

On sampling determinantal point processes

Guillaume Gautier

Ph.D. defense

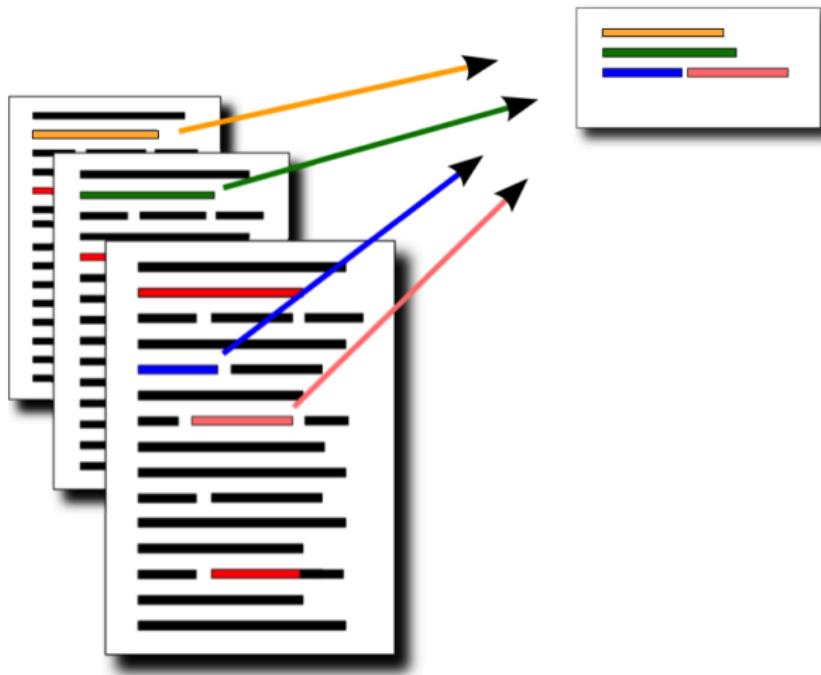


Advisors: Rémi Bardenet, Michal Valko

May 19, 2020

Text summarization

Extract **diverse** sentences of a large corpus to build a representative summary.



Recommendation systems

Two possible sets of answers of an image search engine to the query “bolt”.

relevance



relevance
+
diversity

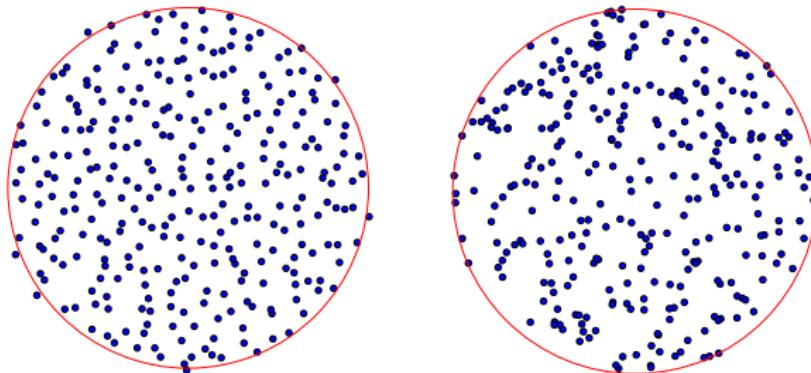


- ▶ Use DPPs to enforce diversity among the recommended items.

Numerical integration

Use random **repulsive** points as quadrature nodes to estimate an integral

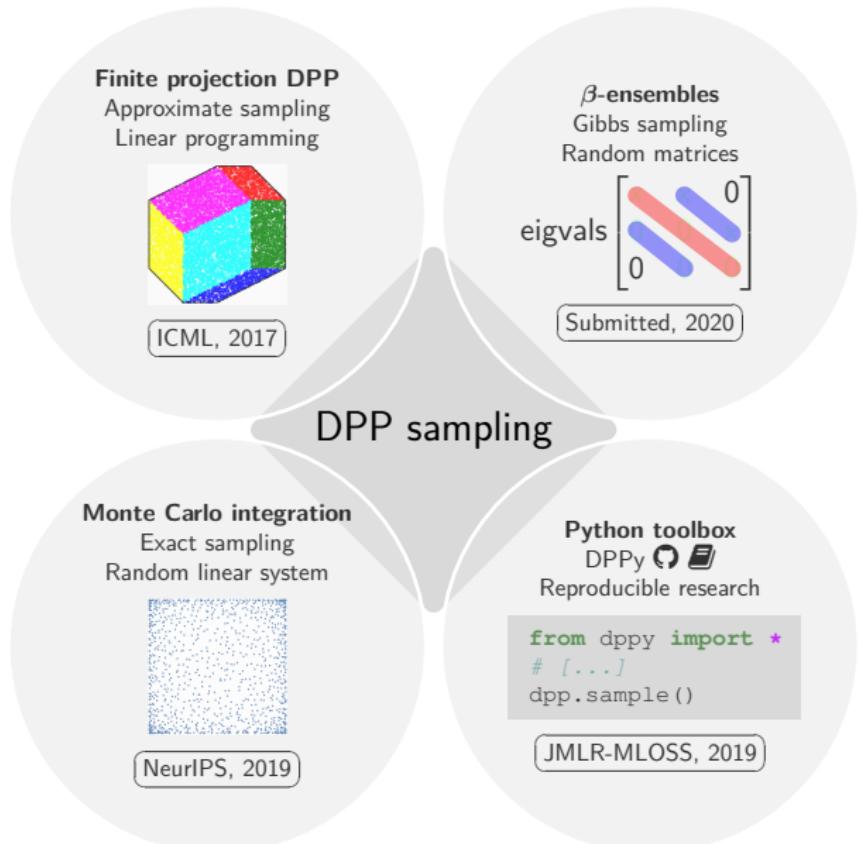
$$\int f(x)\mu(dx) \approx \sum_{n=1}^N \omega_n f(x_n).$$



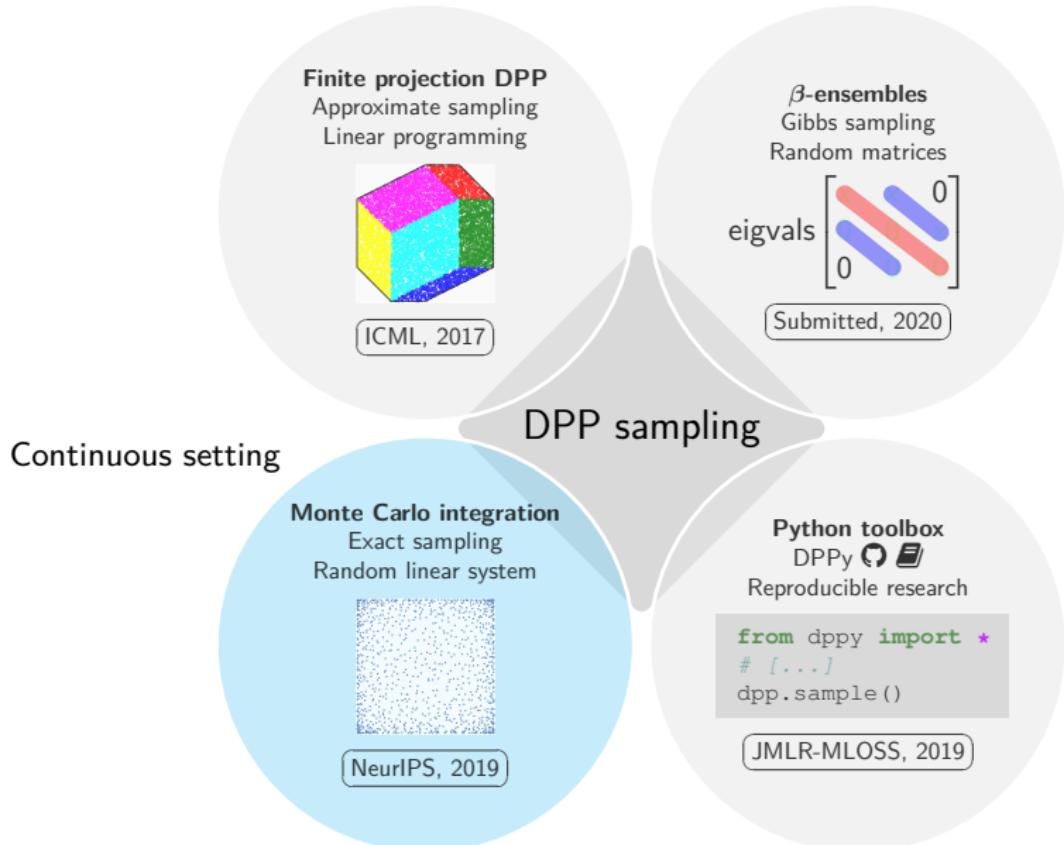
Bardenet and Hardy (2016, 2020)

- ▶ Prove faster rate of convergence with DPP points than i.i.d. points.
- ▶ Efficient sampler to put theory into practice?

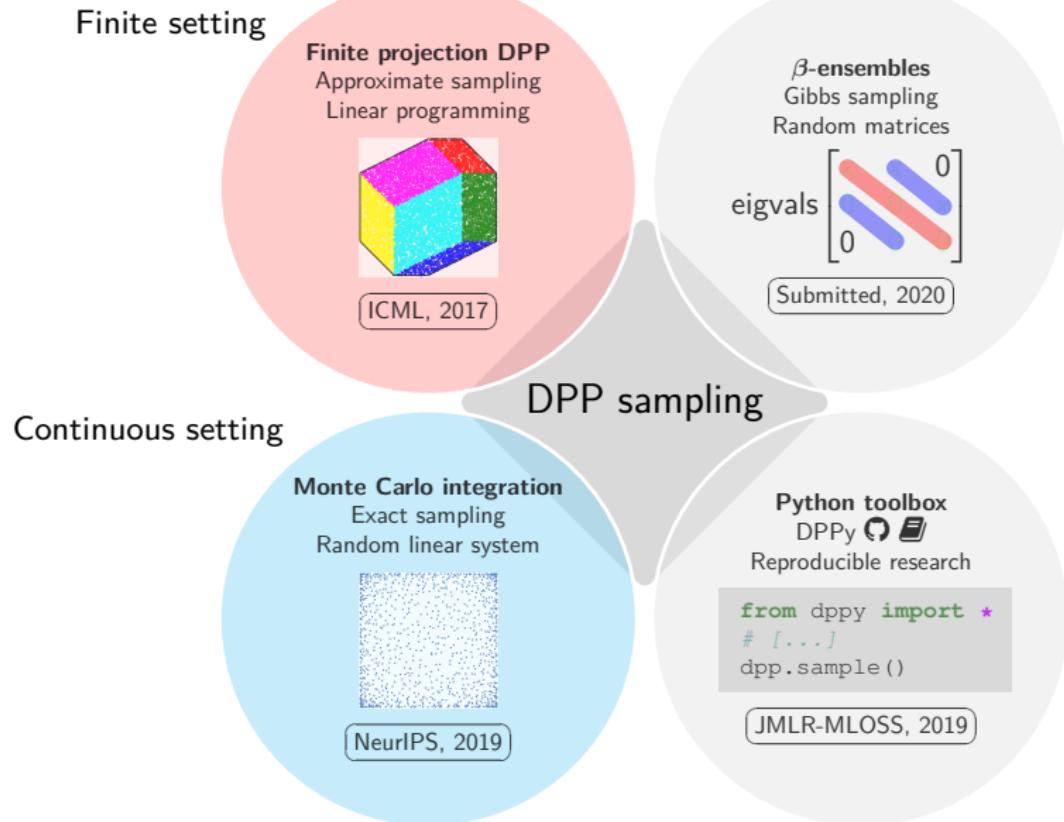
My Ph.D. in a nutshell



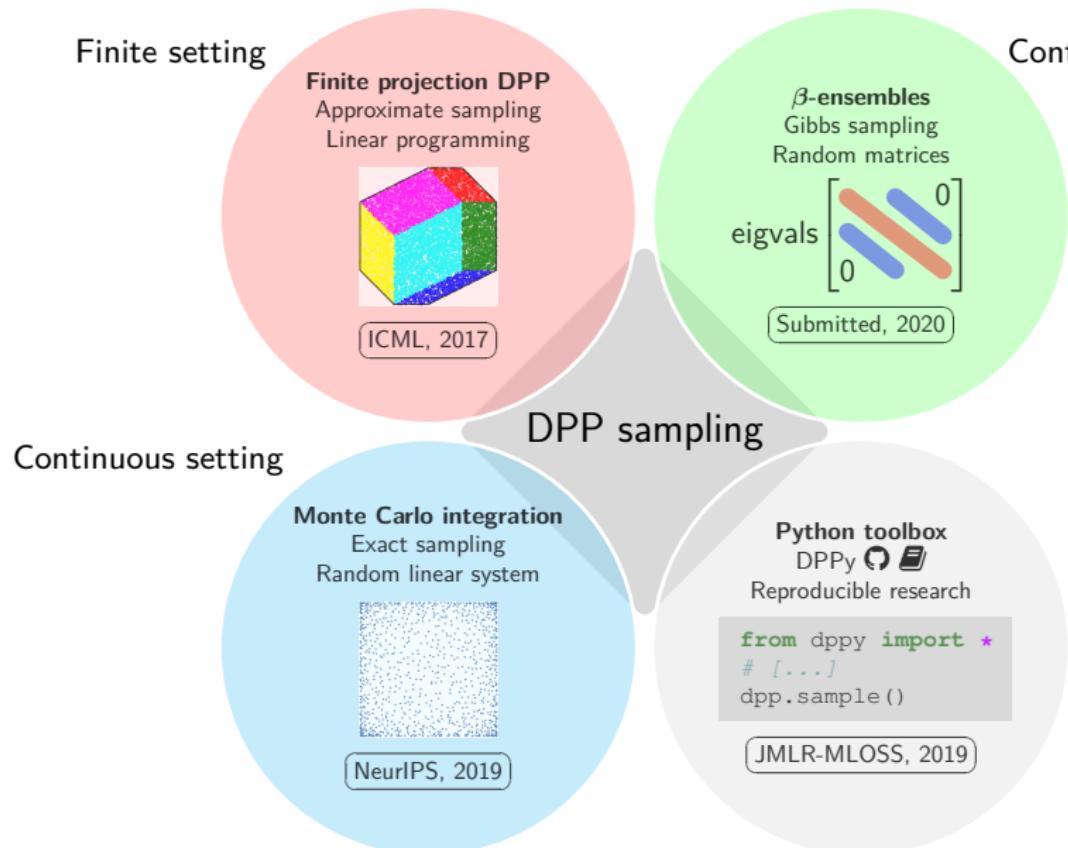
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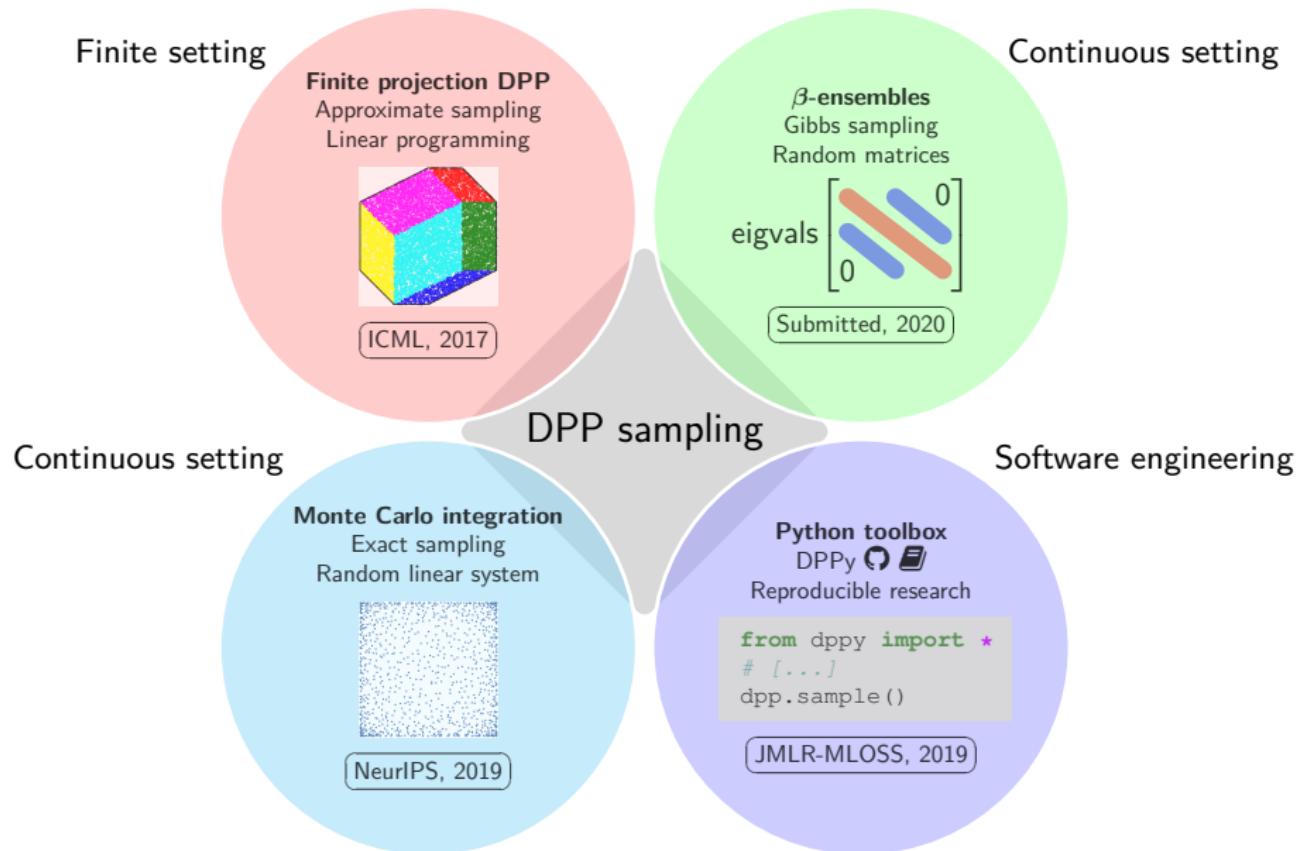
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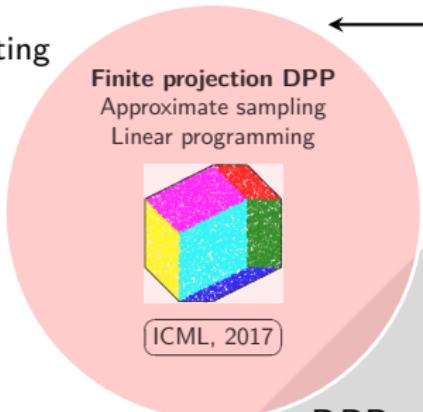


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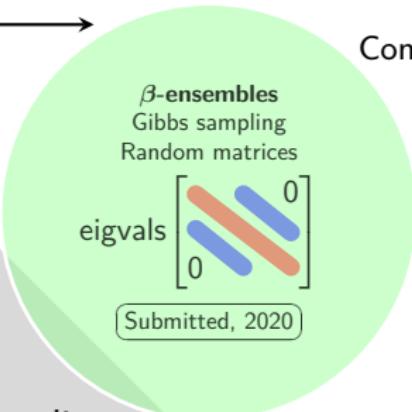


Focus of the presentation

Finite setting

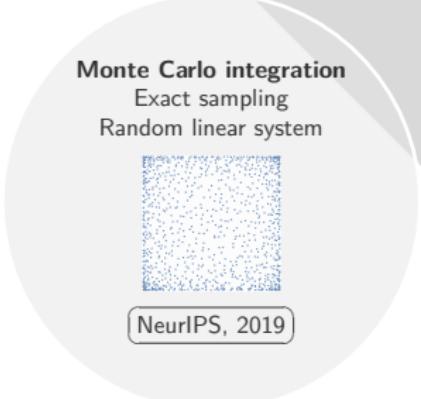


Continuous setting



DPP sampling

Monte Carlo integration
Exact sampling
Random linear system



Python toolbox
DPPy
Reproducible research

```
from dppy import *
# [...]
dpp.sample()
```

JMLR-MLOSS, 2019

Overview

Introduction

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Some insights on finite DPPs

Finite projection DPPs

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Exact sampling from finite projection DPPs

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Summary of contributions

Open questions and perspectives

Some insights on finite DPPs

- ▶ Point process

$$\mathcal{X} \subset \left\{ \begin{array}{c} \text{(Athlete in starting blocks)}, \\ \text{(Athlete running)}, \\ \text{(Cartoon dog)}, \\ \text{(Car)}, \\ \text{(Bolt)}, \\ , \\ , \\ \dots \\ , \\ \text{(Cartoon dog)} \end{array} \right\}$$

Some insights on finite DPPs

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- ▶ Diversity

$$\mathbb{P}[\{\text{(Car)}, \text{(Athlete)}\} \subset \mathcal{X}] \geq \mathbb{P}[\{\text{(Dog)}, \text{(Athlete)}\} \subset \mathcal{X}]$$

Some insights on finite DPPs

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$$\mathcal{X} \subset \left\{ \begin{array}{c} \text{(Athlete in a race)} \\ , \end{array} \begin{array}{c} \text{(Athlete running)} \\ , \end{array} \begin{array}{c} \text{(Cartoon dog)} \\ \text{ROTT} \end{array}, \begin{array}{c} \text{(Car)} \\ , \end{array} \begin{array}{c} \text{(Bolt)} \\ , \end{array} \dots, \begin{array}{c} \text{(Cartoon dog)} \\ \end{array} \right\}$$

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- ▶ Similarity matrix

$$\mathbf{K}$$


Some insights on finite DPPs

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$$\mathcal{X} \subset \left\{ \begin{matrix} \text{(Athlete)} \\ \text{(Athlete)} \\ \text{(Dog)} \\ \text{(Car)} \\ \text{(Bolt)} \\ , \quad , \quad , \quad , \quad , \quad \cdots \quad , \quad \text{(Dog)} \end{matrix} \right\}$$

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$$\mathbf{K} \begin{matrix} \text{(Car)} \\ \text{(Athlete)} \end{matrix}$$

- ▶ Inclusion probabilities

$$\mathbb{P}[\{\text{(Car)}, \text{(Athlete)}\} \subset \mathcal{X}] = \det \begin{bmatrix} \mathbf{K} & \text{(Car)} & \text{(Car)} \\ \text{(Athlete)} & \mathbf{K} & \text{(Athlete)} \\ \text{(Car)} & \text{(Athlete)} & \mathbf{K} \end{bmatrix}$$

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- ▶ Sufficient conditions for existence

$$\mathbf{K}^T = \mathbf{K} \quad \text{and} \quad \mathbf{0} \preceq \mathbf{K} \preceq I.$$

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Finite projection DPP

Definition

Consider $\mathbf{K} \in \mathbb{R}^{M \times M}$ such that $\mathbf{K}^T = \mathbf{K}$ and $\mathbf{K}^2 = \mathbf{K}$.

The point process \mathcal{X} defined by

$$\mathbb{P}[S \subset \mathcal{X}] = \det \mathbf{K}_S, \quad \forall S \subset \{1, \dots, M\},$$

is called a projection DPP with kernel \mathbf{K} .

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$$N \triangleq |\mathcal{X}| = \text{rank } \mathbf{K}.$$

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$$\mathbb{P}[\mathcal{X} = B] = \det \mathbf{K}_B \mathbb{1}_{|B|=N}.$$

Finite projection DPP

Example

Consider the $N \times M$ feature matrix $\Phi = [\phi_1, \dots, \phi_M]$, such that rank $\Phi = N$, and build the kernel

$$\mathbf{K} = \Phi^\top [\Phi \Phi^\top]^{-1} \Phi.$$

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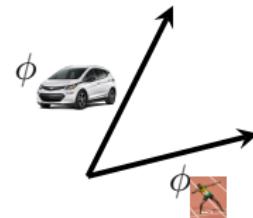
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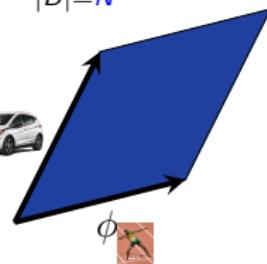
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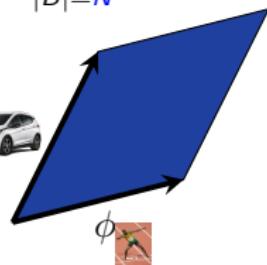
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$$\mathbb{P}[\mathcal{X} = \{\text{car}, \text{truck}\}] \propto \text{volume}^2$$



- ▶ The support is formed by collections of columns of Φ forming a basis of \mathbb{R}^N ,

$$\mathcal{B} \triangleq \{B ; |B| = N, \text{ and } \det \Phi_{:B} \neq 0\}.$$

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Continuous projection DPP

Definition

Let $\phi_0, \dots, \phi_{N-1}$ be orthonormal functions in $L^2(\mathbb{X}, \mu)$ and

$$K(x, y) = \sum_{k=0}^{N-1} \phi_k(x)\phi_k(y).$$

Take (x_1, \dots, x_N) with joint probability distribution

$$\frac{1}{N!} \det[K(x_i, x_j)]_{i,j=1}^N \prod_{n=1}^N \mu(dx_n).$$

Then $\mathcal{X} \triangleq \{x_1, \dots, x_N\} \subset \mathbb{X}$ defines a projection DPP with kernel K .

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Considering

- ▶ $\mathbb{X} = \{1, \dots, M\}$,

- ▶ $\mu = \sum_{m=1}^M \delta_m$,

one recovers the finite case with $\mathbf{K} = \Phi^\top \Phi$ and $\Phi \Phi^\top = I_N$.

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Sequential sampling using the chain rule

The goal is to generate a random **subset** \mathcal{X} , such that

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(Hough et al., 2006; Gillenwater, 2014)

Consider the eigendecomposition $\mathbf{K} = \Phi^T \Phi$, where $\Phi \Phi^T = I_N$.

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$$\mathbb{P}[(x_1, \dots, x_n)] = \frac{1}{N!} \text{volume}^2\{\phi_{x_1}, \dots, \phi_{x_N}\} = \frac{1}{N!} \det \mathbf{K}_{\{x_1, \dots, x_N\}}.$$

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- ▶ The procedure is akin to Gram-Schmidt orthogonalization $\mathcal{O}(MN^2)$.

Illustration of the chain rule ($M = 24, N = 2$)

Text to summarize using $N = 2$ sentences.

*But a dream within a dream?
Is all that we see or seem?
One from the pitiless wave?
O God! can I not save
Them with a tighter clasp?
O God! can I not grasp
While I weep--while I weep!
Through my fingers to the deep,
How few! yet how they creep
Grains of the golden sand--
And I hold within my hand
Of a surf-tormented shore,
I stand amid the roar
Is but a dream within a dream.
All that we see or seem
Is it therefore the less gone?
In a vision, or in none,
In a night, or in a day,
Yet if hope has flown away
That my days have been a dream;
You are not wrong, who deem
Thus much let me avow--
And, in parting from you now,
Take this kiss upon the brow!*

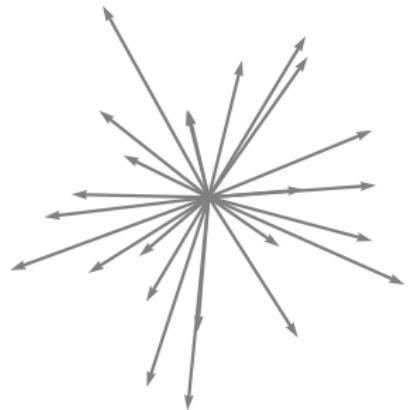
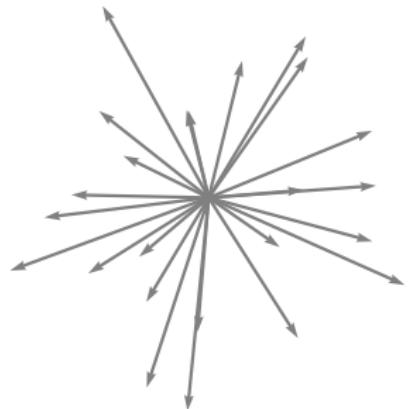


Illustration of the chain rule ($M = 24, N = 2$)

Select the first sentence,

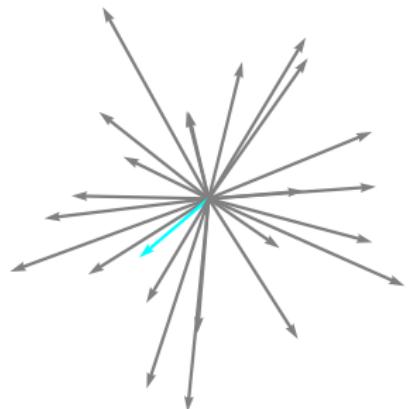
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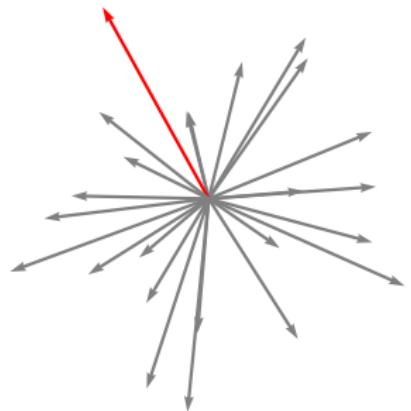
Select the first sentence,



$$\mathbb{P}[x_1 = \textcolor{cyan}{x}] = \frac{1}{2} \|\phi_{\textcolor{cyan}{x}}\|^2.$$

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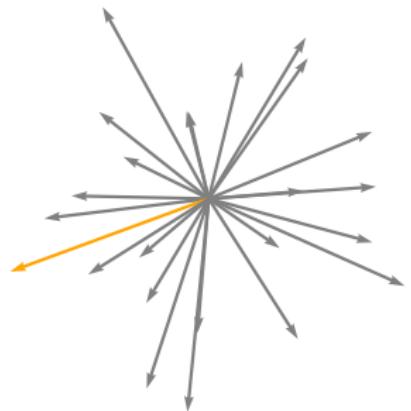
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$$\mathbb{P}[x_1 = \textcolor{red}{x}] = \frac{1}{2} \|\phi_{\textcolor{red}{x}}\|^2.$$

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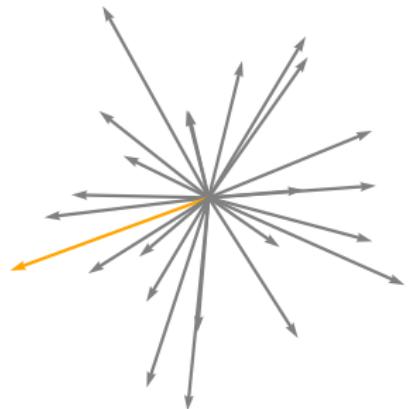
Select the first sentence,



$$\mathbb{P}[x_1 = \textcolor{orange}{x}] = \frac{1}{2} \|\phi_{\textcolor{orange}{x}}\|^2.$$

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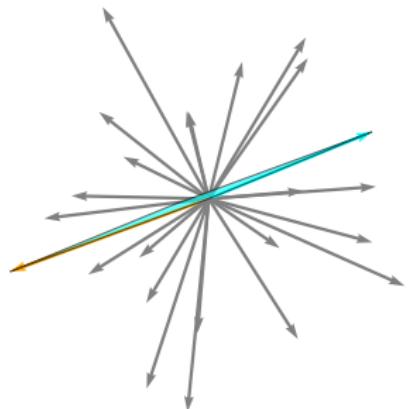
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$$\mathbb{P}[x_2 = x \mid \textcolor{orange}{x_1}] = \text{distance}^2(\phi_x, \text{span}\{\phi_{\textcolor{orange}{x_1}}\}).$$

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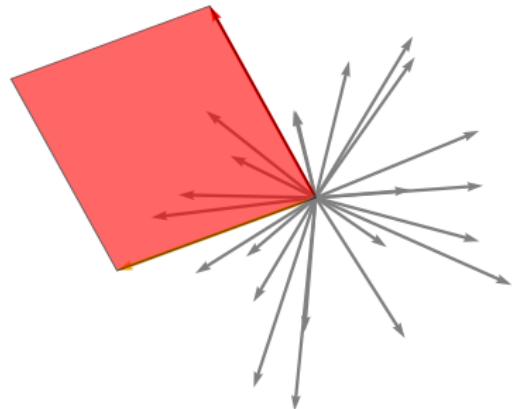
Select the second sentence,



$$\mathbb{P}[x_2 = \textcolor{teal}{x} \mid \textcolor{orange}{x_1}] = \text{distance}^2(\phi_x, \text{span}\{\phi_{x_1}\}).$$

Illustration of the chain rule ($M = 24, N = 2$)

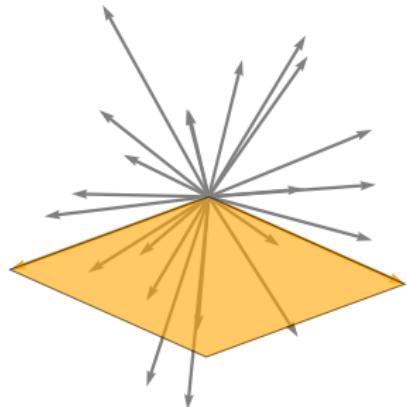
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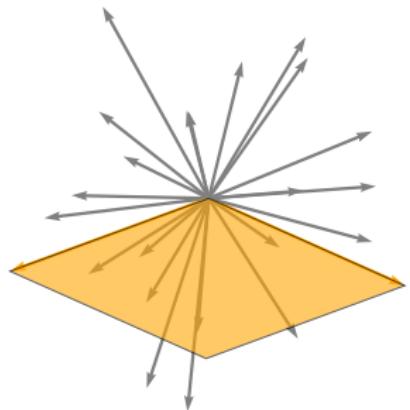


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Illustration of the chain rule ($M = 24, N = 2$)

Output summary.

But a dream within a dream?
Is all that we see or seem
One from the pitiless wave?
O God! can I not save
Them with a tighter clasp?
O God! can I not grasp
While I weep--while I weep!
Through my fingers to the deep,
How few! yet how they creep
Grains of the golden sand-
And I hold within my hand,
Of a surf-tormented shore,
I stand amid the roar
Is but a dream within a dream.
All that we see or seem
Is it therefore the less gone?
In a vision, or in none,
In a night, or in a day,
Yet if hope has flown away
That my days have been a dream;
You are not wrong, who deem
Thus much let me avow-
And, in parting from you now,
Take this kiss upon the brow!



$$\mathbb{P}[\mathcal{X} = \{\mathbf{x}_1, \mathbf{x}_2\}] = \text{volume}^2\{\phi_{\mathbf{x}_1}, \phi_{\mathbf{x}_2}\} = \det \mathbf{K}_{\{\mathbf{x}_1, \mathbf{x}_2\}}.$$

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The basis-exchange walk

Metropolis Hastings kernel

(Feder and Mihail, 1992; Anari, Gharan, and Rezaei, 2016; Li, Jegelka, and Sra, 2016; Hermon and Salez, 2019)

- ▶ Starting from $B_0 \in \mathcal{B}$.

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- ▶ Starting from $B_0 \in \mathcal{B}$.
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$$B \rightarrow \tilde{B} = (B \setminus \{s\}) \cup \{t\},$$

where $s \sim \text{Uniform}(B)$ and $t \sim \text{Uniform}(B^c)$.

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- ▶ Acceptance probability (lazy)

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Illustration of the basis-exchange walk ($M = 24, N = 7$)

$$B_0 = \{1, 3, 9, 12, 13, 18, 24\}$$

But a dream within a dream?

Is all that we see or seem

One from the pitiless wave?

O God! can I not save

Them with a tighter clasp?

O God! can I not grasp

While I weep--while I weep!

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How few! yet how they creep

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Illustration of the basis-exchange walk ($M = 24, N = 7$)

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Conceptual shift: sampling by solving randomized linear programs

Finite projection DPP
Approximate sampling
Linear programming



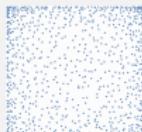
ICML, 2017

β -ensembles
Gibbs sampling
Random matrices

$$\text{eigvals} \begin{bmatrix} \text{red} & \text{blue} & 0 \\ \text{blue} & \text{red} & 0 \\ 0 & 0 & \text{red} \end{bmatrix}$$

Submitted, 2020

Monte Carlo integration
Exact sampling
Random linear system



NeurIPS, 2019

Python toolbox
DPPy
Reproducible research

```
from dppy import *
# [...]
dpp.sample()
```

JMLR-MLOSS, 2019

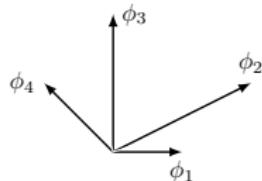
DPP sampling

Continuous embedding of the support

The support of finite projection DPPs, characterized by

$$\mathcal{B} \triangleq \{B ; |B| = N, \text{ and } \det \Phi_{:B} \neq 0\},$$

has the following geometrical representation.

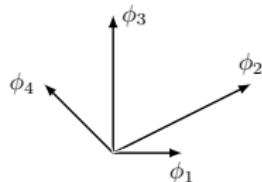
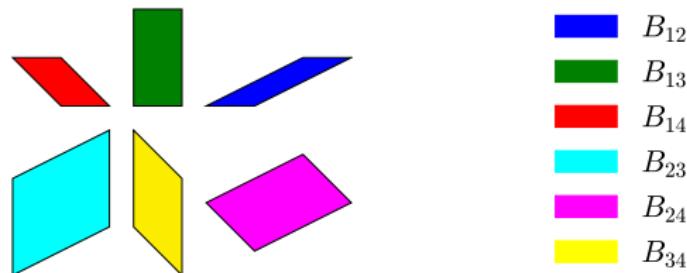


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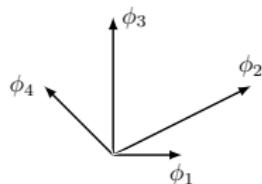
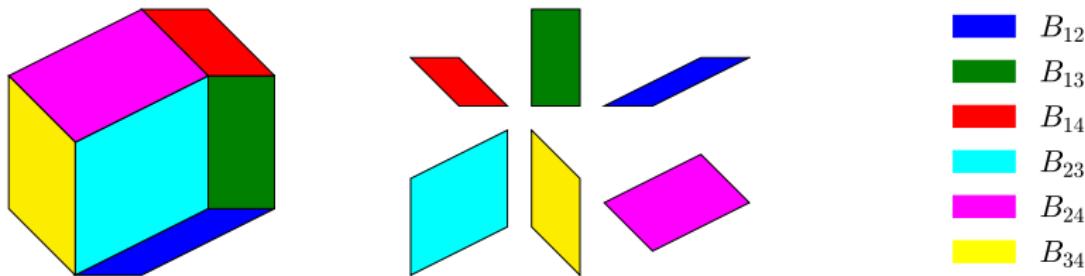


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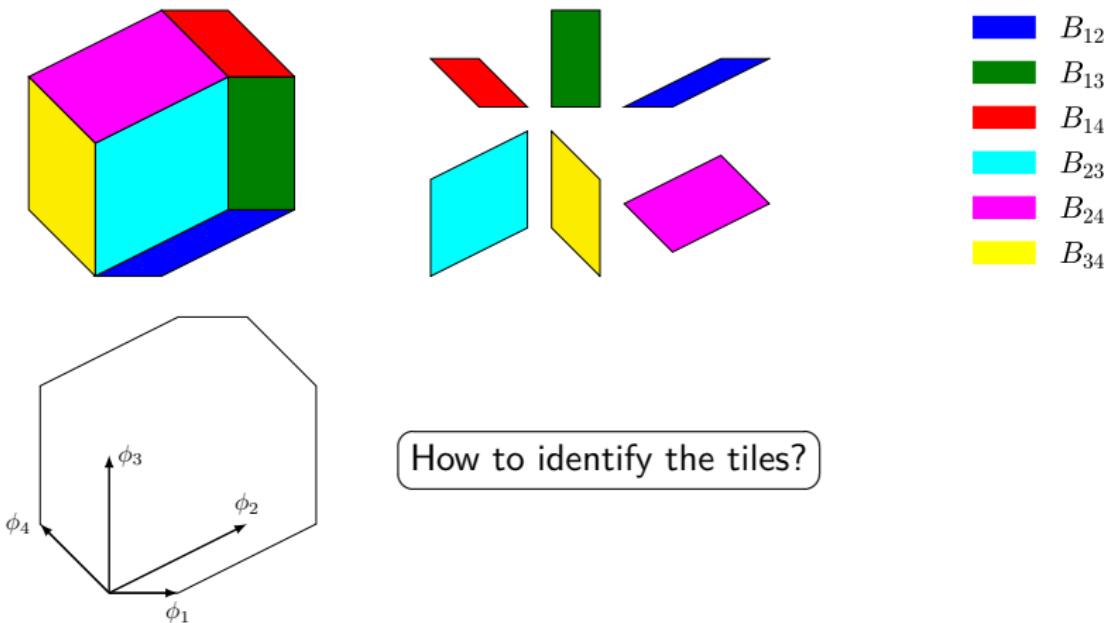


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Tiling of a zonotope (Dyer and Frieze, 1994)

Definition (Zonotope)

Let $\Phi \in \mathbb{R}^{N \times M}$ such that rank $\Phi = N$

$$\mathcal{Z}(\Phi) \triangleq \Phi[0, 1]^M.$$

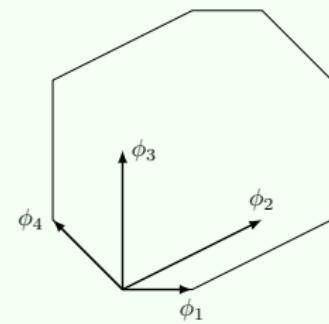
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Example ($M = 4, N = 2$)



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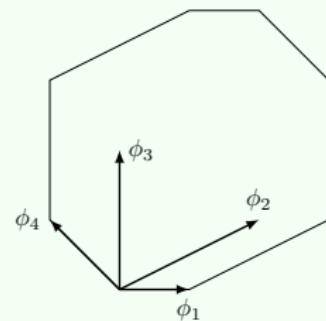
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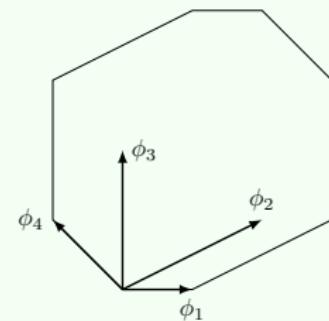
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- ▶ Let $x \in \mathcal{Z}(\Phi)$.
- ▶ Solve the linear program (LP)

$$\begin{array}{ll} \min_{y \in \mathbb{R}^M} & c^T y \\ \text{s.t.} & \Phi y = x \\ & 0 \leq y \leq 1 \end{array}$$

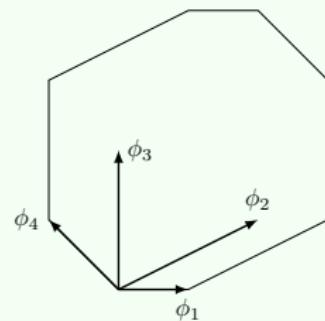
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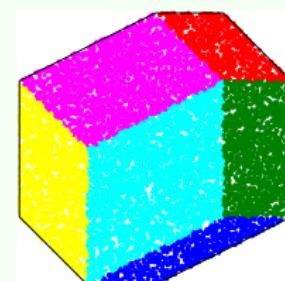
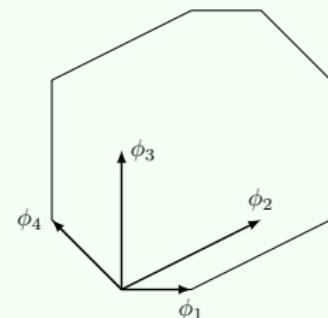
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B_{12}
B_{13}
B_{14}
B_{23}
B_{24}
B_{34}

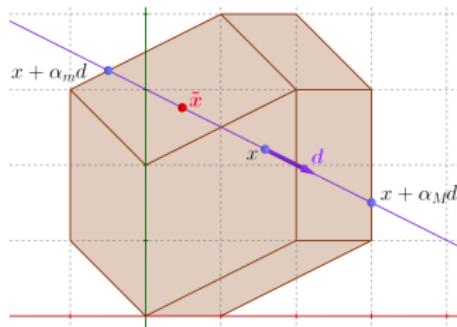
Random walk on zonotope $\xrightarrow{(LP)}$ random walk on tiles

Gautier, Bardenet, and Valko (2017)

Random walk on zonotope $\xrightarrow{(LP)}$ random walk on tiles

Gautier, Bardenet, and Valko (2017)

- ▶ Hit-and-run on $\mathcal{Z}(\Phi)$

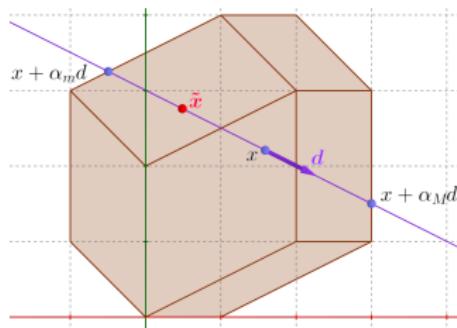


(Lovász and Vempala, 2003; Chen et al., 2018)

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- ▶ Random walk on \mathcal{B}

- ▶ Solve

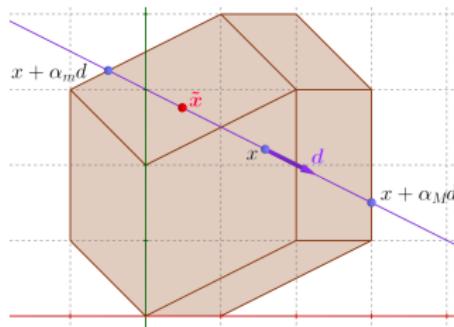
$$\begin{array}{ll} \min_{y \in \mathbb{R}^M} & c^T y \\ \text{s.t.} & \Phi y = x_t \\ & 0 \leq y \leq 1 \end{array}$$

- ▶ $B_{x_t} = \{i ; y_i^* \in]0, 1[\}$
- ▶ Markov Chain $(B_{x_t})_{t \in \mathbb{N}}$

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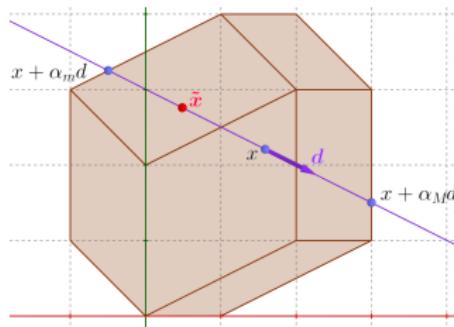
- ▶ Target density on $\mathcal{Z}(\Phi)$

$$\pi(x) = \sum_{B \in \mathcal{B}} C_B \times \mathbb{1}_{\mathcal{Z}(\Phi_B)}(x).$$

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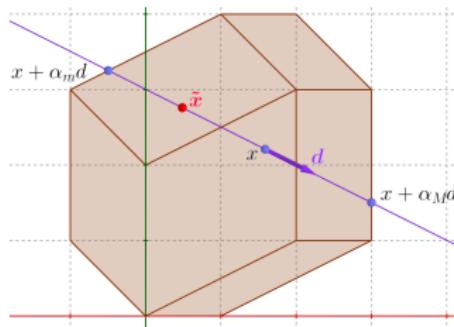
- ▶ Limiting distribution on \mathcal{B}

$$\mathbb{P}[B_x = B] = C_B \times |\det \Phi_{:B}|.$$

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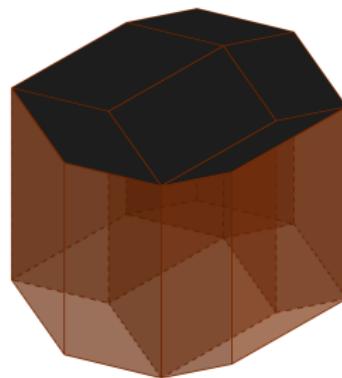
$$\mathbb{P}[B_x = B] = C_B \times |\det \Phi_{:B}|.$$

How to make $\mathbb{P}[B_x = B] \propto (\det \Phi_{:B})^2$?

Hit-and-run with acceptance ratio = 1

The target density on $\mathcal{Z}(\Phi)$ is uniform,

$$\pi(x) \propto \sum_{B \in \mathcal{B}} 1 \times \mathbb{1}_{\mathcal{Z}(\Phi:B)}(x).$$



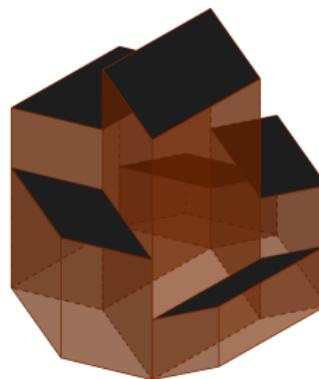
The limiting distribution on \mathcal{B} takes the form

$$\mathbb{P}[B_x = B] \propto 1 \times |\det \Phi_{:B}| = |\det \Phi_{:B}|^1.$$

$$\text{Hit-and-run with acceptance ratio} = \left| \frac{\det \Phi_{:\tilde{B}}}{\det \Phi_{:B}} \right|$$

The target density on $\mathcal{Z}(\Phi)$ is given by

$$\pi(x) \propto \sum_{B \in \mathcal{B}} |\det \Phi_{:B}| \times \mathbb{1}_{\mathcal{Z}(\Phi:B)}(x).$$



The limiting distribution on \mathcal{B} takes the form

$$\mathbb{P}[B_x = B] \propto |\det \Phi_{:B}| \times |\det \Phi_{:B}| = (\det \Phi_{:B})^2.$$

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All that we see or seem
Is it therefore the less gone?
In a vision, or in none,
In a night, or in a day,
Yet if hope has flown away
That my days have been a dream;
You are not wrong, who deem
Thus much let me avow--
And, in parting from you now,
Take this kiss upon the brow!*

Illustration of the zonotope walk ($M = 24, N = 7$)

$$B_0 = \{1, 3, 9, 12, 13, 18, 24\} \quad B_1 = (B_0 \setminus \{1, 9, 13, 18\}) \\ \cup \{6, 8, 14, 17\}$$

But a dream within a dream?
Is all that we see or seem
One from the pitiless wave?
O God! can I not save
Them with a tighter clasp?
O God! can I not grasp
While I weep--while I weep!
Through my fingers to the deep,
How few! yet how they creep
Grains of the golden sand-
And I hold within my hand
Of a surf-tormented shore,
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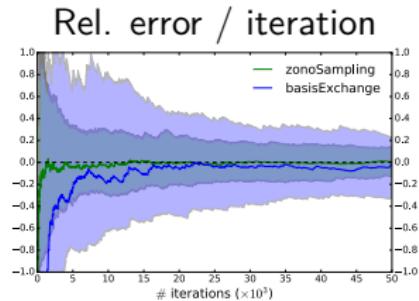
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Comparison of the zonotope and basis-exchange walks

Relative error of the estimation of $\mathbb{P}[\{x_1, x_2, x_3\} \subset \mathcal{X}] = \det \mathbf{K}_{\{x_1, x_2, x_3\}}.$

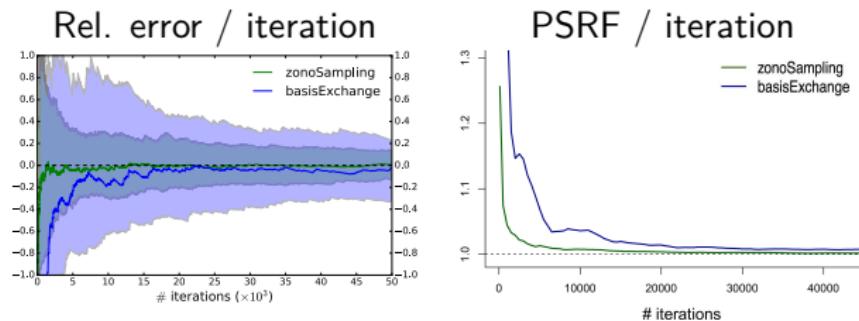
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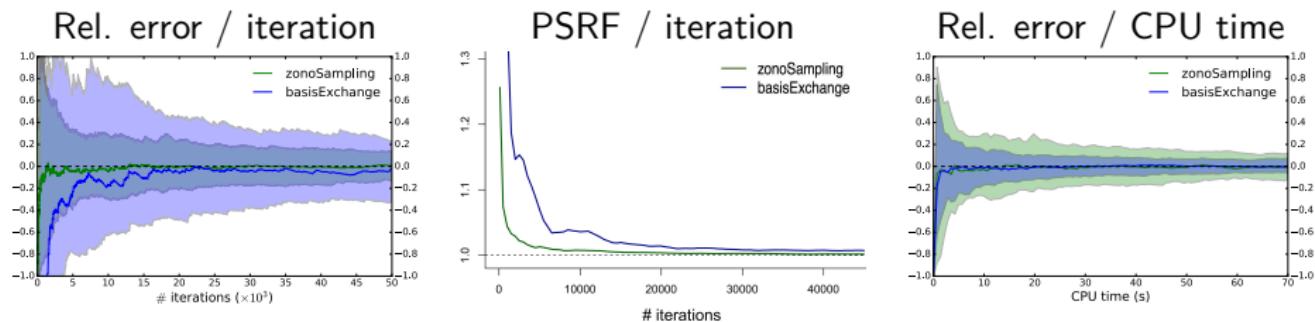
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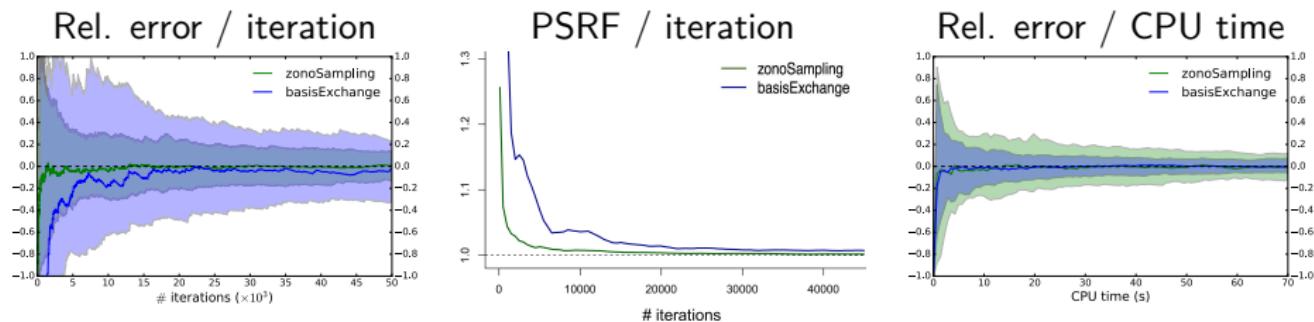
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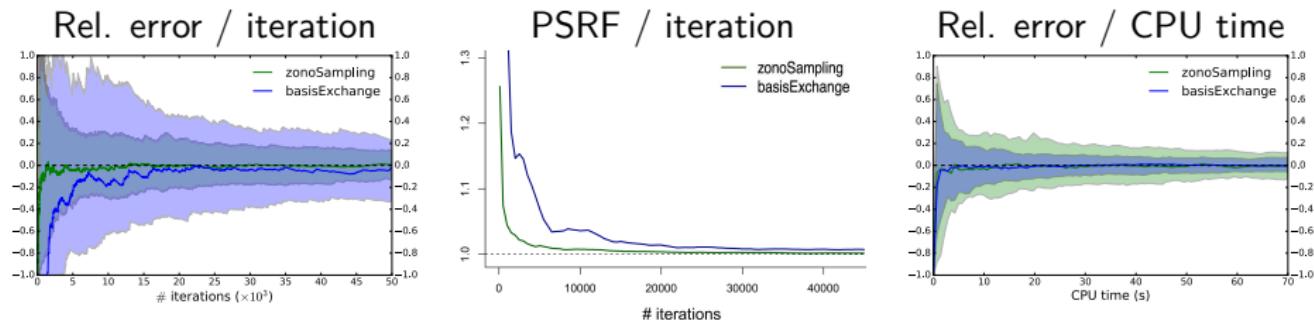
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	basis-exchange	zonotope
Exploration of the support	✓	(lazy)
Empirical mixing	✓✓	✓✓✓
Cost per iteration	✓✓✓	det
Theoretical guarantees	✓✓	poly(M, N)?

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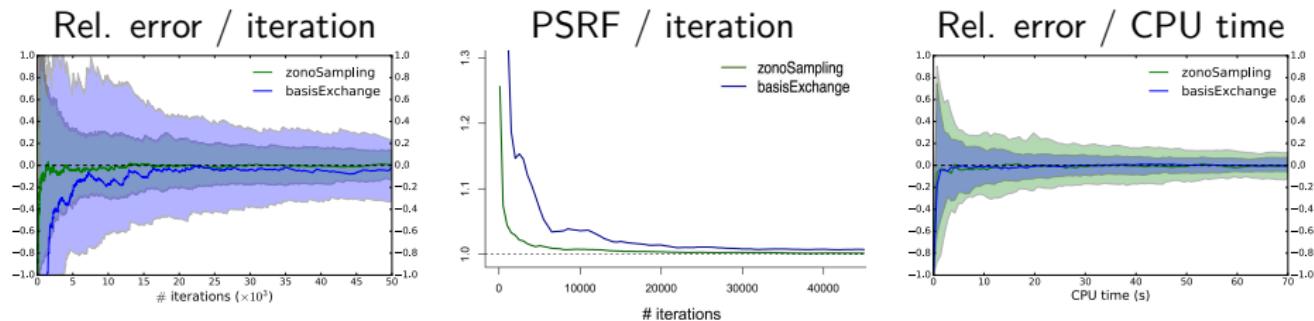


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The zonotope walk is sample efficient.

Comparison of the zonotope and basis-exchange walks

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The zonotope walk is sample efficient.

Can we generalize the idea to the continuous setting?

Overview

Introduction

DPP basics

Some insights on finite DPPs

Finite projection DPPs

Continuous projection DPPs

Exact sampling from finite projection DPPs

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Fast sampling from β -ensembles

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Open questions and perspectives

Sampling by solving randomized linear programs?

Finite case

Linear Programming (LP)

$$\min_y \quad c^T y$$

$$\text{s.t.} \quad \varphi_1^T y = x_1$$

$$\vdots$$

$$\varphi_N^T y = x_N$$

$$0 \leq y \leq 1$$

- ▶ Unique solution y^* 😊
- ▶ Efficient solvers 😊
- ▶ “Support” of the solution
 - ▶ $|i ; 0 < y_i^* < 1| = N$ 😊

Sampling by solving randomized linear programs?

Finite case

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Continuous case (dimension d)

Linear Semi Infinite Programming (LSIP)

$$\begin{aligned} \min_{\nu} \quad & \int c(x)\nu(dx) \\ \text{s.t.} \quad & \int \varphi_1(x)\nu(dx) = m_1 \\ & \vdots \\ & \int \varphi_N(x)\nu(dx) = m_N \\ & "0 \leq \mu \leq 1" \end{aligned}$$

- ▶ Unique solution y^* 😊
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- ▶ No unique solution 😞
- ▶ No efficient solvers 😞
- ▶ Structure of the support of solutions
 - ▶ $\exists \nu^* \text{ s.t. } |\text{supp } \nu^*| \leq N.$

(Goberna and López, 2014)

Sampling by solving randomized linear programs?

Dimension $d > 1$

For polynomials functions c and φ_n

$$\min_{\nu} \quad \int c(x)\nu(dx)$$

$$\text{s.t.} \quad \int \varphi_1(x)\nu(dx) = m_1$$

⋮

$$\int \varphi_N(x)\nu(dx) = m_N$$

- ▶ No unique solution ☹
- ▶ Efficient solvers? (Lasserre, 2010)
 - ▶ hierarchy of SDP relaxations ☺
 - ▶ works for small d and N ☹
- ▶ Structure of the support of solutions
 - ▶ $\exists \nu^*$ s.t. $|\text{supp } \nu^*| \leq N$
 - ▶ unstable support extraction ☹

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Dimension $d = 1$

Truncated moment problem

$$\begin{aligned} \min_{\nu} \quad & \mathbb{E}_{\nu}[X^{2N}] \\ \text{s.t.} \quad & \mathbb{E}_{\nu}[X] = m_1 \\ & \vdots \\ & \mathbb{E}_{\nu}[X^{2N-1}] = m_{2N-1} \end{aligned}$$

- ▶ No unique solution 😐
- ▶ Efficient solvers? (Lasserre, 2010)
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 - ▶ unstable support extraction 😐

- ▶ Unique solution $\nu^* = \sum_{n=1}^N \omega_n \delta_{x_n}$ 😊
- ▶ Unstable support extraction 😐

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How to randomize the moment constraints s.t. $\{x_1, \dots, x_N\} \sim \text{target DPP}$?

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$(\omega_n), (x_n)$ define a quadrature rule

(RyBo15; Dette and Studden, 1997)

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Sampling by computing the eigenvalues of random tridiagonal matrices

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Reparametrize ν^* via the 3-terms recurrence relation \perp polynomials encoded by

$$J_{\mathbf{a}, \mathbf{b}} \triangleq \begin{bmatrix} a_1 & \sqrt{b_1} & & (0) \\ \sqrt{b_1} & a_2 & & \ddots \\ \ddots & \ddots & \ddots & \sqrt{b_{N-1}} \\ (0) & \sqrt{b_{N-1}} & a_N & \end{bmatrix}$$

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- ▶ $\{x_1, \dots, x_N\} = \text{eigvals } J_{\mathbf{a}, \mathbf{b}}$
- ▶ Computational cost $\mathcal{O}(N^2)$

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Sampling 1D continuous projection DPPs may be cheaper than sampling finite projection DPPs!?

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Fast sampling from β -ensembles

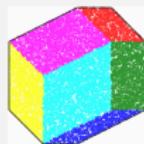
Conclusion

Summary of contributions

Open questions and perspectives

Convert a random matrix analysis tool to a computational tool

Finite projection DPP
Approximate sampling
Linear programming



ICML, 2017

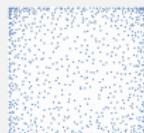
β -ensembles
Gibbs sampling
Random matrices

$$\text{eigvals} \begin{bmatrix} \text{red diagonal} & 0 \\ 0 & \text{blue diagonal} \end{bmatrix}$$

Submitted, 2020

DPP sampling

Monte Carlo integration
Exact sampling
Random linear system



NeurIPS, 2019

Python toolbox
DPPy 
Reproducible research

```
from dppy import *
# [...]
dpp.sample()
```

JMLR-MLOSS, 2019

Definition (β -ensemble)

Let (x_1, \dots, x_N) with distribution proportional to

$$\left| \prod_{i < j} (x_j - x_i) \right|^{\beta} \prod_{n=1}^N e^{-V(x_n)} dx_n,$$

then $\mathcal{X} = \{x_1, \dots, x_N\}$ is called a β -ensemble with potential V .

- ▶ Repulsion characterized by $\prod_{i < j} (x_j - x_i) = \det[x_j^{i-1}]_{i,j=1}^N$.
- ▶ Strength of the repulsion parametrized by $\beta > 0$ (inverse temperature).

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Example ($\beta = 2$, corresponds to a projection DPP)

- ▶ $\mu(dx) = e^{-V(x)} dx$.
- ▶ $K(x, y) = \sum_{k=0}^{N-1} p_k(x)p_k(y), \quad p_k, p_\ell \perp \text{polynomials w.r.t. } \mu.$

Classical β -ensembles and random matrix models

Name	Potential $V(x)$	Support
Hermite	$\frac{1}{2\sigma^2}(x - \mu)^2$	\mathbb{R}
Laguerre	$-(k-1)\log(x) + \frac{1}{\theta}x$	$]0, \infty[$
Jacobi	$-(a-1)\log(x) - (b-1)\log(1-x)$	$]0, 1[$

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$\beta (= 1, 2, 4)$ -ensembles as the eigenvalue distribution of random matrices.

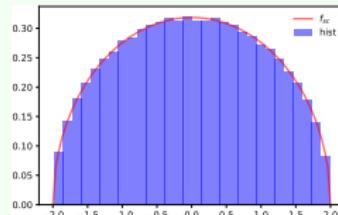
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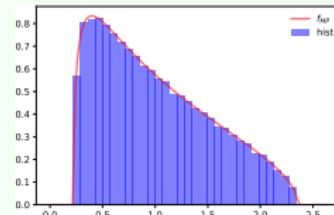
$\beta (= 1, 2, 4)$ -ensembles as the eigenvalue distribution of random matrices.

Example ($\beta = 2$ and $X \sim$ standard complex Gaussian matrix)

- $X \in \mathbb{C}^{N \times N}$
- $\text{eigvals}(X + X^H) \sim \text{Hermite}$



- $X \in \mathbb{C}^{N \times M}$
- $\text{eigvals}(XX^H) \sim \text{Laguerre}$



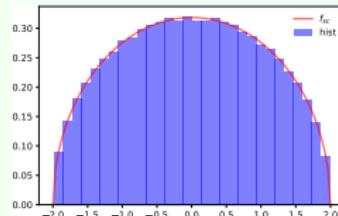
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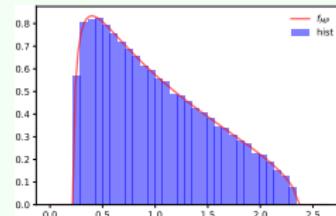
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Random matrix models grant $\mathcal{O}(N^3)$ exact samplers!

Classical β -ensembles and random tridiagonal models

Dumitriu and Edelman (2002) and Killip and Nenciu (2004) derived equivalent random tridiagonal models, valid for $\beta > 0$,

$$\text{eigvals} \begin{bmatrix} a_1 & \sqrt{b_1} & & (0) \\ \sqrt{b_1} & a_2 & \ddots & \\ & \ddots & \ddots & \sqrt{b_{N-1}} \\ (0) & & \sqrt{b_{N-1}} & a_N \end{bmatrix}.$$

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Consider independent $a_n \sim \mathcal{N}(\mu, \sigma^2)$, and $b_n \sim \Gamma\left(\frac{\beta}{2}(N-n), \sigma^2\right)$.

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Dumitriu and Edelman (2002) and Killip and Nenciu (2004) derived equivalent random tridiagonal models, valid for $\beta > 0$,

$$\text{eigvals} \begin{bmatrix} a_1 & \sqrt{b_1} & & (0) \\ \sqrt{b_1} & a_2 & \ddots & \\ & \ddots & \ddots & \sqrt{b_{N-1}} \\ (0) & & \sqrt{b_{N-1}} & a_N \end{bmatrix}.$$

Example (Hermite ensemble, $\beta > 0$, $V(x) = \frac{1}{2\sigma^2}(x - \mu)^2$)

Consider independent $a_n \sim \mathcal{N}(\mu, \sigma^2)$, and $b_n \sim \Gamma\left(\frac{\beta}{2}(N-n), \sigma^2\right)$.

Random tridiagonal models grant $\mathcal{O}(N^2)$ exact samplers!

Extend tridiagonal models to more general β -ensembles ?

How to randomize the entries of the tridiagonal matrix?

Proposition (Krishnapur et. al, 2016)

Consider the random tridiagonal matrix $J_{\mathbf{a}, \mathbf{b}}$ where the entries have joint density

$$\propto e^{-\text{Tr } V(J_{\mathbf{a}, \mathbf{b}})} \prod_{n=1}^{N-1} b_n^{\frac{\beta}{2}(N-n)-1}.$$

Then, the eigenvalues of $J_{\mathbf{a}, \mathbf{b}}$ have joint density

$$\propto \left| \prod_{i < j} (x_j - x_i) \right|^{\beta} \prod_{n=1}^N e^{-V(x_n)}.$$

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Gautier, Bardenet, and Valko (2020 - arXiv)

- ▶ Provide simple and clean proof.

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- ▶ Provide simple and clean proof, starting from

$$\mu^* = \sum_{n=1}^N \omega_n \delta_{x_n}.$$

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- ▶ Provide simple and clean proof, starting from $\mu^* = \sum_{n=1}^N \omega_n \delta_{x_n}$.
- ▶ Extend Krishnapur's result to unify the treatment of classical β -ensembles.
- ▶ Perform empirical study of tridiagonal models for polynomial potential V .

Tridiagonal models for polynomial potentials V

When degree $V = 2$,

- ▶ $(a_n), (b_n)$ are independent ☺
- ▶ have easy-to-sample distribution ☺

Example ($V(x) = \frac{1}{2\sigma^2}(x - \mu)^2$)

$$a_n \sim \mathcal{N}(\mu, \sigma^2), \quad b_n \sim \Gamma\left(\frac{\beta}{2}(N - n), \sigma^2\right).$$

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Example ($V(x) = g_4x^4 + g_2x^2$)

$a_n \mid \mathbf{a}_{\setminus n}, \mathbf{b}$

$$\sim \exp\left[-\left(g_4 a_n^4 + a_n^2 [g_2 + 4g_4(b_{n-1} + b_n)] + 4g_4 a_n (a_{n-1} b_{n-1} + a_{n+1} b_n)\right)\right],$$

$b_n \mid \mathbf{a}, \mathbf{b}_{\setminus n}$

$$\sim b_n^{\frac{\beta}{2}(N-n)-1} \exp\left[-2\left(g_4 b_n^2 + b_n [g_2 + 2g_4(a_n^2 + a_n a_{n+1} + a_{n+1}^2 + b_{n-1} + b_{n+1})]\right)\right].$$

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This suggests a Gibbs sampling strategy!

Combining tridiagonal models with Gibbs sampling

Target: β -ensembles with potentials of the form

$$V(x) = g_6x^6 + \cancel{g_5x^5} + g_4x^4 + g_3x^3 + g_2x^2 + g_1x.$$

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- ▶ Systematic scan Gibbs sampler

for $t = 1$ **to** T

for $n = 1$ **to** N

 sample $a_n | \mathbf{a}_{\setminus n}, \mathbf{b}$

 sample $b_n | \mathbf{a}, \mathbf{b}_{\setminus n}$ **if** $n < N$

$\{x_1^t, \dots, x_N^t\} = \text{eigvals } J_{\mathbf{a}, \mathbf{b}}$

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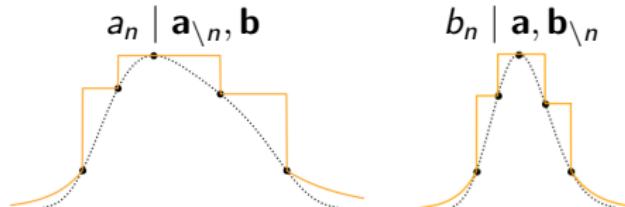
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- ▶ Exact sampling of log-concave conditionals (Devroye, 2012).

 ▶ e.g., $V(x) = \frac{1}{4}x^4$.



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- ▶ Exact sampling of log-concave conditionals (Devroye, 2012).
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- ▶ Metropolis-Hastings kernel (MALA) for **non** log-concave conditionals.
 - ▶ e.g., $V(x) = \frac{1}{6}x^6$.

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How does it perform?

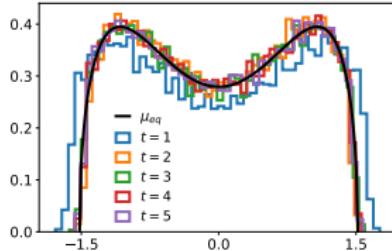
Monitoring of the empirical convergence

Convergence of the empirical marginal distribution to the equilibrium measure.

$$\hat{\mu}_N^t = \frac{1}{N} \sum_{n=1}^N \delta_{x_n^t} \xrightarrow[N, t \rightarrow \infty]{} \mu_{\text{eq}}.$$

- $V(x) = \frac{1}{4}x^4$, exact sampling of the conditionals.

$N = 20$

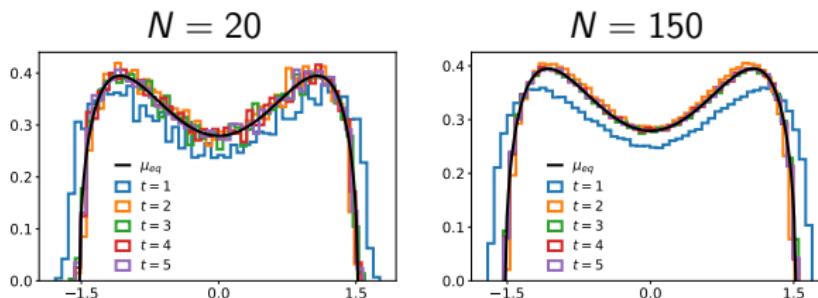


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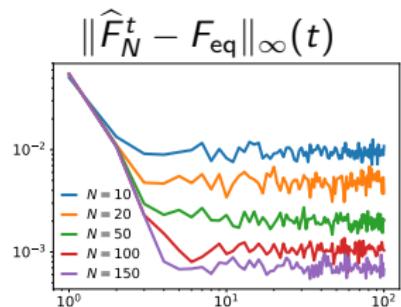
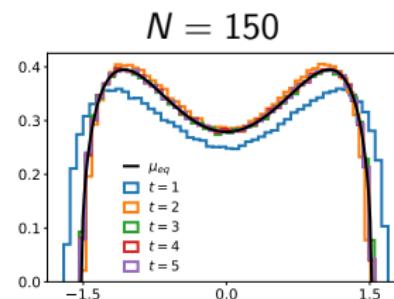
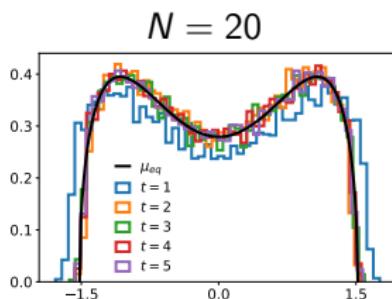


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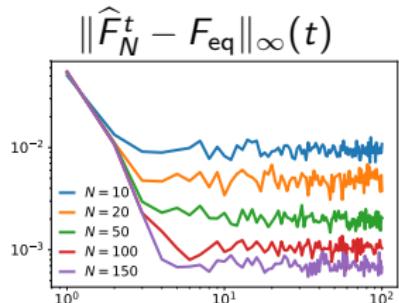
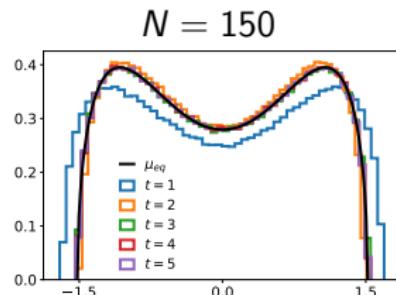
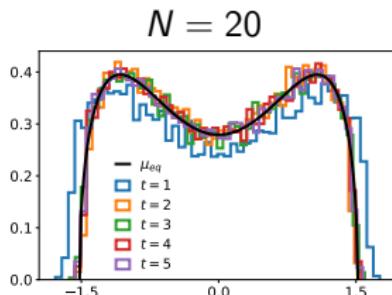


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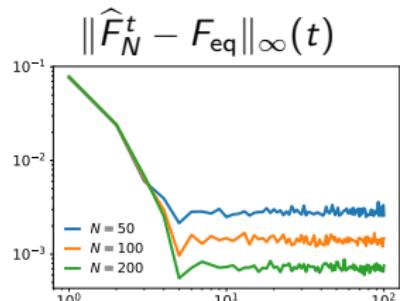
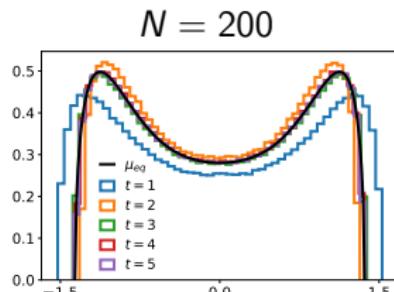
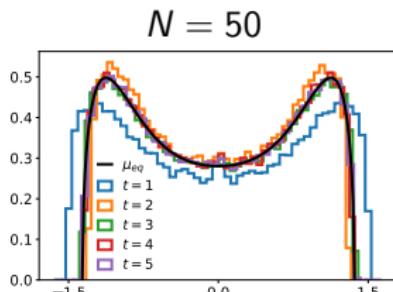
- Good adequation with the theory.
- Empirical convergence within $t \leq 10$ Gibbs passes, **only!**

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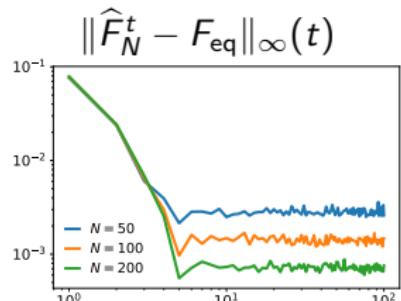
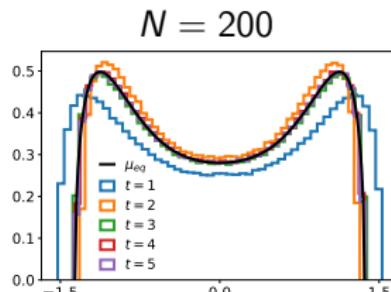
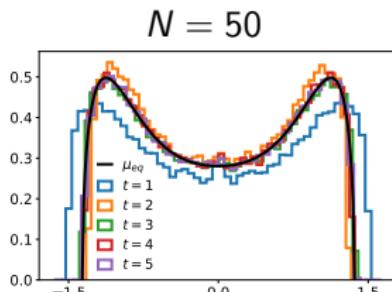
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- Good adequation with the theory.
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Supports the $\mathcal{O}(\log(N))$ mixing time conjecture of Krishnapur et. al (2016).

Overview

Introduction

DPP basics

Some insights on finite DPPs

Finite projection DPPs

Continuous projection DPPs

Exact sampling from finite projection DPPs

Approximate sampling from finite projection DPPs

Contributions

Zonotope sampling for finite projection DPPs

Transition

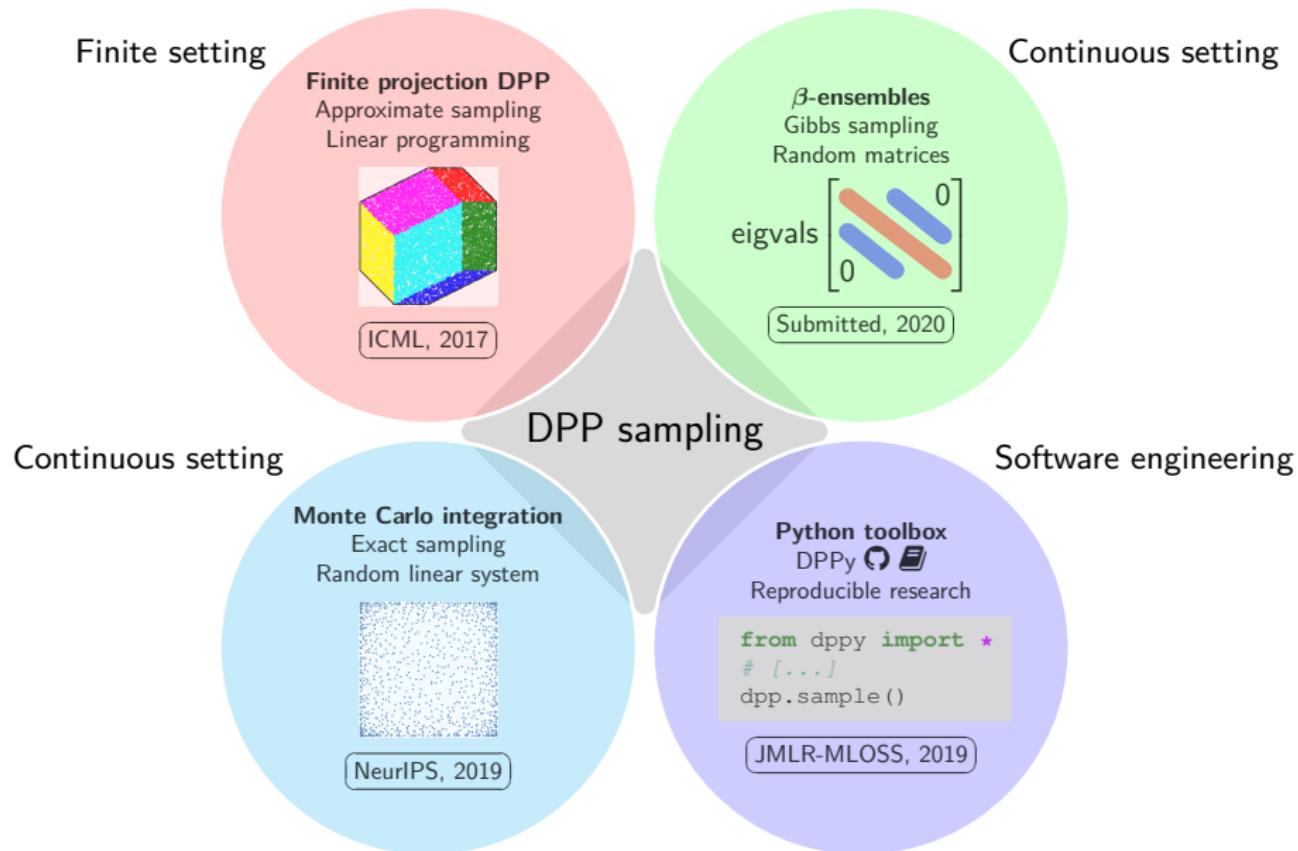
Fast sampling from β -ensembles

Conclusion

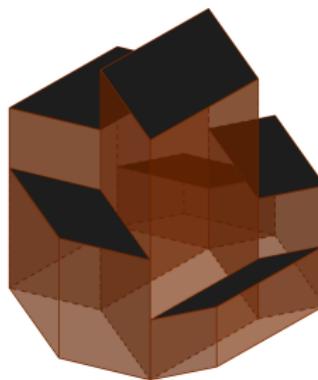
Summary of contributions

Open questions and perspectives

My Ph.D. in a nutshell



Zonotope sampling for finite projection DPPs



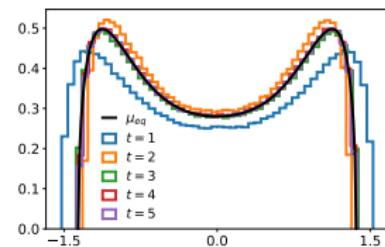
- ▶ New perspective on finite projection DPPs.
- ▶ Combination of geometry, Markov chains and linear programming.
- ▶ Approximate sampler involving randomized linear programs.
- ▶ More efficient exploration of the state space.

ICML, 2017

Tridiagonal models for sampling β -ensembles

eigvals

$$\begin{bmatrix} & & 0 \\ & \text{red cylinder} & \\ & \text{blue cylinder} & \\ 0 & & \end{bmatrix}$$



- ▶ Unified treatment of tridiagonal models for the classical β -ensembles.
- ▶ Combination of a Gibbs sampler with calculation of eigenvalues.
- ▶ Very fast empirical convergence supporting the $\mathcal{O}(\log(N))$ mixing time conjecture.

Submitted to an international journal, 2020

Monte Carlo integration with DPPs

Let $\{x_1, \dots, x_N\} \sim \text{DPP}(K, \mu)$, where $K(x, y) = \sum_{k=0}^{N-1} \phi_k(x)\phi_k(y)$.

$$\int f(x)\mu(dx) \approx \sum_{n=1}^N \omega_n f(x_n),$$

- ▶ Shed light on the estimator of Ermakov and Zolotukhin (1960)
 - ▶ involving a randomized linear system
 - ▶ provide new simple proofs of its properties

$$\text{Var} = \|f\|^2 - \sum_{k=0}^{N-1} \langle f, \phi_k \rangle^2.$$

- ▶ Numerical comparison with the estimator of Bardenet and Hardy (2020)
- ▶ Tailored implementation of the chain rule.

Adapt the kernel K to the basis where f has a smooth/sparse expansion.

DPPy: DPP sampling with Python

guilgautier / DPPy

Used by 8

Unwatch 12

Unstar 94

Fork 24

```
from dppy import *
# [...]
dpp.sample()
```

- ▶ Open source toolbox .
- ▶ Implementation of exact and approximate samplers.
- ▶ Extensive documentation .

JMLR-MLOSS, 2019

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Open questions

- ▶ Zonotope
 - ▶ prove a bound on the mixing time.
 - ▶ extend the LP idea for continuous DPPs.
- ▶ β -ensembles
 - ▶ prove the $\mathcal{O}(\log(N))$ mixing time for the Gibbs sampler.
 - ▶ extend tridiagonal models for multivariate β -ensembles.
- ▶ Efficient sampler for continuous projection DPPs ($d > 1$)?
- ▶ Avoid kernel eigendecomposition for sampling non-projection DPPs?

Perspectives

- ▶ Find a good reparametrization of DPPs where
 - ▶ complex interaction structure vanishes.
 - ▶ efficient sampling can be performed.
- ▶ Continuous extension of sampling by solving linear programs.
- ▶ Sampling by coupling the target DPP with another process.
 - ▶ Decreusefond, Flint, and Low (2013), Launay, Galerne, and Desolneux (2018), and Dereziński, Calandriello, and Valko (2019).
- ▶ Continue developing the DPPy toolbox  .

Thank you!
Ευχαριστώ!
Merci!

"My PhD story"

COMMENT GÉNÉRER DES RECOMMANDATIONS DIVERSES



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L -ensembles and k -DPPs

Definition (L -ensemble)

Let $\mathbf{L} \succeq 0$. The point process defined by

$$\mathbb{P}[\mathcal{X} = S] = \frac{\det \mathbf{L}_S}{\det(I + \mathbf{L})},$$

is called an L -ensemble. It is a DPP with kernel $\mathbf{K} = \mathbf{L}(I + \mathbf{L})^{-1}$.

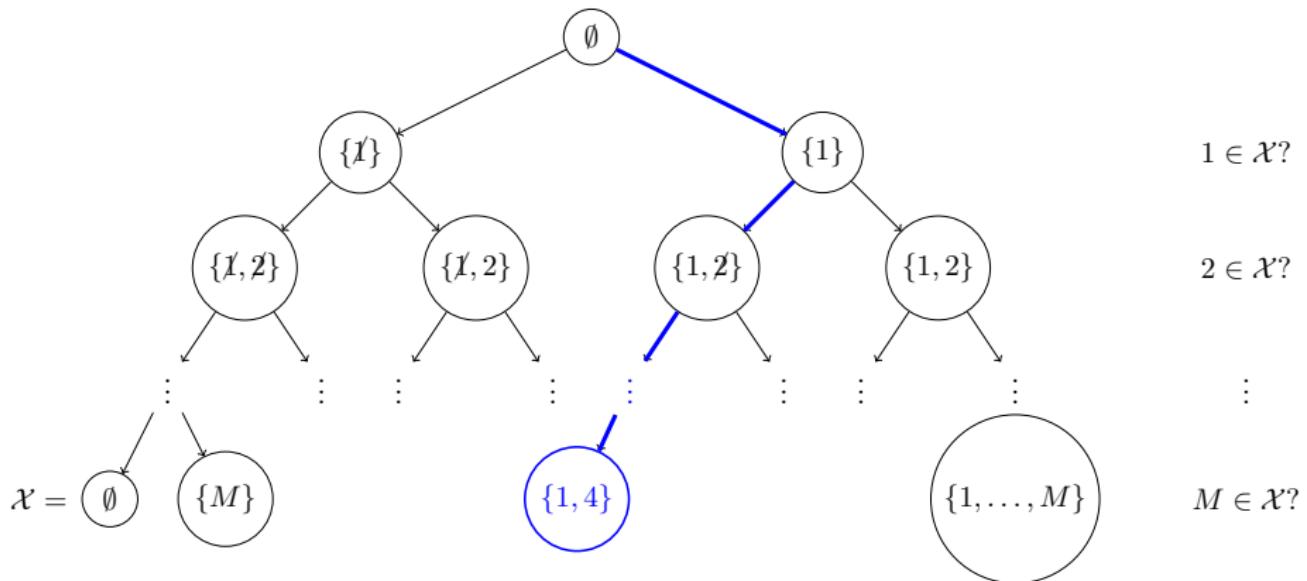
Definition (k -DPP)

Let $\mathbf{L} \succeq 0$ and $k \in \mathbb{N}^*$. The point process defined by

$$\mathbb{P}[\mathcal{X} = S] \propto \det \mathbf{L}_S \mathbb{1}_{|S|=k}.$$

is called a k -DPP.

Chain rule on sets



BH estimator and the multivariate Jacobi ensemble

Natural unbiased estimator of $\int_{\mathbb{X}} f(x) \mu(dx)$

$$\widehat{I}_N^{\text{BH}}(f) = \sum_{n=1}^N \frac{f(x_n)}{K(x_n, x_n)}$$

- ▶ Bardenet and Hardy (2020) show fast CLT, for f essentially C^1

$$\sqrt{N^{1+1/d}} \left(\widehat{I}_N^{\text{BH}}(f) - \int_{[-1,1]^d} f(x) \omega(x) dx \right) \xrightarrow[N \rightarrow \infty]{\text{law}} \mathcal{N}(0, \Omega_{f,\omega}^2),$$

$$\text{with } \Omega_{f,\omega}^2 \triangleq \frac{1}{2} \sum_{k \in \mathbb{N}^d} (k_1 + \cdots + k_d) \mathcal{F} \left[\frac{f \omega}{\omega_{\text{eq}}} \right] (k)^2$$

Theorem (Ermakov and Zolotukhin, 1960)

$$f = \sum_{\ell=0}^{\textcolor{blue}{M}-1} \langle f, \phi_\ell \rangle \phi_\ell, \quad \textcolor{blue}{M} \in \mathbb{N} \cup \{\infty\}$$

1. Sample $\{x_1, \dots, x_{\textcolor{red}{N}}\} \sim \text{DPP}(\mu, K)$ with $K(x, y) = \sum_{k=0}^{\textcolor{red}{N}-1} \phi_k(x)\phi_k(y)$
2. Random linear system

$$\begin{bmatrix} \phi_0(x_1) & \dots & \phi_{\textcolor{red}{N}-1}(x_1) \\ \vdots & & \vdots \\ \phi_0(x_{\textcolor{red}{N}}) & \dots & \phi_{\textcolor{red}{N}-1}(x_{\textcolor{red}{N}}) \end{bmatrix} \begin{bmatrix} y_0 \\ \vdots \\ y_{N-1} \end{bmatrix} = \begin{bmatrix} f(x_1) \\ \vdots \\ f(x_{\textcolor{red}{N}}) \end{bmatrix}$$

- ▶ $\mathbb{E}[y_k] = \langle f, \phi_k \rangle = \int f(x)\phi_k(x)\mu(dx)$
- ▶ $\text{Var}[y_k] = \|f\|^2 - \sum_{\ell=0}^{\textcolor{red}{N}-1} \langle f, \phi_\ell \rangle^2 = \sum_{\ell=\textcolor{red}{N}}^{\textcolor{blue}{M}-1} \langle f, \phi_\ell \rangle^2 = 0 \quad \text{if } \textcolor{blue}{M} \leq \textcolor{red}{N}$
- ▶ $\text{Cov}[y_j, y_k] = 0, j \neq k$

Ermakov and Zolotukhin (1960) estimator

For constant ϕ_0 , e.g., multivariate Jacobi ensemble,

$$\mathbb{E}[y_0] = \phi_0 \int_{\mathbb{X}} f(x) \mu(dx)$$

A direct application of EZ theorem yields

$$\hat{I}_N^{\text{EZ}}(f) \triangleq \frac{y_0}{\phi_0} = \sqrt{\mu([-1, 1]^d)} \frac{\det \Phi_{\phi_0, f}(x_{1:N})}{\det \Phi(x_{1:N})}$$

as an unbiased estimator of $\int f(x) \mu(dx)$

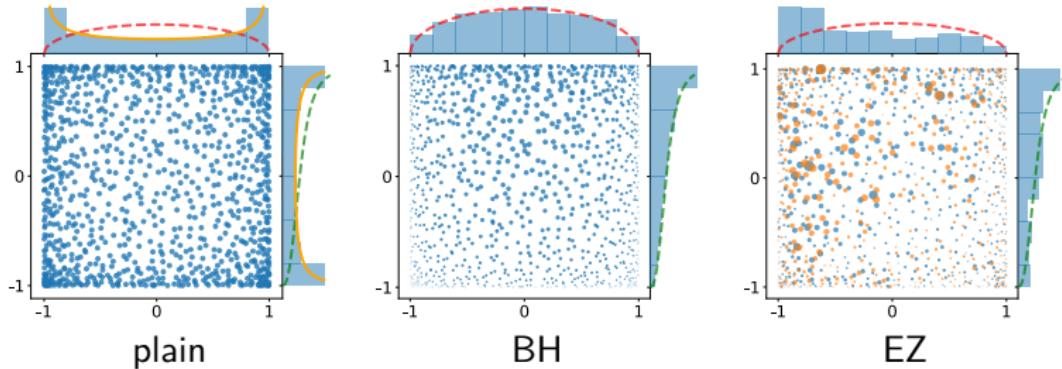
Using $\|\phi_0\| = 1$ and Cramer's rule

$$\Phi_{\phi_0, f} = \begin{bmatrix} f(x_1) & \dots & \psi_{N-1}(x_1) \\ \vdots & & \vdots \\ f(x_N) & \dots & \psi_{N-1}(x_N) \end{bmatrix} \quad \Phi = \begin{bmatrix} \phi_0(x_1) & \dots & \phi_{N-1}(x_1) \\ \vdots & & \vdots \\ \phi_0(x_N) & \dots & \phi_{N-1}(x_N) \end{bmatrix}$$

Comparison weights ω_n BH-EZ

$$\int_{\mathbb{X}} f(x) \mu(dx) \approx \hat{I}_N = \sum_{n=1}^N \omega_n(x_1, \dots, x_N) f(x_n)$$

► weights ω_n



► Non-asymptotic variance

$$\mathbb{V}\text{ar}[\hat{I}_N^{\text{BH}}] = \frac{1}{2} \int_{\mathbb{X}^2} \left(\frac{f(x)}{K(x,x)} - \frac{f(y)}{K(y,y)} \right)^2 K(x,y)^2 \mu(dx)\mu(dy)$$

$$\mathbb{V}\text{ar}[\hat{I}_N^{\text{EZ}}] = \|f\|^2 - \sum_{\ell=0}^{N-1} \langle f, \phi_\ell \rangle^2$$

Timings

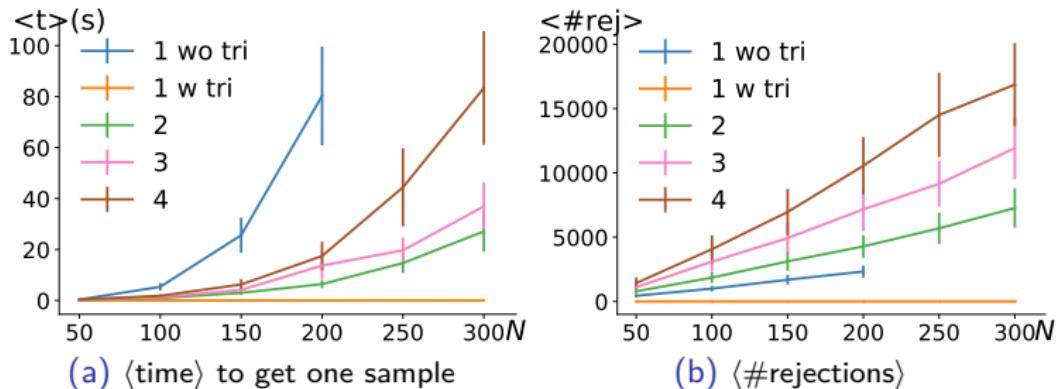


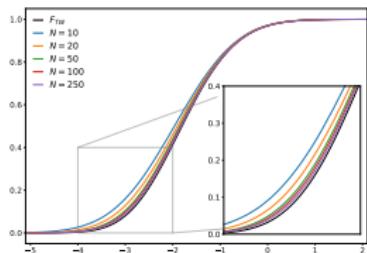
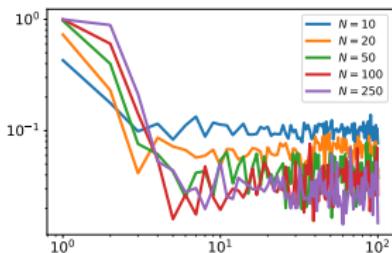
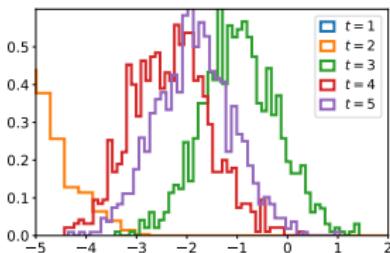
Figure 1: The colors and numbers correspond to the dimension. $a_i, b_i = -1/2$. For $d = 1$, the tridiagonal model (tri) of Killip and Nenciu (BH, 2004) offers tremendous savings, without it is cheaper to get a sample in larger dimension. The number of rejections grows as $N \log(N)2^d$.

Monitoring of the empirical convergence (λ_{\max} , $\beta = 2$)

Convergence of the distribution of the largest eigenvalue to Tracy-Widom.

$$\text{rescaled } \lambda_{\max}^t \xrightarrow[N, t \rightarrow \infty]{\text{law}} \text{TW}_2 .$$

- $V(x) = \frac{1}{4}x^4$, #independant runs = 10^3 .



- $V(x) = \frac{1}{6}x^6$, #independant runs = 10^3 .

