

The Effect of Electoral Systems on Voter Turnout: Evidence from a Natural Experiment*

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I exploit the unique institutional framework of Spanish local elections, where municipalities follow different electoral systems depending on their population size, as mandated by a national law. Using a regression discontinuity design, I compare turnout under closed list proportional representation and under an open list, plurality-at-large system where voters can vote for individual candidates from the same or different party-lists. I find that the open list system increases turnout by between 1 and 2 percentage points. The results suggest that open list systems, which introduce competition both across and within parties, are conducive to more voter turnout.

Voter turnout varies widely across countries. How much of that variation can be explained by differences in the electoral system? What is the causal effect of the electoral system on voter turnout? Do citizens participate more when they are allowed to decide which individual candidates get elected (open list (OL) systems), than in elections where they vote for a party-list (closed list (CL) systems)? These questions have attracted the attention of economists and political scientists for a long time (see Blais 2006; Blais and Aarts 2006; Geys 2006 for reviews of the literature).

There are two main characteristics of electoral systems: the degree of proportionality, determined mainly by the electoral formula and the district magnitude, and the ballot structure (OL versus CL). While the literature has extensively studied the former and generally agreed that more proportionality increases voter turnout (see, e.g., Jackman and Miller 1995; Schram and Sonnemans 1996; Blais and Dobrzynska 1998; Ladner and Milner 1999; Bowler, Donovan and Brockington 2003; Fornos, Power and Garand 2004; Milner and Ladner 2006; Selb 2009; St-Vincent 2013), little is known about the effects of the ballot structure.¹ This is unfortunate, as changes between closed and OLs have been debated in several countries.² Understanding their influence on voter turnout is crucial to enlighten that debate and to the design of electoral systems.

Previous empirical evidence on this issue is mostly based on cross-country regressions, where small sample sizes and the endogeneity of electoral rules raise concerns about the causal interpretation of the estimates, as it is difficult to isolate the effect of the voting systems from other economic, cultural or institutional variables that may also affect

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¹ As Blais and Aarts (2006) put it, “one aspect of electoral systems that has been neglected is the ballot structure” and “we know precious little about the impact of ballot structure on turnout.”

² For example, Japan introduced an open list preference vote in elections for the upper house in 2001, while Italy abandoned an open list proportional representation system in 1993.

voter turnout.³ To the best of my knowledge, this is the first paper to use a quasi-experimental design to estimate how OL systems affect voter turnout. I focus on a previously unexplored setting in which electoral systems are exogenously assigned. In Spain, municipalities with >250 inhabitants elect a city council by CL proportional representation while municipalities with 250 or fewer inhabitants elect a city council in a plurality-at-large, OL election, in which voters can vote for up to four individual candidates from the same or different party-lists. While the two systems differ in both proportionality and ballot structure (OL versus CL), in the Results section I show that the difference in proportionality is small and provide evidence that the effect on turnout is driven by the difference in the ballot structure.

The institutional framework of Spanish local elections is a unique opportunity to study the causal effect of the electoral system on voter turnout for several reasons. First, the electoral system a municipality has to follow is determined by a *national* law as a function of the population size of the municipality, which reduces endogeneity concerns. Second, the number of observations is very high (around 72,000), as there are >8000 municipalities in Spain and election results by municipality are available for nine election years. Furthermore, there are many municipalities with a population size close to the population threshold that separates the two electoral systems (around 700 municipalities in a window of 50 inhabitants around the threshold). Third, all the municipalities under any of the electoral systems follow the exact same electoral system. This is in opposition to cross-country studies, where it is inevitable to pool into the same electoral system a set of systems that are only somewhat similar. Fourth, unlike what is often the case in situations where a policy changes at a municipal population threshold (Campa 2012; Grembi, Nannicini and Troiano 2012; Eggers 2015), no other rule changes at the threshold. Therefore, we can be confident in attributing the differences in outcomes between municipalities at each side of the threshold to the electoral system and not to some other regulation.

To carry out the analysis, I have collected a rich data set with results from all elections held in Spain since the restoration of democracy in 1977. Combining regression discontinuity design and fixed effects estimation, I find that the OL system increases voter turnout with respect to the CL system by between 1 and 2 percentage points. There are many channels that may be conducive to these results—for example, rational-choice calculations about the pivotality of votes or perceived fairness of the systems. I provide evidence that the differences in turnout are at least partially driven by the number of parties that enter competition. A higher number of parties in competition may in turn affect voter turnout by increased aggregate mobilization efforts and by providing voters with a more compelling set of options. I find that the OL system increases by 0.35 the average number of lists in competition. This effect is most likely driven by the fact that the OL system makes it much easier for popular candidates from small parties to get elected, as voters can choose individual candidates from the same or different party-lists.

An issue that requires special attention is that there exists some evidence that municipalities may be able to partially control their population size, as some sorting is observed around the threshold. I carefully study this question in the Robustness section, assessing the validity of the empirical strategy and checking the robustness of the results to donut regression discontinuity design estimation in the spirit of Barreca et al. (2011), by dropping observations within a window where the sorting is most likely to occur.

The rest of the paper is organized as follows. The next section reviews the literature on the impact of electoral systems on voter turnout, analyzing both theoretical predictions and the

³ Countries that use open list systems include Belgium, Brazil, Chile, Colombia, Denmark, Finland, Indonesia, Iraq, Ireland, Italy, Japan, the Netherlands, Slovakia, Sweden and Switzerland.

available empirical evidence. The third section provides background on the Spanish electoral systems. The data and empirical strategy are laid out in the fourth section. The fifth section presents the main results and looks into the mechanisms at work. The robustness of the results is analyzed in the penultimate section. The final section concludes.

LITERATURE REVIEW

This paper contributes to the literature that tries to explain voter turnout (see, e.g., Downs 1957; Riker and Ordeshook 1968) and, in particular, to the literature that studies the link between the electoral system and voter turnout. The electoral systems that I compare in this paper differ mainly in their ballot structure. In this section, I review the theoretical arguments and the previous empirical evidence on the effect of the ballot structure on voter turnout.

In some elections, voters can express a preference for candidates within the party-lists (OL systems), while in others they are limited to choose between different lists (CL systems). There are opposing views in the literature about how the use of OL versus CL systems should affect voter turnout.

On the one hand, it has been argued that OL systems should increase turnout. Mattila (2003) argues that as voters can choose the candidate they wish to vote for, they are likely to feel more satisfied with the act of voting. Along the same lines, Karvonen (2004) indicates that voting for individual candidates makes the election more personal and concrete, and that both elements should provide a stimulus for active electoral participation. Supporting this hypothesis, Hix and Hagemann (2009) find that citizens in EU states who use OL are almost 5 percent more likely to be contacted by candidates or parties than citizens in member states with CL systems. They are also >20 percent more likely to be contacted by candidates or parties and about 15 percent more likely to feel well informed about the elections than citizens in states with CL systems.

On the other hand, Robbins (2010) hypothesizes that turnout should be higher in CL systems. His argument is that in OL systems parties may not exert the same level of resources to solicit support or mobilize voters as in CL systems. Individual candidates, for their part, will appeal to their supporters but will likely avoid mobilization strategies that involve the entire population. In CL systems, on the contrary, “parties place greater emphasis on mobilizing voters everywhere in hopes of soliciting additional support. After all, if they construct the list, then they are responsible for the success of their candidates and will devote more time, energy and resources calling individuals to the polls.”

The empirical evidence for the effect of the ballot structure is scarce, probably due to limited cross-country variation. Hix and Hagemann (2009) find that voters are almost 10 percent more likely to cast their votes on election day in OL systems. Robbins (2010), on the contrary, finds that OL decreases turnout levels. Mattila’s (2003) empirical findings, with data from elections to the European Parliament, show that a variable indicating a CL system is not significant. The empirical analyses in Blais and Aarts (2006), Dos Santos (2007) and Karvonen (2004) also conclude that there is not sufficient evidence to support the hypothesis of a positive correlation between preferential voting and electoral participation. Eggers (2015) uses a discontinuity in the electoral rules of French local elections and finds that a CL proportional representation system leads to more turnout than an OL plurality system, but he attributes the effect to the different proportionality of the systems. In sum, the empirical evidence is non-conclusive: the effect of the ballot structure on turnout is still an open question.

Finally, another strand of literature has studied the effects of OL systems on other outcomes.⁴ Farrell and McAllister (2006) find that preferential voting systems where voters are given more

⁴ This discussion is partially based on the literature review in Negri (2014).

TABLE 1 *Services of Spanish Municipalities by Population*

Municipalities	Services
All	Public lighting, cemeteries, waste collection, street cleaning household water supply, sewerage, access to villages, paving roads, food and beverage control
>5000	Public park, public library, waste treatment, organization of markets
>10,000	Civil protection, social services, prevention and fire-fighting, sports facilities for public use
>20,000	Urban public passenger transport and environmental protection

Source: Law 5/1985 and Campa (2012)

Note: the table shows the services that Spanish municipalities are required to provide, as a function of population size.

freedom in completing the ballot paper lead to higher satisfaction with democracy. Other papers have found that OL systems increase the value of personal reputation with respect to party reputation by enhancing intra-party competition and electoral uncertainty (Carey and Shugart 1995; Chang 2005) or by inducing voters to focus more on candidates' characteristics and less on parties' positions (Shugart, Valdini and Suominen 2005). Ames (Ames 1995a; Ames 1995b) studies the use of OL elections in Brazil and supports these conclusions by highlighting the very weak role played by national parties in the country. Persson, Tabellini and Trebbi (2003) find that OL systems reduce corruption, while Chang and Golden (2007) find the opposite effect. Negri (2014) develops a theoretical model and predicts that, in general, CL proportional representation is associated to lower minority representation within Parliaments than OL systems.

SPANISH ELECTORAL SYSTEMS

Spain is politically decentralized in 17 regions and >8000 municipalities. Each municipality elects a local government in free elections. A national law requires local governments to provide a variety of services, including public lighting, waste collection, cemeteries, street cleaning and road pavement (see Table 1 for a comprehensive list). In addition, municipalities are allowed to provide any other service that they consider useful to the municipality. For example, it is common that they provide touristic information to visitors and organize local festivities. Municipalities can levy a number of taxes and charges (most importantly, a property tax) and receive transfers from regional and national governments to finance some of their expenditures. Approximately 55 percent of their revenues come from the taxes imposed by themselves (Sweeting 2009).

Local elections are held simultaneously in all municipalities in Spain every fourth year.⁵ The elections follow one of three electoral systems, depending upon the population size of the municipality.

Municipalities with >250 inhabitants use a CL proportional representation system.⁶ They elect a city council in a single district, CL election, where each party presents a list of candidates and citizens can vote for one of the party-lists. The size of the council increases with population at certain population cutoffs but all municipalities in the CL system used for identification elect a seven-member council (as the empirical strategy relies on municipalities close to the threshold). The conversion from

⁵ Local elections coincide with regional elections in 13 of the 17 Spanish regions and in one year (1999) they also coincided with elections for the European Parliament. The implications of this are studied in the empirical strategy.

⁶ The population size that determines the electoral system is the official population in the municipal register (*padrón municipal*) the 1st of January of the year before the elections.

votes to seats is done according to the D'Hondt rule.⁷ The council elects a mayor among its members and is entitled to approve the budget, decide on expenditure in various fields, control the governing bodies and to the roll-call vote of confidence on the mayor. The mayor chairs the meetings of the council, casts the decisive vote in the event of a tie, decides on some expenditures, heads the local police and appoints mayoral deputies and cabinet members, among other responsibilities.

Municipalities with a population between 100 and 250 inhabitants follow an OL, plurality-at-large system. Under this system, a council of five members is elected. Political parties can present candidate lists of up to five candidates and voters can vote for up to four candidates from the same or different party-lists. The five most voted candidates are elected members of the council.⁸ As in the CL system, the council elects a mayor among its members and the responsibilities of the council and mayor are identical under the two systems.

Finally, municipalities with <100 inhabitants directly elect a mayor in a simple plurality, first-past-the-post (FPTP) election, that is, each political party can present one candidate and the most voted candidate is elected mayor. These municipalities follow a direct democracy system in which the role of the council is played by open meetings that any citizen in the municipality can attend.⁹

The paper focuses on the 250-inhabitant threshold and compares voter turnout under the CL and OL systems, as only the electoral system changes at that threshold. The results for the 100-inhabitant threshold are discussed in the Results section.

The CL and OL systems differ in two main dimensions. First, the electoral formula differs as under the OL system seats are allocated by plurality to the most voted individual candidates, while under CL they are allocated to parties according to the D'Hondt rule. Both the change in the electoral formula and the increase in the council size (from five to seven members) imply that the CL is a more proportional system *at the party level*. Second, the two systems differ on the ballot structure: while the CL system is a CL system where competition is limited to across-parties competition, the OL system allows voters to express their preferences for individual candidates, making it easier for popular candidates from small parties to get a seat in the council and introducing competition both across and within parties. In the Results section, I provide evidence that the difference in proportionality is small in practice, and that it is the difference in ballot structure what drives the results, with OLs increasing voter turnout.

EMPIRICAL STRATEGY AND DATA

Empirical Strategy

To estimate the effect of the electoral system on voter turnout, I combine regression discontinuity and fixed-effects estimation. In particular, I consider the following estimating equation:

$$y_{mt} = \alpha_m + \gamma_t + \beta D_{mt} + f(x_{mt} - x^*) + u_{mt}, \quad (1)$$

where y_{mt} is the outcome of interest (in the main specifications, voter turnout), D_{mt} a treatment dummy that captures the electoral system municipality m is required to follow in election-year t , x_{mt}

⁷ There is an electoral threshold at 5 percent, that is, parties need to get at least 5 percent of the votes to enter the D'Hondt distribution of seats. However, given the size of the council of the municipalities, we are studying that threshold does not play an important role in these elections.

⁸ The fact that candidates can vote for fewer candidates than there are seats to elect means that it is a *limited voting* system.

⁹ Municipalities with >100 inhabitants could decide to adopt this system. In order to do that, a majority of the citizens of the municipality had to sign a petition and two-thirds of the members of the council and the regional government had to approve it. To the best of my knowledge, no municipality ever used this procedure. Therefore, the regression discontinuity design can be sharp, as the probability of treatment jumps from 0 to 1 at the threshold.

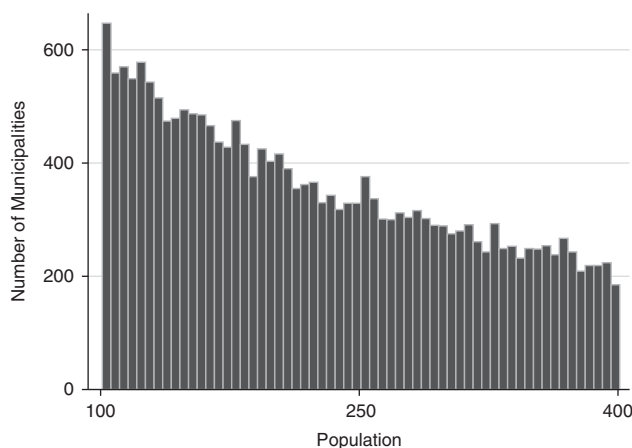


Fig. 1. Histogram of population size

the assignment variable (log population the year before the elections),¹⁰ x^* the log of the population threshold (250 inhabitants), so that treatment status depends on whether x_{mt} is bigger or smaller than x^* , f a smooth function of the assignment variable, α_m a municipality fixed effect, γ_t a year fixed effect and u_{mt} an error term. The parameter of interest is β .¹¹

To estimate f , I use non-parametric estimation (local linear regression). A key ingredient to this approach is the bandwidth. There is a trade-off between precision and bias: a bigger bandwidth increases precision at the cost of more bias. I choose a baseline bandwidth according to the procedure suggested by Imbens and Kalyanaraman (2012) and provide the results at different fractions of that bandwidth to see how sensitive the results are to bandwidth choice. Notice that we may not want the bandwidth to be bigger than 150 inhabitants (remember that municipalities with <100 inhabitants follow a different system), as if that happened we would be mixing outcomes from the three electoral systems in the same specification. For that reason, I restrict the analysis to bandwidths of <150 inhabitants. I use a rectangular kernel, as recommended by Imbens and Lemieux (2008) and Lee and Lemieux (2010). This is equivalent to estimating standard linear regressions over the interval of the selected bandwidth on both sides of the cut-off point. I cluster the standard errors at the municipality level.

A possible cause of concern in a pure regression discontinuity framework would be that a discontinuity in the density of population size is observed at the threshold (Lee and Lemieux 2010) (see Figure 1).¹² To deal with that issue, I add municipality and year fixed effects to the basic regression discontinuity framework. The identification assumption is that there are no

¹⁰ Geys (2006) recommends using log population for turnout studies.

¹¹ Similar strategies to the one used in this paper have been used by Grembi, Nannicini and Troiano (2012), Lemieux and Milligan (2008) and Pettersson-Lidbom (2012). More generally, population thresholds have been widely used in recent years as a way to get a credible estimation of causal effects (Ferraz and Finan 2009; Egger and Koethenbuerger 2010; Fujiwara 2011; Campa 2012; Casas-Arce and Saiz 2012; Pettersson-Lidbom (2012); Arnold and Freier 2015; Bordignon, Nannicini and Tabellini 2013; Brollo *et al.* 2013; Eggers 2015; Gagliarducci and Nannicini 2013; Litschig and Morrison 2013; Hinnerich and Pettersson-Lidbom 2014; Fujiwara 2015).

¹² The discontinuity is significant using McCrary's (2008) test in a pooled cross-section of all the municipality-years, but not if each election-year is considered separately. In the Robustness section, I describe in detail why and how that discontinuity appears.

unobservable factors that may simultaneously affect voter turnout and whether a municipality's population is just above or just below the threshold, *conditional on municipality and year fixed effects*. The identification therefore relies on switchers: intuitively, the regressions do not compare municipalities just above and just below the threshold but municipalities that switch from one system to another with those that remain in the same system.¹³ If the factors that make municipalities sort around the threshold are time invariant, β will identify the average treatment effect of the electoral system on voter turnout for municipalities close to the threshold.¹⁴

Although the identification assumption is fundamentally untestable, in the Robustness section I present three sets of tests to address the validity of the strategy. First, I test whether municipalities at each side of the threshold differ, conditional on the fixed effects, in other variables that may themselves affect the outcomes of interest. Second, I estimate a dynamic model to test for pretrends: in particular, I study whether this period's electoral system has an effect on previous period's turnout. Third, I consider donut regressions to examine the robustness of the results to the exclusion of some observations where any problem of self-selection that might remain after the inclusion of the fixed effects is likely to be concentrated.

Another issue that requires some consideration is that local elections are held on the same day as regional elections in 13 of the 17 Spanish regions. In 1999, local elections also coincided with elections to the European Parliament. This can be analyzed as a measurement error issue. Observed turnout t_{mt} , is measured with error as it differs from the "true" turnout t_{mt}^* , which is the turnout that would have been observed if local elections had been the only elections on election day. Let the error be $e_{mt} = t_{mt} - t_{mt}^*$. If (1) captures the true relationship between the outcome and independent variables, then $y_{mt} = t_{mt}^*$.

Thus, the estimated model is

$$t_{mt} = \alpha_m + \gamma_t + \beta D_{mt} + f(x_{mt} - x^*) + u_{mt} + e_{mt}. \quad (2)$$

If the new error term $u_{mt} + e_{mt}$ is uncorrelated with the regressors, the estimators will be unbiased and consistent. Thus, if e_{mt} is not correlated with the electoral system (and there is no reason to think it is, because the population threshold does not play any role in regional or European elections), measurement error will not affect the unbiasedness and consistency of the estimators.¹⁵

Data

I have collected data from all local elections that have taken place in Spain since the restoration of democracy after General Franco's death in 1975. In this period, local elections have been held in nine years (1979, 1983, 1987, 1991, 1995, 1999, 2003, 2007, 2011) in all municipalities in the country (around 8000).¹⁶ In addition to the data from local elections, I have collected data

¹³ The inclusion of fixed effects implies that municipalities just above and just below the threshold *can* differ in other factors that affect voter turnout.

¹⁴ The use of fixed effects implies that it is not possible to provide a graphical representation of the results, as the comparison of municipalities above and below the threshold is conditioned on the municipality and year fixed effects. (See Pettersson-Lidbom (2012) for a paper that combines regression discontinuity and fixed effects and cannot provide a graphical analysis.)

¹⁵ The estimators will nonetheless be inefficient, but the big sample size helps to overcome that problem.

¹⁶ Due to some missing values or errors in the official elections data, the final data set used in this paper contains observations for 71,780 out of the estimated 73,026 local elections in this period, that is, 98.6 percent of the elections. I do not know the number of total elections because for the first few years I observed and included in the sample only municipalities with election data, and therefore I do not know the number of municipalities

TABLE 2 *Summary Statistics*

	Mean	SD	Minimum	p10	p50	p90	Maximum	Count
All municipalities								
Turnout (%)	75.4	12.1	0.9	59.6	77.4	88.7	100	71,780
Lists (%)	2.9	1.6	1.0	1	3	5	25	71,780
<i>N</i> turnout (%)	77.2	9.0	0.8	66.0	78.4	87.0	100	63,478
<i>N</i> blank (%)	0.9	1.2	0.0	0	0.6	2.0	50	63,478
<i>N</i> spoilt (%)	1.1	1.9	0.0	0	0.6	2.5	84.6	63,478
<i>N</i> right (%)	39.5	19.5	0.0	12.4	40.1	64.5	100	63,478
<i>N</i> left (%)	37.3	15.8	0.0	15.9	37.8	57.5	100	63,478
<i>N</i> far left (%)	3.7	4.5	0.0	0	2.4	8.5	66.7	56,233
OL municipalities (100 ≤ population ≤ 250)								
Turnout (%)	77.4	12	0.9	61.5	79.7	90.1	100	13,496
Lists (%)	2.2	0.9	1.0	1	2	3	6	13,496
<i>N</i> turnout (%)	77.9	8.7	2.1	67.1	79.0	87.6	100	11,941
<i>N</i> blank (%)	1.0	1.5	0.0	0	0.7	2.7	50	11,941
<i>N</i> spoilt (%)	1.2	2.2	0.0	0	0	3.2	45.0	11,941
<i>N</i> right (%)	45.8	20.5	0.0	14.6	48.1	71.3	97.5	11,941
<i>N</i> left (%)	31.2	14.8	0.0	11.9	30.8	50.5	87.9	11,941
<i>N</i> far left (%)	2.7	3.2	0.0	0	1.7	6.5	43.8	10,515
CL municipalities (251 ≤ population ≤ 400)								
Turnout (%)	76.6	14.0	6.5	56.2	80.4	90.3	100	8059
Lists (%)	2.1	0.8	1.0	1	2	3	6	8059
<i>N</i> turnout (%)	77.5	8.9	13.3	66.1	78.9	87.2	99.3	7009
<i>N</i> blank (%)	0.9	1.0	0.0	0	0.6	2.1	11.0	7009
<i>N</i> spoilt (%)	1.1	1.9	0.0	0	0.6	2.7	52.4	7009
<i>N</i> right (%)	40.7	19.5	0.0	11.3	43.1	64.6	91.4	7009
<i>N</i> left (%)	34.6	15.1	0.0	14.0	35.4	53.4	81.0	7009
<i>N</i> far left (%)	2.7	3.0	0.0	0	1.9	6.1	50	6237

Note: the first panel shows summary statistics for all municipalities. The second panel shows summary statistics for municipalities under the open list (OL) system. The third panel, for those under the closed list (CL) system and < 400 inhabitants, which is the biggest bandwidth in the regressions.

at the municipality level from all national Congress elections in that period.¹⁷ These data will be used in the Robustness section to analyze the validity of the identification strategy. All data are from the National Statistics Institute (INE) and are publicly available.

Table 2 presents summary statistics for the variables used in the paper. The main outcome of interest is *Turnout*, defined as the number of votes cast divided by the electoral census. That is, *Turnout* measures the proportion of citizens that cast a vote over the set of potential voters (the electoral census). There is no voter registration in Spain: potential voters are all citizens from Spain, EU countries and countries under Reciprocity Treaties, older than 18 and not disenfranchised by court order.¹⁸ The other outcome variable of interest is *Lists*, defined as the number of party-lists that run for election.

I consider six variables from Congress elections: the percent of voter turnout (*N turnout*), defined analogously to the one for local elections, the percent of blank (*N blank*) and spoilt

(*F*note continued)

that held elections that are not included in my sample. I estimated 73,026 by assuming that the number of municipalities remained constant over time.

¹⁷ Spain is a bicameral parliamentary system. The national elections data are for elections to the Lower House (*Congreso*), which concentrates most of the political power. Throughout the paper, “national elections” refer to these elections.

¹⁸ Disenfranchisement is mostly for disability reasons. In 2011, the figure of disenfranchised individuals was 79,398 (including individuals younger than 18), or around 0.18 percent of the population.

TABLE 3 *Effect of Ballot Structure on Voter Turnout: Open Versus Closed Lists*

	(1)	(2)	(3)	(4)	(5)	(6)
	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout
OL	1.031*** (0.381)	1.237*** (0.395)	1.156*** (0.407)	1.547*** (0.431)	1.844*** (0.491)	1.308** (0.641)
Observations	15,954	13,556	11,226	8798	6404	3976
Municipalities	2826	2597	2343	2079	1803	1512
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman's (2012) bandwidth and 250 inhabitants. OL = open list system.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

(*N* *spoilt*) votes, defined as the number of blank and spoilt votes divided by the number of votes cast, and the share of votes for the three main parties in Spain: the right-wing Popular Party (*N* *right*), the left-wing Socialist Party (*N* *left*) and the far-left United Left (*N* *far left*),¹⁹ defined as the votes for these parties divided by the number of valid votes.²⁰

RESULTS

Main Results

Table 3 presents the main estimates of the impact of the electoral system on voter turnout. The table shows the results of estimating Equation (1) by local linear regression, using *Turnout* as the dependent variable. The treatment variable is *OL*, which takes the value of 1 if the municipality follows the OL system and 0 if it follows the CL system. The coefficient on *OL* therefore captures the effect on turnout of using the OL system in relation to the CL system, that is, it measures the effect on turnout of crossing the threshold from right to left.

Column (1) provides the estimated effect based on the preferred specification, which uses the bandwidth suggested by Imbens and Lemieux (2008). Under that specification, the OL system increases turnout in relation to the CL system by 1 percentage point, and the effect is statistically significant at the 1 percent level. The rest of the columns show the results for different fractions of that bandwidth. The effect is statistically significant and quantitatively similar across bandwidths, ranging from 1 to 2 percentage points, suggesting that the effect is not driven by the choice of the bandwidth. In conclusion, the results imply that the OL system leads to more voter turnout than the CL system.

Additional Results: Mechanism

The OL and CL systems differ in two main dimensions. First, the electoral formula differs as under the OL system seats are allocated by plurality to the most voted individual candidates,

¹⁹ United Left (*Izquierda Unida*) is a coalition of parties created in 1986 whose main party is the Communist Party. For elections before that date, votes for the Communist Party are considered. The sample size is smaller than for the main parties because, in some years, the coalition did not run in some regions.

²⁰ Valid votes include votes for candidates and blank votes, but not spoilt votes. I use this denominator because it is the relevant one for the allocation of seats, as it is used to determine whether parties reach the election threshold to get seats (3 percent in Congress elections). Accordingly, it is the one that is normally reported by the media.

TABLE 4 *Effect of Ballot Structure on Number of Party-Lists and on Voter Turnout When Controlling for Party-Lists*

	(1)	(2)	(3)	(4)	(5)	(6)
	Lists	Lists	Lists	Lists	Lists	Lists
Panel A: party-lists						
OL	0.314*** (0.0292)	0.317*** (0.0304)	0.331*** (0.0321)	0.355*** (0.0341)	0.387*** (0.0402)	0.382*** (0.0511)
Observations	17,693	15,148	12,517	9845	7114	4472
Municipalities	2993	2744	2483	2196	1885	1577
	(1)	(2)	(3)	(4)	(5)	(6)
	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout
Panel B: turnout when controlling for party-lists						
Lists	6.042*** (0.152)	5.979*** (0.169)	5.916*** (0.189)	6.063*** (0.216)	5.937*** (0.256)	5.956*** (0.319)
OL	-0.860** (0.343)	-0.708** (0.359)	-0.857** (0.373)	-0.655 (0.403)	-0.578 (0.453)	-0.944 (0.597)
Observations	15,954	13,556	11,226	8798	6404	3976
Municipalities	2826	2597	2343	2079	1803	1512
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman's (2012) bandwidth and 250 inhabitants. OL = open list system.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

while under the CL system they are allocated to parties according to the D'Hondt rule. Both the change in the electoral formula and the increase in the council size (from five to seven members) imply that the CL is a more proportional system *at the party level*. Second, the two systems differ on the ballot structure: while the CL system is a CL system where competition is limited to competition across political parties, the OL system allows voters to express their preferences for individual candidates, introducing competition both across and within parties. In this section, I provide evidence that the ballot structure difference is more relevant and the one that drives the results.

First, although the CL system is formally a proportional representation system, in practice it is only slightly more proportional than the OL system. Under the OL system, only in 3.5 percent of the elections the most voted party does not get an absolute majority (3/5 or more) of the seats (and therefore most of the political power, including the possibility of appointing the mayor).²¹ But the cases in which no party has an absolute majority of seats in the council are also rare under the CL system, around 7 percent of the elections. That is because the number of seats is small (seven) and the D'Hondt rule is used, which favors the most voted parties when the district size is small.

²¹ To see why that is the case, consider a scenario in which there are three parties, A, B and C, and that 40 percent of the electorate prefers party A, 35 percent party B and 25 percent party C. Assume that voters are indifferent between the two parties other than their preferred party and, to focus on across-parties competition, that they do not care about individual candidates but only about their preferred party getting the mayor. In this case, party A can present a list of four candidates and its supporters give their four votes to those four candidates. This guarantees four seats to party A, no matter how many candidates parties B and C decide to include on their lists. Party B will get the other seat unless it decides to include five candidates on its list. Thus, even with a minor advantage in terms of votes, the leading party gets four out of five seats and therefore the mayor. If the advantage is sufficiently big, it may get the five seats by presenting five candidates if voters split their votes sufficiently across candidates.

Second, I study whether the electoral system also has an effect on the number of party-lists running in the election. Panel A of Table 4 shows that the OL system increases the number of lists by around 0.35, which suggests that at least part of the change in voter turnout might be driven by entry decisions of candidates. A higher number of parties in competition may in turn affect voter turnout by increased aggregate mobilization efforts and by providing voters with a more compelling set of options.²²

Panel B of Table 4 shows the results for voter turnout after a control for the number of lists is included in the main regression. The coefficient on *Lists* indicates that one more party running in the election increases turnout by around 6 percentage points. But the more interesting result is that the estimated effect of the electoral system reverses when the control for the number of lists is introduced: conditional on the number of lists, turnout is higher under the CL system. Of course, this does not mean that the causal effect is actually positive. The number of lists is an endogenous or “bad” control in a regression whose goal is to find the true causal effect, as it is also an outcome of the electoral system (see Angrist and Pischke 2008 on bad controls). However, this exercise reinforces the argument that the reason why the OL system leads to more turnout than the CL system is that it encourages more parties to enter competition.

It is hard to explain the difference in number of parties as being driven by the different proportionality of the systems: if voters care only about party-level competition, the CL system makes it easier for small parties to get some representation so we should expect more parties to run under the CL system.²³ The change in the number of parties is most likely driven by the different ballot structure of the systems, as the OL system makes it easier for individual candidates in small parties to get elected to the council. For example, assume there are three main parties A, B and C, and a small party D that has a very popular leader, and that voters care both about parties and individual candidates. Suppose that if voters can only vote for one party-list, as in the CL system, most individuals prefer to vote for parties A, B and C, so these three parties share the seven seats. That is, although D has a very popular leader, most voters prefer to vote for their preferred parties A, B and C if they can only vote for one party-list. In this scenario, party D will not enter competition if there is a cost of running. On the other hand, under the OL system, voters are allowed to vote for four individual candidates, from the same or different party-lists. Thus, voters could give three of their votes for candidates of their preferred party and one vote for the leader of party D. Under this system, party D will enter competition if the benefit of getting a seat is bigger than the cost of running. This argument suggests that the ballot structure of the OL system, which allows voters to choose among candidates from different party-lists, should lead to more parties in competition.

Third, I compare the OL system with a third electoral system, exploiting a discontinuity in electoral systems at a different threshold: municipalities with 100 or more inhabitants (but with not >250) follow the OL system, while those with <100 inhabitants elect a mayor in a FPTP election, that is, each political party can present one candidate and the most voted candidate is elected mayor. These municipalities follow a direct democracy system in which the role of the council is played by open meetings that any citizen in the municipality can attend.²⁴

²² It has also been argued that a higher number of parties could decrease voter turnout as it usually produces coalition governments, which make elections less decisive (Jackman 1987). However, that is unlikely to play a big role in the case studied in this paper because coalition governments are rare in the two systems that I compare, as is explained above.

²³ Indeed, it has been generally accepted that proportional representation systems increase the number of parties in competition (Blais and Aarts 2006).

²⁴ Sanz (2015) uses the change in the government system at the 100-inhabitant threshold to study how direct democracy, as opposed to representative democracy, affects public finances.

TABLE 5 *Effect of Ballot Structure on Voter Turnout and Party-Lists: Open List Versus FPTP (100-inhabitant threshold)*

	(1)	(2)	(3)	(4)	(5)	(6)
	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout
Panel A: turnout						
OL	2.811*** (0.521)	2.808*** (0.531)	2.566*** (0.548)	2.447*** (0.580)	2.474*** (0.656)	2.243*** (0.807)
Observations	11,161	9662	8020	6297	4604	2901
Municipalities	2621	2349	2054	1714	1384	1057
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK
	(1)	(2)	(3)	(4)	(5)	(6)
	Lists	Lists	Lists	Lists	Lists	Lists
Panel B: party-lists						
OL	0.0761* (0.0390)	0.0764* (0.0408)	0.0808* (0.0433)	0.0462 (0.0491)	0.0764 (0.0533)	0.143* (0.0788)
Observations	6957	5952	4949	3924	2901	1874
Municipalities	1833	1649	1462	1253	1057	835
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman's (2012) bandwidth and 250 inhabitants. OL = open list system.
* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

Although the estimates for this threshold should be interpreted with caution, as in the 100-inhabitant threshold there is a change in the government system in addition to the change in the electoral system, the evidence presented here reinforces some of the conclusions drawn from the main results from the 250-inhabitant threshold. In particular, if it is the ballot structure and not the proportionality of the systems what drives the increased turnout under the OL with respect to the CL system, we would also expect that effect to be present when we compare the OL system with FPTP, as these two systems also differ on the ballot structure but are both plurality (and not proportional representation) systems.

To study turnout at this threshold, I follow the same empirical strategy described in the previous section. The treatment variable is (slightly abusing notation) again *OL*, which takes the value of 1 if the municipality follows the OL system and 0 if it follows the FPTP system. The results in Panel A of Table 5 show that the OL system increases voter turnout by between 2 and 3 percentage points with respect to FPTP.²⁵ It also increases the number of lists in competition, although by a modest amount (0.1) (Panel B). Therefore, the results from this threshold go in the same direction as those obtained from the main

²⁵ For this threshold, I exclude the observations from 1979, 1983 and 2011 because the threshold was not relevant to the electoral system in those years. In 1979 and 1983, the threshold for electing a mayor by FPTP was at 25 inhabitants rather than at 100, so that system was virtually not used (there was only one municipality with a population of <25 inhabitants in those two years). A change in the electoral law in 2011 determined that all municipalities had to elect a council in the 2011 election so no FTPT elections were held that year. I also exclude the observations from 1987. These were the first elections in which all the municipalities with <100 inhabitants were required to follow the FPTP system and in the majority of municipalities only one list ran for the election, probably as a consequence of parties not having time to adapt to the new system. To prevent these transitional effects from affecting the results I do not include observations from 1987 in the estimation. In any case, the results are still significant (and in general even bigger in magnitude) if that year is included in the estimation.

threshold: not only does the OL system lead to more turnout and more lists in competition than the CL system, but the OL system also produces the same effects when we compare it with FPTP.²⁶

ROBUSTNESS

As explained in the Empirical Strategy and Data section, a discontinuity in the density of population sizes is observed at the threshold. In the first part of this section, I review how that discontinuity arises. In the subsequent parts of the section I present three sets of robustness checks to assess the validity of the empirical strategy.

Sorting Around the Threshold: Process

The official population size of a municipality is given by the number of citizens who are registered in the municipal register (*padrón municipal*). Municipalities keep track of all the variations in the population (births, deaths, registrations and unregistrations) in the public register and report periodically the data to the INE. The INE validates the information it receives, checking that there is no fraud—for example, it makes sure that for every sign-in in a municipality there is a corresponding sign-out in another—and yearly makes the final population figures public. Given the mechanics of the process, it is hard to believe that there is outright manipulation of the population figures. Instead, the discontinuity at the threshold arises because individuals that have dwellings in more than one municipality can in practice decide in which of them to register.²⁷ In particular, the mayor or other local politicians may be interested in the electoral system that the municipality follows and may make an effort to persuade some individuals to register in the municipality so that it keeps the preferred electoral system.²⁸ Therefore, some individuals that would have registered in another municipality if there was no change of rules at the threshold, decide to register in the municipality.

Although the identification assumption is fundamentally untestable, I now present three sets of tests to assess the validity of the strategy. First, I test whether municipalities at each side of the threshold differ, conditional on the fixed effects, in other variables that may themselves affect the outcomes of interest. Second, I estimate a dynamic model to test for pretrends: in particular, I study whether this period's electoral system has an effect on previous period's turnout. Third, I consider donut regressions to examine the robustness of the results to the exclusion of some observations where any problem of self-selection that might remain after the inclusion of the fixed effects is likely to be concentrated.

Placebo Tests: Covariates at the Thresholds

The identification strategy requires municipalities that switch into the OL system be comparable with those that stay in the same system and with those that switch out of

²⁶ The robustness checks for the results from this threshold are presented in the Appendix.

²⁷ Although individuals who live in more than one municipality are required to register in the municipality in which they spend more time, that requirement is almost impossible to monitor and not enforced in practice.

²⁸ In particular, the spike of the distribution of population sizes suggests that there is a preference to fall under the CL system. A possible reason for this preference is that the number of seats increases at the threshold, therefore leading to more political positions available. Indeed, similar spikes on the distribution are observed at other thresholds where the council size increases, although it is not possible to know whether the spike is entirely caused by the council size as more regulations change at the other thresholds.

the OL system. In other words, municipalities just above and just below the threshold should not differ, conditional on the fixed effects, in other variables that may themselves have an effect on voter turnout. Here, I use national Congress election variables and test whether they change at the thresholds by including them as dependent variables in Equation (1).²⁹ These covariates are especially suited to this context because if municipalities differ in some unobservable factors that affect voter turnout at local elections, it is likely that those unobservable variable affect turnout at national elections too. Furthermore, Congress elections are the most important elections in Spain and turnout is typically very high (78 percent in the average municipality during the sample period) so they are likely to capture any political differences across municipalities.³⁰

As pointed out by Lee and Lemieux (2010), it is important that the covariates are determined before the present period's realization of the running and treatment variables if they can be affected by treatment.³¹ I merge local elections data with data from the most recent previous Congress elections at the municipality level.³² I consider the percent of voter turnout in (N turnout), blank (N blank) and spoilt votes (N spoilt) in Congress elections and the percent of votes for the three main parties in Spain (N right, N left and N far left).³³ If the empirical strategy is valid, there should be no effect of the electoral system on these variables.

The results from these tests are shown in Table 6. Among the six variables considered, only in one (the share of spoilt votes) there seems to be some difference between municipalities, but the effect is quantitatively small and only significant in some specifications. Overall, the results suggest that, conditional on the fixed effects, municipalities just above and just below the threshold do not differ systematically along these variables: while municipalities operating under open and CLs behave differently in local elections, they do not do so in national elections.

Pretrends in the Outcomes of Interest

The longitudinal structure of the data can be used to estimate dynamic causal effects. This subsection studies the timing of the effect of the electoral system on the outcomes of interest by estimating the following model:

$$y_{mt} = \alpha_m + \gamma_t + \beta_1 D_{mt} + \beta_2 D_{m,t+1} + f_1(x_{mt} - x^*) + f_2(x_{m,t+1} - x^*) + u_{mt}. \quad (3)$$

If, conditional on the fixed effects, next period's electoral system is as good as random for those municipalities sufficiently close to the threshold, next period's electoral system should not help predict this period's outcome variable.³⁴ Thus, the coefficient β_2 serves as a robustness check: $\beta_2 \neq 0$ would suggest an endogeneity problem that may raise concerns about the validity of the

²⁹ Note that the inclusion of the fixed effects implies that to test the robustness of the strategy we need variables that change over time.

³⁰ Spain follows a parliamentary system, so there are no elections for the executive branch. Citizens elect the Congress, which in turn elects the Prime Minister.

³¹ Although the population threshold does not play any role in national elections, behavior at subsequent national elections could be affected if, for example, the electoral system influences voters' political views.

³² Thus, I estimate Equation (1) where instead of turnout at the 1983 (1987, 1991, 1995, 1999, 2003, 2007, 2011) local elections I include turnout at the 1982 (1986, 1989, 1993, 1996, 2000, 2004, 2008) national elections as the dependent variable.

³³ The other parties represented in Congress during the whole period are regional parties that do not run in the whole country.

³⁴ According to Lee and Lemieux (2010), "finding a discontinuity in y_{mt} but not in $y_{m,t-1}$ would be a strong piece of evidence supporting the validity of the RD design."

TABLE 6 *Placebo Tests: Covariates at the Threshold*

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	<i>N</i> Turnout	<i>N</i> Turnout	<i>N</i> Turnout	<i>N</i> Turnout	<i>N</i> Turnout	<i>N</i> Turnout	<i>N</i> Blank	<i>N</i> Blank	<i>N</i> Blank	<i>N</i> Blank	<i>N</i> Blank	<i>N</i> Blank
Panel A: turnout and blank votes												
OL	−0.0562 (0.192)	−0.185 (0.193)	−0.212 (0.199)	−0.172 (0.210)	−0.0113 (0.222)	0.375 (0.250)	0.00401 (0.0337)	0.00638 (0.0347)	0.0306 (0.0365)	0.0259 (0.0394)	0.0651 (0.0451)	0.0531 (0.0522)
Observations	23,482	20,171	16,816	13,393	9787	6044	22,370	19,198	15,983	12,680	9181	5706
Municipalities	3731	3377	2996	2592	2149	1656	3626	3274	2884	2507	2067	1617
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	<i>N</i> Spoilt	<i>N</i> Spoilt	<i>N</i> Spoilt	<i>N</i> Spoilt	<i>N</i> Spoilt	<i>N</i> Spoilt	<i>N</i> Right	<i>N</i> Right	<i>N</i> Right	<i>N</i> Right	<i>N</i> Right	<i>N</i> Right
Panel B: spoilt votes and votes for right-wing party												
OL	0.145** (0.0573)	0.153*** (0.0578)	0.167*** (0.0603)	0.136*** (0.0637)	0.110 (0.0703)	−0.101 (0.0909)	−0.0839 (0.376)	−0.0708 (0.378)	−0.162 (0.382)	−0.285 (0.395)	−0.223 (0.409)	−0.498 (0.480)
Observations	25,484	21,878	18,178	14,475	10,609	6600	19,279	16,604	13,782	10,922	7948	4892
Municipalities	3946	3580	3161	2721	2260	1734	3283	2978	2637	2300	1912	1514
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)
	<i>N</i> Left	<i>N</i> Left	<i>N</i> Left	<i>N</i> Left	<i>N</i> Left	<i>N</i> Left	<i>N</i> Far Left	<i>N</i> Far Left	<i>N</i> Far Left	<i>N</i> Far Left	<i>N</i> Far Left	<i>N</i> Far Left
Panel C: votes for left- and far-left-wing parties												
OL	0.395 (0.272)	0.366 (0.279)	0.300 (0.281)	0.482 (0.296)	0.406 (0.326)	0.482 (0.415)	0.0412 (0.0841)	0.0101 (0.0874)	0.0672 (0.0928)	0.0721 (0.102)	0.127 (0.111)	0.0131 (0.129)
Observations	16,458	14,086	11,664	9146	6645	4150	17,913	15,285	12,806	10091	7347	4547
Municipalities	2957	2679	2393	2065	1742	1402	3367	3018	2706	2344	1940	1525
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman’s (2012) bandwidth and 250 inhabitants.
OL = open list system.
p* < 0.10, *p* < 0.05, ****p* < 0.01.

TABLE 7 *Robustness: Testing for Pretrends in the Outcomes of Interest*

	(1)	(2)	(3)	(4)	(5)	(6)
	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout
Panel A: turnout						
OL	1.628*** (0.410)	1.792*** (0.425)	1.754*** (0.443)	2.014*** (0.471)	2.251*** (0.548)	1.620** (0.762)
OL(+1)	-0.259 (0.437)	-0.302 (0.455)	-0.313 (0.474)	-0.191 (0.495)	-0.150 (0.545)	0.563 (0.713)
Observations	14,136	12,005	9937	7792	5674	3516
Municipalities	2730	2507	2233	1974	1686	1405
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK
	(1)	(2)	(3)	(4)	(5)	(6)
	Lists	Lists	Lists	Lists	Lists	Lists
Panel B: lists						
OL	0.349*** (0.0308)	0.360*** (0.0319)	0.374*** (0.0340)	0.385*** (0.0366)	0.398*** (0.0432)	0.358*** (0.0576)
OL(+1)	-0.0458 (0.0281)	-0.0619** (0.0292)	-0.0776** (0.0308)	-0.0877*** (0.0337)	-0.104*** (0.0382)	-0.0736 (0.0501)
Observations	15,663	13,426	11,084	8723	6312	3962
Municipalities	2904	2649	2383	2087	1767	1466
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman's (2012) bandwidth and 250 inhabitants. OL = open list system.

*p < 0.10, **p < 0.05, ***p < 0.01.

approach of the paper. Alternatively, we can think of (3) as testing whether treatment and control municipalities were on different trends before the realization of this period's variables.³⁵

Table 7 shows the results. The estimates for β_2 are not statistically significant for *Turnout*. For some bandwidth choices, the coefficients for *Lists* are significant but they are small and have the opposite sign to the contemporaneous effect, suggesting that, if there is any difference, OL municipalities had *fewer* lists in the previous election. Moreover, β_1 , the contemporaneous effect of the electoral system on the outcomes, remains similar to the one obtained in the baseline specifications, for the three outcome variables. That implies that previous trends in the outcome variables are not driving the results.

Donut Regressions

This subsection considers donut regressions in the spirit of Barreca et al. (2011). The idea of this technique is to exclude observations very close to the thresholds, where sorting is more likely to occur. Given the nature of the sorting process of Spanish municipalities, it is likely that the sorting of municipalities is limited to a small window around the threshold: as the discontinuity in the density appears not because of outright manipulation of the population data, but as a consequence of registration decisions of citizens, it does not seem plausible that the population size of municipalities that are self-selecting into treatment is much bigger than what is strictly necessary to be above the threshold, or that municipalities with a population size far below the

³⁵ It is necessary to flexibly control for population in both periods (i.e., to include the functions f_1 and f_2) to capture the effect of the electoral system and not of changes in population size.

TABLE 8 *Robustness: Donut Regressions*

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout
Panel A: turnout								
OL	1.031*** (0.381)	0.925** (0.419)	0.920** (0.440)	0.754 (0.494)	1.676*** (0.445)	1.546*** (0.517)	1.689*** (0.583)	1.876*** (0.710)
Observations	15,954	15,531	15,249	14,583	8039	7616	7334	6668
Municipalities	2826	2825	2825	2825	2002	1994	1991	1976
Bandwidth	IK	IK	IK	IK	0.50 × IK	0.50 × IK	0.50 × IK	0.50 × IK
Excluded	0	6	10	20	0	6	10	20
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Lists	Lists	Lists	Lists	Lists	Lists	Lists	Lists
Panel B: lists								
OL	0.314*** (0.0292)	0.307*** (0.0312)	0.288*** (0.0321)	0.282*** (0.0351)	0.361*** (0.0353)	0.357*** (0.0389)	0.333*** (0.0415)	0.353*** (0.0486)
Observations	17,693	17,270	16,988	16,322	9000	8577	8295	7629
Municipalities	2993	2992	2992	2992	2102	2096	2093	2084
Bandwidth	IK	IK	IK	IK	0.50 × IK	0.50 × IK	0.50 × IK	0.50 × IK
Excluded	0	6	10	20	0	6	10	20

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman's (2012) bandwidth and 250 inhabitants. OL = open list system.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

threshold attempt to cross it. Thus, finding a similar effect after municipalities close to the threshold are excluded would reinforce the credibility of the estimates.

For each outcome variable and bandwidth choice, I consider four specifications. The first one is the benchmark regression, in which no observations are excluded. The second excludes observations within a window of six inhabitants around the threshold (three at each side of the threshold), the third excludes those within an interval of ten inhabitants and the fourth those within an interval of 20 inhabitants.

The results (see Table 8) show that excluding observations very close to the threshold does not affect the results, as they are remarkably similar in magnitude to the ones obtained in the baseline specifications. For voter turnout, the results are significant when up to ten inhabitants around the threshold are excluded.³⁶ For a bandwidth of half that length, the results are significant even when municipalities with a population within a 20 inhabitants around the threshold are excluded. For the number of lists, the coefficients remain significant across all specifications. These results suggest that the findings of the paper are not being driven by municipalities of which concerns about self-selection could still persist.

CONCLUSION

By exploiting the unique institutional framework of Spanish local elections, I have shown that the electoral system has an effect on voter turnout. In particular, an OL, plurality-at-large system increases turnout by between 1 and 2 percentage points with respect to a CL proportional representation system. The results suggest that OL systems where parties run in candidate

³⁶ When 20 inhabitants are excluded, the results are not significant but the point estimate remains similar: it is the increased standard errors due the decrease in sample size that makes the coefficient lose significance.

lists but voters can express their preferences for individual candidates are conducive to more voter turnout.

Working on understanding this type of electoral systems better could be a fruitful avenue for future research. Economists and political scientists have extensively worked on voter turnout and the motivations for the decision of voting or abstaining (Ledyard 1984; Palfrey and Rosenthal 1985; Feddersen and Pesendorfer 1996; Feddersen and Sandroni 2006; Battaglini, Morton and Palfrey 2010), and this is still an active area of research (Gerber, Green and Larimer 2008; Nickerson 2008; DellaVigna et al. 2013; Herrera, Morelli and Palfrey 2013; Kartal 2013; Fujiwara, Meng and Vogl 2014). However, to the best of my knowledge, we still do not have a theoretical framework to think about voter turnout in OL elections, in which there is competition both across and within parties.

REFERENCES

- Ames, Barry. 1995a. Electoral Rules, Constituency Pressures, and Pork Barrel: Bases of Voting in the Brazilian Congress'. *The Journal of Politics* 57(4):324–43.
- Ames, Barry. 1995b. Electoral Strategies Under Open-List Proportional Representation'. *American Journal of Political Science* 39(2):406–33.
- Angrist, Joshua D., and Jörn-Steffen Pischke. 2008. *Mostly Harmless Econometrics: An Empiricist's Companion*. Princeton University Press. Available at: <https://books.google.com/books?hl=es&lr=&id=ztXL21Xd8v8C&oi=fnd&pg=PR8&ots=Ui63Xw6OxP&sig=mxvOJztgp5YJJ0lhYSMP0bwqeC0#v=onepage&q&f=false>
- Arnold, Felix, and Ronny Freier. 2015. 'Signature Requirements and Citizen Initiatives: Quasi-Experimental Evidence from Germany'. *Public Choice* 162:43–56.
- Barreca, Alan I., Melanie Guldi, Jason M. Lindo, and Glen R. Waddell. 2011. 'Saving Babies? Revisiting the Effect of Very Low Birth Weight Classification'. *The Quarterly Journal of Economics* 126 (4):2117–123.
- Battaglini, Marco, Rebecca B. Morton, and Thomas R. Palfrey. 2010. 'The Swing Voter's Curse in the Laboratory'. *The Review of Economic Studies* 77(1):61–89.
- Blais, André. 2006. 'What Affects Voter Turnout?'. *Annual Review of Political Science* 9:111–25.
- Blais, André, and Agnieszka Dobrzynska. 1998. 'Turnout in Electoral Democracies'. *European Journal of Political Research* 33(2):239–61.
- Blais, André, and Kees Aarts. 2006. 'Electoral Systems and Turnout'. *Acta Politica* 41(2):180–96.
- Bordignon, Massimo, Tommaso Nannicini, and Guido Tabellini. 2013. 'Moderating Political Extremism: Single Round vs Runoff Elections Under Plurality Rule'. Discussion Paper, IZA. Available at: http://didattica.unibocconi.it/mypage/upload/92884_20130905_083752_DP7561.PDF
- Bowler, Shawn, David Brockington, and Todd Donovan. 2003. *Electoral Reform and Minority Representation: Local Experiments with Alternative Elections*. Ohio State University Press. Available at: <https://books.google.com/books?id=2LU520iehmgC&printsec=frontcover&dq=bowler+donovan&hl=es&sa=X&ved=0CB0Q6AEwAGoVChMI9Xik6v8xwIVgRo-Ch2F-AYP#v=onepage&q=bowler%20donovan&f=false>
- Brollo, Fernanda, Tommaso Nannicini, Roberto Perotti, and Guido Tabellini. 2013. 'The Political Resource Curse'. *The American Economic Review* 103(5):1759–796.
- Campa, Pamela. 2012. 'Gender Quotas, Female Politicians and Public Expenditures: Quasi-Experimental Evidence'. Working Paper No. 157, EONPUBBLICA. Available at: http://www.eonpubblica.unibocconi.it/files/WP_157_2011.pdf
- Carey, John M., and Matthew Soberg Shugart. 1995. 'Incentives to Cultivate a Personal Vote: A Rank Ordering of Electoral Formulas'. *Electoral Studies* 14(4):417–39.
- Casas-Arce, Pablo, and Albert Saiz. 2012. 'Women and Power: Unwilling, Ineffective, or Held Back?', Discussion Paper, IZA. Available at: http://papers.ssrn.com/sol3/papers.cfm?abstract_id=1812530

- Chang, Eric C.C. 2005. 'Electoral Incentives for Political Corruption Under Open-List Proportional Representation'. *The Journal of Politics* 67(3):716–30.
- Chang, Eric C.C., and Miriam A. Golden. 2007. 'Electoral Systems, District Magnitude and Corruption'. *British Journal of Political Science* 37(1):115–37.
- DellaVigna, Stefano, John A. List, Ulrike Malmendier, and Gautam Rao. 2013. 'Voting to Tell Others'. Mimeo, U.C. Berkeley. Available at <http://eml.berkeley.edu/~sdellavi/wp/Turnout15-01-09.pdf>
- Dos Santos, André Marengo. 2007. 'Do Electoral Rules Matter? Electoral List Models and Their Effects on Party Competition and Institutional Performance'. Available at http://socialsciences.scielo.org/pdf/s_dados/v3nse/scs_a06.pdf, accessed 16 September 2015.
- Downs, Anthony. 1957. *An Economic Theory of Democracy*. New York, NY: Harper.
- Eggers, Andrew C. 2015. 'Proportionality and Turnout: Evidence from French Municipalities'. *Comparative Political Studies* 48(2):135–67
- Egger, Peter, and Marko Koethenbueger. 2010. 'Government Spending and Legislative Organization: Quasi-Experimental Evidence From Germany'. *American Economic Journal: Applied Economics*, 200–12.
- Farrell, David M., and Ian McAllister. 2006. 'Voter Satisfaction and Electoral Systems: Does Preferential Voting in Candidate-Centred Systems Make a Difference?'. *European Journal of Political Research* 45(5):723–49.
- Feddersen, Timothy, and Alvaro Sandroni. 2006. 'A Theory of Participation in Elections'. *The American Economic Review* 96(4):1271–282.
- Feddersen, Timothy J., and Wolfgang Pesendorfer. 1996. 'The Swing Voter's Curse'. *The American Economic Review* 86(3):408–24.
- Ferraz, Claudio, and Frederico Finan. 2009. 'Motivating Politicians: The Impacts of Monetary Incentives on Quality and Performance'. Available at http://eml.berkeley.edu/~ffinan/Finan_MPoliticians.pdf, accessed 16 September 2015.
- Fornos, Carolina A., Timothy J. Power, and James C. Garand. 2004. 'Explaining Voter Turnout in Latin America, 1980 to 2000'. *Comparative Political Studies* 37(8):909–40.
- Fujiwara, Thomas. 2011. 'A Regression Discontinuity Test of Strategic Voting and Duverger's Law'. *Quarterly Journal of Political Science* 6(3–4):197–233.
- Fujiwara, Thomas. 2015. 'Voting Technology, Political Responsiveness, and Infant Health: Evidence from Brazil'. *Econometrica* 83:423–64.
- Fujiwara, Thomas, Kyle C. Meng, and Tom Vogl. 2014. 'Estimating Habit Formation in Voting'. Available at http://www.princeton.edu/~fujiwara/papers/fujiwara_et_al_habit_formation_2014.pdf, accessed 16 September 2015.
- Gagliarducci, Stefano, and Tommaso Nannicini. 2013. 'Do Better Paid Politicians Perform Better? Disentangling Incentives from Selection'. *Journal of the European Economic Association* 11:369–98.
- Gerber, Alan S., Donald P. Green, and Christopher W. Larimer. 2008. 'Social Pressure and Voter Turnout: Evidence from a Large-Scale Field Experiment'. *American Political Science Review* 102(1):33–48.
- Geys, Benny. 2006. 'Explaining Voter Turnout: A Review of Aggregate-Level Research'. *Electoral Studies* 25(4):637–63.
- Grembi, Veronica, Tommaso Nannicini, and Ugo Troiano. 2012. 'Policy Responses to Fiscal Restraints: A Difference-in-Discontinuities Design'. Available at http://papers.ssrn.com/sol3/papers.cfm?abstract_id=1852523, accessed 16 September 2015.
- Herrera, Helios, Massimo Morelli, and Thomas R. Palfrey. 2013. 'Turnout and Power Sharing'. *The Economic Journal* 124(574):F131–162.
- Hinnerich, Björn Tyrefors, and Per Pettersson-Lidbom. 2014. 'Democracy, Redistribution, and Political Participation: Evidence from Sweden 1919-1938'. *Econometrica* 82(3):961–93.
- Hix, Simon, and Sara Hagemann. 2009. 'Could Changing the Electoral Rules Fix European Parliament Elections?'. *Politique européenne* 2:37–52.
- Imbens, Guido, and Karthik Kalyanaraman. 2012. 'Optimal Bandwidth Choice for the Regression Discontinuity Estimator'. *The Review of Economic Studies* 79(3):933–59.

- Imbens, Guido W., and Thomas Lemieux. 2008. 'Regression Discontinuity Designs: A Guide to Practice'. *Journal of Econometrics* 142(2):615–35.
- Jackman, Robert W. 1987. 'Political Institutions and Voter Turnout in the Industrial Democracies'. *The American Political Science Review* 81(2):405–23.
- Jackman, Robert W., and Ross A. Miller. 1995. 'Voter Turnout in the Industrial Democracies During the 1980s'. *Comparative Political Studies* 27(4):467–92.
- Kartal, Melis. 2013. 'A Comparative Welfare Analysis of Electoral Systems with Endogenous Turnout'. *The Economic Journal* 125:1369–92.
- Karvonen, Lauri. 2004. 'Preferential Voting: Incidence and Effects'. *International Political Science Review* 25(2):203–26.
- Ladner, Andreas, and Henry Milner. 1999. 'Do Voters Turn Out More Under Proportional Than Majoritarian Systems? The Evidence from Swiss Communal Elections'. *Electoral Studies* 18(2):235–50.
- Ledyard, John O. 1984. 'The Pure Theory of Large Two-Candidate Elections'. *Public Choice* 44(1):7–41.
- Lee, David S., and Thomas Lemieux. 2010. 'Regression Discontinuity Designs in Economics'. *The Journal of Economic Literature* 48(2):281–355.
- Lemieux, Thomas, and Kevin Milligan. 2008. 'Incentive Effects of Social Assistance: A Regression Discontinuity Approach'. *Journal of Econometrics* 142(2):807–28.
- Litschig, Stephan, and Kevin M. Morrison. 2013. 'The Impact of Intergovernmental Transfers on Education Outcomes and Poverty Reduction'. *AEJ: Applied Economics* 5(4):206–40.
- Mattila, Mikko. 2003. 'Why Bother? Determinants of Turnout in the European Elections'. *Electoral Studies* 22(3):449–68.
- McCrary, Justin. 2008. 'Manipulation of the Running Variable in the Regression Discontinuity Design: A Density Test'. *Journal of Econometrics* 142(2):698–714.
- Milner, Henry, and Andreas Ladner. 2006. 'Can PR Voting Serve as a Shelter Against Declining Turnout? Evidence from Swiss Municipal Elections'. *International Political Science Review* 27(1):29–45.
- Negri, Margherita. 2014. 'Minority Representation in Proportional Representation Systems'. Available at <http://margheritanegri.files.wordpress.com/2013/03/minority-representation-in-pr-systems.pdf>, accessed 16 September 2015.
- Nickerson, David W. 2008. 'Is Voting Contagious? Evidence from Two Field Experiments'. *American Political Science Review* 102(1):49–57.
- Palfrey, Thomas R., and Howard Rosenthal. 1985. 'Voter Participation and Strategic Uncertainty'. *The American Political Science Review* 79(1):62–78.
- Persson, Torsten, Guido Tabellini, and Francesco Trebbi. 2003. 'Electoral Rules and Corruption'. *Journal of the European Economic Association* 1(4):958–89.
- Pettersson-Lidbom, Per. 2012. 'Does the Size of the Legislature Affect the Size of Government: Evidence from Two Natural Experiments'. *Journal of Public Economics* 96(3–4):269–78.
- Riker, William H., and Peter C. Ordeshook. 1968. 'A Theory of the Calculus of Voting'. *The American Political Science Review* 62(1):25–42.
- Robbins, Joseph W. 2010. 'The Personal Vote and Voter Turnout'. *Electoral Studies* 29(4):661–72.
- Sanz, Carlos. 2015. 'Direct Democracy and Government Size: Evidence from Spain'. Available at: <http://scholar.princeton.edu/carlossanz>
- Schram, Arthur, and Joepa Sonnemans. 1996. 'Voter Turnout as a Participation Game: An Experimental Investigation'. *International Journal of Game Theory* 25(3):385–406.
- Selb, Peter. 2009. 'A Deeper Look at the Proportionality-Turnout Nexus'. *Comparative Political Studies* 42(4):527–48.
- Shugart, Matthew Soberg, Melody Ellis Valdini, and Kati Suominen. 2005. 'Looking for locals: Voter Information Demands and Personal Vote-Earning Attributes of Legislators Under Proportional Representation'. *American Journal of Political Science* 49(2):437–49.
- St-Vincent, Simon Labbé. 2013. 'An Experimental Test of the Pivotal Voter Model Under Plurality and PR Elections'. *Electoral Studies* 32(4):795–806.
- Sweeting, David. 2009. 'The Institutions of Strong Local Political Leadership in Spain'. *Environment and Planning C: Government and Policy* 27:698–712.

APPENDIX

Robustness Checks for the 100-Inhabitant Threshold

In this appendix, I show robustness checks for the results at the 100-inhabitant threshold. Table A1 tests for pretrends and Table A2 shows the results of the donut regression discontinuity design.

TABLE A1 *100-Inhabitant Threshold Robustness: Pretrends in the Outcomes of Interest*

	(1)	(2)	(3)	(4)	(5)	(6)
	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout
Panel A: turnout						
OL	3.177*** (0.588)	3.195*** (0.600)	3.082*** (0.621)	2.974*** (0.672)	2.796*** (0.764)	2.159** (0.978)
OL (+1)	0.152 (0.544)	0.300 (0.563)	0.235 (0.577)	0.285 (0.626)	0.450 (0.661)	0.192 (0.774)
Observations	8742	7555	6263	4915	3588	2242
Municipalities	2504	2220	1921	1588	1256	931
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK
	(1)	(2)	(3)	(4)	(5)	(6)
	Lists	Lists	Lists	Lists	Lists	Lists
Panel B: lists						
OL	0.100** (0.0439)	0.0927** (0.0463)	0.0827* (0.0496)	0.0538 (0.0555)	0.0679 (0.0657)	0.223** (0.0981)
OL(+1)	-0.0554 (0.0434)	-0.0452 (0.0468)	-0.0592 (0.0486)	-0.0653 (0.0533)	-0.000580 (0.0583)	0.0453 (0.0718)
Observations	5401	4596	3813	2999	2146	1341
Municipalities	1694	1513	1316	1117	911	677
Bandwidth	IK	0.85 × IK	0.70 × IK	0.55 × IK	0.40 × IK	0.25 × IK

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman's (2012) bandwidth and 250 inhabitants. OL = open list system.

* $p < 0.10$, ** $p < 0.05$, *** $p < 0.01$.

TABLE A2 *100-Inhabitant Threshold Robustness: Donut Regressions*

	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout	Turnout
Panel A: turnout								
OL	2.811*** (0.521)	2.947*** (0.578)	2.983*** (0.639)	3.448*** (0.803)	2.487*** (0.595)	2.555*** (0.696)	2.484*** (0.804)	3.679*** (1.074)
Observations	11,161	10,739	10,493	9871	5737	5315	5069	4447
Municipalities	2621	2621	2621	2608	1603	1603	1603	1589
Bandwidth	IK	IK	IK	IK	0.50 × IK	0.50 × IK	0.50 × IK	0.50 × IK
Excluded	0	6	10	20	0	6	10	20
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
	Lists	Lists	Lists	Lists	Lists	Lists	Lists	Lists
Panel B: party-lists								
	0.0778** (0.0390)	0.0818* (0.0475)	0.101** (0.0482)	0.147** (0.0669)	0.0517 (0.0506)	0.0506 (0.0730)	0.0941 (0.0827)	0.187 (0.145)
Observations	6907	6485	6239	5617	3571	3149	2903	2281
Municipalities	1823	1823	1823	1810	1190	1184	1180	1144
Bandwidth	IK	IK	IK	IK	0.50 × IK	0.50 × IK	0.50 × IK	0.50 × IK
Excluded	0	6	10	20	0	6	10	20

Note: standard errors clustered by municipality in parentheses. All regressions include municipality and year fixed effects. IK refers to the minimum of Imbens and Kalyanaraman's (2012) bandwidth and 250 inhabitants. OL = open list system.

*p < 0.10, **p < 0.05, ***p < 0.01.