

Machine Learning, Tutorial 4

Universität Bern

Simon Jenni (simon.jenni@inf.unibe.ch)

Logistic Regression

In Logistic Regression we want to maximize the following log likelihood w.r.t. the model parameters θ

$$\ell(\theta) = \sum_{i=1}^m y^{(i)} \log h_{\theta}(x^{(i)}) + (1 - y^{(i)}) \log (1 - h_{\theta}(x^{(i)})), \quad (1)$$

where the hypothesis function h is given by $h_{\theta}(x) = g(\theta^T x) = \frac{1}{1 + e^{-\theta^T x}}$.
Compute the Hessian of the log likelihood.

Solution:

The Hessian matrix is defined as $H_{ij}(\theta) = \frac{\partial^2 \ell(\theta)}{\partial \theta_i \partial \theta_j}$, i.e., it is the matrix of second partial derivatives. Let us first compute the first partial derivatives of $\ell(\theta)$. Remember that $g(z)$ is the sigmoid function with derivative $g'(z) = g(z)(1 - g(z))$.
The first partial derivatives of $\ell(\theta)$ are given by (see lecture notes):

$$\frac{\partial}{\partial \theta_j} \ell(\theta) = \sum_{i=1}^m (y^{(i)} - h_{\theta}(x^{(i)})) x_j^{(i)}. \quad (2)$$

Let us now compute the Hessian, i.e., the second partial derivatives of $\ell(\theta)$:

$$\begin{aligned} \frac{\partial^2}{\partial \theta_k \partial \theta_j} \ell(\theta) &= \frac{\partial}{\partial \theta_k} \left(\sum_{i=1}^m (y^{(i)} - g(\theta^T x^{(i)})) x_j^{(i)} \right) \\ &= - \sum_{i=1}^m g(\theta^T x^{(i)}) (1 - g(\theta^T x^{(i)})) \frac{\partial}{\partial \theta_k} (\theta^T x^{(i)}) x_j^{(i)} \\ &= - \sum_{i=1}^m h_{\theta}(x^{(i)}) (1 - h_{\theta}(x^{(i)})) x_k^{(i)} x_j^{(i)} \end{aligned} \quad (3)$$

We can right this result more succinctly in vector form:

$$H(\theta) = \sum_{i=1}^m h_{\theta}(x^{(i)}) (h_{\theta}(x^{(i)}) - 1) x^{(i)} (x^{(i)})^T \quad (4)$$

Generalized Linear Models

Consider the Laplace distribution with PDF

$$p(y; \mu, \lambda) = \frac{1}{2\lambda} \exp\left(-\frac{|y - \mu|}{\lambda}\right), \quad (5)$$

where $\mu \in \mathbb{R}$ is a location parameter and $\lambda > 0$ is a scale parameter.

Show that the Laplace distribution, if parametrized only on λ (i.e., with fixed μ), is part of the exponential family.

Solution:

The exponential family is defined as all the distributions that can be written in the form

$$p(y; \eta) = b(y) \exp(\eta^T T(y) - a(\eta)) \quad (6)$$

where η is called natural parameter, $T(y)$ is called sufficient statistic and $a(\eta)$ is called log partition function.

The Laplace distribution, if parametrized only on λ , with μ fixed, can be written in the above form:

$$p(y; \lambda) = \exp\left(-\frac{1}{\lambda}|y - \mu|\right) \frac{1}{2\lambda} = \exp\left(-\frac{1}{\lambda}|y - \mu| - \log(2\lambda)\right), \quad (7)$$

that is, $b(y) = 1$, $\eta = -\frac{1}{\lambda}$, $T(y) = |y - \mu|$ and $a(\eta) = \log\left(-\frac{2}{\eta}\right) = \log(2\lambda)$.

Naive Bayes

Consider a binary classification with one binary output y and two binary features x_1 and x_2 . The Naive Bayes classifier assumes the following distribution for a pair:

$$p(y, x_1, x_2) = p(x_1|y) p(x_2|y) p(y). \quad (8)$$

Let the values of the probabilities be:

$$\begin{array}{ll} p(y = 0) = 0.5 & p(y = 1) = 0.5 \\ p(x_1 = 1|y = 0) = 0.9 & p(x_1 = 1|y = 1) = 0.2 \\ p(x_2 = 1|y = 0) = 0.5 & p(x_2 = 1|y = 1) = 0.5 \end{array}$$

1. What are the simplifying assumptions in the Naive Bayes model?

Solution:

The features x_i are assumed to be conditionally independent given the target y .
For example $p(x_1|y, x_2) = p(x_1|y)$.

2. Compute the posterior values $p(y = 1|x_1 = 1, x_2 = 1)$ and $p(y = 0|x_1 = 1, x_2 = 1)$. How would you classify an example with $x_1 = 1$ and $x_2 = 1$?

Solution:

Using Bayes rule we have:

$$\begin{aligned} p(y = 1|x_1 = 1, x_2 = 1) &\propto p(x_1 = 1|y = 1) p(x_2 = 1|y = 1) p(y = 1) = 0.2 * 0.5 * 0.5 \\ p(y = 0|x_1 = 1, x_2 = 1) &\propto p(x_1 = 1|y = 0) p(x_2 = 1|y = 0) p(y = 0) = 0.9 * 0.5 * 0.5 \end{aligned}$$

Therefore we get $p(y = 1|x_1 = 1, x_2 = 1) = 0.2/1.1 \approx 0.18$

The example should be classified as $y = 0$