

Machine Learning Assignment # 3

Universität Bern

Due date: 09/10/2019

Late submissions will incur a penalty. Submit your answers in ILIAS (as a pdf or as a picture of your written notes if the handwriting is very clear). Submission instructions will be provided via email. You are not allowed to work with others.

Probability theory review

[Total 100 points]

Solve each of the following problems and show all the steps of your working.

1. Show that the covariance matrix is always symmetric and positive semidefinite.

[15 points]

Solution

The $(i, j)^{th}$ element of the covariance matrix Σ is given by

$$\Sigma_{ij} = E[(X_i - \mu_i)(X_j - \mu_j)] = E[(X_j - \mu_j)(X_i - \mu_i)] = \Sigma_{ji}$$

so that the covariance matrix is symmetric.

For an arbitrary vector u ,

$$\begin{aligned} u^T \Sigma u &= u^T E[(X - \mu)(X - \mu)^T] u = E[u^T (X - \mu)(X - \mu)^T u] \\ &= E[(X - \mu)^T u]^T (X - \mu)^T u = E[(X - \mu)^T u]^2 \geq 0 \end{aligned}$$

so that the covariance matrix is positive semidefinite.

2. $X \in R^n$ and $Y \in R^m$ are independent random variables. Their expectations and covariances are $E[X] = 0$, $\text{Cov}[X] = I$, $E[Y] = \mu$ and $\text{Cov}[Y] = \sigma I$, where I is the identity matrix of the appropriate size and σ is a scalar. What is the expectation and covariance of the random variable $Z = AX + Y$, where $A \in R^{m \times n}$?

[20 points]

Solution.

The expectation of Z can be obtained from the definition by applying the linearity of expectation,

$$E[Z] = E[AX + Y] = AE[X] + E[Y] = 0 + \mu = \mu. \quad (1)$$

The covariance of Z is $\text{Cov}[Z] = E[ZZ^T] - E[Z]E[Z]^T = E[ZZ^T] - \mu\mu^T$. Substituting the definition of Z , we get the expression below.

$$E[ZZ^T] = E[(AX + Y)(AX + Y)^T] = \quad (2)$$

$$= E[AXX^T A^T + YX^T A^T + AXY^T + YY^T] = \quad (3)$$

$$= AE[XX^T]A^T + E[YX^T]A^T + \quad (4)$$

$$AE[XY^T] + E[YY^T]. \quad (5)$$

Here we can substitute $E[XX^T] = I$ and $E[YY^T] = \sigma I + \mu\mu^T$. Because X and Y are independent, $E[XY^T] = E[X]E[Y]^T = 0$, similarly $E[YX^T] = 0$. We get $E[ZZ^T] = AA^T + \sigma I + \mu\mu^T$, therefore $\text{Cov}[Z] = AA^T + \sigma I$.

3. Thomas and Viktor are friends. It is Friday night and Thomas does not have a phone. Viktor knows that there is a $2/3$ probability that Thomas goes to the party to downtown. There are 5 pubs in downtown and there is an equal probability of Thomas going to any of them if he goes to the party. Viktor already looked for Thomas in 4 of the bars.

What is the probability of Viktor finding Thomas in the last bar?

[15 points]

Solution.

The sample space is

$$S = \{\text{"home"}, \text{"pub 1"}, \text{"pub 2"}, \text{"pub 3"}, \text{"pub 4"}, \text{"pub 5"}\}, \quad (6)$$

and the probability of the events are $P(\text{"home"}) = 1/3$ and $P(\text{"pub i"}) = 2/15$. We need to compute $P(\text{"pub 5"} | \text{"not in pub 1 ... 4"})$. Using the Bayes rule,

$$P(\text{"pub 5"} | \text{"not in pub 1 ... 4"}) = \quad (7)$$

$$\frac{P(\text{"pub 5"} \cap \text{"not in pub 1 ... 4"})}{P(\text{"not in pub 1 ... 4"})} = \frac{2/15}{7/15} = \frac{2}{7}. \quad (8)$$

4. Derive the mean for the Beta Distribution, which is defined as

[20 points]

$$\text{Beta}(x|a, b) = \frac{1}{B(a, b)} x^{a-1} (1-x)^{b-1} \quad (9)$$

where $B(a, b)$, $\Gamma(a)$ are Beta and Gamma functions respectively:

$$B(a, b) \triangleq \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)} \quad (10)$$

and

$$\Gamma(x) \triangleq \int_0^\infty u^{x-1} e^{-u} du. \quad (11)$$

Hint: Use integration by parts.

Solution.

$$E[X] = \int_0^1 x \frac{x^{a-1} (1-x)^{b-1}}{B(a, b)} dx \quad (12)$$

To solve this problem, we will integrate by parts.

$$\begin{aligned} E[X] &= \frac{1}{B(a, b)} \left[\frac{x^a (-1)(1-x)^b}{b} \Big|_0^1 - \int_0^1 a x^{a-1} (-1) \frac{(1-x)^b}{b} dx \right] = \\ &= \frac{1}{B(a, b)} \int_0^1 \frac{a}{b} x^{a-1} (1-x)^{b-1} (1-x) dx \end{aligned} \quad (13)$$

$$\begin{aligned} E[X] &= \frac{1}{B(a, b)} \int_0^1 \frac{a}{b} x^{a-1} (1-x)^{b-1} (1-x) dx = \\ &= \frac{1}{B(a, b)} \left[\int_0^1 \frac{a}{b} x^{a-1} (1-x)^{b-1} dx - \int_0^1 x \frac{a}{b} x^{a-1} (1-x)^{b-1} dx \right] = \\ &= \frac{a}{b} (1 - E[X]) \end{aligned} \quad (14)$$

$$E[X] = \frac{a}{a+b} \quad (15)$$

5. Let $A \in \mathbb{R}^{n \times n}$ be a positive definite square matrix, $b \in \mathbb{R}^n$, and c be a scalar. Prove that

[20 points]

$$\int_{x \in \mathbb{R}^n} e^{-\frac{1}{2} x^T A x - x^T b - c} dx = \frac{(2\pi)^{n/2} |A|^{-1/2}}{e^{c - \frac{1}{2} b^T A^{-1} b}}.$$

Hint: Use the fact that the integral of the Gaussian probability density function of a random variable with mean μ and covariance Σ is 1.

Solution

$$\begin{aligned} &= \int_{x \in \mathbb{R}^n} \exp\left(-\frac{1}{2} (x + A^{-1}b)^T A (x + A^{-1}b) + \frac{1}{2} b^T A^{-1}b - c\right) dx \\ &= \int_{x \in \mathbb{R}^n} \frac{1}{(2\pi)^{n/2} |A^{-1}|^{1/2}} \exp\left(-\frac{1}{2} (x + A^{-1}b)^T A (x + A^{-1}b) + \frac{1}{2} b^T A^{-1}b - c\right) dx * (2\pi)^{n/2} |A^{-1}|^{1/2} \\ &= \int_{x \in \mathbb{R}^n} \frac{1}{(2\pi)^{n/2} |A^{-1}|^{1/2}} \exp\left(-\frac{1}{2} (x + A^{-1}b)^T A (x + A^{-1}b)\right) dx * \exp\left(\frac{1}{2} b^T A^{-1}b - c\right) * (2\pi)^{n/2} |A^{-1}|^{1/2} \\ &= \frac{(2\pi)^{n/2} |A|^{-1/2}}{\exp(c - \frac{1}{2} b^T A^{-1}b)} \end{aligned}$$

6. From the definition of conditional probability of multiple random variables, show that

[10 points]

$$f(x_1, x_2, \dots, x_n) = f(x_1) \prod_{i=2}^n f(x_i | x_1, \dots, x_{i-1})$$

where x_1, \dots, x_n are random variables and f is a probability density function of its arguments.

Solution

$$\begin{aligned} f(x_1, x_2, \dots, x_n) &= f(x_n | x_1, x_2, \dots, x_{n-1}) f(x_1, x_2, \dots, x_{n-1}) \\ &= f(x_n | x_1, x_2, \dots, x_{n-1}) f(x_{n-1} | x_1, x_2, \dots, x_{n-2}) f(x_1, x_2, \dots, x_{n-2}) \\ &= f(x_1) \prod_{i=2}^n f(x_i | x_1, \dots, x_{i-1}) \end{aligned}$$