

More linear model theory

Recall: our model setup

True model:

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$
$$i=1, \dots, n$$

$$\varepsilon_i \stackrel{\text{ind}}{\sim} N(0, \sigma^2)$$

equivalent

$$y_i \stackrel{\text{ind}}{\sim} N(\beta_0 + \beta_1 x_i, \sigma^2)$$

what assumptions are implicit
in this?

1. Homoscedasticity

Every ϵ_i has SAME
variance

2. Independence of noise/
response

last time showed

BST $\hat{\beta}_0, \hat{\beta}_1$ via Least squares

or equivalently its MLE

for $y_i \stackrel{iid}{\sim} N(\beta_0 + \beta_1 x_i, \sigma^2)$

ASSUME NOTHING

only assure true model
is

$$y_i = \beta_0 + \beta_1 x_i + \varepsilon_i$$

assume $E(\varepsilon_i) = 0$

$\Rightarrow \hat{\beta}_0, \hat{\beta}_1$ unbiased

or

$$E(\hat{\beta}_0) = \beta_0$$

$$E(\hat{\beta}_1) = \beta_1$$

also

$$\text{var}(\hat{\beta}_1) = \frac{\sigma^2}{S_{xx}}$$

$$S_{ab} = \sum_{i=1}^n (a_i - \bar{a})(b_i - \bar{b})$$

Let

$$\hat{\epsilon}_i = y_i - \hat{y}_i \quad \text{"the residual"}$$

$$= y_i - (\hat{\beta}_0 + \hat{\beta}_1 x_i)$$

$\hat{\epsilon}_i$ are usually our estimates
for ϵ_i

Assume homoscedasticity

$$\varepsilon_i \stackrel{\text{ind}}{\sim} [0, \sigma^2]$$

$$\hat{\sigma}_e^2 = \frac{\sum_{i=1}^n \hat{\varepsilon}_i^2}{n-2}$$

$$E(\hat{\sigma}_e^2) = \sigma_e^2$$

Now assume

$$\varepsilon_i \stackrel{iid}{\sim} N(0, \sigma^2)$$

then

$$\frac{\hat{\beta}_{1, LS} - \beta_1}{\hat{\sigma}_e^2 / \sqrt{s_{xx}}} \sim t_{n-2}$$

How to check assumptions
on noise? ϵ_i

Use residuals $\hat{\epsilon}_i$

SOME VECTORS

and

MATRICES

I will try to underline
vectors

Let $\underline{a}, \underline{b} \in \mathbb{R}^n$

By default all vectors are

column vectors ($n \times 1$ matrix)

$$\underline{a} = \begin{pmatrix} a_1 \\ \vdots \\ a_n \end{pmatrix} \quad \underline{b} = \begin{pmatrix} b_1 \\ \vdots \\ b_n \end{pmatrix}$$

$$\underline{a}^T \underline{b} = (a_1 \dots a_n) \begin{pmatrix} b_1 \\ \vdots \\ b_n \end{pmatrix}$$

$$= a_1 b_1 + a_2 b_2 + \dots + a_n b_n$$

$$= \sum_{i=1}^n a_i b_i$$

$$\underline{a}^T \underline{a} = a_1^2 + a_2^2 + \dots + a_n^2$$

$$\sqrt{\underline{a}^T \underline{a}} = \sqrt{a_1^2 + a_2^2 + \dots + a_n^2}$$

$$= \|\underline{a}\|_2$$

$$\underline{a}^T \underline{a} = \|\underline{a}\|_2^2$$

matrix multiplication

$$A B = C$$

$$(n \times m) \cdot (m \times p)$$

$$A \in \mathbb{R}^{n \times m}$$

$$B \in \mathbb{R}^{m \times p}$$

$$C \in \mathbb{R}^{n \times p}$$

Matrices and Random vectors

Let $\underline{Y} \in \mathbb{R}^n$ be

a random vector

$$E(\underline{Y}) = \begin{pmatrix} E(Y_1) \\ \vdots \\ E(Y_n) \end{pmatrix}$$

$$\text{cov}(Y_i, Y_j) = \sigma_{ij}$$

$$\sigma_{ij} = \sigma_{ji} \quad \sigma_{ii} = \text{var}(Y_i)$$

$$\text{cov}(\underline{Y}) = \{\sigma_{ij}\}$$

$$= \begin{pmatrix} \sigma_{11} & \sigma_{12} & \sigma_{13} & \dots & \sigma_{1n} \\ \sigma_{21} & \sigma_{22} & & & \\ \vdots & \ddots & \ddots & \ddots & \\ \sigma_{n1} & \dots & & & \sigma_{nn} \end{pmatrix}$$

Thm A

$\underline{Y} \in \mathbb{R}^n$ random

$\underline{C} \in \mathbb{R}^n$ constant

$A \in \mathbb{R}^{m \times n}$ constant matrix

if $\underline{z} = \underline{c} + A\underline{y}$

then

$$E(\underline{z}) = \underline{c} + A E(\underline{y})$$

take the i th element
of \underline{z} .

$$z_i = c_i + \sum_{j=1}^n a_{ij} y_j$$

$$E(z_i) = c_i + \sum_{j=1}^n a_{ij} E(y_j)$$

$$\Rightarrow E(\underline{z}) = \underline{c} + A E(\underline{y})$$

Theorem B

Same setup

$$\text{if } \underline{z} = \underline{c} + A\underline{y}$$

$$\text{then } \text{cov}(\underline{z}) = A \text{cov}(\underline{y}) A^T$$

$$\text{cov}(z_i, z_j)$$

$$= \text{cov}\left(c_i + \sum_{k=1}^n a_{ik} y_k, c_j + \sum_{\ell=1}^n a_{j\ell} y_\ell\right)$$

$$= \text{cov}\left(\sum_{k=1}^n a_{ik} y_k, \sum_{\ell=1}^n a_{j\ell} y_\ell\right)$$

$$= \sum_{k=1}^n a_{ik} \left(\sum_{l=1}^n a_{jl} \text{cov}(Y_k, Y_l) \right)$$

$$= \sum_{k=1}^n \sum_{l=1}^n a_{ik} a_{jl} \text{cov}(Y_k, Y_l)$$

$$= \sum_{k=1}^n \sum_{l=1}^n a_{ik} a_{jl} \sigma_{kl} \quad \checkmark$$

Theorem C

$$\mathbb{E}(\underline{x}^T A \underline{x}) = \text{trace}(A\Sigma) + \underline{\mu}^T A \underline{\mu}$$

Thanks to Peng Ding for
this proof!

\underline{X} random vector with
mean $\underline{\mu}$ and covariance
 Σ , A is fixed matrix

2 facts

$$\text{cov}(\underline{Y}) = E(\underline{YY^T}) - E(\underline{Y})E(\underline{Y^T})$$

\Rightarrow

$$E(\underline{YY^T}) = \text{cov}(\underline{Y}) + E(\underline{Y})E(\underline{Y^T})$$

$$E(\text{trace}(w)) = \text{trace}(E(w))$$

Pf.

$$E(\underline{x^T A x}) = E(\text{tr}(\underline{x^T A x}))$$

by properties of trace

$$= E(\text{tr}(\underline{A x x^T}))$$

$$= \text{tr}(E(\underline{A x x^T}))$$

$$= \text{tr}(\underline{A E(x x^T)})$$

$$= \text{tr}(A (\Sigma + \underline{\mu \mu^T}))$$

$$= \text{tr}(A \Sigma) + \text{tr}(A \underline{\mu \mu^T})$$

$$= \text{tr}(A\Sigma) + \text{tr}(\underline{\mu}^T A \underline{\mu})$$

$$= \text{tr}(A\Sigma) + \underline{\mu}^T A \underline{\mu}$$

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Let Z_1, Z_2, Z_3, Z_4

be r.v. $\text{var}(Z_i) = 1$

$\text{cov}(Z_i, Z_j) = \rho \quad i \neq j$

Show

$Z_1 + Z_2 + Z_3 + Z_4$ uncor.

$Z_1 + Z_2 - Z_3 - Z_4$

Define

$$\underline{z} = \begin{pmatrix} z_1 \\ z_2 \\ z_3 \\ z_4 \end{pmatrix}$$

$$A = (1 \ 1 \ 1 \ 1)$$

$$Az = z_1 + z_2 + z_3 + z_4$$

$$B = (1 \ 1 -1 -1)$$

$$Bz = z_1 + z_2 - z_3 - z_4$$

$$\Rightarrow \text{cov}(Az, Bz)$$

Theorem D

$$Y = AX \quad Z = BX$$

$$\text{cov}(Y, Z)$$

$$= \text{cov}(AX, BX)$$

$$= A \text{cov}(X) B^T$$

$$\text{cov}(AZ, BZ)$$

$$= A \text{cov}(Z) B^T$$

$$= (1 \ 1 \ 1 \ 1) \begin{pmatrix} \sigma^2 & & & \\ & \sigma^2 & & \\ & & \sigma^2 & \\ & & & \sigma^2 \end{pmatrix} \begin{pmatrix} 1 \\ 1 \\ 1 \\ -1 \end{pmatrix}$$

= 0

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x_1, \dots, x_n

$$\text{var}(x_i) = \sigma^2$$

$$\text{cov}(x_i, x_j) = \rho \sigma^2$$

$\text{Var}(\bar{x})$

$$\underline{x} = \begin{pmatrix} x_1 \\ \vdots \\ x_n \end{pmatrix} \quad \underline{\Sigma} = \begin{pmatrix} 1 & & \\ & \ddots & \\ & & n \end{pmatrix}$$

$$\text{Var}(\bar{X}) = \text{Cov}(\bar{X}, \bar{X})$$

$$= \text{Cov}\left(\frac{1}{n} \mathbf{1}^T \mathbf{X}, \frac{1}{n} \mathbf{1}^T \mathbf{X}\right)$$

use Theorem D

$$= \frac{1}{n} \mathbf{1}^T \text{Cov}(\mathbf{X}) \frac{1}{n}$$

$$= \frac{1}{n} \mathbf{1}^T \begin{pmatrix} \sigma^2 & \rho\sigma^2 \\ \vdots & \vdots \\ \rho\sigma^2 & \sigma^2 \end{pmatrix} \frac{1}{n}$$

$$= \frac{1}{n} \mathbf{1}^T \begin{pmatrix} \frac{1}{n} ((n-1)\rho\sigma^2 + \sigma^2) \\ \vdots \\ \frac{1}{n} ((n-1)\rho\sigma^2 + \sigma^2) \end{pmatrix}$$

$$= \sum_{i=1}^n \frac{1}{n^2} ((n-1)\rho\sigma^2 + \sigma^2)$$

$$= \frac{1}{n} ((n-1)\rho\sigma^2 + \sigma^2)$$

if $\rho = 0$

$$\Rightarrow \text{var}(\bar{x}) = \frac{\sigma^2}{n}$$

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$$\underline{z} = \begin{pmatrix} z_1 \\ z_2 \\ z_3 \\ z_4 \end{pmatrix} \quad \text{cov}(\underline{z}) = \sigma^2 I$$

$$U = z_1 + z_2 + z_3 + z_4$$

$$V = (z_1 + z_2) - (z_3 + z_4)$$

$$A = \begin{pmatrix} 1 & 1 & 1 & 1 \end{pmatrix}$$

$$B = \begin{pmatrix} 1 & 1 & -1 & -1 \end{pmatrix}$$

$$Az = U \quad AB = V$$

$$\text{cov}(U, V)$$

$$= \text{cov}(Az, AB)$$

$$= A \text{cov}(z) B^T$$

$$= A \sigma^2 I \cdot B^T$$

$$= \sigma^2 A B^T$$

$$= \sigma^2 (1 \ 1 \ 1) \begin{pmatrix} 1 \\ -1 \\ -1 \end{pmatrix}$$

$$= 0$$

Q2S

$$y_{i1}, y_{i2}, \bar{y}_i = \frac{(y_{i1} + y_{i2})}{2}$$

$$y_{i1} = \beta_0 + \beta_1 x_i + \epsilon_{i1}$$

$$Y_{i2} = \beta_0 + \beta_1 X_i + \varepsilon_{i2}$$

$$\bar{Y}_i = \frac{Y_{i1} + Y_{i2}}{2}$$

$$= \frac{\beta_0 + \beta_1 X_i + \varepsilon_{i1} + \beta_0 + \beta_1 X_i + \varepsilon_{i2}}{2}$$

$$= \beta_0 + \beta_1 X_i + \frac{\varepsilon_{i1} + \varepsilon_{i2}}{2}$$

$$= \beta_0 + \beta_1 X_i + \varepsilon_i'$$

Solving this with least squares!