol{\theta}^T \cdot \mathbf{x} - y)x_n \end{bmatrix} = \begin{bmatrix} 0 \\ 0 \\ \vdots \\ 0 \end{bmatrix} = \mathbf{0}$$

We need to solve a system of $n+1$ linear equations with $n+1$ variables

$$2(\boldsymbol{\theta}^T \cdot \mathbf{x} - y)x_j = 0 \quad \forall j \in \{0, 1, \dots, n\}$$

Computing the Gradient of MSE (m instances)

In the general case where the dataset \mathcal{D} contains a m instances

$$\nabla \text{MSE}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \frac{2}{m} \left[\sum_{i=1}^m \left(h_{\boldsymbol{\theta}}(\mathbf{x}_i) - y_i \right) \nabla \left(h_{\boldsymbol{\theta}}(\mathbf{x}_i) - y_i \right) \right]$$

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$$\frac{2}{m} \left[(\boldsymbol{\theta}^T \cdot \mathbf{x}_1 - y_1)x_{1,j} + \dots + (\boldsymbol{\theta}^T \cdot \mathbf{x}_m - y_m)x_{m,j} \right] = 0 \quad \forall j \in \{0, \dots, n\}$$

Matrix Notation

$$\mathbf{X} = \underbrace{\begin{bmatrix} x_{1,0} & x_{1,1} & \dots & x_{1,n} \\ x_{2,0} & x_{2,1} & \dots & x_{2,n} \\ \vdots & \vdots & \vdots & \vdots \\ x_{m,0} & x_{m,1} & \dots & x_{m,n} \end{bmatrix}}_{m \times n+1 \text{ feature matrix}} = \begin{bmatrix} -\mathbf{x}_1^T - \\ -\mathbf{x}_2^T - \\ \vdots \\ -\mathbf{x}_m^T - \end{bmatrix}$$

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$n+1$ -dimensional **parameter vector**

$$\mathbf{y} = \underbrace{\begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_m \end{bmatrix}}$$

m -dimensional **target vector**

Vectorized Form of the Optimization Problem

$$h^* = h_{\theta^*} = \operatorname{argmin}_{\theta} \left[\underbrace{\frac{1}{m} \|\mathbf{X} \cdot \boldsymbol{\theta} - \mathbf{y}\|^2}_{\text{MSE}(h_{\theta}, \mathcal{D})} \right]$$

Vectorized Form of the Gradient of MSE

$$\nabla \text{MSE}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \frac{2}{m} \mathbf{X}^T (\mathbf{X} \cdot \boldsymbol{\theta} - \mathbf{y})$$

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$$\boldsymbol{\theta} = \mathbf{X}^\dagger \cdot \mathbf{y}$$

$\mathbf{X}^\dagger = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$ is the **pseudo-inverse** of \mathbf{X}

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 - Therefore, its common to say that this probability is zero

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- OLS is also known as one-step learning as there exists a closed-form (i.e., analytical) solution to the convex optimization problem
- However, other choices of loss functions (even if convex) may need an **iterative** approach to get to a (local) minimum
- Though in general $n \ll m$, computing the inverse of an n -by- n matrix is still a costly operation ($O(n^3)$ time complexity)

Errors vs. Residuals

Subtle yet important difference between **errors** and **residuals**

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MSE is computed from residuals, not unobservable errors!

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- **Linearity** → Linear relationship between the features and the response
 - Only a restriction on the parameters; features themselves can be arbitrarily combined using non-linear transformations
- **Error independence** → Error terms ε_i are uncorrelated with each other
 - Knowing that ε_i is positive (negative) gives no information on the sign of ε_{i+1}

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- **Homoscedasticity** → Different values of the response variable have the same variance in their errors, regardless of the feature values
 - In practice, this does not hold when the response varies over a wide scale
- **No Multicollinearity** → There must not be two or more features whose values are perfectly correlated with each other
 - The feature matrix \mathbf{X} must have full column rank n
 - If \mathbf{X} is full column rank n then $\mathbf{X}^T\mathbf{X}$ is always invertible
 - It can be shown that if $\mathbf{X}^T\mathbf{X}\mathbf{u} = \mathbf{0}$ for some vector \mathbf{u} , then $\mathbf{u} = \mathbf{0}$ (trivial solution)

Checking OLS Assumptions

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- A good way to assess the OLS assumptions hold is to use residual plots
- Plotting residuals against each feature and/or the predicted value may help spot:
 - Non-linearity
 - Correlation between error terms
 - Non-constant variance of error terms (i.e., heteroscedasticity)
 - ...

Assessing the Quality of a Model

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R^2 statistic

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Recall that every observation of the target variable y_i is associated with an error term ϵ_i

$$y_i = \underbrace{\theta_0 x_{i,0} + \theta_1 x_{i,1} + \dots + \theta_n x_{i,n}}_{h_{\boldsymbol{\theta}}(\mathbf{x}_i) \approx f(\mathbf{x}_i)} + \epsilon_i$$

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Even if we were able to find the exact parameters of the true f , we would not be able to perfectly predict y_i from \mathbf{x}_i

Residual Standard Error (RSE)

RSE is an estimate of the standard deviation of ε

$$\text{RSE}(h_{\boldsymbol{\theta}}, \mathcal{D}) = \sqrt{\frac{1}{\underbrace{m - n - 1}_{\text{degrees of freedom}}} \underbrace{\sum_{i=1}^m (\hat{y}_i - y_i)^2}_{\text{RSS}}}$$

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A measure of the **lack** of fit of the model to the data
the lower the better

Degrees of Freedom

$$y_i = \theta_0 + \theta_1 x_{i,1} + \epsilon_i$$

Degrees of Freedom

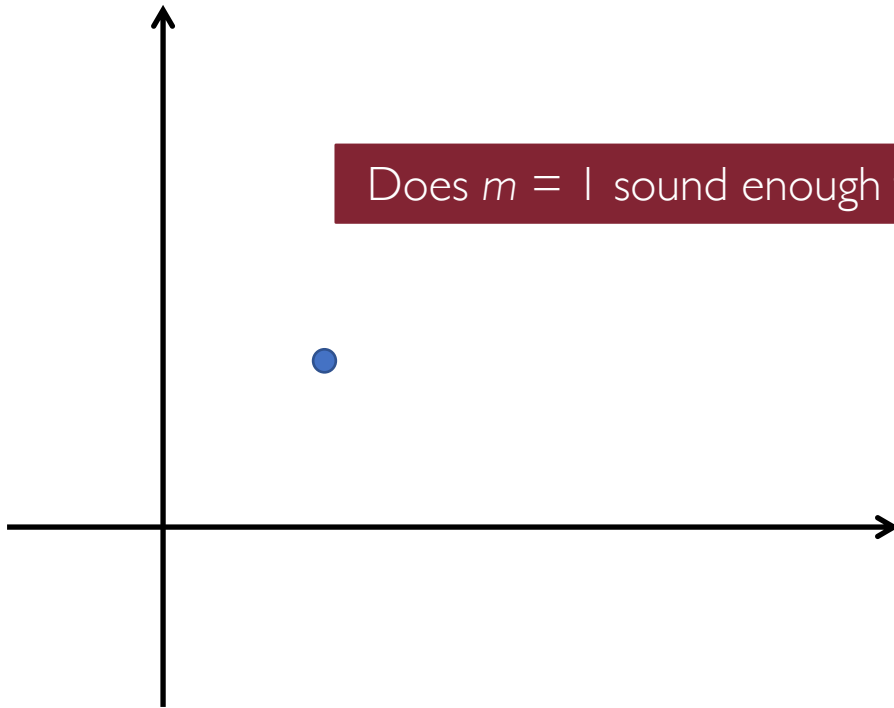
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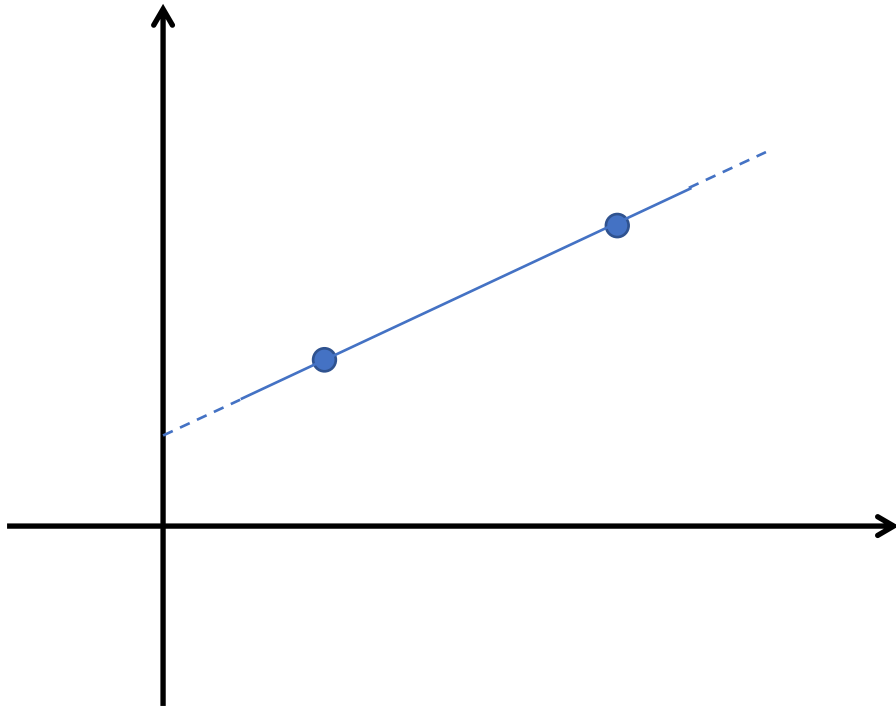
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Does $m = 1$ sound enough to fit a line?

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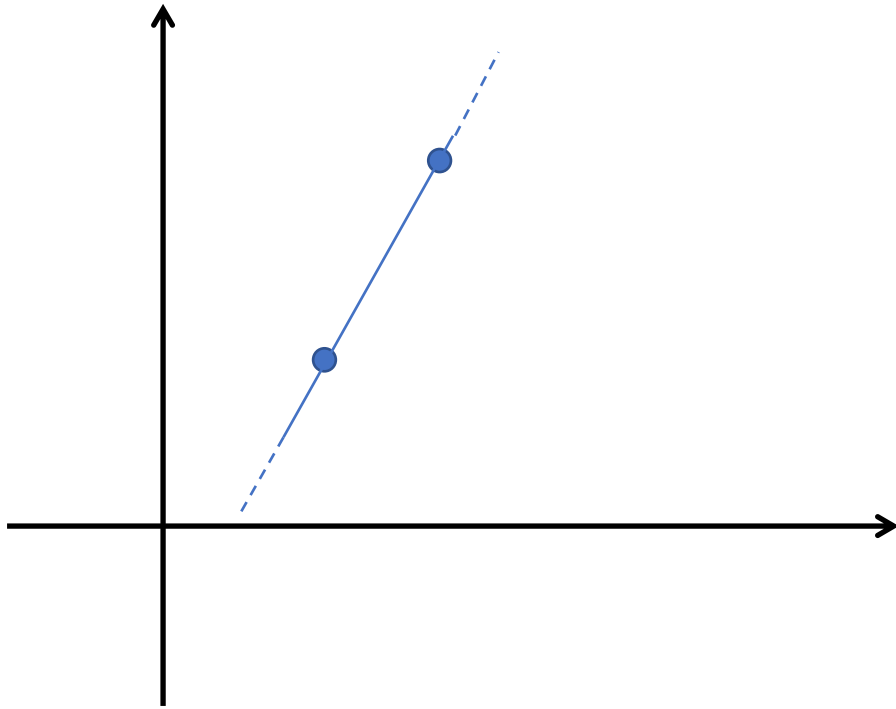


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With 2 data points I am always able to fit a perfect line

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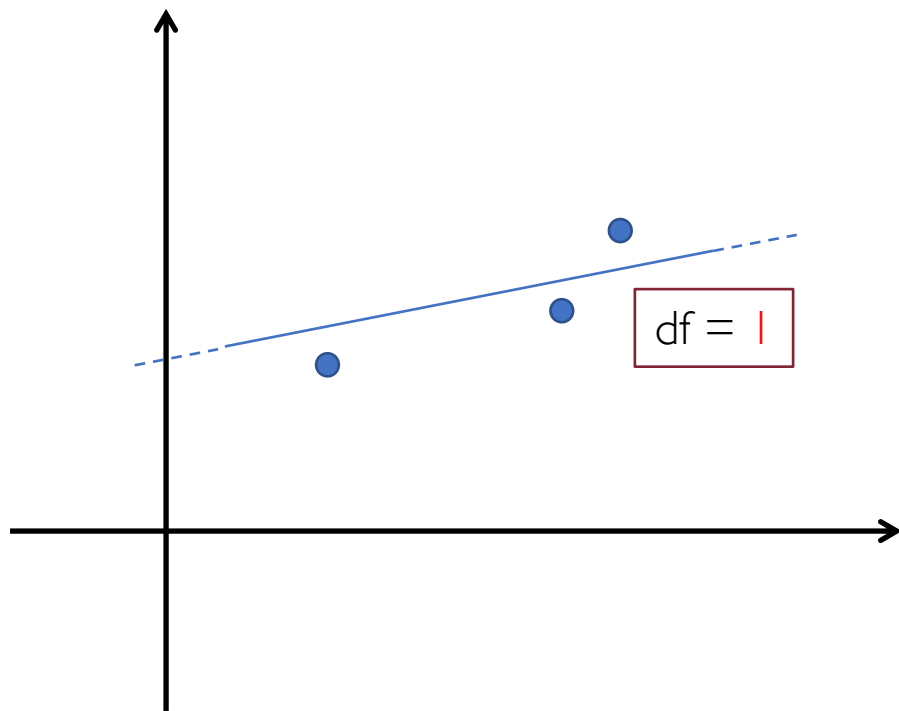
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Problem is that my fitted line may drastically change depending on where the second point is located!

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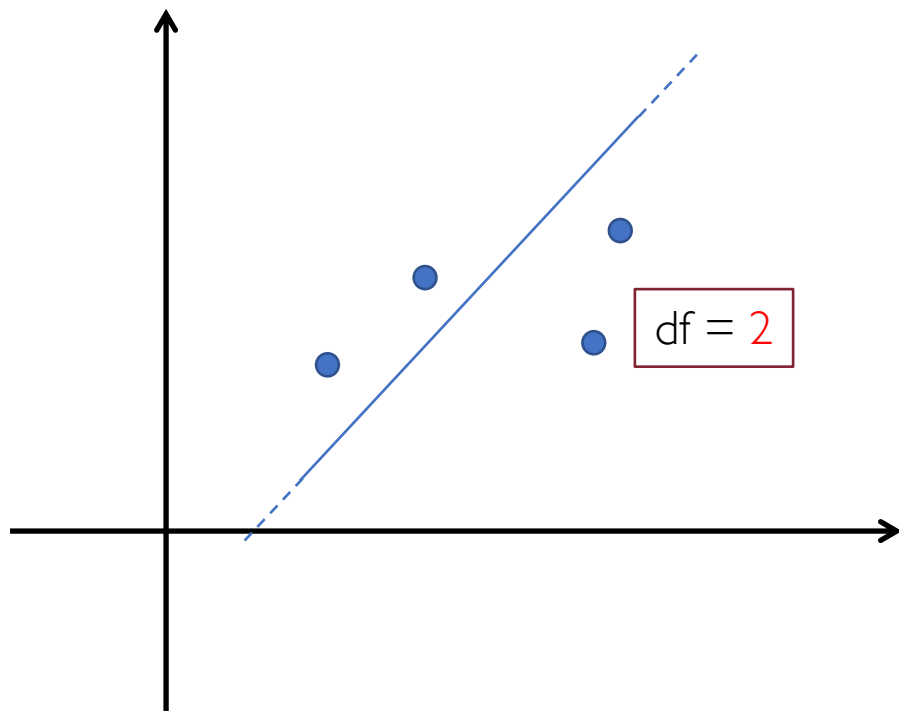
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$$\text{df} = \underbrace{m}_{\text{\#observations}} - \underbrace{n}_{\text{\#features}} - \underbrace{1}_{\text{intercept}}$$

R² Statistic

$$R^2 = 1 - \frac{\text{RSS}}{\text{TSS}} = 1 - \frac{\sum_{i=1}^m (\hat{y}_i - y_i)^2}{\sum_{i=1}^m (y_i - \bar{y})^2}$$

TSS measures the total variance in the response Y before the regression takes place

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RSS measures the amount of variability that is left unexplained after performing the regression

R² measures the proportion of variability in Y that can be explained using X

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- An R^2 statistic that is close to 1 indicates that a large proportion of the variability in the response has been explained by the regression
- The larger R^2 the better is the linear regression model
- R^2 is easier to interpret than RSE as it always ranges between 0 and 1

Adjusted R^2 Statistic

- Fixing the sample size m , RSS decreases (or, at worst, it stays the same) as more variables are added to the fitted model

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- Fixing the sample size m , RSS decreases (or, at worst, it stays the same) as more variables are added to the fitted model
- R^2 always increases as more variables are added (as df decreases!)
- We need a way to adjust for that, otherwise we could get a better model by simply adding useless features to it!

$$R_{\text{adj}}^2 = 1 - \frac{\frac{\text{RSS}}{m-n-1}}{\frac{\text{TSS}}{m-1}}$$

Adjusted R^2 Statistic

- Maximizing the adjusted R^2 is equivalent to minimizing $RSS/(m-n-1)$

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$\lambda > 0; \alpha = 0$ **Ridge** (L2-regularization only)

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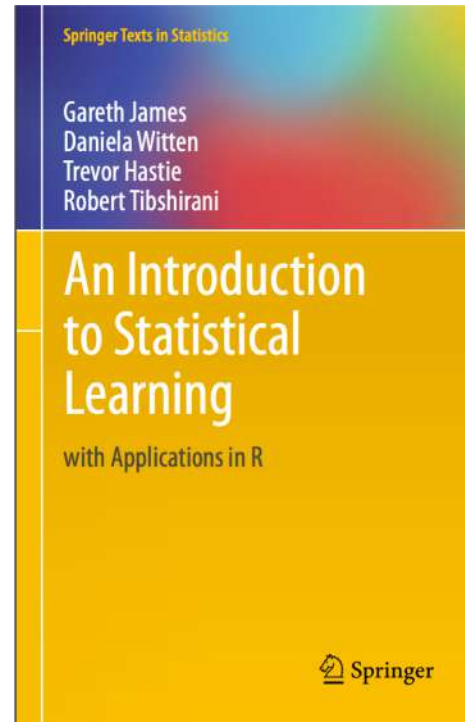
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- Regularization to prevent overfitting: Elastic Net, LASSO, Ridge

Further Readings

An Introduction to Statistical Learning [Chapter 3]



Freely available at:

<https://www.ime.unicamp.br/~dias/Intoduction%20to%20Statistical%20Learning.pdf>