

## Agenda

- 1) Hypothesis Testing  
motivation
- 2) Type 1 and 2 errors
- 3) NP - Lemma
- 4) Problem.

## Intro to Hypothesis Testing

we get some estimate

$\hat{\theta}$ , how can I test  
hypotheses with this estimate?

we've already talked about  
CI's

Another option is

Hypothesis Testing

So we do Null Hypothesis  
Significance Testing.

So we assume a null  
hypothesis and look to  
reject null.

Simple Hypotheses:

Completely specifies the  
distribution when we assume  
hypothesis.

Suppose our model

$$X \sim \text{bin}(10, p).$$

Simple hypothesis  $p=0.5$ .

Composite hypotheses:  
not simple -

in previous example

$$p \in (0.5, 1)$$

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In general       $H_0$ : null hypothesis

$H_1$ : alternative.

Def. The set of values of  
some statistic  $T$  for  
which  $H_0$  is rejected, is  
called rejection region.

The set of values for which  $H_0$  is not rejected  
the acceptance region.

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We never say accept the null. We say fail to reject the null.

How do we pick these regions?

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Let's first see what can go wrong

In framework  $H_0$  vs  $H_1$

We can make two mistakes.

1. We reject  $H_0$  when  
it is actually true.

(Type 1 error)

2. We fail to reject  $H_0$  when  
it is actually false

(Type 2 error)

Which is worse?

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Court room analogy

Rejecting the null  
is conviction.

Failing to reject is  
not convicting.

So essentially in this analogy

Type I error

Sending an innocent person to  
jail

Type II error

Letting a guilty person go  
free

Notation.

$$P_{H_0}(\text{event})$$

“Probability given  $H_0$  is true of the event”

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$$P(\text{Type I error})$$

$$= P_{H_0}(\text{rejecting } H_0) = \alpha$$

This is called the significance level.

Under composite null.

$$\alpha = \max_{H \in H_0} P_H(\text{type I error})$$

$$P_{H_1}(\text{fail to reject } H_0) = \beta$$

we call  $1 - \beta$  the power

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we want  $\alpha, \beta = 0$ , this  
is impossible.

1) Fix a priori  $\alpha$  to small value.

2) Max  $(1 - \beta)$  power while holding constrained  $\alpha$ .

this is Q3 from  
9.11

Suppose  $X \sim \text{bin}(100, p)$

Consider a test that  
rejects  $H_0: p=0.5$   
vs.  $H_1: p \neq 0.5$  for  
 $|X - 50| > 10.$

Use normal approx to  
find  $\alpha$ , graph power  
function.

$$P_{H_0} (\text{rejecting } H_0)$$

$$= P_{H_0} (|x - s_0| > 10)$$

$$= 1 - P_{H_0} (|x - s_0| \leq 10)$$

$$= 1 - P_{H_0} (-10 \leq x - s_0 \leq 10)$$

$$= 1 - P_{H_0} (90 \leq x \leq 60)$$

$$E(x) = np$$

$$\text{Var}(x) = n p(1-p)$$

$$X \sim N(np, np(1-p))$$

$$= 1 - P_{H_0} \left( \frac{50 - np}{\sqrt{np(1-p)}} \leq Z \leq \frac{60 - np}{\sqrt{np(1-p)}} \right)$$

$$Z = \frac{X - np}{\sqrt{np(1-p)}} \sim N(0, 1)$$

$$= 1 - P \left( \frac{50 - s_0}{s} \leq Z \leq \frac{60 - s_0}{s} \right)$$

$$= 1 - P(-z \leq Z \leq z)$$

$$= 1 - (\Phi(z) - \Phi(-z))$$

$\Phi$   $\rightarrow$   
Standard normal CDF

$$\approx 0.0455$$

b) we can get power

$$1 - \beta = 1 - P_{H_0} (\text{fail to reject } H_0)$$

$$= P_{H_1} (\text{reject } H_0)$$

$$= 1 - P_{H_1} \left( \frac{90 - np}{\sqrt{np(1-p)}} \leq Z \leq \frac{60 - np}{\sqrt{np(1-p)}} \right)$$

$$= 1 - P_{H_1} \left( \frac{90 - 100p}{10\sqrt{p(1-p)}} \leq Z \leq \frac{60 - 100p}{10\sqrt{p(1-p)}} \right)$$

All I'm told is

that  $p \neq 0.5$

$$= 1 - \left( \Phi \left( \frac{60 - 100p}{10\sqrt{p(1-p)}} \right) - \Phi \left( \frac{40 - 100p}{10\sqrt{p(1-p)}} \right) \right)$$

at  $p = 0.5$

$$P_{H_0}(\text{reject}) = \alpha$$

How do we actually  
construct a test

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Is there a way to  
find a most powerful  
 $\alpha$  level test easily?

Neyman - Pearson Lemma  
is a way for simple-  
simple tests.

Likelihood Ratio test.

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Ex. of a simple - simple test.  
 $H_0: p=0.5 \quad H_1: p=0.7 \quad (\text{binom})$

Ex of simple - composite

$$H_0: \rho = 0.5 \quad H_1: \rho \neq 0.5$$

Interlude

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$$P_{H_1}(\text{fail to reject } H_0) = \beta$$

we call  $1 - \beta$  the power

$$1 - P_{H_1}(\text{fail to reject } H_0)$$

$$= P_{H_1}(\text{reject } H_0)$$

# Likelihood Ratio Test

Suppose simple-simple hypotheses

We can get a likelihood under both the alternative and the null

Denote  $f_{H_0}(\underline{x})$  or  $f_0(\underline{x})$  as the likelihood under the null. (similar notation alternative)

$$LR(\underline{x}) = \frac{f_{H_0}(\underline{x})}{f_{H_1}(\underline{x})}$$

if  $H_1$  is more "likely"

$$f_{H_1}(x) > f_{H_0}(x)$$

e.g.

$$\frac{f_{H_0}(x)}{f_{H_1}(x)} \text{ is small}$$

### NP-Lemma

Suppose we have a test  
that rejects  $H_0$  when

$$\frac{f_{H_0}(x)}{f_{H_1}(x)} < c(\alpha) \text{ has}$$

Significance Level  $\alpha$ .

Then this is the most powerful  $\alpha$  level test for simple-simple.

### Process

- 1) Compute Likelihood Ratio.
- 2) Find statistic  $T(x)$  that determines when L.R is small
- 3) Determine rejection region for  $T$

4) Choose cutoff to ensure  
 $\alpha$  level test.

Q7 Ch 9.11

Let  $x_1, \dots, x_n$  be a sample from Poisson dist.

Find LRT for

$$H_0: d = d_0$$

$$H_1: d = d_1, \text{ know } d_1 > d_0$$

if  $x_i \stackrel{\text{ind}}{\sim} \text{pois}(d)$

$$P(X_i = x_i) = \frac{e^{-d} d^{x_i}}{x_i!}$$

$$L_n(\lambda) = \prod_{i=1}^n \frac{e^{-\lambda} \lambda^{x_i}}{x_i!}$$

$$= e^{-n\lambda} \cdot \lambda^{\sum x_i}$$

$$\frac{}{\prod x_i!}$$

$$LR(x) = \frac{f_{H_0}(x)}{f_{H_1}(x)} = \frac{L_n(\lambda_0)}{L_n(\lambda_1)}$$

$$= \frac{e^{-n\lambda_0} \lambda_0^{\sum x_i}}{e^{-n\lambda_1} \lambda_1^{\sum x_i}}$$

$$\cancel{\frac{\cdot \lambda_0^{\sum x_i}}{\prod x_i!}}$$

$$\cancel{\frac{\cdot \lambda_1^{\sum x_i}}{\prod x_i!}}$$

$$= e^{-n d_0 - n d_1} \cdot \left( \frac{d_0}{d_1} \right)^{\sum x_i} \quad (1)$$

Let  $T = \sum x_i$

$$= e^{-n d_0 - n d_1} \left( \frac{d_0}{d_1} \right)^T$$

$$d_1 \geq d_0 \Rightarrow \frac{d_0}{d_1} < 1 \quad (2)$$

This  $LR(\chi)$  is small when

$T(\chi)$  is large

So CRT reject when

$$T(x) = \sum x_i \text{ is large} \quad (3)$$

or vice.

$$T(x) = \sum x_i > c(\alpha)$$

(4) we want

$$P_{H_0}(\text{rejecting } H_0) = \alpha$$

$$= P_{H_0} \left( \sum_{i=1}^n x_i > c(\alpha) \right).$$

$$T = \sum_{i=1}^n x_i \sim \text{posS}(n\mu_0)$$

under null

$$= P_{H_0} (T \geq c(\alpha))$$

$$= 1 - P_{H_0} (T \leq c(\alpha)) = \alpha$$

$$= 1 - \sum_{t=0}^{\lfloor c(\alpha) \rfloor} \frac{(n_{do})^t e^{-n_{do}}}{t!} = \alpha$$

Power

$$P_{H_1} (\text{reject } H_0)$$

$$= P_{H_1} (T > c(\alpha))$$

$$= 1 - P_{H_1} (T \leq c(\alpha))$$

now known!

under alt.  $T \sim \text{Pois}(n\lambda_1)$

$$= 1 - \sum_{t=0}^{C(\alpha)} e^{-n\lambda_1} (n\lambda_1)^t / t!$$

imagine  $C(\alpha)$  is an integer

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## Dog Cancer problem

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$$H_0: p = \frac{1}{2} \quad H_1: p > \frac{1}{2}$$

$$X_1, \dots, X_{54} \stackrel{\text{ind}}{\sim} \text{bern}(p)$$

use CRT even  $H_0: p = p_0$

though not simple.  $H_1: p = p_1 \neq p_0$

$$L_n(p; x) = \prod_{i=1}^{54} p^{x_i} (1-p)^{1-x_i}$$

$$= p^{\sum x_i} (1-p)^{54 - \sum x_i}$$

$$LR(x) = \frac{f_{H_0}(x)}{f_{H_1}(x)} = \frac{p_0^{\sum x_i} (1-p_0)^{54 - \sum x_i}}{p_1^{\sum x_i} (1-p_1)^{54 - \sum x_i}}$$

$$= \left( \frac{p_0}{p_1} \right)^{\sum x_i} \left( \frac{1-p_0}{1-p_1} \right)^{54 - \sum x_i}$$

$$= \left( \frac{p_1}{p_0} \right)^{-\sum x_i} \left( \frac{1-p_0}{1-p_1} \right)^{54 - \sum x_i}$$

$$= \left( \frac{p_1}{p_0} \right)^{-54} \left( \frac{p_1}{p_0} \right)^{54 - \sum x_i} \left( \frac{1-p_0}{1-p_1} \right)^{54 - \sum x_i}$$

$$= \left( \frac{p_1}{p_0} \right)^{-54} \begin{pmatrix} p_1 & 1-p_0 \\ p_0 & 1-p_1 \end{pmatrix}^{54 - \sum x_i}$$

under null  $p_1 > p_0$ .

$$\frac{P_1}{P_0} > 1 \quad \frac{(1-P_0)}{1-P_1} > 1$$

$\Rightarrow$  dec function of  $\sum X_i$

So we reject for  
small  $L R(\underline{x})$



reject for large  $T(\underline{x}) = \sum X_i$

under null

$$P_{H_0}(T > c(\alpha)) = \alpha$$

$$= 1 - P_{H_0}(T \leq c(\alpha)) = \alpha.$$

*under null*  
 $T \sim \text{binom}(50, \frac{1}{2})$

Suppose I give  $\alpha = 0.05$ .

$$= 1 - P_{H_0} (\sum X_i \leq c) = 0.05.$$